# ECONOMETRICS IN R LECTURE 4 INTRO TO R PROGRAMMING

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# **REVIEW**

# **AGENDA**

- REGRESSIONS IN R
  - CONTINUOUS VARIABLES
  - CATEGORICAL AND BINARY VARIABLES
- GETTING TO THE RIGHT MODEL
  - VARIABLE TRANSFORMATION
  - FUNCTIONAL FORM
  - CONTROL VARIABLES
  - INTERACTION TERMS

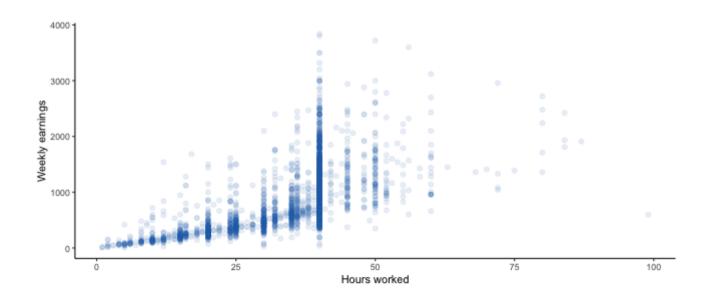
- INFERENCE
- EXPORTING RESULTS
  - REGRESSION TABLES
  - COEFFICIENT PLOTS
- PRACTICING YOUR SKILLS

We are going to work with a portion of the US Current Population Survey for August 2023. Download it from here. We are going to explore some well-known relationships in **labor economics**.

```
cps <- read.csv("../data/04_202308_cps.csv")</pre>
cps10 <- read.csv("../data/04_202308_cps10.csv")
str(cps)
## 'data.frame': 4609 obs. of 13 variables:
             : int 1 2 3 4 5 6 7 8 9 10 ...
##
   $ X
##
   $ region : int 3 3 3 3 3 3 3 3 1 ...
##
   $ state
            : int 111111119 ...
##
  $ age : int 22 50 55 52 59 57 22 38 24 16 ...
  $ sex : chr "Male" "Female" "Male" "Male" ...
##
   S maritl : int
                   6 1 5 5 1 1 6 1 6 6 ...
##
                   "High school" "Bachelor's degree" "No high school" "High school" ...
##
  $ educ : chr
##
   $ race
          : int
                   1 1 1 1 1 1 1 2 2 1 ...
##
   $ status : int 1 1 2 1 1 1 1 1 1 1 ...
            : int 112111111...
##
   $ work
   $ hours : int 40 30 40 40 40 40 20 30 40 12 ...
##
   $ hourrt : num 14 42 20 27.6 40 ...
##
##
   $ earnings: num 560 1260 800 1103 1600 ...
```

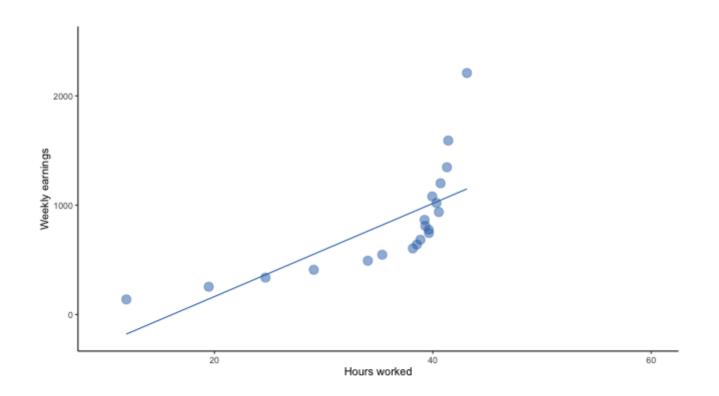
Let's take a look at the relation between hours worked and earnings (measured for a weekly period).

#### How can we best summarize this relationship?



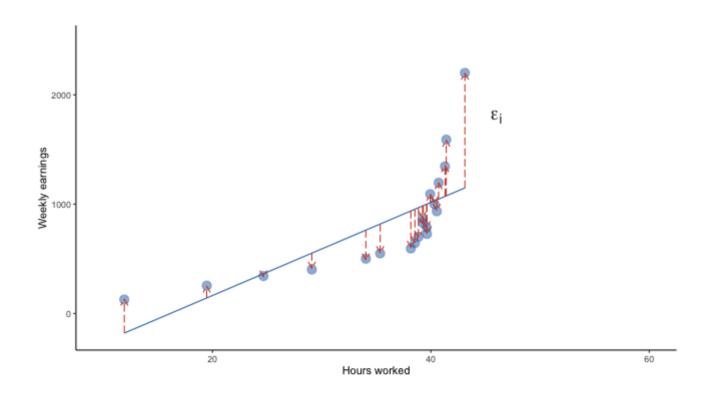
You are already familiar with **univariate regressions**  $y=lpha+eta x+\epsilon$  .

ullet We are looking for the line  $\hat{y}_i=\hat{lpha}+\hat{eta}x_i$  that **minimizes the distance** to the data points



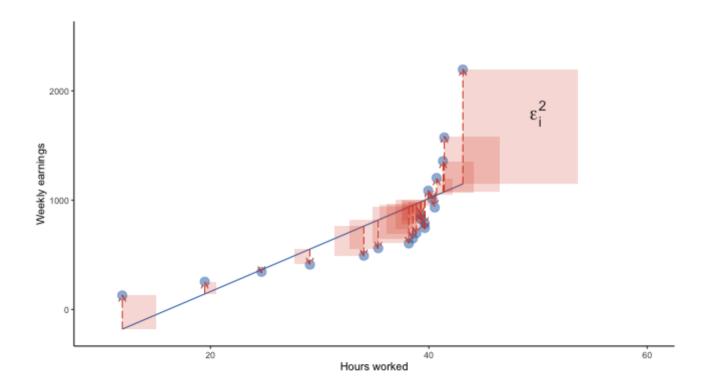
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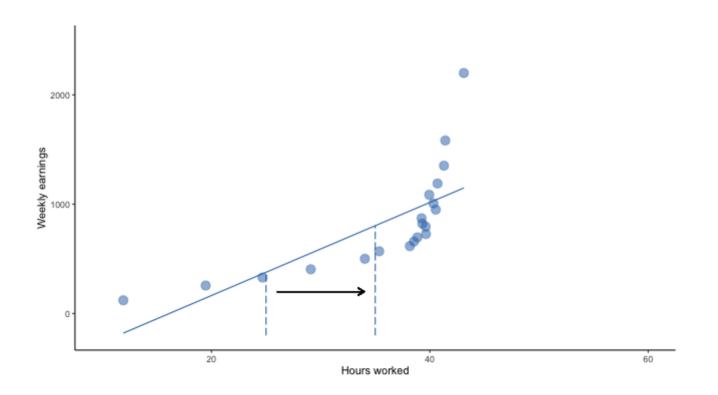
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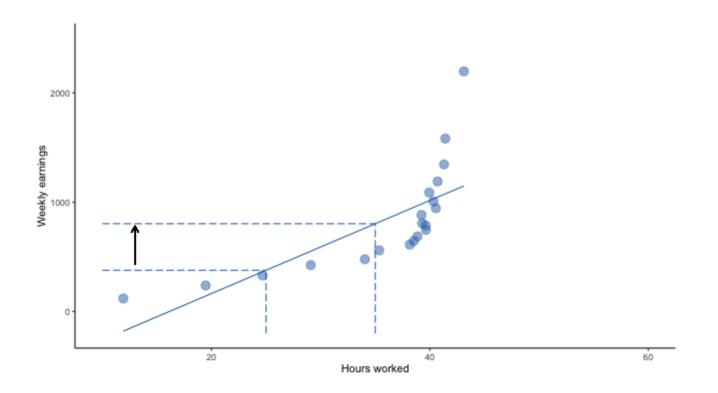
You are already familiar with **univariate regressions**  $y=\alpha+\beta x+\epsilon$ .

• We are looking for the line  $\hat{y}_i = \hat{\alpha} + \hat{\beta} x_i$  that minimizes the distance to the data points:  $\sum \epsilon_i^2$  o Such that for a one unit increase in x,



You are already familiar with **univariate regressions**  $y=lpha+eta x+\epsilon$  .

• We are looking for the line  $\hat{y}_i = \hat{\alpha} + \hat{\beta} x_i$  that **minimizes the distance** to the data points:  $\sum \epsilon_i^2$  o Such that for a **one unit increase** in x,  $\hat{\beta}$  indicates the associated **expected change** in y



In R we can estimate a regression model using the lm() command (Linear Model), which takes the arguments:

- $\bullet$  Formula. Written as  $y \sim x$
- Data

```
lm(earnings ~ hours, cps)

##

## Call:
## lm(formula = earnings ~ hours, data = cps)
##

## Coefficients:
## (Intercept) hours
## -198.81 28.91
```

• To get a more complete description of our regression, we use summary() on the regression object

```
mod <- lm(earnings ~ hours, cps)</pre>
summary(mod)
##
## Call:
## lm(formula = earnings ~ hours, data = cps)
##
## Residuals:
       Min 1Q Median
##
                                 30
                                        Max
## -2066.65 -237.53 -86.26 121.36 2882.47
##
## Coefficients:
##
              Estimate Std. Error t value
                                                  Pr(>|t|)
## (Intercept) -198.8092 21.9152 -9.072 <0.0000000000000000 ***
## hours
         28.9085 0.5911 48.904 <0.0000000000000002 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 393.3 on 4607 degrees of freedom
## Multiple R-squared: 0.3417, Adjusted R-squared: 0.3416
## F-statistic: 2392 on 1 and 4607 DF, p-value: < 0.00000000000000022
```

. . . .

• We get the **command** we used, including the **formula** 

```
mod <- lm(earnings ~ hours, cps)
summary(mod)
##
## Call:
## lm(formula = earnings ~ hours, data = cps)</pre>
```

• A description of the distribution of residuals

```
mod <- lm(earnings ~ hours, cps)
summary(mod)

##
## Call:
## lm(formula = earnings ~ hours, data = cps)
##
## Residuals:
## Min 1Q Median 3Q Max
## -2066.65 -237.53 -86.26 121.36 2882.47
....</pre>
```

• Coefficients and their standard error, t-value, and p-value

```
mod <- lm(earnings ~ hours, cps)</pre>
summary(mod)
##
## Call:
## lm(formula = earnings ~ hours, data = cps)
##
## Residuals:
##
       Min
                1Q Median
                                  30
                                         Max
## -2066.65 -237.53 -86.26 121.36 2882.47
##
## Coefficients:
##
               Estimate Std. Error t value
                                                   Pr(>|t|)
## (Intercept) -198.8092 21.9152 -9.072 <0.0000000000000000 ***
## hours
         28.9085 0.5911 48.904 < 0.0000000000000000 ***
```

• Significance thresholds and their symbols

```
mod <- lm(earnings ~ hours, cps)</pre>
summary(mod)
##
## Call:
## lm(formula = earnings ~ hours, data = cps)
##
## Residuals:
##
       Min 1Q Median
                                 30
                                         Max
## -2066.65 -237.53 -86.26 121.36 2882.47
##
## Coefficients:
##
              Estimate Std. Error t value
                                                   Pr(>|t|)
## (Intercept) -198.8092 21.9152 -9.072 <0.0000000000000000 ***
## hours
         28.9085 0.5911 48.904 < 0.0000000000000000 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
. . . .
```

ullet The **residual standard error**  $\sqrt{\sum (y_i - \hat{y}_i)^2/\mathrm{df}}$  and their symbols

```
mod <- lm(earnings ~ hours, cps)</pre>
summary(mod)
##
## Call:
## lm(formula = earnings ~ hours, data = cps)
##
## Residuals:
           1Q Median
##
       Min
                                30
                                        Max
## -2066.65 -237.53 -86.26 121.36 2882.47
##
## Coefficients:
##
     Estimate Std. Error t value Pr(>|t|)
## (Intercept) -198.8092 21.9152 -9.072 <0.0000000000000000 ***
## hours
         28.9085 0.5911 48.904 < 0.0000000000000000 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 393.3 on 4607 degrees of freedom
. . . .
```

ullet The  $R^2$  and adjusted  $R^2$ 

```
mod <- lm(earnings ~ hours, cps)</pre>
summary(mod)
##
## Call:
## lm(formula = earnings ~ hours, data = cps)
##
## Residuals:
      Min 1Q Median
##
                             30
                                    Max
## -2066.65 -237.53 -86.26 121.36 2882.47
##
## Coefficients:
##
             Estimate Std. Error t value
                                            Pr(>|t|)
## hours
        28.9085 0.5911 48.904 < 0.0000000000000000 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 393.3 on 4607 degrees of freedom
## Multiple R-squared: 0.3417, Adjusted R-squared: 0.3416
. . . .
```

ullet The results of an F-test (  $H_0:eta_k=0 orall k$  )

```
mod <- lm(earnings ~ hours, cps)
summary(mod)
##</pre>
```

```
## Call:
        Formula
                    ## lm(formula = earnings ~ hours, data = cps)
                    ##
                    ## Residuals:
       Residuals
                    ##
                          Min
                                  10 Median
                                                 30
                                                        Max
     distribution
                    ## -2066.65 -237.53 -86.26 121.36 2882.47
                    ##
                    ## Coefficients:
     Coefficients
                                 Estimate Std. Error t value
                    ##
                                                                 Pr(>|t|)
                    ## (Intercept) -198.8092 21.9152 -9.072 <0.0000000000000000 ***
                    ## hours
                             28.9085 0.5911 48.904 < 0.0000000000000000 ***
Significance levels
                    ## ---
                    ## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
                    ##
Residual std. error
                    ## Residual standard error: 393.3 on 4607 degrees of freedom
                    ## Multiple R-squared: 0.3417, Adjusted R-squared: 0.3416
          F-test
```

All these elements are easily accessible using the \$ operator. You can find the definition of most elements in the **Value** section of the summary function documentation ?summary.lm()

• Remember that functions take objects as arguments and produce new objects. **Value** describes the object that is created by a function.

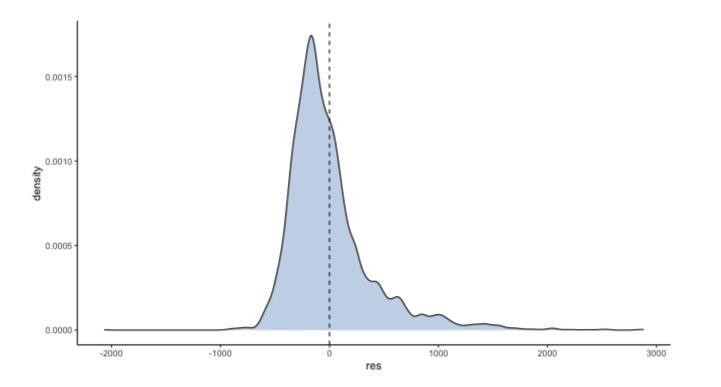
```
sum_mod <- summary(mod)</pre>
str(sum_mod, give.attr = F)
## List of 11
## $ call : language lm(formula = earnings ~ hours, data = cps)
## $ terms :Classes 'terms', 'formula' language earnings ~ hours
## $ residuals : Named num [1:4609] -398 592 -158 146 642 ...
## $ coefficients : num [1:2, 1:4] -198.809 28.908 21.915 0.591 -9.072 ...
## $ aliased : Named logi [1:2] FALSE FALSE
## $ sigma : num 393
## $ df : int [1:3] 2 4607 2
  $ r.squared : num 0.342
   $ adj.r.squared: num 0.342
##
## $ fstatistic : Named num [1:3] 2392 1 4607
   $ cov.unscaled : num [1:2, 1:2] 0.00310512 -0.00008078 -0.00008078 0.00000226
```

Taking the coefficients:

```
sum_mod$coefficients
##
                  Estimate Std. Error t value
                                                                      Pr(>|t|)
## (Intercept) -198.80923 21.9152104 -9.071746 0.0000000000000000001699063
## hours
                 28.90846 0.5911213 48.904453 0.0000000000000000000000000
We can subset this matrix like we would do with a regular data.frame:
sum_mod$coefficients[2,1]
## [1] 28.90846
sum_mod$coefficients[,"Std. Error"]
## (Intercept)
                      hours
   21.9152104 0.5911213
We can use this to compute the fitted values \hat{y} = \alpha + \beta x
alpha <- sum_mod$coefficients[1,1]</pre>
beta <- sum_mod$coefficients[2,1]</pre>
cps$hat_earnings <- alpha + (beta * cps$hours)</pre>
```

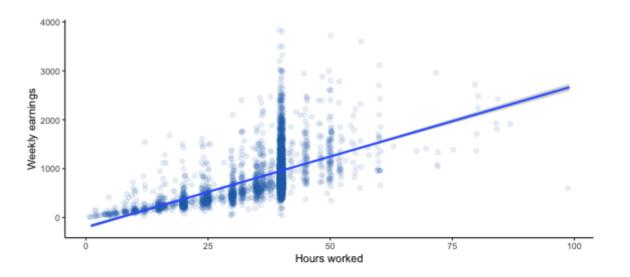
We can easily **plot** the **distribution** of our **residuals** 

```
data.frame(res = sum_mod$residuals) %>%  # Convert the residuals to a data frame
  ggplot(aes(x = res)) +
  geom_density() +
  geom_vline(xintercept = 0, linetype = "dashed")
```



• ggplot() has a dedicated geometry for fitted values geom\_smooth()

```
cps %>%
  ggplot(aes(x = hours, y = earnings)) +
  geom_point() +
  geom_smooth(method = "lm")
```



Check that  $\frac{Im()}{m}$  works fine by computing a regression manually for the model:

$$y = \alpha + \beta x$$

where y is earnings and x is hours worked.

1. Start by creating a variable for  $\hat{eta}$  , then for  $\hat{lpha}$  ,  $\hat{y}_i$  and  $\hat{arepsilon}_i$ 

$$\hat{eta} = rac{Cov(x,y)}{Var(x)}; \quad \hat{lpha} = ar{y} - \hat{eta} imes ar{x}$$

You will need the cov()and var() functions

Review: where do these formulas come from?  $E(u)=0; \quad E(u|x)=0$ 

2. Summarise the data to only display lpha ,  $\hat{eta}$  and the  $R^2$ 

$$R^2 = 1 - rac{\sum (y_i - {\hat y}_i)^2}{\sum (y_i - {ar y}_i)^2} \, .$$

1. Start by creating a variable for  $\hat{eta}$  , then for  $\hat{lpha}$  ,  $\hat{y}_i$  and  $\hat{arepsilon}_i$ 

$$\hat{eta} = rac{Cov(x,y)}{Var(x)}; \quad \hat{lpha} = ar{y} - \hat{eta} imes ar{x}$$

```
cps
#
#
#
#
hours earnings
```

```
## 1
           40
                 560.00
## 2
           30
               1260.00
## 3
           40
                800.00
## 4
           40
               1103.20
## 5
           40
               1600.00
## 6
           40
                840.00
## 7
           20
                280.00
## 8
           30
                 30.00
## 9
                400.00
           40
```

. . . .

1. Start by creating a variable for  $\hat{eta}$  , then for  $\hat{lpha}$  ,  $\hat{y}_i$  and  $\hat{arepsilon}_i$ 

$$\hat{eta} = rac{Cov(x,y)}{Var(x)}; \quad \hat{lpha} = ar{y} - \hat{eta} imes ar{x}$$

```
cps %>%
  mutate(beta = cov(hours, earnings) / var(hours))
#
#
#
```

```
##
        hours earnings
                            beta
## 1
           40
                560.00 28.90846
## 2
               1260.00 28.90846
           30
## 3
               800.00 28.90846
           40
## 4
              1103.20 28.90846
           40
## 5
           40
               1600.00 28.90846
## 6
               840.00 28.90846
           40
## 7
               280.00 28.90846
           20
## 8
           30
               30.00 28.90846
## 9
           40
                400.00 28.90846
. . . .
```

1. Start by creating a variable for  $\hat{eta}$  , then for  $\hat{lpha}$  ,  $\hat{y}_i$  and  $\hat{arepsilon}_i$ 

$$\hat{eta} = rac{Cov(x,y)}{Var(x)}; \quad \hat{lpha} = ar{y} - \hat{eta} imes ar{x}$$

```
##
        hours earnings
                           beta
                                     alpha
## 1
                560.00 28.90846 -198.8092
           40
## 2
               1260.00 28.90846 -198.8092
           30
## 3
           40
               800.00 28.90846 -198.8092
## 4
               1103.20 28.90846 -198.8092
## 5
           40
               1600.00 28.90846 -198.8092
## 6
           40
               840.00 28.90846 -198.8092
## 7
           20
               280.00 28.90846 -198.8092
## 8
           30
                30.00 28.90846 -198.8092
## 9
           40
                400.00 28.90846 -198.8092
. . . .
```

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1. Start by creating a variable for  $\hat{eta}$  , then for  $\hat{lpha}$  ,  $\hat{y}_i$  and  $\hat{arepsilon}_i$ 

$$\hat{eta} = rac{Cov(x,y)}{Var(x)}; \quad \hat{lpha} = ar{y} - \hat{eta} imes ar{x}$$

##		nours	earnings	beta	aipna	y_nat
##	1	40	560.00	28.90846	-198.8092	957.529242
##	2	30	1260.00	28.90846	-198.8092	668.444625
##	3	40	800.00	28.90846	-198.8092	957.529242
##	4	40	1103.20	28.90846	-198.8092	957.529242
##	5	40	1600.00	28.90846	-198.8092	957.529242
##	6	40	840.00	28.90846	-198.8092	957.529242
##	7	20	280.00	28.90846	-198.8092	379.360008
##	8	30	30.00	28.90846	-198.8092	668.444625
##	9	40	400.00	28.90846	-198.8092	957.529242

28 / 120

1. Start by creating a variable for  $\hat{eta}$  , then for  $\hat{lpha}$  ,  $\hat{y}_i$  and  $\hat{arepsilon}_i$ 

$$\hat{eta} = rac{Cov(x,y)}{Var(x)}; \quad \hat{lpha} = ar{y} - \hat{eta} imes ar{x}$$

##		nours	earnings	beta	aıpna	y_nat	res
##	1	40	560.00	28.90846	-198.8092	957.529242	-397.5292415
##	2	30	1260.00	28.90846	-198.8092	668.444625	591.5553752
##	3	40	800.00	28.90846	-198.8092	957.529242	-157.5292415
##	4	40	1103.20	28.90846	-198.8092	957.529242	145.6707585
##	5	40	1600.00	28.90846	-198.8092	957.529242	642.4707585
##	6	40	840.00	28.90846	-198.8092	957.529242	-117.5292415
##	7	20	280.00	28.90846	-198.8092	379.360008	-99.3600080
##	8	30	30.00	28.90846	-198.8092	668.444625	-638.4446248
##	9	40	400.00	28.90846	-198.8092	957.529242	-557.5292415

• • •

1. Summarise the data to only display lpha,  $\hat{eta}$  and the  $R^2=1-rac{\sum(y_i-\hat{y}_i)^2}{\sum(y_i-ar{y}_i)^2}$ 

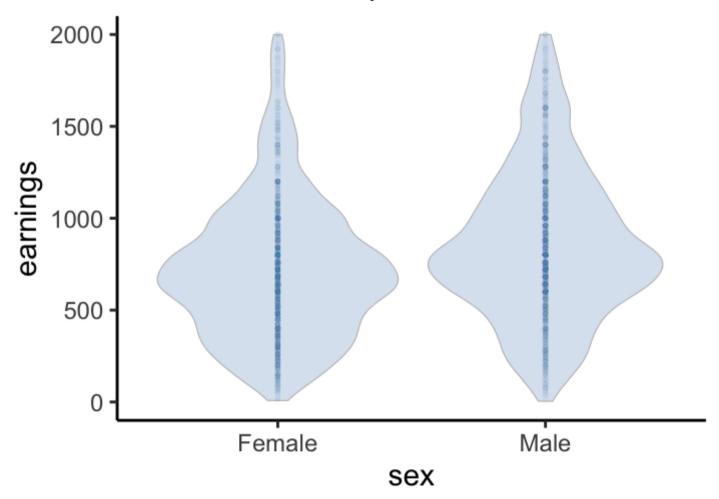
```
cps %>%
  mutate(beta = cov(hours, earnings) / var(hours),
          alpha = mean(earnings) - beta * mean(hours),
          y_hat = alpha + beta * hours,
          res = earnings - y_hat) %>%
   summarise(alpha = first(alpha),
             beta = first(beta),
             r2 = 1 - sum(res^2) / sum((earnings - mean(earnings))^2))
##
         alpha
                   beta
## 1 -198.8092 28.90846 0.3417298
 sum_mod$coefficients[,"Estimate"]
## (Intercept)
                     hours
## -198.80923
                  28.90846
 sum_mod$r.squared
## [1] 0.3417298
```

• We want to know the relationship between people's sex and earnings

```
cps$sex
     [1] "Male" "Female" "Male" "Male" "Female" "Male" "Male"
##
   [9] "Male" "Female" "Female" "Female" "Female" "Female" "Female" "Male"
##
    [17] "Female" "Female" "Female" "Female" "Female" "Female" "Female" "Female"
##
. . . .
Can we just regress earnings on sex even though it's a character variable?
lm(earnings \sim sex, data = cps)
##
## Call:
## lm(formula = earnings \sim sex, data = cps)
##
## Coefficients:
## (Intercept) sexMale
##
        746.3 179.3
```

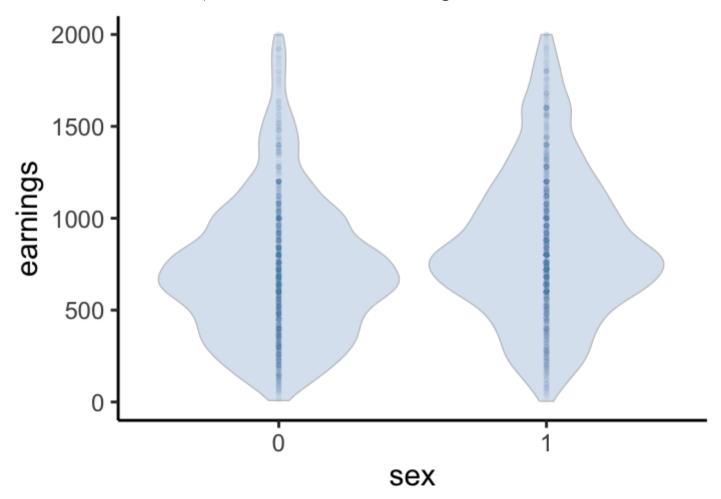
What is going on?

• R implicitely converts character variables into binary variables



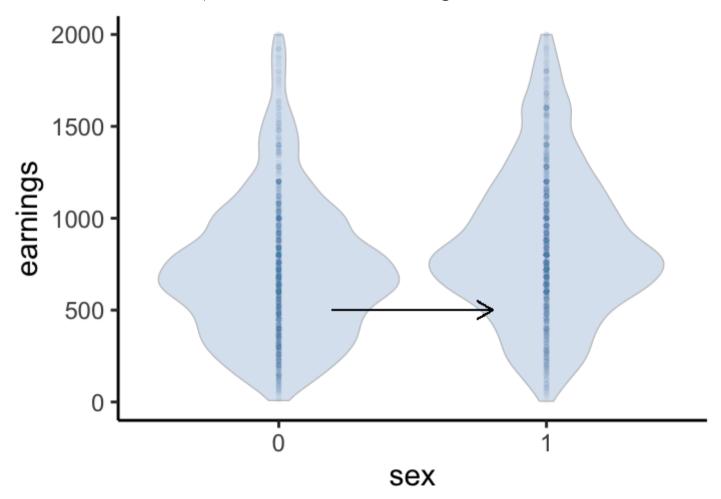
What is going on?

ullet ... to ressemble the **continuous** case, such that we're looking at a 1-unit increase in x



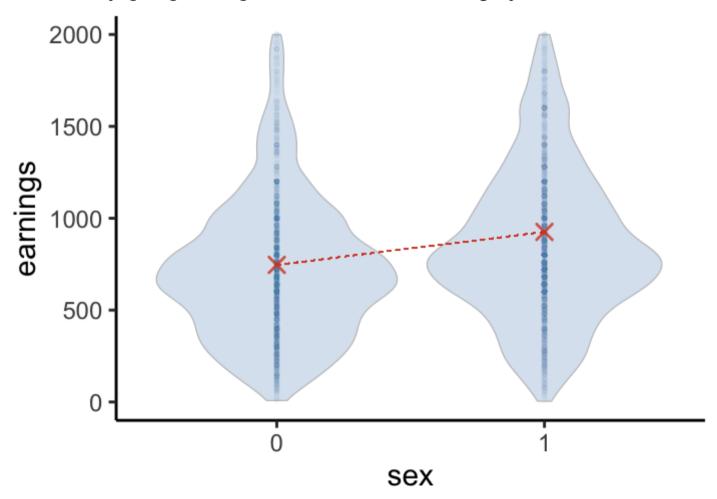
What is going on?

ullet ... to ressemble the **continuous** case, such that we're looking at a 1-unit increase in x



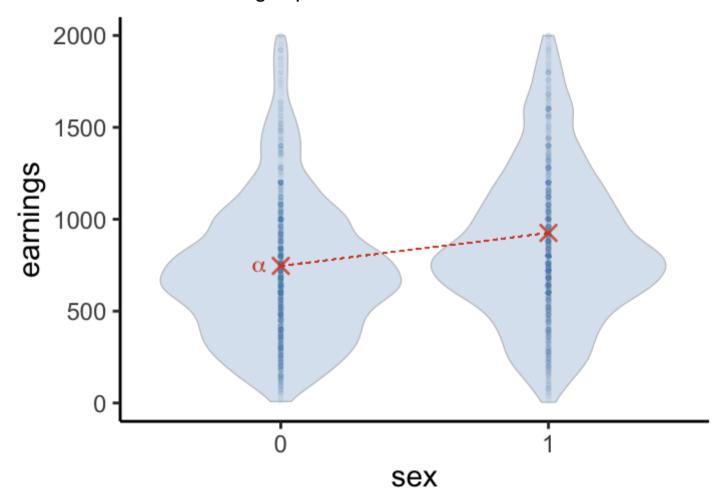
What is going on?

• Also, the **fit** is necessarily going through the **mean of each category** 



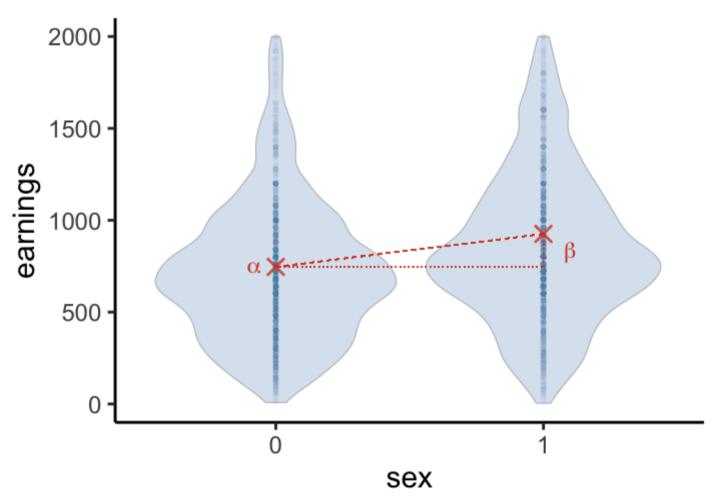
What is going on?

ullet Such that  $\hat{lpha}$  is the **mean of the reference group** 



What is going on?

• And  $\hat{eta}$  is the **difference in means** 



We can verify this easily:

```
lm(earnings ~ sex, data = cps)

##

## Call:
## lm(formula = earnings ~ sex, data = cps)
##

## Coefficients:
## (Intercept) sexMale
## 746.3 179.3
```

Let's complicate it by using categorical variables with more than two values. Let's see how **education** relates to **earnings** 

#### unique(cps\$educ)

```
## [1] "High school" "Bachelor's degree" "No high school"
```

## [4] "Associate degree"

Before, a 2-category variables was equivalent to 1 dummy variable. now an n-category variable is equivalent to n-1 dummy variables

sex	male
Female	0
Female	0
Female	0
Male	1
Male	1
Male	1

educ	Bachelor's degree	High school	No high school
High school	0	1	0
Bachelor's degree	1	0	0
No high school	0	0	1
Associate degree	0	0	0

Once again the constant is the average y for the reference category and the slopes are the relative differences in means

46.43

0

0

0

```
lm(earnings ~ educ, data = cps)
##
                       Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                         904.01
                                    19.47
## educBachelor's degree
                         106.06
                                    24.65 4.30
## educHigh school
                     -97.30
                                    21.58 -4.51
                                    28.19 -11.67
## educNo high school -328.85
cps %>%
  group_by(educ) %>%
  summarise(y_bar = mean(earnings)) %>%
  mutate(dif = y_bar - y_bar[1])
## # A tibble: 4 x 3
##
    educ
                     y_bar
                              dif
                     <dbl> <dbl>
##
    <chr>
## 1 Associate degree
                      904.
## 2 Bachelor's degree 1010.
                            106.
## 3 High school
                      807. -97.3
## 4 No high school
                      575. -329.
```

Note that R always sort character variables by alphabetical order

- Does it make sense that our reference category is "Associate's degree" ?
- We'd prefer to have "No high school" (the least education) as the reference category

We need to talk about factor variables (another class of R objects)

- Variables whose values indicate different groups
- They take different values that are arbitrary group classifiers

```
## [1] 1
## Levels: 1 2 3 4 5
```

• R understands that the different values do not mean anything they are there to differentiate groups only

```
states * 2
```

#### ## [1] NA NA NA NA NA

• We can specify the levels and labels that they take:

## [1] Male Male Female Female
## Levels: Male Female

mutate(educf = relevel(as.factor(educ), ref = "No high school"))

levels(cps\$educf)

Going back to our education variable, we want to order it in a way that makes sense:

```
levels(as.factor(cps$educ)) # Default order is alphabetical
## [1] "Associate degree" "Bachelor's degree" "High school"
## [4] "No high school"
Let's create a factor variable with the right order using factor()
cps <-
  cps %>%
  mutate(educf = factor(educ, levels = c("No high school", "High school",
                                           "Associate degree", "Bachelor's degree")))
levels(cps$educf)
## [1] "No high school" "High school"
                                                "Associate degree"
## [4] "Bachelor's degree"
Or only change the reference category (first level), using relevel()
cps <-
  cps %>%
```

```
lm(earnings ~ educf, data = cps)
##
## Call:
## lm(formula = earnings \sim educf, data = cps)
##
## Coefficients:
              (Intercept)
                            educfAssociate degree educfBachelor's degree
##
##
                    575.2
                                             328.8
                                                                      434.9
##
         educfHigh school
##
                    231.5
We can also modify our educ variable directly in the regression call:
lm(earnings ~ relevel(as.factor(educ), ref = "No high school"), data = cps)
or
lm(earnings ~ factor(educ, levels = c("No high school", "High school",
                                        "Associate degree", "Bachelor's degree")),
    data = cps)
```

# **PRACTICE**

- 1. Load the Current Population Survey data cps,csv in case you haven't
- 2. Regress the earnings on the age variable
- 3. Redo the same regression afer converting the age variable as a factor

# **PRACTICE**

```
lm(earnings ~ age, data = cps) %>%
  summary()
##
## Call:
  lm(formula = earnings ~ age, data = cps)
##
  Residuals:
##
       Min
                    Median
                                 30
                                        Max
   -966.33 -308.32 -75.55
                            207.35 3012.86
##
  Coefficients:
##
               Estimate Std. Error t value
Pr(>|t|)
## (Intercept) 617.8860
                           19.1940
                                      32.19
<0.00000000000000002 ***
## age
                 5.5067
                            0.4534
                                      12.15
<0.000000000000000002 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*'
0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 477.2 on 4607
degrees of freedom
## Multiple R-squared: 0.03103
                                     Adjusted R-
```

```
lm(earnings ~ as.factor(age), data = cps) %>%
   summary()
##
## Call:
## lm(formula = earnings \sim as.factor(age), data
= cps)
##
   Residuals:
##
       Min
                10
                    Median
                                 30
                                        Max
## -995.53 -283.21
                    -75.23
                            191.46 2852.56
##
  Coefficients:
##
                     Estimate Std. Error t value
Pr(>|t|)
## (Intercept)
                                 117.202
                                           2.510
                     294.167
0.012111 *
## as.factor(age)16
                       -9.190
                                 128.533
                                          -0.072
0.943002
## as.factor(age)17
                                           0.056
                       7.128
                                 126.798
0.955174
## as.factor(age)18
                     174.198
                                 126.695
                                           1.375
0.169217
## as.factor(age)19
                     231.065
                                 124.804
                                           1.851
0.064174
```

We can also **one hot encode** the data: converting the categorical variable to **several dummies** so that everything is numeric

Then, we can include them as different variables in the regression using the + sign

```
lm(earnings ~ educ_hs + educ_assoc + educ_bach, data = cps)
```

Why do we need to omit one category?

## Why do we need to omit one category?

ullet We are in the **multivariate case**, so we move from  $\hat{eta}=rac{Cov(x,y)}{Var(x)}$  to  $\hat{eta}=(m{X'X})^{-1}m{X'y}$ 

```
y <- as.matrix(cps$earnings)</pre>
X <- cps %>%
  mutate(constant = 1) %>%
  select(constant, contains("educ_")) %>%
  as.matrix()
dim(y)
                                                       y
## [1] 4609
dim(X)
                                                       X
## [1] 4609
               5
dim(t(X))
                                                       X'
## [1] 5 4609
dim(t(X) %*% X)
                                                       X'X
## [1] 5 5
```

Because of perfect multicollinearity (no explanatory variable is a perfect linear function of other explanatory variables) it will not be posible to invert  $X^{\prime}X$ 

```
solve(t(X) %*% X)
                                                      (X'X)^{-1}
## [1] "Error in solve.default(t(X) %*% X)
system is computationally singular"
X
##
         constant educ_nohs educ_hs educ_assoc educ_bach
##
   [1,]
## [2.]
   [3,]
##
   [4,]
##
##
   [5,]
## [6,]
## [7,]
   [8,]
##
   [9,]
##
```

- constant = educ\_nohs + educ\_hs + educ\_assoc + educ\_bach
- educ\_bach = 1 educ\_nohs educ\_hs educ\_assoc

We need to:

• Remove one category

```
X <- cps %>%
  mutate(constant = 1) %>%
  select(constant, educ_hs,
         educ_assoc, educ_bach) %>%
  as.matrix()
solve(t(X) %*% X) %*% (t(X) %*% y)
##
                  [,1]
## constant 575.1593
## educ hs 231.5476
## educ_assoc 328.8480
## educ_bach 434.9049
lm(earnings ~ educ_hs +
               educ_assoc + educ_bach,
   cps)
```

• Or remove the constant

```
X <- cps %>%
   select(educ_nohs, educ_hs,
         educ_assoc, educ_bach) %>%
   as.matrix()
solve(t(X) %*% X) %*% (t(X) %*% y)
                   [,1]
##
## educ_nohs
               575, 1593
## educ hs 806.7068
## educ_assoc 904.0072
## educ bach 1010.0642
 lm(earnings ~ educ_nohs + educ_hs +
               educ_assoc + educ_bach - 1,
   cps)
```

You can remove the constant in lm() by adding
 to the formula

Actually, if we don't drop anything lm() would still work:

```
lm(earnings ~ educ_nohs + educ_hs + educ_assoc + educ_bach.
   cps)
##
## Call:
## lm(formula = earnings ~ educ_nohs + educ_hs + educ_assoc + educ_bach,
##
       data = cps)
##
## Coefficients:
## (Intercept)
                 educ nohs
                                educ hs educ assoc
                                                        educ bach
##
        1010.1
                    -434.9
                                 -203.4
                                               -106.1
                                                               NA
```

- It will automatically drop one of the categories **but** it might not be the most adequate reference category
- Even if multicollinearity does not break lm() you need to be mindful about it
  - Make sure the your explanatory variables are not redundant
  - Multicollinearity will invalidate statistical inference (inflate standard errors)
- Always be intentional about your reference categories!

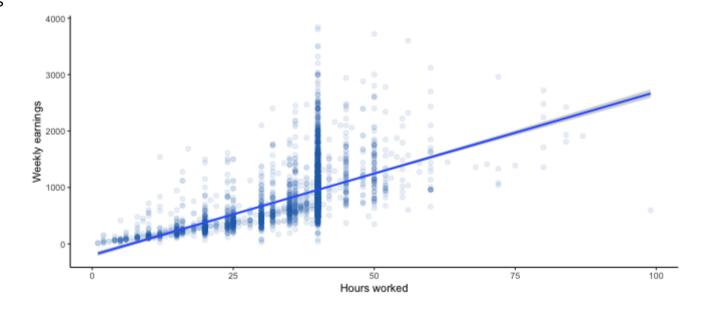
# **AGENDA**

- REGRESSIONS IN R
  - CONTINUOUS VARIABLES
  - CATEGORICAL AND BINARY VARIABLES
- GETTING TO THE RIGHT MODEL
  - VARIABLE TRANSFORMATION
  - FUNCTIONAL FORM
  - CONTROL VARIABLES
  - INTERACTION TERMS

- INFERENCE
- EXPORTING RESULTS
  - REGRESSION TABLES
  - COEFFICIENT PLOTS
- PRACTICING YOUR SKILLS

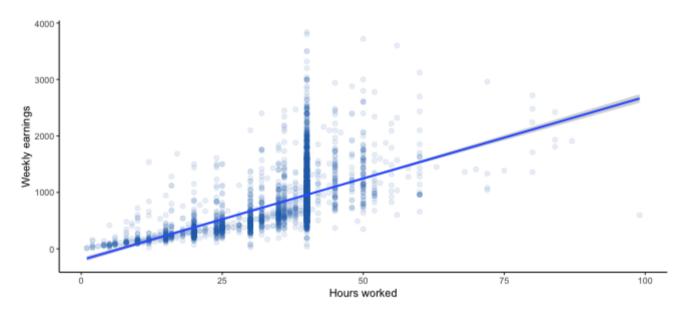
We want to examine the relationship between earnings and hours worked. We might need to add a couple of things before we are happy with our model:

- Variable transformation
- Functional form
- Control variables
- Interaction terms



### Variable transformation

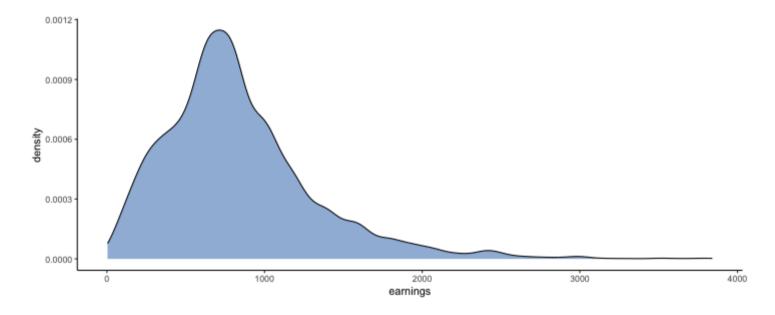
What we have so far:



- Earnings are concentrated below \$2000
- The regression does not fit well on the **y dimension**

#### Variable transformation

```
ggplot(cps, aes(x = earnings)) +
  geom_density()
```



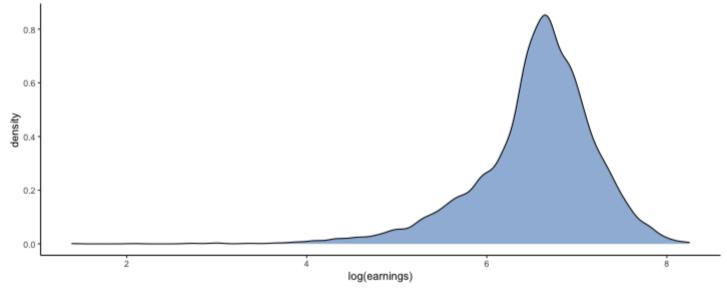
- Earnings are concentrated below \$2000, with fewer and fewer observations after that
- The estimated probability density function has a long right tail

# The distribution is likely log-normal!

(when we take its logarithm it looks like a normal distribution)

#### Variable transformation

```
ggplot(cps, aes(x = log(earnings))) +
  geom_density()
```



The log-transformation is very popular in economics:

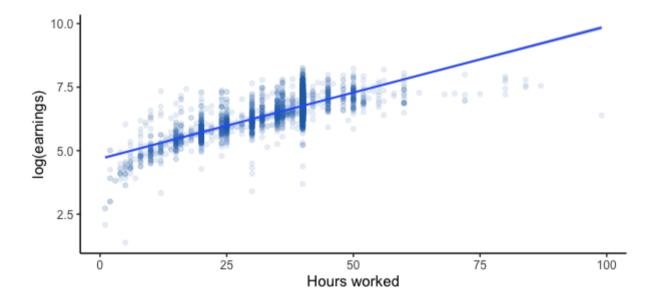
- Several important economic variables (like wages, income, firm size, GDP) are approximately **log-normally distributed**. A normal distribution has desirable properties for our regression.
- Reduces the impact of outliers
- Allow for convenient interpretations in terms of percentage changes of the outcome variable

## Variable transformation

Specification	Outcome var	Regressor	Interpretation of $eta$	Name
Level-level	y	x	$\Delta y = eta \Delta x$	Standard
Level-log	y	log(x)	$\Delta y = rac{eta}{100} \Delta x$	Less common
Log-level	log(y)	x	$\%\Delta y = (100eta)\Delta x$	Semi-elasticity
Log-log	log(y)	log(x)	$\%\Delta y=\%\Deltaeta x$	Elasticity

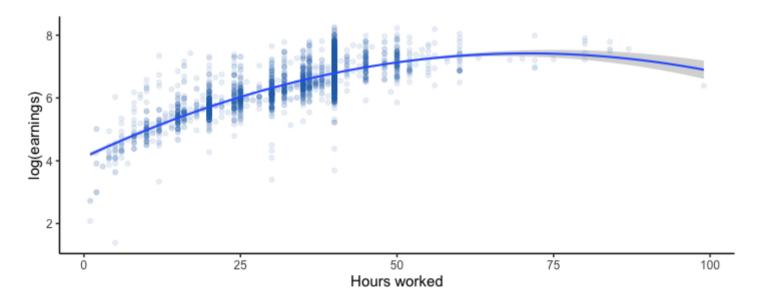
#### Variable transformation

```
cps %>%
  ggplot(aes(x = hours, y = log(earnings) )) +
  geom_point() +
  geom_smooth(method = "lm")
```



• Better! But the fit looks weird for low and high numbers of hours

### **FUNCTIONAL FORM**



#### **FUNCTIONAL FORM**

We can rewrite our model:

$$log(Earnings_i) = lpha + eta_i Hours_i + eta_2 Hours_i^2 + arepsilon_i$$

• Create the necessary variables:

```
cps <- cps %>%
  mutate(logearnings = log(earnings),
     sqhours = hours^2)
```

• Run the new model:

```
lm(logearnings ~ hours + sqhours, cps )
##
## Call:
## lm(formula = logearnings ~ hours + sqhours, data = cps)
##
## Coefficients:
## (Intercept) hours sqhours
## 4.1087277 0.0934110 -0.0006587
```

## FUNCTIONAL FORM Are we missing something?

• This relationship might be **driven** by **something else**. People who work more hours are different than people who work less hours in ways that also affect earnings. **How is this called?** 

#### For example:

- Men tend to work full time more often and earn more
- Higher hours are highly correlated with being a man, being a man is also correlated with higher wages

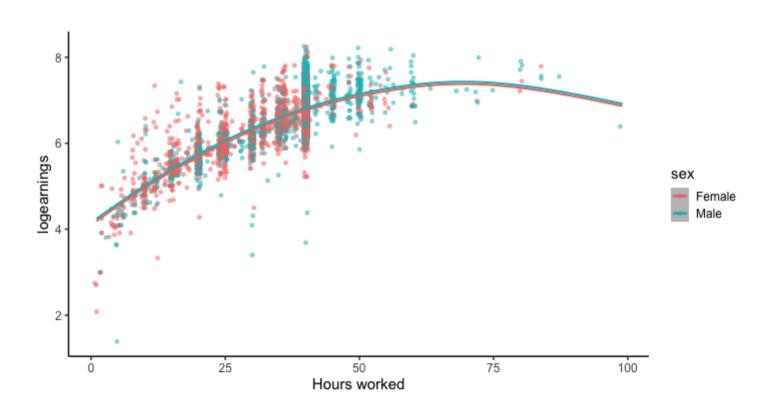
We want to control for this in the regression:

$$log(Earnings_i) = lpha + eta_1 Hours_i + eta_2 Hours_i^2 + eta_3 Male_i + arepsilon_i$$

	Male = 0	Male = 1
Intercept	lpha	$lpha+eta_3$
Slope	$eta_1 + 2eta_2 Hours$	$eta_1 + 2eta_2 Hours$

-> Men and women have the same slope but different
intercepts

## **CONTROL VARIABLES**



#### **CONTROL VARIABLES**

How do the coefficients of both models compare?

```
lm(logearnings ~ hours + sqhours + male, cps)
## (Intercept) hours sqhours male
## 4.0945625390 0.0933176308 -0.0006649689 0.0528821933
lm(logearnings ~ hours + sqhours, cps)
## (Intercept) hours sqhours
## 4.1087276823 0.0934110198 -0.0006587048
```

#### The omitted variable bias was not too big in this case

- Beware of violently adding controls (more on this in Econometrics 3 for a deep dive on identification).
- A Crash Course in Good and Bad Controls

**INTERACTION TERMS** We only allowed the **intercept** to vary by gender. What if we want the **slope** to vary too?

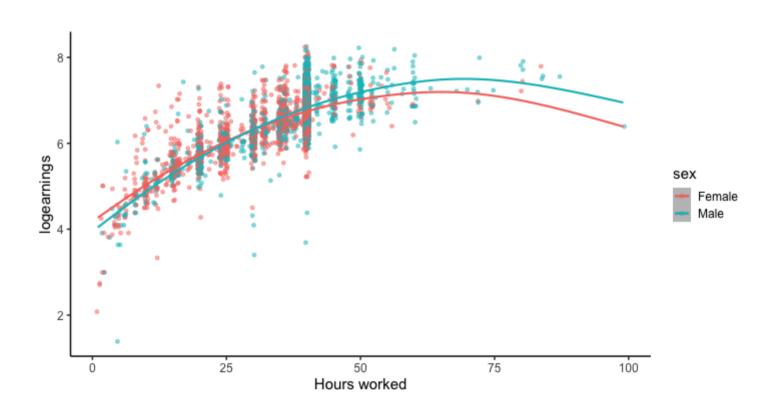
- The relationship between hours and earnings might be heterogeneous between men and women
- ullet We allow eta to vary by gender with an interaction term

$$log(Earnings_i) = lpha + eta_1 Hours_i + eta_2 Hours_i^2 + eta_3 Male_i + eta_4 Hours_i imes Male_i + arepsilon_i$$

	Male = 0	Male = 1
Intercept	lpha	$lpha+eta_3$
Slope	$eta_1 + 2eta_2 Hours$	$eta_1 + 2eta_2 Hours + eta_4$

-> Men and women have different slopes and different intercepts

## **INTERACTION TERMS**



#### INTERACTION TERMS

We specify interaction terms with a \*

```
      mod <- lm(logearnings ~ hours + sqhours + male + male*hours, cps)</td>

      results <- summary(mod)$coefficients</td>

      ## (Intercept)

      4.1714631273 0.0378575672 110.188357 0.0000000000

      ## hours

      0.0930787301 0.0020751060 44.854928 0.0000000000

      ## sqhours

      -0.0007196841 0.0000329061 -21.870832 0.0000000000

      ## male

      -0.2348248075 0.0482515339 -4.866681 0.0000011728

      ## hours:male
```

What can we conclude?

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#### **HYPOTHESIS TESTING**

- Let's assume that we know that  $\hat{\beta}_1$  is equal to .10 using the previous month data. Are we sure that  $\hat{\beta}_1$  is actually below .10?
- How can we test whether or not our test is below from .10?

We want to use our sample estimates to conclude something about the population parameters!

Concept	Description	Example
Null hypothesis	The hypothes to evaluate	$H_0:eta_1\geq .10$
Alternative hypothesis	The statement if the value differs from the null hypothesis	$H_1:eta_1<.10$
Test statistic	The tool (point estimate statistic formula) we use to decide if we reject or not the null hypothesis	$t=rac{\hat{eta}10}{s.e.(\hat{eta})}$
Null distribution	The sampling distribution of the test statistic $\emph{assuming}$ the null $H_0$ is true	t follows a Student's t- distribution
P-value	The prob. of obtaining a test statistic just as extreme or more extreme than the observed test statistic assuming the null is true. How surprised am I of oberving t assuming $H_0$ holds?	
Significance level	A cutof on the p-value: if the p-value does not fall below $\alpha$ we would "fail to reject ${\cal H}_0$	$\alpha = 0.05$ 78

#### HYPOTHESIS TESTING

- 1. State the null and alternative hypotheses  $H_0: eta_1 \geq .10$ ,  $H_1: eta_1 < .10$  2. Choose a test and significance level
  - How many parameters do we have? (one = one-sample test, two = two-sample test)
  - Do we know the population variance? (yes = z-test, no = t-test)
    - 3. Compute the observed test statistic: Computing the t-stat  $t=rac{\hat{eta}-.10}{s.e.(\hat{eta})}$

```
round(results)

## Estimate Std. Error t value Pr(>|t|)

## (Intercept) 4.17146 0.03786 110.18836 0

## hours 0.09308 0.00208 44.85493 0

## sqhours -0.00072 0.00003 -21.87083 0

## male -0.23482 0.04825 -4.86668 0

## hours:male 0.00803 0.00130 6.16499 0

t <- (results[2,1] - .10) / results[2,2]

t

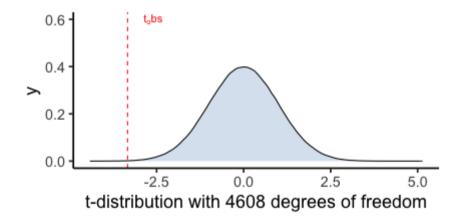
## [1] -3.335381
```

#### HYPOTHESIS TESTING

Our observed test statistic provides a measure of "evidence" against the null hypothesis. In particular, we know that under the null hypothesis, the test statistic follows a  $t_{df}=t_{n-1}$  =  $t_4608$  distribution

```
df <- nrow(cps) - 1
df
```

## [1] 4608



- This distribution represents the distribution of sample evidence given that the null is true
- Our observed test statistic (the dashed red line) shows that the event we observed is unlikely to occur if the null (  $H_0: \beta_1 \geq .10$  ) is true. Now we want to calculate this probability more formally.

#### HYPOTHESIS TESTING

4. Compute the p-value The p-value is the probability of getting sample evidence as or more extreme than what we actually observed given that the null hypothesis is actually true.

Since we are working with a "smaller-than" alternative hypothesis: \$ \text{p-value} =  $P(t{df} < t{obs} | H_0$ \text{ is true}) \$\$

```
pt(t, df = df, lower.tail = TRUE)
```

## [1] 0.0004292814

5. Make a statistical decision and interpret the results:

if p-value 
$$\leq \alpha$$
 reject  $H_0$ 

if p-value 
$$> \alpha$$
 fail to reject  $H_0$ 

• Since  $\alpha$ is the maximum p-value at which we reject  $H_0$ , then we are ensuring that there is at most a  $100 \times \alpha\%$  chance of making a type I error (reject the null when it is true): a 5% chance of mistakenly deciding that  $\beta_1 \geq .10$  when  $\beta_1 < .10$ 

#### HYPOTHESIS TESTING

We can use the linearHypothesis() function from the {car} package to carry out one and two-sample t-tests in R.

- The first argument is the model
- The second argument is the null hypothesis

```
## Linear hypothesis test
##
## Hypothesis:
## male = \theta
##
## Model 1: restricted model
## Model 2: logearnings ~ hours + sqhours + male + hours * male
##
           RSS Df Sum of Sq F 	ext{Pr}(>F)
##
    Res.Df
      4605 773.82
## 1
      4604 769.86 1 3.9604 23.685 0.000001173 ***
## 2
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

### **INFERENCE**

#### HYPOTHESIS TESTING

We can also use it to carry out F-tests for joint significance

```
linearHypothesis(
  lm(logearnings ~ hours + sqhours + male + hours*male, cps).
       c("hours = 0", "sqhours = 0", "male = 0", "hours:male = 0"))
## Linear hypothesis test
##
## Hypothesis:
## hours = 0
## sqhours = 0
## male = \theta
## hours:male = 0
##
## Model 1: restricted model
## Model 2: logearnings ~ hours + sqhours + male + hours * male
##
    Res.Df RSS Df Sum of Sq F
##
                                                     Pr(>F)
## 1
      4608 2056.79
      4604 769.86 4 1286.9 1924 < 0.000000000000000022 ***
## 2
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
mod <- lm(logearnings ~ hours + sghours + male + hours*male. cps)
```

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Eventually, we want to format our regression results into tables to include in our **papers** and **reports**. In general, they look like this:

Table 2
Labor market states of prime-age men and women.

		Men			Women	
	Participate	Emp/Pop	Unem/Pop	Participate	Emp/Pop	Unem/Pop
Prescrip. Rate	-0.053***	-0.055***	0.003*	-0.010***	-0.011**	0.002
	(0.005)	(0.006)	(0.002)	(0.003)	(0.003)	(0.001)
Demand Shock	0.317**	0.689***	-0.311***	-0.456***	-0.202*	-0.210***
	(0.105)	(0.117)	(0.057)	(0.093)	(0.098)	(0.051)
2000 Particip.	0.637***			0.409***		
-	(0.050)			(0.029)		
2000 Emp/Pop		0.517***			0.349***	
		(0.036)			(0.021)	
2000 Unem/Pop			0.263***			0.169***
_			(0.027)			(0.021)
R-sqr	0.09	0.11	0.02	0.06	0.06	0.02
N	6424995	6424995	6424995	6641288	6641288	6641288

All regressions include demographic variables, year, and state fixed effects. Robust standard errors with clustering on coumas.

<sup>\*</sup> p < 0.05, \*\* p < 0.01, \*\*\* p < 0.001.

```
There are many different packages to produce regression tables:
{stargazer} {gtsummary} {huxtable} {fixest} {modelsummary}
We'll work with the <a href="huxreg">huxreg</a>() function from the <a href="huxtable">{huxtable}</a> package. Main arguments:
  • Many models, with a name or automatically numbered
 outreg <- huxreg(Baseline = lm(logearnings ~ hours, cps),</pre>
                    lm(logearnings ~ hours + sqhours, cps),
                    lm(logearnings ~ hours + sqhours + male, cps),
                    lm(logearnings ~ hours + sqhours + male + male*hours, cps),
```

```
There are many different packages to produce regression tables:
{stargazer} {gtsummary} {huxtable} {fixest} {modelsummary}
We'll work with the <a href="huxreg">huxreg</a>() function from the <a href="huxtable">{huxtable}</a> package. Main arguments:
  • Many models, with a name or automatically numbered
  • Which uncertainty statistic to display (std.error, p.value, conf.low, conf.high)
 outreg <- huxreg(Baseline = lm(logearnings ~ hours, cps),</pre>
                   lm(logearnings ~ hours + sqhours, cps),
                   lm(logearnings ~ hours + sqhours + male, cps),
                   lm(logearnings ~ hours + sqhours + male + male*hours, cps),
                   error_format = "({std.error})",
```

```
There are many different packages to produce regression tables:
{stargazer} {gtsummary} {huxtable} {fixest} {modelsummary}
We'll work with the <a href="huxreg">huxreg</a>() function from the <a href="huxtable">{huxtable}</a> package. Main arguments:
  • Many models, with a name or automatically numbered
  • Which uncertainty statistic to display (std.error, p.value, conf.low, conf.high)
  • Where to place the uncertainty statistic (below, same, right),
 outreg <- huxreg(Baseline = lm(logearnings ~ hours, cps),</pre>
                   lm(logearnings ~ hours + sqhours, cps),
                   lm(logearnings ~ hours + sqhours + male, cps),
                   lm(logearnings ~ hours + sqhours + male + male*hours, cps),
                   error_format = "({std.error})",
                   error_pos = "below",
```

```
There are many different packages to produce regression tables:
{stargazer} {gtsummary} {huxtable} {fixest} {modelsummary}
We'll work with the <a href="huxreg">huxreg</a>() function from the <a href="huxtable">{huxtable}</a> package. Main arguments:
  • Many models, with a name or automatically numbered
  • Which uncertainty statistic to display (std.error, p.value, conf.low, conf.high)
  • Where to place the uncertainty statistic (below, same, right),
  • Which general statistics to display (adj.r.squared, df, ...)
 outreg <- huxreg(Baseline = lm(logearnings ~ hours, cps),</pre>
                   lm(logearnings ~ hours + sqhours, cps),
                   lm(logearnings ~ hours + sqhours + male, cps),
                   lm(logearnings ~ hours + sqhours + male + male*hours, cps),
                   error_format = "({std.error})",
                   error_pos = "below",
                   statistics = c(N = "nobs", R2 = "r.squared"),
```

```
There are many different packages to produce regression tables:
{stargazer} {gtsummary} {huxtable} {fixest} {modelsummary}
We'll work with the <a href="huxreg">huxreg</a>() function from the <a href="huxtable">{huxtable}</a> package. Main arguments:
  • Many models, with a name or automatically numbered
  • Which uncertainty statistic to display (std.error, p.value, conf.low, conf.high)
  • Where to place the uncertainty statistic (below, same, right),
  • Which general statistics to display (adj.r.squared, df, ...)
  • The designed significance symbols
 outreg <- huxreg(Baseline = lm(logearnings ~ hours, cps),
                   lm(logearnings ~ hours + sqhours, cps),
                   lm(logearnings ~ hours + sqhours + male, cps),
                   lm(logearnings ~ hours + sqhours + male + male*hours, cps),
                   error_format = "({std.error})",
                   error_pos = "below",
                   statistics = c(N = "nobs", R2 = "r.squared"),
                   stars = c(`***` = 0.01, `**` = 0.05, `*` = 0.1),
```

```
There are many different packages to produce regression tables:
{stargazer} {gtsummary} {huxtable} {fixest} {modelsummary}
We'll work with the <a href="huxreg">huxreg</a>() function from the <a href="huxtable">{huxtable}</a> package. Main arguments:
  • Many models, with a name or automatically numbered
  • Which uncertainty statistic to display (std.error, p.value, conf.low, conf.high)
  • Where to place the uncertainty statistic (below, same, right),
  • Which general statistics to display (adj.r.squared, df, ...)
  • The designed significance symbols
  • What to write in the footnote
 outreg <- huxreg(Baseline = lm(logearnings ~ hours, cps),
                   lm(logearnings ~ hours + sqhours, cps),
                   lm(logearnings ~ hours + sqhours + male, cps),
                   lm(logearnings ~ hours + sqhours + male + male*hours, cps),
                   error_format = "({std.error})",
                   error_pos = "below",
                   statistics = c(N = "nobs", R2 = "r.squared"),
```

stars = c(`\*\*\*` = 0.01, `\*\*` = 0.05, `\*` = 0.1),

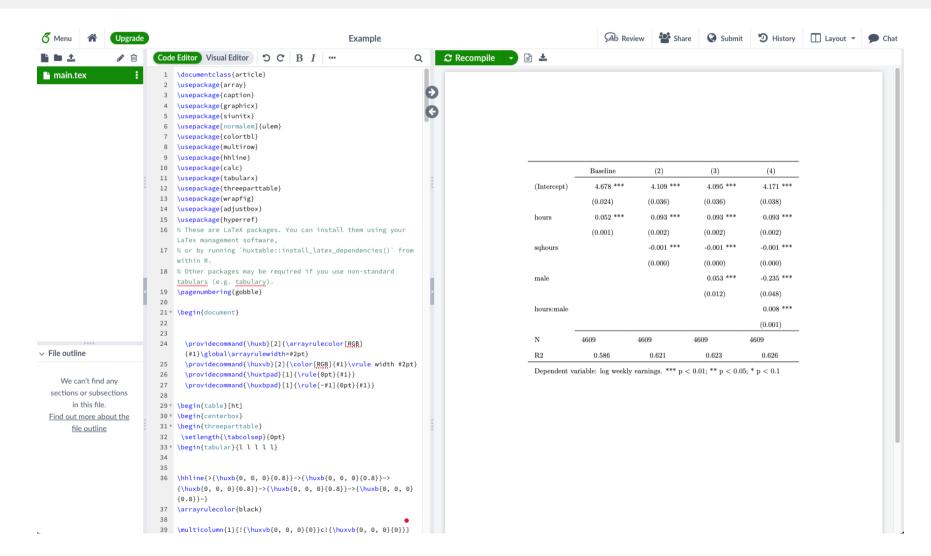
note = "Dependent variable: log weekly earnings. {stars}")

We get somthing like this in the console:

	Baseline	(2)	(3)	(4)
(Intercept)	4.678 ***	4.109 ***	4.095 ***	4.171 ***
	(0.024)	(0.036)	(0.036)	(0.038)
hours	0.052 ***	0.093 ***	0.093 ***	0.093 ***
	(0.001)	(0.002)	(0.002)	(0.002)
sqhours		-0.001 ***	-0.001 ***	-0.001 ***
		(0.000)	(0.000)	(0.000)
male		·	0.053 ***	-0.235 ***
			(0.012)	(0.048)
hours:male			·	0.008 ***
				(0.001)
N	4609	4609	4609	4609
R2	0.586	0.621	0.623	0.626

• We can export it with the functions: quick\_latex(), quick\_html(), quick\_pdf(), quick\_docx(), quick\_xlsx(),

quick\_latex(outreg, file = "./inputs/04\_regtable.tex")



	Baseline	(2)	(3)	(4)
(Intercept)	4.678 ***	4.109 ***	4.095 ***	4.171 ***
	(0.024)	(0.036)	(0.036)	(0.038)
hours	0.052 ***	0.093 ***	0.093 ***	0.093 ***
	(0.001)	(0.002)	(0.002)	(0.002)
sqhours		-0.001 ***	-0.001 ***	-0.001 ***
		(0.000)	(0.000)	(0.000)
male			0.053 ***	-0.235 ***
			(0.012)	(0.048)
hours:male				0.008 ***
				(0.001)
N	4609	4609	4609	4609
R2	0.586	0.621	0.623	0.626
Dependent variable: log weekly earnings. *** p < 0.01; ** p < 0.05; * p < 0.1				

Use the functions huxreg(), insert\_row() and merge\_cells() to reproduce this table and export it to html.

	(1)	(2)	(3)	(4
Hours worked	0.052 ***	0.093 ***	0.093 ***	0.09
	(0.000)	(0.000)	(0.000)	(0.0)
(Hours worked) <sup>2</sup>		-0.001 ***	-0.001 ***	-0.00
		(0.000)	(0.000)	(0.0)
Male			0.053 ***	-0.23
			(0.000)	(0.0)
Hours worked x Male				0.00
				(0.0)
Constant	4.678 ***	4.109 ***	4.095 ***	4.17
	(0.000)	(0.000)	(0.000)	0.0)
N	4609	4609	4609	4609
R2 adj.	0.586	0.621	0.622	0.6



```
huxreg(lm(logearnings ~ hours, cps),
       lm(logearnings ~ hours + sqhours, cps),
       lm(logearnings ~ hours + sqhours + male, cps),
       lm(logearnings ~ hours + sqhours + male + male*hours, cps),
```

```
huxreg(lm(logearnings ~ hours, cps),
       lm(logearnings ~ hours + sqhours, cps),
       lm(logearnings ~ hours + sqhours + male, cps),
       lm(logearnings ~ hours + sqhours + male + male*hours, cps),
       error_format = "({p.value})",
       error_pos = "below",
```

```
huxreg(lm(logearnings ~ hours, cps),
       lm(logearnings ~ hours + sqhours, cps),
       lm(logearnings ~ hours + sqhours + male, cps),
       lm(logearnings ~ hours + sqhours + male + male*hours, cps),
       error_format = "({p.value})",
       error_pos = "below",
       statistics = c(N = "nobs", "R2 adj." = "adj.r.squared"),
       stars = c(`***` = 0.01, `**` = 0.05, `*` = 0.1),
```

```
huxreg(lm(logearnings ~ hours, cps),
       lm(logearnings ~ hours + sqhours, cps),
       lm(logearnings ~ hours + sqhours + male, cps),
       lm(logearnings ~ hours + sqhours + male + male*hours, cps),
       error_format = "({p.value})",
       error_pos = "below",
       statistics = c(N = "nobs", "R2 adj." = "adj.r.squared"),
       stars = c(`***` = 0.01, `**` = 0.05, `*` = 0.1),
        coefs = c("Hours worked" = "hours",
                 "(Hours worked)<sup>2</sup>" = "sqhours",
                 "Male" = "male",
                 "Hours worked x Male" = "hours:male",
                 "Constant" = "(Intercept)"),
```

```
huxreg(lm(logearnings ~ hours, cps),
       lm(logearnings ~ hours + sqhours, cps),
       lm(logearnings ~ hours + sqhours + male, cps),
       lm(logearnings ~ hours + sqhours + male + male*hours, cps),
       error_format = "({p.value})",
       error_pos = "below",
       statistics = c(N = "nobs", "R2 adj." = "adj.r.squared"),
       stars = c(`***` = 0.01, `**` = 0.05, `*` = 0.1),
        coefs = c("Hours worked" = "hours",
                 "(Hours worked)<sup>2</sup>" = "sqhours",
                 "Male" = "male".
                 "Hours worked x Male" = "hours:male",
                 "Constant" = "(Intercept)"),
       note = "P-values in parentheses. {stars}",
       align = "c") %>%
```

```
huxreg(lm(logearnings ~ hours, cps),
       lm(logearnings ~ hours + sqhours, cps),
       lm(logearnings ~ hours + sqhours + male, cps),
       lm(logearnings ~ hours + sqhours + male + male*hours, cps),
       error_format = "({p.value})",
       error_pos = "below".
       statistics = c(N = "nobs", "R2 adj." = "adj.r.squared"),
       stars = c(`***` = 0.01, `**` = 0.05, `*` = 0.1),
        coefs = c("Hours worked" = "hours",
                 "(Hours worked)<sup>2</sup>" = "sqhours",
                 "Male" = "male".
                 "Hours worked x Male" = "hours:male",
                 "Constant" = "(Intercept)"),
       note = "P-values in parentheses. {stars}",
       align = "c") %>%
 insert_row(c("", "Dependent variable: Log weekly earnings", rep("", 3))) %>%
```

```
huxreg(lm(logearnings ~ hours, cps),
       lm(logearnings ~ hours + sqhours, cps),
       lm(logearnings ~ hours + sqhours + male, cps),
       lm(logearnings ~ hours + sqhours + male + male*hours, cps),
       error_format = "({p.value})",
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       statistics = c(N = "nobs", "R2 adj." = "adj.r.squared"),
       stars = c(`***` = 0.01, `**` = 0.05, `*` = 0.1),
        coefs = c("Hours worked" = "hours",
                 "(Hours worked)<sup>2</sup>" = "sqhours",
                 "Male" = "male".
                 "Hours worked x Male" = "hours:male",
                 "Constant" = "(Intercept)"),
       note = "P-values in parentheses. {stars}",
       align = "c") %>%
  insert_row(c("", "Dependent variable: Log weekly earnings", rep("", 3))) %>%
 merge_cells(1, 2:5) %>%
```

```
huxreg(lm(logearnings ~ hours, cps),
       lm(logearnings ~ hours + sqhours, cps),
       lm(logearnings ~ hours + sqhours + male, cps),
       lm(logearnings ~ hours + sqhours + male + male*hours, cps),
       error_format = "({p.value})",
       error_pos = "below".
       statistics = c(N = "nobs", "R2 adj." = "adj.r.squared"),
       stars = c(`***` = 0.01, `**` = 0.05, `*` = 0.1),
        coefs = c("Hours worked" = "hours",
                 "(Hours worked)<sup>2</sup>" = "sqhours",
                 "Male" = "male".
                 "Hours worked x Male" = "hours:male",
                 "Constant" = "(Intercept)"),
       note = "P-values in parentheses. {stars}",
       align = "c") %>%
  insert_row(c("", "Dependent variable: Log weekly earnings", rep("", 3))) %>%
 merge_cells(1, 2:5) %>%
 set_align(1, 2, "center")
```

You can do many more things with <a href="huxreg()">huxreg()</a> and <a href="huxtable()">huxtable()</a>. See here for more details.

	(1)	(2)	(3)	(4)
Hours worked	0.052 ***	0.093 ***	0.093 ***	0.093 ***
	(0.000)	(0.000)	(0.000)	(0.000)
(Hours worked) <sup>2</sup>		-0.001 ***	-0.001 ***	-0.001 ***
		(0.000)	(0.000)	(0.000)
Male			0.053 ***	-0.235 ***
			(0.000)	(0.000)
Hours worked x Male				0.008 ***
				(0.000)
Constant	4.678 ***	4.109 ***	4.095 ***	4.171 ***
	(0.000)	(0.000)	(0.000)	(0.000)
N	4609	4609	4609	4609
R2 adj.	0.586	0.621	0.622	0.625
P-values in narenthe	2000 *** n	< 0 01· **	n < 0 05· *	n < 0 1

PLOT COEFFICIENTS It might be useful to provide a graphical representation of the coefficients

You can do it with dplyr, gplot and extracting elements from the regression object with \$
 -We can use a shortcut using the plot\_summs() function from the {jtools} package. You can:
 Include one or many models

•

\_

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• Specify coefficients to omit from the plot

•

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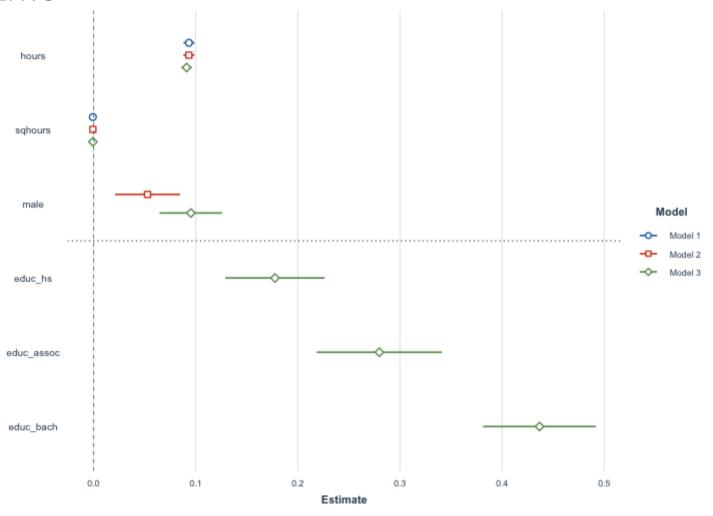
- Include one or many models
- Specify coefficients to omit from the plot
- Change the level of the confidence intervals
- Choose a color palette
- Other functions from ggplot ...

# **AGENDA**

- REGRESSIONS IN R
  - CONTINUOUS VARIABLES
  - CATEGORICAL AND BINARY VARIABLES
- GETTING TO THE RIGHT MODEL
  - VARIABI F TRANSFORMATION
  - FUNCTIONAL FORM
  - CONTROL VARIABLES
  - INTERACTION TERMS

- INFERENCE
- EXPORTING RESULTS
  - REGRESSION TABLES
  - COFFFICIENT PLOTS
- PRACTICING YOUR SKILLS

### **PLOT COEFFICIENTS**



- How can you work on improve your programming skills throughout the year?
- A research project is too demanding and you're not ready with the basics yet! Replicate a paper:-Browse the recent issues (last 3 years) of Economics journals.
  - **General interest:** American Economic Review, The Quarterly Journal of Economics, Econometrica, The Review of Economic Studies, Economic Journal
  - o **Fields:** Journal of Political Economy, Journal of Development Economics, Labor Economics, Journal of Public Economics
- Many of the papers now have a replication package with data and code that you can download- We want to ensure reproduciblity and replicability
  - o Computational reproducibility: the code runs and produces the same results
  - Replicability: ability to replicate results from scratch
- See here for reports on previous replications by the Institute for Replication and papers that need a replicator.
- Sign-up for an edition of the Replication Games organized by the Institute for Replication.



Cooperative Property Rights and Development:

Evidence from Land Reform in El Salvador:

A Comment\*

Anders Kjelsrud<sup>†</sup> Andreas Kotsadam<sup>‡</sup> Ole Rogeberg<sup>§</sup>

March, 2023 (updated version, March 10)\*\*

#### Abstract

Montero (2022) explores a discontinuity in a land reform in El Salvador and reports two main findings. First, relative to outside-owned haciendas operated by contract workers, the productivity of worker-owned cooperatives is higher for staple crops and lower for cash-crop. Second, cooperative property rights increase workers' incomes and compress wage distributions. In this comment, we show that the latter result rests on two mistakes: three-quarters of the observations are duplicates and income inequality is calculated over too few workers to be meaningful. When corrected, the data sources and research design provide no credible evidence regarding the causal effects of ownership structure on income levels and inequality.

#### **Upcoming Games: Registration Open**

**Lyon Replication Games**: October 24th, 2023 at ENS de Lyon, France. Economics and political science papers to be reproduced/replicated. Local organizer is Mathieu Couttenier.

**Stockholm Replication Games**: October 26th, 2023 in Stockholm, Sweden. Behavioral science, economics and political science papers to be reproduced/replicated. Local organizer is I4R's codirector Anna Dreber.

**Brussels Replication Games**: October 27th, 2023 at ULB-ECARES, Belgium. Economics and political science papers to be reproduced/replicated. Local organizers are Paula Gobbi and Joanne Haddad.

#### **Upcoming Games: Registration Closed**

None.

#### **Upcoming Games: Registration Not Open Yet**

**Toronto Replication Games**: February 20th 2024 at the University of Toronto, Canada. This is a collaboration with the University of Toronto's Data Sciences Institute. Local organizer is once again Rohan Alexander.

**Los Angeles Replication Games**: February 28th 2024 at University of California, Los Angeles. Economics and political science papers to be replicated.

**Berkeley Replication Games**: March 7th 2024 at University of California, Berkeley. Social science papers to be replicated. Local organizers are Edward Miguel and Fernando Fernando Hoces de la Guardia.

**Tokyo Replication Games**: Date tbd, at the University of Tokyo campus, Japan. Economics and political science papers to be replicated. Local organizers are Yasuyuki Sawada, Chishio Furukawa and Hiroki Kameyama.

### RESOURCES

- Causal Inference: The Mixtape by Scott Cunningham.
  - Great book on econometrics with a special focus on causal inference (great for us economists!).
    - It has code examples in R, Stata and Python for each chapter.
    - Good for finding the right commands for each method (instrumental variables, regression discontinuity, etc.)
- Mostly Harmless Econometrics by Joshua D. Angrist and Jörn-Steffen Pischke
  - Best book to get "the intuition".
  - Very useful for Econometrics 3
- Introduction to Econometrics with R course by Florian Oswald, Vincent Viers, Jean-Marc Robin, Pierre Villedieu, Gustave Kenedi at Sciences Po
- A crowd-sourced checklist of the top 10 little things that drive us crazy with regression output
  - Things you need to include in your regression tables