

Isotonic subgroup selection

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Collaborators



Henry W. J. Reeve
University of Bristol



Timothy I. Cannings
University of Edinburgh



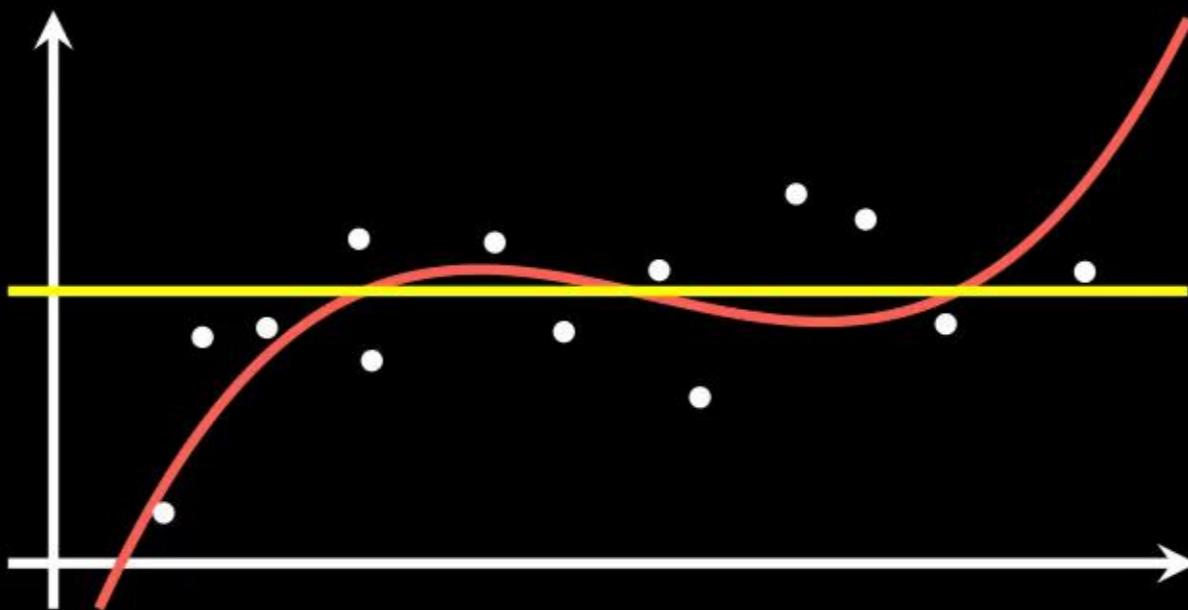
Richard J. Samworth
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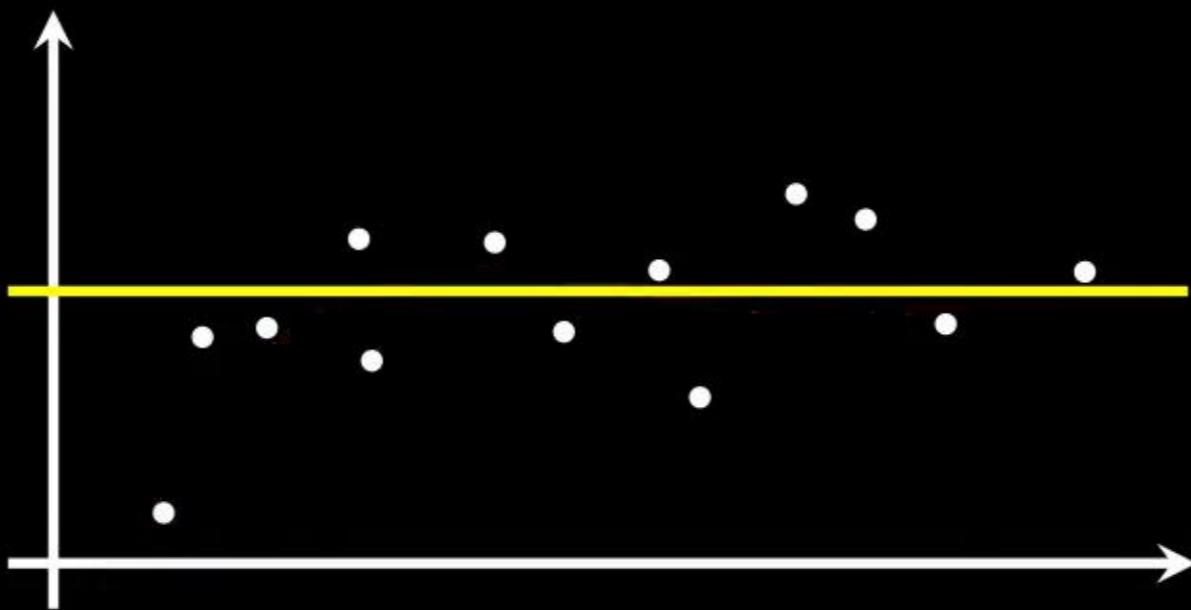
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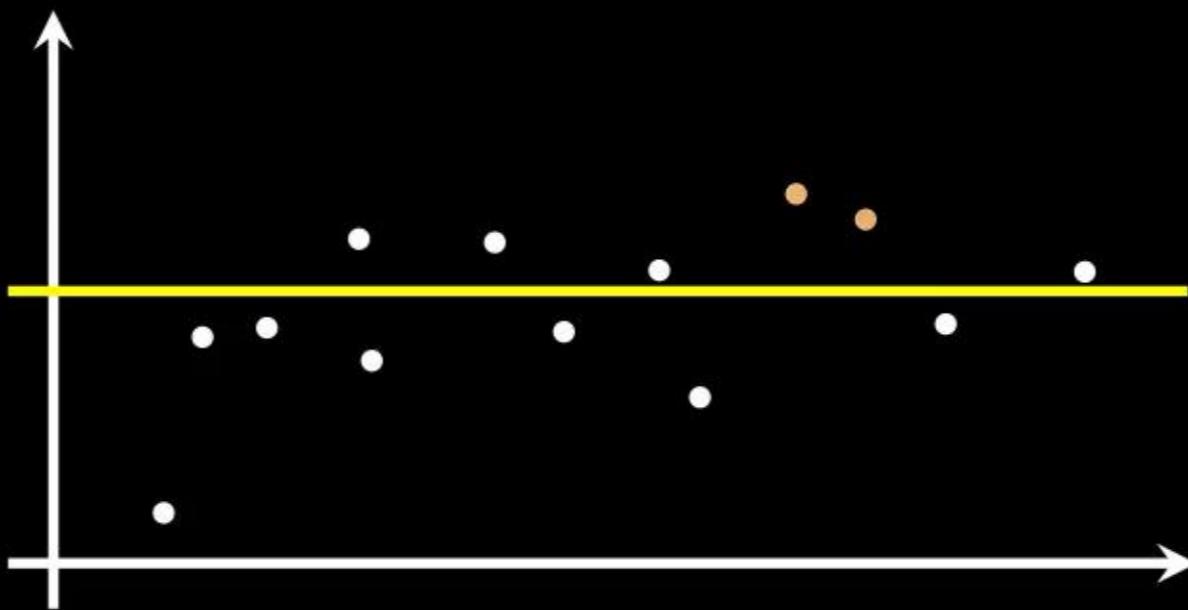
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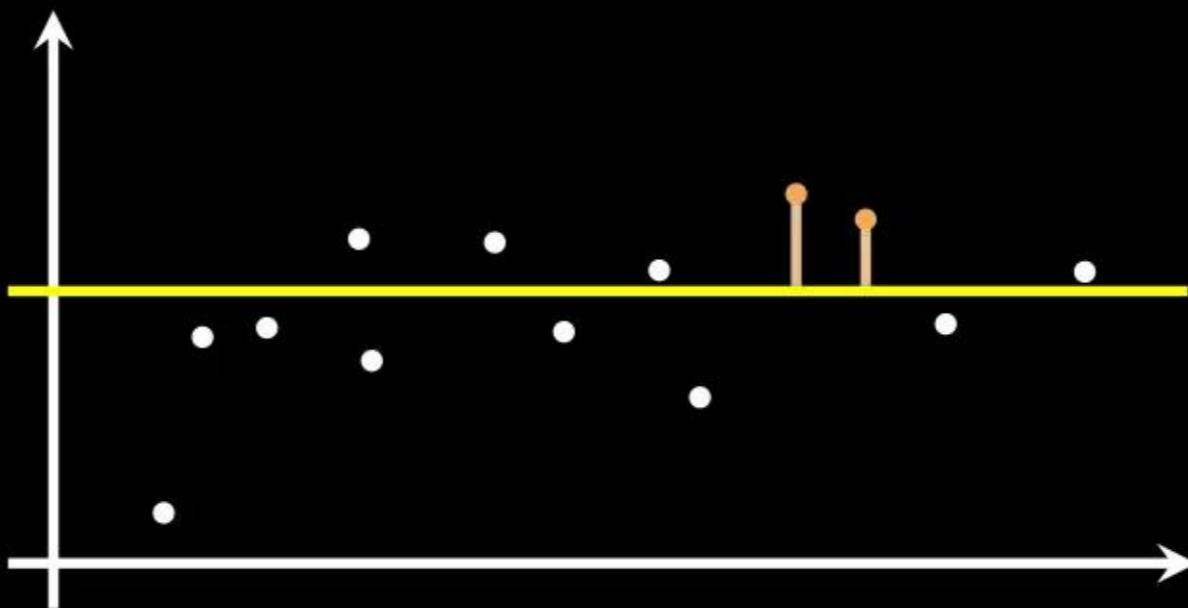
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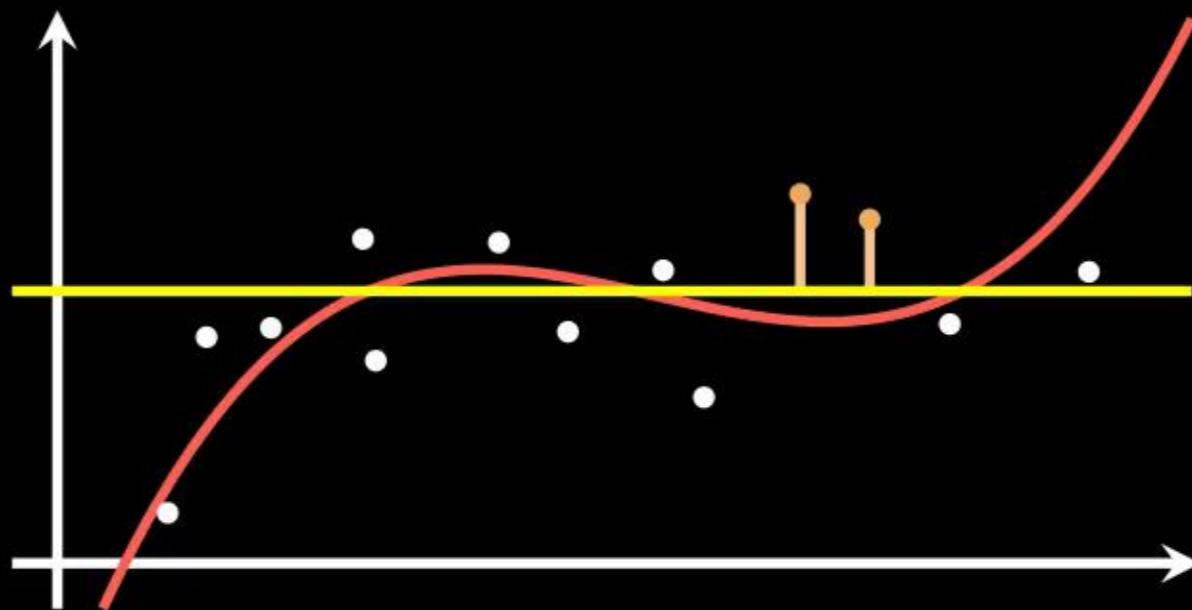
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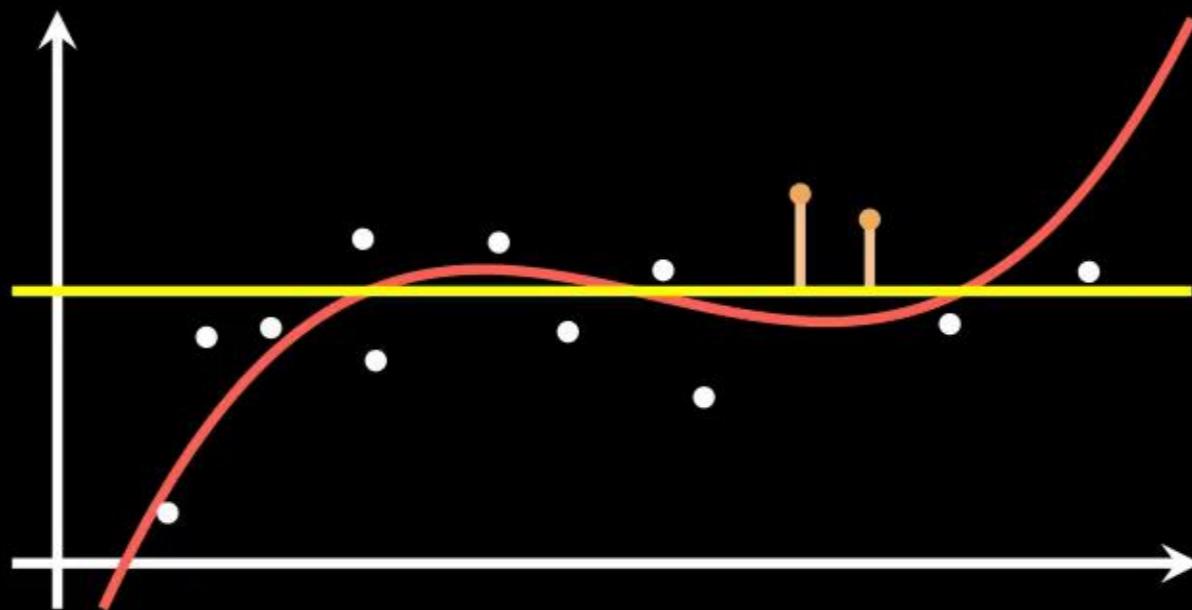
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Example. Efficacy of a new vaccine.

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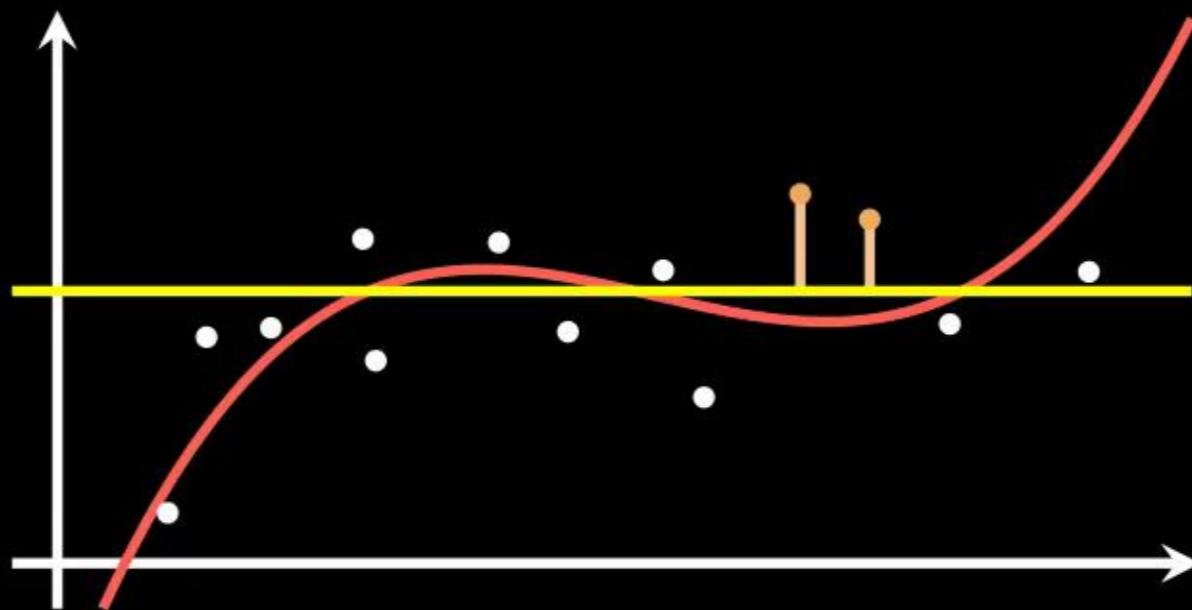
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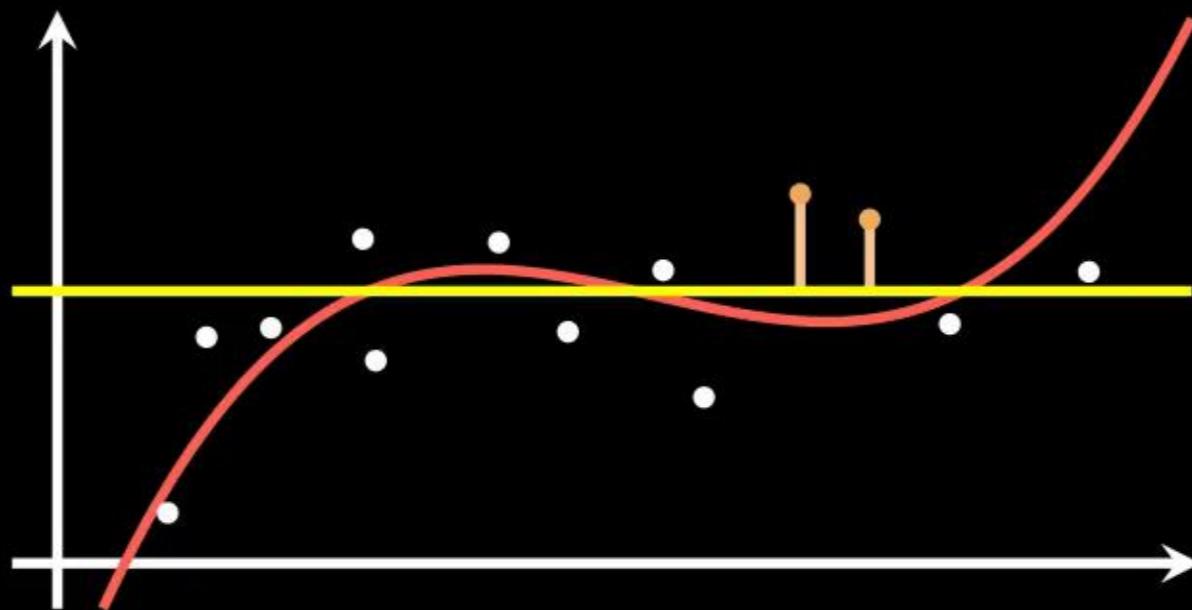


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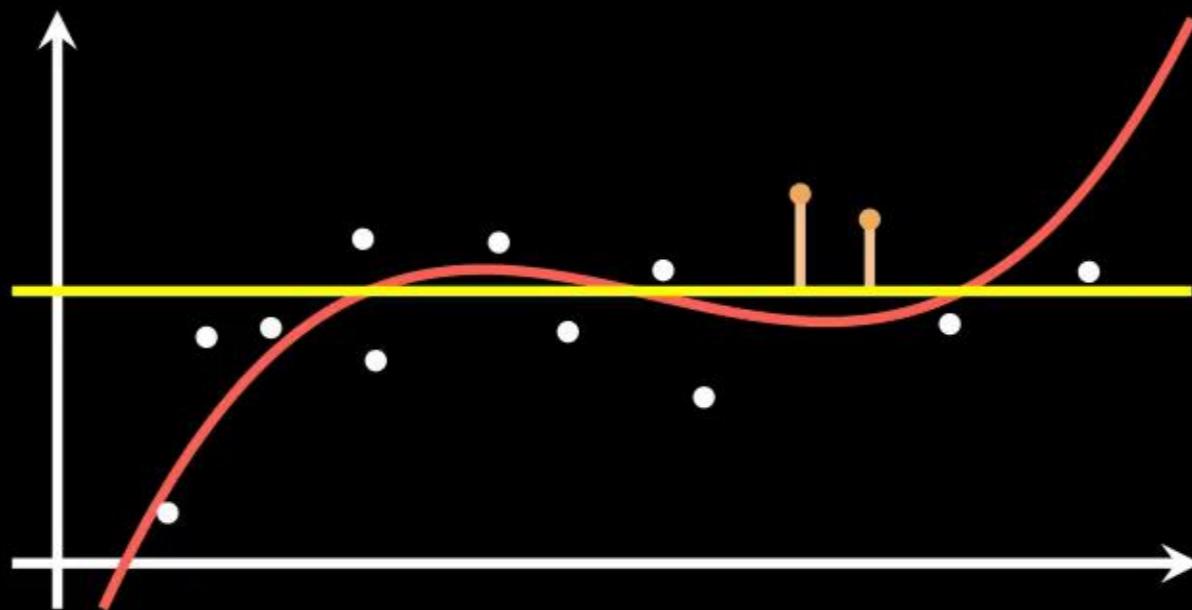
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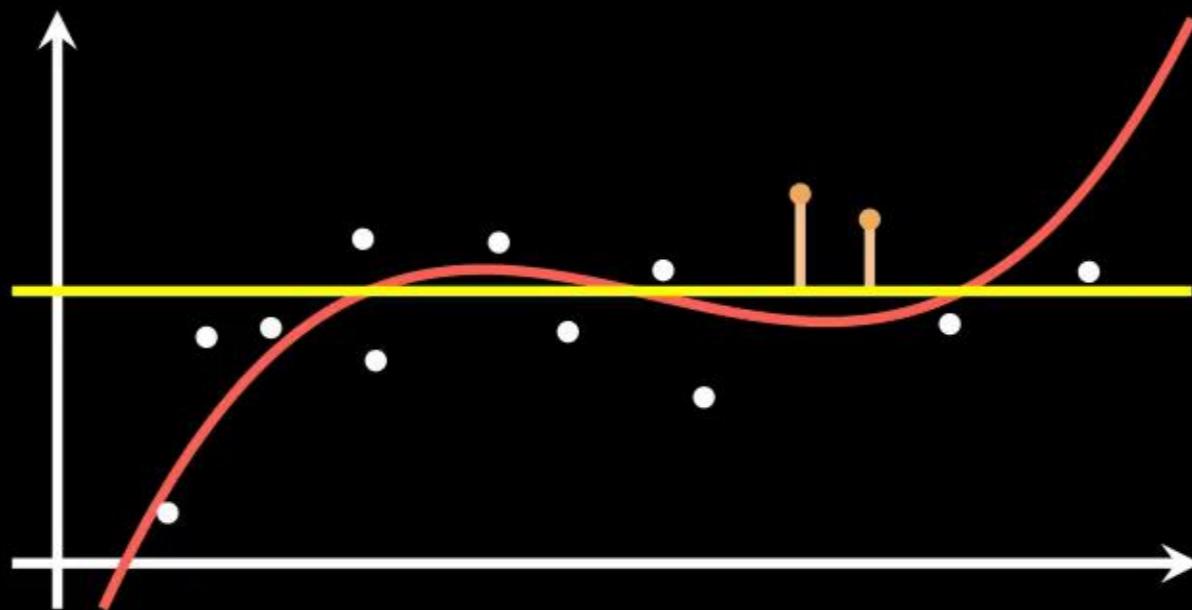
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→ Asymmetry of errors

Statistical setting

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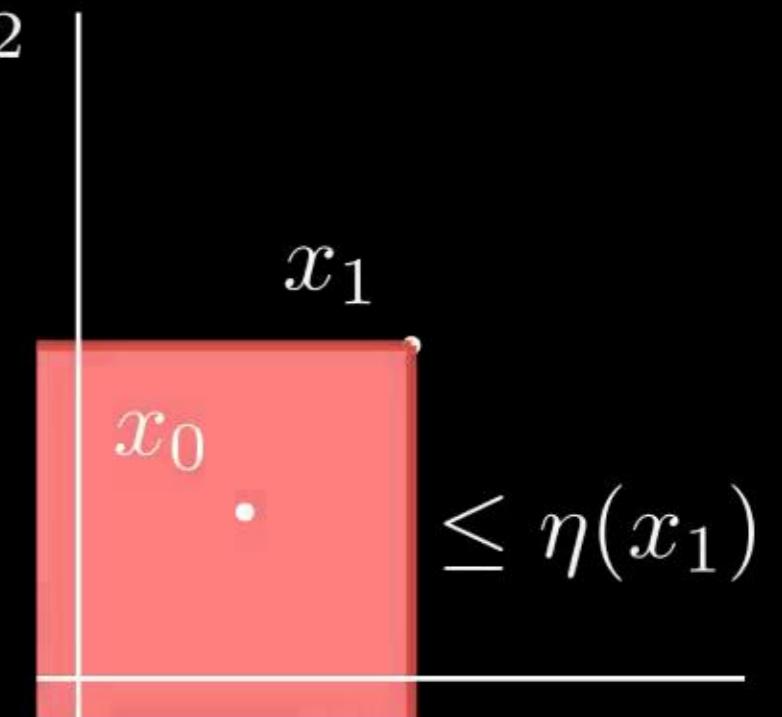
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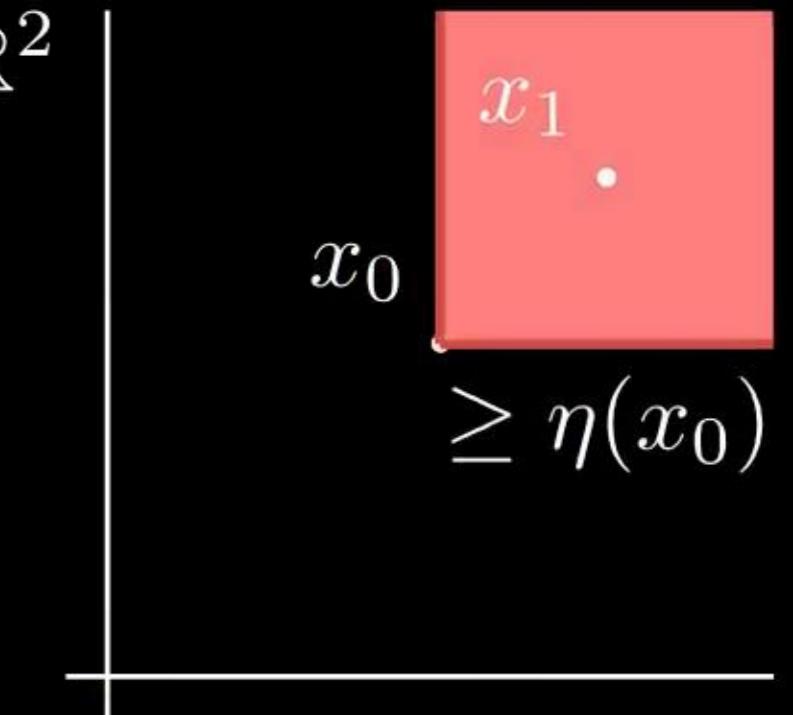


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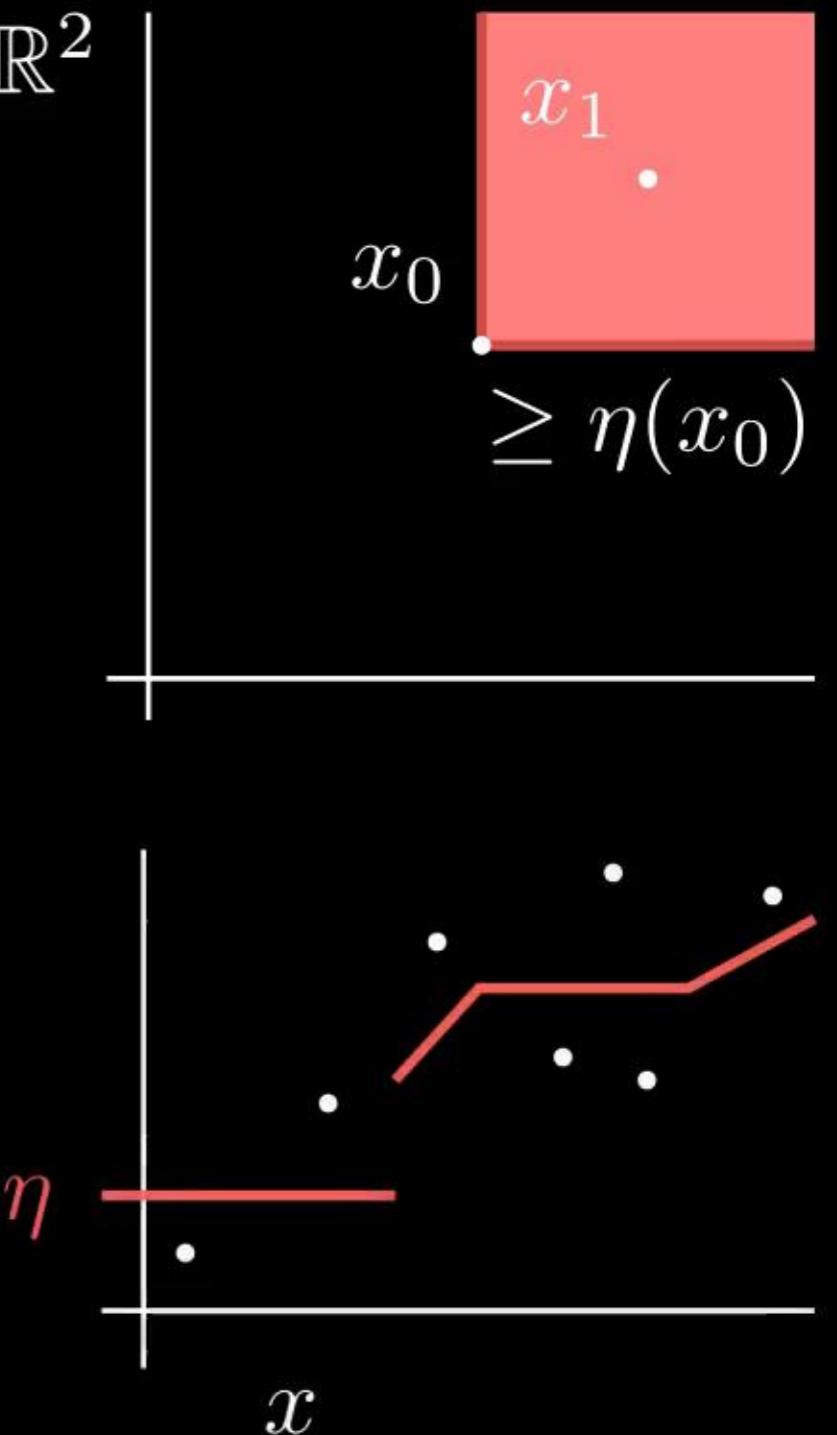


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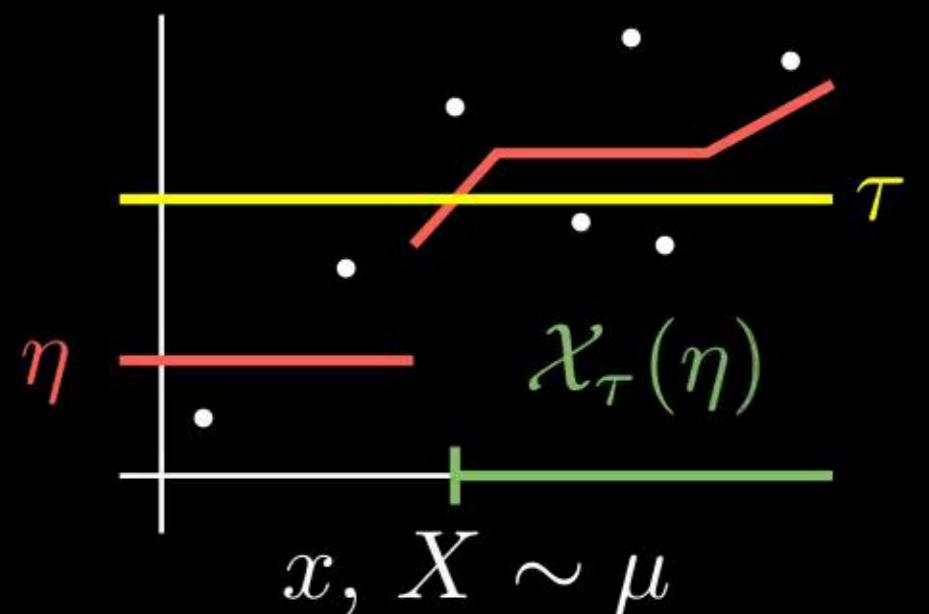
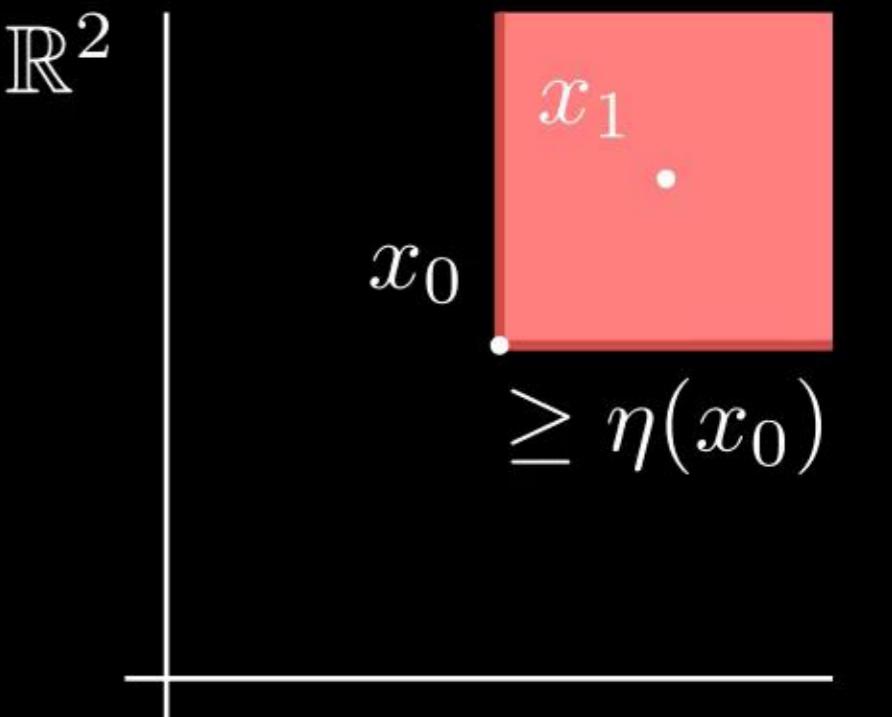
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Notation:

- Fix $\tau \in \mathbb{R}$. Define τ -superlevel set by

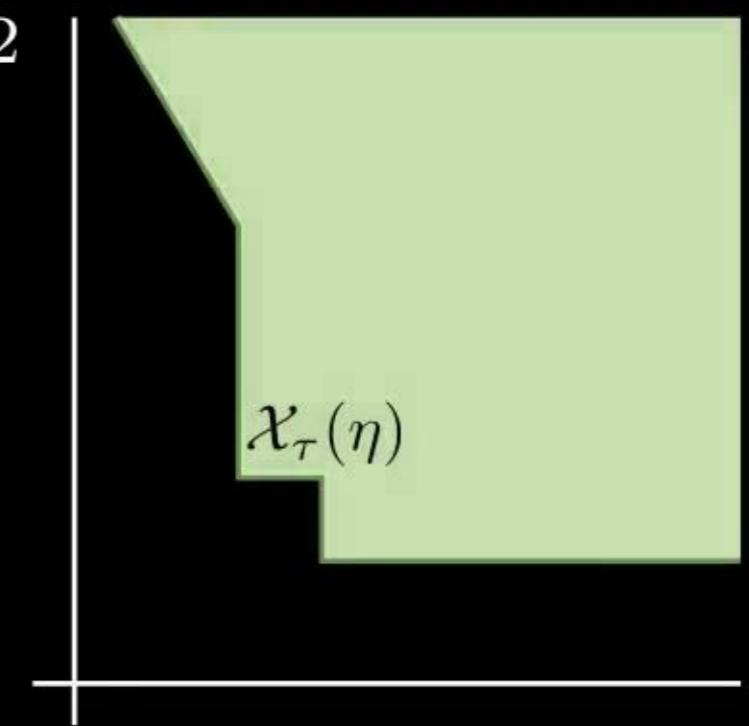
$$\mathcal{X}_\tau(\eta) := \{x \in \mathbb{R}^d : \eta(x) \geq \tau\}$$

- Denote the marginal distribution of X by μ .



Goal

Writing $\mathcal{D} := \left((X_1, Y_1), \dots, (X_n, Y_n) \right) \sim P^n$, we want
 $\hat{A} : \mathcal{D} \mapsto \hat{A}(\mathcal{D}) \subseteq \mathbb{R}^d$ such that:

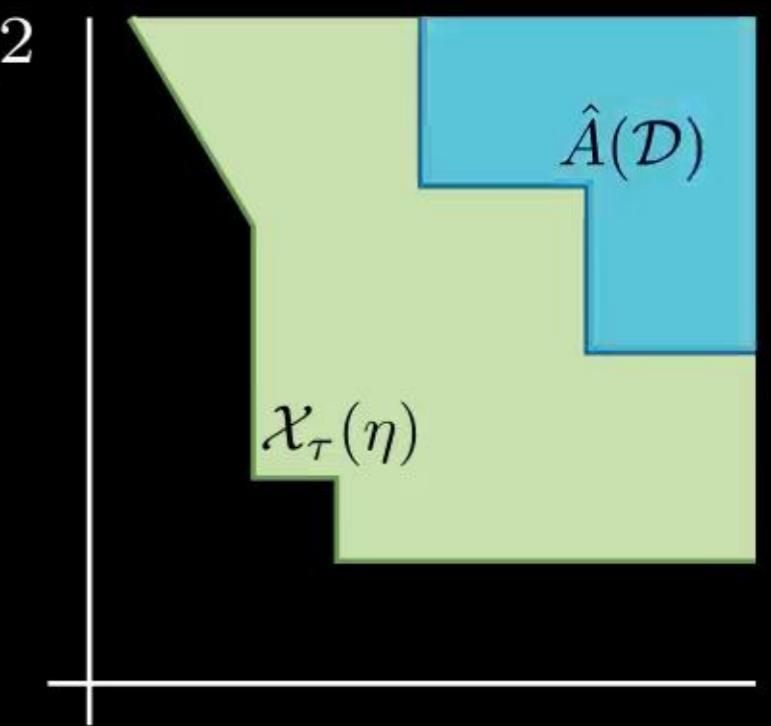


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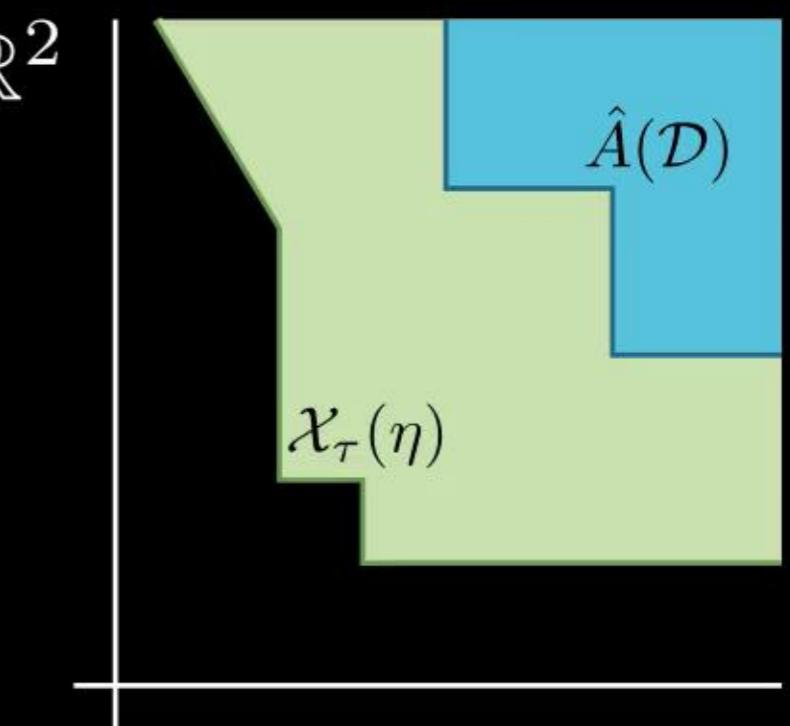
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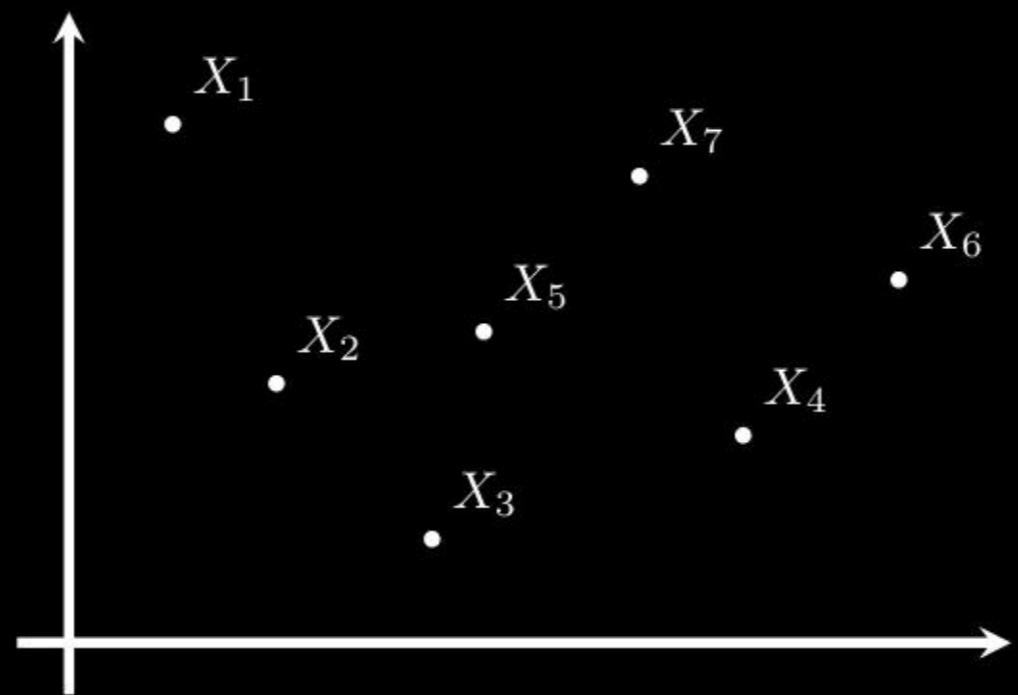
Power: Want small expected regret

$$R_\tau(\hat{A}) := \mathbb{E}\{\mu(\mathcal{X}_\tau(\eta) \setminus \hat{A}(\mathcal{D}))\}.$$



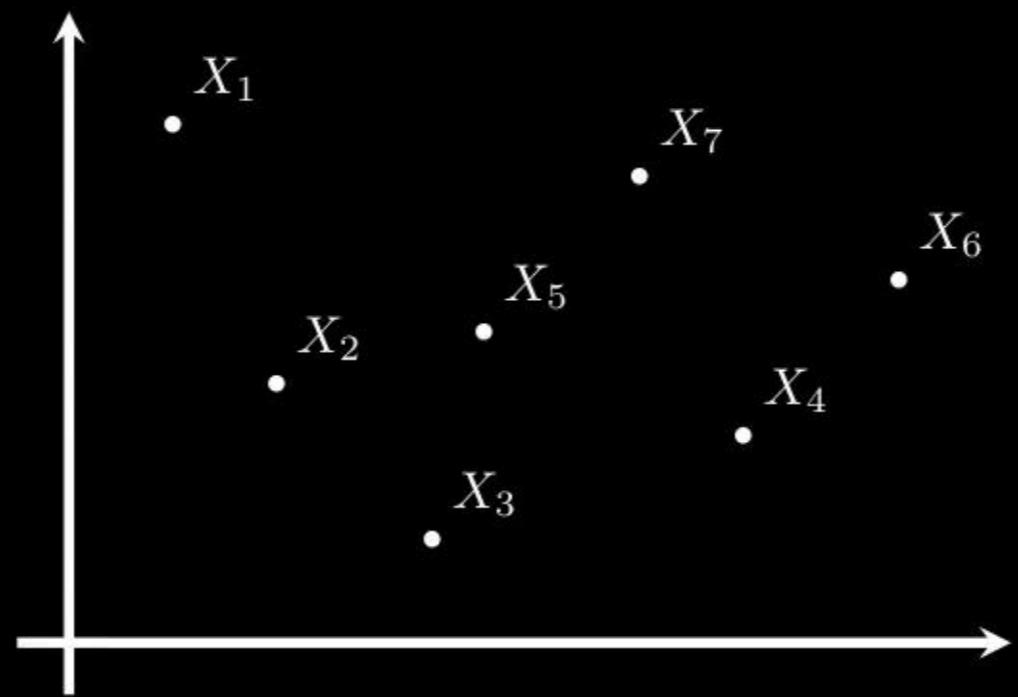
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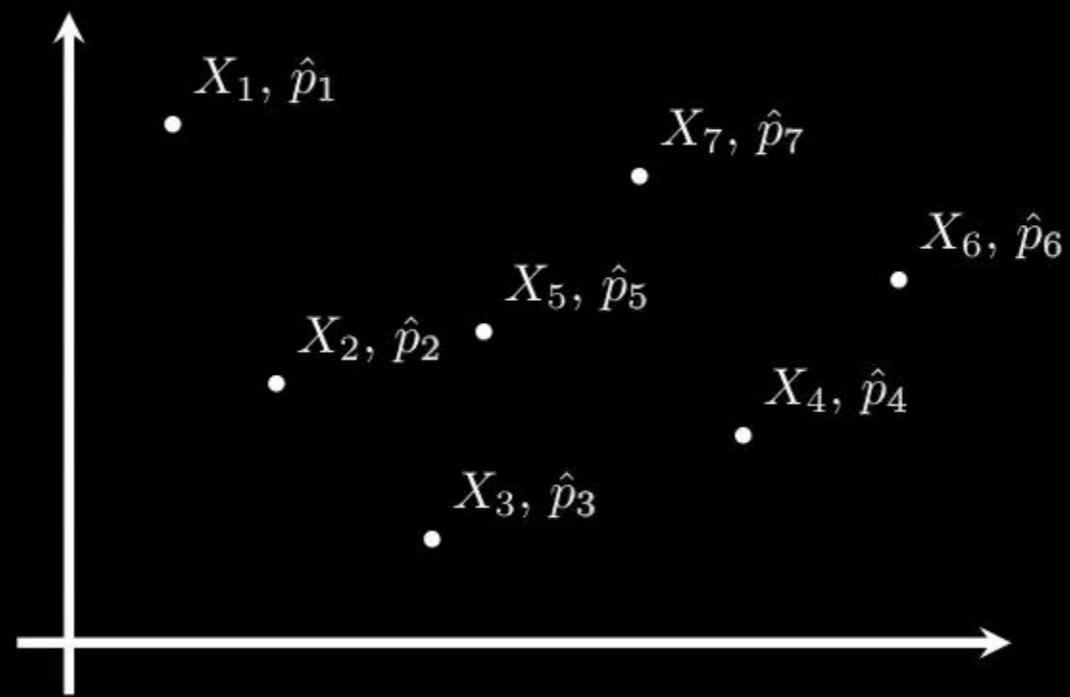
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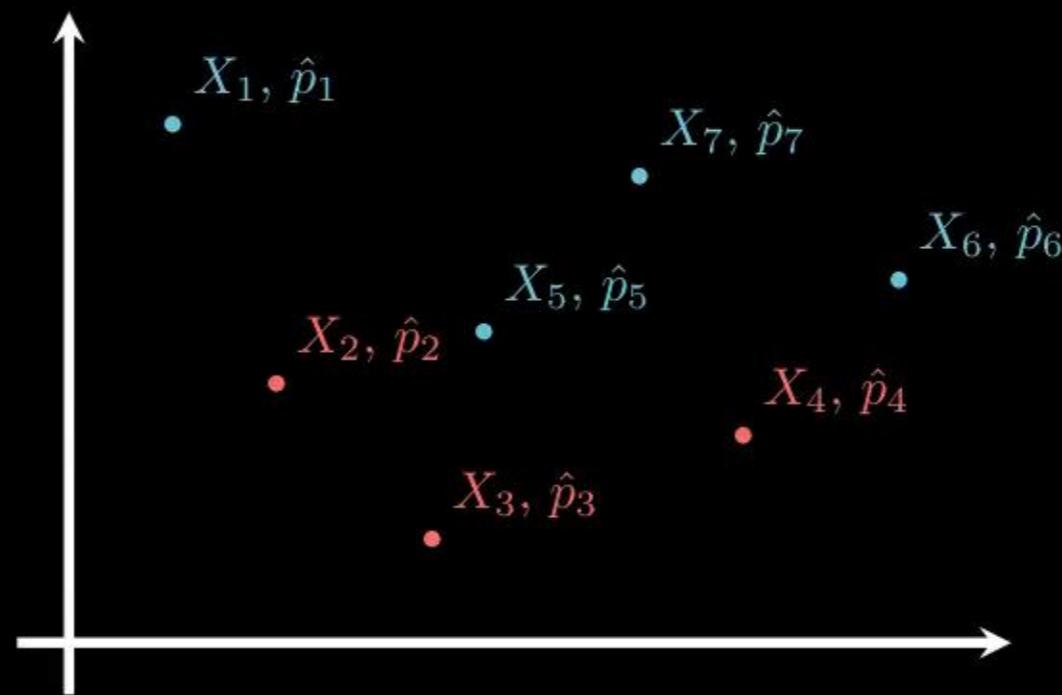


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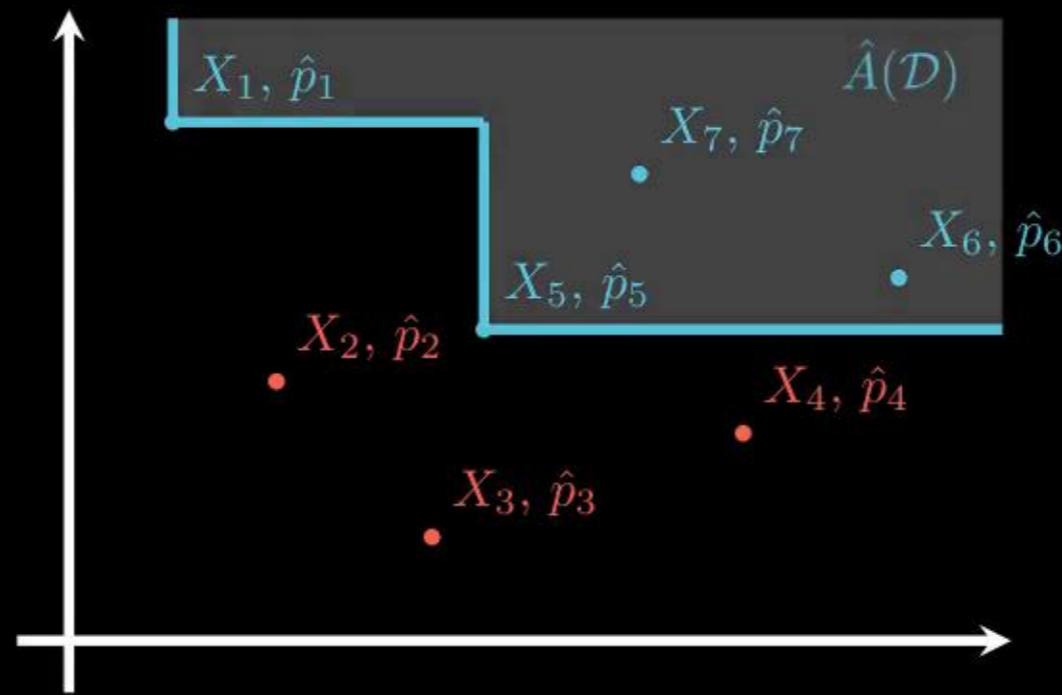


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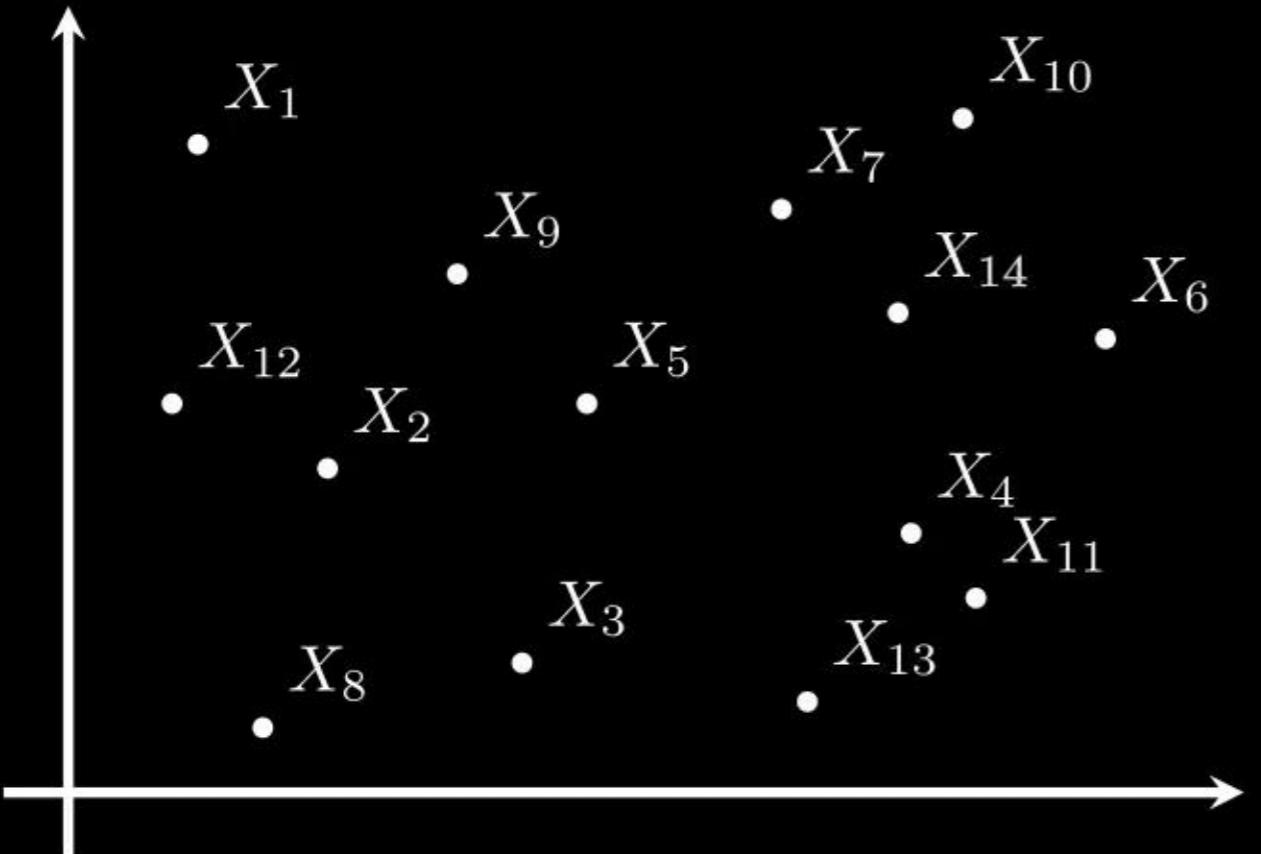
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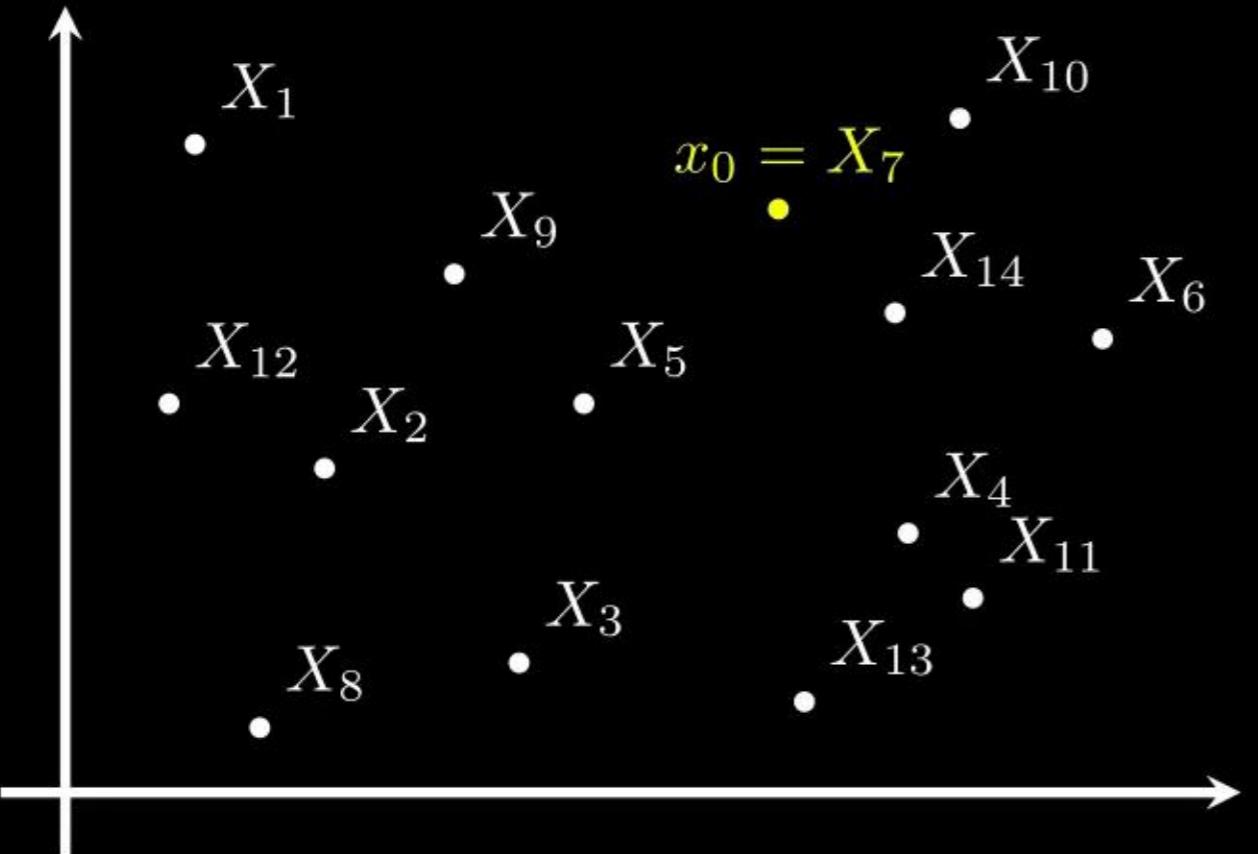
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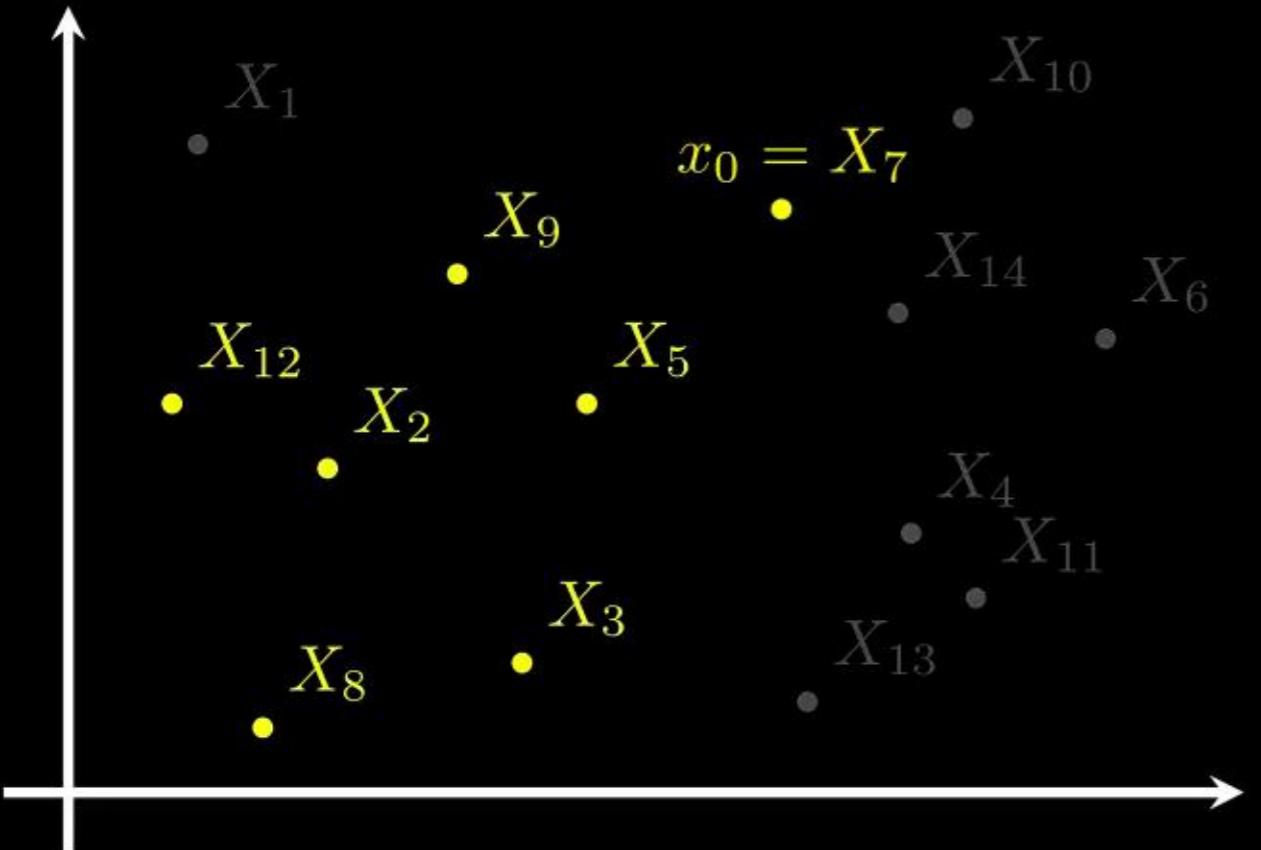
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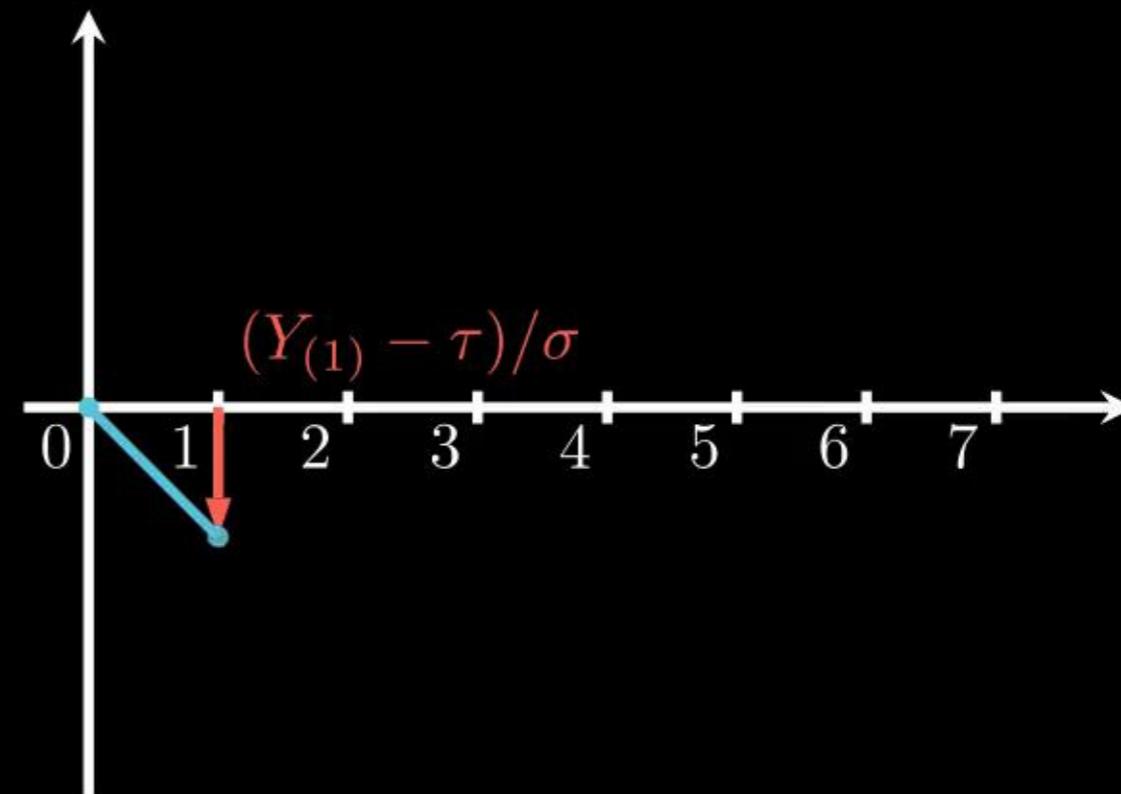
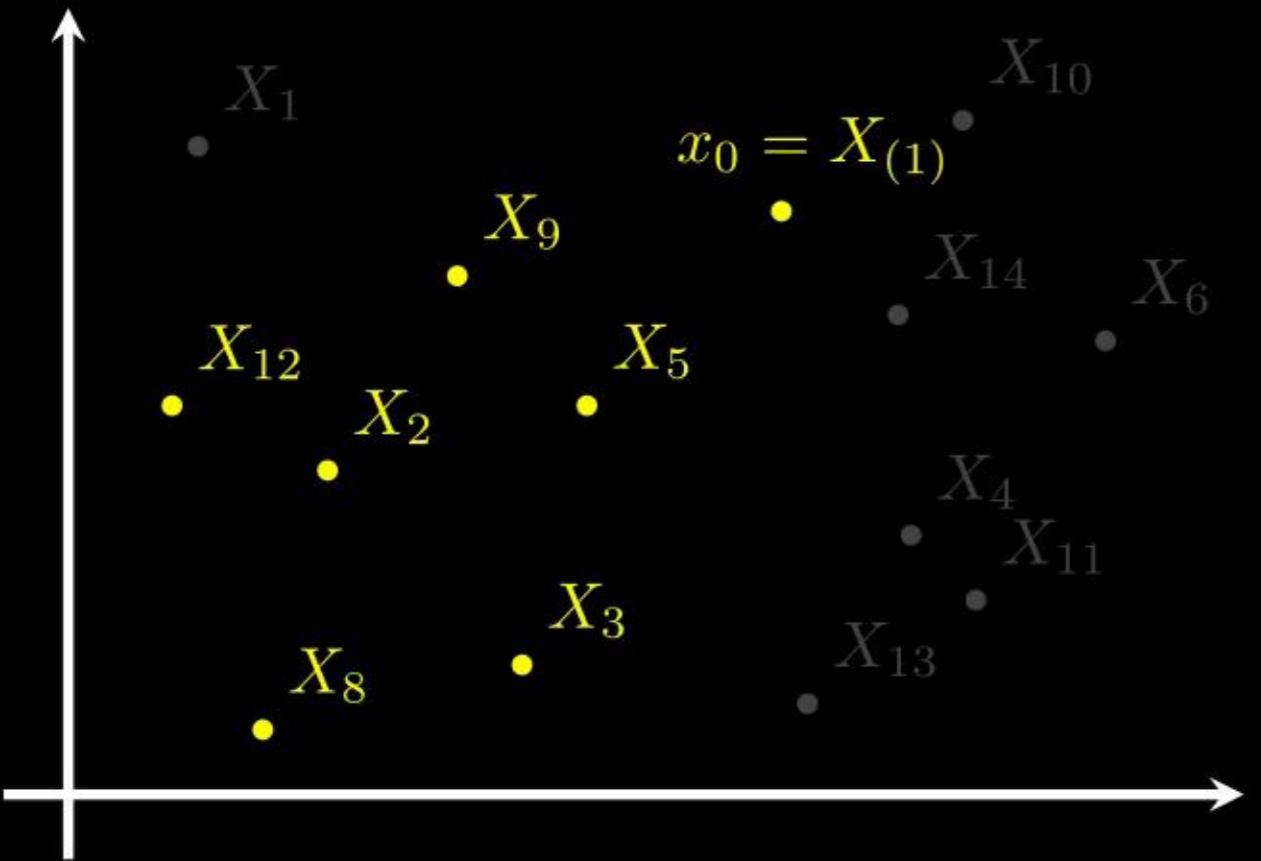
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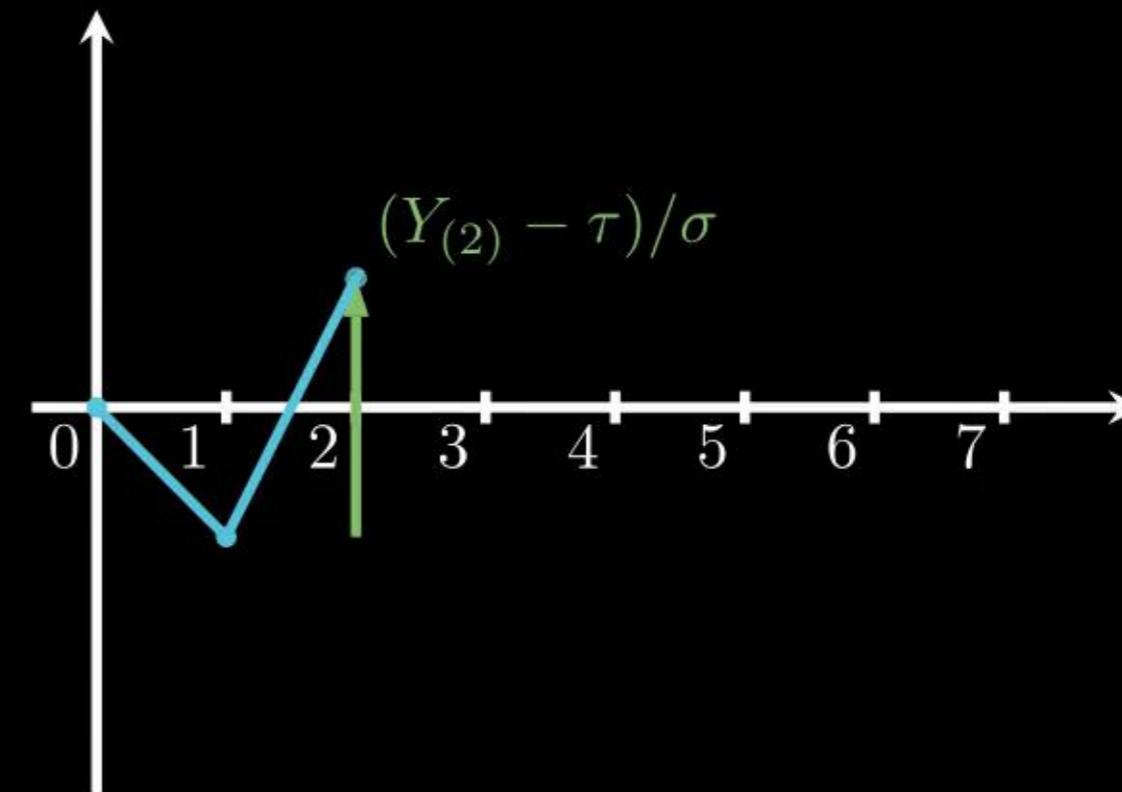
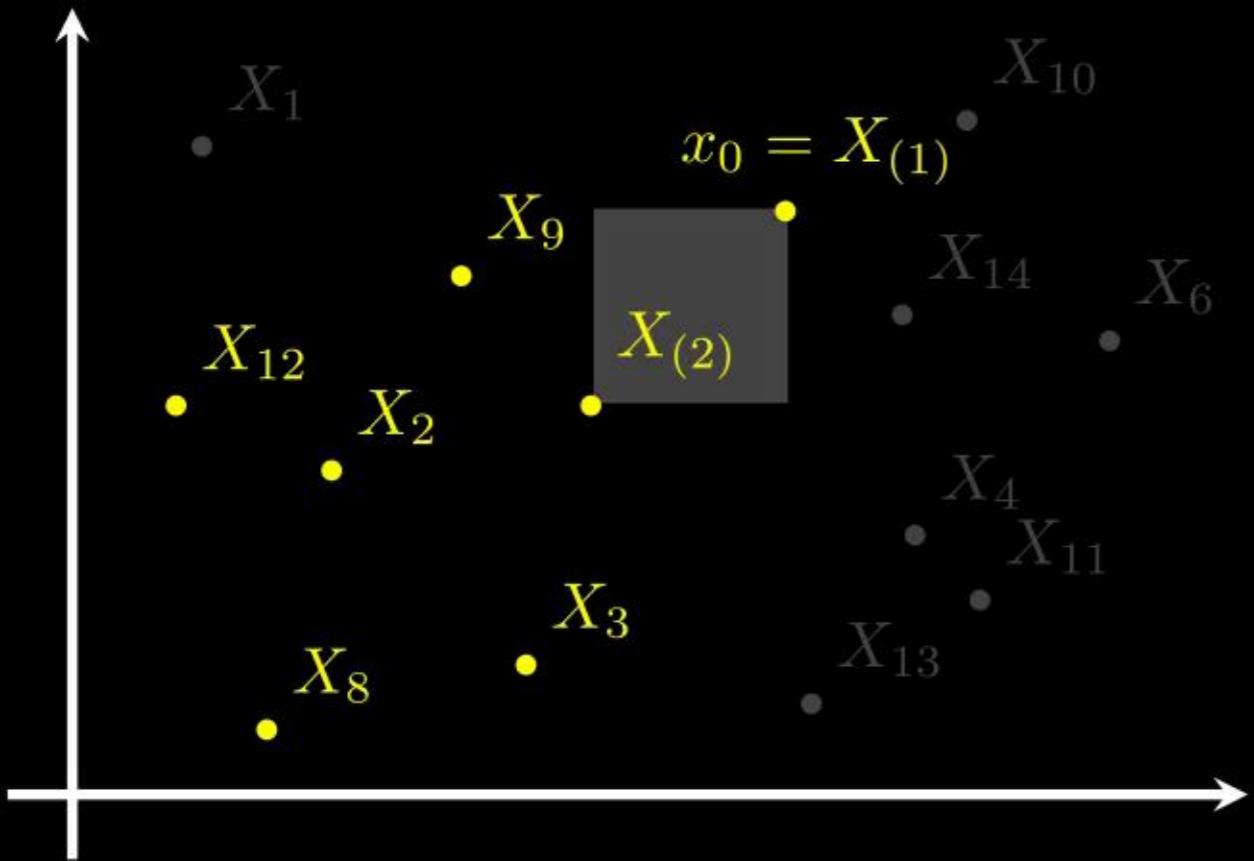
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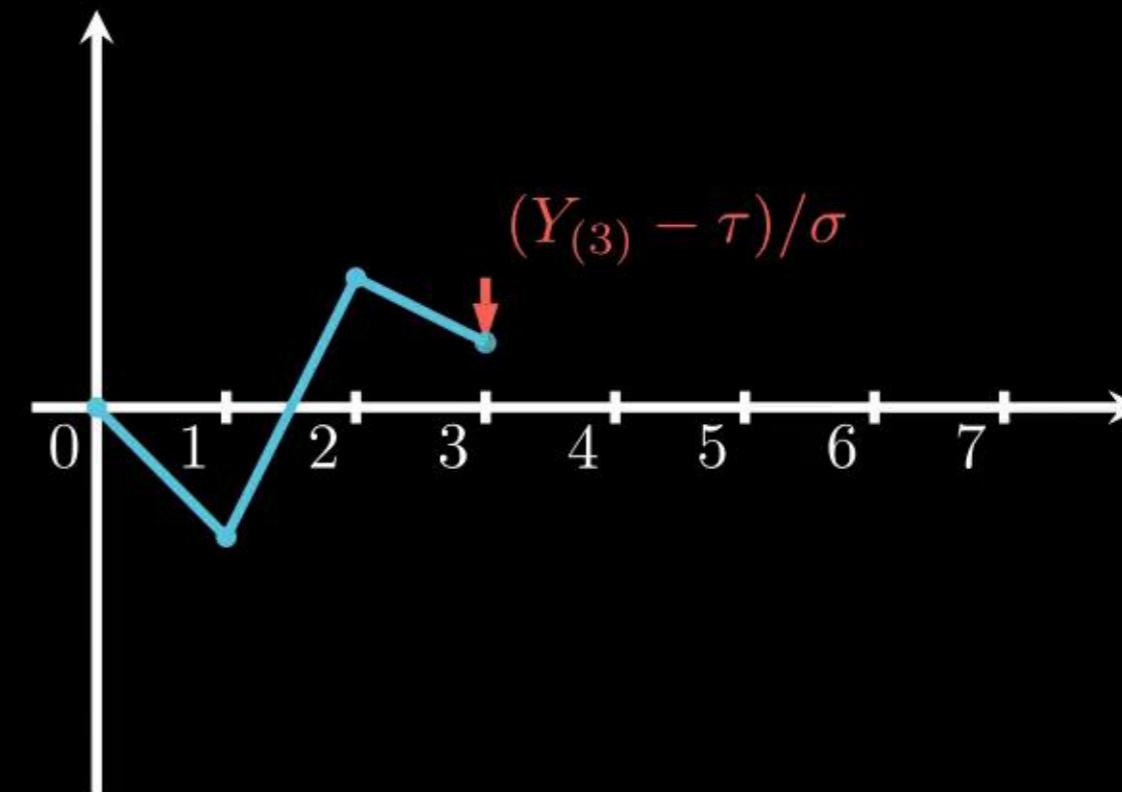
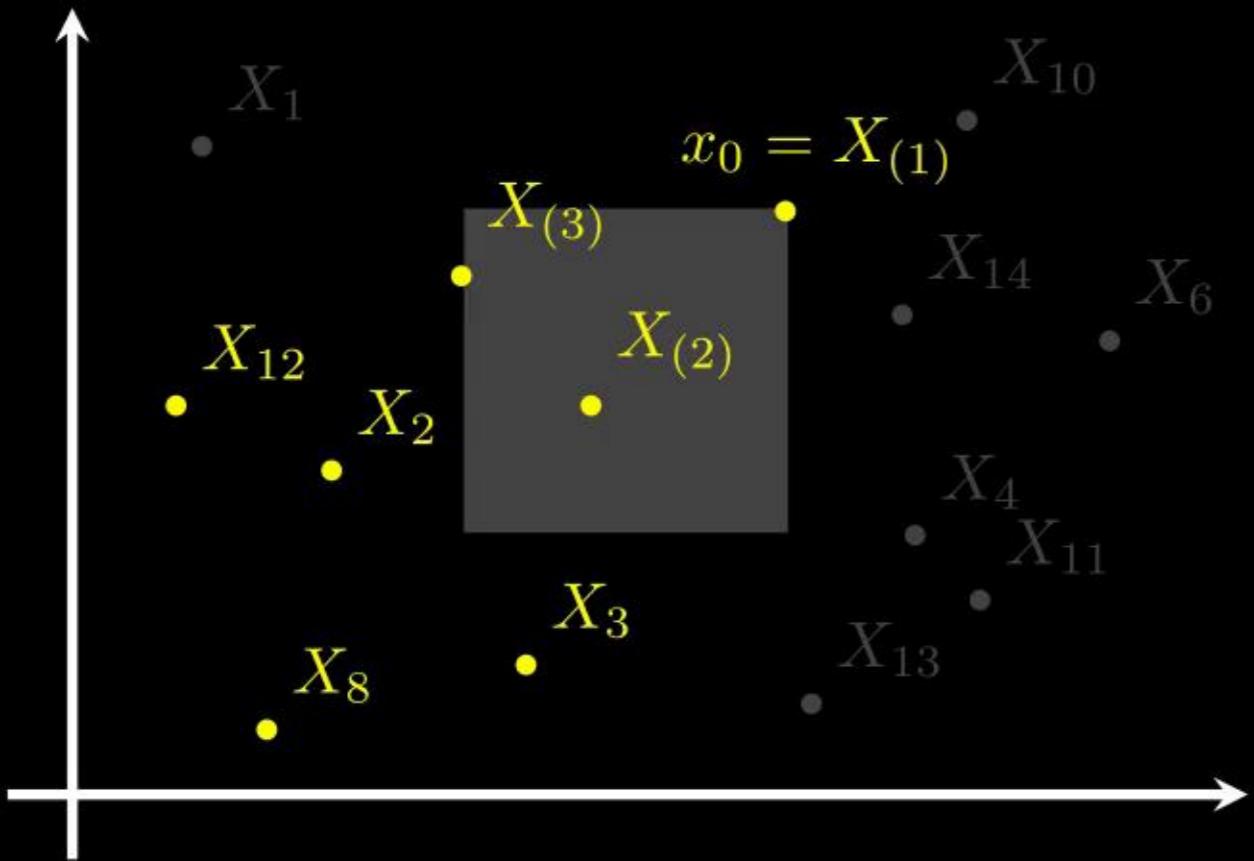
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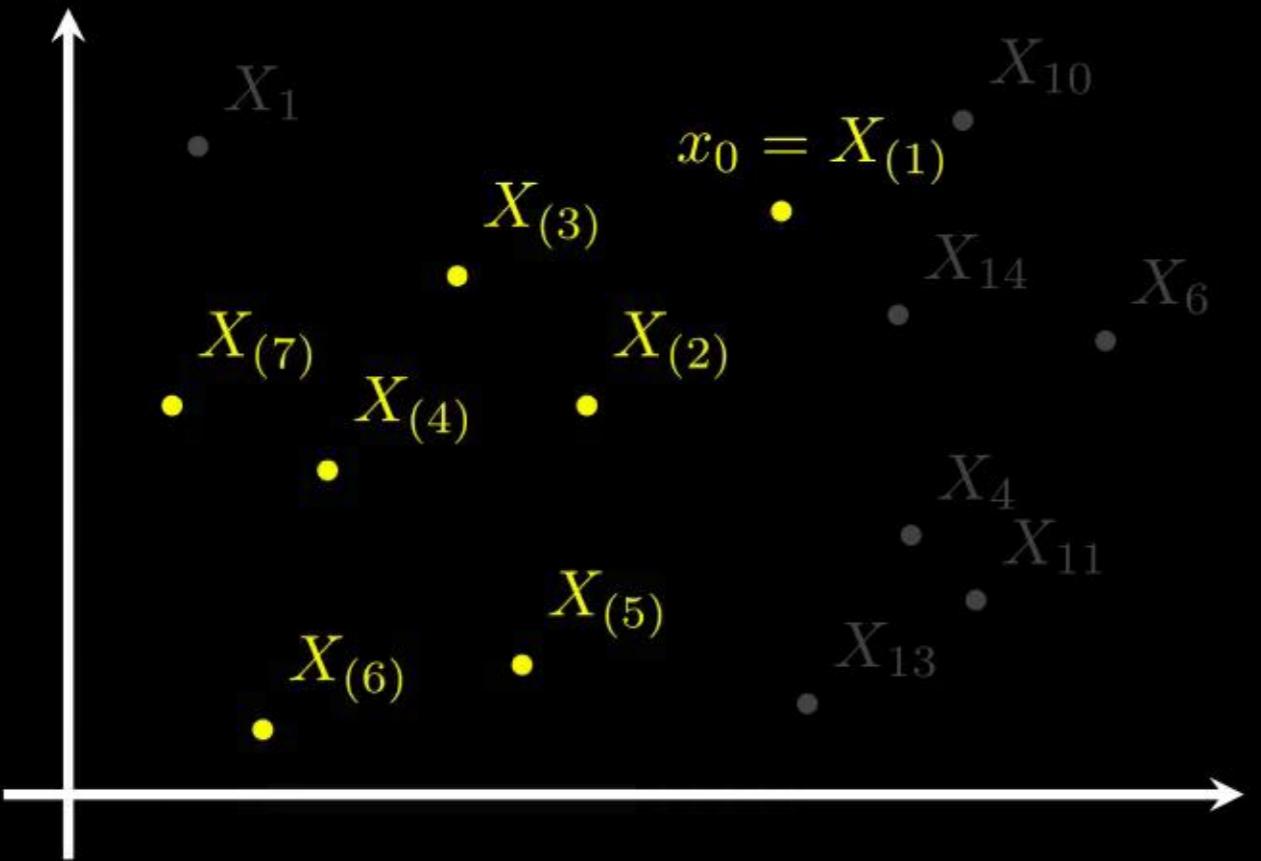
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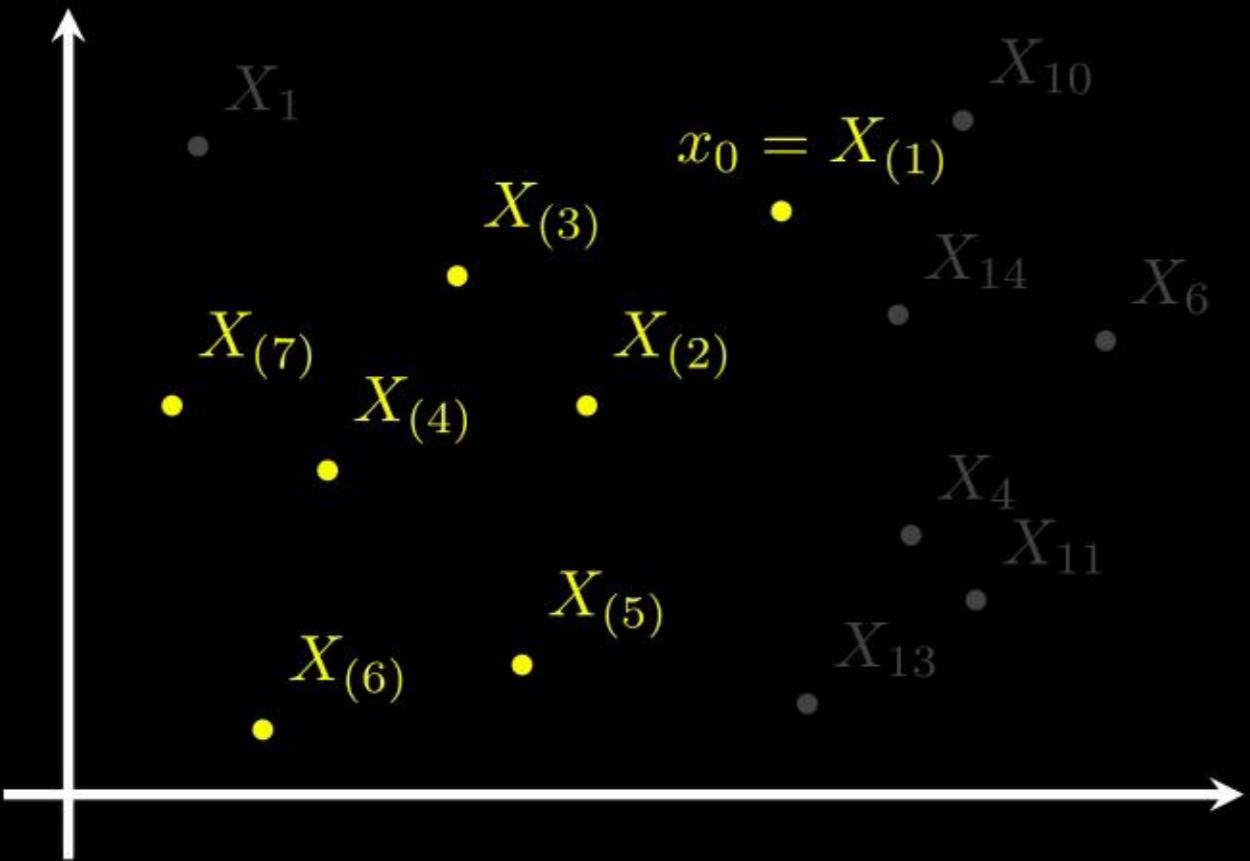
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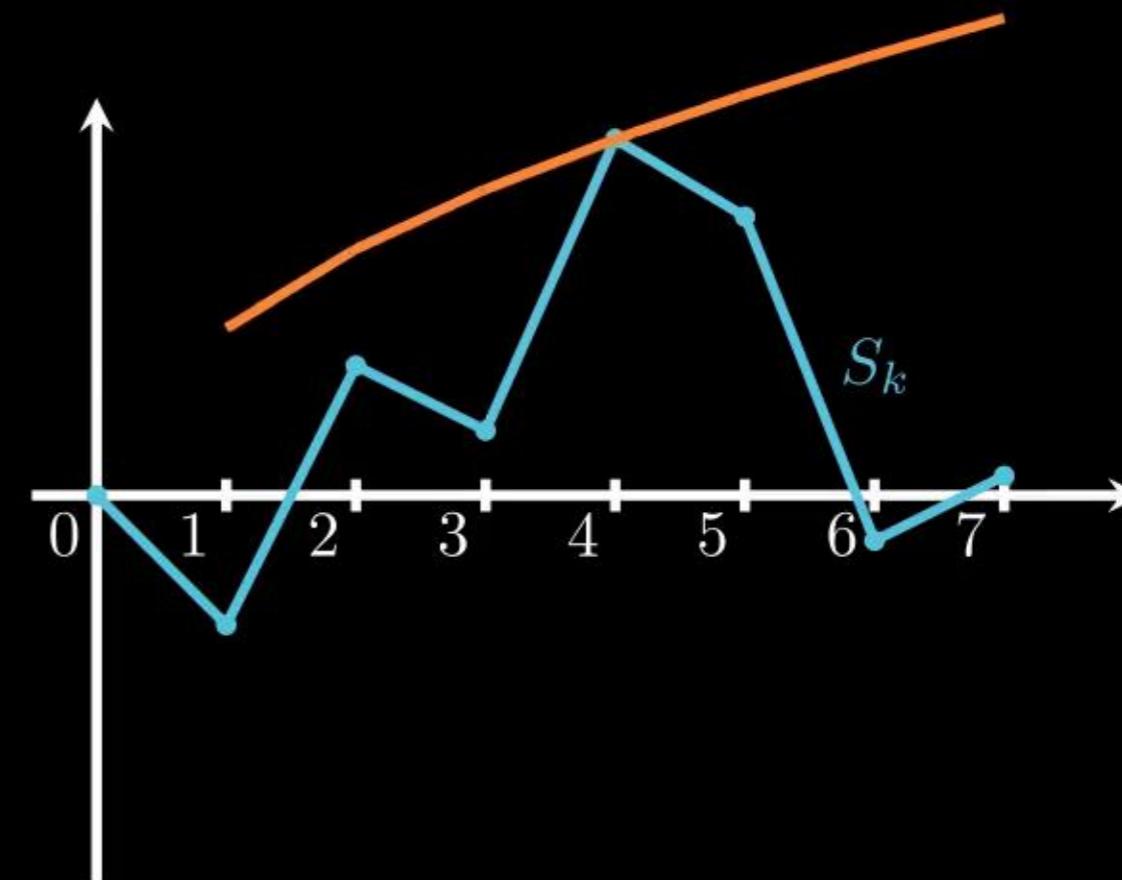
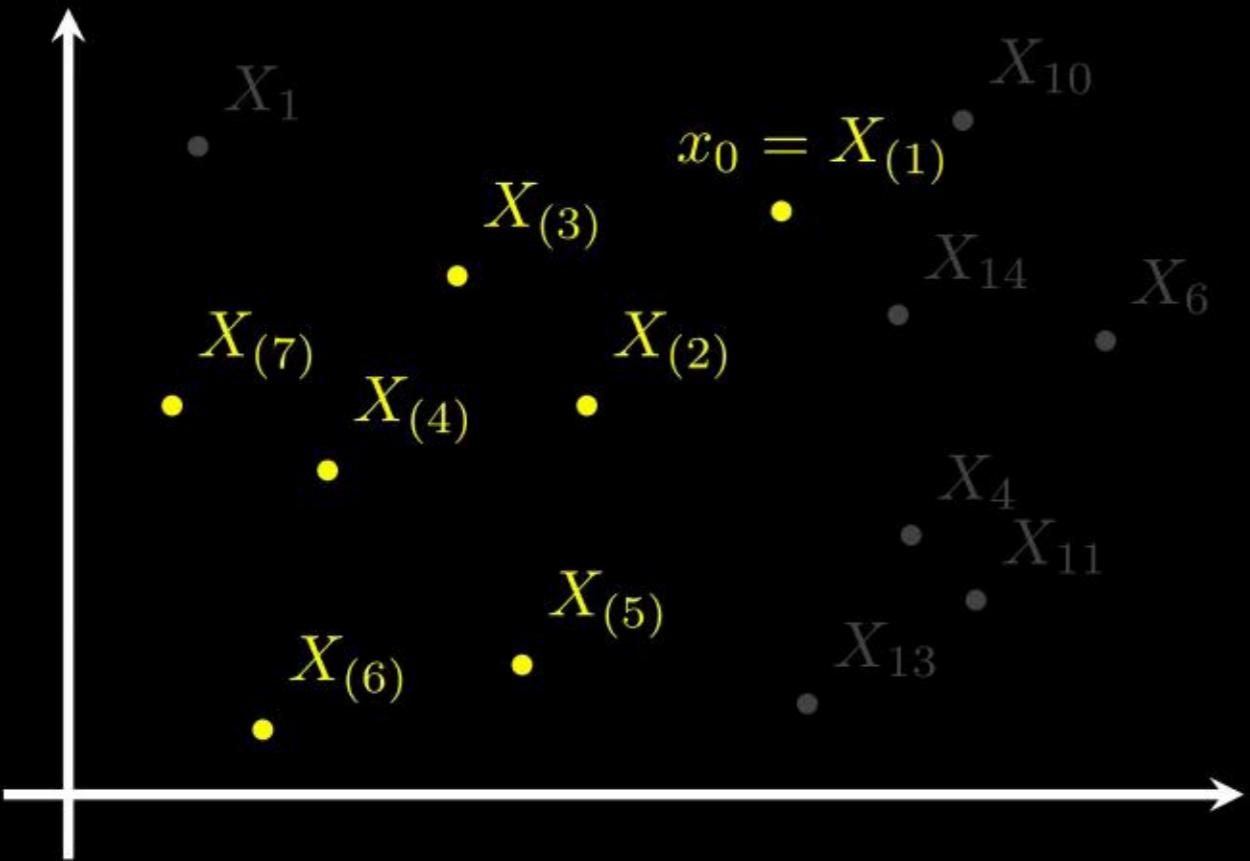
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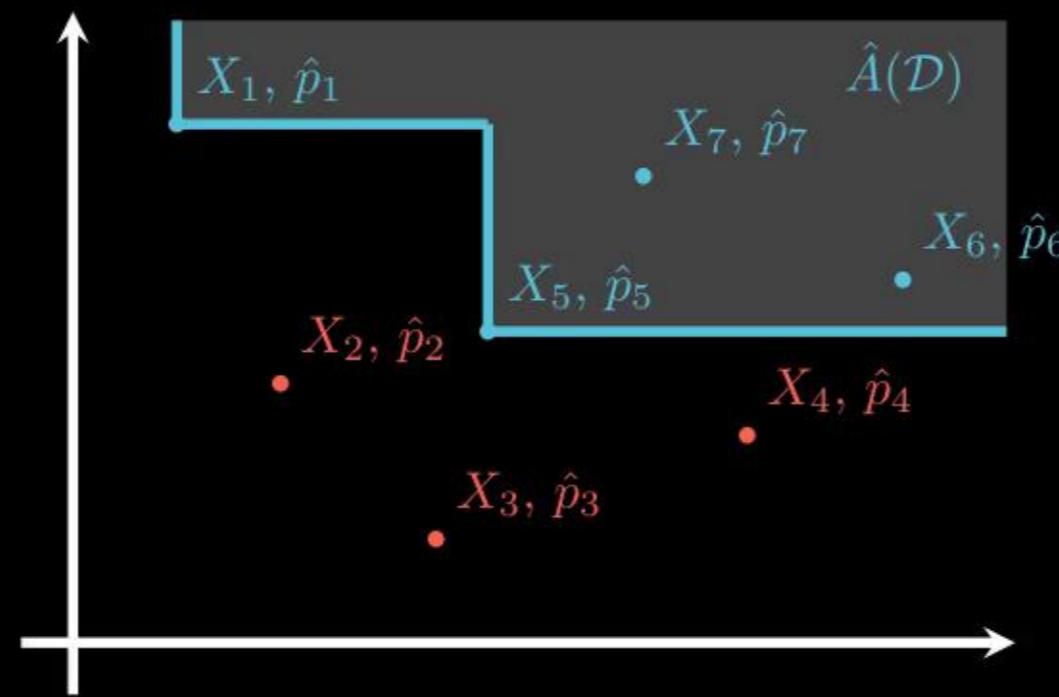
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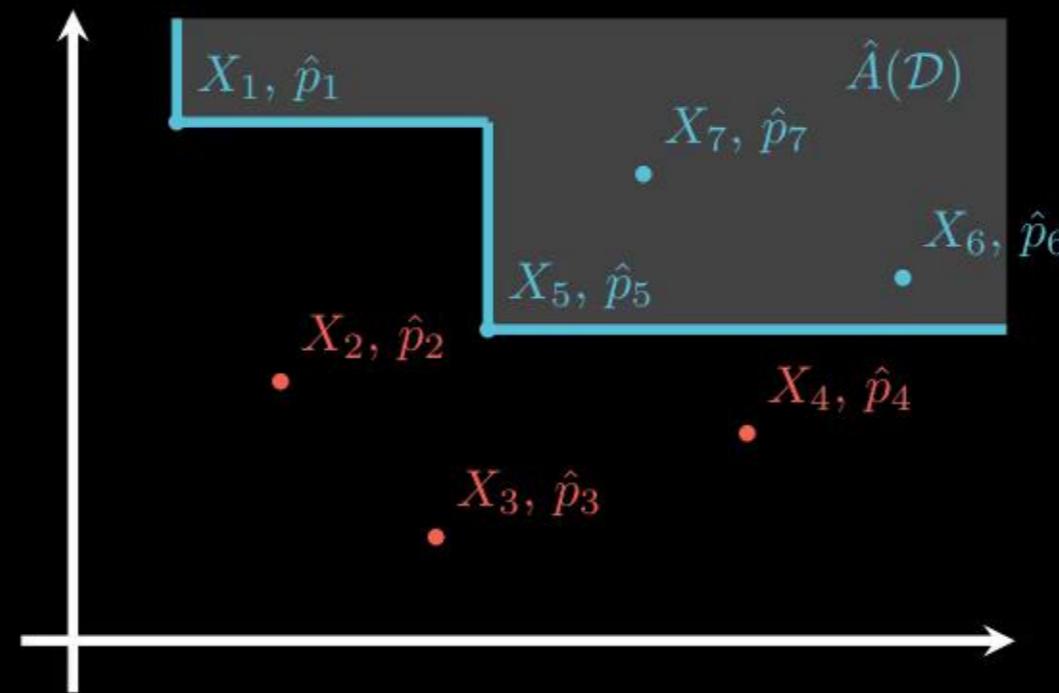


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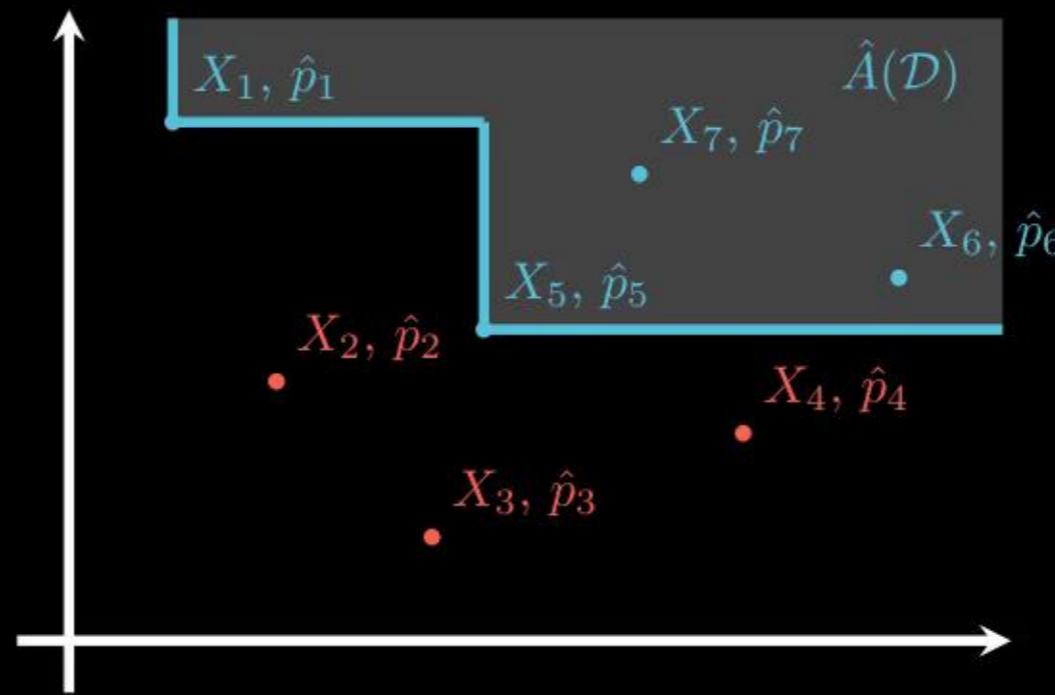


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Theorem. For any $n \geq 1$, $\alpha \in (0, 1)$, $\sigma > 0$, we have

$$\inf_{P \in \mathcal{P}_{\text{Mon}, d}(\sigma)} \mathbb{P}\left(\hat{A}^{\text{ISS}} \subseteq \mathcal{X}_\tau(\eta) \mid \mathcal{D}_X\right) \geq 1 - \alpha.$$

Power and Optimality

Let $\sigma, \gamma, \lambda > 0$, $\theta > 1$. $\mathcal{P}_{\text{MonReg}} \equiv \mathcal{P}_{\text{MonReg},d}(\sigma, \tau, \gamma, \lambda, \theta)$ contains $P \in \mathcal{P}_{\text{Mon},d}(\sigma)$ with:

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Theorem. $\exists C(d, \theta) \geq 1$, such that for any $n \geq 1$ and $\alpha \in (0, 1)$,

$$\sup_{P \in \mathcal{P}_{\text{MonReg}}} R_\tau(\hat{A}^{\text{ISS}}) \leq 1 \wedge C(d, \theta) \left\{ \left(\frac{\sigma^2}{n\lambda^2} \log_+ \left(\frac{n \log_+ n}{\alpha} \right) \right)^{1/(2\gamma+d)} + \left(\frac{\log_+ n}{n} \right)^{1/d} \right\},$$

where $\log_+ x := \log(x \vee e)$ for $x \in \mathbb{R}$.

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$$\sup_{P \in \mathcal{P}_{\text{MonReg}}} R_\tau(\hat{A}^{\text{ISS}}) \leq 1 \wedge C(d, \theta) \left\{ \left(\frac{\sigma^2}{n\lambda^2} \log_+ \left(\frac{n \log_+ n}{\alpha} \right) \right)^{1/(2\gamma+d)} + \left(\frac{\log_+ n}{n} \right)^{1/d} \right\},$$

where $\log_+ x := \log(x \vee e)$ for $x \in \mathbb{R}$.

Theorem. $\exists c(d, \gamma) \in (0, 1)$, such that for any $n \geq 1$ and $\alpha \in (0, 1)$,

$$\inf_{\hat{A} \in \hat{\mathcal{A}}_n(\tau, \alpha, \mathcal{P}_{\text{MonReg}})} \sup_{P \in \mathcal{P}_{\text{MonReg}}} R_\tau(\hat{A}) \geq c(d, \gamma) \left[1 \wedge \left\{ \left(\frac{\sigma^2}{n\lambda^2} \log_+ \left(\frac{1}{5\alpha} \right) \right)^{1/(2\gamma+d)} + \frac{1}{n^{1/d}} \right\} \right].$$

Power and Optimality

Let $\sigma, \gamma, \lambda > 0$, $\theta > 1$. $\mathcal{P}_{\text{MonReg}} \equiv \mathcal{P}_{\text{MonReg},d}(\sigma, \tau, \gamma, \lambda, \theta)$ contains $P \in \mathcal{P}_{\text{Mon},d}(\sigma)$ with:

1. η growing at least as quickly as $r \mapsto \lambda \cdot r^\gamma$ on the superlevel set near its boundary,
2. μ resembling the Lebesgue measure up to a constant θ .

Theorem. $\exists C(d, \theta) \geq 1$, such that for any $n \geq 1$ and $\alpha \in (0, 1)$,

$$\sup_{P \in \mathcal{P}_{\text{MonReg}}} R_\tau(\hat{A}^{\text{ISS}}) \leq 1 \wedge C(d, \theta) \left\{ \left(\frac{\sigma^2}{n\lambda^2} \log_+ \left(\frac{n \log_+ n}{\alpha} \right) \right)^{1/(2\gamma+d)} + \left(\frac{\log_+ n}{n} \right)^{1/d} \right\},$$

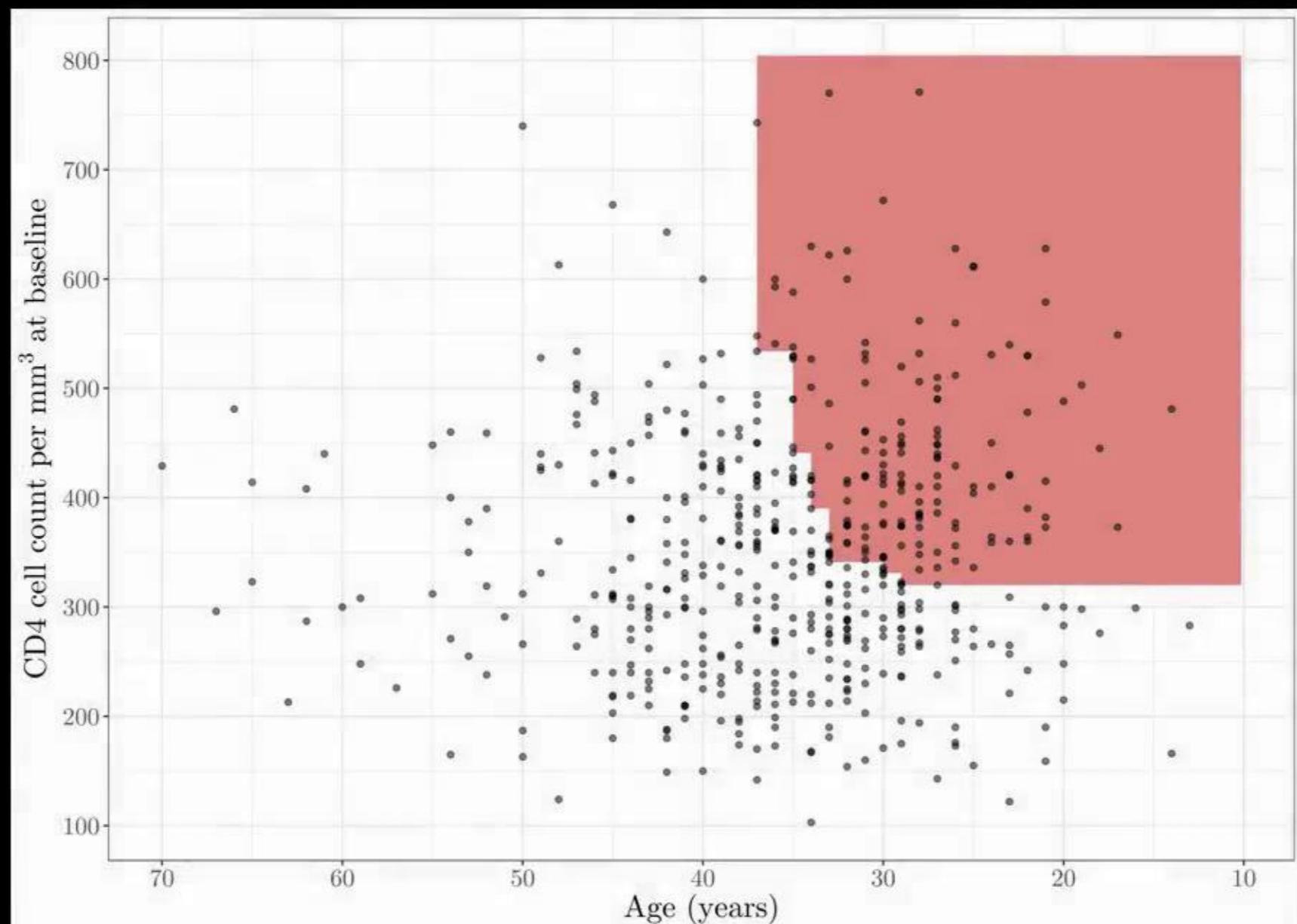
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$$\inf_{\hat{A} \in \hat{\mathcal{A}}_n(\tau, \alpha, \mathcal{P}_{\text{MonReg}})} \sup_{P \in \mathcal{P}_{\text{MonReg}}} R_\tau(\hat{A}) \geq c(d, \gamma) \left[1 \wedge \left\{ \left(\frac{\sigma^2}{n\lambda^2} \log_+ \left(\frac{1}{5\alpha} \right) \right)^{1/(2\gamma+d)} + \frac{1}{n^{1/d}} \right\} \right].$$

Application

Primary endpoint: reduction of the CD4 cell count by 50%, development of AIDS, or death, with median follow-up duration 143 weeks (Hammer et al., 1996). Let $\alpha = 0.05$ and $\tau = 1/2$.



Extensions in the paper

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Take-home messages

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In common situations, no smoothing-parameters have to be specified.

References and acknowledgement

- Duan, B., Ramdas, A., Balakrishnan, S., and Wasserman, L. (2020). Interactive martingale tests for the global null. *Electronic Journal of Statistics*, 14(2):4489–4551.
- Goeman, J. J. and Solari, A. (2010). The sequential rejection principle of familywise error control. *The Annals of Statistics*, 38(6):3782–3810.
- Hammer, S. M., Katzenstein, D. A., Hughes, M. D., Gundacker, H., Schooley, R. T., Haubrich, R. H., Henry, W. K., Lederman, M. M., Phair, J. P., Niu, M., et al. (1996). A trial comparing nucleoside monotherapy with combination therapy in HIV-infected adults with CD4 cell counts from 200 to 500 per cubic millimeter. *New England Journal of Medicine*, 335(15):1081–1090.
- Howard, S. R., Ramdas, A., McAuliffe, J., and Sekhon, J. (2021). Time-uniform, nonparametric, nonasymptotic confidence sequences. *The Annals of Statistics*, 49:1055–1080.
- Meijer, R. J. and Goeman, J. J. (2015). A multiple testing method for hypotheses structured in a directed acyclic graph. *Biometrical Journal*, 57(1):123–143.
- O'Mahony, C., Jichi, F., Pavlou, M., Monserrat, L., Anastasakis, A., Rapezzi, C., Biagini, E., Gimeno, J. R., Limongelli, G., McKenna, W. J., et al. (2014). A novel clinical risk prediction model for sudden cardiac death in hypertrophic cardiomyopathy (HCM risk-SCD). *European Heart Journal*, 35(30):2010–2020.
- Wasserman, L., Ramdas, A., and Balakrishnan, S. (2020). Universal inference. *Proceedings of the National Academy of Sciences*, 117(29):16880–16890.

I am grateful for receiving the travel award that allowed me to present at ICSDS 2023.

Thank you!

Main reference:

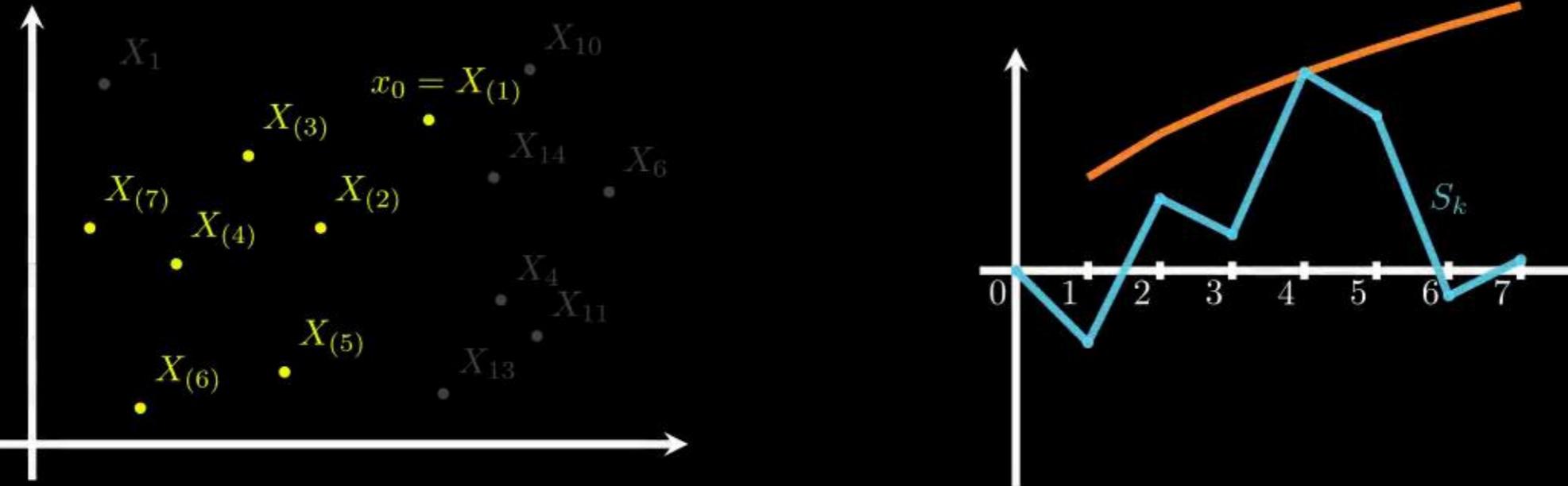
Müller, M. M., Reeve, H. W. J., Cannings, T. I. and Samworth, R. J. (2023) Isotonic subgroup selection. *arXiv preprint arXiv:2305.04852*.

See manuelmmueller.github.io for data and R-code.

Appendix

Construct p -values \hat{p}_i for $H_0(X_i)$, $i \in [n]$

Given $x_0 \in \mathbb{R}^d$, we seek a p -value for $H_0(x_0) := \{P \in \mathcal{P}_{\text{Mon},d}(\sigma) : \eta(x_0) < \tau\}$.

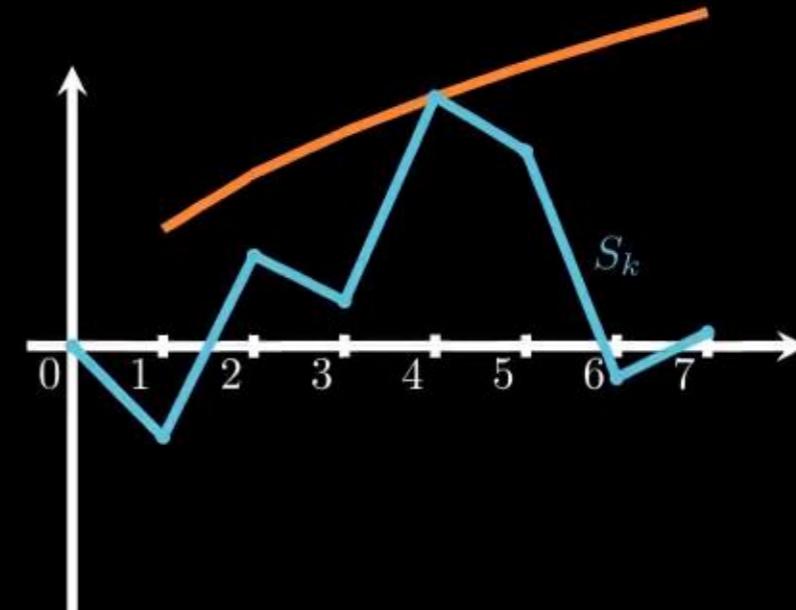
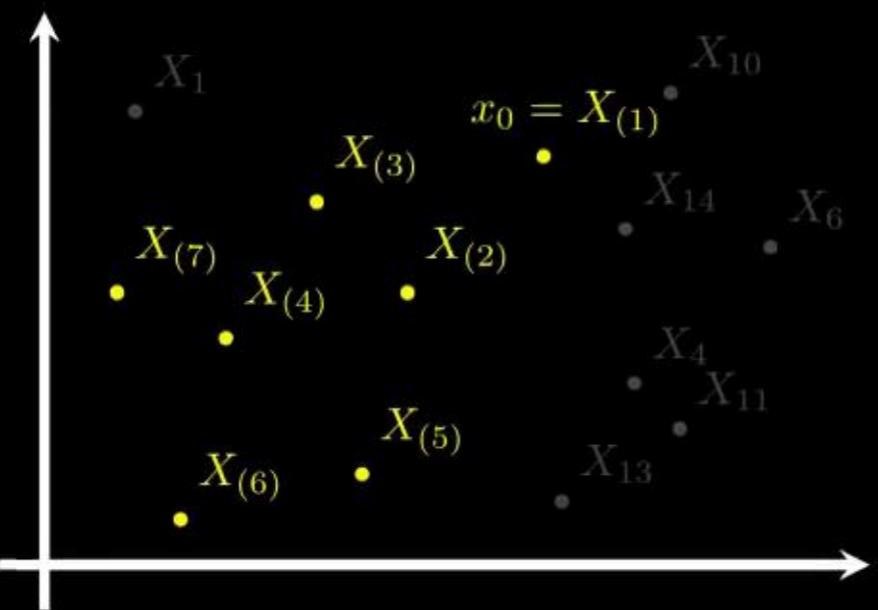


Denote $\mathcal{I}(x_0) := \{i \in [n] : X_i \preceq x_0\}$, $n(x_0) := |\mathcal{I}(x_0)|$.

Let $X_{(j)}$ be the j th nearest neighbour of x_0 among X_i , $i \in \mathcal{I}(x_0)$, in sup-norm and let $Y_{(j)}$ be the corresponding response. Let $S_k := \sum_{j=1}^k (Y_{(j)} - \tau)/\sigma$.

Then S_k is a supermartingale under $P \in H_0(x_0)$. Combination with time-uniform bounds by Howard et al. (2021) gives p -values from this martingale test (Duan et al., 2020).

Construct p -values \hat{p}_i for $H_0(X_i)$, $i \in [n]$

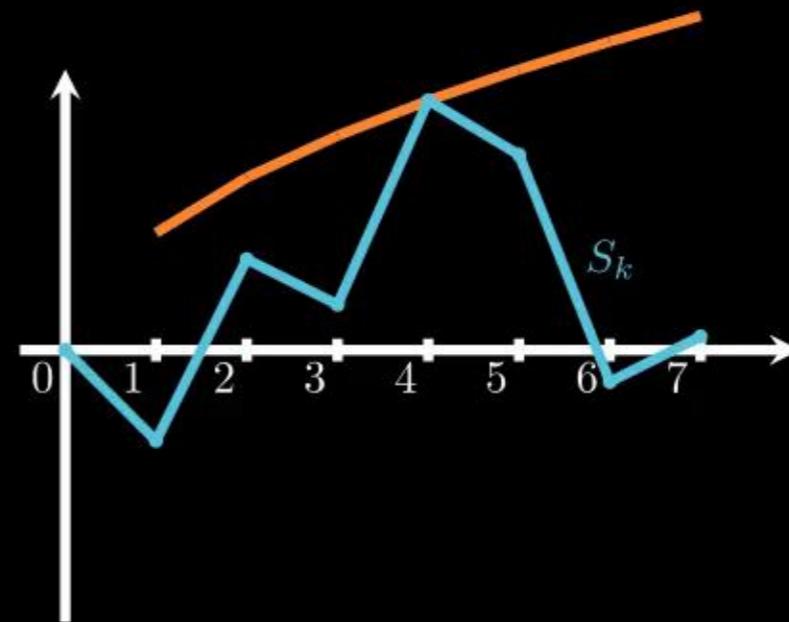
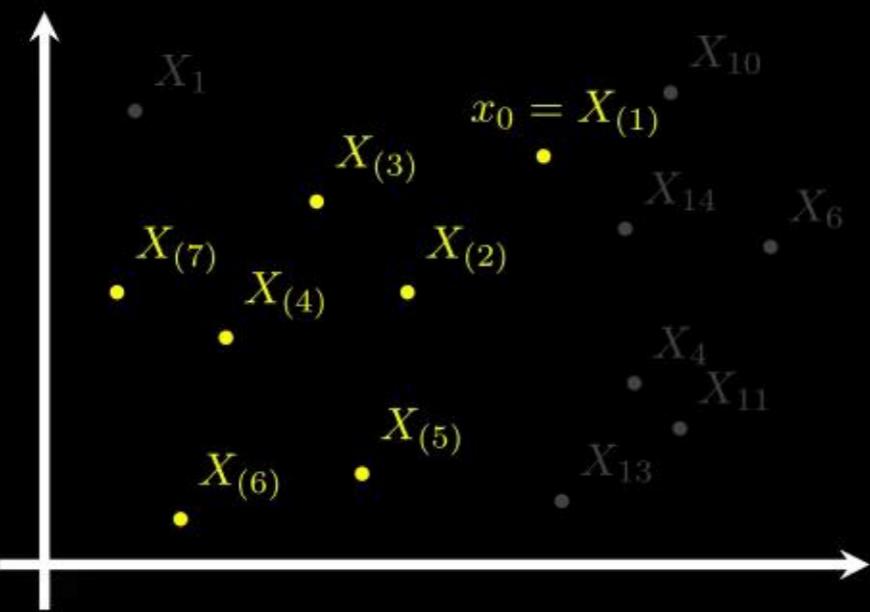


Definition. For $\sigma > 0$, $x \in \mathbb{R}^d$, let

$$\hat{p}_{\sigma,\tau}(x) := 1 \wedge \min_{k \in [n(x)]} 5.2 \exp \left\{ -\frac{(S_k \vee 0)^2}{2.0808k} + \frac{\log \log(2k)}{0.72} \right\},$$

whenever $n(x) > 0$, and $\hat{p}_{\sigma,\tau}(x) := 1$ otherwise.

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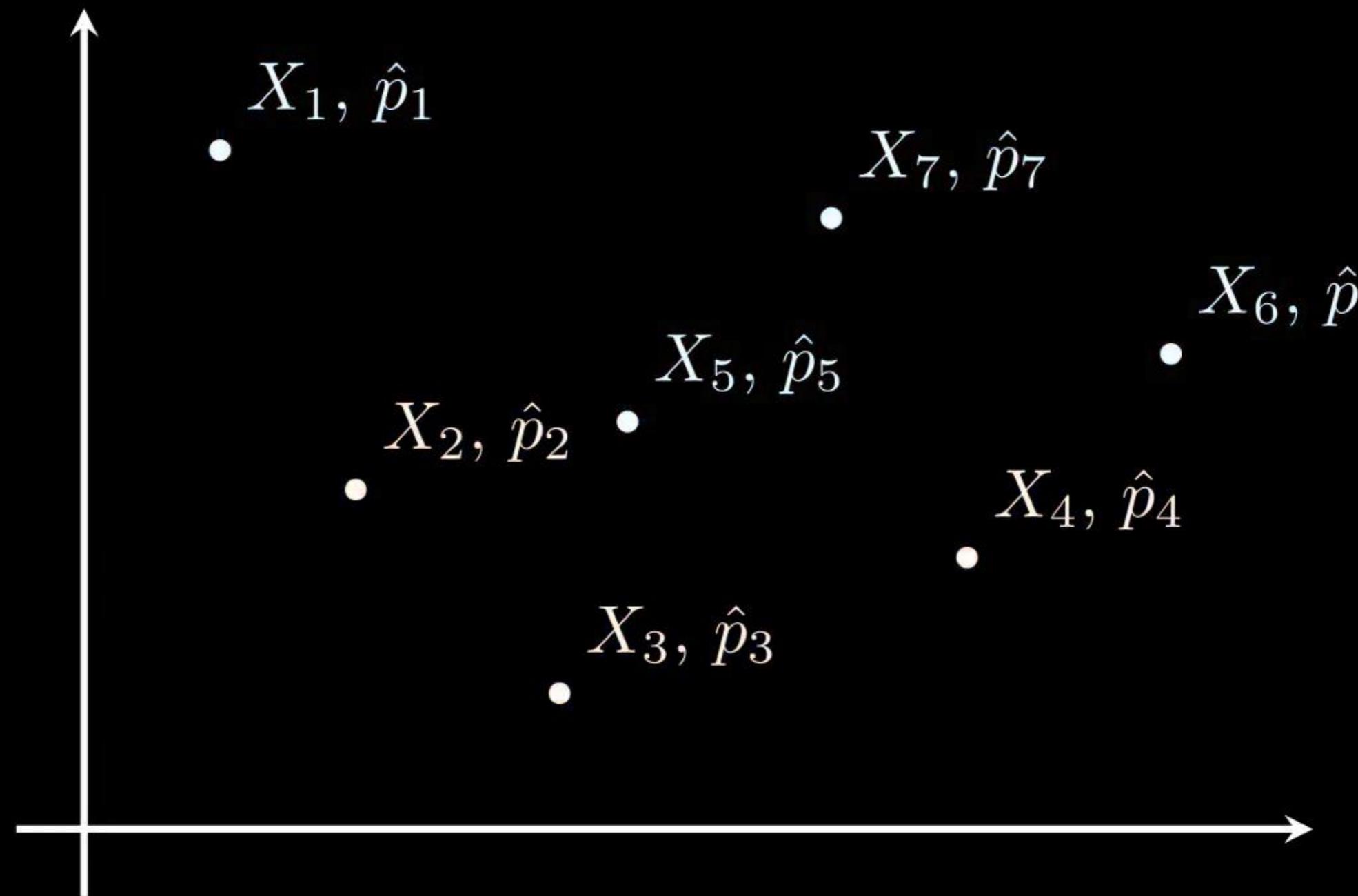
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whenever $n(x) > 0$, and $\hat{p}_{\sigma,\tau}(x) := 1$ otherwise.

Lemma. When $\eta(x) < \tau$, we have $\mathbb{P}\{\hat{p}_{\sigma,\tau}(x) \leq t \mid (X_i)_{i \in [n]}\} \leq t$ for all $t \in (0, 1)$.

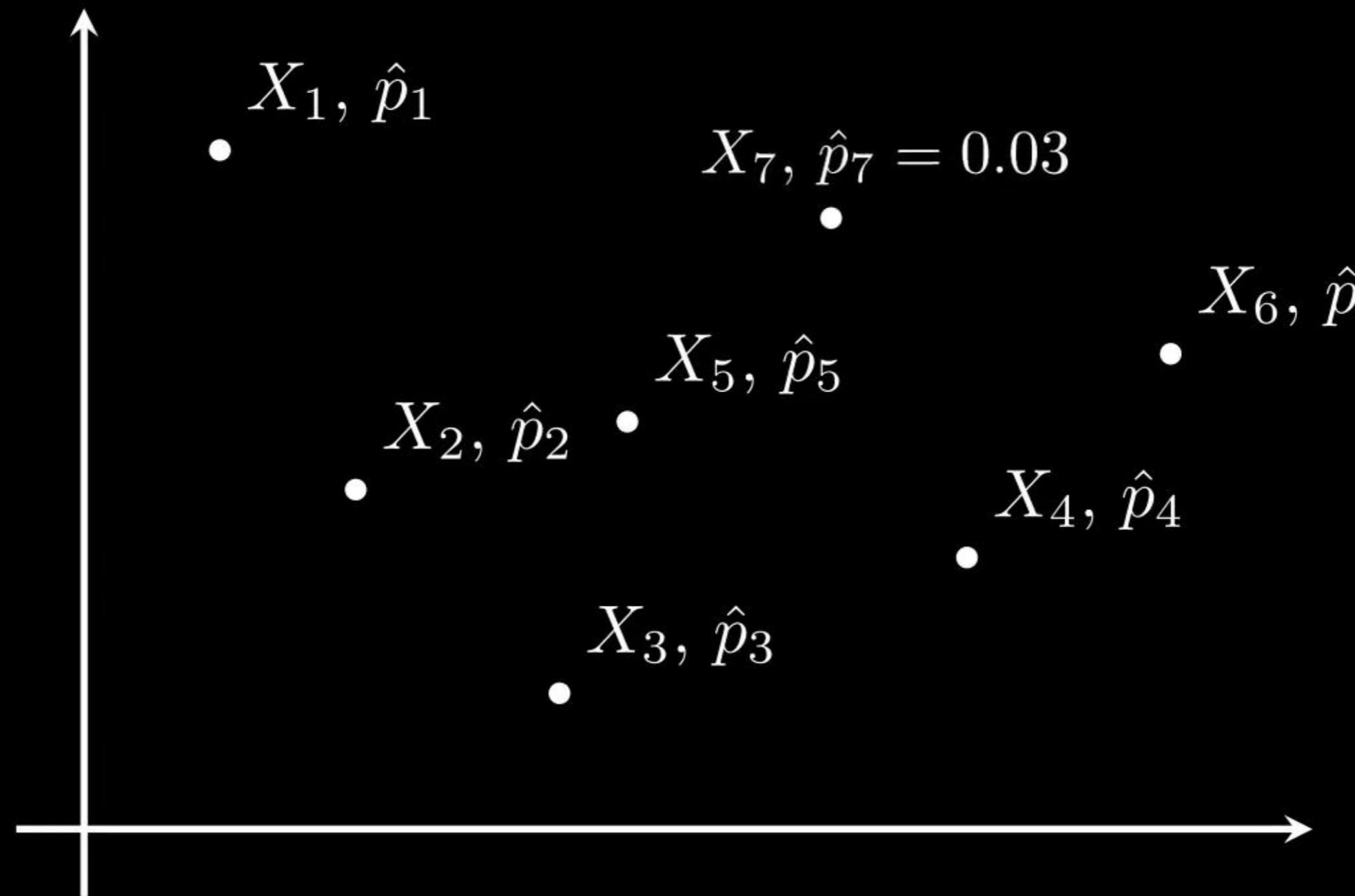
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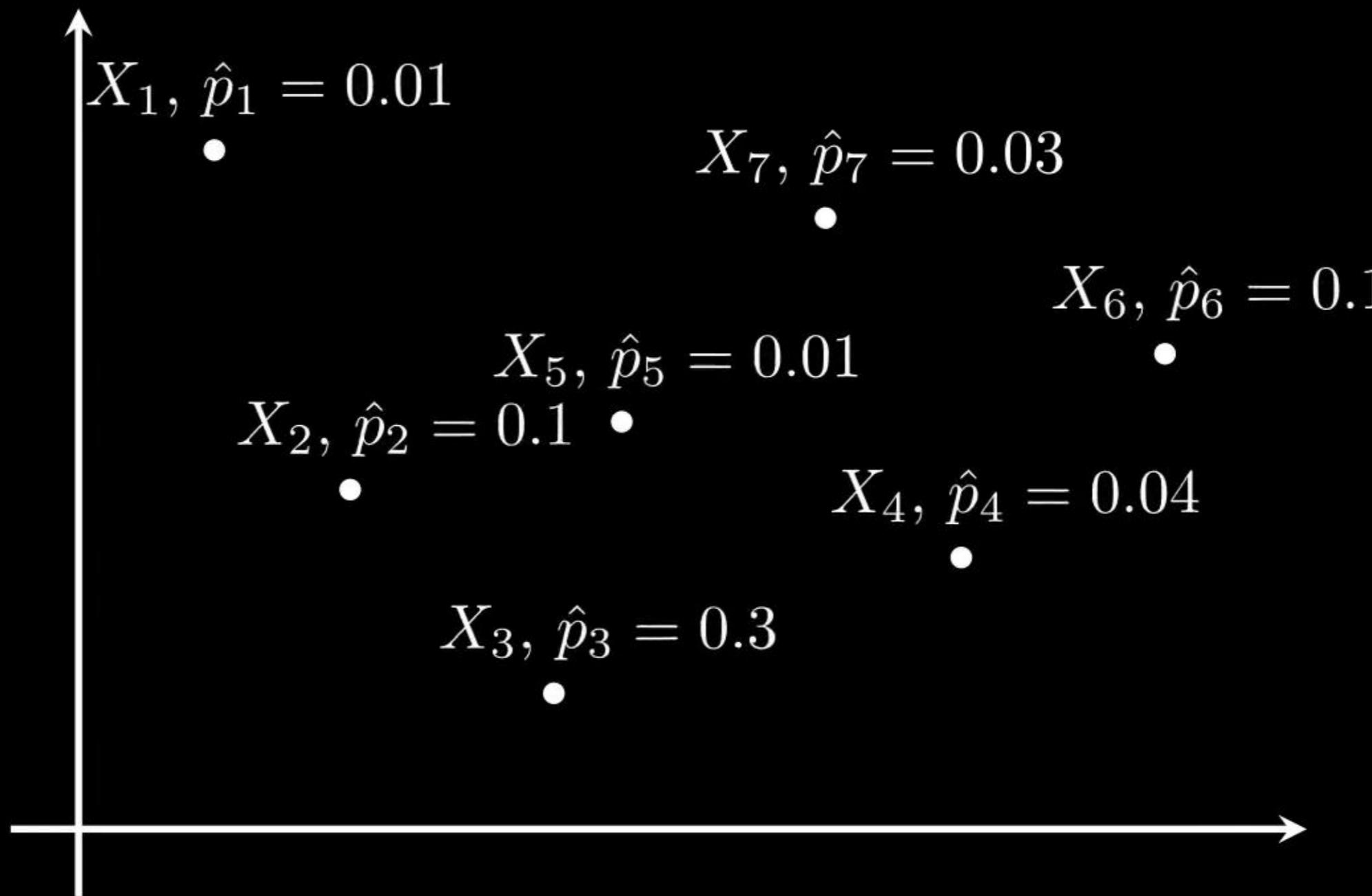
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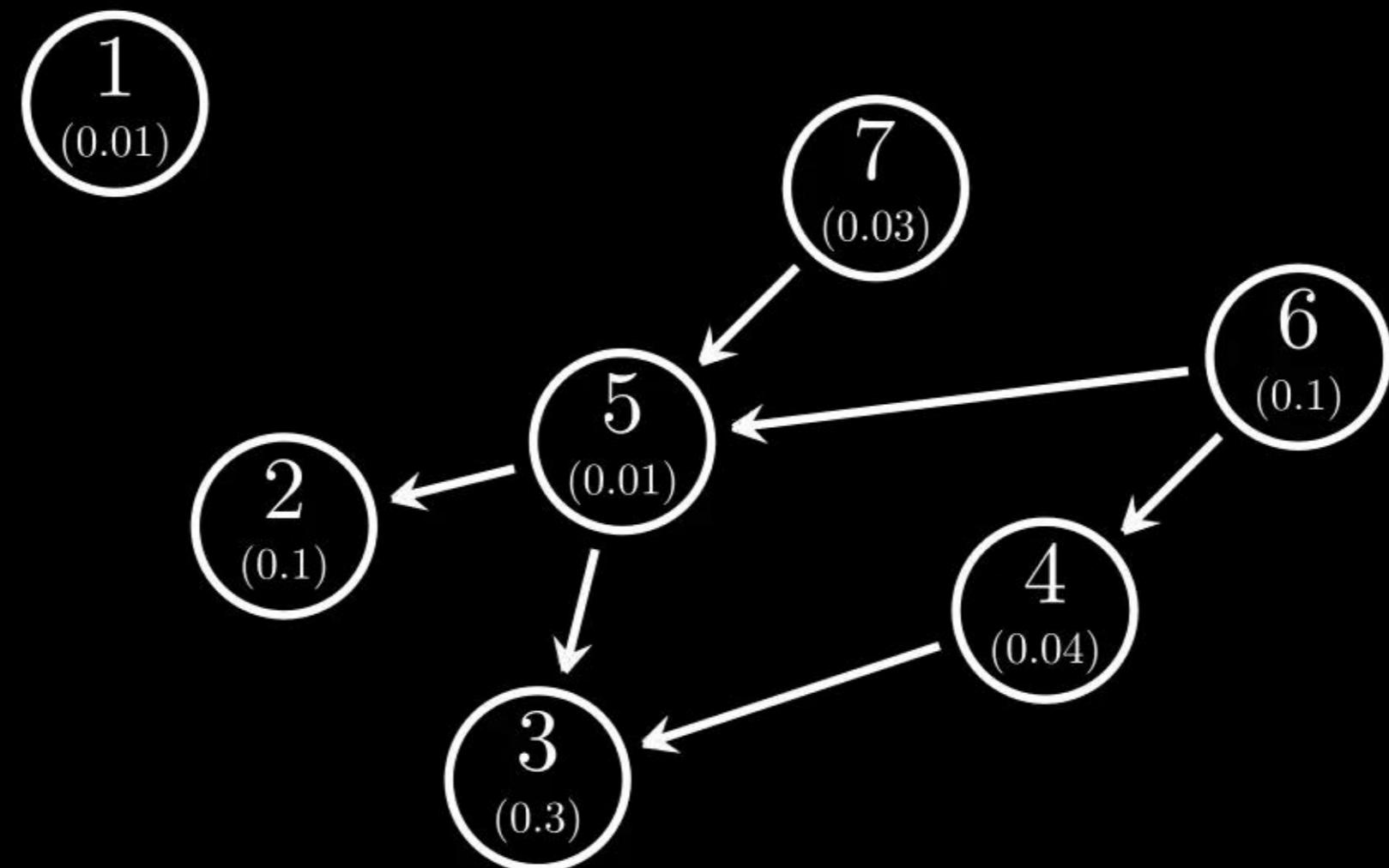
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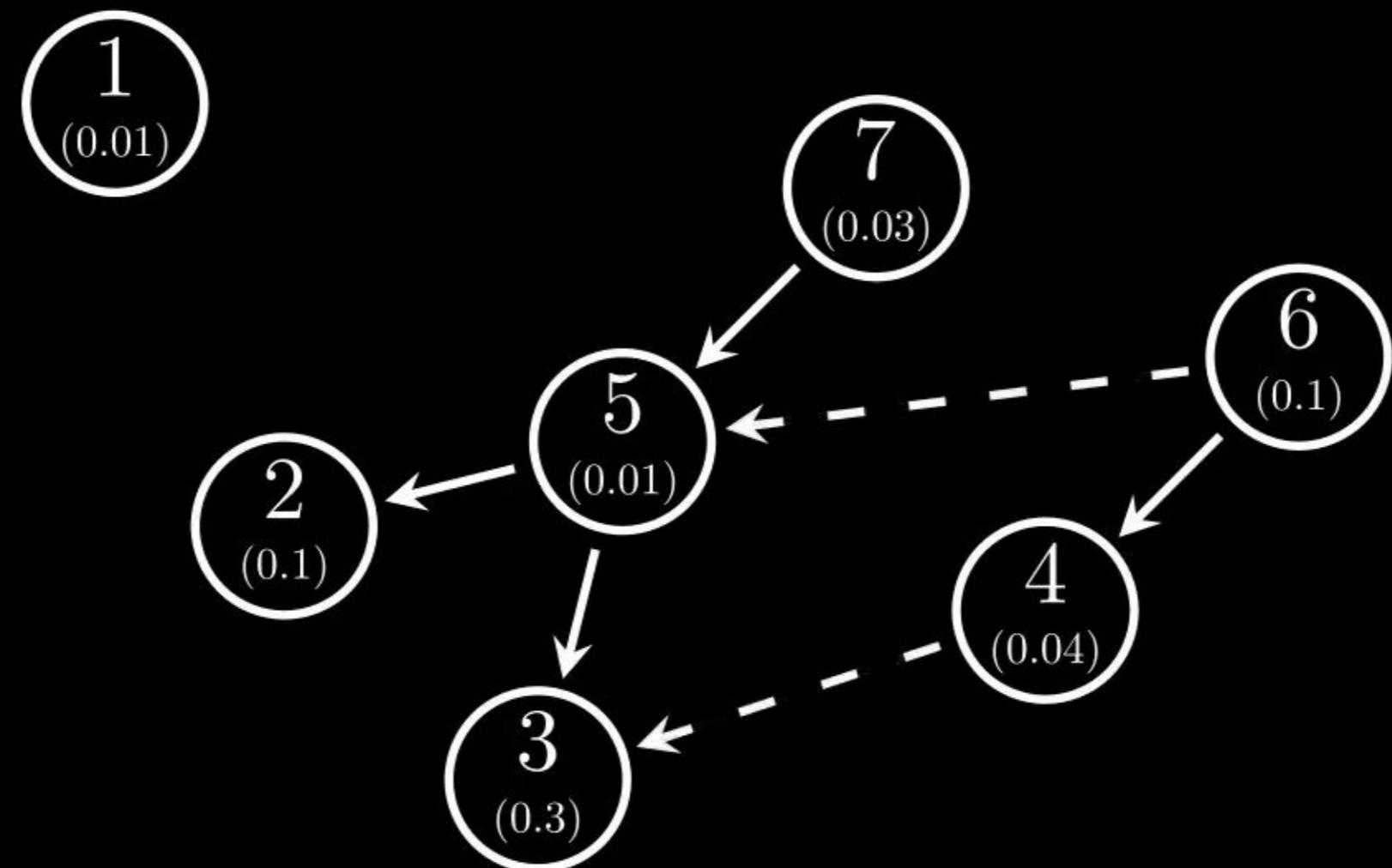
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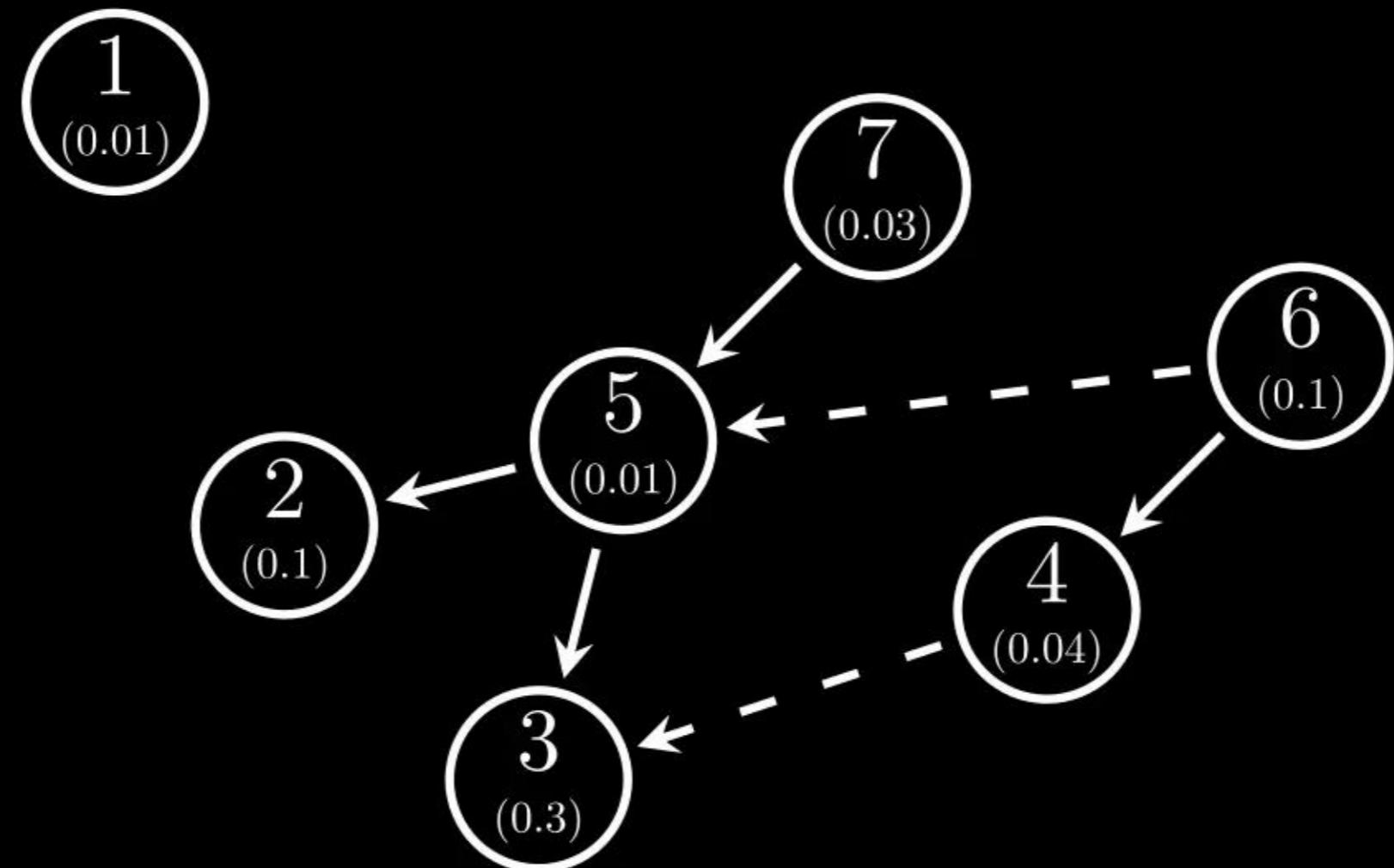
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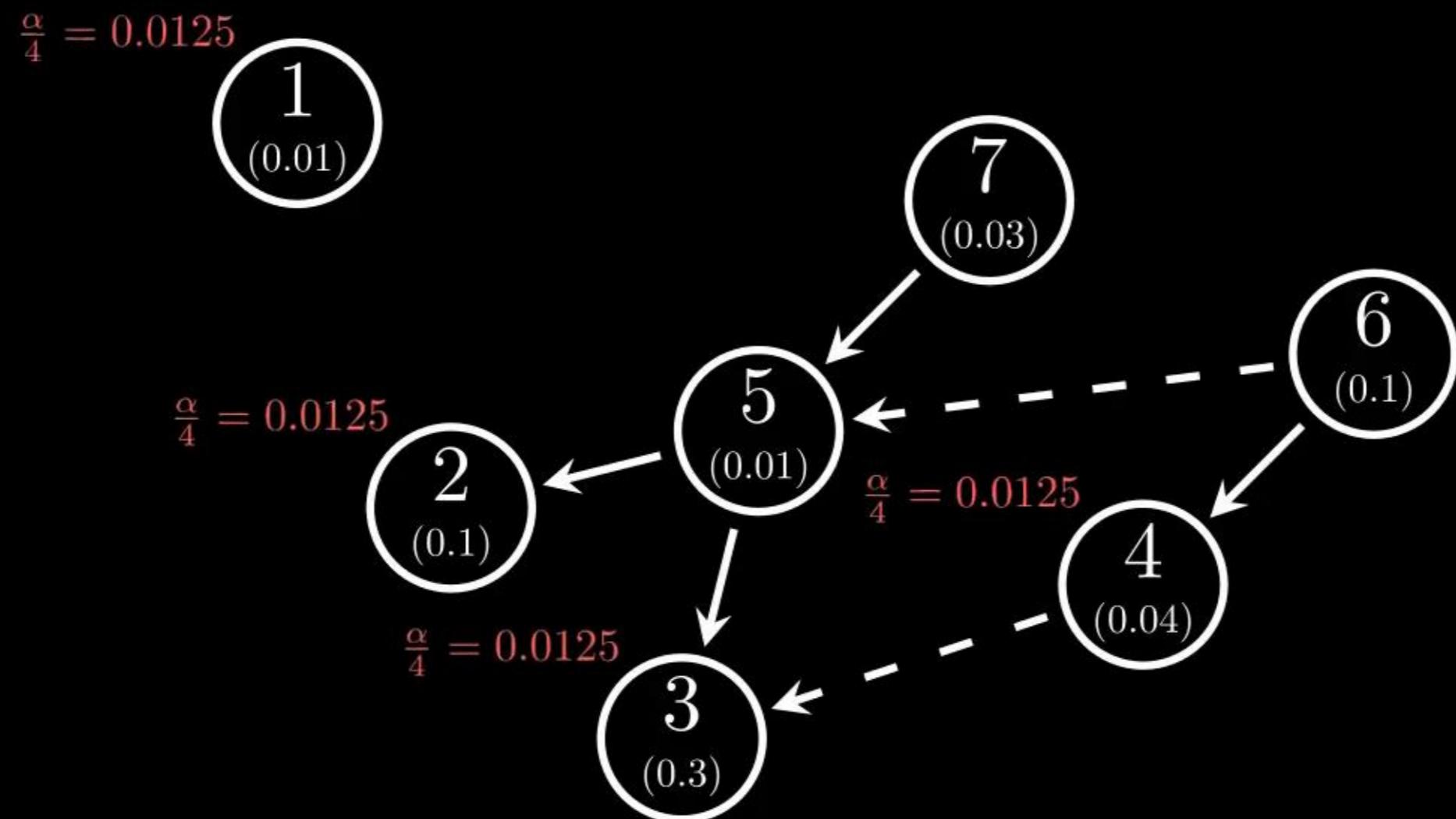
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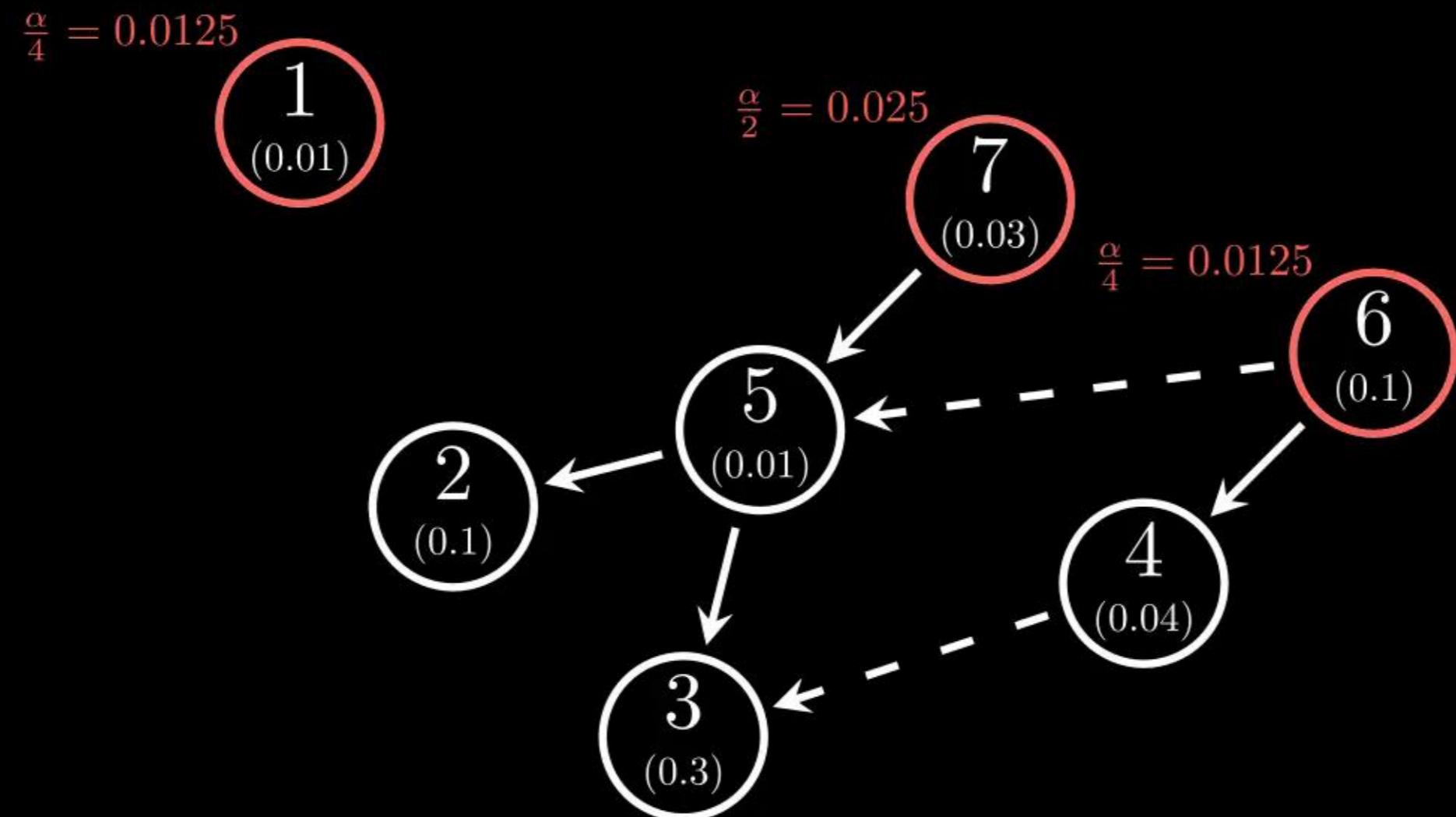
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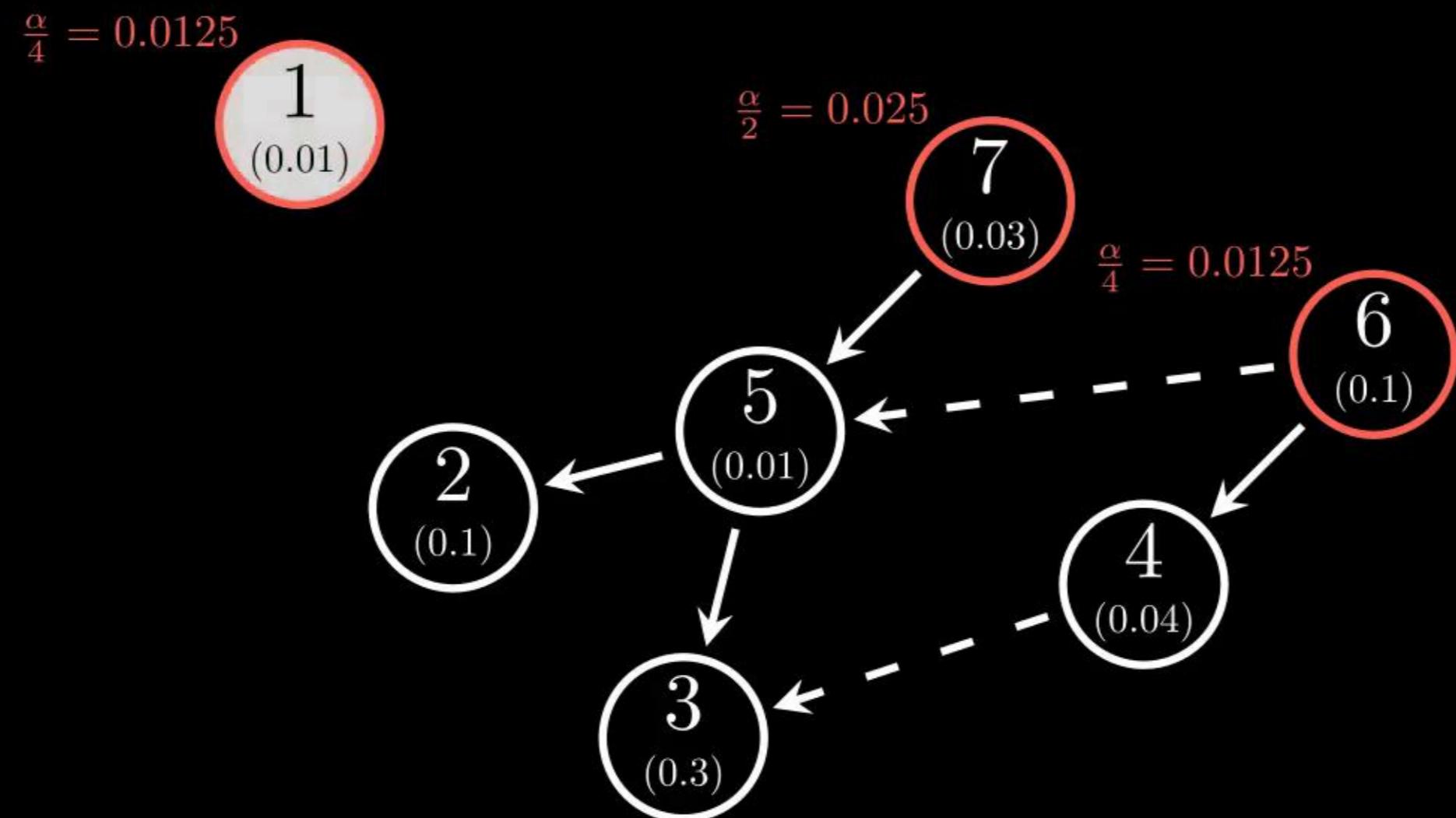
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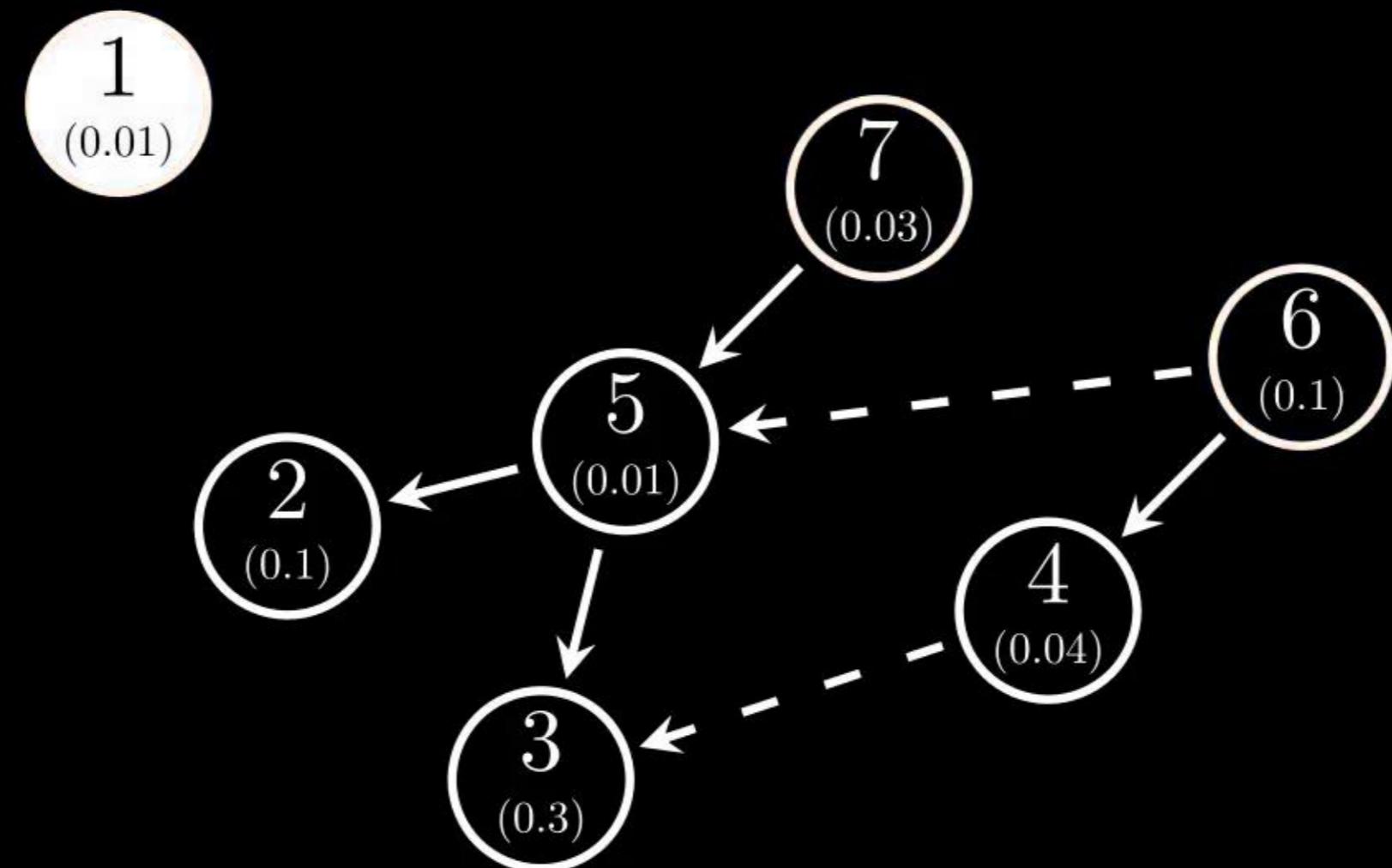
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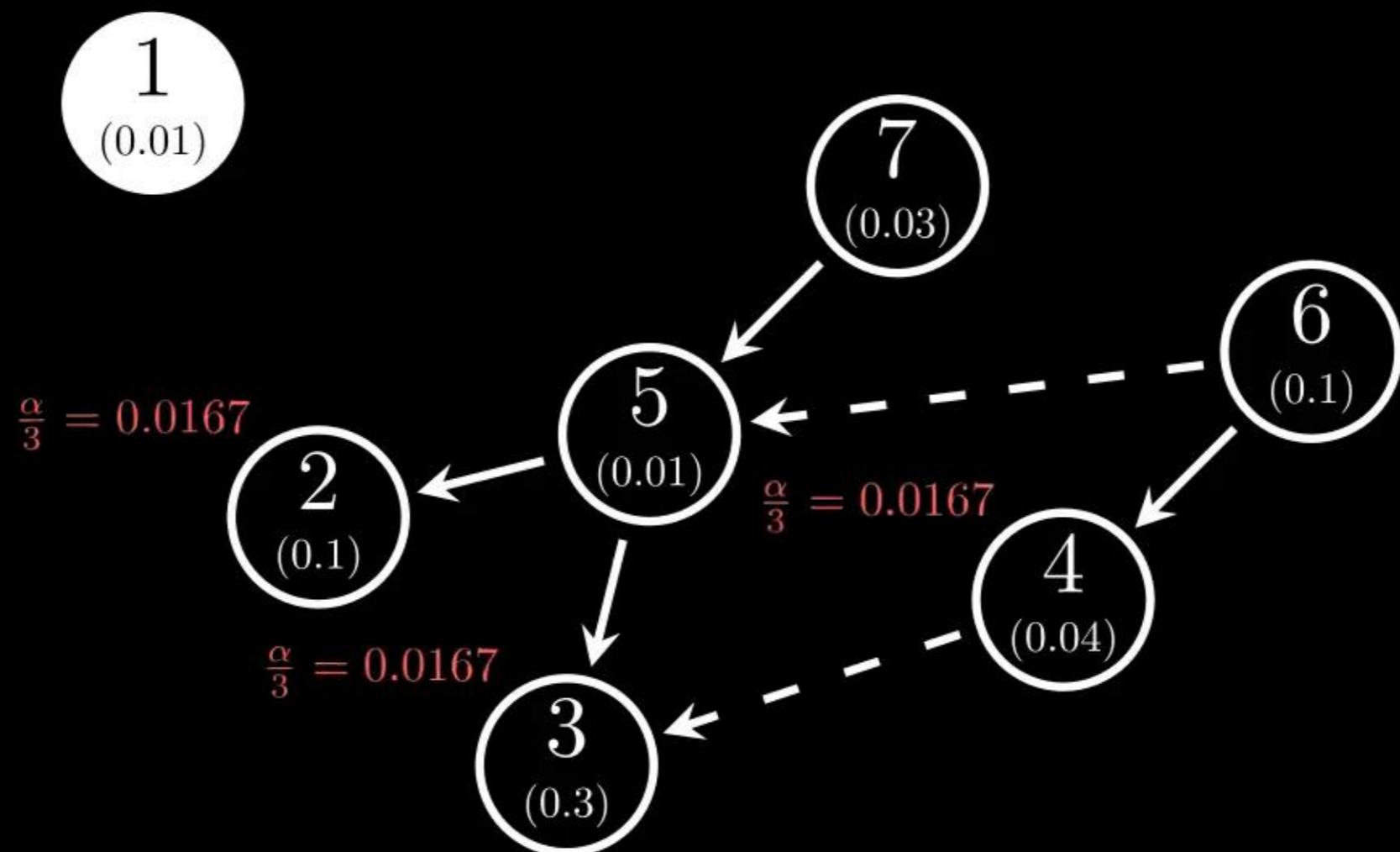
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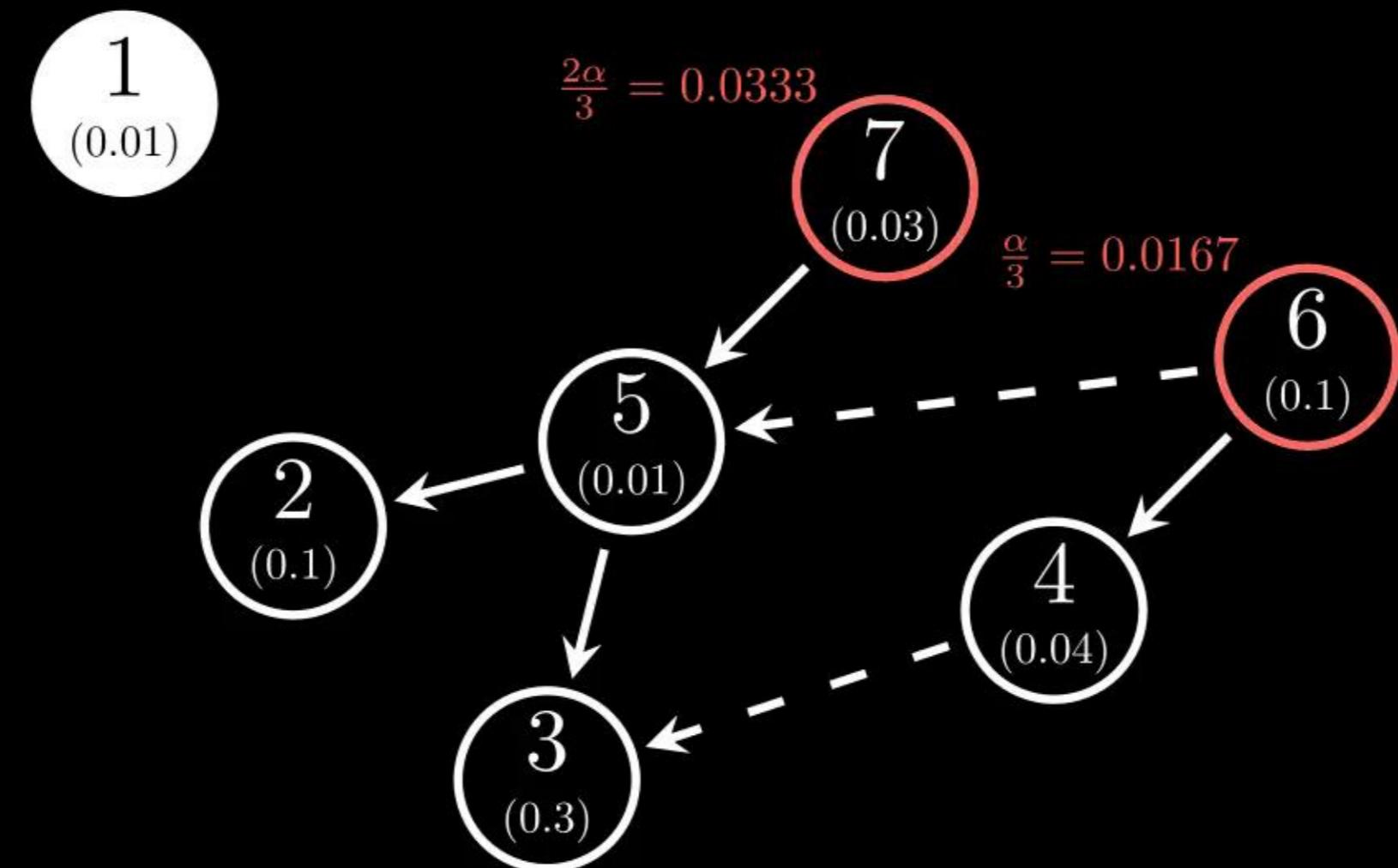
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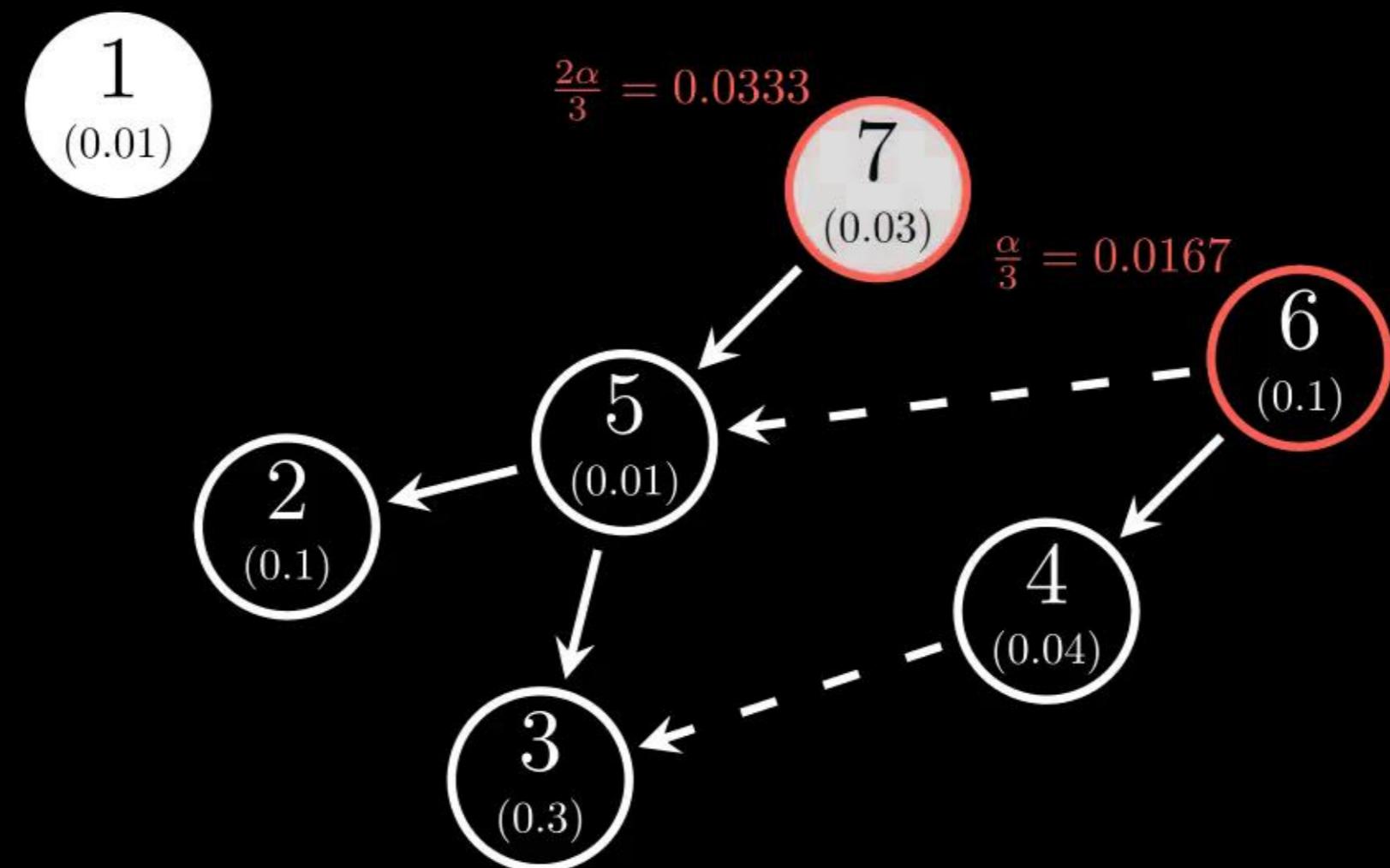
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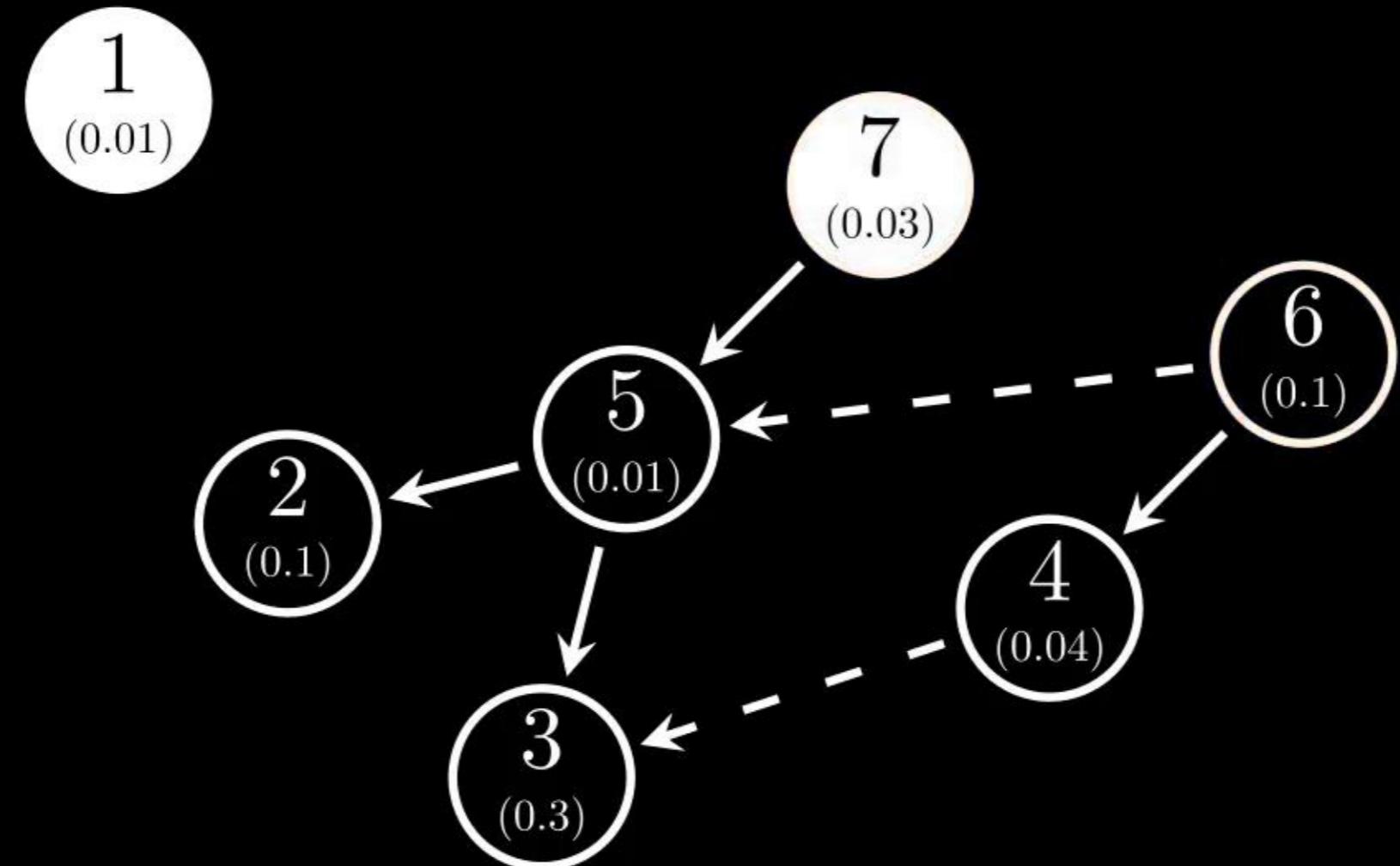
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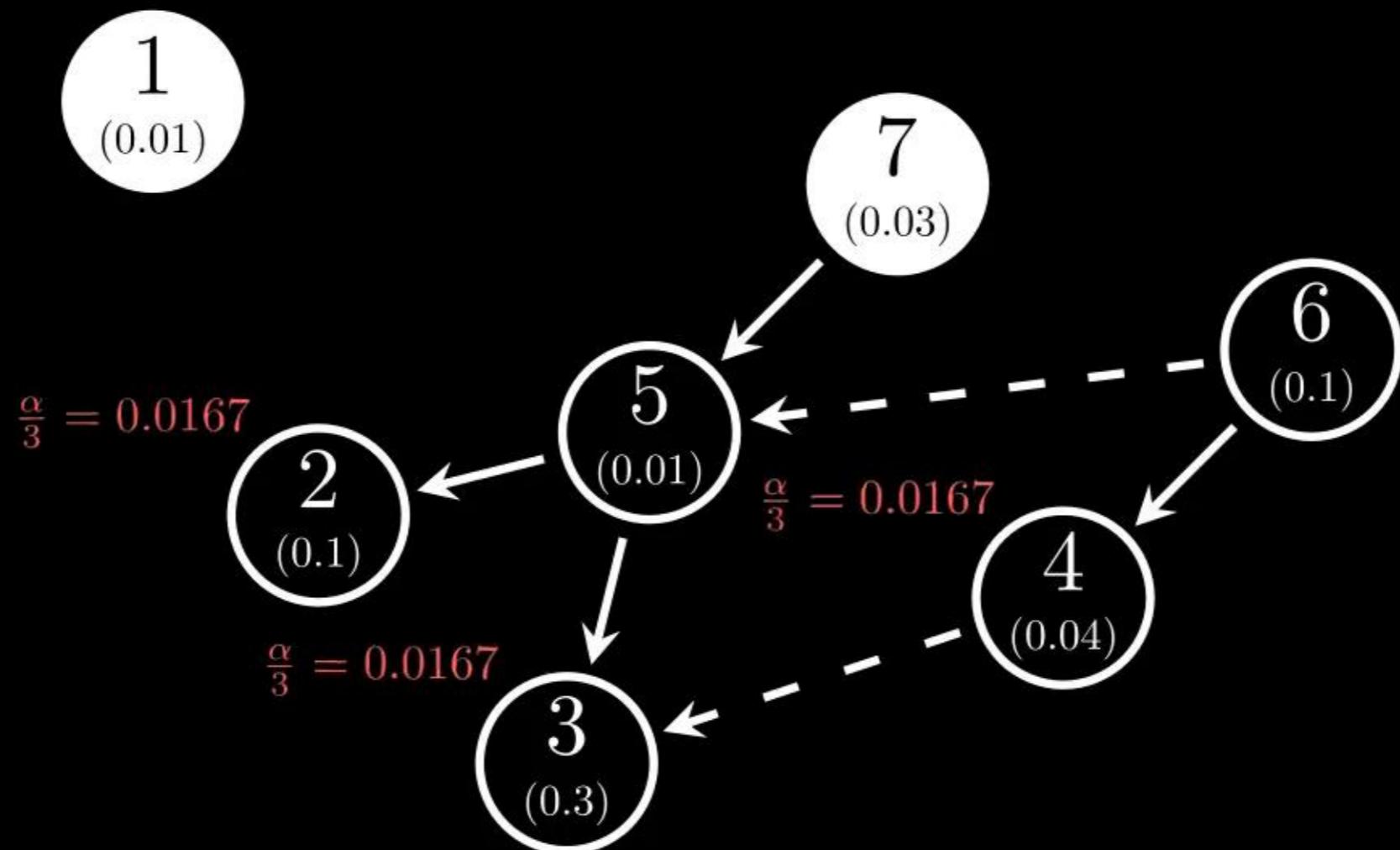
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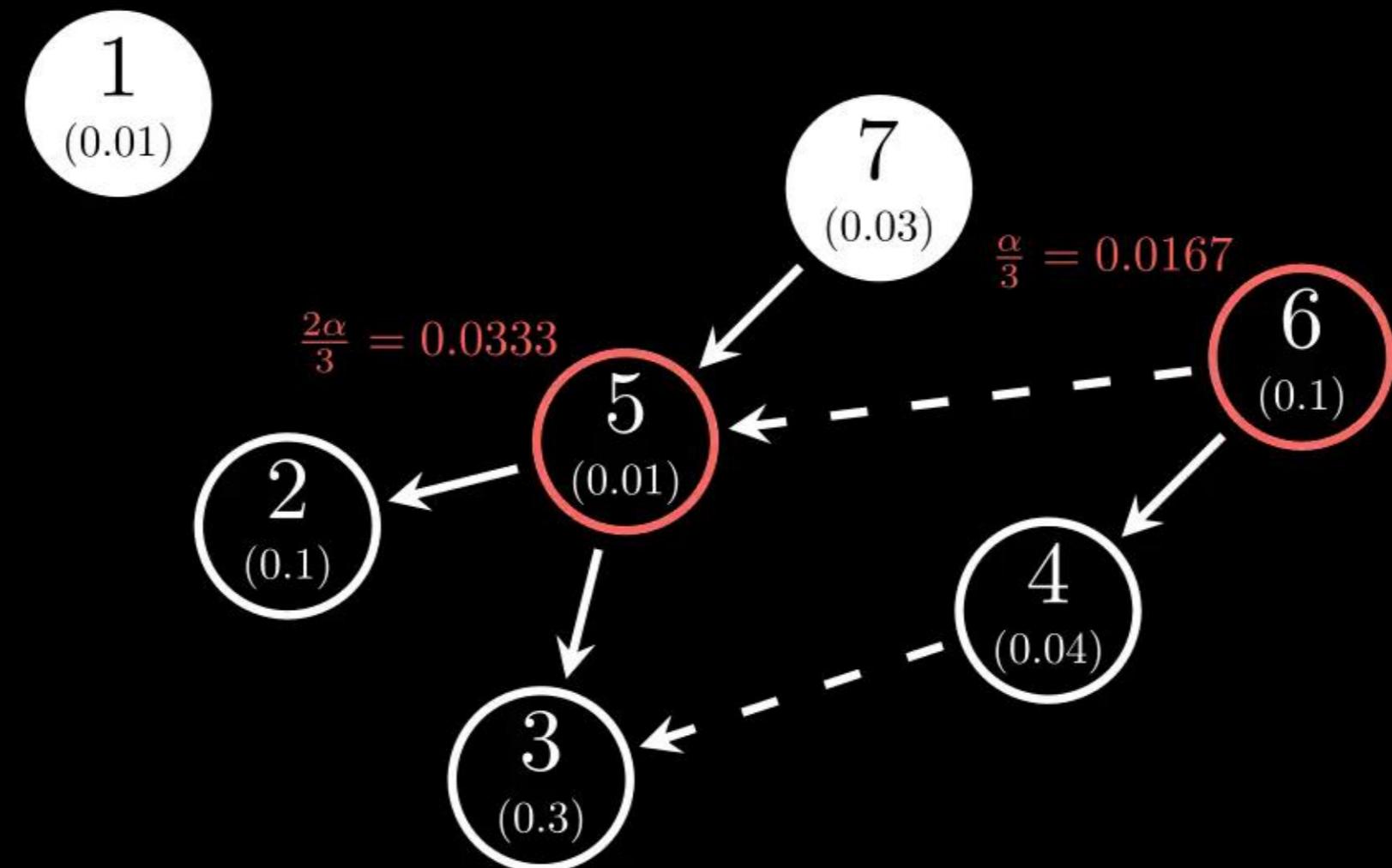
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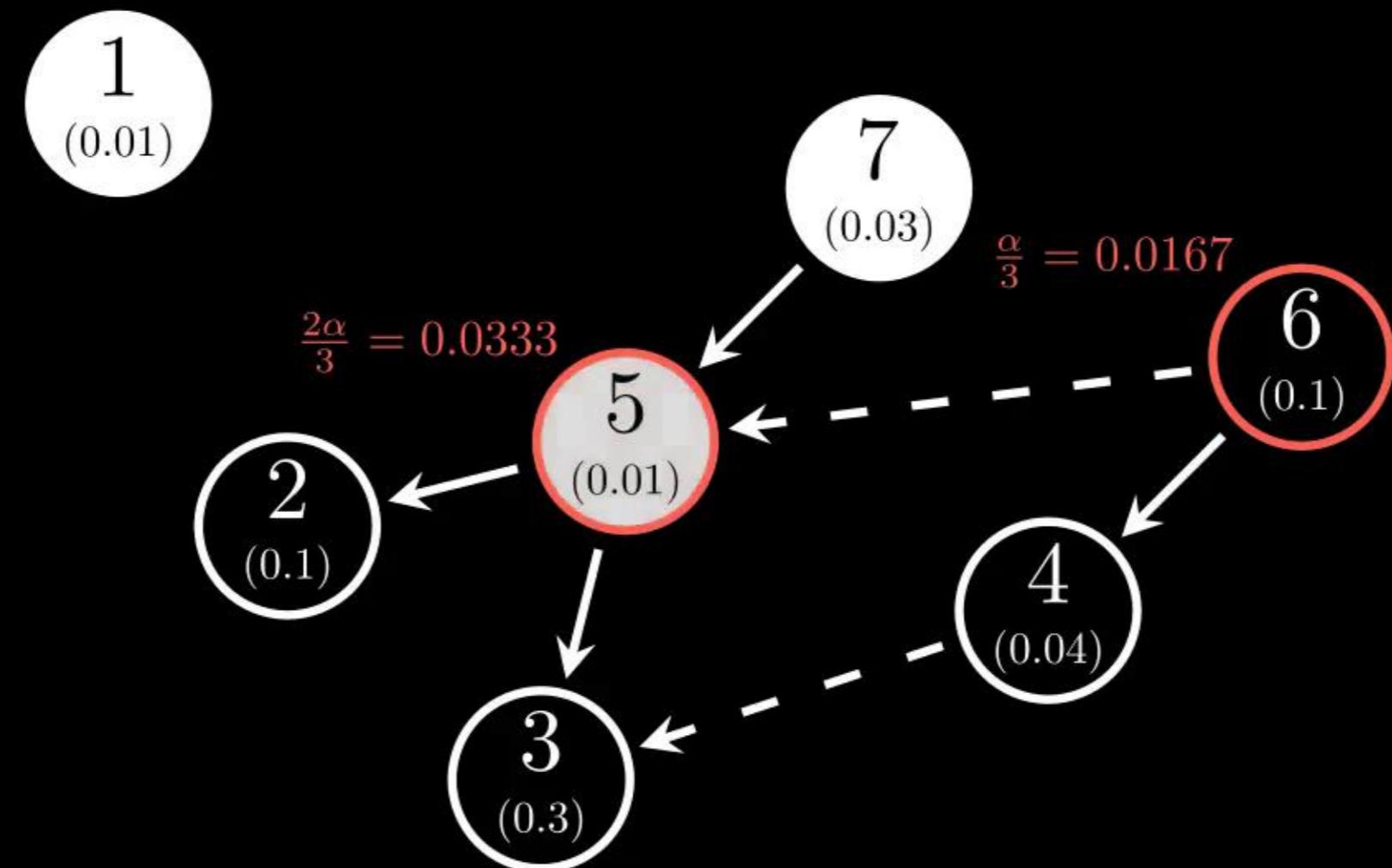
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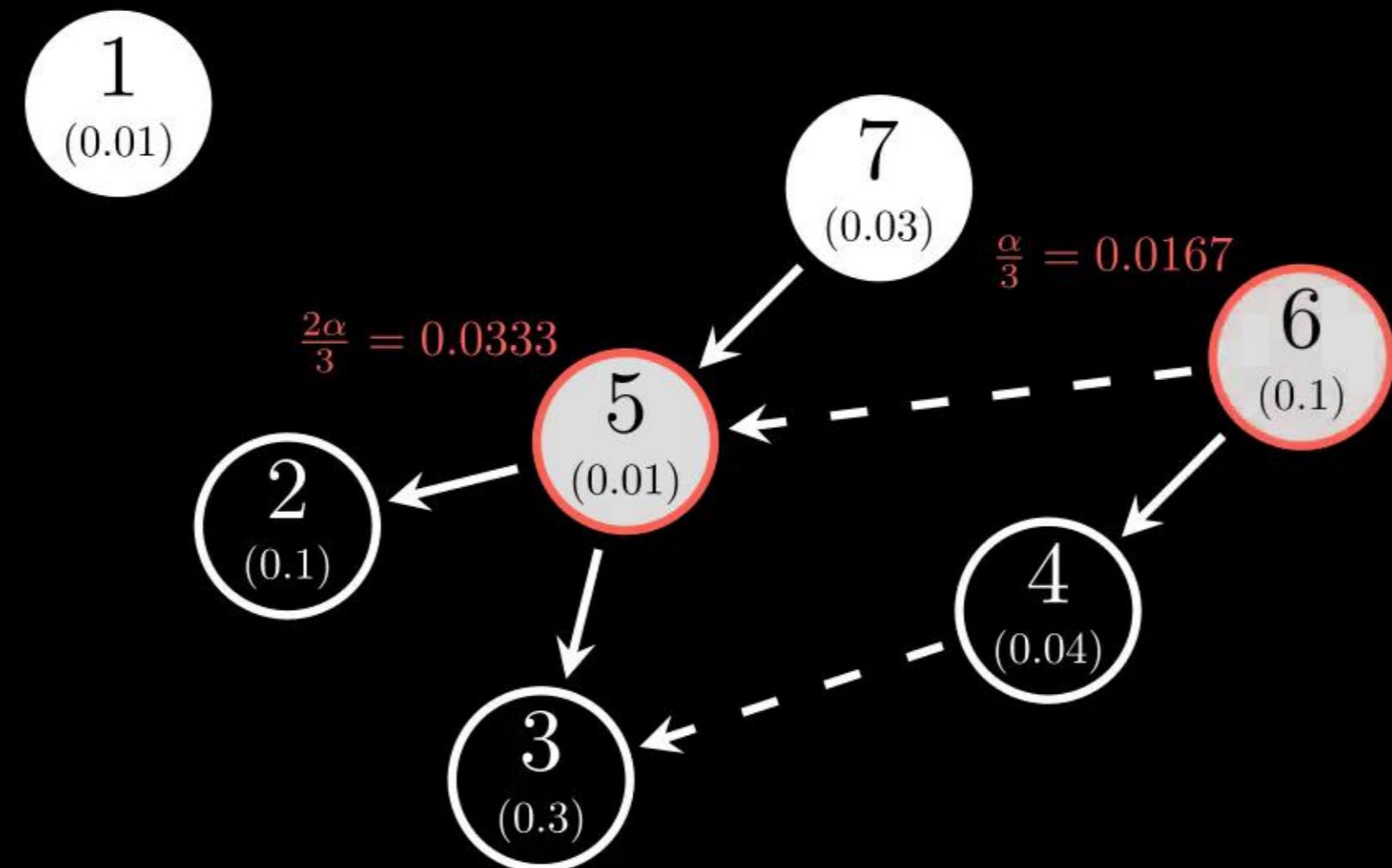
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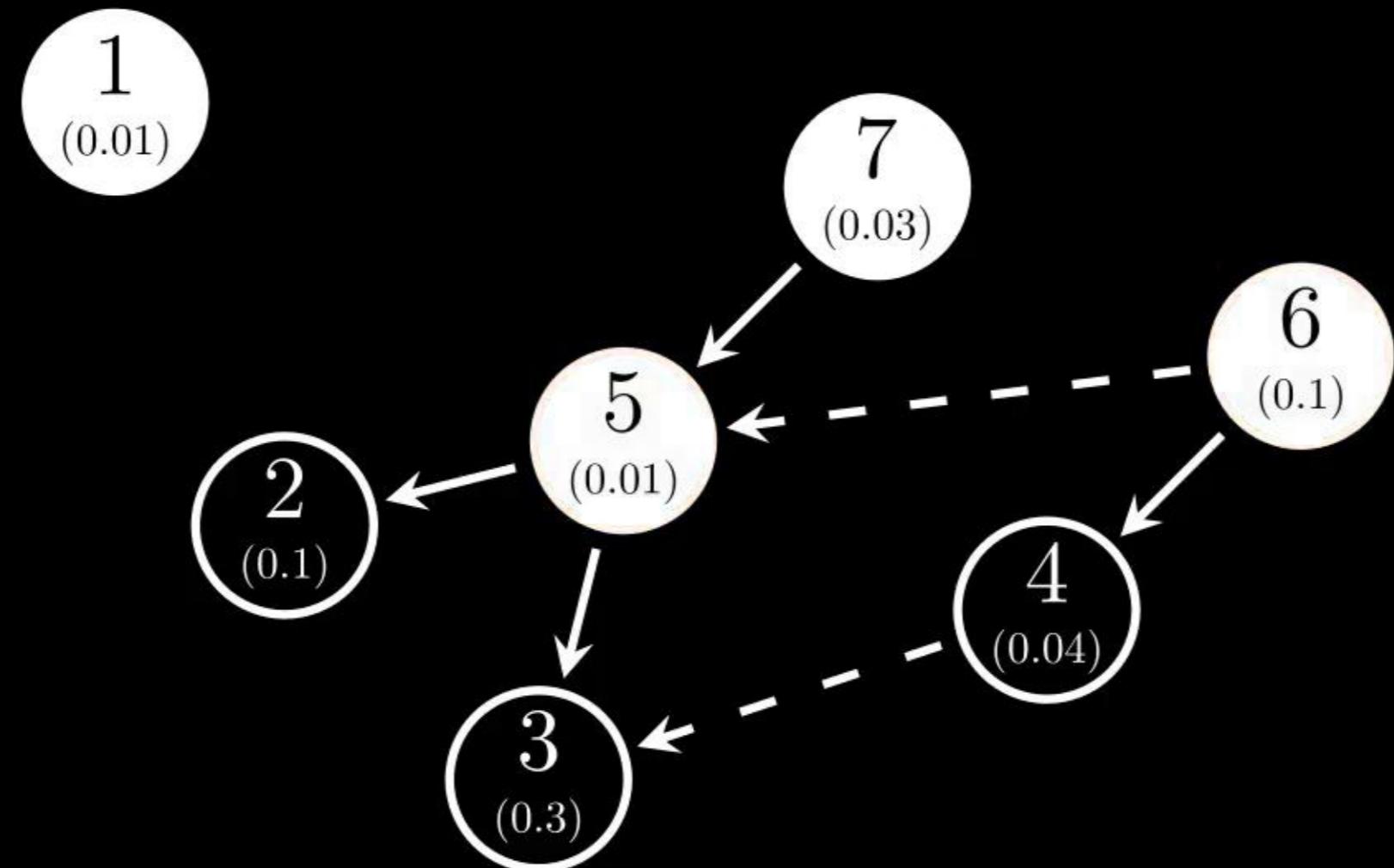
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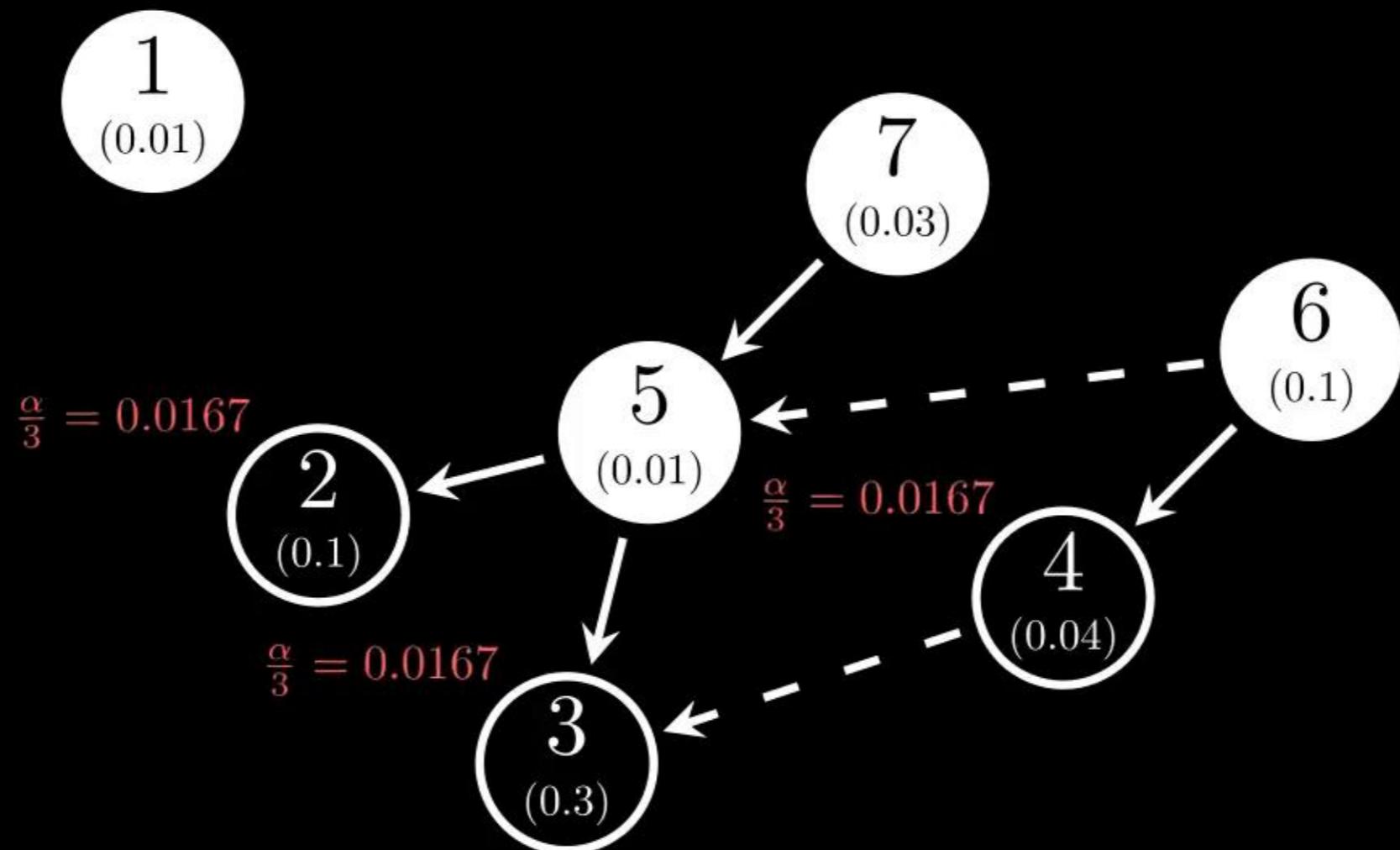
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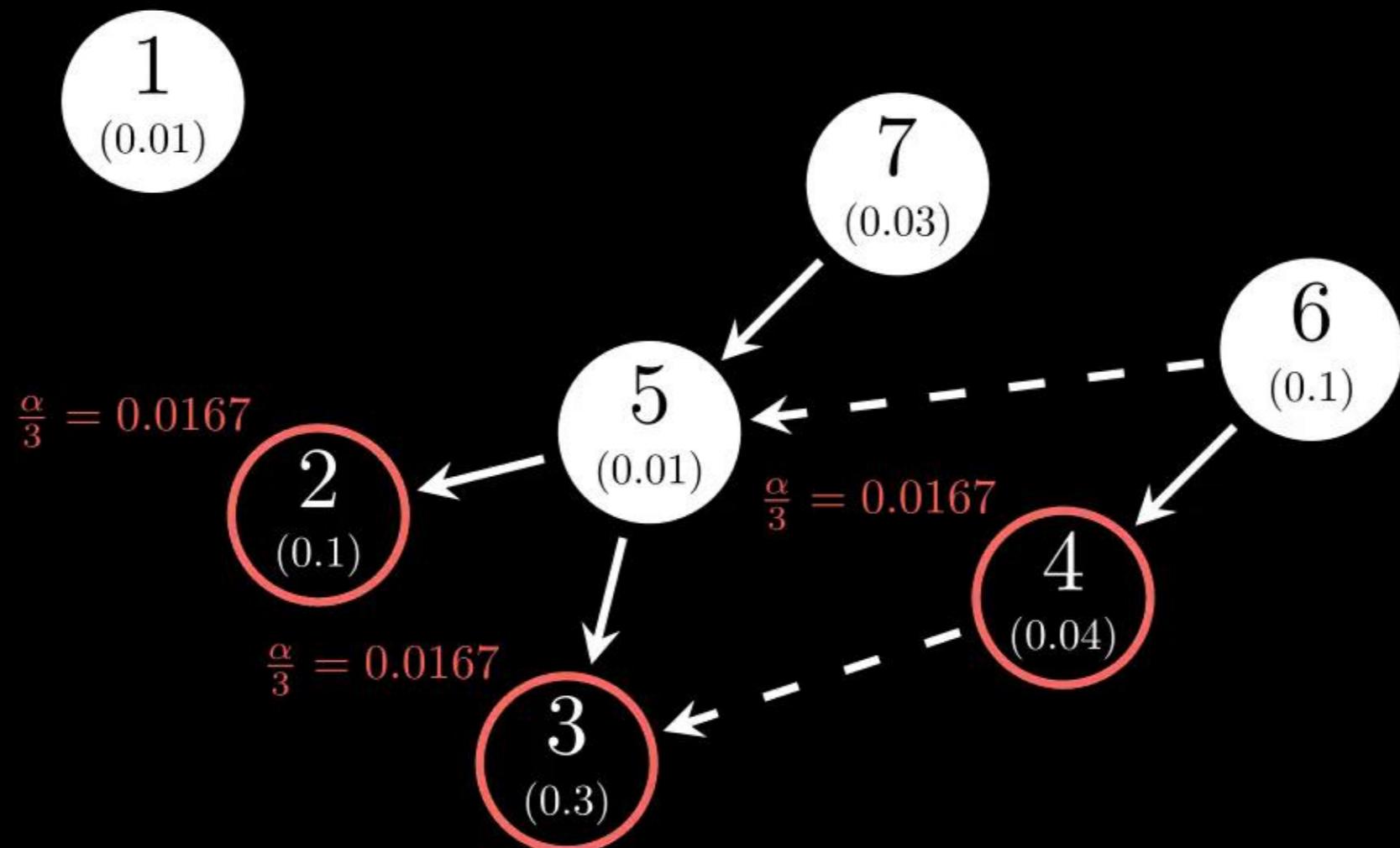
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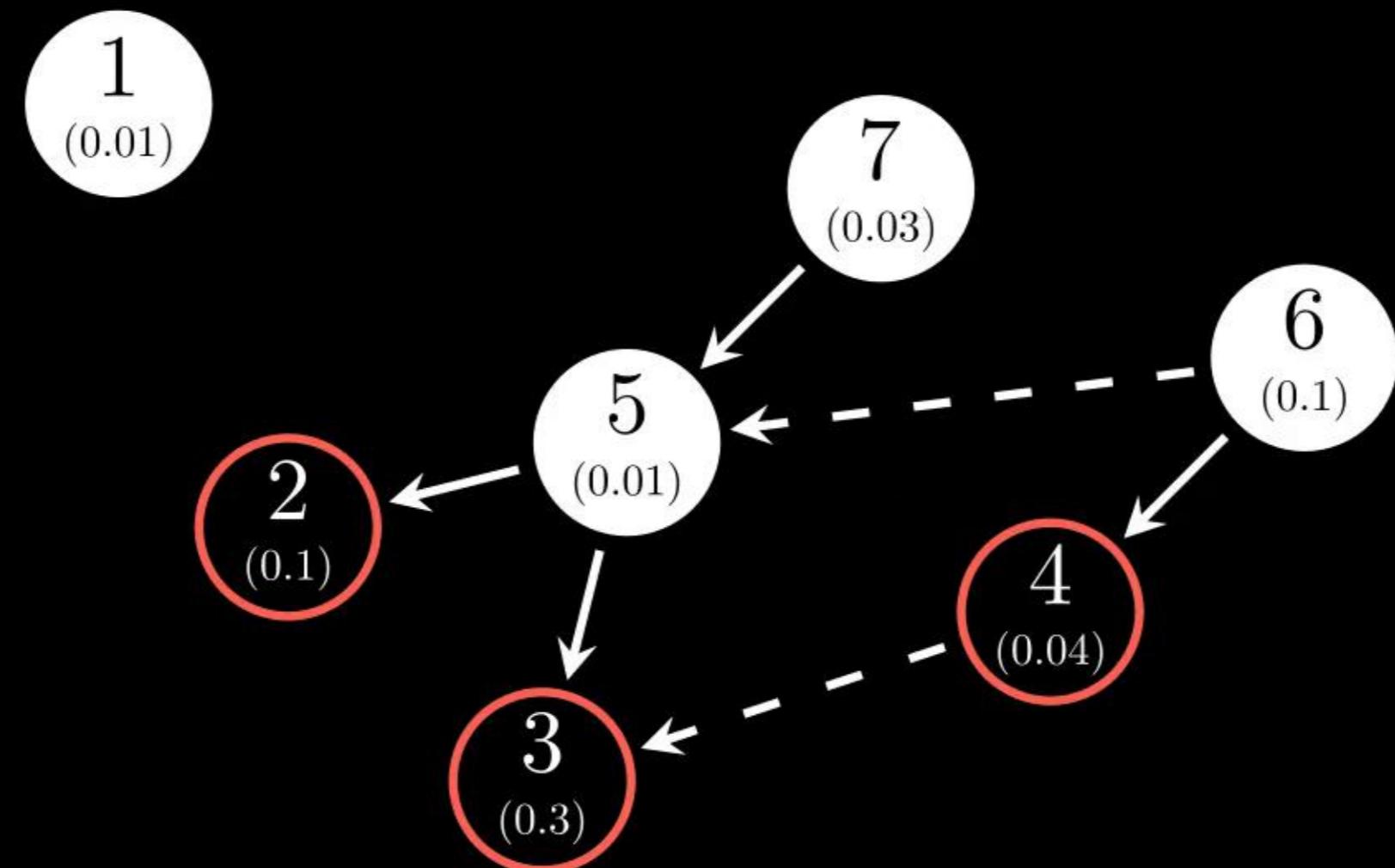
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Here: $\alpha = 0.05$. The procedure terminates with $\mathcal{R}_\alpha = \{1, 5, 6, 7\}$.

Further conditions are needed for power

Let $\hat{\mathcal{A}}_n(\tau, \alpha, \mathcal{P})$ denote the set of *data-dependent selection sets* controlling Type I error over \mathcal{P} . Recall $R_\tau(\hat{A}) := \mathbb{E}\{\mu(\mathcal{X}_\tau(\eta) \setminus \hat{A})\}$.

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Proposition. For any $n \in \mathbb{N}$, $\alpha \in (0, 1)$, $\sigma > 0$,

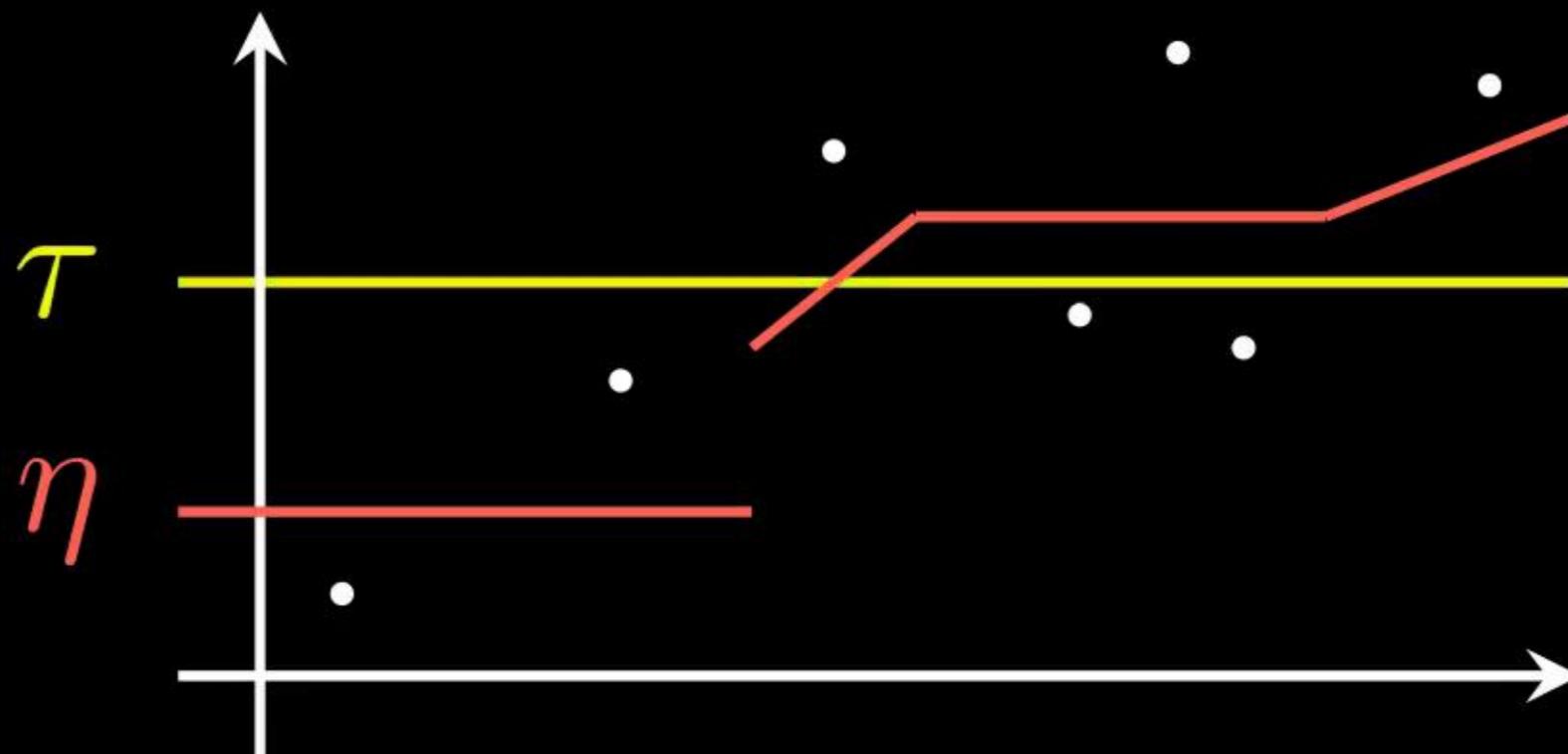
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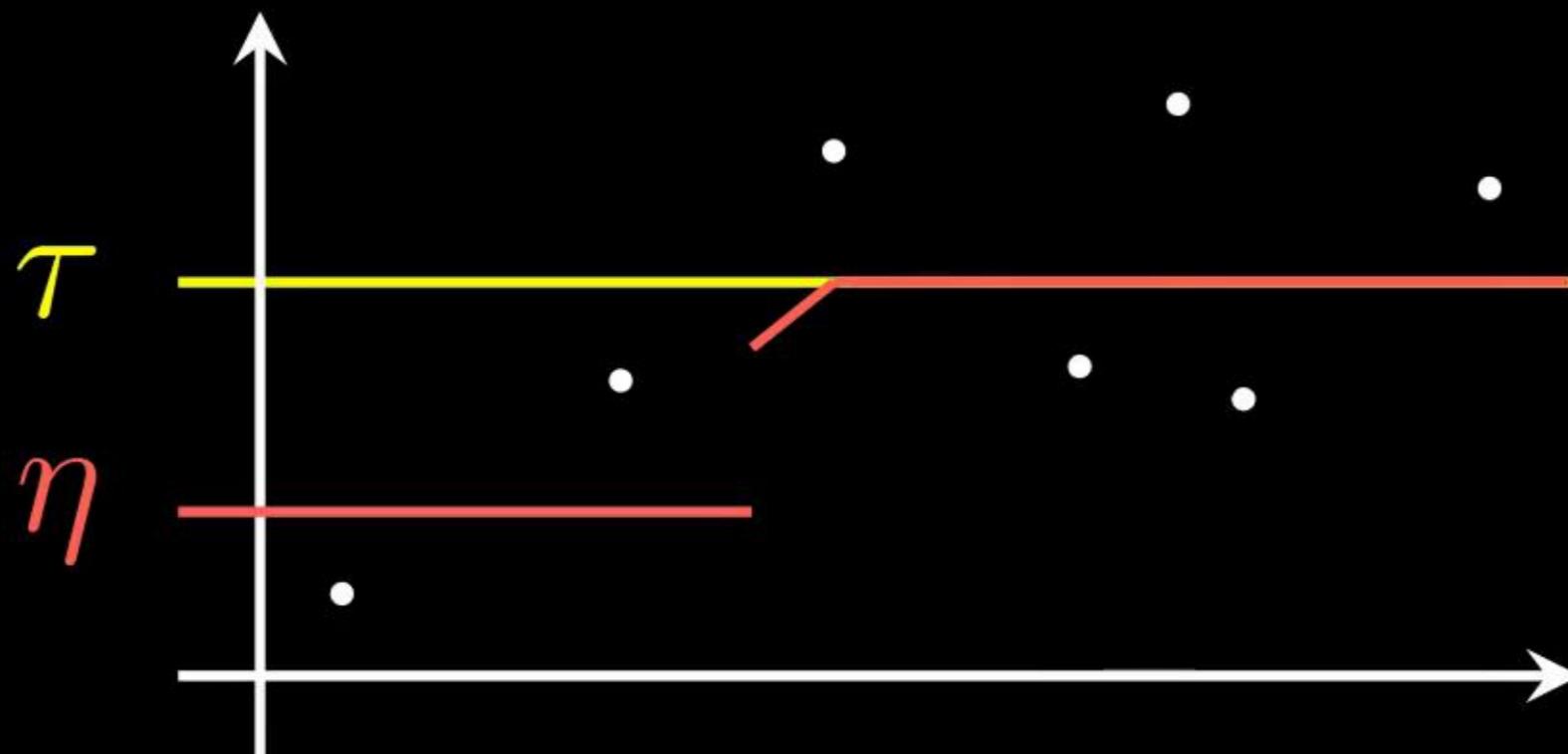


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Definition. For $\sigma, \gamma, \lambda > 0$ and $\theta > 1$, let $\mathcal{P}_{\text{MonReg},d}(\sigma, \tau, \gamma, \lambda, \theta) \subseteq \mathcal{P}_{\text{Mon},d}(\sigma)$ denote the class of distributions $P \in \mathcal{P}_{\text{Mon},d}(\sigma)$ for which additionally

1. $\eta(x + r(1, \dots, 1)^\top) \geq \tau + \lambda \cdot r^\gamma$ for all $x \in \mathcal{X}_\tau(\eta) \cap \text{supp}(\mu)$ and $r \in (0, 1]$;
2. $\theta^{-1} \cdot r^d \leq \mu(\{y \in \mathbb{R}^d : \|y - x\|_\infty \leq r\}) \leq \theta \cdot (2r)^d$ for all $x \in \mathcal{X}_\tau(\eta) \cap \text{supp}(\mu)$, $r \in (0, 1]$.

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Theorem. Let $\sigma, \gamma, \lambda > 0$ and $\theta > 1$. There exists $C \geq 1$, depending only on (d, θ) , such that for any $n \geq 1$ and $\alpha \in (0, 1)$,

$$\sup_{P \in \mathcal{P}_{\text{MonReg},d}(\sigma, \tau, \gamma, \lambda, \theta)} R_\tau(\hat{A}^{\text{ISS}}) \leq 1 \wedge C \left\{ \left(\frac{\sigma^2}{n\lambda^2} \log_+ \left(\frac{n \log_+ n}{\alpha} \right) \right)^{1/(2\gamma+d)} + \left(\frac{\log_+ n}{n} \right)^{1/d} \right\},$$

where $\log_+ x := \log(x \vee e)$ for $x \in \mathbb{R}$.

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Hence, if we compute $\mathcal{D} = ((X_1, Y_1), \dots, (X_n, Y_n)) \sim P^n$, then $\mathbb{P}(\hat{A}^{\text{ISS}}(\mathcal{D}) \subseteq \mathcal{X}_\tau(\eta)) \geq 1 - \alpha$ whenever $P \in \mathcal{P}_{\text{Mon}, d}(\sigma)$.

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Definition. Let $\hat{\sigma}_{0,k}^2 := \frac{1}{k} \sum_{j=1}^k (Y_{(j)} - \tau)_+^2$ and $\bar{Y}_{1,k} := \frac{1}{k} \sum_{j=1}^k Y_{(j)}$ for $k \in [n(x)]$ and $\hat{\sigma}_{1,k}^2 := \frac{1}{k} \sum_{j=1}^k (Y_{(j)} - \bar{Y}_{1,k})^2$ for $k \in \{2, \dots, n(x)\}$. Denote $\bar{Y}_{1,0} := 0$, and $\hat{\sigma}_{1,k}^2 := 1$ for $k \in \{0, 1\}$. For $k \in [n(x)]$, define

$$\bar{p}_\tau^k(x) := \frac{1}{\hat{\sigma}_{0,k}^k e^{k/2}} \cdot \prod_{j=1}^k \hat{\sigma}_{1,j-1} \exp \left\{ \frac{(Y_{(j)} - \bar{Y}_{1,j-1})^2}{2\hat{\sigma}_{1,j-1}^2} \right\},$$

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Lemma. When $\eta(x) < \tau$, we have $\mathbb{P}\{\bar{p}_\tau(x) \leq t | \mathcal{D}_X\} \leq t$ for all $t \in (0, 1)$.

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where for $z \in [0, 1]$ and $a, b > 0$, we write $\text{B}(z; a, b) := \int_0^z t^{a-1} (1-t)^{b-1} dt$ for the *incomplete beta function*.

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Lemma. If $\tau \in [0, 1)$, $\eta(x) < \tau$, then $\mathbb{P}\{\check{p}_\tau(x) \leq t | \mathcal{D}_X\} \leq t$ for $t \in (0, 1)$.

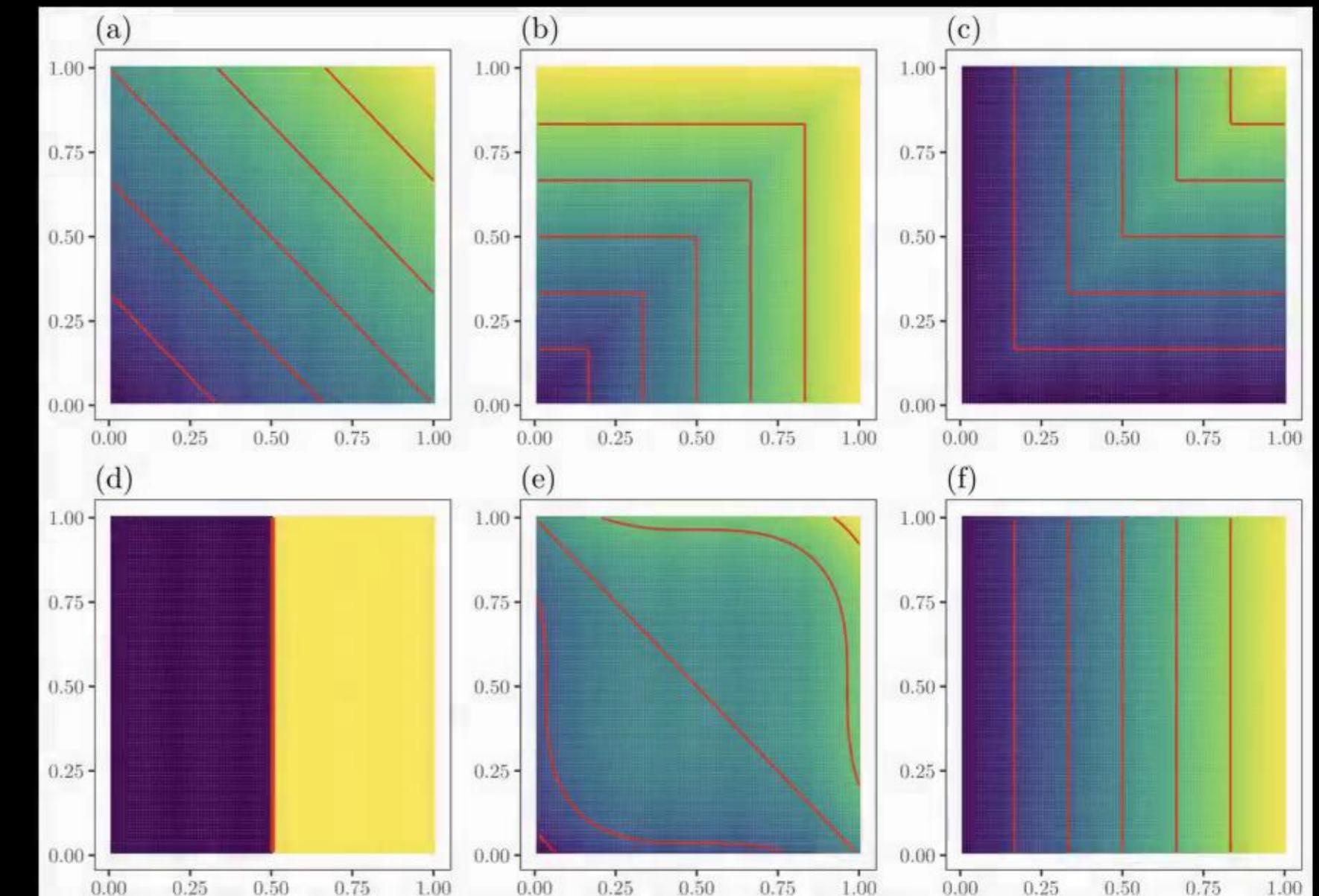
Simulations

We conduct a simulation study to compare with other choices of multiple testing procedure. We take $\mu = \text{Unif}([0, 1]^d)$, $Y - \eta(X)|X \sim \mathcal{N}(0, \sigma^2)$ and our regression functions η are obtained by rescaling f to $[0, 1]$ on $[0, 1]^d$:

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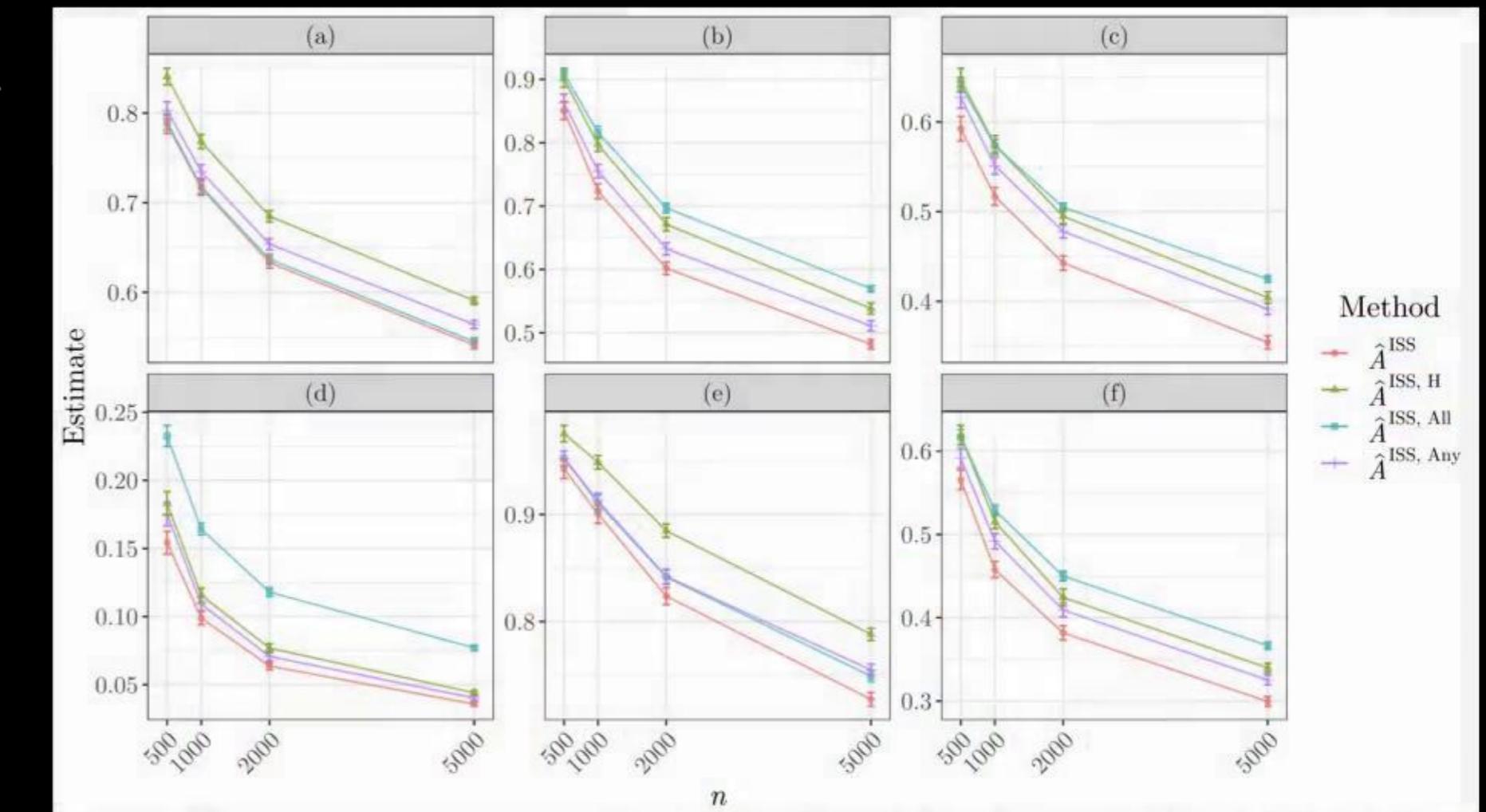
Label	Function f	τ	$\gamma(P)$
(a)	$\sum_{j=1}^d x^{(j)}$	$1/2$	1
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(c)	$\min_{1 \leq j \leq d} x^{(j)}$	$1 - 1/2^{1/d}$	1
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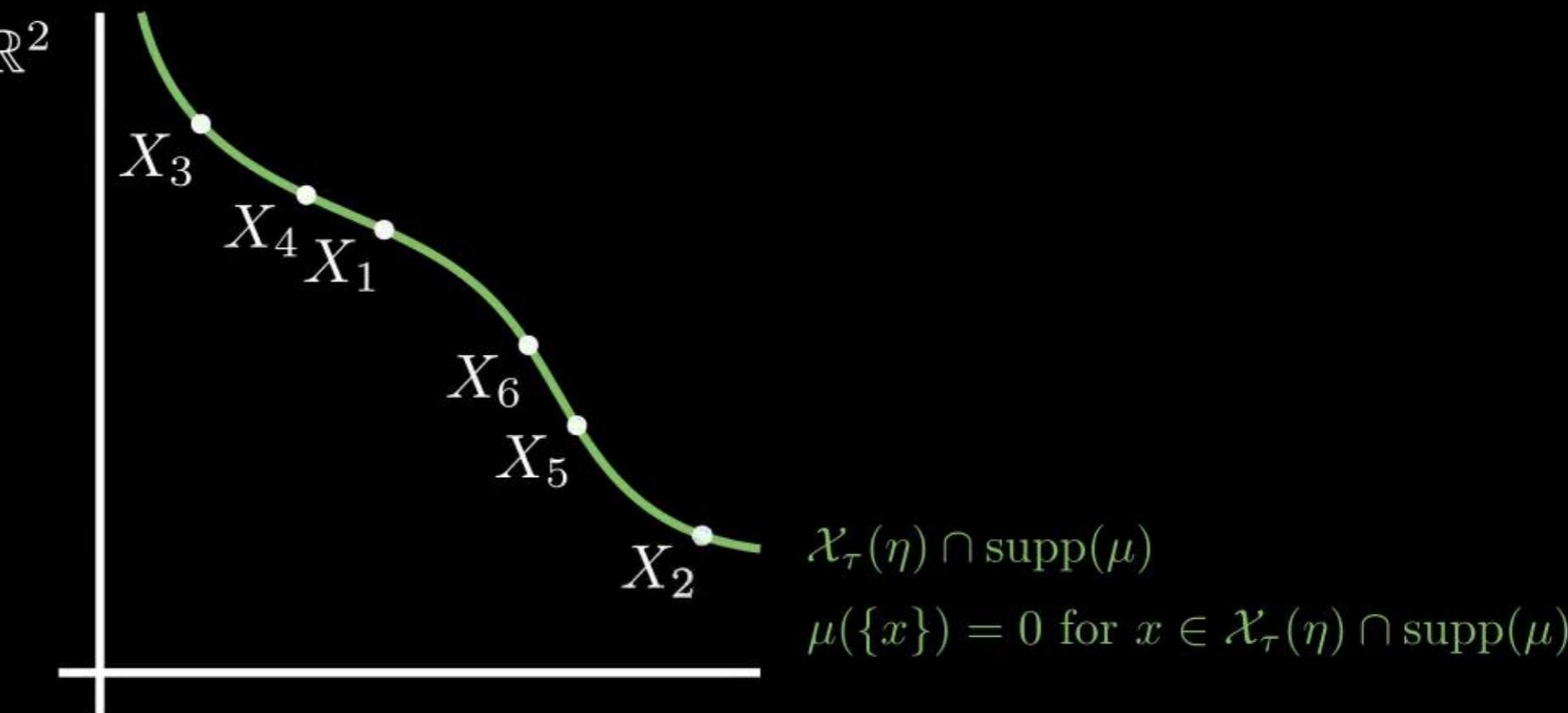
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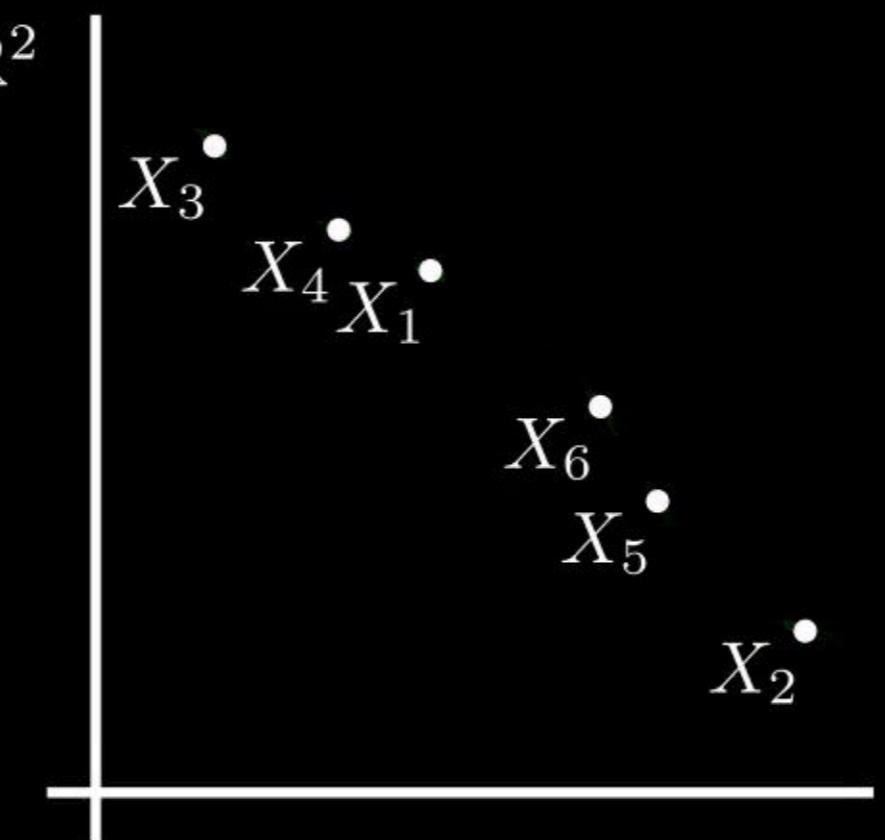
Here, $d = 2$, $\sigma = 1/4$.

See also Meijer and Goeman (2015).

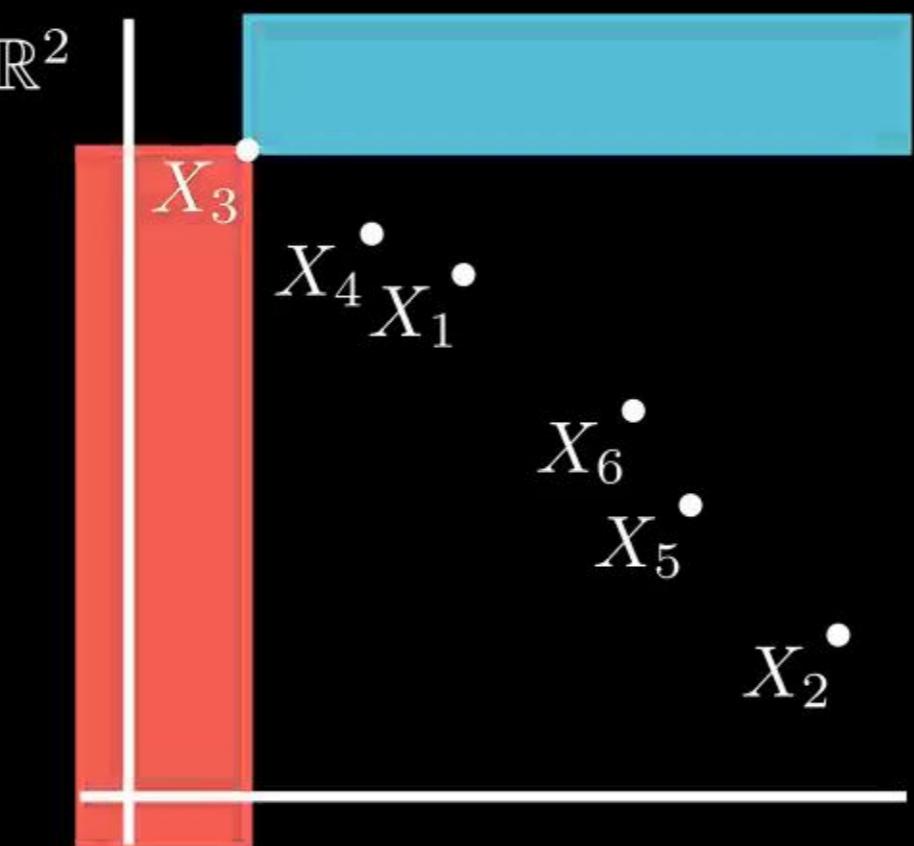
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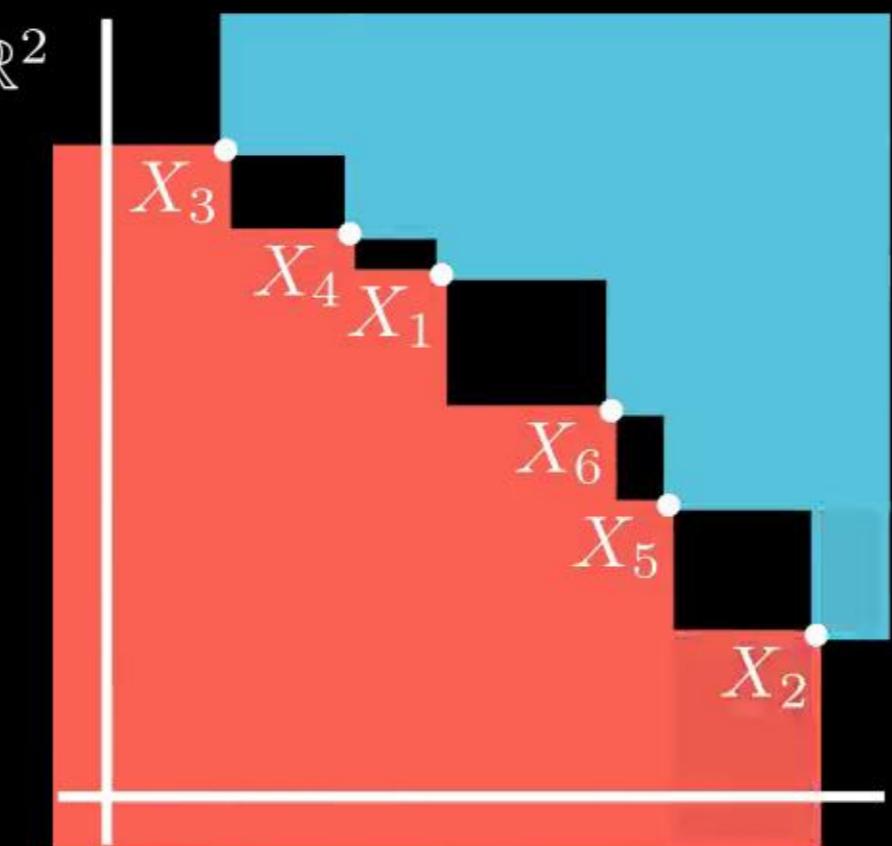
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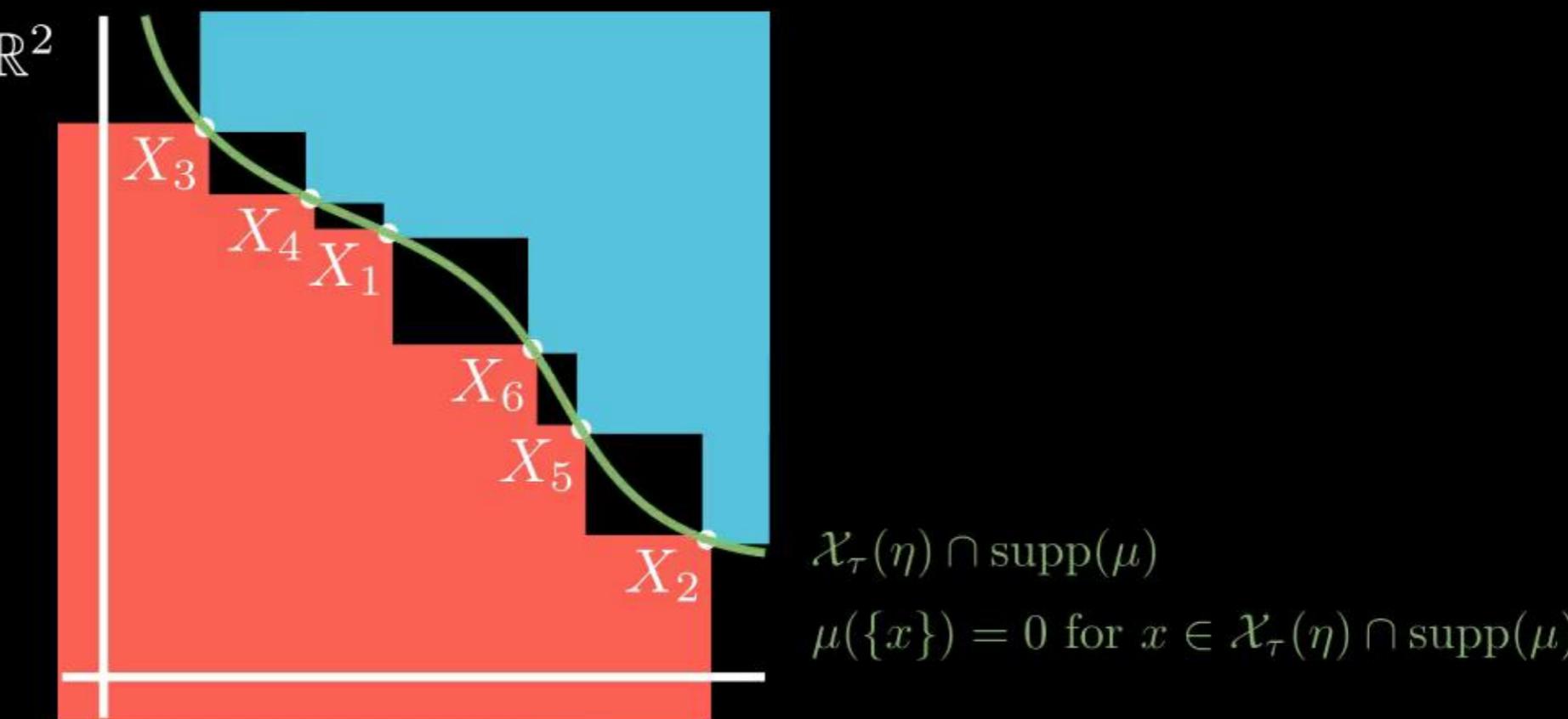
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Proof of the upper bound (for $m \in [n]$)

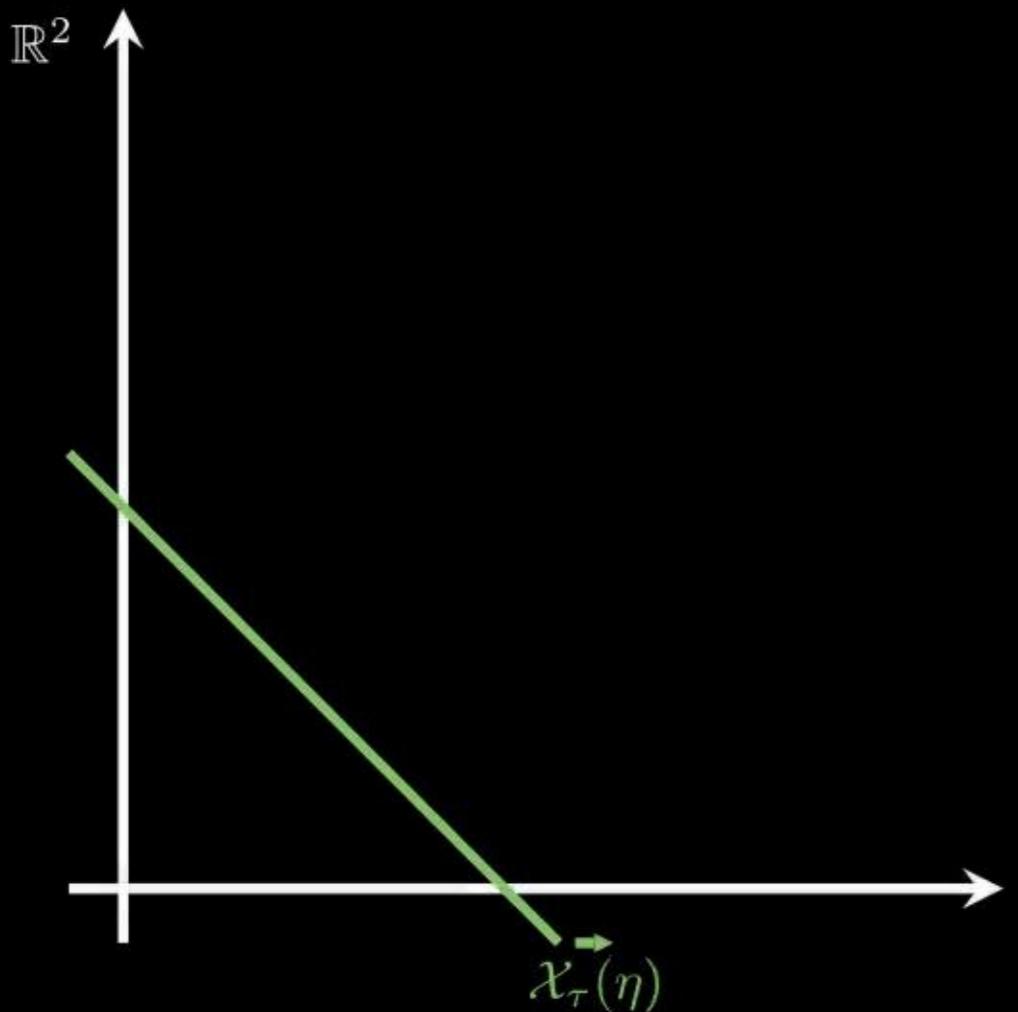
Theorem. There exists $C \geq 1$, depending only on (d, θ) , such that for any $P \in \mathcal{P}_{\text{MonReg}, d}(\sigma, \tau, \gamma, \lambda, \theta)$, $\alpha \in (0, 1)$ and $m \in [n]$,

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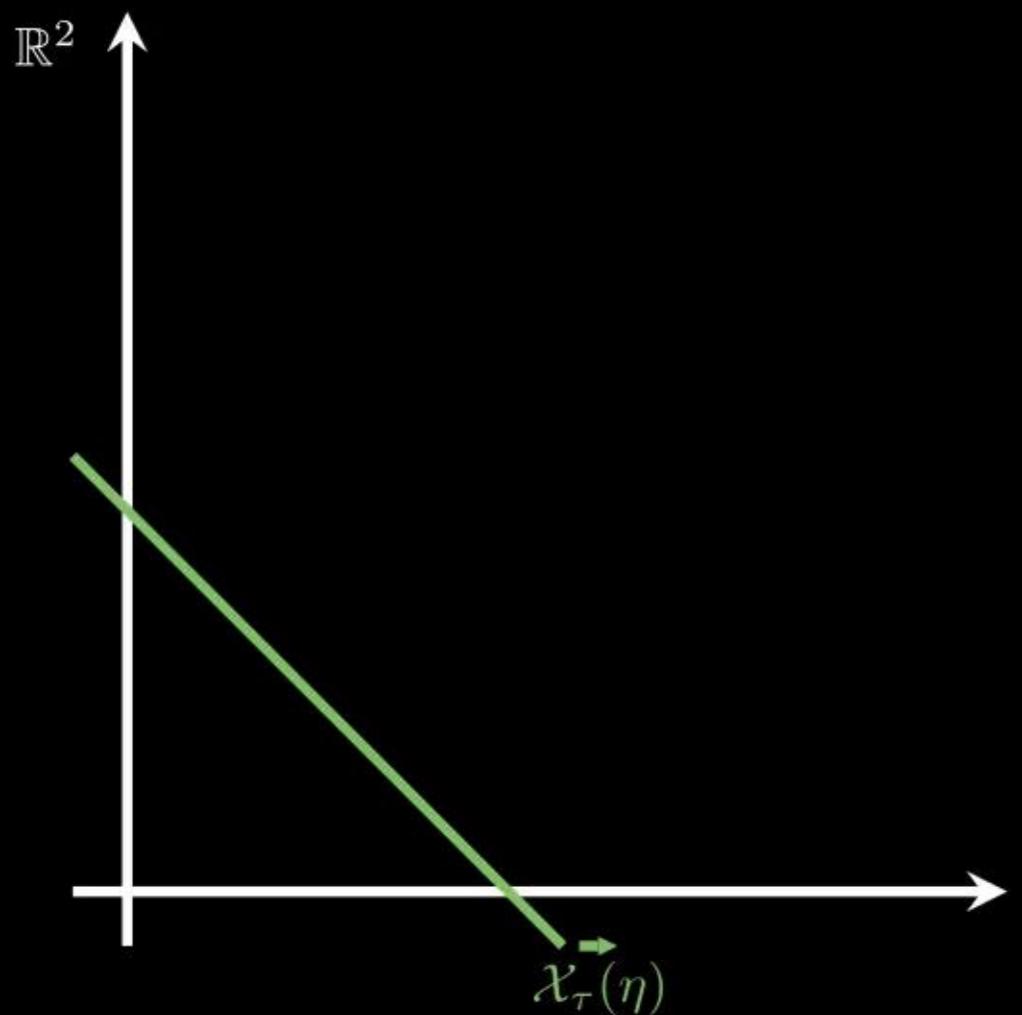
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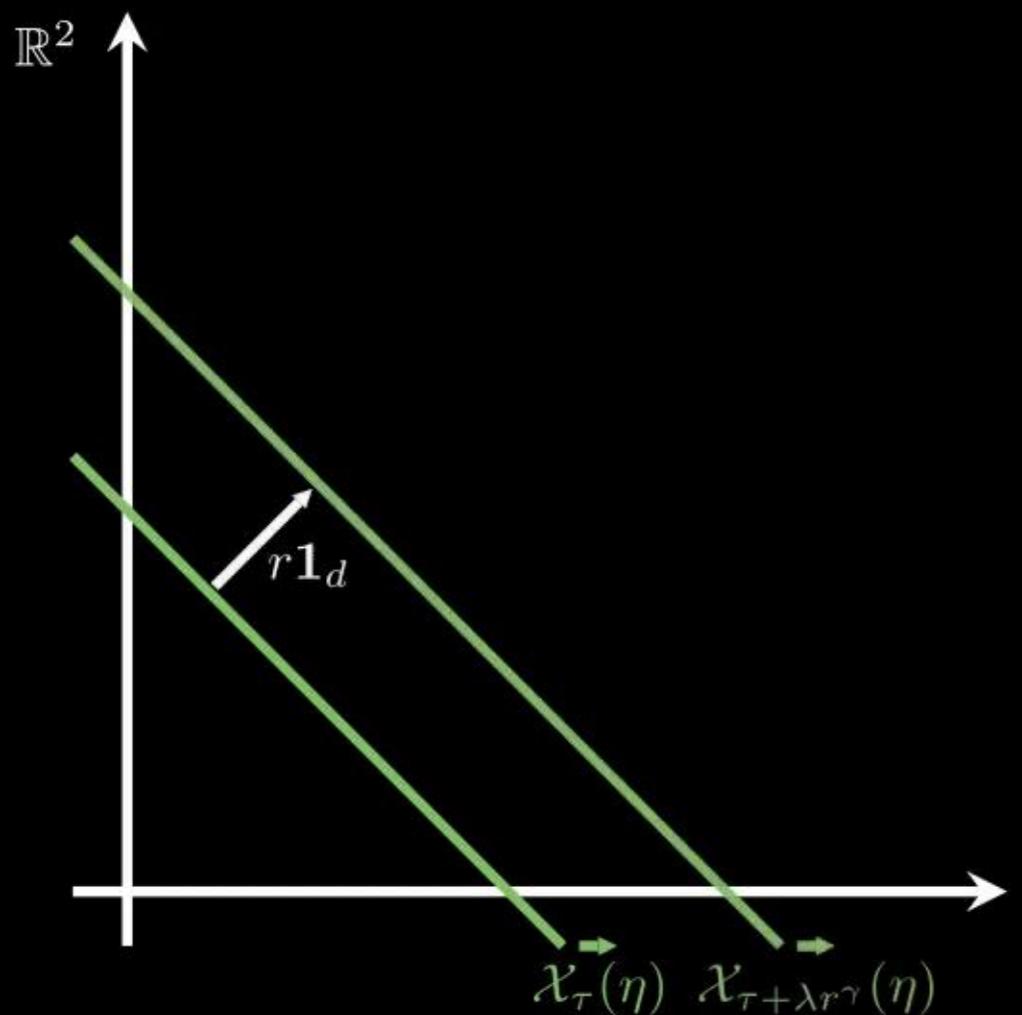


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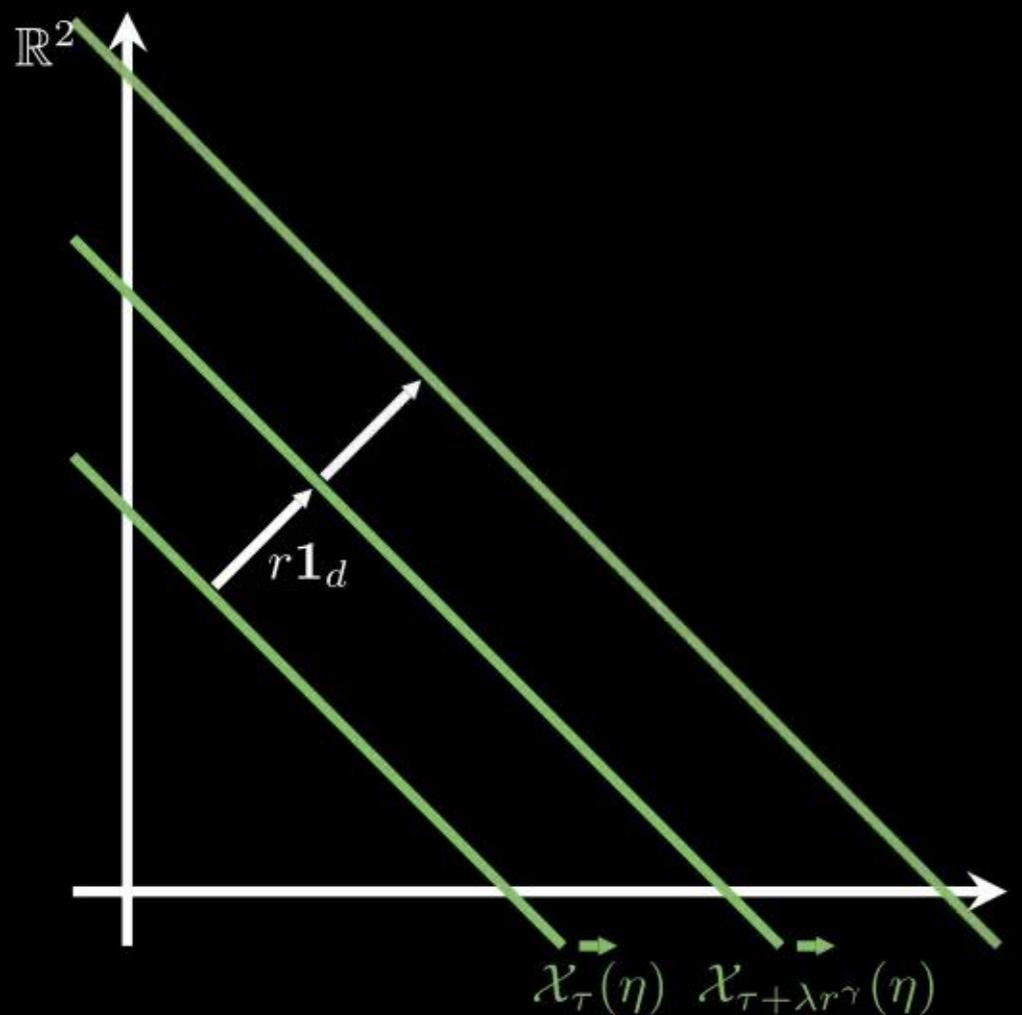


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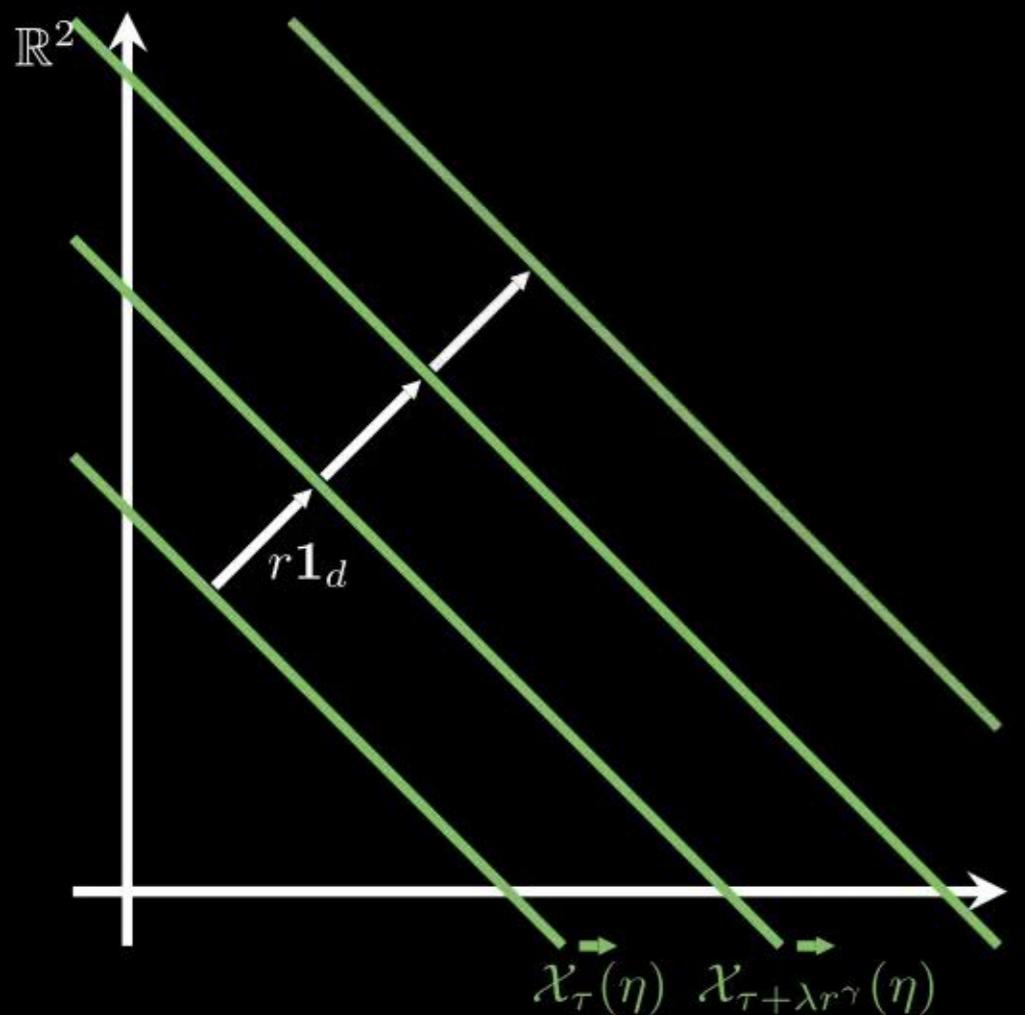


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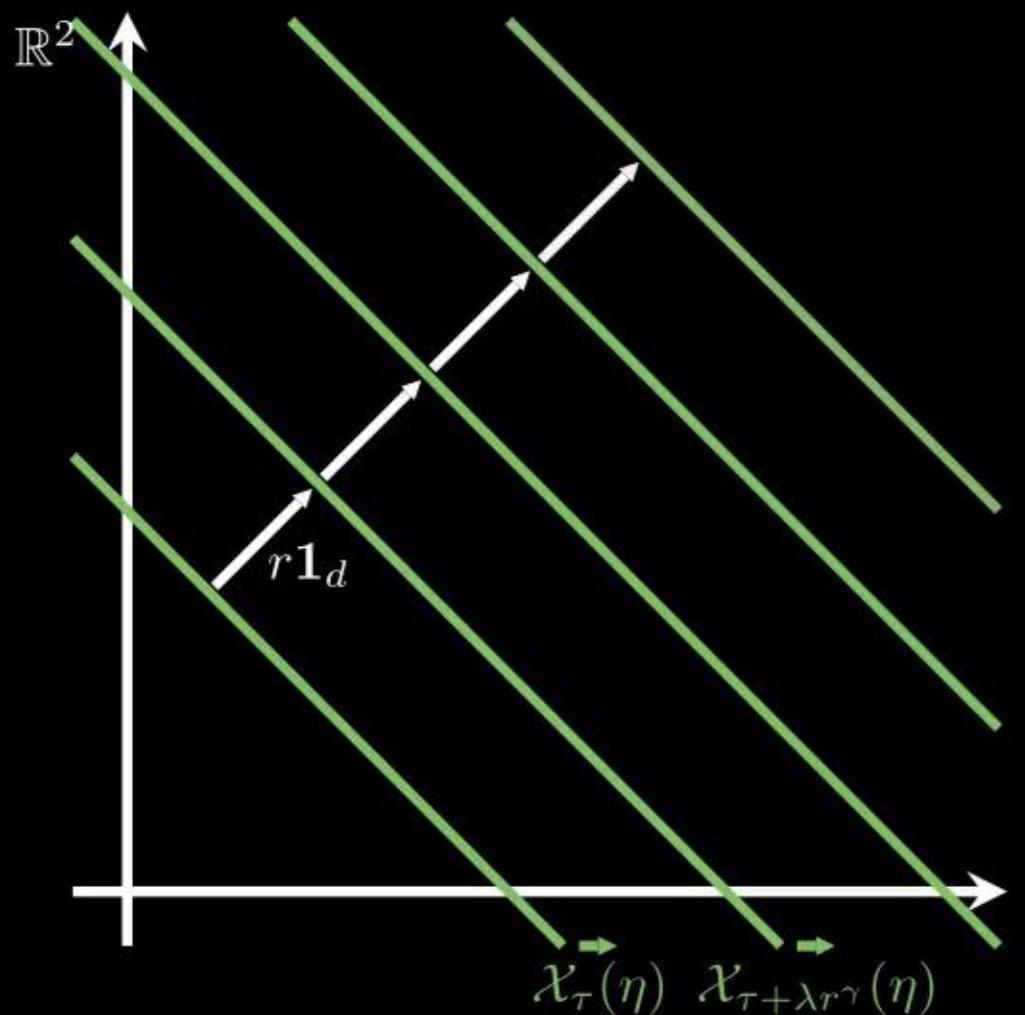


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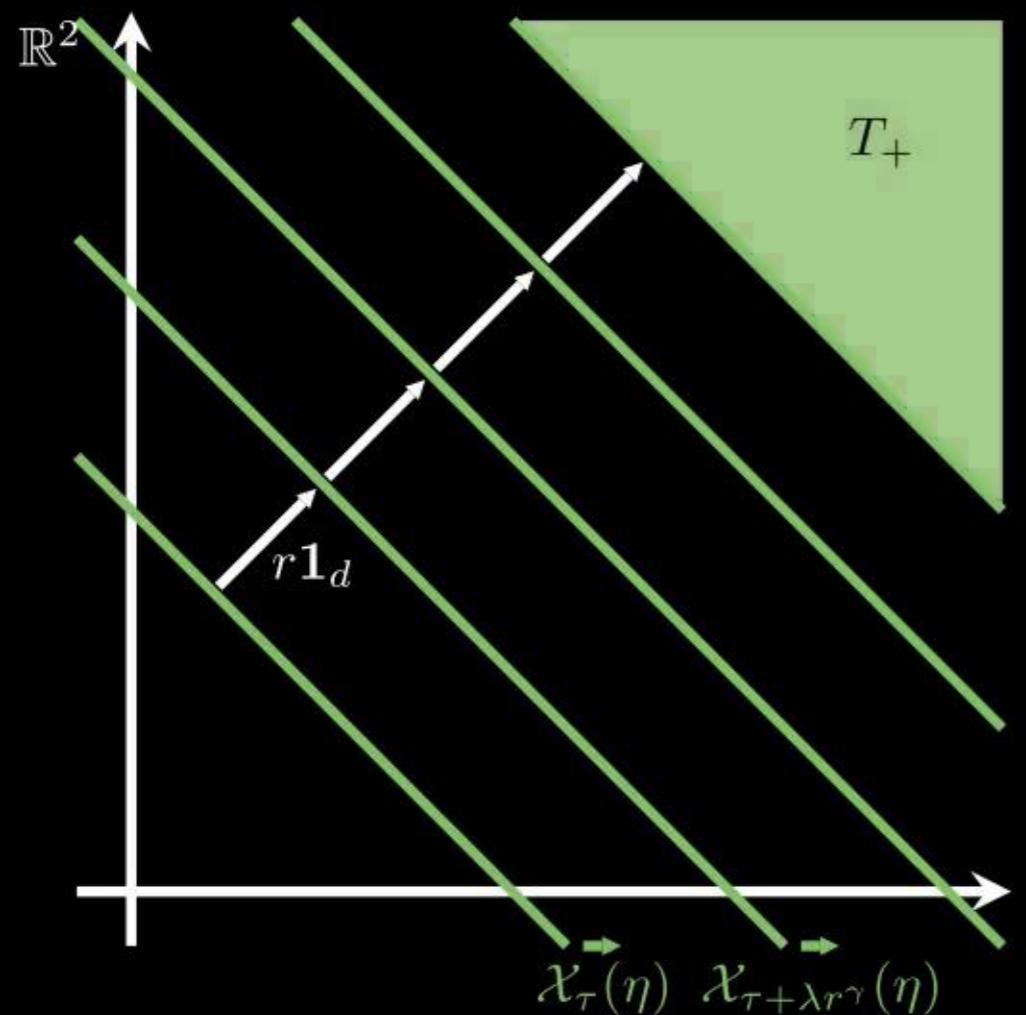


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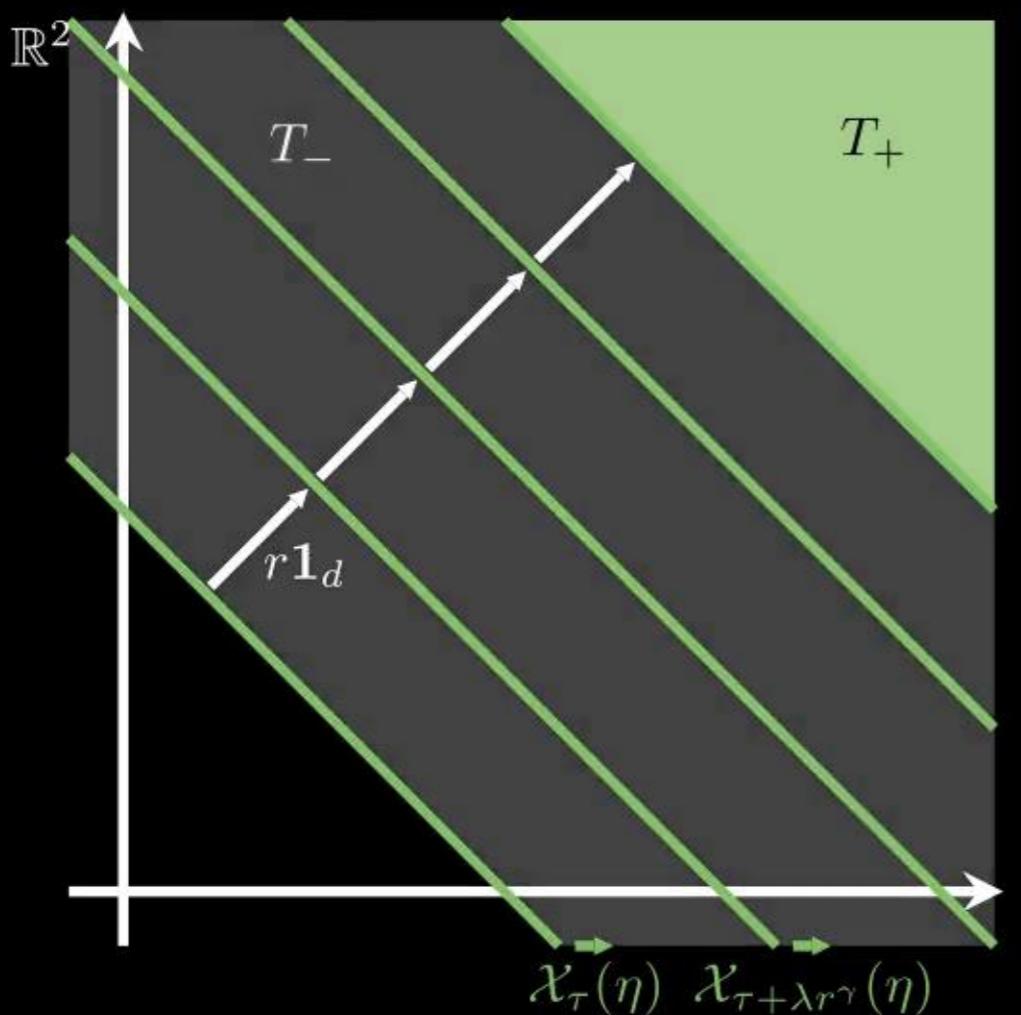


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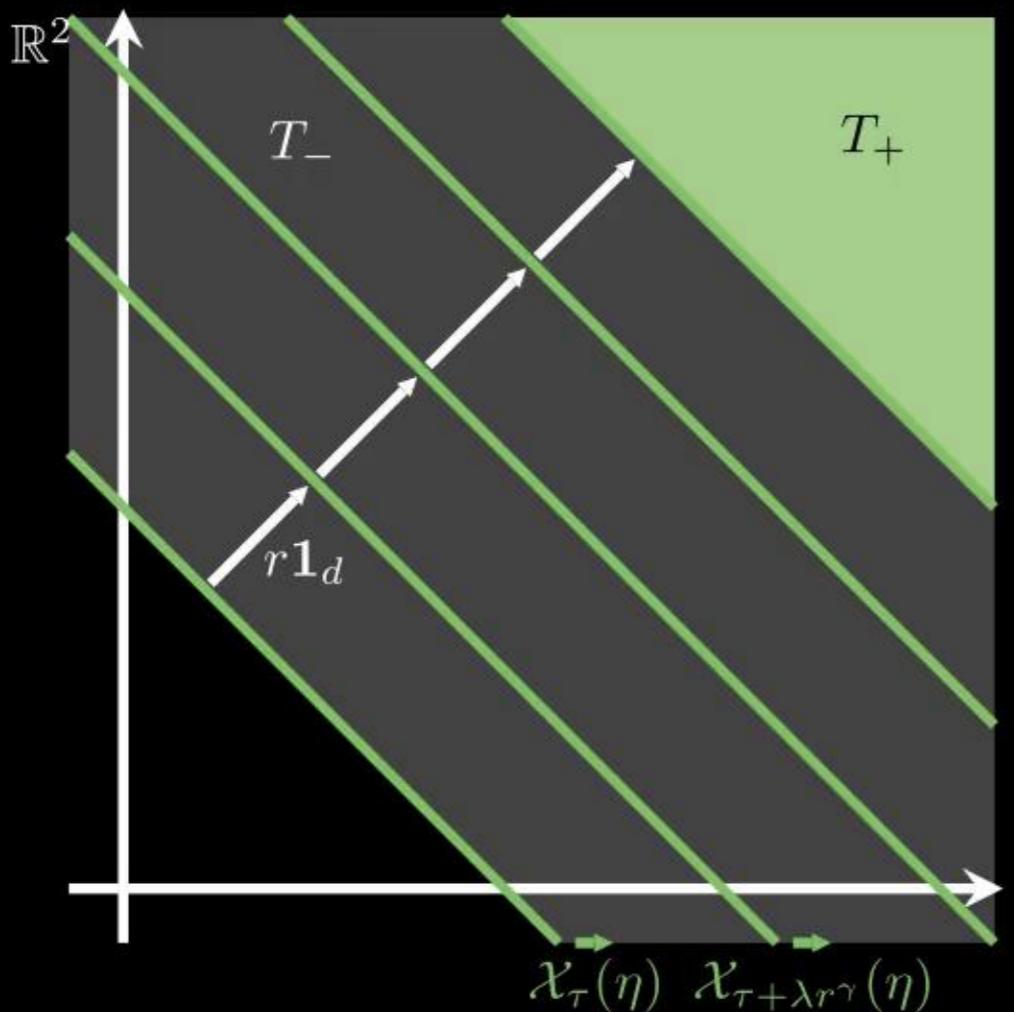


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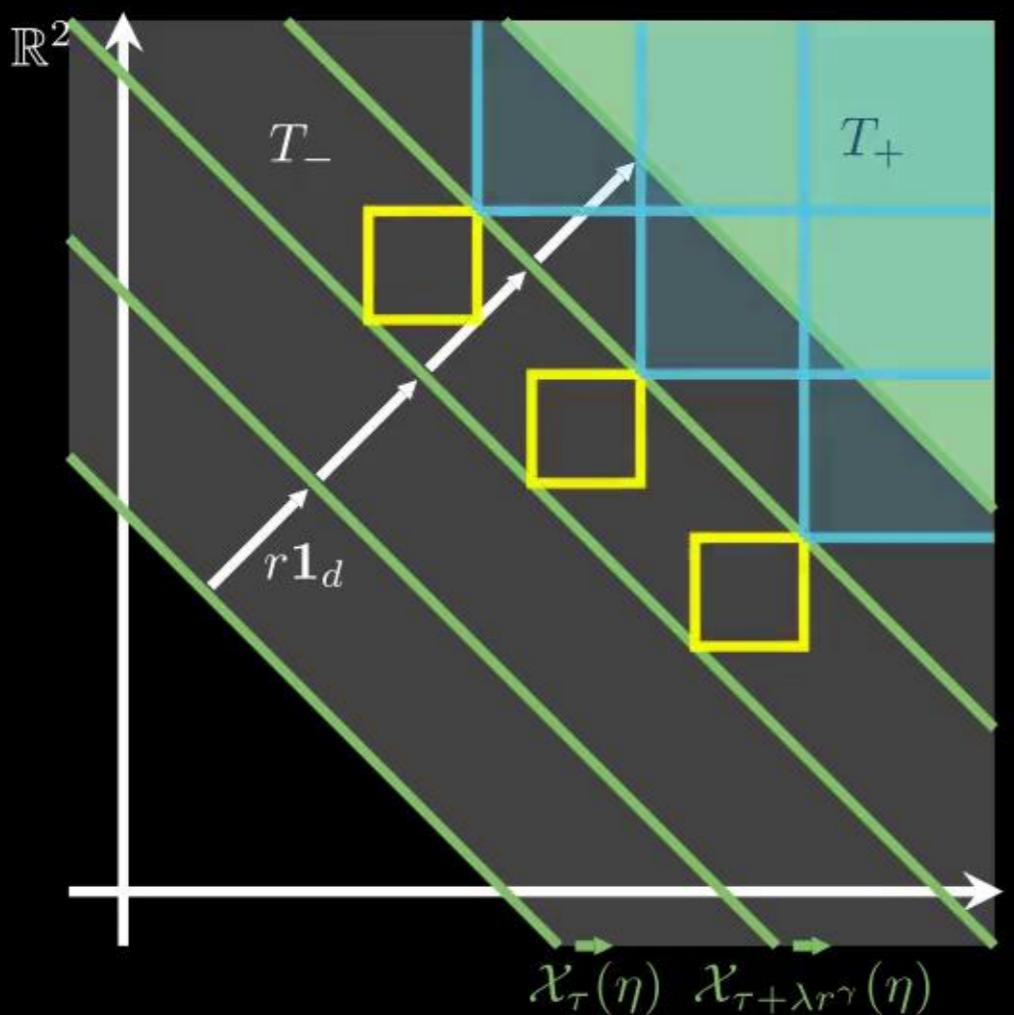


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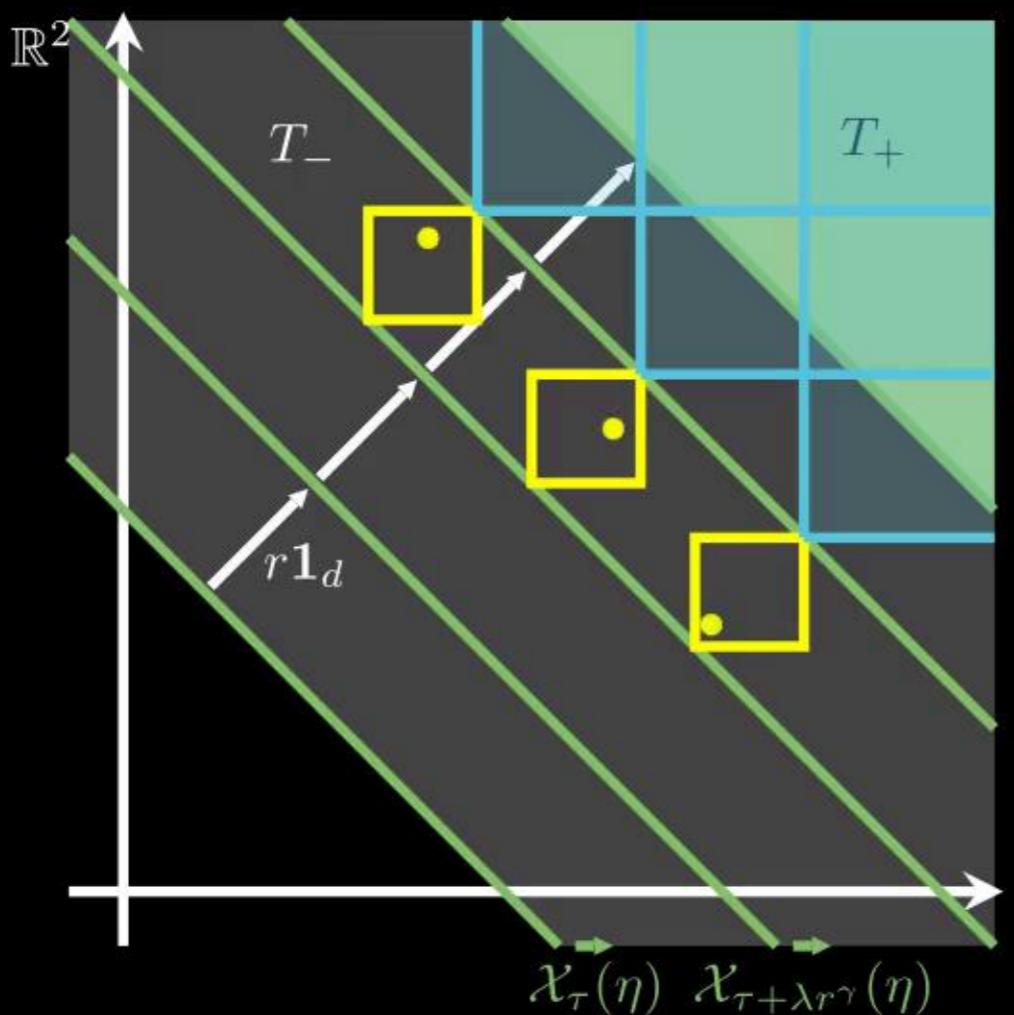


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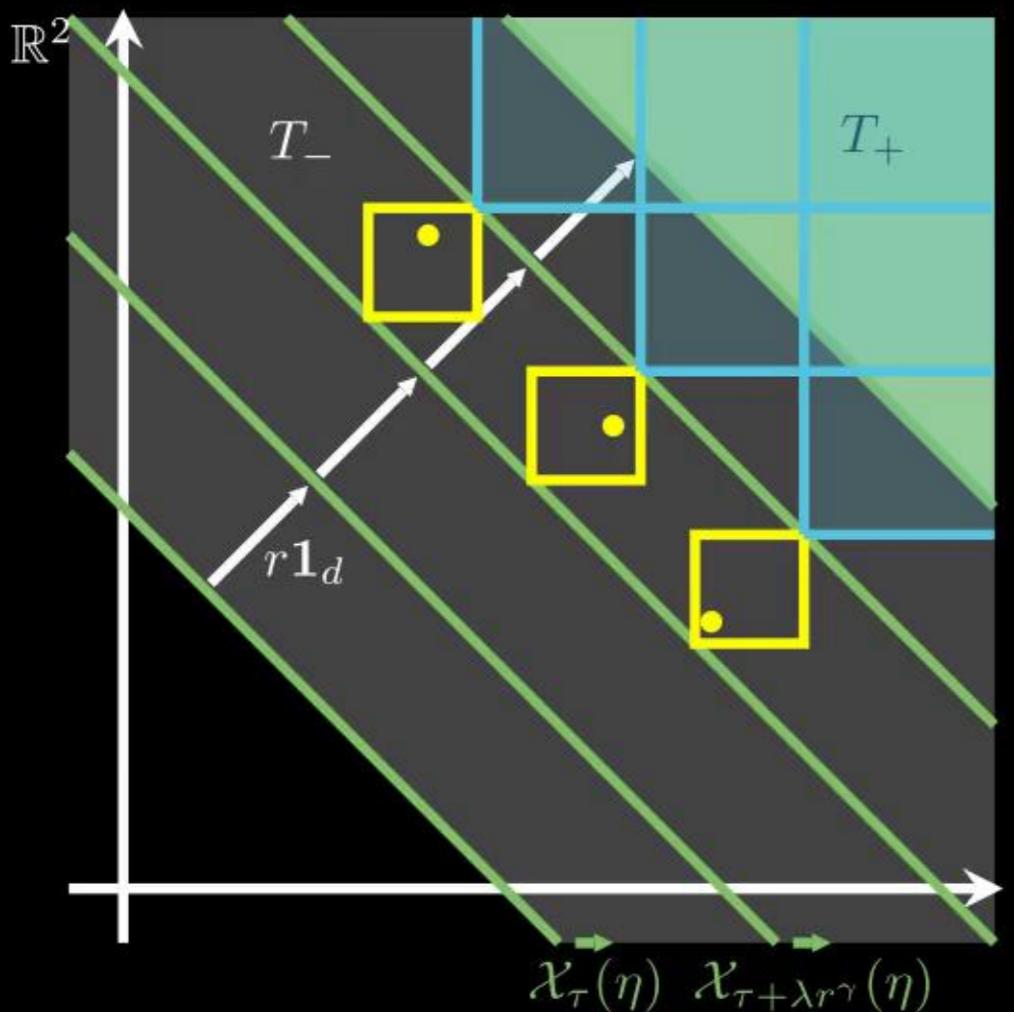


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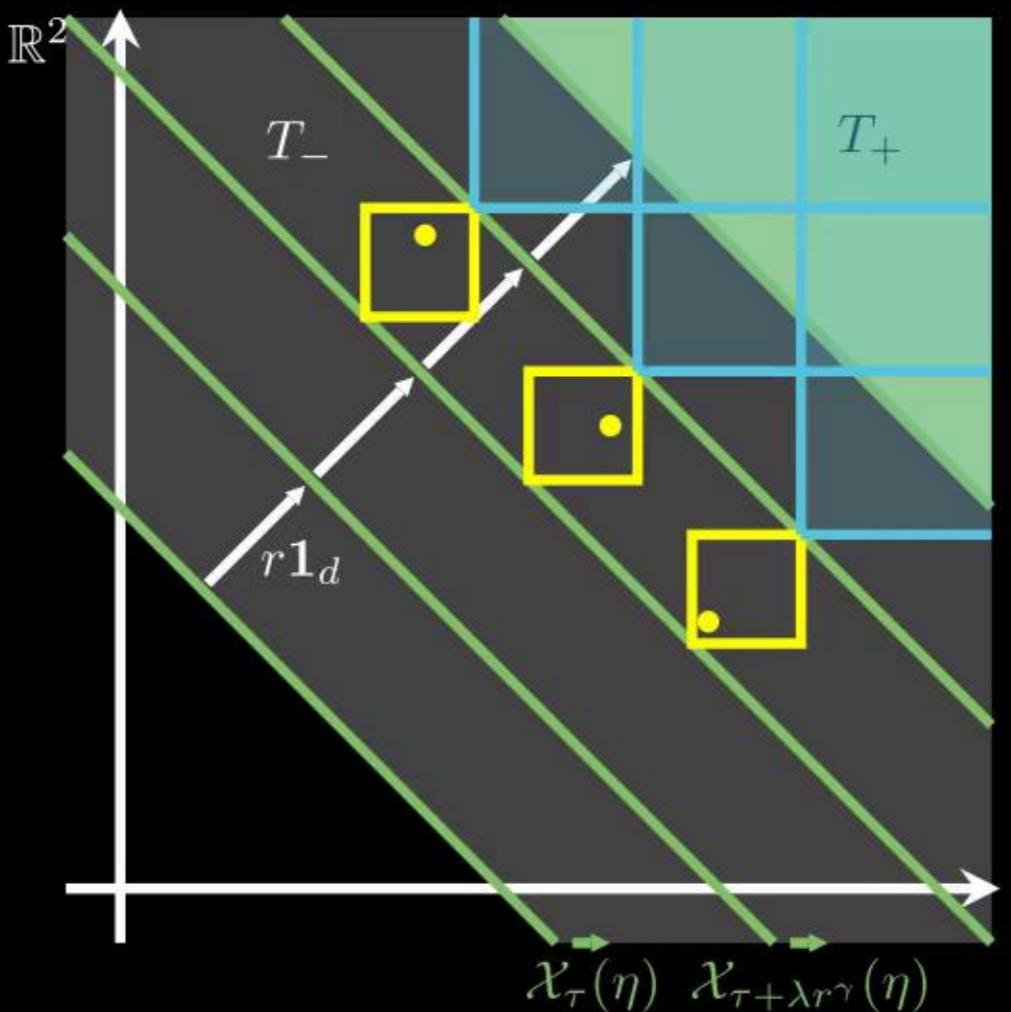


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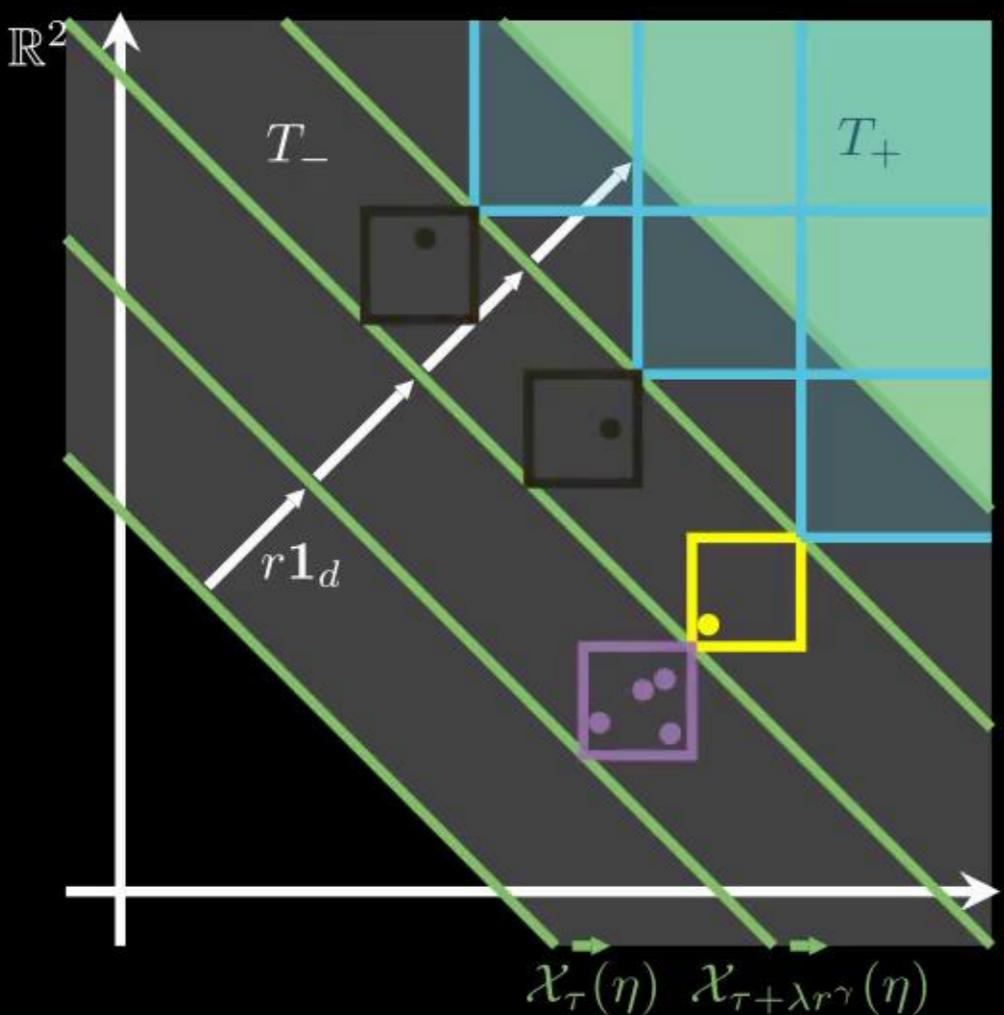
Out of $(X_i)_{i \in [m]}$ roughly mr^d in each box.

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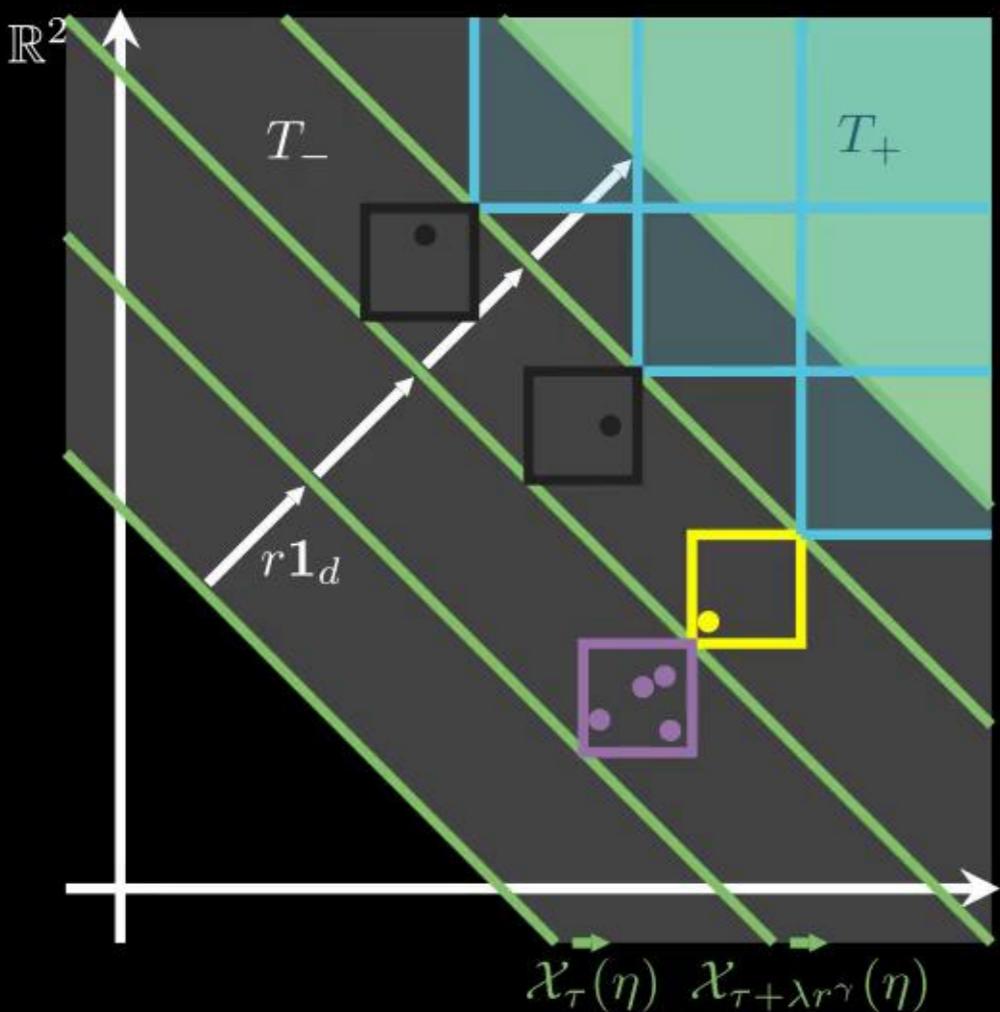
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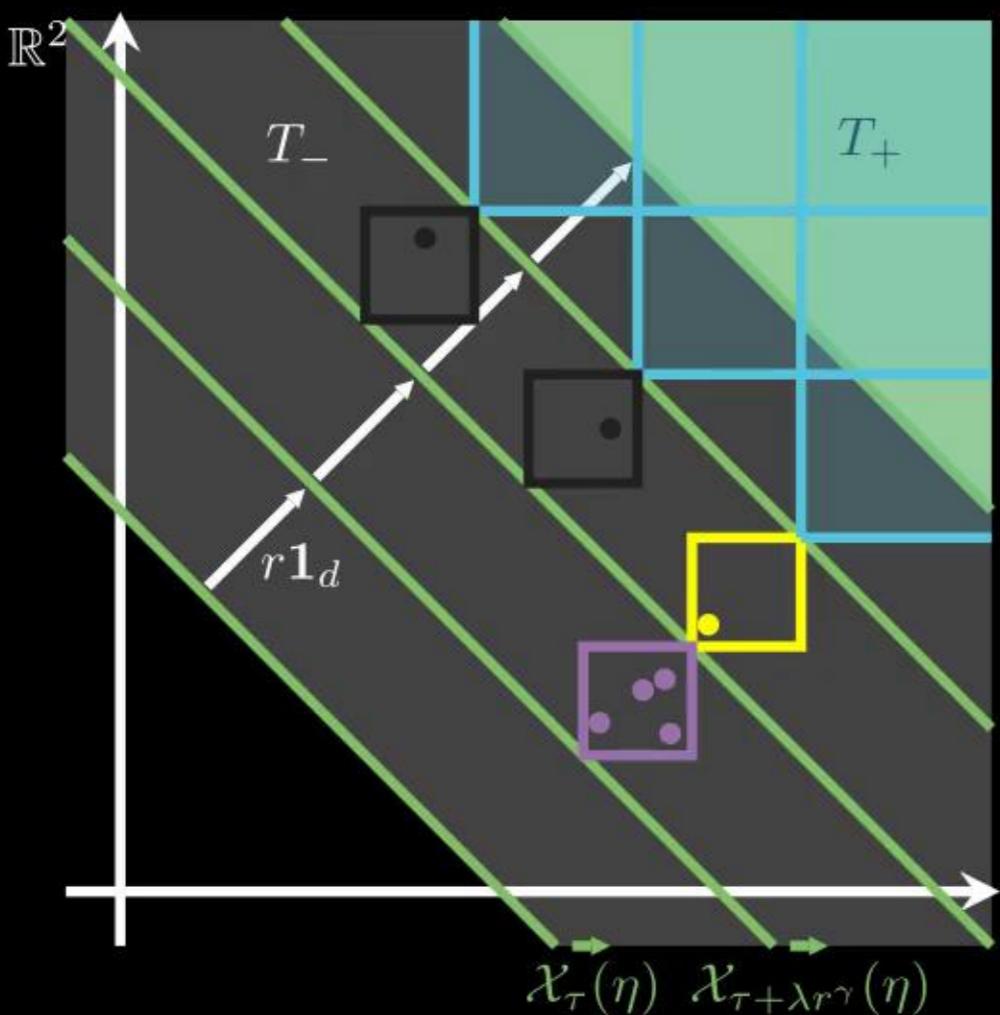
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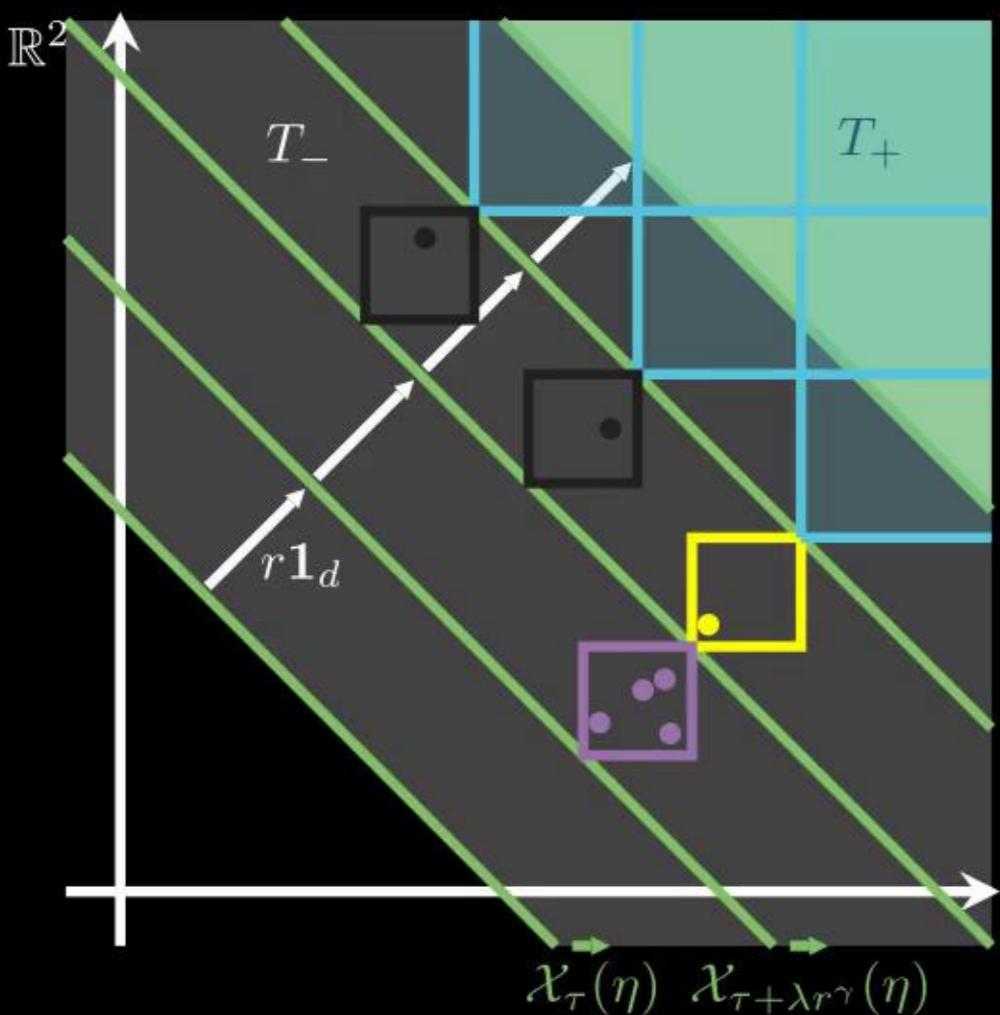
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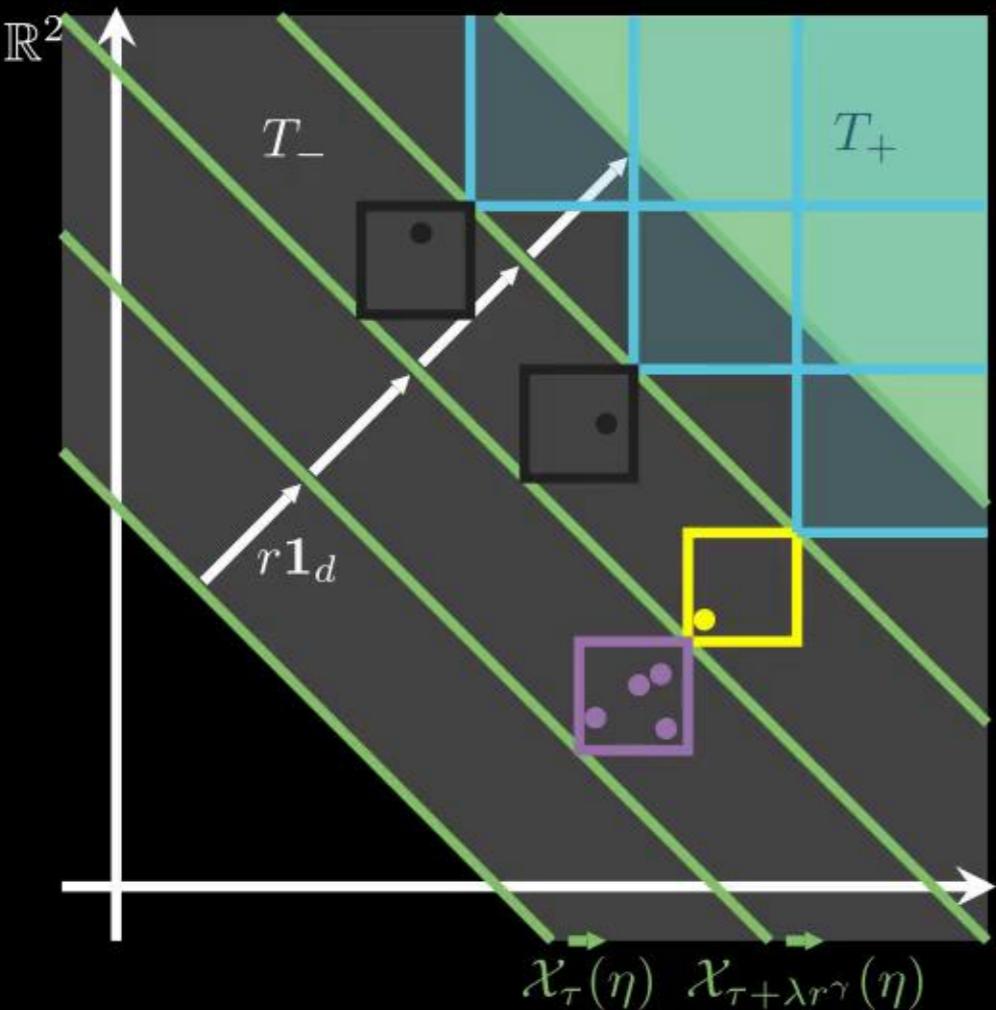
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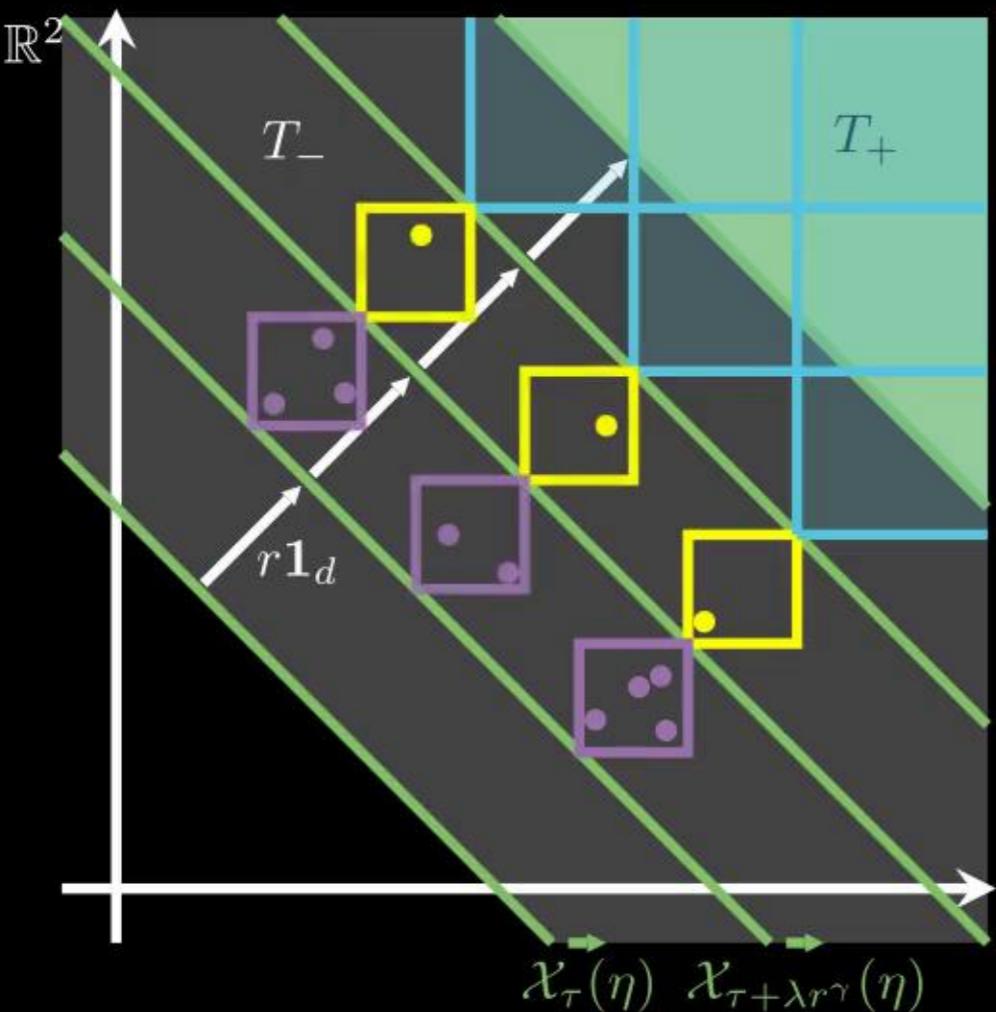
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