BEE552 Biometry Week 4

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My Learning Journey

Over the last week, I participated in Biometry in the following ways:

- I asked / answered **5** questions posed in class.
- I asked **0** questions in Slack.
- I answered **2** questions posed by other students on Slack.
- I came to Heather's office hours: **Yes**
- I came to Jose's office hours: **No**
- I met with Heather or Jose separately from office hours: **No**

Anything not falling into one of the above categories?

> **No**

On a scale of 1 (no knowledge) to 10 (complete expert), how would I rate my comfort with R programming after this week?

> **6**

Any topics from last week that you are still confused about?

> **Doing fine for now**

Problem Set

Part I

We will be using a recently published dataset on ancient rings and ribs that may have been used as early forms of money. These data come from the paper "The origins of money: Calculation of similarity indexes demonstrates the earliest development of commodity money in prehistoric Central Europe" by M.H.G. Kuijpers and C. Popa (PLoS ONE, January 20, 2021).

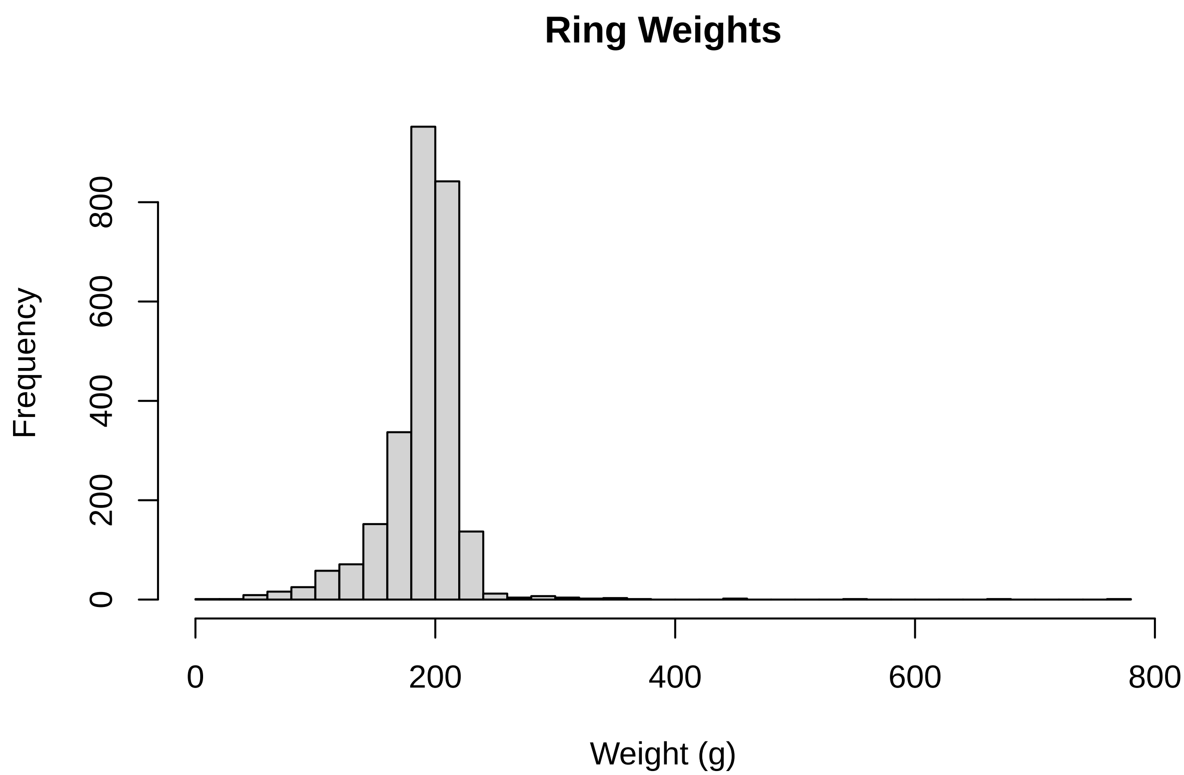
Download the data in the file "rings_and_ribs.csv".

```r
money <- read_excel("journal.pone.0240462.s002.xlsx",
                    sheet = 1
                    )

# Rename the weight column to make it tidy
money <- rename(money, Weight = `Weight (g)`)
```

Table 1: A selection of data from Kuijpers and Popa 2021.

Location	Country	Zone	Date	Weight	Type	Museum	Source
Carsdorf (Pegau)	Germany	2	EBA	201.0	Ring	Leipzig	Lenerz- de Wilde documentation
Obereching	Austria	1	EBA	202.0	Rib	Salzburg	Moolsteiner and Maoesta 1988
Unknown (Linz)	Austria	1	EBA	84.0	Rib	Linz	Lenerz- de Wilde documentation
Waging am See	Germany	1	EBA	101.0	Rib	Munich	Lenerz- de Wilde documentation
St. Pölten - Spratzern	Austria	1	EBA	231.0	Ring	Vienna	Lenerz- de Wilde documentation
Radostice	Czech Republic	1	EBA	175.0	Ring	Ceské Budejovice	Moucha 2005
Schrobenhausen-Freinhausen	Germany	1	EBA	201.0	Ring	Schrobenhausen	Lenerz- de Wilde documentation
Znojmo	Czech Republic	1	EBA	197.0	Ring	Znojmo	Lenerz- de Wilde documentation
München-Luitpoldpark	Germany	1	EBA	180.0	Rib	Munich	Lenerz- de Wilde documentation
Mikulov	Czech Republic	1	EBA	210.0	Ring	Brno	Lenerz- de Wilde documentation
Obereching	Austria	1	EBA	200.0	Rib	Salzburg	Moolsteiner and Maoesta 1988
Harmannsdorf	Austria	1	EBA	114.0	Rib	Vienna	Lenerz- de Wilde documentation
München-Luitpoldpark	Germany	1	EBA	178.0	Rib	Munich	Lenerz- de Wilde documentation
Obereching	Austria	1	EBA	205.5	Rib	Salzburg	Moolsteiner and Maoesta 1988
Bubenec	Czech Republic	1	EBA	159.0	Ring	Prague 2	Moucha 2005
Pilszcz	Poland	2	EBA	210.0	Ring	Wroclaw	Lenerz- de Wilde documentation
Obereching	Austria	1	EBA	204.0	Rib	Salzburg	Moolsteiner and Maoesta 1988
Kilb	Austria	1	EBA	183.0	Ring	Asparn	Lauermann and Pernicka 2013; Lenerz- de Wilde documentation
Mauthausen	Germany	1	EBA	203.5	Ring	Bad Reichenhall	Lenerz- de Wilde documentation; Menke 1987/1979
Mauthausen	Germany	1	EBA	151.0	Ring	Bad Reichenhall	Lenerz- de Wilde documentation; Menke 1987/1979
München-Luitpoldpark	Germany	1	EBA	152.0	Rib	Munich	Lenerz- de Wilde documentation
Mauthausen	Germany	1	EBA	192.5	Ring	Bad Reichenhall	Lenerz- de Wilde documentation; Menke 1987/1979
Schrobenhausen-Freinhausen	Germany	1	EBA	184.0	Ring	Schrobenhausen	Lenerz- de Wilde documentation
Obereching	Austria	1	EBA	215.0	Rib	Salzburg	Moolsteiner and Maoesta 1988
Amselfing	Germany	1	EBA	152.0	Ring	Straubing	Lenerz- de Wilde documentation
Suchohrdly	Czech Republic	1	EBA	179.0	Ring	Brno	Lenerz- de Wilde documentation
München-Luitpoldpark	Germany	1	EBA	173.0	Rib	Munich	Lenerz- de Wilde documentation
Wildendürnbach	Austria	1	EBA	181.0	Ring	Asparn	Lenerz- de Wilde documentation
Bernhaupten	Germany	1	EBA	168.0	Ring	Munich	Lenerz- de Wilde documentation
Gammersham	Germany	1	EBA	257.0	Ring	Munich	Lenerz- de Wilde documentation

Question 1 *For now, let's group all the locations together. Make a histogram of ring weight.*



Question 2 *What is the FORMULA for the sample mean and standard deviation (in other words, what is the formula you would want to use if you wanted to estimate the population mean [the μ parameter assuming a normal distribution] and the population variance [σ^2 if we assume a normal distribution] from a sample that represented a random subset of the entire population)?*

$$\overline{X} = \frac{\sum_{i=1}^{n} X_i}{n}$$

$$S = \sqrt{\frac{\sum_{i=1}^{n}(X_i - \overline{X})^2}{n-1}}$$

Question 3 *What is the population about which we are trying to make inference?*

The population is the total collection of rings and ribs ever used for currency in the context of Germany, Austria, the Czech Republic, and Poland during this time period.

Question 4 *Using R, what are the mean and the standard deviation of ring weight?*

The mean of the ring weights is 172.633. The standard deviation of ring weights is 48.296.

Question 5 *What is the formula for the standard error of the mean (heretofore s.e.)?*

$$SEM = \sqrt{\frac{S^2}{n}}$$

Question 6 *Describe the difference between the s.e. (of the mean) and the s.d.*

The standard deviation is a measure of the variance of your sample The standard error of the mean describes how close your sample mean approximates the population mean.")

Question 7 *Finish the sentence: The standard error is the standard deviation of...*

...the mean.

Question 8 *Use the MASS package's 'fitdistr' function to fit a normal distribution to the ring weight data. Do you get the roughly same answer as above?*

```
ringWeightFit <- fitdistr(money$Weight[money$Type == "Ring"],
                          "normal")

# Calculate difference in calculated and fitdistr() means
meandif <- unname(ringWeightFit$estimate[1]) - ringMean

# Calculate difference in calculated and fitdistr() standard deviations
sddif <- unname(ringWeightFit$estimate[2]) - ringSD
```

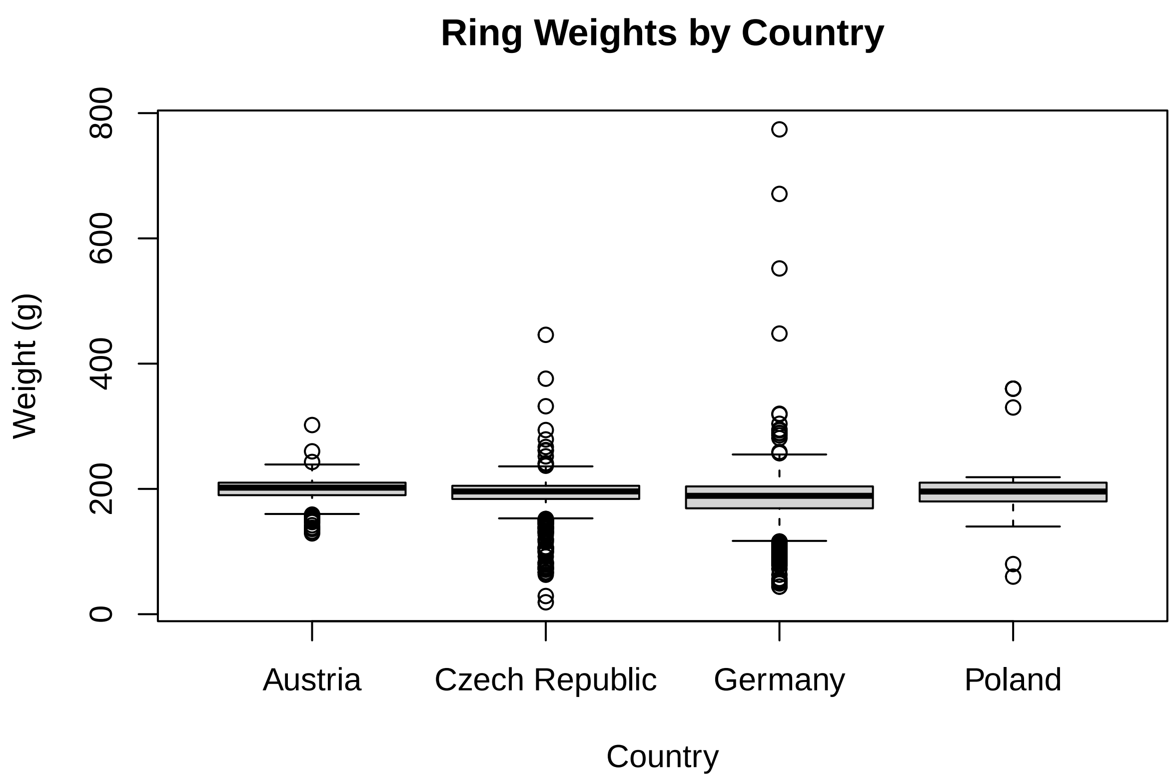
The difference between the calculated mean and the mean from fitdistr() is 0 The difference between the calculated standard deviation and the standard deviation from fitdistr() is -0.00677.

Question 9

Why might it not be valid to group the different locations when summarizing the data?

Though there might be differences between the rings used as currency in each country, all were considered to have monetary value and might have changed hands outside of country/political entity lines. Since they are all from the same time period, if you want to characterize how the earliest money worked, you must use all the material that was used as money.

Question 10 *Using the R command 'boxplot', create a boxplot to compare ring weight across different countries.*



Question 11 *Make the case (in words and/or mathematically) that average ring weight in Germany is or is not statistically different from the average ring weight in Austria.*

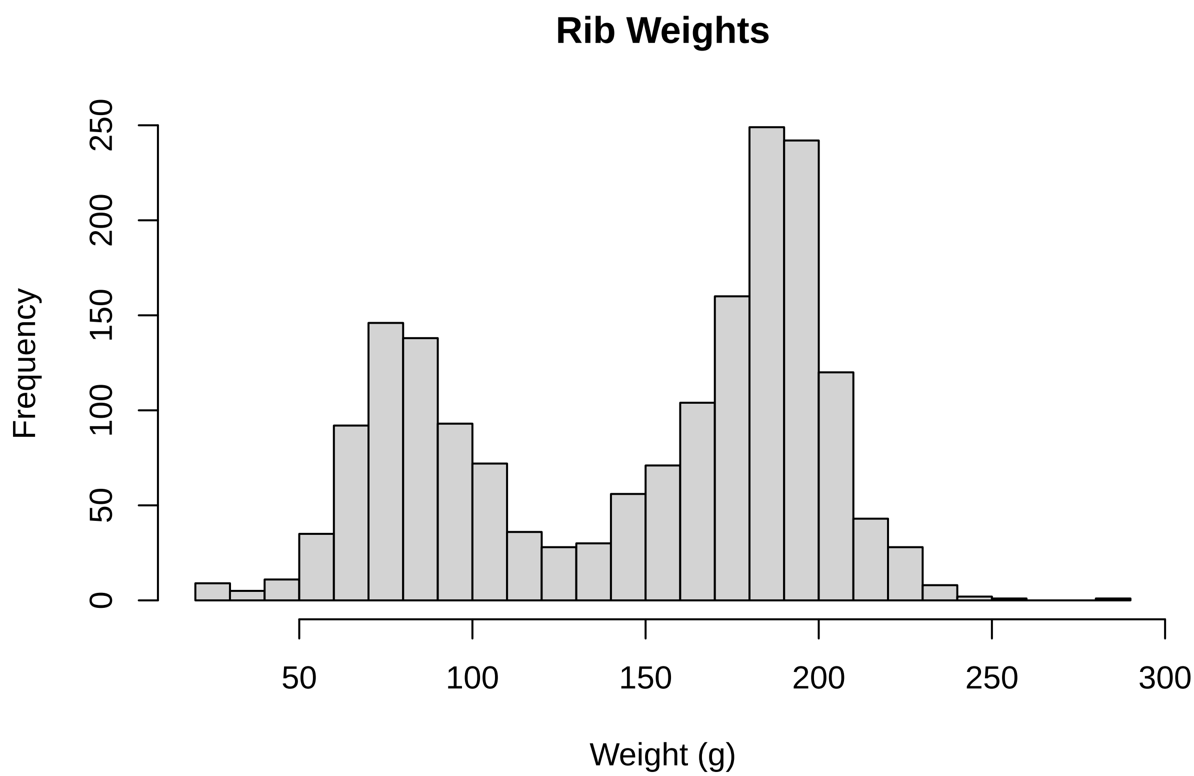
```r
# Calculate and store the mean and standard deviation of German and Austrian
# rings
ringsGerman <- list("mean" = mean(money$Weight[money$Country == "Germany"
                                               & money$Type == "Ring"]
                                  ),
                    "sd" = sd(money$Weight[money$Country == "Germany"
                                           & money$Type == "Ring"]
                              )
                    )

ringsAustrian <- list("mean" = mean(money$Weight[money$Country == "Austria"
                                                 & money$Type == "Ring"]
                                    ),
                      "sd" = sd(money$Weight[money$Country == "Austria"
                                             & money$Type == "Ring"]
                                )
                      )
```

Table 2: Summary Statistics of German and Austrian Ring Weights (g)

	Germany	Austria
mean	184.207140238313	198.870917573872
sd	44.097469474065	18.7458110028249

Since the mean weight of Austrian rings is within 1.92 standard deviations (or within the 95% confidence interval) of the German ring mean weight, then I conclude that there is no statistical difference between the German and Austrian ring samples.

Question 12 *So far, we've only been focused on ring weight. Let's go back and make a histogram of rib weight. Why would testing a hypothesis about average rib weight be harder?*



Making predictions or testing hypotheses about rib weight would be difficult because they appear to follow a bimodal distribution.

Part II

Assume an experiment in which the number of plants in 16 experimental plots is counted as:

```
plants <- c(8, 0, 3, 3, 4, 3, 2, 3, 3, 2, 4, 2, 3, 1, 4, 1)
```

We want to model the number of plants in each plot as being distributed according to a Poisson distribution.

Question 1 *Starting with the probability density function for the Poisson, manually derive the maximum likelihood estimate for λ, the parameter for the Poisson distribution.*

$$
\begin{aligned}
f(x \mid \lambda) &= \frac{e^{-\lambda}\lambda^{x}}{x!} \\
f(X_1, X_2 ..., X_n \mid \lambda) &= \prod_{i=1}^{n} \frac{e^{-\lambda}\lambda^{X_i}}{X_i!} \\
L(\lambda \mid X_1, X_2, ..., X_n) &= \prod_{i=1}^{n} \frac{e^{-\lambda}\lambda^{X_i}}{X_i!} \\
LL &= \sum_{i=1}^{n} \left(ln \frac{e^{-\lambda}\lambda^{X_i}}{X_i!} \right) \\
LL &= \sum_{i=1}^{n} (ln\ e^{-\lambda} + ln\ \lambda^{X_i} - ln\ X_i!) \\
LL &= \sum_{i=1}^{n} (-\lambda + X_i\ ln\ \lambda - ln\ X_i!) \\
NLL &= \sum_{i=1}^{n} (\lambda - X_i\ ln\ \lambda + ln\ X_i!) \\
\frac{\partial NLL}{\partial \lambda} &= \sum_{i=1}^{n} \left(1 - \frac{X_i}{\lambda}\right) \\
0 &= \sum_{i=1}^{n} \left(1 - \frac{X_i}{\hat{\lambda}}\right) \\
0 &= n - \sum_{i=1}^{n} \left(\frac{X_i}{\hat{\lambda}}\right) \\
n &= \sum_{i=1}^{n} \left(\frac{X_i}{\hat{\lambda}}\right) \\
n &= \frac{1}{\hat{\lambda}} \sum_{i=1}^{n} X_i \\
\hat{\lambda} &= \frac{\sum_{i=1}^{n} X_i}{n}
\end{aligned}
$$

Question 2 *Write an R function to calculate the negative log-likelihood for this data as described by the Poisson distribution.*

From question 1, we know the negative log-likelihood of a Poisson distribution is $NLL = \sum_{i=1}^{n} (\lambda - X_i\ ln\ \lambda + ln\ X_i!)$

```
neg.ll <- function(x, lambda){
        result <- length(x)*lambda - sum(x)*log(lambda) + sum(log(factorial(x)))
        return(result)
}
```

Question 3 *Using your function for the negative log-likelihood, calculate the MLE for λ and the 95th percentile confidence interval. What would the 99th percentile confidence interval be?*

```
# Define test lambda values
lambdaVals <- seq(min(plants),
                  max(plants),
                  by = 0.05
                  )

# Create vector to catch log likelihood values of test lambdas
plantLogLiks <- rep(0, length(lambdaVals))

# Run neg.ll() over all lambdaVals
for (i in 1:length(lambdaVals)){
        plantLogLiks[i] <- neg.ll(plants, lambdaVals[i])
}

# Use optimize() to determine the minimum value
plantMinLogLik <- optimize(f = neg.ll, x = plants, interval = lambdaVals)

# 95th percentile confidence interval
# Collect the positions of the lambda values whose negative log-likelihoods are
# within 1.92 of the minimum
vals95 <- which(plantLogLiks < plantMinLogLik$objective + 1.92)

# Use the first and last to select the corresponding lambda value
LL95 <- lambdaVals[vals95[1]]
UL95 <- lambdaVals[vals95[length(vals95)]]

# 99th percentile confidence interval
# Collect the positions of the lambda values whose negative log-likelihoods are
# within 3.32 of the minimum
vals99 <- which(plantLogLiks < plantMinLogLik$objective + 3.32)

# Use the first and last to select the corresponding lambda value
LL99 <- lambdaVals[vals99[1]]
UL99 <- lambdaVals[vals99[length(vals99)]]
```

The calculated lambda for the plants per plot is 2.875. The optimized lambda value with the lowest log-likelihood is 2.875, NLL = 30.39.

95$^{\text{th}}$ percentile confidence interval: (2.15, 3.75)

99$^{\text{th}}$ percentile confidence interval: (1.95, 4.1)

Question 4 *Plot the likelihood over a range of parameter values and plot the boundaries of the 95^{th} and 99^{th} percentile confidence interval. Remember: The confidence interval is a range of parameter values, it is NOT the likelihood values itself.*