

Multiple Linear Regression

Simple linear regression: One predictor — X

Multiple linear regression: Several predictors — X_1, \dots, X_k

Linear (regression) model:

$E(Y|X_1 = x_1, \dots, X_k = x_k) = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k$ — models mean response as a function of predictors

Examples:

linear model?

- $E(Y|x) = \beta_0 + \beta_1 x$ — *Y.M.*

- $E(Y|\mathbf{x}) = \beta_0 + \beta_1 x + \beta_2 x^2$ — *Y.M.*

- $E(Y|\mathbf{x}) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 (x_1 * x_2)$ — *Y.M.*

- $E(\log(Y)|x) = \beta_0 + \beta_1 \log(x)$ — *Y.M.*

- $E(Y|x) = \beta_0 + (\beta_1 x)^{-1} = \beta_0 + \frac{1}{\beta_1} \cdot \frac{1}{x}$ — *NO, but can be made a linear model by taking $\delta_1 = \frac{1}{\beta_1}$*

- $E[Y|x] = \frac{1}{\beta_0 + \beta_1 x}$ — *No*

Note: "Linear" refers to linear in regression coefficients

Linear model: $E(Y|\mathbf{x}) = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k$

Y — response,
 x_1, \dots, x_k — predictors

Interpretation of $k + 1$ regression coefficients:

- $\beta_0 = E(Y|\mathbf{x} = \mathbf{0})$ — intercept
- $\beta_j = E(Y|x_1, \dots, \underline{x_j + 1}, \dots, x_k) - E(Y|x_1, \dots, \underline{x_j}, \dots, x_k)$
 — slope of x_j , i.e., change in mean response when j th predictor increases by 1, while keeping other predictors fixed, $j = 1, \dots, k$. $\rightarrow \beta_j$ represents the effect of the j th predictor after adjusting for the effects of the other predictors.

Data: n independent subjects, i th subject gives $(Y_i, X_{1i}, X_{2i}, \dots, X_{ki})$, $i = 1, \dots, n$.

Linear model for data: For $i = 1, \dots, n$,
 $E(\underline{Y_i} | x_{i1}, \dots, x_{ik}) = \beta_0 + \beta_1 x_{i1} + \dots + \beta_k x_{ik}$

Subject #	Y	X_1	...	X_k
1	Y_1	X_{11}	...	X_{1k}
2	Y_2	X_{21}	...	X_{2k}
\vdots	\vdots	\vdots	\vdots	\vdots
i	Y_i	X_{i1}	...	X_{ik}
\vdots	\vdots	\vdots	\vdots	\vdots
n	Y_n	X_{n1}	...	X_{nk}

$- E[Y_i | x_i]$

Alternative form: $Y_i = \underbrace{\beta_0 + \beta_1 x_{i1} + \dots + \beta_k x_{ik}}_{E[Y_i | x_i]} + \epsilon_i$

Assumptions:

- $E(\epsilon_i) = 0$, $\text{var}(\epsilon_i) = \sigma^2$, and ϵ_i are independent. — as before.
- $k + 1 < n$ — i.e., have more observations than the number of regression coefficients
- The predictors are considered fixed and are measured without error

These imply:

- $E(Y_i | x_{i1}, \dots, x_{ik}) = \beta_0 + \beta_1 x_{i1} + \dots + \beta_k x_{ik}$
 - $\text{var}(Y_i) = \sigma^2$
 - Y_1, \dots, Y_n are independent.
- } — as before.

Linear Model in Matrix Notation

Define:

$$Y = \begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix}_{n \times 1}, X = \begin{bmatrix} 1 & x_{11} & \dots & x_{1k} \\ 1 & x_{21} & \dots & x_{2k} \\ \vdots & \vdots & \vdots & \vdots \\ 1 & x_{n1} & \dots & x_{nk} \end{bmatrix}_{n \times (k+1)}, \beta = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_k \end{bmatrix}_{(k+1) \times 1}, \epsilon = \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{bmatrix}_{n \times 1}$$

Handwritten notes:
 - Y : response vector
 - X : design matrix w/ intercept
 - β : regression coeff. vector
 - ϵ : error vector

- $Y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_k x_{ik} + \epsilon_i = (\text{i-th row of } X) \times \beta + \epsilon_i$

- $Y = X\beta + \epsilon$

- $E(Y|X) = X\beta + E[\epsilon] = X\beta, \quad E[\epsilon] = \begin{bmatrix} E[\epsilon_1] \\ \vdots \\ E[\epsilon_n] \end{bmatrix}$

- rank of X is full, i.e., $(X'X)^{-1}$ exists.

- $\hat{\beta}$ = estimator of β

- $\hat{Y} = X\hat{\beta}$ = fitted (or predicted) response

Predicted response when $x = \underline{x_0}$: $\hat{Y}_0 = x_0' \hat{\beta}$

$$x_0 = \begin{bmatrix} x_{01} \\ \vdots \\ x_{0k} \end{bmatrix}_{(k+1) \times 1}$$

Estimated mean response from the model at the given predictor values

Least Squares Estimation of β

$\beta_0, \beta_1, \dots, \beta_k$

As before: Minimize $\sum_{i=1}^n \epsilon_i^2$ with respect to $\beta_0, \beta_1, \beta_k$ to get $\hat{\beta}$

- Least squares estimator: $\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$ — solution of the system of equations
- Minimum value of $\sum_{i=1}^n \epsilon_i^2$ is
 $\sum_{i=1}^n (Y_i - \hat{Y}_i)^2 = (\mathbf{Y} - \hat{\mathbf{Y}})'(\mathbf{Y} - \hat{\mathbf{Y}}) = SS_{\text{ERR}}$ — error (or residual) sum of squares

$$\frac{\partial \sum_{i=1}^n \epsilon_i^2}{\partial \beta_j} = 0, \quad j=1, \dots, k$$

Properties of $\hat{\beta}$:

- Linear in \mathbf{Y}

- Unbiased, i.e., $E[\hat{\beta}] = E[(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}] = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'E(\mathbf{Y}) = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'(\mathbf{X}\beta) = \beta$

- $\text{var}(\hat{\beta}) = \sigma^2(\mathbf{X}'\mathbf{X})^{-1}$

- $\text{var}(\hat{\beta}_0) = \sigma^2 \times \text{first diagonal element of } (\mathbf{X}'\mathbf{X})^{-1}$

- $\text{var}(\hat{\beta}_j) = \sigma^2 \times (j+1)\text{th diagonal element of } (\mathbf{X}'\mathbf{X})^{-1}$

- $\hat{\sigma}^2 = SS_{\text{ERR}}/(n - k - 1) = MS_{\text{ERR}}$ is unbiased for σ^2 .

unknown parameters = $(k+1) \text{ reg coeff} + 1 = (k+2)$.

ANOVA table

As before:

- $SS_{TOT} = \sum_{i=1}^n (Y_i - \bar{Y})^2 = (\mathbf{Y} - \bar{\mathbf{Y}})'(\mathbf{Y} - \bar{\mathbf{Y}})$, where

$$\bar{\mathbf{Y}} = \bar{Y} \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix}$$

ANOVA identity:
 $SS_{REG} + SS_{ERR} = SS_{TOT}$

- $SS_{REG} = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 = (\hat{\mathbf{Y}} - \bar{\mathbf{Y}})'(\hat{\mathbf{Y}} - \bar{\mathbf{Y}})$

- $SS_{ERR} = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2 = (\mathbf{Y} - \hat{\mathbf{Y}})'(\mathbf{Y} - \hat{\mathbf{Y}})$

Source	SS	d.f.	MS	F
Model	SS_{REG}	k	$MS_{REG} = \frac{SS_{REG}}{k}$	$\frac{MS_{REG}}{MS_{ERR}}$
Error	SS_{ERR}	$n - k - 1$	$MS_{ERR} = \frac{SS_{ERR}}{n - k - 1}$	
Total	SS_{TOT}	$n - 1$		

$n - \# \text{ reg. coeff.}$