

# BS2280 - Econometrics 1

## Lecture 3 - Part 2: Bias, Efficiency and Interpreting Coefficients

by Dr Yichen Zhu

# Outline

- 1 Review: 6 OLS assumptions
- 2 Unbiasedness and Efficiency
- 3 Interpretation
- 4 Prediction (post-lecture video)
- 5 Goodness of Fit (post-lecture video, very important)

## Intended Learning Outcomes

- Interpreting coefficients of regressions
- Deriving and understanding  $R^2$

# Review: Simple Linear Regression Model - Population vs. Sample



Population	Sample
$Y_i = \beta_1 + \beta_2 X_i + u_i$	$Y_i = \hat{Y}_i + \hat{u}_i = \hat{\beta}_1 + \hat{\beta}_2 X_i + \hat{u}_i$
parameters $\beta_1$ and $\beta_2$	coefficients $\hat{\beta}_1$ and $\hat{\beta}_2$
$u_i$ disturbance term	$\hat{u}_i$ residual

OLS estimates  $\hat{\beta}_1$  and  $\hat{\beta}_2$  have certain desirable properties, but these properties rely on a set of assumptions we need to make!!!

# Review: OLS Assumption 1

**Model is linear in parameters and correctly specified**

Which one describes Assumption 1?

A)  $Y_i = \beta_1 + \beta_2 X_i + u_i$

B)  $Y_i = \beta_1 X_i^{\beta_2} + u_i$

## Review: OLS Assumption 2

**There is some variation in the  $X$  variable**

Which one describes Assumption 2?

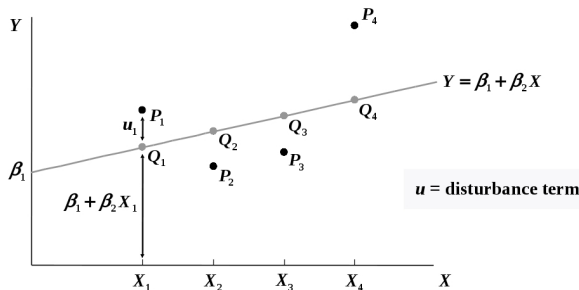
- A)  $X$  can be constant
- B)  $X$  cannot be constant

# Review: OLS Assumption 3

## Disturbance term has zero expectation

Which one describes Assumption 3?

- A)  $E(u_i) = 0$  for all  $i$  or  $E(u_i|X_1, X_2, \dots, X_n) = 0$  for all  $i$
- B)  $E(u_i) \neq 0$  for all  $i$  or  $E(u_i|X_1, X_2, \dots, X_n) \neq 0$  for all  $i$



## Review: OLS Assumption 4

### The disturbance term is homoscedastic

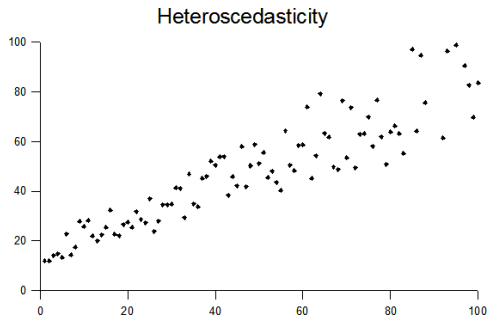
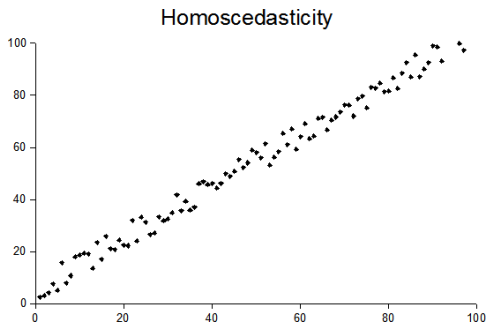
Which one describes Assumption 4?

- A) We assume that the error term has a constant variance,  
 $E(u_i^2) = \sigma_u^2$  for all  $i$
- B) We assume that the error term has a non-constant variance,  
 $E(u_i^2) \neq \sigma_u^2$  for all  $i$



# Review: OLS Assumption 4

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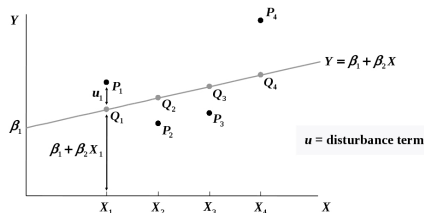


# Review: OLS Assumption 5

## Values of disturbance term have independent distributions

Which one describes Assumption 5?

- A) We assume that the error terms are absolutely **independent** of each other.  $u_i$  is independently distributed from  $u_j$  for all  $i \neq j$
- B) We assume that the error terms are absolutely **dependent** of each other.  $u_i$  is dependently distributed from  $u_j$  for all  $i \neq j$



## Review: OLS Assumption 6

**The disturbance term has a normal distribution**

Which one describes Assumption 6?

- A) We assume the error terms come from a **normal** distribution
- B) We assume the error terms come from a **T** distribution

## Check the 'quality' of OLS estimated $\hat{\beta}_1$ and $\hat{\beta}_2$

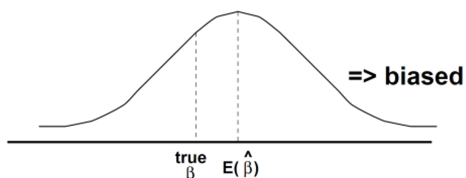
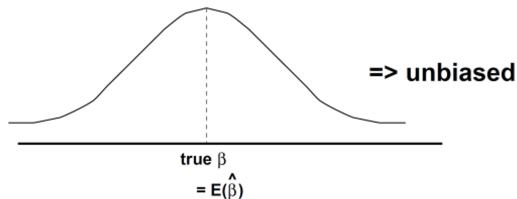
- Unbiasedness
- Consistency
- Efficiency / Precision
- If all our OLS assumptions hold, we get Best (most efficient) Linear Unbiased Estimators (BLUE)  $\hat{\beta}_1$  and  $\hat{\beta}_2$

# Unbiasedness

## Unbiasedness

An estimator is unbiased when expected value equals population value. E.g. for  $\hat{\beta}_2$ :  $E(\hat{\beta}_2) = \beta_2$ , otherwise, the estimator is said to be biased

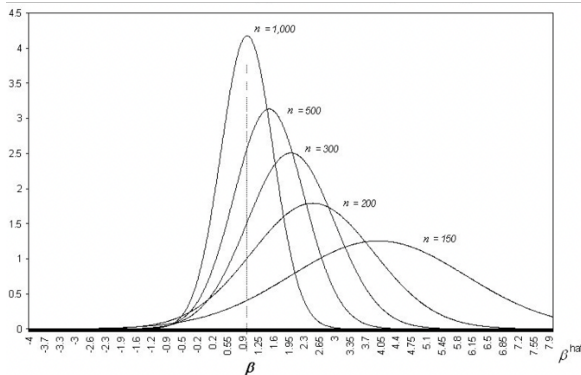
- This will be the case for OLS because we have Assumptions 1,2 and 3  
 $E(u_i) = 0$
- See Dougherty chapter 2 for mathematical prove



# Consistency

## Consistency

The larger the sample the closer our estimators ( $\hat{\beta}_1$  and  $\hat{\beta}_2$ ) should be to the population value ( $\beta_1$  and  $\beta_2$ )

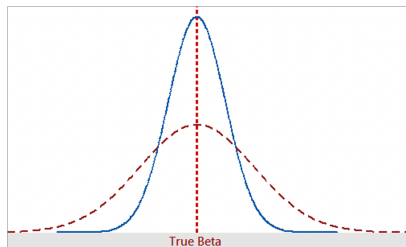


# Efficiency / Precision

## Efficiency / Precision

For the OLS estimator to be best (most efficient) it needs to have a lower variance than all the other estimators within the class.

- Recall that our interest lies in knowing if  $\hat{\beta}_1$  and  $\hat{\beta}_2$  are close estimates of true but unknown  $\beta_1$  and  $\beta_2$
- An estimator has higher precision if its variance is small compared to one with a larger variance



## Efficiency / Precision

- Variances of  $\hat{\beta}_1$  and  $\hat{\beta}_2$ , i.e.  $\sigma_{\hat{\beta}_1}^2$  and  $\sigma_{\hat{\beta}_2}^2$
- Lower variance  $\sigma_{\hat{\beta}_1}^2$ , our estimates are more efficient / precise
- What factors will affect the value of variance:  $\sigma_{\hat{\beta}_1}^2$  and  $\sigma_{\hat{\beta}_2}^2$  ?
- We will use  $\sigma_{\hat{\beta}_2}^2$  to explain, but same intuition works for  $\sigma_{\hat{\beta}_1}^2$



# Efficiency / Precision

- Variance of  $\beta_2$  is:

$$\sigma_{\hat{\beta}_2}^2 = \frac{\sigma_{u_i}^2}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sigma_{u_i}^2}{n \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2}$$

- we define  $MSD(X)$  as:

$$MSD(X) = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$$

- Therefore

$$\sigma_{\hat{\beta}_2}^2 = \frac{\sigma_{u_i}^2}{n MSD(X)}$$

# Efficiency / Precision

$$\sigma_{\hat{\beta}_2}^2 = \frac{\sigma_{u_i}^2}{nMSD(X)}$$

Efficiency will depend on three factors:

- $\sigma_{u_i}^2$ : Larger residuals' variance reduce the precision of estimates (see earlier example above)
- $n$ : Larger the number of observations ( $n$ ), the smaller  $\sigma_{\hat{\beta}_2}^2$
- $MSD(X)$ : Larger  $MSD(X)$  leads to smaller  $\sigma_{\hat{\beta}_2}^2$ 
  - Regression analysis - understand how variations in  $X_i$  explain variations in  $Y_i$ . Variations in  $u_i$  pick up the rest
  - Low variations in  $MSD(X)$  means low variations in  $X_i \rightarrow$  more variations from  $u_i \rightarrow$  higher variance (lower precision) in  $\hat{\beta}_2$

# Efficiency / Precision

```
> summary(earnfit)
Call:
lm(formula = EARNINGS ~ S, data = EAWE21.simple)

Residuals:
Min      1Q  Median      3Q      Max
-20.079  -6.726  -2.203   3.451  79.037

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)   0.7647     2.8038   0.273   0.785
S              1.2657     0.1855   6.824 2.58e-11 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 11.36 on 498 degrees of freedom
Multiple R-squared:  0.08551, Adjusted R-squared:  0.08368
F-statistic: 46.57 on 1 and 498 DF,  p-value: 2.579e-11
```

# BLUE!!!

- If all our 6 assumptions mentioned above hold, then our OLS estimator is BLUE
- Best Linear Unbiased Estimator
- The Best property is particular of interest: the Gauss-Markov Theorem states that provided the assumptions hold, OLS estimators will have the lowest variance amongst all possible estimators (proof of this in Dougherty textbook Appendix 2.1)
- An estimator with low variance means it comes from a sampling distribution where most of the values are concentrated around the true but unknown population value of the characteristic it is estimating

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# Interpretation of Coefficients I

- Last week we analysed the relationship between education and income.
- We used OLS regression to estimate  $\beta_1$  and  $\beta_2$

```
> earnfit <- lm(EARNINGS~S, data=EAW21.simple)
> earnfit
```

Call:

```
lm(formula = EARNINGS ~ S, data = EAW21.simple)
```

Coefficients:

```
(Intercept)          S
0.7647         1.2657
```

Regression equation:

$$\widehat{EARNINGS}_i = 0.765 + 1.266S_i$$



## Interpretation of Coefficients II: $\beta_2$

$$\widehat{EARNINGS}_i = 0.765 + 1.266S_i$$

- What do results actually mean?
- Interpretation depends on units of variables!
- To interpret  $\beta_2$  we think at the margin: A one unit change in X leads to a  $\beta_2$  unit change in Y.
- In our example:
  - Increasing education by 1 year will increase hourly earnings by \$1.27
  - Increasing education by 3 years will increase hourly earnings by  $3 \times \$1.27$

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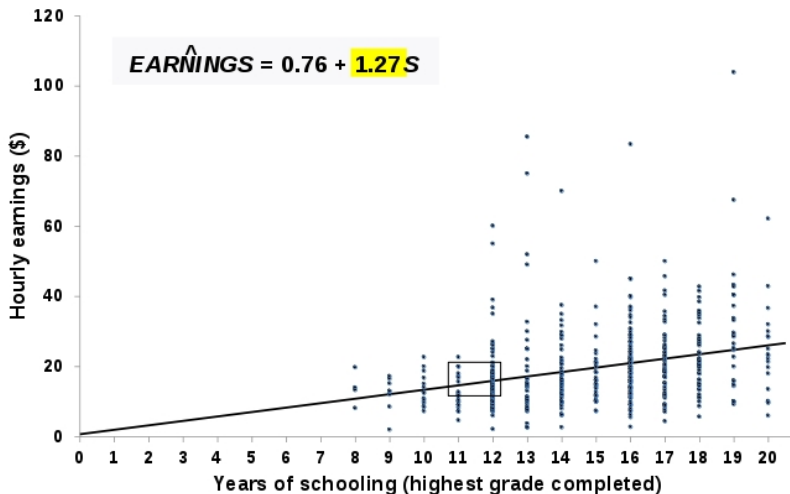
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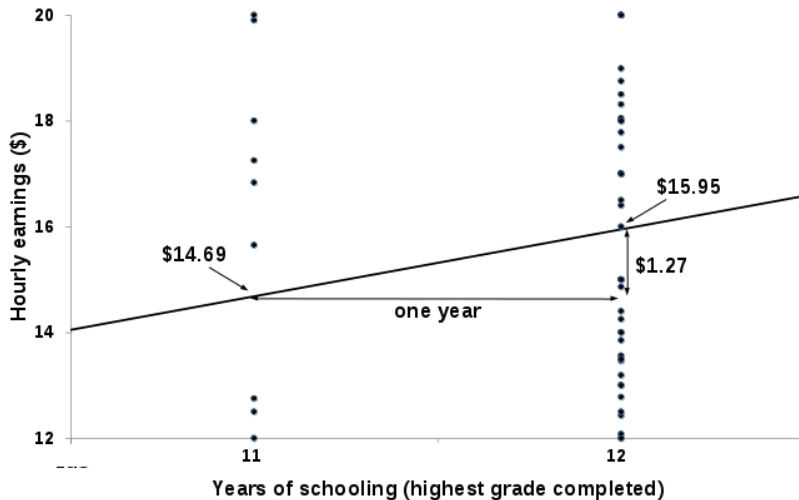
Hourly earnings (\$)

Years of schooling (highest grade completed)

# Interpretation of Coefficients II: $\beta_2$

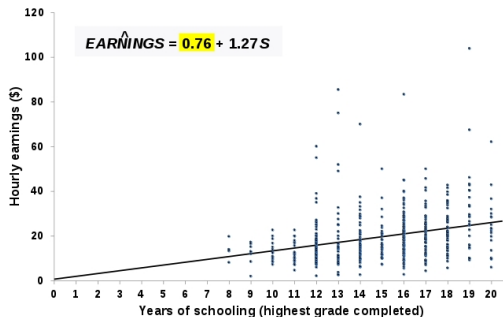


## Interpretation of Coefficients II: $\beta_2$



# Interpretation of Coefficients III: $\beta_1$

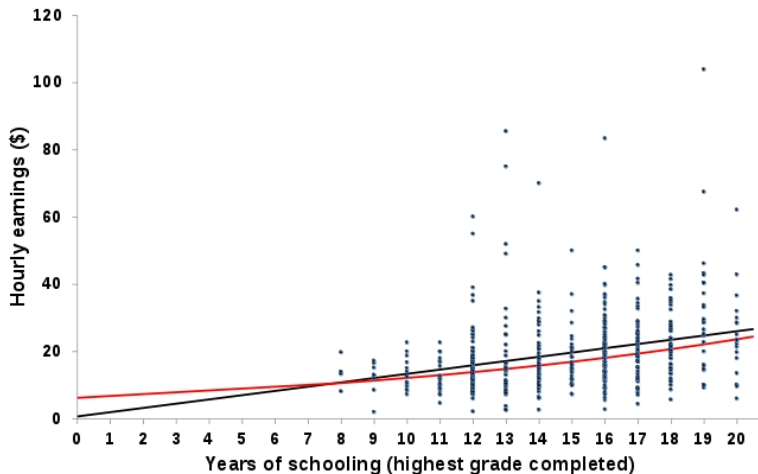
- $\beta_1$ : intercept, it shows value of Y when X = 0.
- A person with no education would have an hourly salary of \$0.76.
- Is such a low wage realistic? No!
- Statistically, it makes sense. But, economically, it does not make sense.
- Solution: Limit interpretation to a specific range





# Interpretation of Coefficients III: $\beta_1$

- Alternative solution: use non-linear specification. More on that later...



## Student Task:

- We estimate the impact of temperature on ice cream sales over a period of 21 days.
- Temperature is measured in degree Celsius and ice cream sales in thousand GBP
- The results are:  $\widehat{Sales}_t = -0.184 + 0.111 Temp_t$
- Interpret carefully the sign and the size of both the  $\beta_1$  and  $\beta_2$  coefficients.

# Prediction

- We can use our regression line for predictions!
- Our model predicts the average hourly wages for different levels of education.
- For example, what is the predict level of average wages?

Person	Education Years	Predicted Earnings
Person 1	$S_1 = 0$	$\widehat{EARNINGS}_1 = 0.76 + 1.27 \times 0 = \$0.76$
Person 3	$S_3 = 5$	$\widehat{EARNINGS}_3 = 0.76 + 1.27 \times 5 = \$7.11$
Person 7	$S_7 = 10$	$\widehat{EARNINGS}_7 = 0.76 + 1.27 \times 10 = \$13.46$

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- The results were:  $\widehat{Sales}_t = -0.184 + 0.111 Temp_t$
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# Goodness of Fit

- Regression analysis aims to understand whether observed variations in Y can be explained by observed variations in X.
- Variance of a variable is a measure of its dispersion, i.e. how far away the observations are from their mean value.
- Dispersion in the Y variable is given by the sum of squared deviation from the mean (Total Sum of Squares, TSS)

$$TSS = \sum (Y_i - \bar{Y})^2$$

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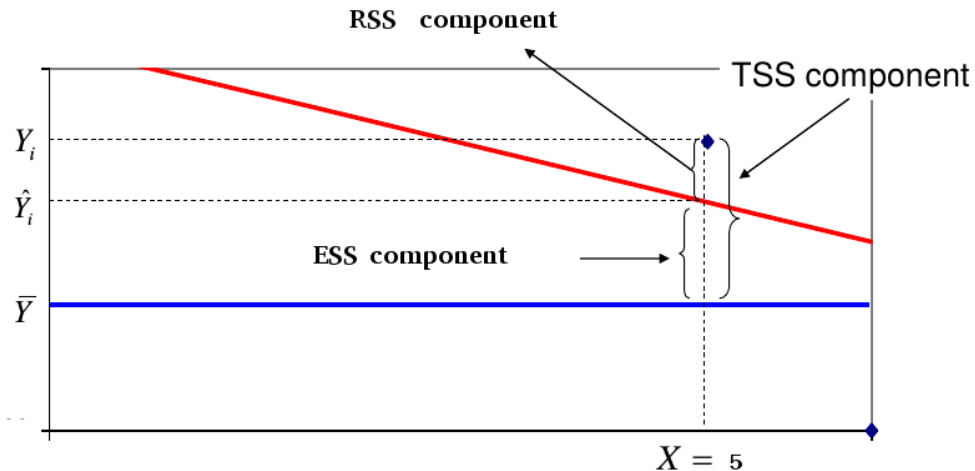
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$Y_i$	$= \beta_1 + \beta_2$	$X_i$	+	$u_i$
Total variations, variations in $Y$		Explained variations, variations in $X$		Residual variations, variations in $u_i$
Total Sum of Squares, TSS		Explained sum of squares, ESS		Residual sum of squares, RSS
$TSS = \sum (Y_i - \bar{Y})^2$		$ESS = \sum (\hat{Y}_i - \bar{Y})^2$		$RSS = \sum (Y_i - \hat{Y}_i)^2 = \sum u_i^2$ (1)

- Total variation in  $Y$  can be decomposed into variations arising from the  $X$  variable and variations arising from the residuals  $u_i$

$$\sum_{i=1}^n (Y_i - \bar{Y})^2 = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 + \sum_{i=1}^n (Y_i - \hat{Y}_i)^2$$

$$TSS = ESS + RSS$$



# Calculating $R^2$

- Create a measure of the proportion of ESS in TSS: the  $R^2$  - the coefficient of determination

$$R^2 = \frac{\text{Explained Variations in } X}{\text{Total Variations in } Y} = \frac{ESS}{TSS} = \frac{\sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2}{\sum_{i=1}^n (Y_i - \bar{Y})^2}$$

- $0 \leq R^2 \leq 1$ : The  $R^2$  is bound between 0 and 1
- $R^2 = 0$ : variations in X cannot explain any variations in Y
- $R^2 = 1$ : variations in X can fully explain variations in Y
- A high  $R^2$  means that the independent variable X is good at predicting Y
- **Example:**  $R^2 = 0.75$ , it means that 75% of the variations in Y can be explained by variations in the X variable.
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- $R^2 = 0$ : variations in X cannot explain any variations in Y
- $R^2 = 1$ : variations in X can fully explain variations in Y
- A high  $R^2$  means that the independent variable X is good at predicting Y
- **Example:**  $R^2 = 0.75$ , it means that 75% of the variations in Y can be explained by variations in the X variable.
- Remaining 25% comes from unobservables through the residual term

## Calculating $R^2$

- Create a measure of the proportion of ESS in TSS: the  $R^2$  - the coefficient of determination

$$R^2 = \frac{\text{Explained Variations in } X}{\text{Total Variations in } Y} = \frac{ESS}{TSS} = \frac{\sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2}{\sum_{i=1}^n (Y_i - \bar{Y})^2}$$

- $0 \leq R^2 \leq 1$ : The  $R^2$  is bound between 0 and 1
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- A high  $R^2$  means that the independent variable X is good at predicting Y
- **Example:**  $R^2 = 0.75$ , it means that 75% of the variations in Y can be explained by variations in the X variable.
- Remaining 25% comes from unobservables through the residual term

```
> summary(earnfit)
Call:
lm(formula = EARNINGS ~ S, data = EAWE21.simple)

Residuals:
    Min       1Q   Median       3Q      Max
-20.079  -6.726  -2.203   3.451  79.037

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)   0.7647     2.8038   0.273   0.785
S             1.2657     0.1855   6.824 2.58e-11 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 11.36 on 498 degrees of freedom
Multiple R-squared:  0.08551, Adjusted R-squared:  0.08368
F-statistic: 46.57 on 1 and 498 DF,  p-value: 2.579e-11
```

$R^2 = 0.0855$ , only 8.55% of the **variations in earnings** can be explained by **variations in schooling**. 91.45% of the variations in earnings is left unexplained.

# Calculating $R^2$

- The `anova` command allows us to identify **ESS** and **RSS**:

```
> anova(earnfit) # use regression output as argument
```

Analysis of Variance Table

Response: EARNINGS

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
S	1	6014	6014.0	46.568	2.579e-11 ***
Residuals	498	64315	129.1		

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

- $ESS = 6014$
- $RSS = 64315$
- so  $TSS = ESS + RSS = 70329$

- $$R^2 = \frac{ESS}{TSS} = \frac{6014}{70329} = 0.0855$$



## Student Task

- Using the simplified R output below, interpret the coefficients and comment on the overall fit of the model.

Call:

```
lm(formula = lexp ~ gnppc, data = lifeexp)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
(Intercept)	69.44836458	0.54798213	126.735	< 2e-16	***
gnppc	0.00032342	0.00004012	8.061	3.45e-11	***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.36 on 61 degrees of freedom

(5 observations deleted due to missingness)

Multiple R-squared: 0.5158, Adjusted R-squared: 0.5079

F-statistic: 64.98 on 1 and 61 DF, p-value: 3.452e-11

- where *lexp* is the life expectancy at birth in years and *gnppc* the gross national product per capita measured in USD.

# Student Task

- Using the Anova output below, calculate  $R^2$  and explain how well the weight (wt, Weight in 1000lbs) of a car can explain the fuel consumption (mpg) of a car.

```
Response: mpg
Df Sum Sq Mean Sq F value    Pr(>F)
wt          1  847.73    847.73  91.375 1.294e-10 ***
Residuals  30  278.32     9.28
---
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

## What to do next:

- Attempt homework 3
- Revise basic R commands from R Workshop 1
- Read chapter 2.1 - 2.7 of Dougherty