Lab 4_MCM

April 23, 2018

```
In [3]: import numpy as np
    import matplotlib.pyplot as plt
    from scipy.special import gamma as gamma
    import scipy.stats as stats

%matplotlib inline
```

1 Lab 4: Heirarchial Models in Meta-Study Analysis

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```
In [46]: def posterior(fx):
    mu = fx[0]
    tau_sq = fx[1]
# print (mu, tau_sq)

    odds_ratio = [1.06, -0.1, 0.62, 0.02, 1.07, -0.02, -0.12, -0.38, 0.51, 0.0, 0.38, 0.63]
    sigma_i = [0.37, 0.11, 0.22, 0.11, 0.12, 0.12, 0.22, 0.23, 0.18, 0.32, 0.20, 0.25]

    prior = np.log(stats.norm.pdf(mu, 0, 1000.*tau_sq))+ np.log(stats.gamma.pdf(1/tau_scored))

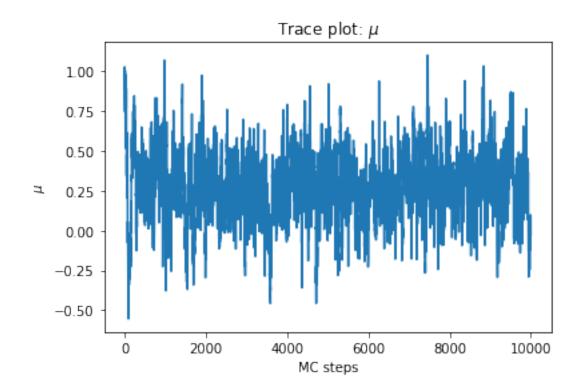
    overall_likelihood = 0
    for i in range(len(odds_ratio)):
        exp_part1 = (mu*sigma_i[i]**2 + odds_ratio[i]*tau_sq)**2/(sigma_i[i]**2 + tau_scored)
        exp_part_all = (-mu**2*sigma_i[i]**2 - odds_ratio[i]**2*tau_sq + exp_part1)/(2*
        olstep_i = np.log(1/np.sqrt(2*np.pi*(sigma_i[i]**2 + tau_sq))) + exp_part_all
        overall_likelihood = overall_likelihood + olstep_i

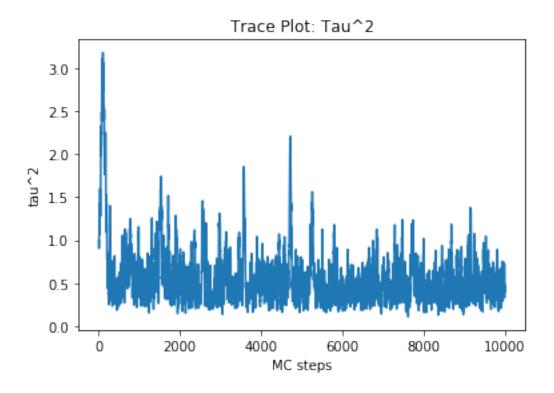
    return overall_likelihood + prior
```

```
In [48]: # proposal
    def proposal(xc, delta):
        # change tau_sq and mu with delta
        tau_sq = -1
        mu = np.random.normal(xc[0], delta[0])
        while tau_sq < 0:</pre>
```

```
tau_sq = np.random.normal(xc[1], delta[1])
             return [mu, tau_sq]
In [52]: # acceptance test
         def accept(x_old, x_new, posterior):
             accept_r = False
             # calculate posterior_new/posterior_old
             ratio = np.exp(posterior(x_new) - posterior(x_old))
             # if posterior_ratio is greater than some random value, accept proposal
             if ratio > 1.:
                 accept_r = True
             elif ratio > np.random.rand():
                 accept_r = True
             else:
                 accept_r = False
             # return acceptance (bool)
             return accept_r
In [50]: # metropolis driver
         def metropolis_MCMC(nmcs, delta, init):
             x = np.zeros((nmcs, 2))
             xc = init
             Acc = 0.0
             # for loop n steps (nmcs)
             for i in range(nmcs):
                 # log current state
                 x[i] = xc
                 # new state, proposed
                 xn = proposal(xc, delta)
                 # if accept,
                 if (accept(xc, xn, posterior)):
                     # current state = new state
                     xc = xn
                     # inc number of sucesses
                     Acc = Acc + 1
             # return all states, and ratio of success
             return x, Acc/nmcs
In [75]: x, acc_percent = metropolis_MCMC(10000, [0.1, 0.1], [1.0, 1.0])
         acc_percent
/home/celine/anaconda2/lib/python2.7/site-packages/ipykernel_launcher.py:10: RuntimeWarning: div
```

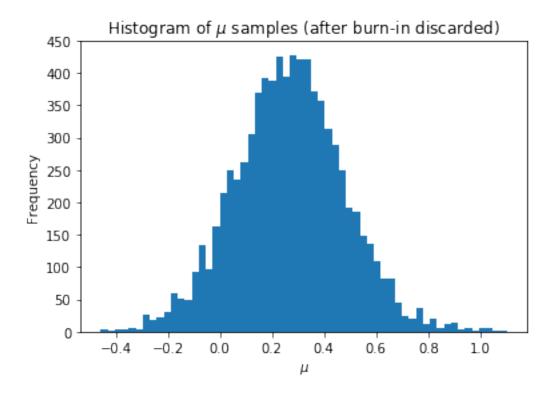
Remove the CWD from sys.path while we load stuff.





```
In [86]: # discarded burn-in

plt.hist(x[2000:, 0], bins='auto')
plt.ylabel("Frequency")
plt.xlabel("$\mu$")
plt.title("Histogram of $\mu$ samples (after burn-in discarded)")
plt.show()
```



Out [76]: 0.26882826910603042

The expected value of μ is 0.2688. The "true" log odds ratio for the population is μ , and the logarithm of the odds ratio ranges between $-\infty$ through 0, to ∞ . Since the estimation of μ turns out to be closest to 0, this follows that the trait produces no enhancement in disease (no effect). Although the expected value is not exactly 0, you still have to factor in Measurement error and Sampling Error