

# Econometrics II

## Lecture 4: Experiments

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# Literature

- 1 **"Mostly Harmless Econometrics"**, Angrist and Pischke  
Chapter 2 [introduction]; Chapter 3.2.3 [bad controls ]
- 2 **"Causal Inference for Statistics, Social and Biomedical Sciences: An Introduction"**, Imbens and Rubin  
Chapter 4 [introduction]; Chapter 5 [Fisher inference]; Chapter 7.5 [controls]

All mistakes are mine.

# Plan for Today

- 1 Unbiased Estimation
- 2 Balance
- 3 Stratification/Paired Experiments
- 4 Power
- 5 Control/Bad Control
- 6 Attrition
- 7 Canonical Experimental Designs

# Unbiased Estimation

Why randomize?

⇒ Balanced distribution of potential outcomes.

Randomized experiment guarantees, by design, *ex ante*:

$$\mathbb{E}[Y_i(0)|D_i = 1] - \mathbb{E}[Y_i(0)|D_i = 0] = 0$$

Therefore:

$$\begin{aligned} & \mathbb{E}[Y_i(1)|D_i = 1] - \mathbb{E}[Y_i(0)|D_i = 0] \\ &= \mathbb{E}[Y_i(1)|D_i = 1] - \mathbb{E}[Y_i(0)|D_i = 1] + \mathbb{E}[Y_i(0)|D_i = 1] - \mathbb{E}[Y_i(0)|D_i = 0] \\ &= \mathbb{E}[Y_i(1) - Y_i(0)|D_i = 1] \equiv ATT \end{aligned}$$

and in fact:

$$= \mathbb{E}[Y_i(1) - Y_i(0)] \equiv ATE(\text{'average treatment effect'})$$

# How to Randomize?

## Why should we think about how to randomise?

Purely random treatment assignment:

- 1 Suboptimal fractions of T/C.
  - 1 Ex-ante random assignment: unnecessarily low **power**.
- 2 Imbalanced distribution of potential outcomes across T/C.
  - 1 Ex-post random assignment: problematic **causal inference**.
  - 2 Ex-ante random assignment: unnecessarily low **power**.

How should we randomize optimally then?

# Set Fraction of T and C

**Assignment Mechanism:** Randomisation conditional on  $N_1$  and  $N_0$ .

Under this assignment mechanism:  $V(\hat{\beta}|N_0, N_1) = \frac{\sigma_1^2}{N_1} + \frac{\sigma_0^2}{N-N_1}$ .

How to minimise  $V(\hat{\beta})$ ?

With  $\sigma_0^2 = \sigma_1^2$ :  $N_1^* = N/2$ .

With  $\sigma_0^2 \neq \sigma_1^2$ : More observations for noisy outcome.

**Budget & Costs:** With costs  $c_0$  and  $c_1$ , and a fixed budget  $B$ ,  
then  $\min_{N, N_1} V(\hat{\beta})$  s.s.  $(N - N_1)c_0 + N_1c_1 \leq B$ .  
More observations for cheap outcome.

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# Balance Tests: How to judge imbalances ex-post?

In practice, treatment arms are unlikely “balanced”.

Imbalances in potential outcomes show up in covariates!

## How to detect imbalances?

- Often t-test for each covariate shown.
  - Conceptually problematic.
  - Statistical significance is not what matters.<sup>1</sup>
- What is useful then?
  - Focus on **size of differences for covariates that impact outcome!**<sup>2</sup>
  - Estimate size of difference using same specification as for differences in outcomes.
  - To check randomisation was implemented correctly, use omnibus test.

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<sup>1</sup>Altman (1985) notes that such tests amount to assessing the probability of something having occurred by chance when you know that it did occur by chance. “Such a procedure is clearly absurd”.

<sup>2</sup>Imbens and Rubin, 2015

# Forcing Balance

Without further information:

Randomisation conditional on  $N_1$  is best we can do  
to achieve  $(Y(0), Y(1)) \perp D$  in sample.

But generally have more information:

$$\mathbb{E}[Y_i(D_i)] = f(X_i^+, X_i^-, D_i) \quad (1)$$

where  $X_i^+$  are observable and  $X_i^-$  non-observable covariates.

Cannot force balance of potential outcomes, but...  
...can try to force balance of  $X_i^+$ . Reduce  $V(\hat{\beta})$ .

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# Stratification: Simple Example

**Idea:** Do not leave imbalance of important covariates to chance.

**Formally:** Restrict assignment mechanisms.

Example 1: Potential outcomes determined as  $Y_i(D_i) = g_i + 1 \times D_i$ , where  $g_i \in \{1, 2\}$  is only covariate,  $N = 4$ ,  $N_1 = 2$  and  $i = 1, 2$  only have  $g_i = 1$ . Note: true  $\beta = 1$ .

Assignment	$d_1$	$d_2$	$d_3$	$d_4$	$\hat{\beta}$
#1	0	1	0	1	1
#2	0	1	1	0	1
#3	1	0	0	1	1
#4	1	0	1	0	1
#5	1	1	0	0	0
#6	0	0	1	1	2

$\Rightarrow V(\hat{\beta})$  lower by excluding last two assignments! **Stratification.**

# Stratification: Practicalities

Extends to several categorical variables. **How?**

⇒ In Problem Set 2 you are asked to simulate effect of stratification on the variance of the estimator.

## Implementation in STATA:

```
set seed 20230323
gen random = runiform()
sort cat_var1 cat_var2 ... random
gen treatment = mod(_n,2)
```

# Stratification: Practicalities

**Question 1:** Which variables should we stratify on?

General recommendation: covariates strongly related to the outcome.<sup>3</sup> Why?

**Question 2:** And when covariates are continuous, or cells sparse?

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<sup>3</sup>See Bruhn and McKenzie, 2009; Glennerster and Takavarasha, 2013.

# Stratification: Implementations

## What do people do? (Bruhn and McKenzie, 2009)

- 1 Pure Randomisation.
- 2 Re-Randomisation.
  - Subjectively decide whether to make another draw; or
  - Re-randomise until some statistic of balance is achieved: or
  - Choose assignment with best balance amongst N draws.
- 3 Matched-Pair ('blocking'): stratified randomisation with 2 units in each stratum.

### Recent insight:

Matched-Pair designs optimal (under some conditions, in some sense).

# Stratification: Matched-Pair Designs

**Matched-Pair Design** is stratified randomisation with two units in each stratum.

① Bai (AER, 2022):

*“Optimality of Matched-Pair Designs in Randomized Controlled Trials”*

- Shows optimality (in MSE sense) of a specific matched-pair design: Calculate  $\mathbb{E}[Y_i(1)|X_i] + \mathbb{E}[Y_i(0)|X_i]$ , match adjacent units on that “simple” scalar function.
- Problem: we do not know that sum! Can be estimated using pilot data with large class of estimators, including machine learning techniques.

② Barrios (2015) – special case:

*“Optimal Stratification in Randomized Experiments”*

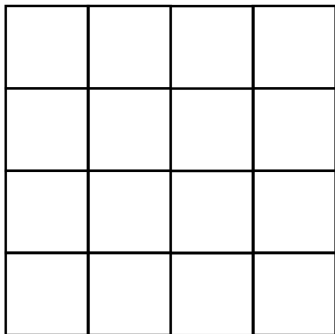
- With homogeneous treatment effects, best way to choose matched-pairs is to match on  $\mathbb{E}[Y_i(0)|X_i]$ .
- Baseline, but no pilot data needed.



# Stratification: Taking it to the Extreme

**Next level:** No randomization.

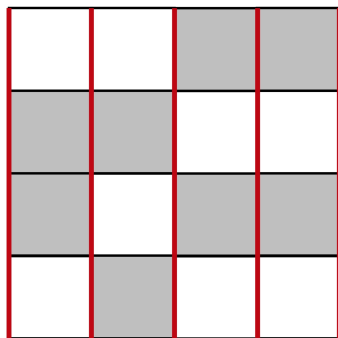
Consider this example, a 'field'. *What is optimal?*



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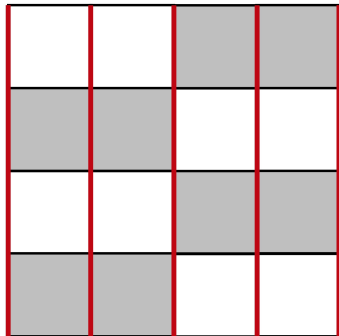


Stratification by  $x_1$ ? Potentially unbalanced marginal distribution of  $x_2$ .

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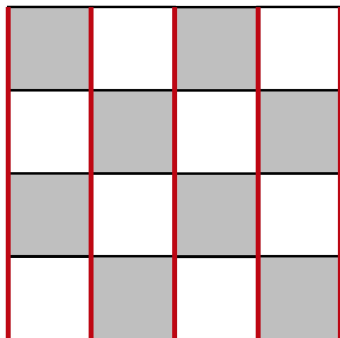


Something like this? Not maximizing power.

# Stratification: Taking it to the Extreme

**Next level:** No randomization.

Consider this example, a 'field'. *What is optimal?*



Balance joint distribution of  $x_1, x_2$ .

But only two assignments achieve that! Randomization?

# Stratification: Taking it to the Extreme

*Given any prior on how potential outcomes are generated...*

...argument is very general! (Kasy, 2016)

- For any assignment  $\mathbf{D}$ , calculate loss function (MSE, V, ...).
- Expected loss is average across potential assignments.
- Find subset of assignments such that loss is minimised.
- Typically those are two.

Controlled Trial, not Randomized Controlled Trial

- Stratification special case where subset larger than 2.

# Benefits

**Notice:** Stratification has benefits ex-post, and ex-ante!

**After Kasy:** *Why do we randomize then?*

- Optimal solution might be hard to find.
- Might set up the decision problem differently.
- Randomization Inference!

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# Power: The Concept

Talked about experimental design choices that lead to lower variance of the estimator. Closely related concept: **Power**.

- ① Eventually want to test hypothesis  $H_0$  vs alternative  $H_1$ .
- ② Two types of errors we can make:
  - ① Type I error: reject  $H_0$  when  $H_0$  is true.  
Probability of type I error chosen by setting  $\alpha$ .
  - ② Type II error: fail to reject  $H_0$  when  $H_0$  is false.  
Probability of type II error depends on true effect size, experimental set-up and estimator/test statistics..



# Power: How to increase it?

*How to increase power?*

- 1 Reduce the variance of the estimator:
  - Stratification (see before).
  - Baseline Controls (see later).
  - Sample Size.<sup>4</sup>
  - Measurement!
- 2 Choice of test statistic.

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<sup>4</sup>When starting a project, make sure it is 'powered' to detect expected effects with reasonably high probability. Funders want to see this in grant applications.

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# Analysis: Control Variables

- Control variables: pre-determined, observed variables.
- If simple random assignment: no need to include control to guarantee **unbiasedness**...
- ... but might make sense to increase **power**.
- If data is generated by *conditional random assignment*, need to hold variables you conditioned on in randomisation constant in the analysis. (Only *conditional* independence assumption will surely hold, will come back to this.)

# Analysis: Bad Control

Control variables can not be outcomes themselves!

Mathematical reason:

$$\mathbb{E}[Y_i(0)|D_i = 1] = \mathbb{E}[Y_i(0)|D_i = 0]$$

does not imply

$$\mathbb{E}[Y_i(0)|D_i = 1, X_i = x] = \mathbb{E}[Y_i(0)|D_i = 0, X_i = x]$$

where  $X_i$  refers to the (observed) outcomes for the covariate for unit  $i$ . Such an  $X_i$  would be a **bad control**.

# Analysis: Bad Control

Example: Think of  $Y$  as achievement,  $D$  as class size (0 if large/1 if small), and  $X$  as parental help (1 if help/0 if not).  $D_i$  assigned by classical randomized experiment. Plausibly parents' help responds to variation in  $D$ , so  $X$  is an outcome. Define  $X_i$  as potential outcome analogous to how we defined  $Y_i$ .

Suppose 'control for  $X$ ' in difference in outcomes:

$$\begin{aligned} & \mathbb{E}[Y_i | D_i = 1, X_i = 1] - \mathbb{E}[Y_i | D_i = 0, X_i = 1] \\ &= \mathbb{E}[Y_i(1) | D_i = 1, X_i = 1] - \mathbb{E}[Y_i(0) | D_i = 0, X_i = 1] \\ &= \mathbb{E}[Y_i(1) | X_i(1) = 1] - \mathbb{E}[Y_i(0) | X_i(0) = 1] \quad | \text{ by Random Assignment} \\ &= \underbrace{\mathbb{E}[Y_i(1) - Y_i(0) | X_i(1) = 1]}_{\text{'some' causal effect}} + \underbrace{\mathbb{E}[Y_i(0) | X_i(1) = 1] - \mathbb{E}[Y_i(0) | X_i(0) = 1]}_{\text{selection bias}} \end{aligned}$$

## Analysis: Bad Control

$$\underbrace{\mathbb{E}[Y_i(1) - Y_i(0) | X_i(1) = 1]}_{\text{'some' causal effect}} + \underbrace{\mathbb{E}[Y_i(0) | X_i(1) = 1] - \mathbb{E}[Y_i(0) | X_i(0) = 1]}_{\text{selection bias}}$$

**Causal effect:** for population that helps when class size is small.

**Compositional bias:** control sample *shouds be* those that help when class size is small, but *it is* those that help when class size is large - who are probably more than those who help when size is small. Probably special group!

Same argument applies *generally* when restricting to subgroup with some outcome.  
Example: effect of training on wages – which are only observed for employed.

Extensive margin: easy; intensive margin: infeasible.

# Analysis: Estimate only reduced form of effects

Example: Educational production function  $Y = f(D, X, \theta)$ . Suppose that  $X$  responds to  $D$ .

Comparing outcomes across treatment arms, we estimate:

$$\Delta Y = \frac{\partial f}{\partial D} \Delta D + \frac{\partial f}{\partial X} \frac{\partial X}{\partial D} \Delta D \quad (2)$$

- Sometimes that is what you are interested in.
- But for other questions you might need  $\frac{\partial f}{\partial D}$  or  $\frac{\partial f}{\partial X}$ .

P. Fredriksson: "One reason for the slight dismay in certain quarters over the experimental approach."

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# Attrition: Practical Advise

**Attrition** is common in experiments:

cannot obtain follow-up data for some treated and some un-treated observations.

- Generally good to avoid!
- Unproblematic when attrition is unrelated to potential outcomes.

Unfortunately impossible to know!

- Commonly accepted test: attrition rates unrelated to treatment.

Might still be that attrition is selected differently in T and C.

- When attrition rates differ across T and C:

Some bounding exercise, commonly [Lee \(2009\) bounds](#).

## Attrition: Lee (2009) Bounds

Suppose the attrition rate is higher in C than in T.

**Worry:** Those with 'best'  $Y_i(0)$  cannot be observed in Control. Comparison of T with C exaggerates ATE.

**Lee Bounds** are an extreme bounding exercise:

- Drop the 'best' from T, s.t. T and C have same attr. rate, and rerun analysis.
- Drop the 'worst' from T, s.t. T and C have same attr. rate, and rerun analysis.
- Report results from both exercises as bounds.

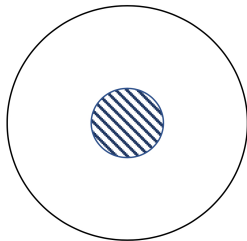
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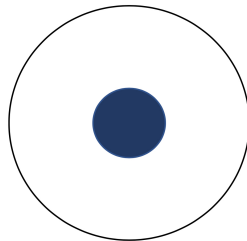
## Spill-Overs: Miguel and Kremer (2004)

Suppose you believe your treatment might affect other units, i.e. has **spill-over effects**. *How can you estimate those?*

Village A

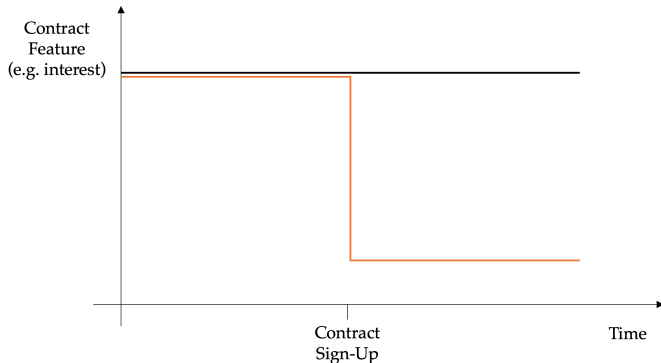


Village B



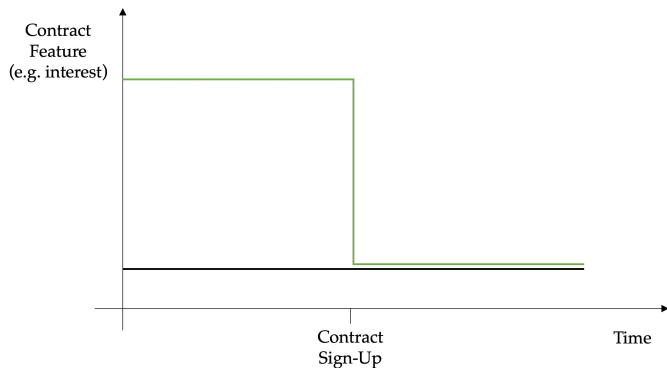
# Moral Hazard and Adverse Selection: *Karlan and Zinman (2009)*

Distinguishing **Moral Hazard** and **Adverse Selection** *empirically* is generally hard. Karlan and Zinman figured out one way:



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# Randomized Experiments

- Experiments important to understand on their own right.

Today just a small introduction.

- Modern econometric approaches approximate that ideal.
- The experimental ideal will often be very useful to think in *observational studies* whether you estimate causal effects.
- Issues we discussed will come up, more or less explicitly: assignment mechanism, conditional independence, balance, biases, reduced form estimation, average treatment effects for subpopulations...

Questions?



# References

- 1 Angrist and Pischke (2008): Chapter 2 [introduction]; Chapter 3.2.3 [bad controls]
- 2 Imbens and Rubin (2015): Chapter 4 [introduction]; Chapter 5 [Fisher inference], Chapter 7.5 [controls]
- 3 Kasy, Maximilian (2016) "Why Experimenters Might Not Always Want to Randomize, and What They Could Do Instead" Political Analysis: 1-15. [stratification]
- 4 Lee, D. S. (2009) "Training, Wages, and Sample Selection: Estimating Sharp Bounds on Treatment Effects". Review of Economic Studies 76: 1071-1102. [attrition]
- 5 Bai, Yuehao (2022) "Optimality of Matched-Pair Designs in Randomized Controlled Trials". Working Paper. [matched-pair designs]
- 6 Barrios, Thomas (2014) "Optimal Stratification in Randomized Experiments". Working Paper. [matched-pair design]
- 7 Bruhn, M. and D. McKenzie (2009) "In Pursuit of Balance: Randomization in Practice in Development Field Experiments". American Economic Journal: Applied Economics 1:4, 200-232.
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- 9 Miguel and Kremer (2004) "Worms: Identifying Impacts on Education and Health in the Presence of Treatment Externalities". 72:1, 159-217.