

# Bootstrap and Subsampling

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Applied Econometrics II

# Bootstrap

- ▶ Bootstrap takes a different approach.
  - Instead of estimating  $\hat{\theta}$  and then using a first-order Taylor Approximation...
  - What if we directly tried to construct the **sampling distribution** of  $\hat{\theta}$ ?
- ▶ Our data  $(X_1, \dots, X_n) \sim P$  are drawn from some measure  $P$ 
  - We can form a **nonparametric estimate**  $\hat{P}$  by just assuming that each  $X_i$  has weight  $\frac{1}{n}$ .
  - We can then simulate a new sample  $X^* = (X_1^*, \dots, X_n^*) \sim \hat{P}$ .
    - Easy: we take our data and construct  $n$  observations by **sampling with replacement**
  - Compute whatever statistic of  $X^*$ ,  $S(X^*)$  we would like.
    - Could be the OLS coefficients  $\beta_1^*, \dots, \beta_k^*$ .
    - Or some function  $\beta_1^*/\beta_2^*$ .
    - Or something really complicated: estimate parameters of a game  $\hat{\theta}^*$  and now find Nash Equilibrium of the game  $S(X^*, \hat{\theta}^*)$  changes.
  - Do this  $B$  times and calculate at  $Var(S_b)$  or  $CI(S_1, \dots, S_b)$ .

## Bootstrap: Bias Correction

The main idea is that  $\hat{\theta}^{1*}, \dots, \hat{\theta}^{B*}$  approximates the **sampling distribution** of  $\hat{\theta}$ . There are lots of things we can do now:

- We already saw how to calculate  $\text{Var}(\hat{\theta}^{1*}, \dots, \hat{\theta}^{B*})$ .

$$\frac{1}{B-1} \sum_{b=1}^B (\hat{\theta}_{(b)}^* - \bar{\theta}^*)^2$$

- Calculate  $E(\hat{\theta}_{(1)}^*, \dots, \hat{\theta}_{(B)}^*) = \bar{\theta}^* = \frac{1}{B} \sum_{b=1}^B \hat{\theta}_{(b)}^*$ .

## Bootstrap: Bias Correction

- We can use the estimated bias to **bias correct** our estimates

$$\begin{aligned} \text{Bias}(\hat{\theta}) &= \mathbb{E}[\hat{\theta}] - \theta \\ \text{Bias}_{bs}(\hat{\theta}) &= \overline{\theta^*} - \hat{\theta} \end{aligned}$$

Recall  $\theta = \mathbb{E}[\hat{\theta}] - \text{Bias}[\hat{\theta}]$ :

$$\hat{\theta} - \text{Bias}_{bs}(\hat{\theta}) = \hat{\theta} - (\overline{\theta^*} - \hat{\theta}) = 2\hat{\theta} - \overline{\theta^*}$$

- Correcting bias isn't for free - variance tradeoff!
- Linear models are (hopefully) unbiased, but most nonlinear models are **consistent but biased**.

## Bootstrap: Confidence Intervals

There are actually three ways to construct bootstrap CI's:

1. Obvious way: sort  $\hat{\theta}^*$  then take  $CI: [\hat{\theta}_{\alpha/2}^*, \hat{\theta}_{1-\alpha/2}^*]$ .
2. Asymptotic Normal:  $CI: \hat{\theta} \pm 1.96\sqrt{V(\hat{\theta}^*)}$ . (CLT).
3. Better Way: let  $W = \hat{\theta} - \theta$ . If we knew the distribution of  $W$  then:  $Pr(w_{1-\alpha/2} \leq W \leq w_{\alpha/2})$ :

$$CI: [\hat{\theta} - w_{1-\alpha/2}, \hat{\theta} - w_{\alpha/2}]$$

We can estimate with  $W^* = \hat{\theta}^* - \hat{\theta}$ .

$$CI: [\hat{\theta} - w_{1-\alpha/2}^*, \hat{\theta} - w_{\alpha/2}^*] = [2\hat{\theta} - \theta_{1-\alpha/2}^*, 2\hat{\theta} - \theta_{\alpha/2}^*]$$

Why is this preferred? Bias Correction!

## Bootstrap: Why do people like it?

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- ▶ Econometricians like the bootstrap because under certain conditions it is **higher order efficient** for the confidence interval construction (but not the standard errors).
  - Intuition: because it is non-parametric it is able to deal with more than just the first term in the Taylor Expansion (actually an **Edgeworth Expansion**).
  - Higher-order asymptotic theory is best left for real econometricians!
- ▶ Practitioner's like the bootstrap because it is easy.
  - If you can estimate your model once in a reasonable amount of time, then you can construct confidence intervals for most parameters and model predictions.

## Bootstrap: When Does It Fail?

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- ▶ Bootstrap isn't magic. If you are constructing standard errors for something that isn't asymptotically normal, don't expect it to work!
- ▶ The Bootstrap exploits the notion that your sample is IID (by sampling with replacement). If IID does not hold, the bootstrap may fail (but we can sometimes fix it!).
- ▶ Bootstrap depends on asymptotic theory. In small samples weird things can happen. We need  $\hat{P}$  to be a good approximation to the true  $P$  (nothing missing).

## Bootstrap: Variants

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The bootstrap I have presented is sometimes known as the **nonparametric bootstrap** and is the most common one.

**Parametric Bootstrap** ex: if  $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$  then we can estimate  $(\hat{\beta}_0, \hat{\beta}_1)$  via OLS.

Now we can generate a bootstrap sample by drawing an  $x_i$  at random with replacement  $\hat{\beta}_0 + \hat{\beta}_1$  and then drawing **independently** from the distribution of estimated residuals  $\hat{\epsilon}_i$ .

**Wild Bootstrap** Similar to parametric bootstrap but we rescale  $\epsilon_i$  to allow for **heteroskedasticity**

**Block Bootstrap** For correlated data (e.g.: time series). Blocks can be overlapping or not.



## Bootstrap vs Delta Method

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- ▶ Delta Method works best when working out Jacobian  $D(\theta)$  is easy and statistic is well approximated with a linear function (not too curvy).
- ▶ I would almost always advise Bootstrap unless:
  - Delta method is trivial e.g.:  $\beta_1/\beta_2$  in linear regression.
  - Computing model takes many days so that 10,000 repetitions would be impossible.
- ▶ Worst case scenario: rent time on Amazon EC2!
  - I “bought” over \$1,000 of standard errors recently.
- ▶ But neither is magic and both can fail!

The bootstrap has a close cousin **subsampling**.

- ▶ In practice it looks similar, but the underlying theory is quite different.
- ▶ It relies on weaker assumptions and works even in some cases where the bootstrap fails.
- ▶ Again “fails” means that the 95% confidence interval has coverage that isn't very close to 95%.

## Subsampling: How does it work?

1. Draw a **smaller** sample  $X^* = (X_1^*, \dots, X_{a_n}^*)$  **without replacement** of size  $a_n$  where as  $n \rightarrow \infty$  we have  $a_n \rightarrow \infty$  and  $\frac{a_n}{n} \rightarrow 0$ .
  - e.g.  $a_n = \log n$  or  $a_n = \sqrt{n}$ . Note that  $a_n/10$  doesn't work.
2. Compute the relevant statistic  $\theta(X^*)$  or  $g(\theta(X^*))$ .
3. Repeat this  $B$  times and construct the CDF:

$$L_n(t) = \frac{1}{B} \sum_{b=1}^B \mathbf{I} \left( \sqrt{a_n}(\hat{\theta}_b - \hat{\theta}_n) \leq t \right)$$

4. Calculate the quantiles of the CDF and CI:

$$\hat{t}_{\alpha/2} = L_n^{-1}(\alpha/2), \quad \hat{t}_{1-\alpha/2} = L_n^{-1}(1 - \alpha/2)$$

$$C_n = \left[ \hat{\theta}_n - \frac{\hat{t}_{1-\alpha/2}}{\sqrt{n}}, \hat{\theta}_n - \frac{\hat{t}_{\alpha/2}}{\sqrt{n}} \right]$$

## Subsampling: Caveats

- ▶ The proof for why subsampling works is complicated. See <https://web.stanford.edu/~doubleh/lecturenotes/lecture13.pdf>.
- ▶ Downsides:
  - Subsampling really leans on  $n \rightarrow \infty$  more than bootstrap. People often use bootstrap to understand finite sample performance (is this a good idea though?).
  - Choice of  $a_n$  is difficult. Calculating the optimal value can be quite complicated and there aren't great rules of thumb.
- ▶ But if you're in a weird case where bootstrap fails (parameter on the boundary, etc.) try subsampling and see!

**Thanks!**

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