Unicare zero inflated admissions counts models summary

Summary remarks

- We compare inpatient admissions for Unicare Study vs all the other members
- 270 window for inpatient admissions.
- The zero inflated models appear to be a good approximation of the data generating process.
- The various models are in agreement as to the effect of ibis.
- Statistical significance for coefficient for cohort-ibis vs control is achieved in some cases.
- The difference in mean admissions is more that 50% reduced for the ibis vs control. This is borne out by the models as well.
- On the other hand the same coefficient value gives statistical signifiance for less than 10% percent increase in the probability of zero admissions for ibis cohort, depending on values of covariates.
- Additional covariates were selected based on correlations and other considerations. The inclusion of these can reduce the variance of estimates, even for a randomized control study where there is complete balance across cohorts.

Models

Based on earlier model results we consider Bayes and frequentist zero inflated Poisson and negative binomials models.

For the zero inflated Poisson model we use the following model specification:

$$\begin{aligned} \text{counts}_i &\sim ZIP(\pi_i, \mu_i) \\ &\log \mu_i = \beta_0 + \beta_1 \text{cohort}_i + \beta_2 \text{chf}_i + \beta_3 \text{age}_i + \beta_4 \text{afib}_i \\ &\log \frac{\pi_i}{1 - \pi_i} = \gamma_0 + \gamma_1 \text{age}_i \end{aligned}$$

Note that the zero inflated negative binomial models model the mean the same way. The expected value for these models, which is the average inpatient count, is

$$(1-\pi_i)\mu_i$$

We scale the age variable as

$$age o rac{age - 60}{10}$$

for interpretability and for numerical stability.

Effects sizes.

Presently we are not modeling the π using cohort. So we can compare cohorts with the same values of other predictors-whatever they are-simply using the ratios of the means μ_i .

We cannot do across the board comparison of the effect on the probability of zero admissions as it varies depending on the values of the other covariates as well as the zero inflation probability π . But we can compare for given values of the covariates.

We do these comparisons below.

Summary statistics

We consider admissions occurring within 270 day window of observation.

```
# A tibble: 2 x 2
  cohort count
  <chr>     <int>
1 control 881
2 ibis 82
```

Mean inpatient counts

Overall

By cohort

So this is more than 50% reduction.

Proportion if zero admissions

Overall

By cohort

Mean for patients with one or more admissions

Overall

By cohort

Model results

Frequentist zero inflated Poisson model

```
Call:
pscl::zeroinfl(formula = inpatient_count ~ cohort + age + chf + atrial_fibrillation |
   age, data = patients_events, dist = "poisson")
Pearson residuals:
  Min
          1Q Median
                       3Q
                             Max
-0.681 -0.455 -0.402 -0.312 16.120
Count model coefficients (poisson with log link):
                   Estimate Std. Error z value Pr(>|z|)
(Intercept)
                    0.2698
                             0.1113 2.42 0.0153 *
cohortibis
                   -0.7187
                               0.3239
                                       -2.22 0.0265 *
                               0.0608 -4.04 5.4e-05 ***
                   -0.2456
age
                    0.4644
                               0.1443
                                        3.22 0.0013 **
chf
                               0.1447 2.97 0.0030 **
atrial_fibrillation
                    0.4300
Zero-inflation model coefficients (binomial with logit link):
           Estimate Std. Error z value Pr(>|z|)
                                9.14 < 2e-16 ***
(Intercept) 1.3549 0.1482
```

```
-0.4835
                        0.0975 -4.96 7.1e-07 ***
age
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Number of iterations in BFGS optimization: 16
Log-likelihood: -726 on 7 Df
Frequentist zero inflated negative binomial model
Call:
pscl::zeroinfl(formula = inpatient_count ~ cohort + age + chf + atrial_fibrillation |
    age, data = patients_events, dist = "negbin")
Pearson residuals:
   Min
          10 Median
                        3Q
-0.579 -0.398 -0.372 -0.291 15.937
Count model coefficients (negbin with log link):
                   Estimate Std. Error z value Pr(>|z|)
(Intercept)
                    -0.4663
                                0.2364
                                        -1.97 0.04856 *
cohortibis
                    -0.6441
                                0.3477
                                        -1.85 0.06398 .
                                0.0913 -3.15 0.00165 **
age
                    -0.2874
chf
                     0.7563
                                0.2015
                                         3.75 0.00017 ***
atrial_fibrillation
                     0.7098
                                         3.57 0.00036 ***
                                0.1990
Log(theta)
                    -0.3909
                                0.3327
                                        -1.17 0.24010
Zero-inflation model coefficients (binomial with logit link):
           Estimate Std. Error z value Pr(>|z|)
              0.308
                                          0.44
(Intercept)
                         0.397
                                  0.77
             -0.656
                         0.154
                               -4.25 2.1e-05 ***
age
___
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Theta = 0.676
Number of iterations in BFGS optimization: 16
Log-likelihood: -698 on 8 Df
```

Bayesian zero inflated Poisson model

Estimate Est.Error Q2.5 Q97.5

Intercept	0.285	0.1094	0.0685	0.49130
zi_Intercept	1.336	0.1382	1.0700	1.61402
cohortibis	-0.511	0.2656	-1.0393	0.00693
age	-0.209	0.0601	-0.3253	-0.09031
chf	0.413	0.1372	0.1452	0.68345
atrial_fibrillation	0.376	0.1397	0.1004	0.64854
zi age	-0.387	0.0845	-0.5544	-0.22165

Model fit and predictions

We can see that these models are in agreement, and the results do fit the data well.

Expected number of inpatient admissions; probabilities of zero admissions

If we take the mean of the cohortibis count coefficients for the two models as an estimate, the models suggest a mean admission count for ibis cohort is 0.54 times that of the control, assuming the same values for other predictors, which amounts to a 46 percent reduction.

zero admissions probabilities

We can compare prevalences for outcomes in the data vs those predicted by the model. This was done in the zer0_infl_admissions_models_2025-05_short.html. We give a graphical illustration below with the Bayes model. Presently, we will compare what the models give as probabilities of admissions count outcomes for a single patient with the same age and conditions covariates, with one being ibis cohort and the other control.

For 60 year old patient, without no chf or afib, for ibis vs control,On the other hand, for the probability of zero admissions, the zero inflated Poisson model gives,

```
cohort age chf atrial_fibrillation 0 1 2 3 4
1 control 0 0 0 0.850 0.0725 0.0475 0.02072 0.006785
2 ibis 0 0 0 0.903 0.0691 0.0221 0.00469 0.000749
5
1 1.78e-03
2 9.56e-05
```

while for patients with both chronic conditions we have

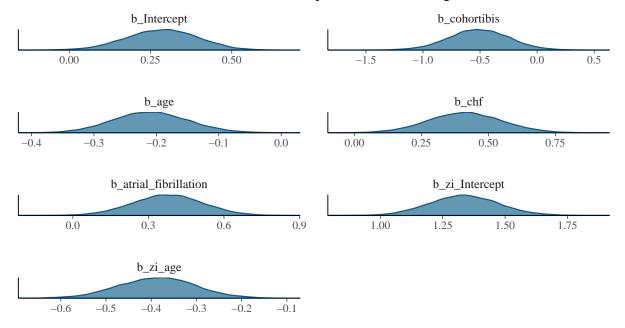
```
cohort age chf atrial_fibrillation
                                                    1
                                                            2
                                                                   3
                                             0
1 control
                                      0 0.763 0.0942 0.0768 0.0417 0.01700
            1
                 1
                 1
                                      0 0.838 0.1059 0.0421 0.0111 0.00221
2
     ibis
            1
         5
1 0.005541
2 0.000351
```

In either case we get a little under 6-10% increase. This agrees with the difference in proportion of zero admits for ibis vs control above.

Post predictive checks

We use Bayes model here. The coefficients are similar, but we can get a more complete picture of both the distributions of the coefficients, and the how well the model reproduces the observations, as well as model variability.

Posterior distributions for zero-inflated model parameters; scaled age



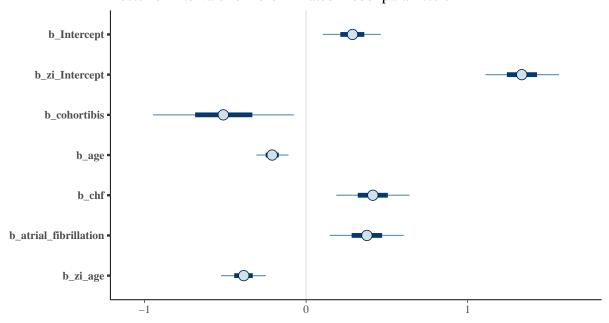
Credible intervals

These are 50% and 90% "credible intervals".

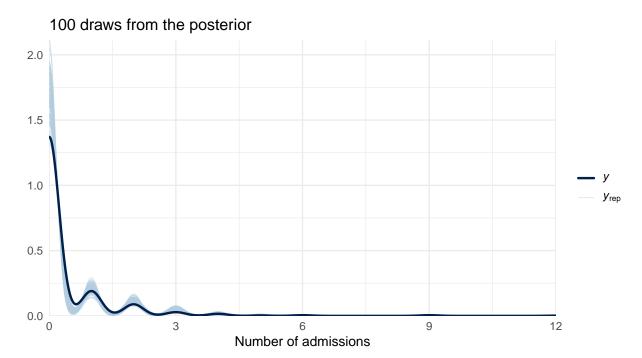
Remark These are the middle quantiles in the distribution for the various coefficients. We can interpret this as a probability the coefficient lies within the interval, according to the model. By contrast, a frequentist confidence interval does not admit a probabilistic interpretation this

way. In particular, we can say that the probability the true parameter is less than the upper bound of a 90% credible interval is 95%, because of the area in the left tail. In that sense, a 90% credible interval is comparable to the upper limit of a 95% frequentist confidence interval, if, as in this case, we are concerned with that upper limit being greater than zero or not.

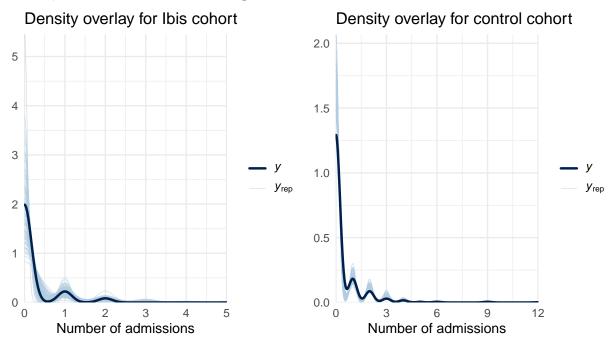
Posterior intervals for zero-inflated model parameters



If anything, the model may inflate the zeros too much.

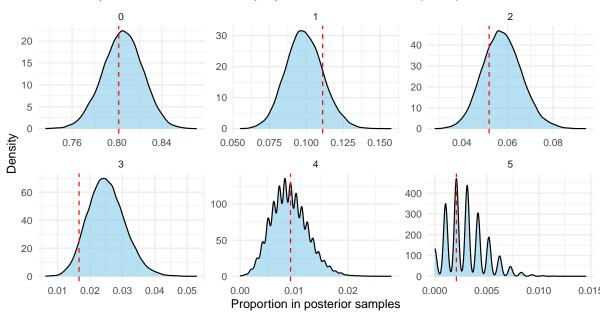


We plot posterior draws by cohort. Not surprisingly there is more uncertainty with the Ibis cohort, which has much smaller sample size.



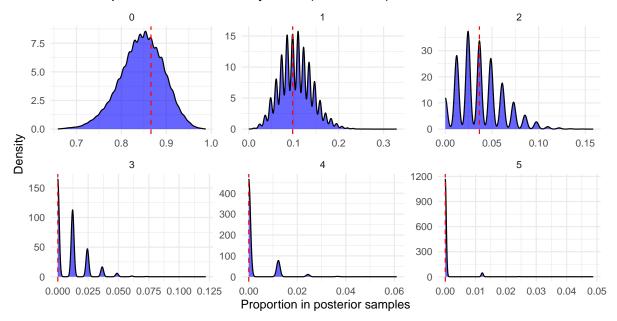
We can plot proportions of various counts in the posterior and compare with actual proportions in the data.

Posterior predictive distribution of proportion for each count (0-10)



And also do this by cohort. There is not enough data

Posterior predictive distribution by count (IBIS cohort)



Posterior predictive distribution by count (Control cohort)

