





DECIDE

Introduction to Health Interventions, Policy and Services

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Introduction to Health Interventions, Policy and Services

Methods in Evidence Synthesis – Part III

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Summary

- Meta-analysis for binary and continuous outcomes using R:
 - metafor package:
 - Calculating effect sizes
 - Performing meta-analysis
 - Interpreting and plotting results
 - Exploring sources of heterogeneity
 - meta package:
 - Calculating effect sizes
 - Performing meta-analysis
 - Interpreting and plotting results
 - Exploring sources of heterogeneity







metafor package

We will use metafor as the main package to perform metaanalysis for binary and continuous outcomes

> install.packages("metafor")
library(metafor)





Meta-analysis for binary outcomes: BCG dataset

Includes 13 primary studies assessing the effectiveness of the BCG vaccine against tuberculosis

Two groups (vaccinated versus non-vaccinated) being compared

Binary outcome (tuberculosis diagnosis - positive vs negative)

	Tuberculosis + Tuberculosis -	
Vaccinated group	tpos	tneg
Control group	cpos	cneg









Meta-analysis for binary outcomes: BCG dataset

	Tuberculosis + Tuberculosis -	
Vaccinated group	tpos	tneg
Control group	cpos	cneg

Effect measures that could be used to assess the effectiveness of BCG vaccine: risk ratio, odds ratio, risk difference...







Meta-analysis for binary outcomes: Effect size calculation

We will perform meta-analysis to estimate the pooled risk ratio

As we only have raw data, we firstly need to calculate the effect size and sampling variance for each primary study.

For that, we will use the escalc() function

Let's start to check its arguments!

> ?escalc









```
escalc(measure, formula, ai, bi, ci, di, nli, n2i, xli, x2i, tli, t2i, mli, m2i, sdli, sd2i, xi, mi, ri, ti, sdi, r2i, ni, yi, vi, sei, data, slab, subset, add=1/2, to="only0", drop00=FALSE, vsvpe="LS", var.names=c("yi","vi"), add.measure=FALSE, append=TRUE, replace=TRUE, digits=4, ...)
```

The options for the measure argument are then:

- "RR" for the log risk ratio.
- "OR" for the log odds ratio.
- "RD" for the risk difference.
- "AS" for the arcsine square root transformed risk difference (Rücker et al., 2009).
- "PETO" for the log odds ratio estimated with Peto's method (Yusuf et al., 1985).

Measures for Dichotomous Variables

In various fields (such as the health and medical sciences), the response or outcome variable measured is often dichotomous (binary), so that the data from a study comparing two different groups can be expressed in terms of a 2x2 table, such as:

where ai, bi, ci, and di denote the cell frequencies (i.e., the number of people falling into a particular category) and nli and n2i the row totals (i.e., the group sizes).

	Tuberculosis +	Tuberculosis -
Vaccinated group	tpos	t <mark>n</mark> eg
Control group	cpos	cneg







Meta-analysis for binary outcomes: Effect size calculation

The bcg dataset now contains two more variables:

- yi computed effect sizes
- vi computed sampling variances

Please note that in this dataset we have no group with zero events. Otherwise, we would have to include additional arguments...







Effect size calculation

The existence of cells with zero counts is particularly troublesome:

	Event	No event
Experimental	0	b
Control	С	d

$$RR = \frac{\frac{0}{0+b}}{\frac{c}{c+d}} = 0 \ln(0) = -\infty \qquad SE_{lnRR} = \sqrt{\frac{b}{0} + \frac{d}{c(c+d)}}$$

$$SE_{lnRR} = \sqrt{\frac{b}{0} + \frac{d}{c(c+d)}}$$

$$OR = \frac{0 \times d}{b \times c} = 0 \ln(0) = -\infty \qquad SE_{lnOR} = \sqrt{\frac{1}{0} + \frac{1}{b} + \frac{1}{c} + \frac{1}{d}}$$

$$SE_{lnOR} = \sqrt{\frac{1}{0} + \frac{1}{b} + \frac{1}{c} + \frac{1}{d}}$$

To circumvent this issue, it is possible to add a small increment – known as continuity correction - so that there are no cells with zero counts left.







Effect size calculation

In the escalc() function, the add= and to= arguments are used to define continuity corrections:

- add= indicates the amount to add to zero cells. The default value is 0.5
- to= indicates to which cells the continuity correction is to be added:
 - "only0": Default option. Continuity correction is only added to cells with zero counts
 - "all": Continuity correction is added to all cells
 - "if0all" : Continuity correction is added to all cells if there are any with zero counts
 - "none": Continuity correction is not added







Performing meta-analysis

To perform meta-analysis, we apply rma function.

We only need to define the following arguments:

- data: Dataset
- yi: Effect size variable
- vi: Variance variable
- slab: Primary studies identification
- method: Defines the model and, in case of a randomeffects model, the method for estimating heterogeneity

- method="FE" = fixed-effects model
- method="DL" = DerSimonian-Laird estimator
- method="HE" = Hedges estimator
- method="HS" = Hunter-Schmidt estimator
- method="SJ" = Sidik-Jonkman estimator
- method="ML" = maximum-likelihood estimator
- method="REML" = restricted maximum-likelihood estimator
- method="EB" = empirical Bayes estimator
- method="PM" = Paule-Mandel estimator
- method="GENQ" = generalized Q-statistic estimator





Meta-analysis for binary outcomes: Meta-analysis

```
# Meta-analysis following a random-effects model with the
restricted maximum-likelihood estimator method for the amount of
heterogeneity
> ma01 <- rma(yi=yi, vi=vi, data=bcg, slab=paste(author, year,
sep=", "), method="REML")</pre>
```

```
# In fact, we could have done it all in just one step!
> ma01 <- rma(measure="RR", ai=tpos, bi=tneg, ci=cpos, di=cneg, data=bcg, slab=paste(author, year, sep=", "), method="REML")</pre>
```







```
> ma01
Random-Effects Model (k = 13; tau<sup>2</sup> estimator: REML)
 logLik deviance AIC BIC AICC
-12.2024 24.4047 28.4047 29.3746 29.7381
tau^2 (estimated amount of total heterogeneity): 0.3132 (SE = 0.1664)
tau (square root of estimated tau^2 value): 0.5597
I^2 (total heterogeneity / total variability): 92.22%
H^2 (total variability / sampling variability): 12.86
Test for Heterogeneity:
Q(df = 12) = 152.2330, p-val < .0001
Model Results:
estimate se zval pval ci.lb ci.ub
 -0.7145 0.1798 -3.9744 <.0001 -1.0669 -0.3622
                                                 ***
```





```
> ma01
                                          Number of primary studies
Estimates for the log(risk ratio) and for the corresponding lower
and upper bounds of the confidence interval. To convert these
values back into the natural scale, the exp() function must be applied.
O(df = 12) = 152.2330
estimate
                   p-value
                                          -1.0669
                                                     -0.3622
                                                                * * *
```







Meta-analysis for binary outcomes: Results interpretation

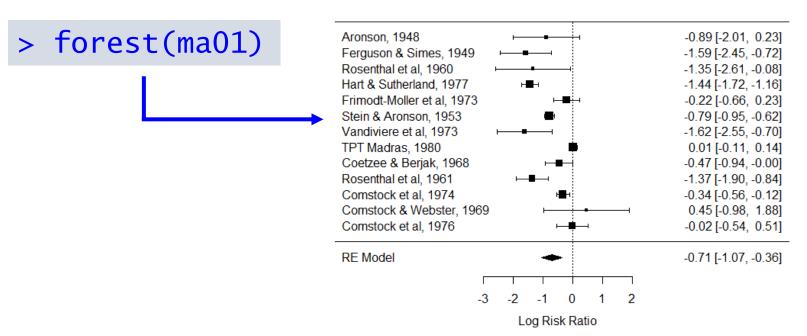
```
# For obtaining the pooled estimate and the respective
confidence interval bounds transformed into the natural scale:
> exp(c(ma01$beta, ma01$ci.lb, ma01$ci.ub))
[1] 0.4894209 0.3440743 0.6961661
```

We observed that the BCG associates with decreased risk of tuberculosis (RR=0.5; 95%CI=0.3-0.7; p<0.001)



Meta-analysis for binary outcomes: Forest plot

The **forest()** function creates a forest plot. The only mandatory argument is the meta-analysis object.



What can be done to improve this plot?







Meta-analysis for binary outcomes: Forest plot

```
# "alim" corresponds to the limit of the xx axis of the forest plot;
"xlim" concerns the width of the plot; "at" locates the xx axis marks;
"refline" locates the reference line.
> forest.rma(ma01, transf=exp, showweights=TRUE, alim =c(0,2), xlim=c(-
2.5,4.5), at=c(0,0.5,1,1.5,2), order = order(bcg\$year), refline=1)
# To add a top legend
text(-2.5,15, "Author(s) and Year", pos=4)
text(3,15, "Weight", pos=2)
text(4.5,15, "RR [95% CI]", pos=2)
# To add text for the effect size p-value and for heterogeneity
text(3.7, -1.85, pos=4, cex=0.8, bquote(paste("p <0.001")))
text(-2.5, -1.85, pos=4, cex=0.8, bquote(paste("Q Cochran p<0.001", ";
", I^2, " = ", .(formatC(ma01$I2, digits=1, format="f")), "%")))
```







Author(s) and Year		Weight	RR [95% CI]	
Aronson, 1948	⊢ ■	5.06%	0.41 [0.13, 1.26]	
Ferguson & Simes, 1949	⊢■		0.20 [0.09, 0.49]	
Stein & Aronson, 1953	■ 4	10.10%	0.46 [0.39, 0.54]	
Rosenthal et al, 1960	⊢ •	4.44%	0.26 [0.07, 0.92]	
Rosenthal et al, 1961	₽	8.37%	0.25 [0.15, 0.43]	
Coetzee & Berjak, 1968	⊢ ■	8.74%	0.63 [0.39, 1.00]	
Comstock & Webster, 1969	-	3.82%	1.56 [0.37, 6.53]	
Frimodt-Moller et al, 1973	⊢ ■÷	8.87%	0.80 [0.52, 1.25]	
Vandiviere et al, 1973	H■────────────────────────────────────	6.03%	0.20 [0.08, 0.50]	
Comstock et al, 1974	⊦⊞⊸	9.93%	0.71 [0.57, 0.89]	
Comstock et al, 1976	- <u>=</u>	8.40%	0.98 [0.58, 1.66]	
Hart & Sutherland, 1977	■H	9.70%	0.24 [0.18, 0.31]	
TPT Madras, 1980	· •	10.19%	1.01 [0.89, 1.14]	
RE Model	•	100.00%	0.49 [0.34, 0.70]	
Q Cochran p<0.001; $I^2 = 92.2\%$			p < 0.001	
	0 0.5 1 1.5 2			
Risk Ratio				

forestplot package can be used to further personalize the presentation of forest plots – you can find an online script guiding its use









Meta-analysis for binary outcomes: Heterogeneity

```
> ma01
                                                                I^2 statistic
I^2 (total heterogeneity / total variability):
Q(df = 12) = 152.2330,
                        p-val < .0001
                                              p-value of the Cochran Q test
```









Meta-analysis for binary outcomes: Sensitivity analysis

Leave-one-out sensitivity analysis is a special form of sensitivity analysis consisting in removing each study at a time. The leavelout function allows for such type of sensitivity analysis.

```
#The leavelout function can be applied to the meta-analysis model:
```

> leavelout(ma01)

Each output line corresponds to the results obtained when the respective study was not included in the meta-analysis







Meta-analysis for binary outcomes: Meta-regression

To perform meta-regression, we need to introduce covariates after the effect size and the character ~

```
# Univariable meta-regressions with different covariates
(latitude, year and allocation):
> mreg01_lat <- rma(yi=yi~ablat, vi=vi, data=bcg, method="REML")
> mreg01_year <- rma(yi=yi~year, vi=vi, data=bcg, method="REML")
> mreg01_alloc <- rma(yi=yi~alloc, vi=vi, data=bcg, method="REML")</pre>
```







```
# Results of the univariable meta-regression analysis with the latitude as covariate:
> mreg01_lat
Mixed-Effects Model (k = 13; tau<sup>2</sup> estimator: REML)
tau^2 (estimated amount of residual heterogeneity):
                                                       0.0764 \text{ (SE} = 0.0591)
tau (square root of estimated tau^2 value):
                                                       0.2763
I^2 (residual heterogeneity / unaccounted variability): 68.39%
H^2 (unaccounted variability / sampling variability):
                                                       3.16
R^2 (amount of heterogeneity accounted for):
                                                       75.62%
Test for Residual Heterogeneity:
QE(df = 11) = 30.7331, p-val = 0.0012
Test of Moderators (coefficient(s) 2):
QM(df = 1) = 16.3571, p-val < .0001
Model Results:
        estimate se zval pval ci.lb ci.ub
intrcpt 0.2515 0.2491 1.0095 0.3127 -0.2368 0.7397
ablat -0.0291
                 0.0072 - 4.0444 < .0001 - 0.0432
                                                    -0.0150
```





```
# Results of the univariable meta-regression analysis with the latitude as covariate:
> mreg01_lat
Mixed-Effects Model (k = 13; tau<sup>2</sup> estimator: REML)
tau^2 (estimated amount of residual heterogeneity):
tau (square root of estimated tau^2 value):
   Exponentials of meta-regression coefficients can be interpreted
   as odds ratio (RR). In this case, RR=0.97 (95%CI=0.96-0.99)
Test of Moderators (coefficies
QM(df = 1) = 16.3571, p-val < .0001
         -0.0291 0.0072
                                     <.0001
ablat
                                                       -0.0150
                                                                ***
```









Meta-analysis for binary outcomes: Subgroup analysis

To perform subgroup analysis, we need to define the argument subset= in the rma() function

```
# Subgroup analysis in relation to the allocation type:
> ma01_random <- rma(yi=yi, vi=vi, data=bcg, method="REML",
slab=paste(author, year, sep=", "), subset=alloc=="random")
> ma01_alternate <- rma(yi=yi, vi=vi, data=bcg, method="REML",
slab=paste(author, year, sep=", "), subset=alloc=="alternate")
> ma01_system <- rma(yi=yi, vi=vi, data=bcg, method="REML",
slab=paste(author, year, sep=", "), subset=alloc=="systematic")</pre>
```





```
# Results for the subgroup of primary studies with random allocation:
> ma01_random
Random-Effects Model (k = 7; tau<sup>2</sup> estimator: REML)
tau^2 (estimated amount of total heterogeneity): 0.3925 (SE = 0.3029)
tau (square root of estimated tau^2 value):
                                          0.6265
I^2 (total heterogeneity / total variability): 89.93%
H^2 (total variability / sampling variability): 9.93
Test for Heterogeneity:
Q(df = 6) = 110.2133, p-val < .0001
Model Results:
estimate se zval pval ci.lb ci.ub
 -0.9710 0.2760 -3.5186 0.0004 -1.5118 -0.4301
```









Meta-analysis for binary outcomes: Funnel plot

The possibility of publication bias can be assessed by means of the rank correlation test of Begg ranktest() and of funnel plot visual inspection.

> ranktest(ma01)

Rank Correlation Test for Funnel Plot Asymmetry Kendall's tau = 0.0256, p = 0.9524









```
# Obtention of the funnel plot
> funnel(ma01)
# Obtention of the funnel plot with trim-and-fill
> (ma01_tf <- trimfill(ma01))</pre>
  funnel(ma01_tf)
Estimated number of missing studies on the right side: 1 \text{ (SE = 2.4528)}
Random-Effects Model (k = 14; tau<sup>2</sup> estimator: REML)
tau^2 (estimated amount of total heterogeneity): 0.3313 (SE = 0.1701)
tau (square root of estimated tau^2 value):
                                           0.5756
I^2 (total heterogeneity / total variability): 92.14%
H^2 (total variability / sampling variability): 12.72
Test for Heterogeneity:
Q(df = 13) = 154.6750, p-val < .0001
Model Results:
estimate se zval pval ci.lb ci.ub
 -0.6571 0.1785 -3.6805 0.0002 -1.0070 -0.3072 ***
```

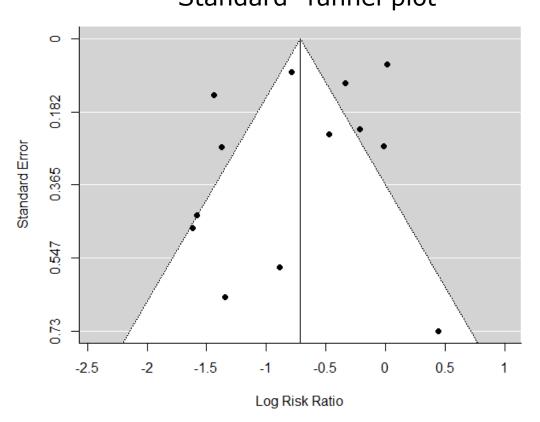




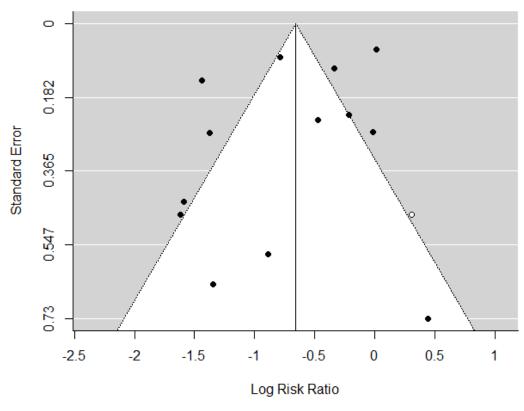


Meta-analysis for binary outcomes: Funnel plot





Funnel plot with trim-and-fill









Meta-analysis for continuous outcomes: stroke dataset

Includes 9 primary studies comparing the length of stay between stroke patients receiving specialized care *versus* those receiving routine care

Two groups (specialized versus routine care) being compared

Continuous outcome: Length of stay (LOS)

Effect measures that could be used: mean difference, standardized mean difference...





Meta-analysis for continuous outcomes: Effect size calculation

As we only have raw data, we firstly need to calculate the effect size and sampling variance for each primary study.

Once again, we will use the escalc() function

Let's check again its arguments!

> ?escalc









```
escalc(measure, formula, ai, bi, ci, di, nli, n2i, xli, x2i, tli, t2i, mli, m2i, sdli, sd2i, xi, mi, ri, t1, sdi, r2i, ni, yi, vi, sei, data, slab, subset, aid=1/2, to="only0", drop00=FALSE, vtype="LS", var.names=c("yi","vi"), add.measure=FALSE, append=TRUE, replace=TRUE, digits=1, ...)
```

The options for the measure argument are then:

- "MD" for the raw mean difference.
- "SMD" for the standardized mean difference.
- "SMDH" for the standardized mean difference with heteroscedastic population variances in the two groups (Bonett, 2008, 2009).
- "ROM" for the log transformed ratio of means (Hedges et al., 1999; Lajeunesse, 2011).

Measures for Quantitative Variables

When the response or dependent variable assessed in the individual studies is measured on some quantitative scale, it is customary to report certain summary statistics, such as the mean and standard deviation of the scores. The data layout for a study comparing two groups with respect to such a variable is then of the form:

mean standard deviation group size

group 1 mli	sdli	nli
group 2 m2i	sd2i	n2i

	Mean LOS	Standard-deviation	Group size
Specialized care	los_spec	sd_spec /	n_spec
Routine care	los_rout 🗲	sd_rout 🗡	n_rout *







Meta-analysis for continuous outcomes: Effect size calculation

```
# Computation of the effect size (mean difference) and of
the corresponding sampling variance for each primary study
> stroke <- escalc(measure="MD", n1i=n_spec, m1i=los_spec,
sd1i=sd_spec, n2i=n_rout, m2i=los_rout, sd2i=sd_rout,
data=stroke)</pre>
```

The stroke dataset now contains two more variables:

- yi computed effect sizes
- vi computed sampling variances







Meta-analysis for continuous outcomes: Meta-analysis

```
# Meta-analysis following a random-effects model with the restricted
maximum-likelihood estimator method for the amount of heterogeneity
> ma02 <- rma(yi=yi, vi=vi, data=stroke, slab=paste(study, year,
sep=", "), method="REML")</pre>
```

```
# Once again, we could have done it all in just one step!
> ma02 <- rma(measure="MD", n1i=n_spec, m1i=los_spec, sd1i=sd_spec,
n2i=n_rout, m2i=los_rout, sd2i=sd_rout, data=stroke, slab=paste(study,
year, sep=", "), method="REML")</pre>
```









```
> ma02
Random-Effects Model (k = 9; tau<sup>2</sup> estimator: REML)
tau^2 (estimated amount of total heterogeneity): 684.6462 (SE = 359.7541)
tau (square root of estimated tau^2 value):
                                               26.1657
I^2 (total heterogeneity / total variability):
                                               98.97%
H^2 (total variability / sampling variability): 97.21
Test for Heterogeneity:
Q(df = 8) = 238.9158, p-val < .0001
Model Results:
estimate se zval pval ci.lb ci.ub
-15.1060 8.9466 -1.6885 0.0913 -32.6409 2.4289
```

Specialized care associates with a trend towards decreased LOS (pooled meta-analytic estimate: -15.1 days; 95%CI=-32.6;2.43; p=0.091). However, severe heterogeneity was observed ($p<0.001; I^2=99.0\%$)







Meta-analysis for continuous outcomes: Forest plot

```
# Forest plot presenting relative weights, p-value for the effect size and
heterogeneity statistics.
> forest.rma(ma02, showweights=TRUE, order = order(stroke$year), xlim=c(-
350,350))
text(-350,10.5, "Author(s) and Year", pos=4)
text(150,10.5, "Weight", pos=2)
text(350,10.5, "RR [95% CI]", pos=2)
text(265, -1.85, pos=4, cex=0.8, bquote(paste("p = ", .(formatC(ma02$pval,
digits=3, format="f")))))
text(-350, -1.8, pos=4, cex=0.8, bquote(paste("Q Cochran p<0.001", "; ",
I^2, " = ", .(formatC(ma01$I2, digits=1, format="f")), "%")))
```

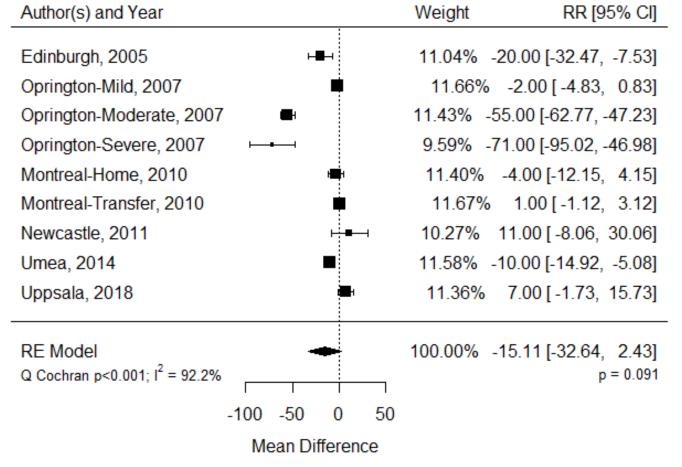








Meta-analysis for continuous outcomes: Forest plot









Other packages: meta package

There are other packages to perform "classical" meta-analysis.

We will redo our analyses using meta package.

```
> install.packages("meta")
library(meta)
```

You can also find an online script guiding the use of rmeta.







Meta-analysis for binary outcomes: Meta-analysis

Again, we will use bcg dataset to perform meta-analysis to estimate the pooled risk ratio

Given the raw data, meta-analysis of binary outcomes can be performed using the metabin() function

Let's start to check its arguments!

> ?metabin





```
metabin(event.e, n.e, event.c, n.g, studlab,
        level=gs("level"), level.comb=gs("level.comb"),
        comb.fixed=gs("comb.fixed"), comb.random=gs(
        hakn=gs("hakn"),
                is.na(charmatch(tolower(method), "glmm", nomatch
                                       evel.predict=gs("level.predict"),
        label.e=gs("label.e"), label.c=gs("label.c")
        label.left=gs("label.left"), label.right=gs("label.ri
        byvar, bylab, print.byvar=gs("print.byvar"),
        byseparator = gs("byseparator"),
        print.CMH=gs("print.CMH"),
        keepdata=gs("keepdata"),
        warn=gs("warn"),
        control=NULL.
```

	Tuberculosis +	Tuberculosis -	Total
Vaccinated group	tpos	tneg	tpos + tneg
Control group	cpos	cneg	cpos + cneg

Pooling method

Effect measure

Method to estimate tau – DerSimonian and Laird is the default method. REML and other methods rely on metafor package.

With meta package, fixed effects and random effects meta-analysis results are simultaneously presented







If there are cells with zero counts, the following function may be defined:

• incr= indicates the amount to add to zero cells. It can be defined as a numerical value or as "TACC" (treatment arm continuity correction - to correct for unbalanced sample sizes in the experimental *versus* control group). The default value is 0.5.

By default, continuity correction is added only to zero cells. This can be modified by setting as TRUE one of the following arguments:

- allincr= If TRUE, continuity correction is added to all cells if there are any zero counts.
- addincr= If TRUE, continuity correction is added to all cells.







Meta-analysis for binary outcomes: Meta-analysis

```
# Meta-analysis weighting by the inverse variance and with the restricted maximum-likelihood estimator method for the amount of heterogeneity. The prediction argument allows for obtention of estimates of a hypothetical future study.
```

```
> ma04 <- metabin(data=bcg, event.e=tpos, n.e=tpos+tneg,
event.c=cpos, n.c=cpos+cneg, method="Inverse", sm="RR",
method.tau="REML", prediction=TRUE, studlab=paste(author,year))</pre>
```





```
> ma04
                               RR 95%-CI %W(fixed) %W(random)
                           0.4109 [0.1343; 1.2574] 0.5
Aronson 1948
                                                                      5.1

      Ferguson & Simes 1949
      0.2049 [0.0863; 0.4864]
      0.8
      6.4

      Rosenthal et al 1960
      0.2597 [0.0734; 0.9186]
      0.4
      4.4

Hart & Sutherland 1977 0.2366 [0.1793; 0.3121] 8.2 9.7
Frimodt-Moller et al 1973 0.8045 [0.5163; 1.2536] 3.2 8.9
Number of studies combined: k = 13
                          RR 95%-CI z p-value
Fixed effect model 0.6503 [0.6007; 0.7040] -10.62 < 0.0001
Random effects model 0.4894 [0.3441; 0.6962] -3.97 < 0.0001
Prediction interval [0.1342; 1.7848]
Quantifying heterogeneity:
tau^2 = 0.3132; H = 3.56 [2.93; 4.34]; I^2 = 92.1\% [88.3\%; 94.7\%]
Test of heterogeneity:
      Q d.f. p-value
152.23 12 < 0.0001
Details on meta-analytical method:
- Inverse variance method
- Restricted maximum-likelihood estimator for tau∧2
```



```
> ma04
metafor package
                                               Estimates for the risk ratio and for the
                                           corresponding confidence interval and p-value
             95%-CI
                         z p-value
 Fixed effect model
                    0.6503 [0.6007; 0.7040] -10.62 < 0.0001
Random effects model 0.4894 [0.3441; 0.6962] -3.97 < 0.0001
                                                                   Estimates for the risk
Prediction interval
                           0.1342; 1.7848
                                                                   ratio of a future study
Quantifying heterogeneity:
tau^2 = 0.3132; H = 3.56 [2.93; 4.34]; I<sup>2</sup> = 92.1% [88.3%; 94.7%]
Test of heterogeneity:
      O d.f. p-value
 Details on meta-analytical method:
                               Estimates of the I^2 and of p-value for the Q Cochran test
```



Meta-analysis for binary outcomes: Forest plot

The function forest() creates a forest plot.

The "common=FALSE" expression is used to hide fixed effects model results. To hide random effects model results, we would use the expression "random=FALSE". The "prediction" argument orders the presentation of the predicted interval. The number of decimal places is set by the "digits" argument.

> forest(ma04, common=FALSE, prediction = TRUE, digits = 1)







	Experimental Co		Control						
Study	Events	Total	Events	Total	Risk F	Ratio	RR	95%-CI	Weight
Aronson 1948	4	123	11	139		_	0.4	[0.1; 1.3]	5.1%
Ferguson & Simes 1949	6	306	29	303	-			[0.1; 0.5]	6.4%
Rosenthal et al 1960	3	231	11	220				[0.1; 0.9]	4.4%
Hart & Sutherland 1977	62	13598	248	12867	-			[0.2; 0.3]	9.7%
Frimodt-Moller et al 1973	33	5069	47	5808	_	_		[0.5; 1.3]	8.9%
Stein & Aronson 1953	180	1541	372	1451	-			[0.4; 0.5]	10.1%
Vandiviere et al 1973	8	2545	10	629				[0.1; 0.5]	6.0%
TPT Madras 1980	505	88391	499	88391				[0.9; 1.1]	10.2%
Coetzee & Berjak 1968	29	7499	45	7277	-		0.6	[0.4; 1.0]	8.7%
Rosenthal et al 1961	17	1716	65	1665				[0.1; 0.4]	8.4%
Comstock et al 1974	186	50634	141	27338	-		0.7	[0.6; 0.9]	9.9%
Comstock & Webster 1969	5	2498	3	2341	+	-	1.6	[0.4; 6.5]	3.8%
Comstock et al 1976	27	16913	29	17854	-	—	1.0	[0.6; 1.7]	8.4%
Random effects model		191064		166283	\Diamond		0.5	[0.3; 0.7]	100.0%
Prediction interval								[0.1; 1.8]	
Heterogeneity: $I^2 = 92\%$, $\tau^2 =$	0.3132, <i>p</i>	< 0.01			1 1	I I			
					0.1 0.5 1	2 10			







Meta-analysis for binary outcomes: Meta-regression

To perform meta-regression, we need to apply the metareg() function to an already existent meta-analysis object

```
# Univariable meta-regressions with different covariates (latitude, year
and allocation):
> mreg04_lat <- metareg(ma04, ablat)
> mreg04_year <- metareg(ma04, year)
> mreg04_alloc <- metareg(ma04, alloc)
# Multivariable meta-regression with latitude and year as covariates:
> mreg04_multi <- metareg(ma04, ablat+year)</pre>
```









```
# Results of the univariable meta-regression analysis with the latitude as covariate:
> mreg04_lat
Mixed-Effects Model (k = 13; tau<sup>2</sup> estimator: REML)
tau^2 (estimated amount of residual heterogeneity):
                                                        0.0764 \text{ (SE} = 0.0591)
tau (square root of estimated tau^2 value):
                                                        0.2763
I^2 (residual heterogeneity / unaccounted variability): 68.39%
H^2 (unaccounted variability / sampling variability):
                                                        3.16
R^2 (amount of heterogeneity accounted for):
                                                        75.62%
Test for Residual Heterogeneity:
QE(df = 11) = 30.7331, p-val = 0.0012
Test of Moderators (coefficient(s) 2):
QM(df = 1) = 16.3571, p-val < .0001
                                             meta relies on metafor to perform meta-regression
Model Results:
         estimate se zval pval ci.lb ci.ub
intrcpt 0.2515 0.2491 1.0095 0.3127 -0.2368 0.7397
ablat -0.0291
                  0.0072 - 4.0444 < .0001 - 0.0432
                                                     -0.0150
```

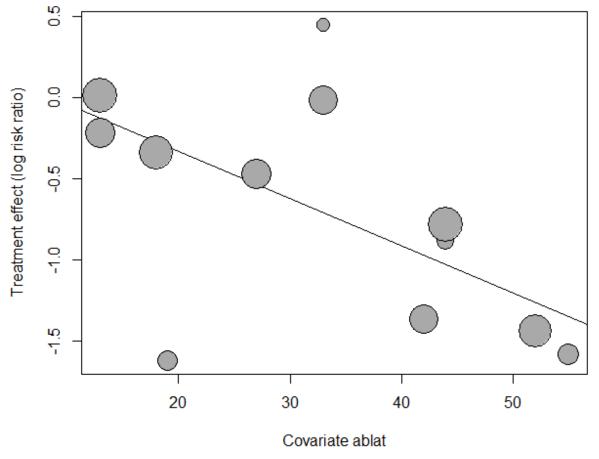








- # Plotting the results of a meta-regression for a continuous covariate:
- > bubble(mreg04_lat)











Meta-analysis for binary outcomes: Subgroup analysis

To perform subgroup analysis, we use the byvar argument

```
# Subgroup analysis (based on the allocation type) in a meta-
analysis de novo:

> ma04_sub <- metabin(data=bcg, event.e=tpos, n.e=tpos+tneg,
event.c=cpos, n.c=cpos+cneg, method="Inverse", sm="RR",
method.tau="REML", prediction=TRUE, studlab=paste(author,year),
byvar=alloc)

# Update of an existent meta-analysis
> ma04_sub <- update(ma04, byvar=alloc)</pre>
```





```
# Results for the subgroup analysis based on the allocation:
> ma04_sub
                              95%-CI %W(fixed) %W(random) alloc
                                                          5.1 random
Aronson 1948
              0.4109 [0.1343; 1.2574]
                                                0.5
Ferguson & Simes 1949
                      0.2049 [0.0863; 0.4864] 0.8 6.4
                                                                 random
Results for subgroups (random effects model):
                                  95%-CT
                                             0 tau^2 I^2
                       RR
alloc = random 7 0.3787 [0.2205; 0.6504] 110.21 0.3925 94.6%
alloc = alternate 2 0.5823 [0.3353; 1.0112] 5.56 0.1326 82.0%
alloc = systematic 4 0.6543 [0.3233; 1.3243] 16.59 0.4003 81.9%
Test for subgroup differences (random effects model):
                 Q d.f. p-value
Between groups 1.86
                   2 0.3943
```

In the output, both fixed effects and random effects meta-analysis results are presented









Plotting the results of a subgroup analysis:

> forest(ma04_sub, prediction=FALSE, common=FALSE,
xlim=c(0.05,10), digits=1, digits.sd=1)

Experimental Control										
Study	Events	Total	Events	Total	Risk Ratio	RR	95%-CI	Weight		
alloc = random										
Aronson 1948	4	123	11	139		0.4	[0.1; 1.3]	5.1%		
Ferguson & Simes 1949	6	306	29	303	-		[0.1; 0.5]	6.4%		
Rosenthal et al 1960	3	231	11	220	-		[0.1; 0.9]	4.4%		
Hart & Sutherland 1977	62	13598	248	12867	-		[0.2; 0.3]	9.7%		
Vandiviere et al 1973	8	2545	10	629	-	0.2	[0.1; 0.5]	6.0%		
TPT Madras 1980	505	88391	499	88391		1.0	[0.9; 1.1]	10.2%		
Coetzee & Berjak 1968	29	7499	45	7277	-	0.6	[0.4; 1.0]	8.7%		
Random effects model		112693		109826	←	0.4	[0.2; 0.7]	50.5%		
Heterogeneity: $I^2 = 95\%$, $\tau^2 =$	0.3925, p	< 0.01								
alloc = alternate										
Frimodt-Moller et al 1973	33	5069	47	5808		8.0	[0.5; 1.3]	8.9%		
Stein & Aronson 1953	180	1541	372	1451		0.5	[0.4; 0.5]	10.1%		
Random effects model		6610		7259	-	0.6	[0.3; 1.0]	19.0%		
Heterogeneity: $I^2 = 82\%$, $\tau^2 =$	0.1326, <i>p</i>	= 0.02								
alloc = systematic										
Rosenthal et al 1961	17	1716	65	1665		0.3	[0.1; 0.4]	8.4%		
Comstock et al 1974	186	50634	141	27338	<u> </u>	0.7	[0.6; 0.9]	9.9%		
Comstock & Webster 1969	5	2498	3	2341	 *	1.6	[0.4; 6.5]	3.8%		
Comstock et al 1976	27	16913	29	17854		1.0	[0.6; 1.7]	8.4%		
Random effects model		71761		49198	-	0.7	[0.3; 1.3]	30.5%		
Heterogeneity: $I^2 = 82\%$, $\tau^2 =$	0.4003, p	< 0.01								
Random effects model		191064		166283	*	0.5	[0.3; 0.7]	100.0%		
Heterogeneity: $I^2 = 92\%$, $\tau^2 =$				l	1 1 1	I				
Residual heterogeneity: $I^2 = 9$	92%, p < 0	.01		0.0	05 0.1 0.5 1 2	10				







Meta-analysis for binary outcomes: Funnel plot

The possibility of publication bias can be assessed by means of hypothesis tests (metabias() function) and of funnel plot visual inspection.





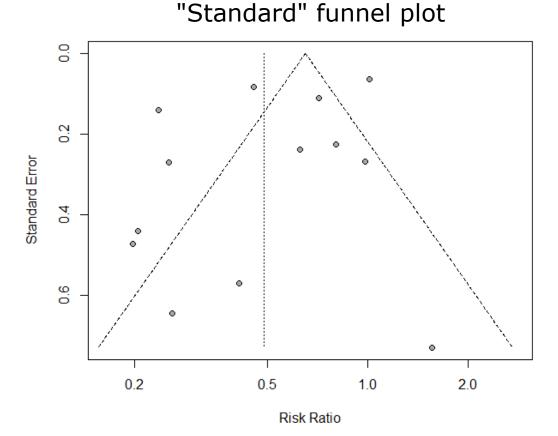




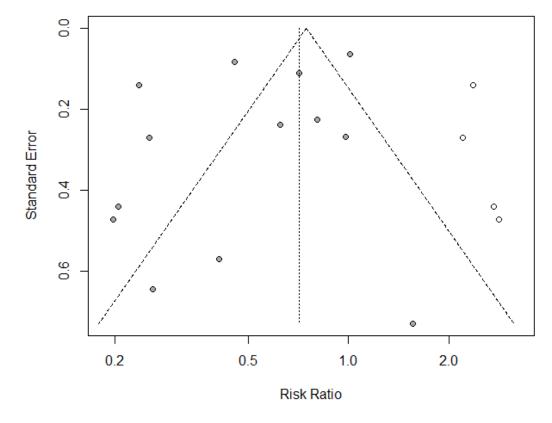
```
# To obtain a funnel plot
> funnel(ma04, common=FALSE)
# To obtain a funnel plot with trim-and-fill
> (ma04_tf <- trimfill(ma04, common=FALSE))</pre>
  funnel(ma04_tf)
Filled: Rosenthal et al 1961 2.2018 [1.2965; 3.7394]
                                                           6.3
Filled: Hart & Sutherland 1977 2.3620 [1.7901; 3.1166]
                                                           6.7
Filled: Ferguson & Simes 1949 2.7274 [1.1489; 6.4747]
                                                           5.4
Filled: Vandiviere et al 1973 2.8260 [1.1199; 7.1309]
                                                           5.3
Number of studies combined: k = 17 (with 4 added studies)
                        RR 95%-CI z p-value
Random effects model 0.7102 [0.4593; 1.0981] -1.54 0.1237
Prediction interval [0.1100; 4.5860]
Quantifying heterogeneity:
tau^2 = 0.7164; H = 4.05 [3.46; 4.75]; I<sup>2</sup> = 93.9% [91.6%; 95.6%]
Test of heterogeneity:
     Q d.f. p-value
 262.73 	 16 < 0.0001
```



Meta-analysis for binary outcomes: Funnel plot



Funnel plot with trim-and-fill









Meta-analysis for continuous outcomes: Meta-analysis

Again, we will use stroke dataset to perform meta-analysis to estimate the pooled mean difference

Given the raw data, meta-analysis of continuous outcomes can be performed using the metacont() function

Let's start to check its arguments!

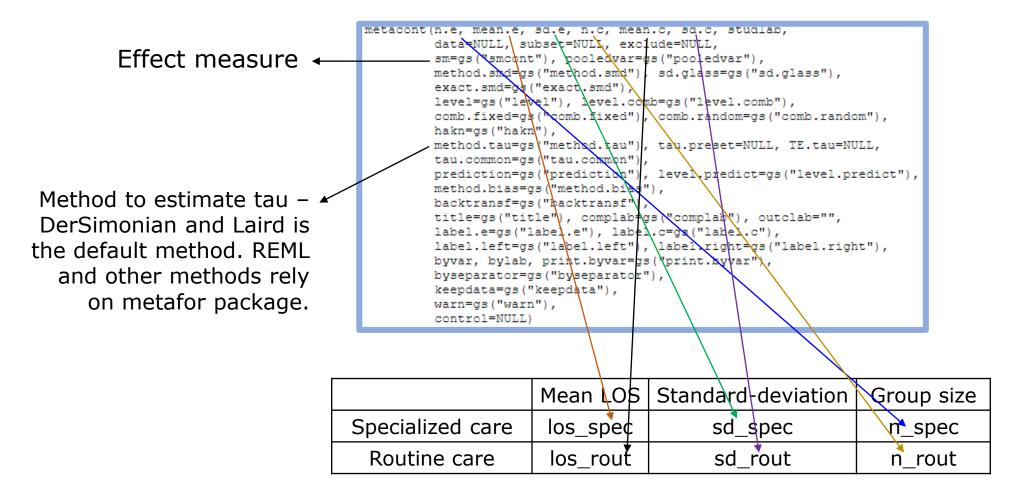
> ?metacont











In meta package, fixed effects and random effects meta-analysis results are simultaneously presented







Meta-analysis for continuous outcomes: Meta-analysis

```
# Meta-analysis for the mean difference with the amount of heterogeneity estimated by the restricted maximum-likelihood method. The prediction argument allows for obtention of estimates of a hypothetical future study.
```

```
> ma05 <- metacont(data=stroke, n.e=n_spec, mean.e=los_spec,
sd.e=sd_spec, n.c=n_rout, mean.c=los_rout, sd.c=sd_rout, sm="MD",
method.tau="REML", prediction=TRUE, studlab=paste(study,year))</pre>
```









```
> ma05
                              95%-CI %W(fixed) %W(random)
                          MD
Edinburgh 2005 -20.0000 [-32.4744; -7.5256] 1.4
Oprington-Mild 2007 -2.0000 [ -4.8271; 0.8271] 28.1
                                                              11.7
Oprington-Moderate 2007 -55.0000 [-62.7656; -47.2344] 3.7 11.4
Oprington-Severe 2007 -71.0000 \ [-95.0223; -46.9777] 0.4 9.6
Number of studies combined: k = 9
              95%-CI z p-value
Fixed effect model -3.4636 \ \lceil -4.9626; -1.9646 \rceil -4.53 < 0.0001
Random effects model -15.1060 [-32.6409; 2.4289] -1.69
                                                   0.0913
Prediction interval [-80.4949; 50.2829]
Quantifying heterogeneity:
tau^2 = 684.6462; H = 5.46 [4.54; 6.58]; I<sup>2</sup> = 96.7% [95.2%; 97.7%]
Test of heterogeneity:
     o d.f. p-value
238.92 8 < 0.0001
```

Specialized care associates with a trend towards decreased LOS (pooled meta-analytic estimate: -15.1 days; 95%CI=-32.6;2.4; p=0.091). However, severe heterogeneity was observed (p<0.001; I²=96.7%)







Meta-analysis for continuous outcomes: Forest plot

The function forest() creates a forest plot.

The "common=FALSE" argument is used to hide fixed effects model results. In this example, the predicted interval will not be presented ("prediction=FALSE"). The "xlim" argument sets the limits of the xx axis of the forest plot. The number of decimal places is set by the "digits" arguments.

> forest(ma05, common=FALSE, prediction=FALSE, xlim=c(-75,50), digits=1,
digits.sd=1)







	Ex	perime	ental		Co	ntrol				
Study	Total	Mean	SD	Total	Mean	SD	Mean Difference	MD	95%-CI	Weight
Edinburah 2005	155	EE O	47.0	156	75.0	640		20.0	[22 E: 7 E]	11 00/
Edinburgh 2005	155	55.0		156	75.0				[-32.5; -7.5]	11.0%
Oprington-Mild 2007	31	27.0	7.0	32	29.0	4.0		-2.0	[-4.8; 0.8]	11.7%
Oprington-Moderate 2007	75	64.0	17.0	71	119.0	29.0		-55.0	[-62.8; -47.2]	11.4%
Oprington-Severe 2007	18	66.0	20.0	18	137.0	48.0		-71.0	[-95.0; -47.0]	9.6%
Montreal-Home 2010	8	14.0	8.0	13	18.0	11.0	-	-4.0	[-12.2; 4.2]	11.4%
Montreal-Transfer 2010	57	19.0	7.0	52	18.0	4.0	+	1.0	[-1.1; 3.1]	11.7%
Newcastle 2011	34	52.0	45.0	33	41.0	34.0	+	11.0	[-8.1; 30.1]	10.3%
Umea 2014	110	21.0	16.0	183	31.0	27.0		-10.0	[-14.9; -5.1]	11.6%
Uppsala 2018	60	30.0	27.0	52	23.0	20.0	 	7.0	[-1.7; 15.7]	11.4%
Random effects model	548			610				-15.1	[-32.6; 2.4]	100.0%
Heterogeneity: $I^2 = 97\%$, τ^2 :	= 684.6	462, p <	< 0.01							
,						-1	00-80 -60 -40 -20 0 20 40			







Meta-analysis for the generic inverse variance method: stroke2 dataset

In meta package, metagen() function allows to perform metaanalysis based on the generic inverse variance method :

```
> ma06 <- metagen(data=stroke2, TE=mean_diff, seTE=sem,
sm="MD", prediction=TRUE, studlab=paste(study, year))</pre>
```







```
> ma06
                                95%-CI %W(fixed) %W(random)
Edinburgh 2005 -10.5000 [-14.4199; -6.5801]
                                           20.8
                                                     20.8
Newcastle 2011 -11.0000 [-15.5079; -6.4921] 15.7 15.7
Umea 2014 -9.6000 [-13.1279; -6.0721] 25.6 25.6
Uppsala 2018 -8.3000 [-14.1799; -2.4201] 9.2 9.2
oslo 2019 -11.4000 [-14.7319; -8.0681] 28.7
                                                     28.7
Number of studies combined: k = 5
                               95%-CI z p-value
Fixed effect model -10.4035 [-12.1891; -8.6179] -11.42 < 0.0001
Random effects model -10.4035 [-12.1891; -8.6179] -11.42 < 0.0001
Prediction interval [-13.3028; -7.5041]
Quantifying heterogeneity:
tau^2 = 0; H = 1.00 [1.00; 1.15]; I^2 = 0.0\% [0.0\%; 24.7\%]
Test of heterogeneity:
   Q d.f. p-value
1.10 4 0.8936
```









```
> forest(ma06, common=FALSE, prediction=TRUE,
xlim=c(-16,5), digits=1, digits.se = 1)
```

Study	TE se	TE	Mean D	Difference	• N	ΙD	95%-CI	Weight
Edinburgh 2005 Newcastle 2011 Umea 2014 Uppsala 2018	-11.0 -9.6	2.0 — 2.3 — 1.8 — 3.0 —			-1° -9	1.0 9.6	[-14.4; -6.6] [-15.5; -6.5] [-13.1; -6.1] [-14.2; -2.4]	20.8% 15.7% 25.6% 9.2%
Oslo 2019		1.7					[-14.7; -8.1]	28.7%
Fixed effect mode Prediction interval Heterogeneity: $I^2 = 0$	al		<u></u>	1			[-12.2; -8.6] [-13.3; -7.5]	100.0%
		-15	-10	-5 () 5			



