

1.

Table 1: The joint p.m.f. of  $X, Y$

$X \downarrow Y \rightarrow$	0	1	2	3
1/2	1/12	1/8	1/8	1/12
1	0	1/12	1/9	1/9
6	1/12	1/12	0	1/9

$$(a) P_X(x) = \sum_{y=0}^3 P_{XY}(x, y) = \frac{5}{12}$$

$$P_X(0) = \frac{1}{12} + \frac{1}{8} + \frac{1}{8} + \frac{1}{12} = \frac{5}{12}$$

$$P_X(1) = \frac{1}{12} + \frac{1}{12} + 0 + \frac{1}{9} = \frac{7}{18}$$

$$P_X(6) = \frac{1}{12} + \frac{1}{12} + 0 + \frac{1}{9} = \frac{7}{18}$$

$$A = \{x_1, x_2, x_3\}, \quad P_A(a) = \sum_{y=0}^3 P_{XY}(a, y) = \frac{5}{12}$$

$$P_A(0) = \frac{1}{12} + \frac{1}{8} + \frac{1}{8} + \frac{1}{12} = \frac{5}{12}$$

$$P_A(1) = \frac{1}{12} + \frac{1}{12} + 0 + \frac{1}{9} = \frac{7}{18}$$

$$P_A(6) = \frac{1}{12} + \frac{1}{12} + 0 + \frac{1}{9} = \frac{7}{18}$$

$$P_X(x) = \begin{cases} \frac{5}{12}, & k=0 \\ \frac{7}{18}, & k=1 \\ \frac{7}{18}, & k=2 \\ \frac{5}{12}, & k=3 \end{cases}$$

$$(b) P_{XY}(x, y) = \frac{1}{12} \neq \frac{5}{12} \cdot \frac{5}{12} = P_X(x) \cdot P_Y(y) \quad \therefore X, Y \text{ are not independent.}$$

$$(c) G_X(X(t)) = E[X(t)] - \mu_X \mu_Y$$

$$\mu_X = \frac{1}{2} + \frac{1}{6} + \frac{1}{6} + \frac{1}{2} = \frac{17}{12}, \quad \mu_Y = \frac{1}{2} + \frac{1}{6} + \frac{1}{6} + \frac{1}{2} = \frac{17}{12}$$

$$\begin{aligned} E[XY] &= \sum_{x,y} xy P_{XY}(x, y) = \frac{1}{12} \left( \frac{1}{2} P_X(1/2) + \frac{1}{6} P_X(1) + \frac{1}{6} P_X(6) + \frac{1}{2} P_X(3) \right) \\ &= \frac{1}{12} \left( P_X(1/2) + 2 P_X(1) + 2 P_X(6) + 3 P_X(3) \right) + \frac{1}{12} \left( P_X(1/2) + 2 P_X(1) + 2 P_X(6) + 3 P_X(3) \right) \\ &= \frac{1}{12} \left( \frac{1}{2} + \frac{1}{3} + \frac{1}{3} + \frac{1}{2} \right) + \frac{1}{12} \left( \frac{1}{2} + \frac{1}{3} + \frac{1}{3} + \frac{1}{2} \right) = \frac{5}{12} + \frac{5}{12} = \frac{10}{12} \end{aligned}$$

$$\therefore \frac{5}{12} + \frac{5}{12} + \frac{1}{12} - \frac{17}{12} \cdot \frac{17}{12} = -\frac{1495}{144}$$

2.

$$E[D(X_1, X_2)] = E[(X_1 - X_2)^2 + (Y_1 - Y_2)^2]$$

$$= E[X_1^2 + X_2^2 - 2X_1X_2 + Y_1^2 + Y_2^2 - 2Y_1Y_2]. \text{ Since } D \text{ is symmetric and } X_1, X_2 \text{ and } Y_1, Y_2 \text{ are independent, } E[D(X_1, X_2)] = E[D(X_1, X_2)] + E[D(Y_1, Y_2)]$$

$$E[D(X_1, X_2)] = \int_0^1 \int_0^1 (x_1 - x_2)^2 \cdot \frac{1}{4} dx_1 dx_2 = \frac{1}{4} \int_0^1 \int_0^1 (x_1^2 - 2x_1x_2 + x_2^2) dx_1 dx_2 = \frac{1}{4} \int_0^1 \left( \frac{x_1^3}{3} - x_1^2 x_2 + \frac{x_2^3}{3} \right) dx_1 dx_2 = \frac{1}{4} \int_0^1 \left( \frac{1}{3} - x_2 \right) dx_2 = \frac{1}{4} \left( \frac{1}{3} - \frac{1}{2} \right) = \frac{1}{24}$$

9.

Exercise 8.16. Let  $E[X] = 1$ ,  $E[X^2] = 3$ ,  $E[XY] = -4$  and  $E[Y] = 2$ . Find  $\text{Cov}(X, 2X + Y - 3)$ .

$$\text{Cov}(X, 2X + Y - 3) = 2\text{Cov}(X, X) + \text{Cov}(X, Y)$$

$$= 2(E[X^2] - E[X]^2) + (E[XY] - E[X]E[Y])$$

$$= 2(3 - 1^2) + (-4 - 1 \cdot 2) = 4 - 6 = -2 \quad \therefore -2$$

6.



0.25

$X$  := number of shirts that a color with at least one of the two shirts adjacent to it

$$X = \sum_{i=1}^n \mathbb{1}_{\{X_i\}}, \quad \mathbb{1}_{\{X_i\}} \text{ is an indicator such that } \mathbb{1}_{\{X_i\}} = \begin{cases} 1, & \text{color with at least one of the two shirts adjacent to it} \\ 0, & \text{otherwise} \end{cases} \quad \text{and} \quad P(\mathbb{1}_{\{X_i\}} = 1) = \frac{2}{3}, \quad \mathbb{1}_{\{X_i\}} = 0$$

$$E[X] = E\left[\sum_{i=1}^n \mathbb{1}_{\{X_i\}}\right] = \sum_{i=1}^n E[\mathbb{1}_{\{X_i\}}] = \sum_{i=1}^n P(\mathbb{1}_{\{X_i\}} = 1) = n \cdot \frac{2}{3} \quad \therefore E[X] = \frac{2}{3}n$$

$$\text{Var}(X) = \text{Var}\left(\sum_{i=1}^n \mathbb{1}_{\{X_i\}}\right) = \sum_{i=1}^n \text{Var}(\mathbb{1}_{\{X_i\}}) + 2 \sum_{1 \leq i < j \leq n} \text{Cov}(\mathbb{1}_{\{X_i\}}, \mathbb{1}_{\{X_j\}})$$

let's denote  $\bar{i} := 1 - i$ , so that  $\bar{i}$  is the complement of  $i$ . Then  $\text{Cov}(\mathbb{1}_{\{X_i\}}, \mathbb{1}_{\{X_{\bar{i}}\}}) = \text{Cov}(\mathbb{1}_{\{X_i\}}, \mathbb{1}_{\{X_{\bar{i}}\}})$ , now let's consider the cases ①  $|k \bmod n - j \bmod n| = 1$  ②  $|k \bmod n - j \bmod n| \geq 3$  ③  $|k \bmod n - j \bmod n| = 2$

For case ①, we know adjacent  $t$ -shirts are dependent.

For case ②, let's consider the sets of  $t$ -shirts, let  $T_i := T$ -shirt at position  $i$ ,  $(T_i, T_j, T_{j+1})$ ,  $(T_{k+1}, T_k, T_{k+2})$  where  $|k \bmod n - j \bmod n| \geq 3$  then every  $t$ -shirt from the first set is at least 3 positions different from the second set, hence they are independent.

For case ③, let's consider the set of  $t$ -shirts  $(T_i, T_j, T_{j+1}, T_{j+2}, T_{j+3})$

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For the set  $\{j+1, j_2, j_{2n}, j_{2n+1}, j_{2n+2}\}$ , let's see if positions  $j$  and  $j+2$  are independent.

Since we're calculating  $Cov(j_2, j_{2n})$  we see if  $P(+\text{shirt } j \text{ doesn't share same color } j+1, j+2) = P(+\text{shirt } j_2 \text{ is different color from } j+1, j+2) = P(+\text{shirt } j \text{ doesn't share same color } j+1, j+2) \cdot P(+\text{shirt } j_2 \text{ is different color from } j+1, j+2)$

$P(+\text{shirt } j \text{ doesn't share same color } j+1, j+2)$  and  $+ \text{shirt } j_2 \text{ is different color from } j+1, j+2 \Rightarrow$  let's say  $j_{2n}$  is G then  $j_2, j_{2n}$  are not G and  $j_2, j_{2n}$  are not the same as  $j_2, j_{2n}$  respectively.

$$\text{then } p = 3 \cdot \frac{1}{2} \cdot \left(\frac{2}{3}\right) \cdot \left(\frac{2}{3}\right) = \left(\frac{2}{3}\right)^2$$

$$P(+\text{shirt } j \text{ doesn't share same color } j+1, j+2) = 3 \cdot \left(\frac{1}{3}\right) \cdot \left(\frac{2}{3}\right) \cdot \left(\frac{2}{3}\right) = P(+\text{shirt } j_2 \text{ is different color from } j+1, j+2)$$

So since both LHS and RHS have probability  $\left(\frac{2}{3}\right)^2$  they are independent.

$$\text{Therefore, } Var(X) = \sum_{i=1}^n Var(S_i) + 2 \sum_{i=1}^n Cov(r_{i \text{ and } n}, r_{i+1 \text{ and } n})$$

$$= n \cdot Var(S_i) + 2 \cdot n \cdot Cov(r_i, r_n)$$

$$= n \cdot P(S_i=1) \cdot (1 - P(S_i=1)) + 2 \cdot n \cdot (P(r_i=1 | n | r_n=1) - P(r_i=1) \cdot P(r_n=1))$$

$$= n \cdot \frac{2}{3} \cdot \frac{1}{3} + 2 \cdot n \cdot \left(\frac{2}{3} \cdot \frac{1}{3} - \frac{1}{3} \cdot \frac{1}{3}\right)$$

$$= \frac{2}{3}n + \frac{2}{3}n = \frac{4}{3}n = \frac{4}{3}n$$

$$\therefore \frac{4}{3}n$$

$$9. r_1, \dots, r_n \text{ be i.i.d each having pdf } f(y) = \begin{cases} \frac{1}{y^2}, & y \geq 1 \\ 0, & y < 1 \end{cases}$$

$$E(X) = E\left[\sum_{i=1}^n 1_{A_i}\right] = \sum_{i=1}^n E[1_{A_i}] = n \cdot P(X \geq r_1), \text{ for } B = \{y_1, y_2 : y_1 \geq r_1\}, \int_0^{\infty} \int_{r_1}^{\infty} f_2(x, y) dy dx = \int_0^{\infty} \int_{r_1}^{\infty} \frac{1}{x^2} \cdot \frac{1}{y^2} dy dx = \int_0^{\infty} \left[ -\frac{1}{y} \right]_{r_1}^{\infty} dx = \int_0^{\infty} \frac{1}{r_1} dx = \left[ \frac{x}{r_1} \right]_0^{\infty} = \frac{1}{r_1}$$

$$\therefore E(X) = n \left(\frac{1}{2}\right)$$

$$Var(X) = Var\left(\sum_{i=1}^n 1_{A_i}\right) = \sum_{i=1}^n Var(1_{A_i}) + 2 \sum_{1 \leq i < j \leq n} Cov(1_{A_i}, 1_{A_j}) \longrightarrow \text{if } |k-j| \geq 2, \text{ then events } 1_{A_i} \text{ and } 1_{A_j} \text{ are independent so } P(1_{A_i} | 1_{A_j}) = P(1_{A_i} \geq r_1 | 1_{A_j} \geq r_2, r_1) = P(1_{A_i} \geq r_1) P(1_{A_j} \geq r_2, r_1)$$

$$= n \cdot Var(1_{A_i}) + 2 \cdot \sum_{i=1}^{n-1} Cov(1_{A_i}, 1_{A_{i+1}})$$

$$= n \cdot Var(1_{A_i}) + 2 \cdot (n-1) \cdot Cov(1_{A_i}, 1_{A_{i+1}})$$

$$= n \cdot Var(1_{A_i}) + 2 \cdot (n-1) \cdot \{P(A_i | A_{i+1}) - P(A_i)P(A_{i+1})\}$$

$$= n \cdot \{E[1_{A_i}^2] - E[1_{A_i}]E[1_{A_i}]\} + 2 \cdot (n-1) \cdot \{P(A_i | A_{i+1}) - P(A_i)P(A_{i+1})\}$$

$$= n \cdot \{E[1_{A_i}] \cdot (1 - E[1_{A_i}])\} + 2 \cdot (n-1) \cdot \{P(A_i | A_{i+1}) - P(A_i)P(A_{i+1})\}$$

$$= n \cdot \left(\frac{1}{2}\right) \cdot \left(1 - \frac{1}{2}\right) + 2 \cdot (n-1) \cdot \{P(A_i | A_{i+1}) - P(A_i)P(A_{i+1})\}$$

$$P(A_i | A_{i+1}) = P\left(\frac{1}{2} \geq r_1 \geq r_2 \geq r_1 \geq r_2\right) = \int_0^{\infty} \int_{r_1}^{\infty} \int_{r_2}^{\infty} \frac{1}{x^2} \cdot \frac{1}{y^2} \cdot \frac{1}{z^2} dz dy dx = \int_0^{\infty} \int_{r_1}^{\infty} \frac{1}{x^2} \cdot \frac{1}{y} dy dx = \int_0^{\infty} \left[ -\frac{1}{y} \right]_{r_1}^{\infty} dx = \int_0^{\infty} \frac{1}{r_1} dx = \left[ \frac{x}{r_1} \right]_0^{\infty} = \frac{1}{r_1}$$

$$= n \cdot \left(\frac{1}{2}\right) \cdot \left(1 - \frac{1}{2}\right) + 2 \cdot (n-1) \cdot \left(\frac{1}{2} - \frac{1}{4}\right)$$

$$\therefore E(X) = n \cdot \frac{1}{2}, Var(X) = n \cdot \frac{1}{4} \cdot \left(1 - \frac{1}{2}\right) + 2 \cdot (n-1) \cdot \left(\frac{1}{4} - \frac{1}{16}\right)$$