ARIMA models: Introduction

EC 361-001

Prof. Santetti Spring 2024

Materials

Required readings:

- Hyndman & Athanasopoulos, ch. 9
 - sections 9.1—9.4.

Motivation

Motivation

Along with exponential smoothing, **ARIMA models** are the most **widely used** approach for time series forecasting.

That said, these two techniques should not be seen as **competing**, but as **complementary** to each other.

One **key distinction** between these two approaches is that, while exponential smoothing focuses on time series **features** (*error*, *trend*, and *seasonality*), ARIMA models focus on the presence of **autocorrelation** in the data.

We will start studying ARIMA models by looking at time series **stationarity**.

In short:

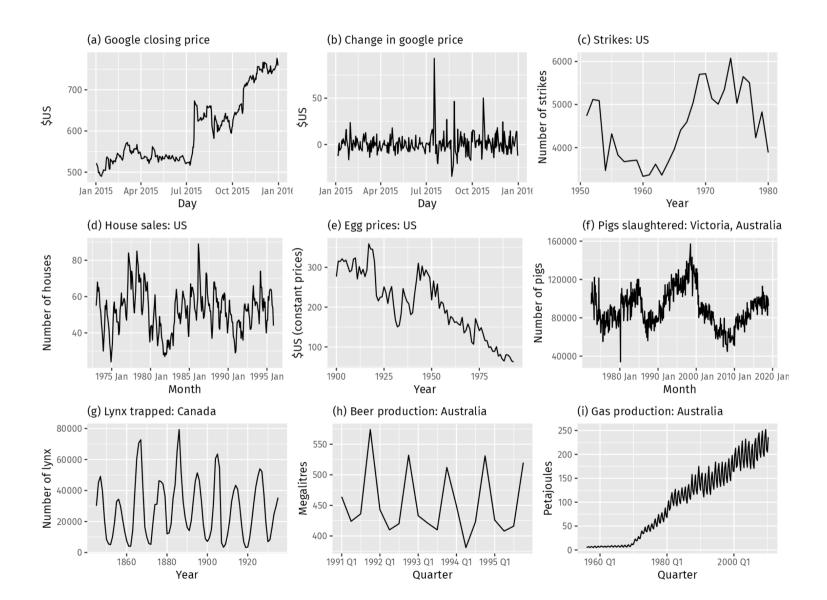
A **stationary** time series is one whose statistical properties **do not** depend on the time at which the series is observed.

In other words, a time series that contains relevant features—such as *trend* and *seasonality*—will not be **stationary**, since its behavior changes over time.

On the other hand, a time series following a **white noise** process is stationary.

Whenever you observe the series, it will look pretty much the same.

In **visual** terms, the *time plot* of a stationary time series will be roughly **horizontal** (which can include *cyclic* behavior) and with a *constant variance*.



Recall that one **key** characteristic of a stationary time series is **constant variance**.

A few weeks ago, we learned that log-transformations help to stabilize a series' variance over time.

But in order for a time series to be stationary, its **mean** should also be stable over time (i.e., with no trend).

A technique that helps to stabilize the **mean** of a time series is known as **differencing**.

A **differenced series** is the change between consecutive observations in the original series, and can be written as

$$y_t^\prime = y_t - y_{t-1}$$

When the differenced series is **white noise**, the model for the original series can be written as:

$$y_t - y_{t-1} = arepsilon_t$$

where ε_t denotes a white noise process (i.e., with no autocorrelation).

Rearranging the above equation, we end up with

$$y_t = y_{t-1} + arepsilon_t$$

which, unsurprisingly, is a **random walk process**.

$$y_t = y_{t-1} + arepsilon_t$$

Random walk models are widely used for **non-stationary** data, as they typically show:

- long periods of apparent *trends* up or down;
- sudden and unpredictable changes in direction.

As future movements in a random walk process are **unpredictable**, forecasts from these processes are usually equal to the **last observation**.

This fact generates the naïve forecast method.

Whenever the **differences** between y_t and y_{t-1} are, on average, different from zero, one can account for this in the following way:

$$y_t - y_{t-1} = c + arepsilon_t$$

where c accounts for the possibility of y_t drifting upwards or downwards.

This model is behind the **drift forecast method**.

When a time series shows **seasonality**, one option is to work with **seasonal differencing**.

What this implies is that we can take the difference between an observation y_t and the *previous* observation from the **same season**:

$$y_t^\prime = y_t - y_{t-m}$$

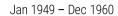
where m denotes the number of seasonal periods.

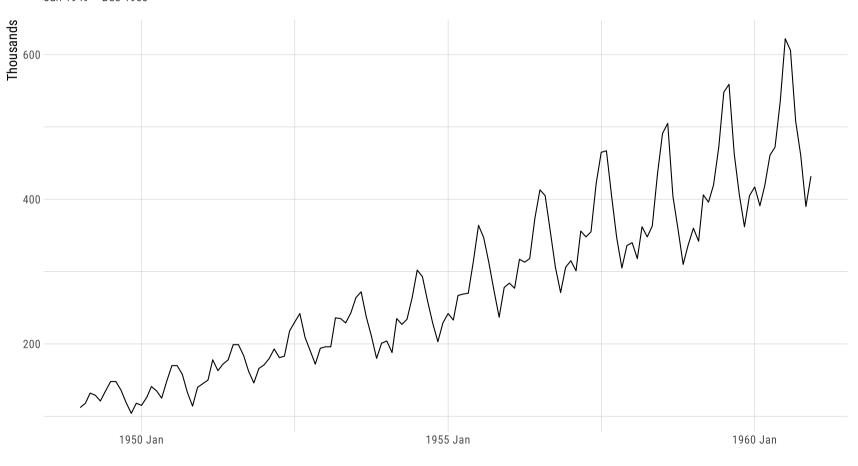
In case the differenced data from seasonal differencing is **white noise**, then one can write:

$$y_t = y_{t-m} + arepsilon_t$$

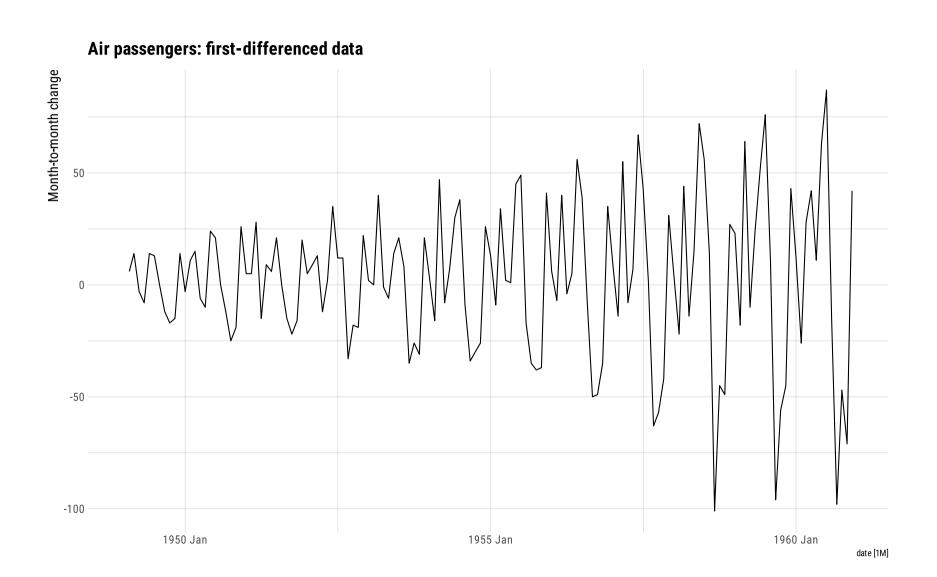
And this model gives origin to the **seasonal naïve forecast method**.

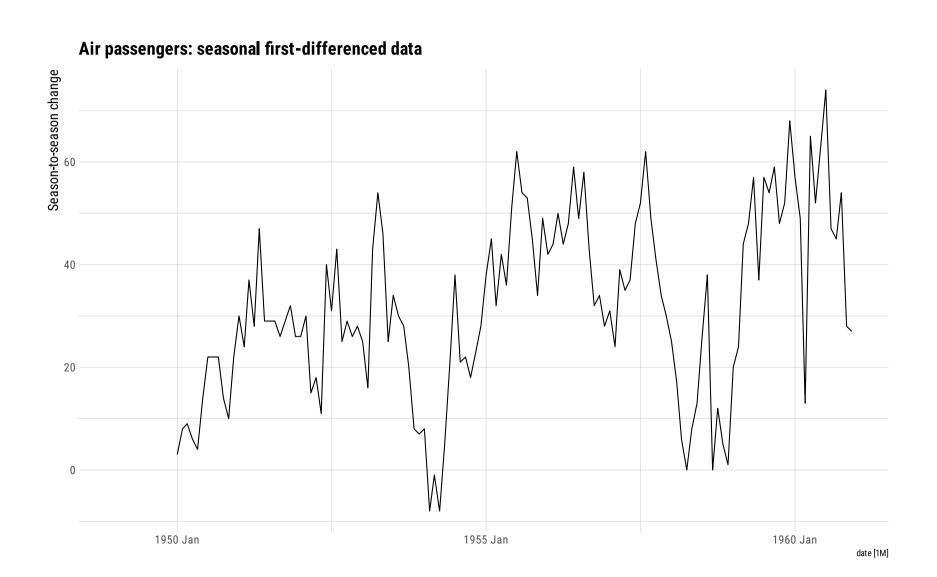
International airline passengers

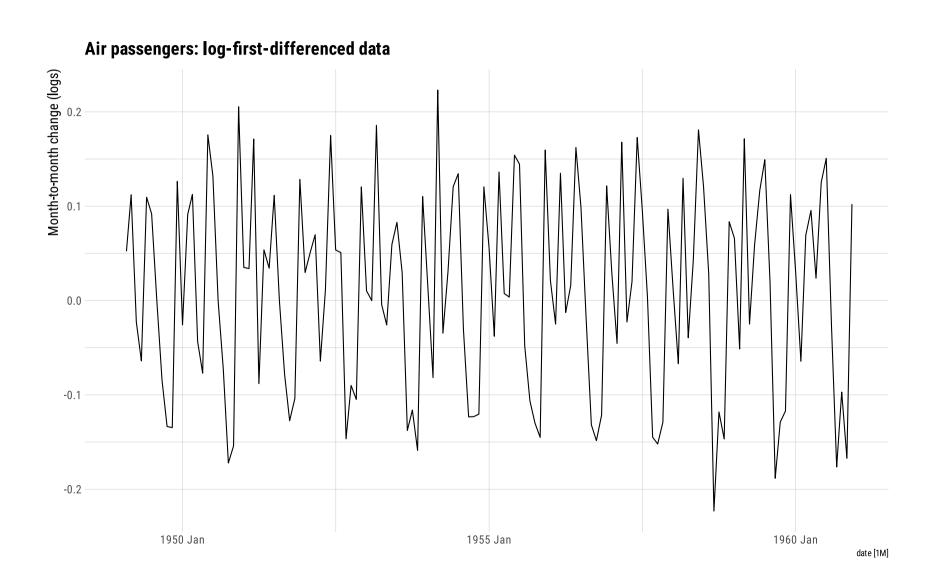


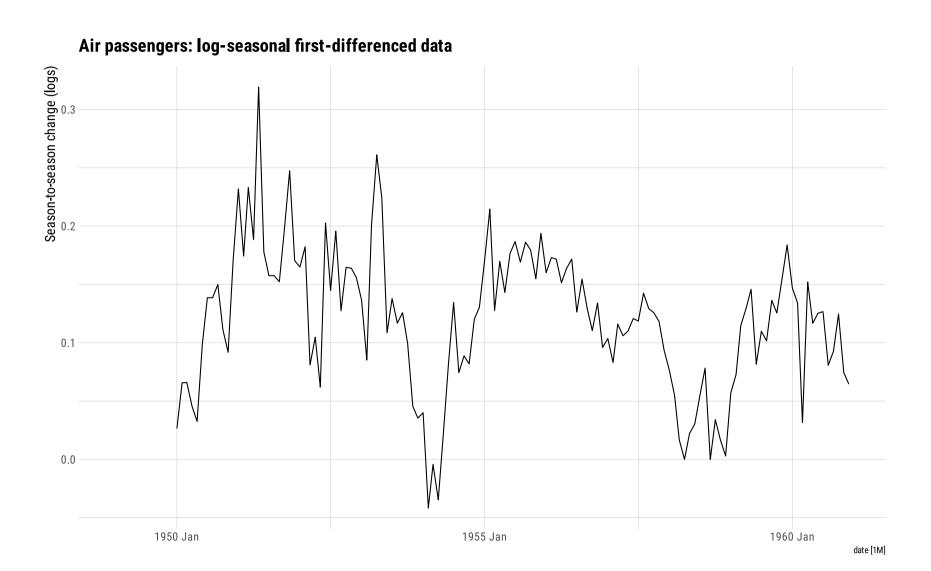


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air ts ▷
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          seas diff passengers = difference(passengers, lag = 12),
          diff_log_passengers = difference(log(passengers), lag = 1),
          seas diff log passengers = difference(log(passengers), lag = 12))
#> # A tsibble: 144 x 6 [1M]
#>
          date passengers diff passengers seas diff passengers diff log passengers seas diff log passengers
         <mth>
                    <dbl>
                                     <dbl>
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   1 1949 Jan
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   2 1949 Feb
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                                         6
   3 1949 Mar
                      132
                                                                              0.112
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#> # i 134 more rows
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When applying differencing techniques to our data, there is no definitive approach.

There is always a **degree of subjectivity** when choosing the best way of dealing with the data we have at hand.

What has to be kept in mind, though, is interpretability.

- First differences are the change between one observation and the next;
- Seasonal differences are the change between one year to the next;
- Other lags are unlikely to make much interpretable sense and should be avoided.

Unit root tests

Unit root tests

Don't ever undermine your subjectivity/gut feeling when it comes to analyzing time series!

That said, a way to be more objective when it comes to determining whether a time series is stationary or not is by using **unit root testing**.

Unit root tests are statistical **hypothesis tests** of stationarity that are designed for determining whether differencing is required.

One of these tests is the Kwiatkowski-Phillips-Schmidt-Shin (KPSS) test.

Its **null hypothesis** states that the data are *stationary*, and in case the latter is rejected, we have evidence in favor of **differencing** our data.

Unit root tests

#> # A tibble: 1 × 2

#> 1 0.0146

#>

0.1

features(difference(passengers, lag = 1), unitroot_kpss)

```
air_ts D
features(passengers, unitroot_kpss)

#> # A tibble: 1 × 2
#> kpss_stat kpss_pvalue
#> <dbl> <dbl>
#> 1 2.74 0.01
air_ts D
```

Backshift notation

Backshift notation

As we will progressively explore autocorrelations in our time series, working with **lagged values** will become more necessary.

To make this easier in terms of **notation**, we can introduce the **backshift operator**, B:

$$By_t = y_{t-1}$$

In words, B shifts back the data one period in time.

In case we want to shift back the series two periods, we may write:

$$B^2y_t=y_{t-2}$$

Backshift notation

For **differencing** purposes, the backward shift operator can be used as follows:

$$y_t' = y_t - y_{t-1} = y_t - By_t = (1 - B)y_t$$

In case we use **second differencing**, we can write:

$$y_t^{\prime\prime}=(1-B)^2y_t$$

In general, a **dth-order difference** can be written as

$$y_t' = (1 - B)^d y_t$$

Autoregressive models

Autoregressive models

When we estimate a **regression model**, our aim is to explain the behavior of a *dependent* variable in terms of a **linear combination** of *independent* (*predictor*) variables.

When these predictor variables are past values of the variable of interest, we have the following:

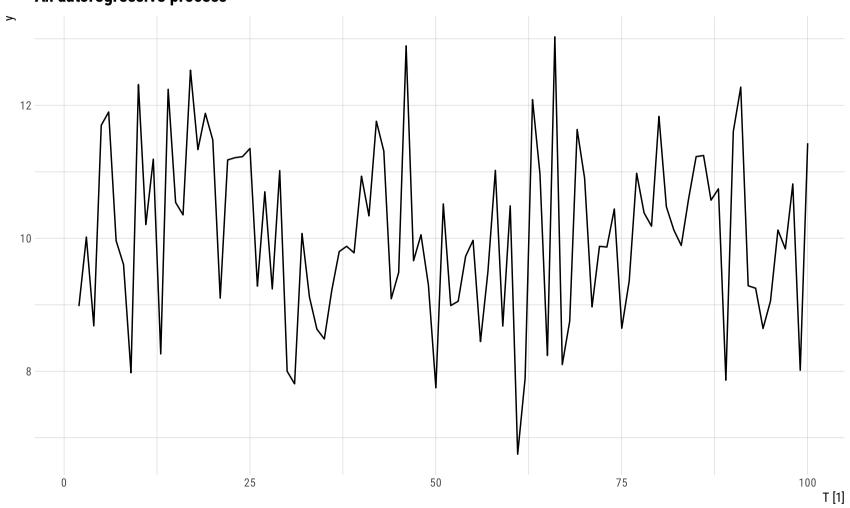
$$y_t = c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \phi_3 y_{t-3} + \dots + \phi_p y_{t-p} + \varepsilon_t$$

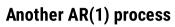
where ε_t is white noise.

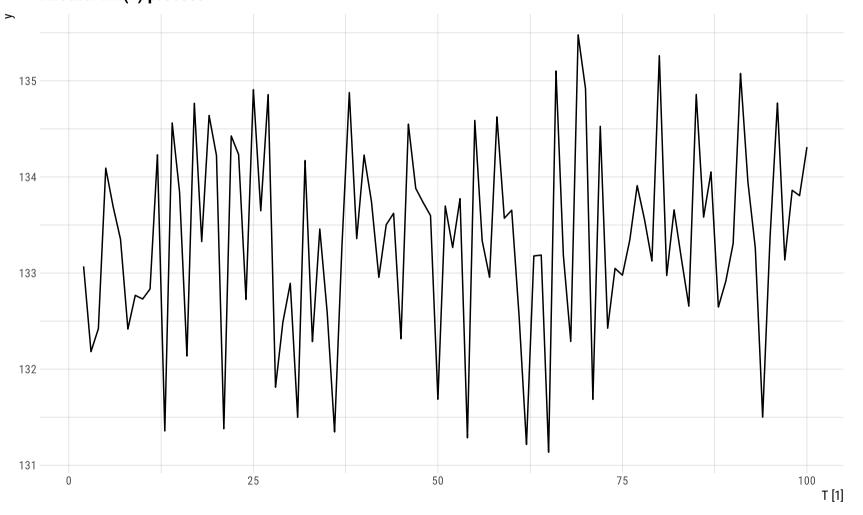
The above model is known as an **autoregressive model** of order *p*.

• In short, we may write this as an AR(p) model.

An autoregressive process

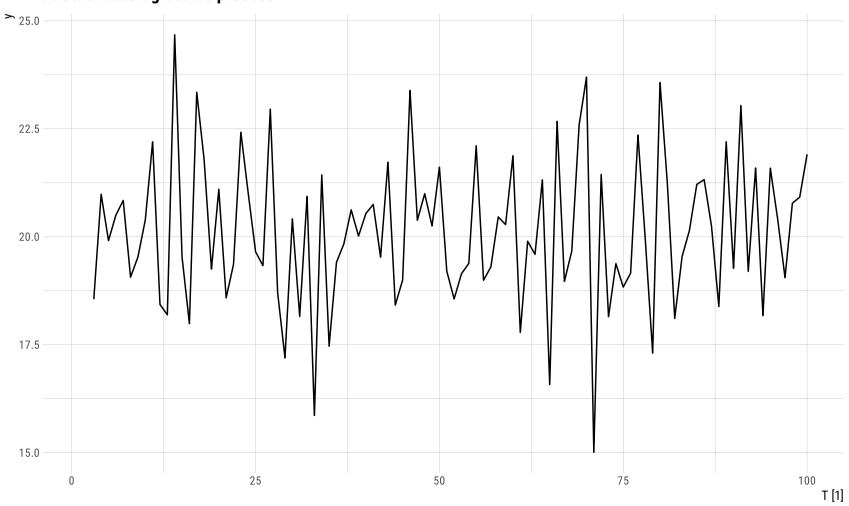






$$y_t = 8 + 1.3y_{t-1} - 0.7y_{t-2} + \varepsilon_t$$

Another autoregressive process



Autoregressive models

A few **remarks**:

- For an **AR(1)** model: $y_t = c + \phi_1 y_{t-1} + \epsilon_t$:
 - \circ If $\phi_1=0$ and c=0, y_t is white noise;
 - \circ If $\phi_1=1$, y_t follows a random walk process;
 - \circ If $\phi_1=1$ and c
 eq 0, y_t is follows a random walk with drift.

Moving average models

Moving average models

A **Moving Average (MA) process** of order *q* can be expressed as follows:

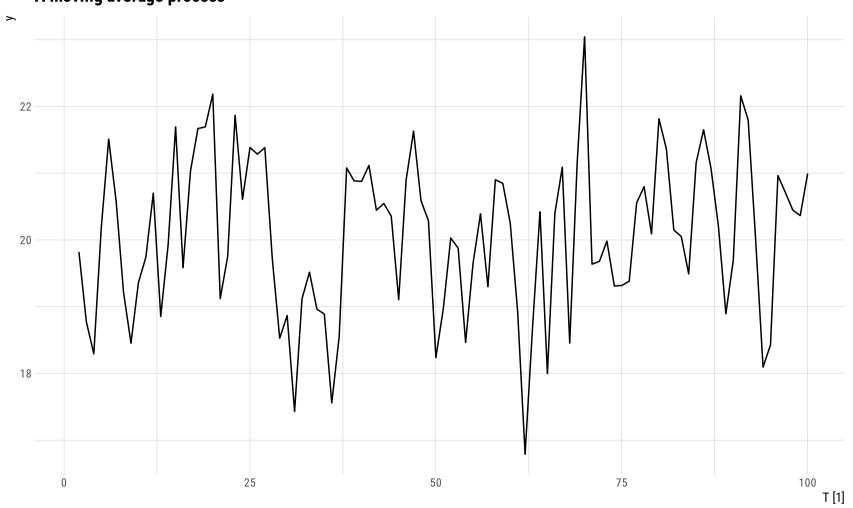
$$y_t = c + arepsilon_t + heta_1 arepsilon_{t-1} + heta_2 arepsilon_{t-2} + heta_3 arepsilon_{t-3} + \cdots + heta_q arepsilon_{t-q}$$

where ε_t is white noise.

In words, a moving average model uses a linear combination of **past forecast errors** to explain the current behavior of a time series y_t .

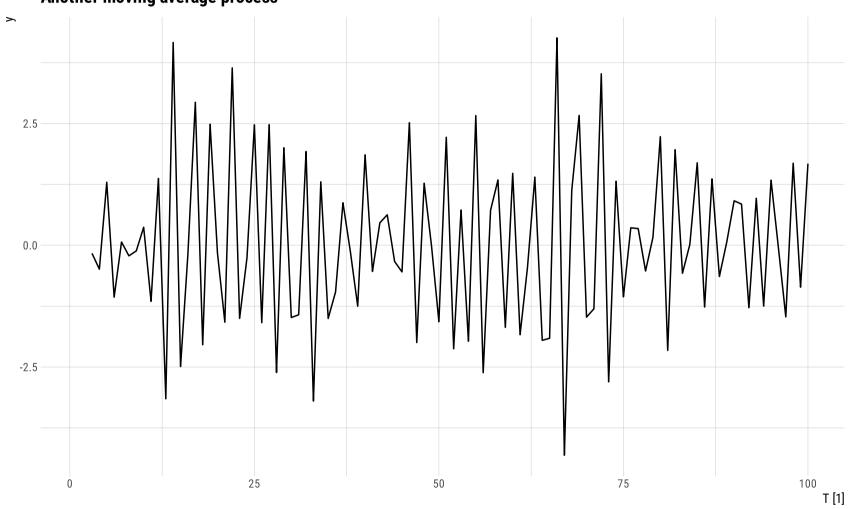
$$y_t = 20 + arepsilon_t + 0.8arepsilon_{t-1}$$

A moving average process



$$y_t = arepsilon_t - arepsilon_{t-1} + 0.8arepsilon_{t-2}$$

Another moving average process



Next time: More on ARIMA models