

# Catch-Up Growth and Inter-Industry Productivity Spillovers in Open Economies

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## Abstract

Countries tend to export more skill-intensive products as they become more productive. This paper proposes a tractable quantitative framework to examine the role of inter-industry productivity spillovers in this development process. I document that a country's comparative advantage tends to increase in industries that employ occupations that are used most intensively in current exports. The model rationalizes these findings by incorporating occupation-specific dynamic scale economies into a multi-sector gravity framework. I estimate the model using cross-sector heterogeneity in foreign demand shocks and find that scale economies are relatively large in high-skilled production. As a result, productivity spillovers tend to be larger in richer countries, and access to foreign markets allows developing countries to shift labor into sectors that contribute more to aggregate productivity growth. Counterfactual exercises suggest that spillovers play a quantitatively substantial role in accounting for slow cross-country convergence and increase the gains from trade, especially in economies with a comparative advantage in manufacturing.

*JEL Classification:* F1; F4; F6; O1; O3; O4;

*Keywords:* Productivity; Convergence; Spillovers; Dynamic scale economies; Comparative advantage; Exports

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\*First version: December 30, 2018. Department of Economics, University of Toronto, 150 St. George Street, Toronto, Ontario, Canada M5S 3G7. Email: [marijn.bolhuis@mail.utoronto.ca](mailto:marijn.bolhuis@mail.utoronto.ca) I thank Xiaodong Zhu, Kevin Lim and Diego Restuccia for their advice and encouragement and for extensive discussions. I am also indebted to Douglas Gollin for his support, as well as the Department of Economics and the Centre for the Study of African Economies at the University of Oxford for their hospitality while writing this paper. This paper has benefited from discussions with Stephen Ayerst, Loren Brandt, Murat Celik, Dylan Gowans, Ian Herzog, Allan Hsiao, Torsten Jaccard, Jonathan Lehne, Andrei Levchenko, Yuhei Miyauchi, Jordi Mondria, Peter Morrow, Andreas Moxnes, Ezra Oberfield, Serdar Ozkan, Caroline Pitchik, Simon Quinn, Baxter Halm Robinson, Todd Schoelman, Lidia Smitkova, Tomasz Swiecki, Dan Trefler, along with participants at the Warwick Economics PhD Conference 2019, the North East Universities Development Consortium Conference 2019 at Northwestern University, and several seminars at the University of Toronto. Financial support from the Ontario Trillium Foundation is gratefully acknowledged. All remaining errors are mine.

# 1. Introduction

Developing countries that catch up successfully to the global economic frontier tend to experience rapid growth in labor productivity. What is the role of a country's production structure as a driver of this productivity growth? The notion that what a country produces matters for aggregate productivity has a long history in macroeconomics and international trade, dating back to Marshall's concept of external economies of scale (Marshall, 1890).<sup>1</sup> In a dynamic setting, the sectoral composition of production matters for growth in the presence of dynamic scale economies that are heterogeneous across sectors (Krugman, 1987). A well-established theoretical literature has elaborated several mechanisms through which these scale economies can manifest themselves, such as learning-by-doing (e.g. Young, 1991) and human capital spillovers (Lucas Jr, 1988; Stokey, 1991). More recently, empirical work on this question has focused on how external demand conditions and a country's structure of comparative advantage interact to affect income and sector-level productivity growth. A common finding in this literature is that 'what you export matters' (Hausmann et al., 2007): on average, countries that export in technologically more advanced sectors tend to experience faster income growth (e.g. Bartelme et al., 2019b).

The empirical literature on dynamic scale economies has focused on whether countries' production structures affect growth heterogeneously across sectors, without providing a structural framework to interpret these findings. The theoretical literature provides explanations for these patterns but does not assess their quantitative significance. The goal of this paper is to narrow the gap between theory and data. I develop a quantitative general equilibrium, multi-sector trade model that is tractable enough to estimate dynamic scale economies and quantify their welfare implications. My main empirical finding is that dynamic scale economies are largest in skill-intensive sectors. As a result, the dynamic gains from international trade vary across countries, and reflect countries' initial patterns external demand and comparative advantage. This heterogeneity has implications for cross-country convergence in productivity and the gains from trade,

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<sup>1</sup>In recent years, economists have revived the literature on Marshallian externalities by incorporating external economies of scale into otherwise standard quantitative trade models (Kucheryavyi et al., 2020; Bartelme et al., 2019a). In contrast to this literature, this paper examines the dynamic productivity effects of countries' production structures.

which I explore through the lens of the model.

I start out by documenting two empirical facts that motivate my structural framework. First, I show that as developing economies become more productive, they tend to experience a shift in comparative advantage from low- to high-skill intensive sectors, which suggests that they become relatively more productive in high-skill intensive production. Second, I document that a country's comparative advantage tends to increase in industries that employ occupations that are used most intensively in current exports. This evidence of spillovers in productivity is mainly present in high-skilled manufacturing industries.<sup>2</sup>

Motivated by these empirical patterns, I develop a tractable quantitative model that generates inter-industry spillovers through dynamic scale economies at the level of occupational groups ('tasks'). I model dynamic scale economies as a process involving the combination of new ideas with existing insights (Buera and Oberfield, 2020). Workers doing the same task accumulate knowledge through learning-by-doing and by adopting new ideas from others. As all sectors in the economy combine different combinations of tasks in production, these interactions give rise to varying degrees of connectedness between sectors. The resulting law of motion of task productivity has three distinctive features. First, conditional on the mass of workers in a task, productivity growth can differ across tasks due to different degrees of increasing returns to scale in idea creation. Second, conditional on the pace of idea creation, productivity growth is faster in tasks that employ more workers, as spillovers are more likely. Finally, task productivity growth is always subject to convergence (or 'fishing-out'), as more productive workers are less likely to find new, productivity-improving ideas.

My framework remains tractable enough to estimate model parameters and perform counterfactuals using closed form solutions. I combine the dynamic part of my model with a standard multi-sector Ricardian economy. The model equilibrium can be summarized as a series of static equilibria connected by a law of motion of task produc-

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<sup>2</sup>These empirical results are in line with recent empirical work: while countries' productivity (or comparative advantage) in individual sectors tends to exhibit strong convergence (e.g. Levchenko and Zhang, 2016) relative to other sectors, countries tend to experience relatively faster productivity growth in sectors that are closely related to those in which they are initially most competitive (Bahar et al., 2019). They are also corroborated by recent work of Helm (2019) who finds that employment spillovers from other tradable industries' trade shocks are larger for industries employing similar workers and are generated mainly by shocks to technologically sophisticated industries.

tivity growth that only depends on previous sectoral employment shares and levels of task productivity. Conditional on estimates of supply and demand side parameters, the model can be solved in counterfactual changes using exact hat algebra ([Dekle et al., 2007](#)) without relying on estimates of initial productivity levels and trade costs.

I estimate the elasticities governing dynamic scale economies in two steps. First, I use the model's implied gravity equation to estimate sector-specific unit cost levels across countries. Richer countries tend to be more competitive in advanced sectors that use high-skilled tasks intensively. Second, I estimate the task-specific spillover parameters by relating changes in sector-specific unit costs to countries' export structures. To address endogeneity issues arising from supply side factors, I consider only variation in export structure induced by foreign demand shocks ([Bartelme et al., 2019b](#)). I implement my approach on UN COMTRADE data from 1962 to 2000, and use detailed occupation-specific data from O\*NET to assign occupations to groups based on their task content.

As the main empirical result, I find substantial heterogeneity in the extent of dynamic scale economies of different tasks. Spillovers are generally increasing in the skill level of the occupational groups. As a consequence, allocating labor to sectors that use high-skilled labor more intensively has a greater effect on aggregate productivity growth. Across clusters of sectors, spillovers are lowest in agriculture and highest in advanced manufacturing.<sup>3</sup> When I use the estimated parameters to assess model fit, the framework performs well at predicting cross-country differences in aggregate and sector-level labor productivity growth for the period 1970-2000. The model-implied sector-level accumulated spillovers explain more than 20 percent of the variation in long run changes in effective unit costs among a sample of 60 tradable sectors. In terms of aggregate labor productivity growth over the same period, the model explains 20 percent of the variation in changes in real GDP growth.

I then explore the consequences of this heterogeneity through the lens of my model. Combined with the estimated scale economies, the model implies that spillovers are larger in advanced economies as these tend to export and consume relatively more

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<sup>3</sup>These sectoral differences are in line with a recent paper of [Brummitt et al. \(2020\)](#), who use machine learning techniques to identify what distinguishes export baskets of fast-growing countries. One of their main findings is that countries with a larger share of exports in machinery relative to agriculture tend to grow faster.

high-skill intensive goods (Caron et al., 2014). As a result, inter-industry spillovers could potentially account for the lack of catch up in levels of aggregate labor productivity between developing and advanced economies during the 1970 to 2000 period (Johnson and Papageorgiou, 2019). I assess to what extent spillovers can account for slow unconditional (beta) convergence by exploring a counterfactual in which I set any dynamic scale economies to zero. Indeed, without spillovers a typical country at one tenth of the frontier in 1970 experiences 0.39 to 1.23 percentage points per year faster catch up to the frontier.

As a second exercise, I assess how inter-industry spillovers affect the gains from trade and to what extent these gains depends on a country's initial patterns of comparative advantage. Given that domestic demand in poorer countries tends to be concentrated in the technologically least advanced sectors, the model implies that the availability of foreign demand for goods from sectors with high spillovers is crucial for achieving catch-up to the frontier. In particular, trade integration leads to both static and dynamic gains if it shifts countries' exports towards high spillover sectors while integrating with trade partners that can provide its preferred imports at lower cost.<sup>4</sup> I explore a series of counterfactuals in which I keep a country's implied trade costs at their 1970 level, and construct the ensuing counterfactual path of productivity levels. In most countries, dynamic gains of trade are substantial and equal roughly one third of the average static gains. I find considerable heterogeneity across countries in terms of the dynamic gains from trade. Countries with a comparative advantage in agriculture tend to have lower dynamic gains, which is not surprising given that estimated spillovers are low in this sector. At the same time, estimated gains are generally higher in countries with an initial comparative advantage in low-skilled manufacturing. These results suggest that labor abundant countries gain more from trade integration than commodity exporters, as low-skilled manufacturing serves as a stepping stone towards the production of more technologically advanced goods.<sup>5</sup>

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<sup>4</sup>This intuition is borne out by empirical evidence on the exports of 20th century East Asian growth miracles that have been relatively technologically advanced. In the case of Korea, for example, the United States and Japan provided large foreign demand for steel and ships in the 1980s, and cars and electronics in the 1990s. In the case of China, the U.S. and Europe have formed the largest foreign markets in the form of toys and simple electronics in the 1990s and 2000s, and machinery, TVs and personal computers in the 2010s. Indeed, Rodrik (2006) and Schott (2008) argue that Chinese exports have been considerably more technologically sophisticated than exports of developing economies with similar income levels.

<sup>5</sup>This is in line with the recent empirical work of Hanson (2017), who documents that labor-abundant

**Related Literature and Contributions.** This paper adds to several strands of the trade and growth literature. First, it is particularly related to recent work on the role of static external economies of scale in open economies (Bartelme et al., 2019a; Kucheryavyy et al., 2020). These papers incorporate industry-level returns to scale in otherwise standard quantitative trade models, relating sector- and country-specific differences in productivity *levels* to factor allocations. This study complements this literature by developing a new framework in which industries are subject to dynamic scale economies, relating differences in productivity *growth* to factor allocations. In fact, in supplementary material I show that the dynamic scale economies in this paper are isomorphic to static external economies of scale when comparing productivity levels across countries. More generally, this paper contributes to the literature dynamic scale economies, which argues that trade may increase or decrease disparities between countries due to the existence of dynamic scale economies that differ by sector. Potential reasons for such divergence are sector-specific learning-by-doing (e.g. Young, 1991), human capital externalities (e.g. Lucas Jr, 1988), as well as trade-induced differences in incentives to accumulate physical capital (e.g. Krugman, 1981) and technology (e.g. Matsuyama, 2019). This paper contributes to this literature by quantifying the importance of dynamic scale economies for cross-country convergence.

Second, this study responds to the broad literature on the importance of trade and idea flows for endogenous growth (e.g. Sampson, 2016). In recent years, this literature has built on the seminal work of Eaton and Kortum (2001) and Eaton and Kortum (2002) by introducing innovation, imitation, and idea diffusion into standard multi-country general equilibrium models.<sup>6</sup> Theoretically, this paper is most closely related to Buera and Oberfield (2020), who develop a tractable model of idea diffusion through international trade, which allows them to quantify the contribution of trade barriers to TFP differences across countries and over time. I contribute to this literature by developing a new multi-sector framework with inter-industry spillovers that remains tractable

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East Asian countries tend to cycle through ever more skill- and capital-intensive offshoring industries, while these patterns are not present for commodity exporters.

<sup>6</sup>On the theoretical side, Alvarez et al. (2013), Sampson (2016) and Perla et al. (2015) build on this work by examining how free trade not only encourages the selection of efficient producers but also facilitates the diffusion of ideas between the most efficient exporters. Somale (2017), Cai and Li (2012), Cai and Li (2015) and Santacreu (2015) study how R&D is shaped by international trade.

enough to identify spillovers and to perform counterfactuals under only limited assumptions.

Third, this paper is nested in the growth literature that documents factor efficiency differences between countries.<sup>7</sup> A robust conclusion in this literature is that skilled labor tends to be relatively more productive in richer countries. Through the lens of this paper's framework, these factor efficiency differences are the result of dynamic scale economies that are stronger for high-skilled production. As such, this paper connects the empirical literature on factor efficiency differences with the theoretical literature on multi-factor dynamic general equilibrium models.<sup>8</sup>

Finally, this paper contributes to the literature on dynamic comparative advantage that examines how and why countries' patterns of comparative advantage evolve (e.g. [Hanson et al., 2018](#)). This paper contributes to this literature by offering a new endogenous growth theory of dynamic comparative advantage based on changes in occupational-specific productivity levels that are driven by dynamic scale economies.

## 2. Motivating Facts

I first highlight two key facts about structural transformation of skill-intensive production and comparative advantage. As a first fact, I document that countries' tradable production becomes more skill-intensive as aggregate productivity increases. I use newly classified cross-country data on sector-specific employment to document skill-biased structural change ([Buera et al., 2015](#)) in a broad panel of mainly developing economies. First, labor moves out of agriculture into mining and low-skilled manufacturing sectors. As countries develop further, these sectors contract too as labor further shifts into sectors intensive in the use of high-skilled labor. This pattern holds in a new cross-country employment panel from IPUMS International (1960-2011) and within the United States for the period 1850-2010. Second, I also document a strong positive association between countries' GDP per capita and their revealed comparative advantage in skill-

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<sup>7</sup>[Caselli \(2016\)](#) provides a good survey of this literature.

<sup>8</sup>There exists a literature that uses multi-country quantitative general equilibrium models to study how trade and structural change affect skill premia (e.g. [Parro, 2013](#); [Burstein et al., 2013](#); [Burstein and Vogel, 2017](#); [Cravino and Sotelo, 2019](#)). While the static component of this paper's quantitative framework is similar, the focus of this paper is on how trade and structural change affect income differences *across*, rather than *within*, countries.



intensive sectors. While skill-biased structural change in value added and employment could be driven entirely by non-homothetic preferences (e.g. [Kongsamut et al., 2001](#)) or an increase in the relative price of skill-intensive goods ([Ngai and Pissarides, 2007](#)),<sup>9</sup> this second pattern suggests that countries' relative price of skill-intensive tradable goods decreases as they become richer.<sup>10</sup> I confirm this cross-country pattern by observing the same shift in revealed comparative advantage within a sample of fast-growing East Asian countries.

As a second fact, I document that countries' revealed comparative advantage (RCA) -a measure of sector-specific inverse unit costs- tends to shift towards occupationally similar sectors, while also exhibiting convergence over time. This first pattern suggests that current production (in sectors with a high RCA) tends to foster above average productivity growth in sectors with a similar production structure.<sup>11</sup> At the same time, the second pattern confirms fast 'churning' of comparative advantage ([Hanson et al., 2015](#); [Darwich et al., 2019](#)) such that sector-specific unit costs tend to exhibit mean reversion.

## 2.1 Fact 1

**Fact 1:** *As countries become more productive, employment and comparative advantage shift from low- to high-skill intensive production.*

### Employment

To facilitate exposition, I aggregate tradable sectors into three clusters: (i) Agriculture and Food (agriculture, forestry, fisheries, food, beverages, and tobacco) (ii) Low-Skilled Manufacturing and Mining (textiles, clothing, leather, footwear, wood products, furniture, recycling, and mining), and (iii) High-Skilled Manufacturing (minerals, fuels, metals, rubbers, plastics, paper, printing, chemicals, machinery, transport and electronic equipment).<sup>12</sup> I use new internationally comparable census data from IPUMS Interna-

<sup>9</sup>For the latter, it is also necessary that goods from different sectors are complements for consumers.

<sup>10</sup>See also [Malmberg \(2017\)](#). Note that this result does not necessarily conflict with [Buera et al. \(2015\)](#), as I only consider tradable sectors (agriculture, mining and manufacturing). [Buera et al. \(2015\)](#) focus on services.

<sup>11</sup>See also [Bahar et al. \(2019\)](#), who explore different empirical channels through which countries diversify their exports over time.

<sup>12</sup>The pattern of skill-biased structural change in tradable sectors described here also holds at a more granular level of sector classification. See Appendix A for plots of sector-level tradable employment against log GDP p.w., where sectors are classified at the ISIC 3.0 2 digit level. The elasticity of the employ-



tional to document structural transformation in *employment* from low- to high-skilled sectors.<sup>13</sup> These data cover a wide range of countries and time periods, and -in contrast to the usual evidence- advanced economies are underrepresented.<sup>14</sup> Figure 1 plots employment shares in the three tradable clusters against log real GDP per worker. As poorer countries develop, labor shifts monotonically out of agriculture and food into industrial sectors. At first, low-skilled manufacturing and mining experience an increase in employment, but these sectors contract as a country's income increases further. Employment in more skill-intensive manufacturing expands monotonically.

A potential concern about these new employment data is that they only document a cross-sectional relationship between employment shares and GDP per capita. To address this, I examine structural transformation in employment *within* the United States from 1850 to 2010 using census data. Figure 1 plots the corresponding series. As in the cross-sectional IPUMS International sample, structural transformation of employment within the U.S. involves shifting labor from agriculture low-skilled industrial sectors and subsequently into more knowledge-intensive manufacturing.

### Comparative Advantage

Employment (nor value added) data do not provide a clear insight into sector-specific productivity trends. I use international trade data to shed light on the latter. An advantage of using trade data in this setting is that it is recorded at a detailed industry level across a wide range of countries and historical time periods. At the same time, measures of comparative advantage reflect differences in inverse unit costs when interpreted through the lens of any gravity model.<sup>15</sup>

First, I define the Revealed Comparative Advantage (RCA) (Balassa, 1965) of country

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ment share w.r.t. to GDP p.w. is lowest in agriculture (-.44) and broadly increases with skill intensity in the following order: textiles/clothing (.42), furniture/recycling/n.e.c (.50), leather/footwear (.52), wood products (.56), food/beverages (.65), non-metallic minerals (.69), mining (.73), fuels (.99), metals (1.1), rubbers/plastics (1.2), paper/printing (1.2), chemicals (1.3), machinery (1.4), transport equipment (1.4) and electronic equipment (1.4).

<sup>13</sup>A recent paper by Duernacker and Herrendorf (2016) uses the same data to examine structural change in 'service' and 'goods' occupations.

<sup>14</sup>In total, IPUMS International covers 94 countries, 365 censuses, and over 1 billion person records, from 1960 to 2013. Several large advanced economies have no or only limited public census records, such as Japan, Germany, United Kingdom, Italy, Korea, Russia, and Australia.

<sup>15</sup>For other papers that use trade data to infer productivity differences across countries and over time, see Levchenko and Zhang (2016), and Malmberg (2017).

$n$  in sector  $k$  at time  $t$  as:

$$RCA_{n,t}^k = \frac{X_{n,t}^k / X_{n,t}}{X_t^k / X_t} \quad (1)$$

where  $X_{n,t}^k$ ,  $X_{n,t}$ ,  $X_t^k$ , and  $X_t$  denote a country's exports in sector  $k$ , its total exports, global exports in sector  $k$ , and global total exports. RCA thus measures how specialized a country is in a given sector relative to the global mean. I construct the RCA of the three clusters for a wide range of countries using trade data from World Trade Flows (Feenstra et al., 2005) for the period 1970-2000.

In a world without trade costs, constant returns to scale, and homogeneous sector-specific Cobb-Douglas preferences, RCA reflects relative sector-specific unit costs.<sup>16</sup> Figure 2 plots (log) RCA for the three clusters against GDP per worker. While poorer countries tend to be relatively more productive in agriculture and food, their RCA shifts towards low-skilled manufacturing and mining as they increase their aggregate productivity. In turn, their RCA in this cluster tends to peak and decline at the expense of higher RCA in high-skilled manufacturing. Through the lens of a gravity model, these patterns suggest that, as countries become richer, they tend to experience below average productivity growth in agriculture and food. At the same time, their productivity growth in high-skilled manufacturing tends to accelerate, while relative productivity growth in low-skilled manufacturing is hump-shaped.

In order to go beyond this cross-sectional pattern, I plot the evolution of RCA against GDP per capita for *within* fast-growing East Asian countries (Korea, China, Japan, Thailand, Malaysia, Indonesia, Singapore, and Vietnam). These plots are presented in Figure 2. Reassuringly, these patterns are very similar to the ones previously documented. Again, RCA is monotonically decreasing in income per worker for agriculture and food, hump-shaped for low-skilled manufacturing and mining, and monotonically increasing for high-skilled manufacturing.

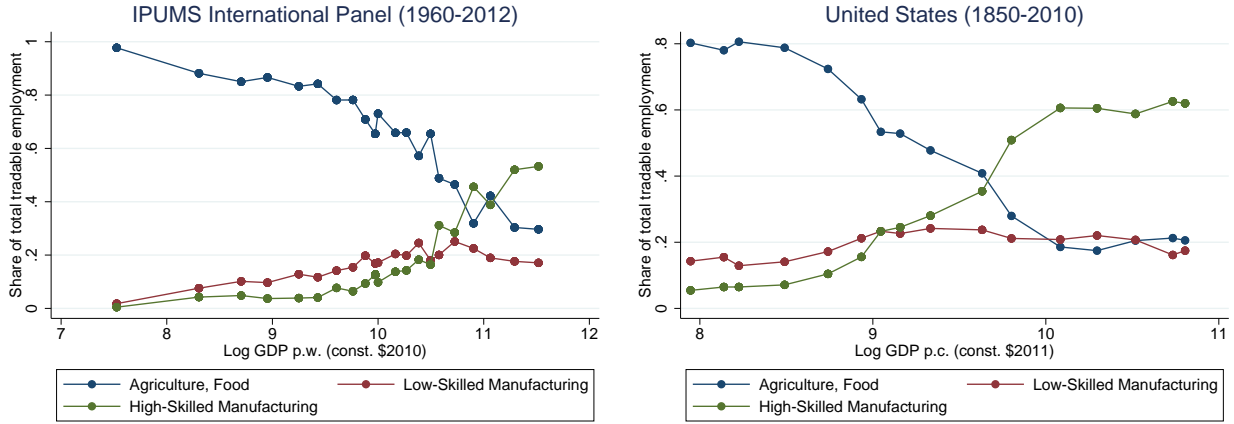
## 2.2 Fact 2

**Fact 2:** *Comparative advantage exhibits convergence and tends to shift into occupationally similar sectors.*

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<sup>16</sup>See the definition of static equilibrium in section 3.1.4 and exporter fixed effects in gravity equation estimates (section 4.1.1). Appendix B contains details.

Figure 1: Structural Transformation of Employment in Tradable Clusters



**Notes:** The left figure presents a binned scatterplot of countries' share of tradable employment for the three tradable clusters against log GDP p.w. for the IPUMS International Panel from 1970 to 2012. The right figures presents a similar binned scatterplot for the United States for each decade from 1850 to 2010. Employment data for the U.S. are based on micro-data from IPUMS USA censuses. Real GDP per worker (constant \$2010) data for the cross-country sample are from Penn World Tables 9.0. Estimates of real GDP per capita (constant \$2011) data for the U.S. are from the Maddison Project Database.

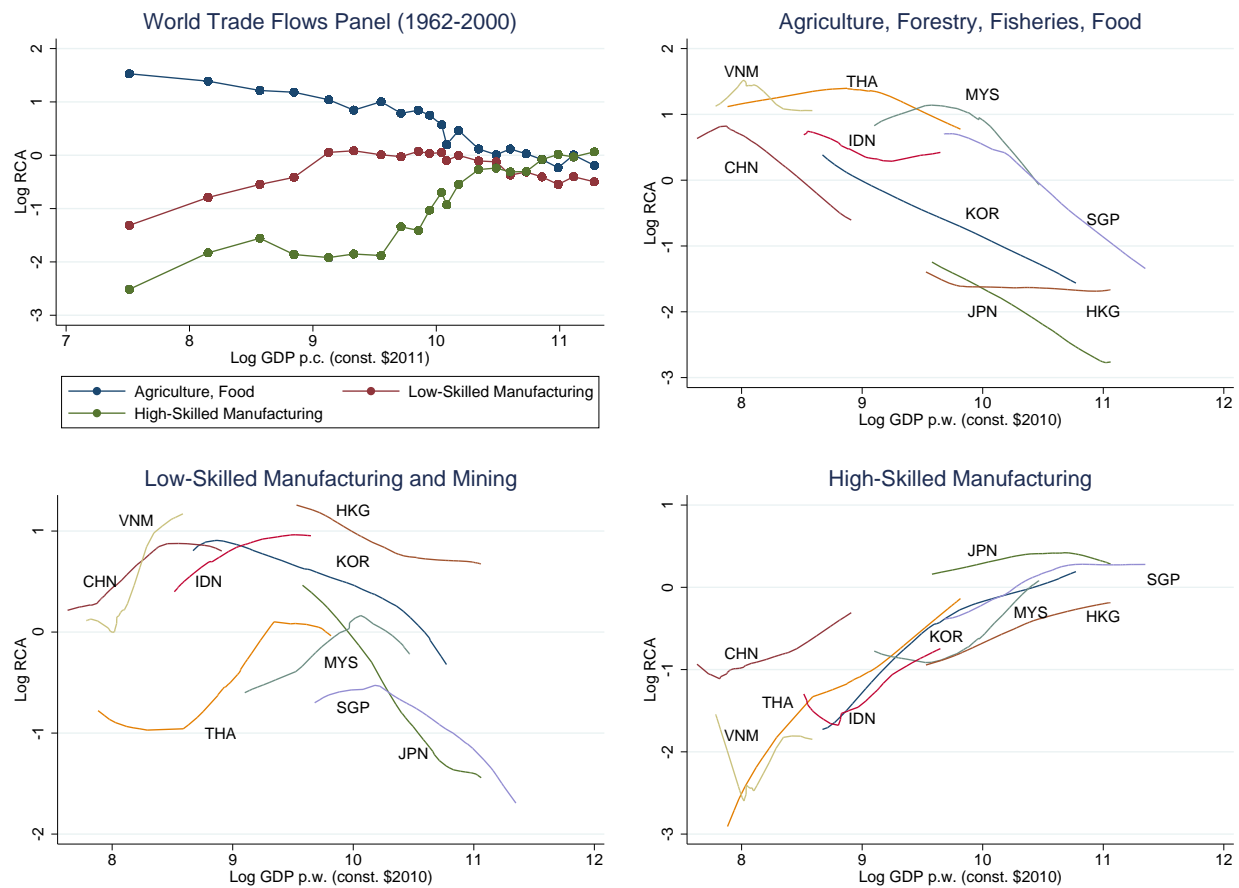
In order to explore to what extent comparative advantage (or relative productivity) tends to spill over between industries, I need a sector-specific measure of 'related' comparative advantage in similar sectors, which in turn requires a notion of similarity. Given that sectors that are similar in skill intensity tend to have similar levels of RCA (Fact 1), I posit a simple production function in which a firm in sector  $k$  combines inputs  $t_{n,t}^{a,k}$  from occupations (denoted by  $a$ ) with different skill levels:

$$Q_{n,t}^k = \prod_{a=1}^A (t_{n,t}^{a,k})^{\zeta_a^k} \quad ; \quad \sum_{a=1}^A \zeta_a^k = 1 \quad (2)$$

where  $\zeta_a^k$  is sector  $k$ 's input intensity of occupation  $a$ , which has an empirical equivalent as the occupation  $a$ 's share of wages or employment in sector  $k$ . For details on the occupational classification and data used, see section 5.

I can now define 'Revealed Occupational Advantage' (ROA) as the share of exports

Figure 2: Structural Transformation of Comparative Advantage in Tradable Clusters



**Notes:** The upper left figure presents a binned scatterplot of countries' revealed comparative advantage for the three tradable clusters against log GDP p.w. for the World Trade Flows sample from 1970 to 2000. The other three figures presents series of revealed comparative advantage in the three clusters for fast-growing East Asian countries (Vietnam, China, Thailand, Indonesia, Malaysia, Korea, Japan, Singapore, and Hong Kong SAR), smoothed using a lowess smoother. Real GDP per worker (constant \$2010) data are from Penn World Tables 9.0.

attributed to an occupation  $a$  in country  $n$  relative to the global average:

$$ROA_{n,t}^a = \frac{\sum_{k=1}^K \zeta_a^k X_{n,t}^k / X_{n,t}}{\sum_{k=1}^K \zeta_a^k X_t^k / X_t} \quad (3)$$

In turn, I construct a sector's 'related' RCA as a Cobb-Douglas aggregate of a country's ROA terms, with a sector's occupation cost share as exponents. For example, if a country has a high RCA in chemicals and aircrafts then it will also have a high 'related RCA' in office machinery as these sectors all use high-skill occupations relatively intensively. Formally, 'related RCA'  $RR_{n,t}^k$  of sector  $k$  in country  $n$  is defined as:

$$RR_{n,t}^k = \Pi_{a=1}^A (ROA_{n,t}^a)^{\zeta_a^k} \quad (4)$$

The final estimation equation becomes:

$$\Delta \ln RCA_{n,t}^k = \beta_0 + \beta_1 \ln RCA_{n,t-1}^k + \beta_2 \ln RR_{n,t-1}^k + \delta_{n,k} + \delta_{n,t} + \delta_{k,t} + \epsilon_{n,t}^k \quad (5)$$

where  $\Delta \ln RCA_{n,t}^k$  is a sector's log 10 year difference in RCA.  $\delta_{n,k}$ ,  $\delta_{n,t}$ , and  $\delta_{k,t}$  are country-sector, country-time and sector-time fixed effects, respectively, and  $\epsilon_{n,t}^k$  is an error term.<sup>17</sup> I estimate 5 using OLS on the WTF panel.<sup>18,19</sup>

Table 1 reports the associated regressions. In the first four specifications, the unconditional convergence coefficient on  $\ln RCA_{n,t-1}^k$  is negative and significant around -.25, which is close to the average estimates in [Levchenko and Zhang \(2016\)](#). The coefficient on related RCA is positive, significant, and around .24 with or without including sector-year fixed effects. This indicates substantial inter-industry spillovers, as a 10 percent increase in initial related RCA is associated with a roughly 2.3 percent higher subsequent growth in RCA. However, this average masks substantial heterogeneity. In columns (2) and (4), I interact a sector's related RCA with a dummy for the cluster to which it belongs. Most of the average inter-industry spillovers are driven by high-skilled manufacturing sectors, for which the coefficient on related RCA edges around 1, implying a 10 percent

<sup>17</sup>Under certain conditions, equation 5 corresponds to the estimating equation of endogenous changes in sector-specific unit costs in the model developed in section 3.2. See Appendix B for details.

<sup>18</sup>For details on sectors and countries covered, see Appendix F.

<sup>19</sup>I consider only country-year cells that contain at least 30 sectors after dropping observations that are within the 0.1 % tails in terms of  $\Delta \ln RCA_{n,t}^k$ .

Table 1: Inter-Industry Spillovers of Revealed Comparative Advantage

	(1)	(2)	(3)	(4)	(5)	(6)
	$\Delta \ln RCA_{n,t}^k$	$\Delta \ln RCA_{n,t}^k$	$\Delta \ln RCA_{n,t}^k$	$\Delta \ln RCA_{n,t}^k$	$\Delta \ln RCA_{n,t}^k$	$\Delta \ln RCA_{n,t}^k$
$\ln RCA_{n,t-1}^k$	-0.211*** (0.00438)	-0.233*** (0.00475)	-0.220*** (0.0131)	-0.235*** (0.00493)	-0.877*** (0.0190)	-0.883*** (0.00858)
$\ln RR_{n,t-1}^k$	0.240*** (0.0200)		0.245*** (0.0499)		0.431*** (0.115)	
$HSM \cdot \ln RR_{n,t-1}^k$		0.986*** (0.0712)		1.123*** (0.0781)		1.218*** (0.122)
$LSM \cdot \ln RR_{n,t-1}^k$		0.131** (0.0566)		0.0585 (0.0621)		0.916*** (0.107)
$AG \cdot \ln RR_{n,t-1}^k$		0.145*** (0.0192)		0.145*** (0.0191)		0.159*** (0.0449)
Country-year FE	Y	Y	Y	Y	Y	Y
Sector-year FE	N	N	Y	Y	Y	Y
Country-sector FE	N	N	N	N	Y	Y
Observations	123502	123502	123502	123502	123450	123450
Adjusted $R^2$	0.359	0.366	0.385	0.392	0.720	0.720

Standard errors, clustered at country-year level, in parentheses

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

**Notes:** This table reports estimates of different specifications of estimating equation 5. Columns (1) and (2) includes only country-year fixed effects, with and without interaction effects. Columns (3) and (4) add sector-year fixed effects. Columns (5) and (6) add country-sector fixed effects. Interaction effects include a dummy interacted with lagged related revealed comparative advantage (RR) for the three clusters of tradable sectors: High-Skilled Manufacturing (HSM), Low-Skilled Manufacturing and Mining (LSM) and Agriculture and Food (AG).

higher subsequent RCA growth rate. At the same time, the coefficient on related RCA is close to zero for the other two low-skilled clusters.

## 2.3 Preliminary Conclusions

Taken together, these two motivating facts suggest that (i) economies become relatively more productive in skill-intensive tradable sectors as they get richer, (ii) sector-specific productivity -relative to the country mean- exhibits convergence over time, and (iii) producing in a given sector leads to (above average) productivity growth in occupationally similar sectors, and (iv) these inter-industry spillovers are mainly present in high-skilled

manufacturing industries.<sup>20</sup>

### 3. Theoretical Framework

The model developed in this section has two components. The static component entails a multi-country multi-sector GE model. The main difference with the canonical model of [Caliendo and Parro \(2015\)](#) is its production structure. Rather than combining capital and labor, firms employ different combinations of occupations. In this sense, the static part of the model is very similar to that of [Lee \(2015\)](#).

The dynamic component of the model endogenizes the evolution of an economy's aggregate occupational productivity levels. In particular, I model this type of productivity growth in the form of dynamic scale economies. Within a period, production is constant returns to scale, but over time, countries endogenously increase their productivity in different tasks through learning-by-doing or human capital spillovers.

The model is tractable enough to be able to perform counterfactuals with closed form solutions. Nevertheless, estimating inter-industry productivity spillovers requires only four key assumptions: (i) bilateral trade takes a gravity form at the sectoral level<sup>21</sup>, (ii) goods and factor markets are competitive, (iii) sector-specific Hicks-neutral TFP terms are orthogonal to occupation-specific cost shares, and (iv) agents do not internalize any of their effects on future productivity. The gravity equation and competitive market assumptions guarantee bilateral trade flows reflect effective unit costs. The last two assumptions ensure that inter-industry productivity spillovers reflect the mapping from initial export structure to sector-specific changes in effective unit costs.

The model can be extended in several ways. In the appendix, I present three exten-

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<sup>20</sup>Note that these four patterns may shed a light on a puzzle in the growth literature highlighted by [Rodrik \(2012\)](#), who documents that, even though aggregate labor productivity tends to exhibit extremely slow unconditional (cross-country) convergence, sector-specific labor productivity converges much faster. Heterogeneous inter-industry productivity spillovers may play a role here. If spillovers are stronger in high-skill intensive sectors -in which richer countries tend to have a comparative advantage- then poorer countries may catch up at the sector-level but not in the aggregate. I explore the role of inter-industry spillovers in accounting for cross-country convergence in more detail in section 7.1.

<sup>21</sup>In the model below I micro-found the gravity equation by using a multi-sector set-up of [Eaton and Kortum \(2002\)](#) as developed by [Costinot et al. \(2011\)](#). However, this specification is not necessary for any of the empirical results of the paper. Alternatively, one could use the supply side of several other models that deliver a gravity equation, such as those of [Anderson and Van Wincoop \(2003\)](#) or [Krugman \(1979\)](#).



sions. First, I show that the model changes little when allowing for capital as an additional input in production. The static component of the model is isomorphic to a model with capital. Second, I incorporate production and trade of intermediate goods in the model, which magnifies any effects of trade integration (as in [Caliendo and Parro \(2015\)](#)) but does not change the dynamics of the model. Third, I show how the labor market of the model can be micro-founded using a Roy model of worker assignment. To facilitate exposition, I will now present the model in the main text without these extensions.

### 3.1 Static Framework

#### 3.1.1 Environment

The world consists of  $N$  countries indexed  $n \in \mathbf{N} = \{1, \dots, N\}$ . In each country, there are  $K$  tradable sectors indexed  $k \in \mathbf{K} = \{1, \dots, K\}$ . In turn, each sector is composed of a continuum of product varieties indexed  $\omega_k \in \Omega = \{1, \dots, +\infty\}$ . The production of a variety entails combining services from different types of occupations  $a \in \mathbf{A} = \{1, \dots, A\}$ . Finally, in every country there is a continuum of households -measure  $L_{n,t}$ - that supply labor inelastically.

#### 3.1.2 Demand

Households consume a bundle of sector aggregates  $\{C_{n,t}^k\}_{k=1}^K$ . Preferences are Cobb-Douglas with country- and time-specific weights  $\alpha_{n,t}^k$ :

$$U(\{C_{n,t}^k\}_{k=1}^K) = \prod_{k=1}^K (C_{n,t}^k)^{\alpha_{n,t}^k} \quad \sum_{k=1}^K \alpha_{n,t}^k = 1 \quad (6)$$

Households have two sources of income: wages,  $w_{n,t}$  and deficits,  $D_{n,t}/L_{n,t}$ .<sup>22</sup> To maximize welfare, a typical household picks expenditure shares  $\alpha_{n,t}^k$  s.t.

$$\alpha_{n,t}^k = \frac{P_{n,t}^k C_{n,t}^k}{w_{n,t} + D_{n,t}/L_{n,t}} \quad (7)$$

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<sup>22</sup>Trade deficits are necessary to exactly match observed data when computing equilibria in counterfactual changes. Throughout the paper, they serve as a source of exogenous income for households.

### 3.1.3 Production

In a given sector, productions entails combining services of different types of occupations. In the rest of the paper, I will refer to these services as *tasks*. There are three types of firms:

- **Task Producers** use labor  $L_{n,t}^a$  to produce task services.
- **Variety Producers** combine task services to produce quantity  $q_{n,t}^k(\omega_k)$  of variety  $\omega_k$  in sector  $k$  of country  $n$ .
- **Sector Aggregators** combine varieties from the lowest cost producers of  $\omega_k$  to produce the sector aggregate  $Q_{n,t}^k$ . These varieties can be imported from any country.

#### Task Producers

Tasks are indexed  $a \in \mathbf{A} = \{1, \dots, A\}$ . The task production market is perfectly competitive and market prices of tasks are denoted by  $p_{n,t}^a$ . A firm in county  $n$  producing services of task  $a$  hires  $L_{n,t}^a$  units of labor with mean task productivity  $T_{n,t}^a$  at a market price  $w_{n,t}^a$  per effective unit. A typical task producer solves the problem:

$$\max_{L_{n,t}^a \geq 0} p_{n,t}^a T_{n,t}^a L_{n,t}^a - w_{n,t}^a T_{n,t}^a L_{n,t}^a$$

#### Variety Producers

The market for varieties is perfectly competitive. A firm in country  $n$  producing variety  $\omega_k$  in sector  $k$  hires  $A$  different task inputs  $\{t_{n,t}^a(\omega_k)\}_{a=1}^A$  to produce quantity  $q_{n,t}^k(\omega_k)$ . The TFP of a variety producer  $z_{n,t}^k(\omega_k)$  is a random draw from a Frechet distribution with location and dispersion parameters  $T_{n,t}^k$  and  $\theta > 1$ . The variety's production function is

$$q_{n,t}^k(\omega_k) = z_{n,t}^k(\omega_k) \prod_{a=1}^A (t_{n,t}^a(\omega_k))^{\zeta_a^k} \quad (8)$$

where factor shares  $\zeta_a^k$  are uncorrelated with the sector-specific location parameter of TFP,  $T_{n,t}^k$ .<sup>23</sup>  $\zeta_a^k$  are crucial parameters as they capture the similarity of production func-

<sup>23</sup>The assumption of independence between factor shares and sector-specific TFP ensures that a linear fitted relationship between sector-specific unit costs and factor shares reflects differences in relative task productivity levels. See also [Malmberg \(2017\)](#).

tions in any two given sectors, thereby governing the strength of inter-industry productivity spillovers. A typical variety producer solves the problem

$$\max_{\{t_{n,t}^a(\omega_k) \geq 0\}_{a=1}^A} p_{n,t}^k(\omega_k) q_{n,t}^k(\omega_k) - \sum_{a=1}^A p_{n,t}^a t_{n,t}^a(\omega_k)$$

### Sector Aggregators

The representative aggregator firm of sector  $k$  in country  $n$  combines varieties from the lowest cost producers of  $\omega_k$  to produce the sector aggregate. Varieties can be imported from any country. The firm sells the sector aggregate to consumers and material producers in country  $n$ .

Aggregators face sector-specific trade costs that vary by importer-exporter pair and are denoted by  $\tau_{nm,t}^k$  for importer  $n$  and exporter  $m$ . The trade costs take the usual iceberg form and satisfy the triangle inequality.

A typical aggregator solves the problem

$$\begin{aligned} \max_{\{q_{n,t}^k(\omega_k)\}_{\omega_k \in \Omega}} & P_{n,t}^k Q_{n,t}^k - \int_0^1 p_{n,t}^k(\omega_k) q_{n,t}^k(\omega_k) \\ \text{s.t. } Q_{n,t}^k &= \left[ \int_0^1 (q_{n,t}^k(\omega_k))^{\frac{\xi-1}{\xi}} d\omega_k \right]^{\frac{\xi}{\xi-1}} \quad ; \quad \xi > 0 \end{aligned}$$

where  $p_{n,t}^k(\omega_k)$  is the (unique) minimum price at which a firm in sector  $k$  producing variety  $\omega_k$  can deliver that variety in country  $n$ , i.e.  $p_{n,t}^k(\omega_k) = \min\{\frac{c_{i,t}^k(\omega_k) \tau_{ni,t}^k}{z_{i,t}^k(\omega_k)}; i = 1, \dots, N\}$ .

### Households

Before being hired by a firm, households are homogeneous with respect to their task-specific productivity levels. Given task wages  $\{w_{n,t}^a\}_{a=1}^A$  a household maximizes its income by sorting into a specific task  $a$ . This problem can be summarized as

$$\max_{\{\tilde{a}_a\}_{a=1}^A} \sum_{a=1}^A \tilde{a}_a w_{n,t}^a T_{n,t}^a \quad \text{s.t. } \sum_{a=1}^A \tilde{a}_a = 1; \quad \{\tilde{a}_a\}_{a=1}^A \in \{0, 1\}^A$$

### 3.1.4 Equilibrium

#### Static Level Equilibrium

Given country-specific fundamentals, a static level equilibrium consists of a vector of prices

$$\{w_{n,t}, \{P_{n,t}^k\}_{k=1}^K, \{w_{n,t}^a, p_{n,t}^a\}_{a=1}^A\}_{n=1}^N \text{ s.t.}$$

- Task producers minimize unit costs:

$$p_{n,t}^a = w_{n,t}^a \quad (9)$$

- Variety producers minimize unit costs:<sup>24</sup>

$$c_{n,t}^k = \Gamma_k \prod_{a=1}^A (p_{n,t}^a)^{\zeta_a^k}$$

- Aggregators source from lowest cost producers such that expenditure shares  $\pi_{ni,t}^k$  and prices of sector aggregates  $P_{n,t}^k$  equal:<sup>25</sup>

$$P_{n,t}^k = \Lambda^k [\sum_{i=1}^N T_{i,t}^k (c_{i,t}^k \tau_{ni,t}^k)^{-\theta}]^{-1/\theta}$$

$$\pi_{ni,t}^k = \frac{T_{i,t}^k [c_{i,t}^k \tau_{ni,t}^k]^{-\theta}}{\sum_{i'=1}^N T_{i',t}^k [c_{i',t}^k \tau_{ni',t}^k]^{-\theta}}$$

- Households maximize income and utility:

$$w_{n,t}^a T_{n,t}^a = w_{n,t}^{a'} T_{n,t}^{a'} = w_{n,t} \quad \forall a, a' \in \mathbf{A} \quad (10)$$

$$\alpha_{n,t}^k = \frac{P_{n,t}^k C_{n,t}^k}{w_{n,t} + D_{n,t}/L_{n,t}} \quad (11)$$

- Trade is balanced for each country:

$$-\sum_{k=1}^K \sum_{i=1}^N \alpha_{n,t}^k (w_{n,t} L_{n,t} + D_{n,t}) \pi_{ni,t}^k = \sum_{k=1}^K \sum_{i=1}^N \alpha_{i,t}^k (w_{i,t} L_{i,t} + D_{i,t}) \pi_{in,t}^k + D_{n,t}$$

- Goods and labor markets clear in each country

### Static Counterfactual Equilibrium

Let  $\{w_{n,t}, \{P_{n,t}^k\}_{k=1}^K, \{w_{n,t}^a, p_{n,t}^a\}_{a=1}^A\}_{n=1}^N$  denote an equilibrium under a set of country-specific fundamentals, and let  $\{w'_{n,t}, \{(P_{n,t}^k)'\}_{k=1}^K, \{(w_{n,t}^a)', (p_{n,t}^a)'\}_{a=1}^A\}_{n=1}^N$  denote an equilibrium under a different set of country-specific fundamentals. We can now define an equilibrium in relative changes, where a variable with a hat ( $\hat{x}$ ) represents the relative change of that variable ( $\hat{x} = x'/x$ ). A static counterfactual equilibrium consists of a vector of relative counterfactual prices  $\{\hat{w}_{n,t}, \{\hat{P}_{n,t}^k\}_{k=1}^K, \{\hat{w}_{n,t}^a, \hat{p}_{n,t}^a\}_{a=1}^A\}$  s.t.

<sup>24</sup> $\Gamma_k$  is a sector-specific constant. It does not play any role in this paper.

<sup>25</sup> $\Lambda_k$  is a sector-specific constant. It does not play any role in this paper.

- $\hat{p}_{n,t}^a = \hat{w}_{n,t}^a$
- $\hat{c}_{n,t}^k = \prod_{a=1}^A (\hat{p}_{n,t}^a)^{\zeta_a^k}$
- $\hat{P}_{n,t}^k = [\sum_{i=1}^N \pi_{ni,t}^k \hat{T}_{i,t}^k (\hat{c}_{i,t}^k \hat{\tau}_{ni,t}^k)^{-\theta}]^{-1/\theta}$
- $\hat{\pi}_{ni,t}^k = \hat{T}_{i,t}^k [\frac{\hat{c}_{i,t}^k \hat{\tau}_{ni,t}^k}{\hat{P}_{n,t}^k}]^{-\theta}$
- $\hat{w}_{n,t}^a \hat{T}_{n,t}^a = \hat{w}_{n,t}^{a'} \hat{T}_{n,t}^{a'} = \hat{w}_{n,t} \quad \forall (a, a')$
- $\sum_{k=1}^K \sum_{i=1}^N \alpha_{n,t}^k (w_{n,t} L_{n,t} \hat{w}_{n,t} + D_{n,t}) \pi_{ni,t}^k \hat{\pi}_{ni,t}^k = \sum_{k=1}^K \sum_{i=1}^N \alpha_{i,t}^k (w_{i,t} L_{i,t} \hat{w}_{i,t} + D_{i,t}) \pi_{in,t}^k \hat{\pi}_{in,t}^k + D_{n,t}$

## 3.2 Dynamic Framework

### 3.2.1 Endogenous Task Productivity Growth

#### Task Production

The production of a task entails the completion of a continuum of *subtasks*  $\omega_a \in [0, 1]$ .<sup>26</sup> A task producer hires a measure  $L_{n,t}^a$  of ex ante homogeneous workers at the effective wage rate  $w_{n,t}^a T_{n,t}^a$ . After being hired, each worker is uniformly assigned to do one of the subtasks and produces this subtask at productivity  $z_{n,t}(\omega_a)$ . A worker's productivity  $z_{n,t}(\omega_a)$  is drawn from a productivity distribution  $G_{n,t}^a$  and represents the state-of-the-art technology or idea about how to produce subtask  $\omega_a$ . Together, these workers produce a quantity  $q_{n,t}^a$  of task  $a$  using a CES aggregator:

$$q_{n,t}^a = T_{n,t}^a L_{n,t}^a = \left[ \int_0^1 (z_{n,t}^a(\omega_a))^{\frac{\chi-1}{\chi}} d\omega_a \right]^{\frac{\chi}{\chi-1}} L_{n,t}^a \quad \chi > 0 \quad (12)$$

where  $T_{n,t}^a$  captures the average productivity of a worker.  $\chi$  is the elasticity of substitution between different subtasks.<sup>27</sup>

#### New Ideas

<sup>26</sup>The concept of a *subtask* is analogous to a task as a variety is to a sector in Buera and Oberfield (2020). It is similar to the *trade* of a craftsman in De la Croix et al. (2017), who model historical labor productivity growth as 'learning on the shopfloor' through personal contact between a designated 'master' and an apprentice. The concept is not strictly necessary for any of the results in the paper but facilitates exposition.

<sup>27</sup>The value of  $\chi$  is irrelevant for any of the empirical results as it affects only the absolute level, but not the relative level or growth of  $T_{n,t}^a$ .

Within each time period, a worker assigned to subtask  $\omega_a$  receives  $n_{n,t}^a = n \cdot \tilde{n}_{n,t}^a$  new, random ideas with productivity  $z$  from exogenous distribution with CDF  $H(z)$ . This distribution has a Pareto right tail with exponent  $\theta_H$  such that  $\lim_{z \rightarrow \infty} (1 - H(z))/z^{-\theta_H} = 1$ . Moreover, I assume that the initial knowledge frontier follows a Frechet distribution.<sup>28</sup>

### Productivity Spillovers

Each worker combines an original random idea with insights from others. Every time a worker receives random idea, it meets others with probability  $p_{n,t}^a = p \cdot \tilde{p}_{n,t}^a$ , so the number of successful meetings of a worker of type  $a$  follows a Binomial distribution with parameters  $(n_{n,t}^a, p_{n,t}^a)$ . As  $n \rightarrow \infty$  while  $n \cdot p$  remains constant, this process converges to a Poisson distribution with an arrival rate of  $\eta_{n,t}^a = n_{n,t}^a p_{n,t}^a$ . When the two workers meet, they exchange ideas and potentially engage in technology adoption. I assume that this form of learning is external to any individual firm and/or worker.

If the worker producing variety  $\omega_a$  chooses to adopt, the actual productivity of the new technology is  $(z_{n,t}^a(\omega'_a))^\beta (z_{n,t}^a(H))^{1-\beta}$ .  $\beta \in [0, 1)$  is an adoption parameter that captures the importance of ideas of others. A worker only adopts if the new technology is better than the old one, i.e. if  $(z_{n,t}^a(\omega'_a))^\beta (z_{n,t}^a(H))^{1-\beta} > z_{n,t}^a(\omega_a)$ .

Under these assumptions, the state-of-the-art productivity levels of a given variety  $\omega_a$  are distributed Frechet (Buera and Oberfield, 2020) with CDF  $F_{n,t}^a = \exp(-\tilde{T}_{n,t}^a z^{-\theta_a})$ , where  $\theta_a = \frac{\theta_H}{1-\beta}$ . The latter is assumed to be invariant across tasks. The location parameter  $\tilde{T}_{n,t}^a$  of this distribution follows the law of motion:

$$\frac{\tilde{T}_{n,t}^a}{\tilde{T}_{n,t-1}^a} = n_{n,t}^a p_{n,t}^a \int_0^\infty x^{\beta\theta_a} dG_{n,t}^a(x) = n_{n,t}^a p_{n,t}^a \Gamma(1-\beta) (\tilde{T}_{n,t-1}^a)^{\beta-1} \quad (13)$$

where  $\Gamma(\cdot)$  is the Gamma function.

Expected task productivity,  $T_{n,t}^a$  equals:

$$T_{n,t}^a = (\tilde{T}_{n,t}^a)^{1/\theta_a} \left[ \Gamma\left(1 - \frac{\chi-1}{\chi} \frac{1}{\theta_a}\right) \right]^{\frac{\chi}{\chi-1}} \quad (14)$$

<sup>28</sup>To satisfy Assumption 1 in Buera and Oberfield (2020), the *initial* frontier distribution of knowledge also needs to have a sufficiently thin right tail. This additional assumption is satisfied if the initial frontier of knowledge follows a Frechet distribution.

which follows the law of motion (in logs):

$$\Delta \ln T_{n,t}^a = \frac{1}{\theta_a} [\ln n_{n,t}^a p_{n,t}^a + \ln \Gamma(1 - \beta) + (\beta - 1) \ln T_{n,t-1}^a] \quad (15)$$

where  $\ln n_{n,t}^a p_{n,t}^a$  captures the arrival of new, successfully adopted ideas,  $\ln \Gamma(1 - \beta)$  captures an exogenous, time- and task-invariant component of productivity growth, and  $(\beta - 1) \ln T_{n,t-1}^a$  captures a 'fishing-out' effect, as more productive ideas become harder to find when the knowledge frontier expands (as  $\beta - 1 < 0$ ).

### Dynamic Scale Economies

There is a long-standing literature in trade and growth examining the importance of dynamic scale economies for economic convergence and the dynamic welfare gains from trade. In a large set of models the growth in productivity  $T_{n,t}^x$  of a given sector or factor  $x$  takes the form:

$$\Delta \ln T_{n,t}^x = \beta_x + \tilde{\eta}_x \ln L_{n,t-1}^x + \phi \ln T_{n,t-1}^x; \quad \tilde{\eta}_x > 0; \quad \phi \leq 0 \quad (16)$$

where  $\beta_x$  is a constant, and  $\phi \ln T_{n,t-1}^x$  captures convergence in productivity.  $\tilde{\eta}_x \ln L_{n,t-1}^x$  reflect dynamic scale economies, i.e. dynamic increasing returns to the use of a production factor. The exact source of dynamic scale economies can differ depending on the setting, although most papers posit them either as a result of learning-by-doing (e.g. [Krugman, 1987](#); [Lucas Jr, 1988](#); [Matsuyama, 1992](#); [Redding, 1999](#); [Mendoza, 2010](#)) or human capital spillovers (e.g. [Lucas Jr, 1988](#); [Stokey, 1991](#); [Lucas, 2004](#); [Lucas Jr, 2015](#)).<sup>29</sup>

Consider the framework that generates the law of motion of task productivity growth, equation 15. Suppose the arrival rate of ideas is constant ( $n_{n,t}^a = n \forall a, n, t$ ). Whether a meeting is successful (with probability  $p_{n,t}^a$ ) depends on the share of workers engaged in the production of the task and task-specific learning spillovers. The latter manifest as increasing returns to scale in the production of task-specific knowledge. Specifically, the extent of learning spillovers depends on the required *team size* for a successful meeting,

<sup>29</sup>A good example is the learning-by-doing model in [Matsuyama \(1992\)](#), later used by [Mendoza \(2010\)](#). If production is Cobb-Douglas, sector-specific TFP evolves according to  $\Delta \ln T_{n,t}^k = \beta_k + \zeta \ln L_{n,t-1}^k$  where  $\zeta \in (0, 1)$  is the Cobb-Douglas exponent and  $L_{n,t}^k$  the share of labor employed in sector  $k$ .



$\tilde{\eta}_a \geq 0$ . I treat  $\tilde{\eta}_a$  as a continuous variable in the rest of the paper.<sup>30</sup> In every time period, a worker of type  $a$  is randomly assigned to  $\tilde{\eta}_a - 1$  other workers of the same type. The worker's meeting is successful if and only if all members of its team meet another worker of the same type, so  $p_{n,t}^a = (L_{n,t}^a/L_{n,t})^{\tilde{\eta}_a}$ , where  $L_{n,t}^a/L_{n,t}$  is the share of workers in country  $n$  that are engaged in the production of task  $a$ .

I will now consider two extreme cases of these dynamic scale economies: pure learning-by-doing ( $\beta = 0$ ) and pure human capital spillovers ( $\beta \rightarrow 1$ ).

### Pure Learning-by-doing

If there is no diffusion of ideas between workers ( $\beta = 0$ ), successfully arrived ideas for active workers are the only source of productivity growth. Any differences in task productivity growth between countries  $n$  and  $n'$  are driven by differences in the allocation of labor and initial task productivity, i.e.

$$\Delta \ln T_{n,t}^a/T_{n',t}^a = \frac{1}{\theta_a} [\tilde{\eta}_a \ln \frac{L_{n,t-1}^a/L_{n',t-1}^a}{L_{n,t-1}/L_{n',t-1}} + \ln T_{n,t-1}^a/T_{n',t-1}^a] \quad (17)$$

### Pure Human Capital Spillovers

If active workers do not receive any new ideas ( $\beta \rightarrow 1$ ), existing ideas from other workers are the only source of productivity growth to contribute meaningfully to any cross-country differences. As a consequence, any differences in task productivity growth between countries  $n$  and  $n'$  are approximately driven by differences in the allocation of labor only, i.e.

$$\Delta \ln T_{n,t}^a/T_{n',t}^a \approx \frac{1}{\theta_a} [\tilde{\eta}_a \ln \frac{L_{n,t-1}^a/L_{n',t-1}^a}{L_{n,t-1}/L_{n',t-1}}] \quad (18)$$

### Combination of Mechanisms

In the rest of the paper, I do not take a stance on the relative importance of learning-by-doing or human capital spillovers, but posit the arrival rate of new ideas as:

$$\ln n_{n,t}^a p_{n,t}^a = \tilde{\eta}_a \ln L_{n,t-1}^a/L_{n,t-1} + \epsilon_{n,t}^a \quad (19)$$

<sup>30</sup>The continuous interpretation is necessary in order to map the model to the data. Alternatively, one could interpret the estimated rate  $\hat{\eta}_a$  in section IV as an average treatment effect of time- and/or country-dependent  $\tilde{\eta}_{a,t}^n$ .

where  $\epsilon_{n,t}^a$  captures the arrival of new ideas independent from the scale of production, and  $\tilde{\eta}_a$  is the elasticity of total new ideas with respect to the scale of production,  $L_{n,t}^a$ . As a consequence, the law of motion for task productivity growth now becomes:

$$\Delta \ln T_{n,t}^a = \frac{1}{\theta_a} [\ln \Gamma(1 - \beta) + \tilde{\eta}_a \ln L_{n,t-1}^a / L_{n,t-1} + (\beta - 1) \ln T_{n,t-1}^a + \epsilon_{n,t}^a] \quad (20)$$

### Relation to Static External Economies of Scale

In the appendix (section C) I present a mapping between the dynamic scale economies in this paper and the external economies of scale in the recent papers by [Bartelme et al. \(2019a\)](#) and [Kucheryavyy et al. \(2020\)](#). In particular, I show that in a steady state, dynamic scale economies are isomorphic to static external economies of scale as specified in [Kucheryavyy et al. \(2020\)](#).

## 3.2.2 Equilibrium

### Dynamic Level Equilibrium

Given country-specific fundamentals and initial task productivity levels  $\{T_{n,t=1}^a\}_{a=1}^A$ , a dynamic level equilibrium consists of a vector of prices

$$\{\{w_{n,t}, \{P_{n,t}^k\}_{k=1}^K, \{w_{n,t}^a, p_{n,t}^a\}_{a=1}^A\}_{n=1}^N\}_{t=1}^T \text{ s.t.}$$

- In each period  $t = \tau$ , given country-specific fundamentals and task productivity levels  $\{T_{n,\tau}^a\}_{a=1}^A$ , the price vector  $\{w_{n,\tau}, \{P_{n,\tau}^k\}_{k=1}^K, \{w_{n,\tau}^a, p_{n,\tau}^a\}_{a=1}^A\}_{n=1}^N$  solves a static level equilibrium, as defined in section 3.1.4.
- Between two periods  $t = \tau$  and  $t = \tau - 1$ , task productivity levels  $\{\{T_{n,\tau}^a\}_{a=1}^A\}_{t=1}^T$  satisfy

$$\Delta \ln T_{n,t}^a = \frac{1}{\theta_a} [\ln \Gamma(1 - \beta) + \tilde{\eta}_a \ln L_{n,t-1}^a / L_{n,t-1} + (\beta - 1) \ln T_{n,t-1}^a + \epsilon_{n,t}^a]$$

### Dynamic Counterfactual Equilibrium

A dynamic counterfactual equilibrium consists of a vector of relative counterfactual prices

$$\{\{\hat{w}_{n,t}, \{\hat{P}_{n,t}^k\}_{k=1}^K, \{\hat{w}_{n,t}^a, \hat{p}_{n,t}^a\}_{a=1}^A\}_{n=1}^N\}_{t=1}^T \text{ s.t.}$$

- In each period  $t = \tau$ , given counterfactual task productivity levels  $\{\{\hat{T}_{n,t}^a\}_{a=1}^A\}_{t=1}^T$ , the price vector  $\{\hat{w}_{n,t}, \{\hat{P}_{n,t}^k\}_{k=1}^K, \{\hat{w}_{n,t}^a, \hat{p}_{n,t}^a\}_{a=1}^A\}_{n=1}^N$  solves a static counterfactual equilibrium, as defined in 3.1.4.

- Between two periods  $t = \tau$  and  $t = \tau - 1$ , counterfactual task productivity levels

$\{\{\hat{T}_{n,\tau}^a\}_{a=1}^A\}_{t=1}^T$  satisfy

$$\hat{T}_{n,\tau}^a = [\hat{L}_{n,\tau-1}^a]^{\tilde{\eta}_a} (\hat{T}_{n,\tau-1}^a)^{\beta-1}$$

$$\hat{L}_{n,\tau-1}^a = \frac{\sum_{k=1}^K \zeta_a^k \sum_{i=1}^N (X_{in,\tau-1}^k)' \frac{1}{\tilde{w}_{n,\tau-1}}}{\sum_{k=1}^K \zeta_a^k \sum_{i=1}^N X_{in,\tau-1}^k}$$

### 3.2.3 Balanced Growth

A global balanced growth path exists in two cases. In both cases, the task-specific residual arrival rate  $\epsilon_{n,t}^a$  must grow at an exponential rate ( $\epsilon_{n,t}^a = \tilde{n}_0 \exp(\gamma_n^a t)$ ) to offset the fishing-out effect of a higher task productivity level (as  $\theta_a(\beta - 1) < 0$ ). Task productivity  $T_{n,t}^a$  grows at the rate  $\frac{\gamma_n^a}{1-\beta}$ .<sup>31</sup>

In the first case, all countries are closed ( $\tau_{ni,t}^k \rightarrow \infty \forall n \neq i$ ) and preferences and sector-specific productivity levels are constant over time. In this case, welfare grows at the country-specific rate  $\frac{1}{1-\beta} \sum_{k=1}^K \alpha_n^k \sum_{a=1}^A \zeta_a^k \gamma_n^a$ . Note that relatively closed economies (e.g. the United States) tend to approximately follow a balanced growth path. A second balanced growth path exists in a symmetric world in which all countries are same in terms of fundamentals (preferences, trade costs, and productivity levels) but any degree of (symmetric) trade integration can exist. In this case, welfare grows at the general rate  $\frac{1}{1-\beta} \sum_{k=1}^K \alpha^k \sum_{a=1}^A \zeta_a^k \gamma^a$ .

In general, a balanced growth path does not exist if economies differ in terms of fundamentals and there is some degree of trade integration.<sup>32</sup> However, a unique equilibrium path exists for any sequence of fundamentals as a sequence of static equilibria defined in section 3.1.4. These equilibria are connected by the law of motion of task productivity levels in equation 15. As firms and workers do not internalize any of the learning spillovers, there is no forward-looking behavior and the sequence of task productivity levels is solely determined by initial productivity levels and the sequence of fundamentals.

<sup>31</sup>The necessity of the exponential growth of idea arrival rates mirrors traditional endogenous growth models (see Jones (1995), Bloom et al. (2017), Jones (2017), Jones (2019)).

<sup>32</sup>Note that this result is not a consequence of the lack of forward-looking behavior, but rather due to sectoral asymmetry in terms of dynamic scale economies. Empirically, balanced growth paths are seldom observed for developing economies. Since this group is the focus of this paper, I do not view the absence of a guaranteed balanced growth path as being inconsistent with the data.

### 3.3 Welfare Implications

#### 3.3.1 Static and Dynamic Gains from Trade

Counterfactual welfare in period  $t$  is given by

$$\ln \hat{w}_{n,t} / \hat{P}_{n,t} = - \sum_{k=1}^K \frac{\alpha_{n,t}^k}{\theta} \ln \hat{\pi}_{nn,t}^k + \sum_{k=1}^K \alpha_{n,t}^k \sum_{a=1}^A \zeta_a^k \tilde{\eta}_a \ln \hat{L}_{n,t-1}^a \quad (21)$$

This decomposition describes the welfare effects of integration as the product of two terms. The first term covers the usual static gains from trade (Arkolakis et al., 2012): the more open -as measured by domestic expenditure shares- a country becomes (smaller  $\hat{\pi}_{nn}^k$ ), the larger the welfare gains from trade. These gains are especially large if integration reduces a country's domestic expenditure shares more in sectors from which it consumes more (higher  $\alpha_{n,t}^k$ ).

The second term is new and summarizes the dynamic effects from trade integration. In contrast to the first term, the dynamic term can contribute to lower welfare gains and could potentially lead to welfare losses if it is larger than the static gains. The direction of this effect depends on the changes in labor allocation at period  $t = 1$ . If the trade shock induces countries to shift labor into tasks ( $\ln \hat{L}_{n,t-1}^a > 0$ ) with a high diffusion rate  $\tilde{\eta}_a$ , gains are higher. These gains are particularly elevated when labor reallocation leads to higher endogenous productivity growth in tasks that are used intensively (higher  $\zeta_k^a$ ) in sectors from which the country consumes more.

#### 3.3.2 Planner Solution in a Closed Economy

Welfare in equation 21 is not necessarily optimized in the case of free trade nor autarky. To maximize welfare, a country must shift labor into sectors that use high spillover tasks intensively, while minimizing associated static welfare losses.<sup>33</sup>

Consider a closed 2-period economy with  $K$  sectors and  $A$  tasks. Normalize  $L_{n,t} = 1$ . To optimize the sum of the static and dynamic welfare components in equation 21, a planner would solve the allocation of sector-specific task employment shares  $\{\{L_{n,t}^{a,k}\}_{a=1}^A\}_{k=1}^K$

<sup>33</sup>I do not discuss the case for industrial policy in detail here. The implied policy of encouraging production in technologically advanced sectors is in line with policies pursued by the Asian Miracles (Hong Kong SAR, Korea, Singapore, and Taiwan PoC) (Cherif and Hasanov, 2019).

in two steps.

First, as task-specific labor from any sector are perfect substitutes in their contribution to future productivity (equation 15), the share of labor of task  $a$  allocated to sector  $k$  equals its relative contribution to current marginal utility, i.e.

$$L_{n,t}^{a,k} = \frac{\alpha_n^k \zeta_a^k}{\sum_{k'=1}^K \alpha_n^{k'} \zeta_{a'}^{k'}} L_{n,t}^a \quad (22)$$

where  $L_{n,t}^a$  is the share of total labor employed in task  $a$ . In turn, this share equals its relative contribution to the sum of discounted marginal utility over the two periods, i.e.

$$L_{n,t}^a = \frac{(1 + \rho \tilde{\eta}_a) \sum_{k=1}^K \alpha_n^k \zeta_a^k}{\sum_{a'=1}^A (1 + \rho \tilde{\eta}_{a'}) \sum_{k'=1}^K \alpha_n^{k'} \zeta_{a'}^{k'}} \quad (23)$$

where  $\rho \in (0, 1)$  is a discount factor. In a decentralized equilibrium, on the other hand:

$$L_{n,t}^a = \frac{\sum_{k=1}^K \alpha_n^k \zeta_a^k}{\sum_{a'=1}^A \sum_{k'=1}^K \alpha_n^{k'} \zeta_{a'}^{k'}} \quad (24)$$

Note that the decentralized and planner equilibrium coincide if the strength of diffusion  $\tilde{\eta}_a$  is the same across tasks, or when  $\alpha_n^k \zeta_a^k$  is the same across sectors.<sup>34</sup> In the end, whether the decentralized allocation is optimal depends on the extent to which workers and firms internalize learning spillovers. To facilitate the analysis, I assume an extreme case in which there is not internalization whatsoever. An interesting question for further research is how the model's predictions would change if one loosens this assumption.

### 3.3.3 Gains from Trade in a Small Open Economy

While the decentralized allocation is not necessarily efficient under trade (nor autarky), this does not imply that the gains from trade are negative for countries that have a comparative advantage in low spillover sectors. First, even if there are dynamic efficiency losses from specializing in these kinds of sectors, trade integration always features static efficiency gains through lower prices. Second, if trade integration gives a country access to a larger export market in high spillover sectors, it can achieve productivity gains by

<sup>34</sup>The efficient equilibrium can be implemented using sector-specific transfers (e.g. explicit sector-specific wage subsidies paid for by a lump-sum tax on households), as in [Fajgelbaum and Gaubert \(2018\)](#).

exporting in these sectors while achieving static efficiency gains from cheaper imports.

Consider an open economy that is small in the sense that it does not impact other countries' productivity levels, wages and prices. Preferences differ by country but are constant over time. Over two periods, welfare in counterfactual changes takes the form

$$\ln \hat{w}_{n,t}/\hat{P}_{n,t} + \rho \ln \hat{w}_{n,t+1}/\hat{P}_{n,t+1} = - \sum_{k=1}^K [\frac{\alpha_n^k}{\theta} \ln \hat{\pi}_{nn,t}^k + \rho \frac{\alpha_n^k}{\theta} \ln \hat{\pi}_{nn,t+1}^k] + \sum_{k=1}^K \alpha_n^k \sum_{a=1}^A \zeta_a^k \tilde{\eta}_a \ln \hat{L}_{n,t}^a \quad (25)$$

where  $\rho \in (0, 1)$  is a discount factor.

Suppose that a country's trade costs change by a factor  $\tau$  that is common to all country-sector-pairs ( $\tau_{ni,t}^k = \tau \tilde{\tau}_{ni,t}^k$ ). In the short-run, the change only impacts the static welfare gains from trade. Total differentiation of the short-run impacts in equation 25 gives

$$\frac{d \ln \frac{w_{n,t}}{P_{n,t}}}{d\tau} = - \sum_{k=1}^K \frac{\alpha_n^k}{\theta} \frac{\partial \ln \pi_{nn,t}^k}{\partial \tau} = -\frac{1}{\tau} + \frac{\partial \ln w_{n,t}}{\partial \tau} [1 - \sum_{k=1}^K \alpha_n^k \frac{\partial \ln P_{n,t}^k}{\partial w_{n,t}}] \quad (26)$$

where the first term is the direct effect of the change in trade costs on the price index and the second term captures the general equilibrium effect of the change on the country's wage. In the long-run (at  $t + 1$ ), the change in trade costs also affects the country's productivity:

$$\frac{d \ln \frac{w_{n,t+1}}{P_{n,t+1}}}{d\tau} = - \sum_{k=1}^K \alpha_n^k [\frac{1}{\theta} \frac{\partial \ln \pi_{nn,t+1}^k}{\partial \tau} - \sum_{a=1}^A \zeta_a^k \frac{\partial \ln T_{n,t+1}^a}{\partial \tau}] = \frac{\partial \ln w_{n,t+1}}{\partial \tau} - \sum_{k=1}^K \alpha_n^k \frac{\partial \ln P_{n,t}^k(\tau, \{\ln T_{n,t+1}^a\}_{a=1}^A)}{\partial \tau} \quad (27)$$

which is equivalent to

$$-\frac{1}{\tau} + \frac{\partial \ln w_{n,t+1}}{\partial \tau} [1 - \sum_{k=1}^K \alpha_n^k \frac{\partial \ln P_{n,t+1}^k}{\partial w_{n,t+1}}] - \sum_{k=1}^K \alpha_n^k \sum_{a=1}^A \frac{\partial \ln P_{n,t}^k}{\partial \ln T_{n,t+1}^a} \frac{\partial \ln T_{n,t+1}^a}{\partial \tau} \quad (28)$$

where the first two terms are the short-term effects identified earlier, and the third term captures dynamic productivity effects of trade integration.

Dynamic scale economies affects welfare through  $\frac{\partial \ln T_{n,t+1}^a}{\partial \tau}$ . Unpacking this term further gives:

$$\frac{\partial \ln T_{n,t+1}^a}{\partial \tau} = \tilde{\eta}_a [\frac{\sum_{i=1}^N w_{i,t} \sum_{k=1}^K \alpha_{i,t}^k \zeta_a^k \frac{\partial \pi_{in,t}^k}{\partial \tau}}{L_{n,t}^a w_{n,t}} + \frac{\partial \ln w_{n,t}}{\partial \tau} [\frac{\partial \ln L_{n,t}^a w_{n,t}}{\partial \ln w_{n,t}} - 1]] \quad (29)$$

where the first term is the effect of reallocation of labor caused directly by the change in trade costs, and the second term captures the general equilibrium effect of the change on reallocation through a change in domestic income.

The first term reflects the effect of foreign demand on task productivity growth. If the change in the import share of foreign countries  $\frac{\partial \pi_{in,t}^k}{\partial \tau}$  is stronger for richer ones with greater demand in sectors that use this particular task intensively, the effect on task productivity growth is stronger. As a result, if trade integration leads a country to export to countries that consume more in sectors that use high spillover tasks intensively than its own consumption share, the exporting country achieves higher productivity growth under trade than under autarky while it also benefits from lower import prices.

## 4. Estimation Strategy

### 4.1 Productivity Parameters

This section outlines the procedure for estimating and calibrating the following productivity parameters: the dispersion parameters of the Frechet distributions of task productivity ( $\theta_a$ ) and sectoral TFP ( $\theta$ ), the adoption parameter ( $\beta$ ) that governs convergence of task productivity, the sectoral output elasticities of different tasks ( $\zeta_a^k$ ), and, ultimately, the diffusion parameters ( $\tilde{\eta}_a$ ) that govern the extent of dynamic scale economies for different tasks.

I calibrate the dispersion parameters externally using estimates from the literature. Specifically, I set the trade elasticity  $\theta$  to 4 (Simonovska and Waugh, 2014) and  $\theta_a$  to 1.13 (Burstein et al., 2015). Moreover, I calibrate  $\zeta_a^k$  to the sector- and occupation-specific wage share of tradable sectors in the United States in 1970.<sup>35</sup> This leaves the rest of this section to the estimation of  $\beta$  and  $\tilde{\eta}_a$ .

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<sup>35</sup>For details on occupational groups and micro-data used for wage sector- and occupation-specific wage shares, see section 5 and Appendix F



### 4.1.1 Comparative Advantage

Several international GE models -including the EK-style model outlined above- deliver a gravity equation of the form:

$$\ln \pi_{nm,t}^k = \delta_{m,t}^k + \mu_{n,t}^k - \theta \ln \tau_{nm,t}^k \quad (30)$$

where  $\pi_{nm,t}^k$  is the import share of country  $n$  from country  $m$  in sector  $k$ .  $\delta_{m,t}^k$  is an exporter-specific fixed effect in sector  $k$ ,  $\mu_{n,t}^k$  is an importer-specific fixed effect, and  $\theta \ln \tau_{nm,t}^k$  covers the sector-specific effect of trade costs  $\tau_{nm,t}^k$  between the two countries that affect the trade share with elasticity  $\theta$ .

If  $\tau_{ni,t}^k = \tau_{in,t}^k \forall n \neq i$  and  $\tau_{nn',t}^k = 1 \forall n = n'$  (Head and Ries, 2001), one can recover trade costs using<sup>36</sup>

$$\tau_{ni,t}^k = \left[ \frac{\pi_{ni,t}^k}{\pi_{nn,t}^k} \frac{\pi_{in,t}^k}{\pi_{ii,t}^k} \right]^{-\frac{1}{2\theta}} \quad (31)$$

In an Eaton-Kortum type model, the exporter- and importer-specific fixed effects take the form

$$\delta_{m,t}^k = -\theta \ln c_{m,t}^k / \bar{c}_t^k + \ln T_{n,t}^k / \bar{T}_t^k \quad (32)$$

$$\mu_{n,t}^k = \ln \Phi_{n,t}^k / \bar{\Phi}_{n,t}^k \quad (33)$$

where a bar over a variable refers to its global (unweighted) mean. The exporter-specific fixed effect thus captures *effective* unit costs of country  $m$  in sector  $k$ , which reflect the country's comparative advantage in that sector.<sup>37</sup> The importer-specific fixed effect captures the 'multilateral resistance' in sector  $k$  of country  $n$  such that  $\Phi_{n,t}^k = (\sum_{i=1}^N T_{i,t}^k (c_{i,t}^k \tau_{ni,t}^k)^{-\theta})^{-1}$ . Hanson et al. (2015), following Eaton et al. (2012), recast this gravity equation to allow for zero trade flows by assuming that in each industry-country pair only a finite number

<sup>36</sup>I choose to specify trade costs symmetrically to be consistent with the counterfactual exercise in section 7.1. An alternative would be to specify trade costs as a sector- and year-specific log-linear function of distance variables (e.g. Levchenko and Zhang, 2016; Bartelme et al., 2019b). Using this specification, estimated comparative advantage terms or spillover parameters are not qualitatively different.

<sup>37</sup>Note that in a world without trade costs ( $\tau_{ni,t}^k = 1 \forall n, i \in N$ ), the exporter fixed effects are a measure of revealed comparative advantage (RCA), i.e. for any two countries  $n$  and  $m$  and any two sectors  $k$  and  $k'$ ,  $\ln \frac{RCA_{n,t}^k / RCA_{n,t}^{k'}}{RCA_{m,t}^k / RCA_{m,t}^{k'}} = \frac{\delta_{n,t}^k / \delta_{n,t}^{k'}}{\delta_{m,t}^k / \delta_{m,t}^{k'}}$ . See also Appendix section B.

of firms make productivity draws. As a result, equation 30 holds only in expectation:

$$\mathbb{E}[\pi_{nm,t}^k] = \frac{\exp[\delta_{m,t}^k - \theta \ln \tau_{nm,t}^k]}{\exp[\mu_{n,t}^k]} \quad (34)$$

Combining the gravity and unit cost equations, and expressing them relative to the global mean yields an estimating equation

$$\ln \mathbb{E}[\pi_{nm,t}^k] = \underbrace{\ln T_{m,t}^k / \bar{T}_t^k - \theta \ln c_{m,t}^k / \bar{c}_t^k}_{\text{exporter fixed effect } \delta_{m,t}^k} + \underbrace{\Phi_{n,t}^k / \bar{\Phi}_{n,t}^k}_{\text{importer fixed effect } \mu_{n,t}^k} - \frac{1}{2} \ln \frac{\pi_{ni,t}^k}{\pi_{nn,t}^k} \frac{\pi_{in,t}^k}{\pi_{ii,t}^k} + \nu_{nm,t}^k \quad (35)$$

where  $\nu_{nm,t}^k$  is a mean-zero misspecification term. I estimate equation 35 separately for each sector and year under the constraint that the exporter and importer fixed effects each sum up to zero. I use both log-linear OLS and PPML (Silva and Tenreyro, 2006).<sup>38</sup>

#### 4.1.2 Spillovers

Changes in a country's sector-specific *effective* unit costs  $\tilde{c}_{n,t}^k$  can be expressed as

$$\Delta \ln \tilde{c}_{n,t}^k = \Delta \ln c_{n,t}^k - \Delta \ln T_{n,t}^k \quad (36)$$

Using the definition of unit costs and arbitrage in the labor market gives

$$\Delta \ln \tilde{c}_{n,t}^k = \Delta \ln w_{n,t} - \sum_{a=1}^A \zeta_a^k \Delta \ln T_{n,t}^a - \Delta \ln T_{n,t}^k \quad (37)$$

In the endogenous growth model outlined above, we can express the growth of the task productivity level  $T_{n,t}^a$  in log changes as

$$\Delta \ln T_{n,t}^a = \beta_0 + \frac{1}{\theta_a} \tilde{\eta}_a \ln L_{n,t-1}^a + (\beta - 1) \ln T_{n,t}^a + \epsilon_{n,t}^a \quad (38)$$

where  $\beta_0$  is a constant,  $\eta_{n,t}^a$  is a diffusion parameter that differs by task, country and time.  $\epsilon_{n,t}^a$  is a mean-zero shock specific to task  $a$  in country  $n$  at time  $t$ . Plugging this expression into equation 37 on the right-hand-side, and the definition of the exporter

<sup>38</sup>The correlation between the estimated fixed effects of these two methods is always higher than 0.99. In the rest of the paper I use the OLS estimates whenever referring to results or computations based on these fixed effects

fixed effect on the left-hand-side, while expressing all relative to the global unweighted mean, yields an estimating equation:

$$\frac{1}{\theta} \Delta \delta_{n,t}^k = \underbrace{-\beta_0}_{\text{constant}} - \frac{1}{\theta_a} \sum_{a=1}^A \tilde{\eta}_a \zeta_a^k \ln L_{n,t-1}^a / L_{n,t-1} + (\beta - 1) \frac{1}{\theta} \delta_{n,t-1}^k + \gamma_{n,t} + \tilde{\epsilon}_{n,t}^k \quad (39)$$

where  $\gamma_{n,t}$  is a fixed effect that captures  $\Delta \ln w_{n,t} + (1 - \beta) \ln w_{n,t-1}$ , and  $\tilde{\epsilon}_{n,t}^k$  is an error term capturing  $\sum_{a=1}^A \zeta_a^k \epsilon_{n,t}^a - \Delta \ln T_{n,t}^k + (1 - \beta) \ln T_{n,t-1}^k$ .<sup>39</sup> The only unknown parameters of interest are  $\tilde{\eta}_a$ , and  $\beta$ .  $\frac{1}{\theta} \Delta \delta_{n,t}^k$  and  $\frac{1}{\theta} \delta_{n,t-1}^k$  can be constructed using a value of  $\theta$  and gravity estimates. I first estimate equation 39 for 10 year periods using OLS.

This naive estimation method raises endogeneity issues, however. A potential concern could be that producers in fast-growing sectors preemptively increase production in period  $t - 1$ , anticipating productivity growth between  $t$  and  $t - 1$ . Such anticipatory behavior would generate a correlation between  $\zeta_a^k L_{n,t-1}^a$  and  $\zeta_a^k \epsilon_{n,t}^a$  or between  $\zeta_a^k L_{n,t-1}^a$  and  $\Delta \ln T_{n,t}^k$ . To circumvent these kind of supply side concerns, I rely on foreign demand shocks that generate variation in the task employment share  $L_{n,t}^a / L_{n,t}$ .

Note that the task employment share has a model-implied equivalent in the weighted average of exports, i.e.

$$L_{n,t}^a / L_{n,t} = \frac{\sum_{k=1}^K \zeta_a^k \sum_{n'=1}^N X_{n'n,t}^k}{\sum_{k=1}^K \sum_{n'=1}^N X_{n'n,t}^k} \quad (40)$$

Using the gravity equation in section 9, exports of country  $n$  in sector  $k$  can be expressed as (relative to the global mean):

$$\sum_{n' \neq n}^N X_{n'n,t}^k = \delta_{n,t}^k \cdot \underbrace{\sum_{n' \neq n}^N E_{n',t}^k \mu_{n',t}^k (\tau_{n'n,t}^k)^{-\theta}}_{FMA_{n,t}^k} \quad (41)$$

where  $\delta_{n,t}^k$  captures effective unit costs relative to the global mean and  $E_{n',t}^k$  is total expenditure of country  $n'$  in sector  $k$ . While the exporter fixed effect reflects the effect of supply side factors in country  $n$  on its exports, the second term captures foreign demand. I will refer to this sum of foreign demand factors as Foreign Market Access (FMA) (Bartelme et al., 2019b).

<sup>39</sup>Note that under the assumption that both  $T_{n,t}^k$  and  $\epsilon_{n,t}^a$  are orthogonal to  $\zeta_a^k$  the regression coefficient on  $\zeta_a^k \ln \frac{L_{n,t-1}^a}{L_{n,t-1}}$  yields consistent estimates of  $\frac{1}{\theta_a} \tilde{\eta}_a$  using OLS.

I can now construct a synthetic measure of task employment share that is only driven by demand shocks. Removing supply side variation by setting effective unit costs to the global mean:

$$(L_{n,t}^a/L_{n,t})^{FMA} = \frac{\sum_{k=1}^K \zeta_a^k \cdot FMA_{n,t}^k}{\sum_{k=1}^K FMA_{n,t}^k} \quad (42)$$

I also estimate equation 39 using this demand shock driven measure of task employment. Finally, I run IV 2SLS using  $(L_{n,t}^a/L_{n,t})^{FMA}$  as instruments for  $L_{n,t}^a/L_{n,t}$ .

## 4.2 Demand Parameters

On the demand side, I need to estimate preference parameters  $\alpha_{n,t}^k$  in order to perform counterfactuals and construct measures of welfare. I do this by calibrating  $\alpha_{n,t}^k$  as a country's expenditure share in sector  $k$  at time  $t$ , i.e.

$$\alpha_{n,t}^k = \frac{\sum_{i=1}^N X_{ni,t}^k}{\sum_{k=1}^K \sum_{i=1}^N X_{ni,t}^k} \quad (43)$$

## 5. Data

This section briefly describes data sources and implementation strategy. See Appendix F for details on steps taken.

### 5.1 Occupations

A challenge for the estimation of the parameters governing task productivity growth is that there are potentially hundreds of occupations and thus potentially hundreds of parameters that can be estimated using only limited variation in the trade data. To reduce the number of estimable parameters, I assign occupations into four groups based on the task content of their work.<sup>40</sup>

<sup>40</sup>Although there are some standard classifications that assign occupations to broad groups, these tend to be based on the sectors in which occupations are mainly active, and not the occupations task content. As a result, these classifications do a poor job capturing occupational linkages between sectors. The International Standard Classification Occupations ISCO-08, for example, assigns all agricultural workers to one of eight groups, and most service and sales workers to another. The BLS Standard Occupational Classification (SOC) assigns almost all occupations to its major groups based on sector, for example, Healthcare Support Occupations and Farming, Fishing, and Forestry Occupations. Moreover, SOC's highest level of aggregation includes 23 major groups, without an obvious method of clustering these into

I use data on tasks ('Work Activities') in O\*NET to assign each occupation in the BLS Standard Occupational Classification (SOC) to an occupational group, which I call task groups. First, I reduce the task space from 41 to 5 tasks by assigning each task to one of five groups (Table A1): (i) information-intensive tasks (e.g. "Analyzing Data or Information"), (ii) planning-intensive tasks (e.g. "Performing Administrative Activities"), (iii) equipment-intensive tasks (e.g. "Inspecting Equipment, Structures or Material"), (iv) mechanical tasks (e.g. "Performing General Physical Activities"), and (v) contact tasks (e.g. "Establishing and Maintaining Interpersonal Relationships").

Next, I normalize each occupations' score for a task ('task score') and rank all occupations according to each task score. Finally, I assign an occupation to the task for which it has the highest rank. Table A2 shows the five highest ranked occupations and sectors (by share of total wages) for each of the information, planning, equipment and mechanical tasks. In the subsequent empirical analysis, I focus on these four non-contact tasks.<sup>41</sup>

## 5.2 Productivity Parameters

Estimation of comparative advantage, and, ultimately, spillover parameters requires data on (i) bilateral sectoral trade flows  $X_{ni,t}^k$ , and (ii) task-level sectoral wage shares  $\zeta_a^k$ , (iii) as well as estimates of task employment shares  $L_{n,t}^a/L_{n,t}$ .

Data on bilateral sectoral trade flows are from the World Trade Flows (WTF) database developed by Feenstra et al. (2005). These cover UN COMTRADE bilateral trade between country pairs at the disaggregated four digit SITC2 level for the years 1962-2000. I aggregate goods to the level of industries from the 1950 Census Bureau industrial classification system to be able to calibrate task-level sectoral wage shares. Table A3 in the Appendix contains the corresponding list of sectors. The rest of my procedure follows Hanson et al. (2015), leaving me with 87 countries and 61 tradable sectors. See Appendix F for details.

Micro-data on sector- and occupation-specific wages are from the IPUMS USA 1970 Census. I consider only wage income and consider both part-time and full-time employment. I then calculate the wage share as total wages paid to an occupational group smaller groups.

<sup>41</sup>An alternative method for classifying would be to assign shares of working time of each occupation to tasks. Unfortunately, O\*NET contains only ordinal data on the relative importance of tasks for a given occupation.

in a sector as a share of total wages paid in that same sector.

I consider two measures of task employment shares  $L_{n,t}^a/L_{n,t}$ . As I lack domestic expenditure data at a detailed sector level, I use only export data to construct shares. First, I construct a share based on reported exports, i.e.

$$L_{n,t}^a/L_{n,t} = \frac{\sum_{k=1}^K \zeta_a^k \sum_{n' \neq n}^N X_{n'n,t}^k}{\sum_{k=1}^K \zeta_a^k \sum_{n' \neq n}^N X_{n'n,t}^k} \quad (44)$$

Second, I construct shares based on estimated foreign market access (equation 42).

## 6. Results

### 6.1 Productivity

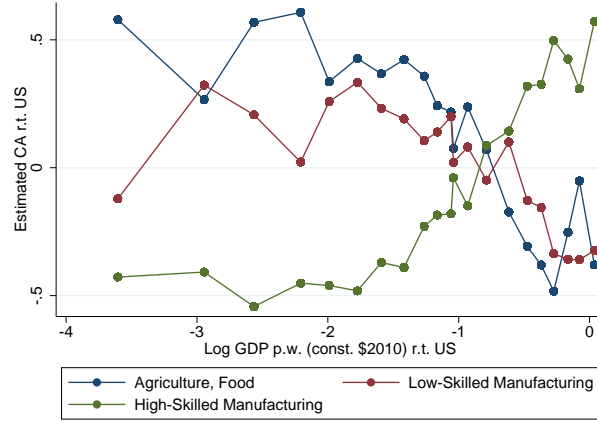
#### 6.1.1 Effective Unit Costs

Figure 3 shows a binned scatterplot with estimates of countries' comparative advantage (or inverse effective unit costs) in the three tradable clusters plotted against GDP per worker, where both are expressed relative to the United States. I present these inverse effective unit cost terms as a weighted average (by exports) of the exporter fixed effects in equation 35. In line with the evidence presented in section 2, poorer countries tend to have lower effective unit costs in agriculture, which increases relative to the U.S. as they become more productive. Comparative advantage in low-skilled manufacturing and mining is hump-shaped in GDP per worker, increasing for poorer countries but declining for advanced economies. In contrast, effective unit costs in high-skilled manufacturing are relatively flat for the poorest economies in the sample but decrease steeply as countries catch up to the frontier.

#### 6.1.2 Spillovers

Table 2 presents the estimation results for spillovers from estimating equation 39. Column (1) and (2) report estimates using exports- and FMA-based task employment shares, and column (3) reports estimates of an IV 2SLS procedure with the FMA-based shares as an instruments for the exports-based shares.

Figure 3: Estimated Comparative Advantage (Inverse Effective Unit Costs) and GDP per worker r.t. U.S.



**Notes:** This figure presents a binned scatterplot of countries' estimated comparative for the three tradable clusters against log GDP p.w. for the World Trade Flows sample from 1970 to 2000. These inverse effective unit cost terms are a country's weighted average (by exports) of exporter fixed effects in a cluster as estimated in equation 35. Real GDP per worker (constant \$2010) data are from Penn World Tables 9.0.

There is considerable and significant heterogeneity in the coefficients of spillover parameters  $\tilde{\eta}_a$  of different tasks. For all specifications, the point estimates of the spillover parameter is highest for planning tasks (1.2 to 5.1) and also large for information tasks (0.5 to 0.7). These contrast with estimated spillovers for equipment (0.2 to 0.6) and mechanical tasks (0.1 to 0.3). Hence overall, estimated spillovers are higher for tasks generally performed by high-skilled labor than for those performed more often by low-skilled labor.

Another apparent feature of the results is the robust and substantially negative estimate of the convergence parameter,  $\beta - 1$ . It is relatively precisely estimated around -0.45, indicating that, over a 10 year period, comparative advantage at one tenth of the global frontier catches up to that same frontier by about 11 percent per year. Through the lens of the model in section 3.2, an 0.55 estimate of  $\beta$  suggests a roughly equal importance of learning-by-doing forces and human capital spillovers. It is only slightly lower than the calibrated estimate of 0.6 in Buera and Oberfield (2020).

In the rest of the paper, I will use the estimates in column (2) of Table 2 to perform quantitative exercises. A first implication of the heterogeneity in spillovers concerns how they aggregate to the sector level. Producing in sectors that use high-skill planning



Table 2: Estimated Spillovers

	(1)	(2)	(3)
	Exports	Foreign demand shocks	IV 2SLS
$\tilde{\eta}_{planning}$ (high-skill)	1.202*** (0.181)	1.842*** (0.466)	5.066 *** (0.448)
$\tilde{\eta}_{information}$ (high-skill)	0.731*** (0.0431)	0.724*** (0.121)	0.528*** (0.0642)
$\tilde{\eta}_{equipment}$ (medium-skill)	0.238*** (0.0337)	0.629** (0.110)	0.118** (0.0385)
$\tilde{\eta}_{mechanical}$ (low-skill)	0.142*** (0.0194)	0.164** (0.0491)	0.350** (0.0263)
$-\phi$ (convergence)	-0.467*** (0.00903)	-0.449*** (0.00893)	-0.476*** (0.00611)
Country-year FE	Y	Y	Y
Observations	91626	91626	91626
Adjusted $R^2$	0.464	0.455	

Standard errors in parentheses, clustered at country-year level.

2SLS uses 1000 bootstrap replications to compute standard errors.

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

**Notes:** This table reports estimates of different specifications of estimating equation 39. All specifications include country-year fixed effects. Column (1) reports estimates using task employment shares based on export data (equation 44). Column (2) reports estimates using task employment shares based on foreign market access estimates (equation 42). Column (3) reports estimates for an IV 2SLS procedure with employment shares from Column (2) as instruments for those in Column (1).

and information tasks intensively, has a larger positive effect on aggregate productivity because spillovers are higher in these high-skill tasks. Table A5 in the Appendix reports the relative effect of allocating labor to sectors on aggregate productivity. In general, inter-industry spillovers are highest in technologically advanced sectors (e.g. electrical and optical equipment, chemicals) at intermediate levels for low-skilled manufacturing sectors (e.g. textiles, furniture, toys) and mining, and lowest in agriculture. See section G.1 in the Appendix for details.

## 6.2 Model Fit

### 6.2.1 Sectoral Productivity Growth

I will now assess the extent to which the model can account for measured changes in sectoral unit costs. Using the gravity estimates from 35, we can express the change in log unit costs from 1970 to 2000 relative to the (unweighed) global mean as:

$$\sum_{t=1980}^{2000} \Delta \ln \frac{c_{n,t}^k}{\bar{c}_t^k} = \sum_{t=1971}^{2000} \left[ \Delta \ln \frac{w_{n,t}}{\bar{w}_t} - \Delta \ln \frac{T_{n,t}^k}{\bar{T}_{n,t}^k} - \frac{1}{\theta_a} \sum_{a=1}^A (\zeta_a^k \tilde{\eta}_a \ln \frac{L_{n,t-1}^a}{\bar{L}_{n,t-1}^a} + \epsilon_{n,t}^a - \bar{\epsilon}_{n,t}^a) + (\beta - 1) \ln \frac{c_{n,t-1}^k}{\bar{c}_{n,t-1}^k} \right] \quad (45)$$

where bars refer to variables expressed in terms of the global mean. Out of the five sources of relative changes in unit costs (wages, sectoral TFP, spillovers, convergence, and residual country-specific growth) I can construct an estimate of the contribution of spillovers and convergence using estimates of  $\tilde{\eta}_a$ ,  $\beta$ ,  $L_{n,t-1}^a$  and  $c_{n,t}^k$ . Formally, I construct 'sector-level' accumulated spillovers as:

$$\sum_{t=1980}^{2000} \left[ \frac{1}{\theta_a} \sum_{a=1}^A (\zeta_a^k \tilde{\eta}_a \ln \frac{L_{n,t-1}^a}{\bar{L}_{n,t-1}^a} + (\beta - 1) \ln \frac{c_{n,t-1}^k}{\bar{c}_{n,t-1}^k} \right]$$

Figure 4 shows a binned scatterplot of these accumulated spillovers against estimated relative changes in effective unit costs,  $\sum_{t=1971}^{2000} \Delta \ln \frac{c_{n,t}^k}{\bar{c}_t^k}$ . The model performs well in this regard. Indeed, on average the model predicts about 23 % of the variation in measured relative changes in unit costs for the entire sample. Quantitatively, the model predicted accumulated spillovers are on average somewhat more extreme.

### 6.2.2 Aggregate Productivity Growth

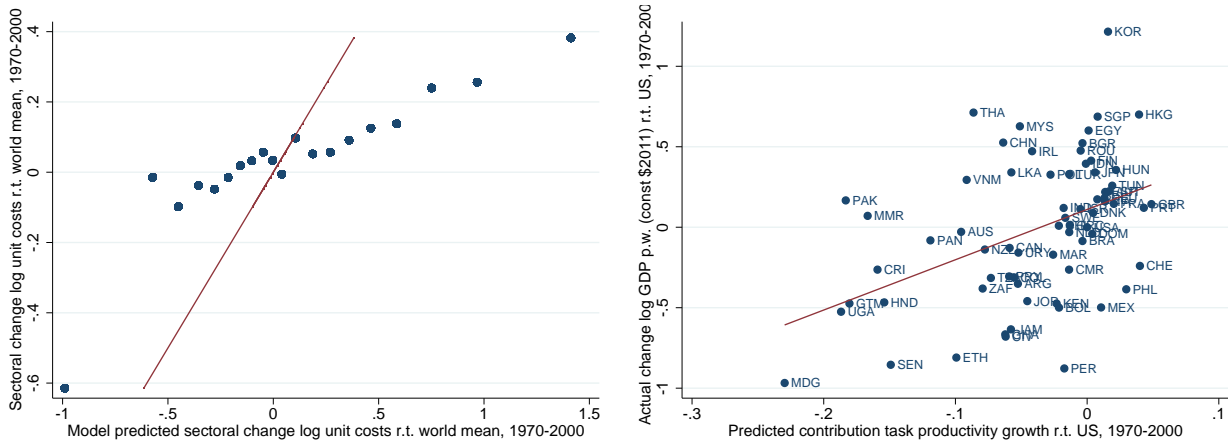
Aggregate labor productivity  $y_{n,t}$  is the weighted average of value added labor productivity  $y_{n,t}^k$  in individual sectors:

$$y_{n,t} = \sum_{k=1}^K L_{n,t}^k y_{n,t}^k \quad (46)$$

where  $L_{n,t}^k$  is a sector's employment share. Up to a first order, aggregate labor productivity growth can be expressed as (Rodrik, 2011):

$$\Delta \ln y_{n,t} \approx \sum_{k=1}^K \theta_{n,t-1}^k \sum_{a=1}^A \zeta_a^k \Delta \ln T_{n,t}^a + \sum_{k=1}^K \theta_{n,t-1}^k [\Delta \ln P_{n,t}^k + \Delta \ln T_{n,t}^k] + \sum_{k=1}^K \frac{y_{n,t-1}^k}{y_{n,t-1}} \Delta L_{n,t}^k \quad (47)$$

Figure 4: Model Predicted Accumulated Spillovers and Actual Sectoral and Aggregate Productivity Growth



**Notes:** The left figure plots changes in estimated sector-level effective unit costs between 1970 and 2000 (equation 35) against sector-level accumulated spillovers  $\sum_{t=1980}^{2000} \left[ \frac{1}{\theta_a} \sum_{a=1}^A (\zeta_a^k \tilde{\eta}_a \ln \frac{L_{n,t-1}^a}{L_{n,t-1}^a} + (\beta-1) \ln \frac{c_{n,t-1}^k}{c_{n,t-1}^k} \right]$  predicted by the model. Both are expressed relative to the (unweighted) global mean in each sector. The right figure plots changes in real GDP p.w. (national prices) between 1970 and 2000 against aggregate accumulated spillovers  $\sum_{t=1980}^{2000} \theta_{US,t-1}^k \sum_{a=1}^A \left[ \frac{1}{\theta_a} \sum_{a=1}^A (\zeta_a^k \tilde{\eta}_a \ln \frac{L_{n,t-1}^a}{L_{US,t-1}^a} \right]$  predicted by the model.

where  $\theta_{n,t-1}^k$  is the value added share of sector  $k$  in period  $t - 1$ . Plugging in the evolution task productivity growth and ignoring the second reallocation term:

$$\Delta \ln y_{n,t} \approx \sum_{k=1}^K \theta_{n,t-1}^k \sum_{a=1}^A \zeta_a^k \tilde{\eta}_a \ln L_{n,t-1}^a + \sum_{k=1}^K \theta_{n,t-1}^k [\Delta \ln P_{n,t}^k + \Delta \ln \tilde{T}_{n,t}^k + \sum_{a=1}^A \zeta_a^k [\beta_0 + \phi \ln T_{n,t-1}^a + \epsilon_{n,t}^a]] \quad (48)$$

Similar to the sector-level case, I can now construct 'aggregate' accumulated spillovers. Aggregating sectors using United States value added shares and expressing task productivity growth relative to the U.S. gives a measure of these spillovers:<sup>42</sup>

$$\sum_{t=1980}^{2000} \theta_{US,t-1}^k \sum_{a=1}^A \left[ \frac{1}{\theta_a} \sum_{a=1}^A (\zeta_a^k \tilde{\eta}_a \ln \frac{L_{n,t-1}^a}{L_{US,t-1}^a}) \right]$$

Figure 4 plots these accumulated spillovers against changes in real GDP per worker (at national prices) between 1970 and 2000. Again, the model performs well in this regard. On average the model predicts about 22 % of the variation in real GDP per worker. Quantitatively, however, the model predicted contribution of spillovers to aggregate productivity growth is smaller than the data.<sup>43</sup>

## 7. Quantitative Implications

In this section, I use the quantitative framework to study two things. First, I explore the role of inter-industry spillovers in accounting for cross-country convergence by shutting down the mechanism in the model. Second, I assess the dynamic gains from trade through the lens of the model. In these exercises, I use the version of the model outlined in section 3.

<sup>42</sup>In order to compare aggregate accumulated spillovers across countries, it is necessary to use a common aggregator across sectors, in this case the value added share in the U.S. Doing so will depress the explanatory power of the model, since this aggregation captures only differences in aggregate productivity growth driven by differences in spillovers *within* sectors. It does not capture cross-country differences in sectoral value added shares interacted with high or low average spillovers. I therefore interpret the estimated explanatory power of the model as a lower bound.

<sup>43</sup>This is not surprising given that model-predicted spillovers are likely to be correlated with other factors that positive affect convergence to the frontier. Other explanations for the 20th century convergence successes (e.g. Japan, Korea, China) put forward in the literature are learning-by-exporting (e.g. Studwell, 2013; Buera and Oberfield, 2020), undervalued exchange rates (Rodrik, 2008), land reforms (e.g. Paldam, 2003), human capital endowments (e.g. Ciccone and Papaioannou, 2009), as well as higher saving rates and industrial policy (e.g. Stiglitz, 1996).

## 7.1 Inter-Industry Spillovers and Economic Convergence

What is the role of inter-industry spillovers in accounting for cross-country convergence in aggregate productivity? I explore this counterfactual by setting  $\tilde{\eta}_a = 0$  for all tasks, shutting down the effect of dynamic scale economies on productivity growth. Between two periods  $t = \tau$  and  $t = \tau - 1$ , counterfactual task productivity levels  $\{\{\hat{T}_{n,\tau}^a\}_{a=1}^A\}_{t=1}^T$  now satisfy:

$$\hat{T}_{n,\tau}^a = [L_{n,\tau-1}^a]^{-\tilde{\eta}_a} (\hat{T}_{n,\tau-1}^a)^{\beta-1} \quad (49)$$

Countries that allocate labor to high-skill intensive sectors thus experience a relative reduction in their productivity growth as spillovers tend to be higher in those sectors. I simulate counterfactual changes for the World Trade Flows panel for the period 1970-2000, with the counterfactual dynamic equilibrium defined in section 3.2.2.

The counterfactual change in welfare  $U_{n,t}$  (real GDP at PPP) per worker can then be expressed in two ways. First, if nominal incomes are deflated using international prices, it is simply the weighted geometric average of changes in sector-specific prices:

$$\Delta \ln \hat{U}_{n,t} = \ln \hat{w}_{n,t} - \sum_{k=1}^K \alpha_{n,t}^k \ln \hat{P}_{n,t}^k \quad (50)$$

Second, if real GDP is measured using prices of a domestically produced bundle:

$$\Delta \ln \hat{U}_{n,t} = \ln \hat{w}_{n,t} + \sum_{k=1}^K \alpha_{n,t}^k \sum_{a=1}^A \zeta_a^k \ln \hat{T}_{n,t}^a \quad (51)$$

I show the importance of spillovers for counterfactual convergence using both measures. While the first is a more accurate measure of welfare, the second is a better reflection of changes in productivity.

I perform an unconditional (beta) convergence regression for 1970-2000 of counterfactual real GDP per worker growth on real GDP per worker in 1970. Table 3 summarizes these regressions. After accounting for the effect of spillovers, the convergence coefficient for unconditional convergence turns robustly negative. While there is no unconditional cross-country convergence in aggregate productivity (column 3),<sup>44</sup> counterfactual changes in aggregate productivity are significantly higher for countries that were

<sup>44</sup>See also Rodrik (2012).

initially poorer in 1970. With an international price deflator (column 1), a country at one tenth of the frontier in 1970 grows 0.39 percentage points per year faster in a counterfactual without spillovers. Using prices of a domestically produced bundle, this difference increases to 1.23 percentage points per year. Putting these numbers into perspective, in the same sample countries experienced *divergence* in real GDP per worker over the 1970-2000 period, with a typical country at one tenth of the frontier growing 0.43 percentage points slower than the frontier.

Figures 5(a)-(b) highlights these effects of spillovers on cross-country convergence. Shutting down spillovers increases real GDP growth relative to the frontier (the United States) for poorer countries in Sub-Saharan Africa (e.g. Ethiopia, Tanzania, Madagascar, Uganda) but lowers growth for some Eastern European and East Asian fast-growing countries (e.g. Romania, Korea, Hong Kong SAR). At the same time, in the counterfactuals growth is slower in Western European countries (e.g. France, Italy, Denmark, Sweden) that catch up to the United States during this period. These countries export relatively more in skill intensive sectors. All in all, relatively faster growth for initially poorer countries turns divergence in the data into convergence in the counterfactual. Convergence is stronger when deflating GDP using domestic prices, which are a more direct reflection of endogenous changes in productivity.

## 7.2 Dynamic Gains from Trade and Initial Comparative Advantage

How do inter-industry spillovers affect the dynamic gains from trade? What is the role of initial comparative advantage in mediating these dynamic gains? I examine these questions by exploring a series of counterfactuals in which I keep one country's (symmetric) trade costs remain at the 1970 level, and construct the ensuing dynamic counterfactual equilibrium from 1970 to 2000. This equilibrium is summarized in section 3.2.2. Remember that the gains from trade in 2000 are summarized by

$$\ln \hat{w}_{n,t} / \hat{P}_{n,t} = - \sum_{k=1}^K \frac{\alpha_{n,t}^k}{\theta} \ln \hat{\pi}_{nn,t}^k + \sum_{k=1}^K \alpha_{n,t}^k \sum_{a=1}^A \zeta_a^k \ln \hat{T}_{n,t}^a \quad (52)$$

where the first term covers the standard static gains from trade (Arkolakis et al., 2012), and the second term covers the dynamic gains from trade.

Table 3: Counterfactual Convergence of Real GDP per worker (1970-2000)

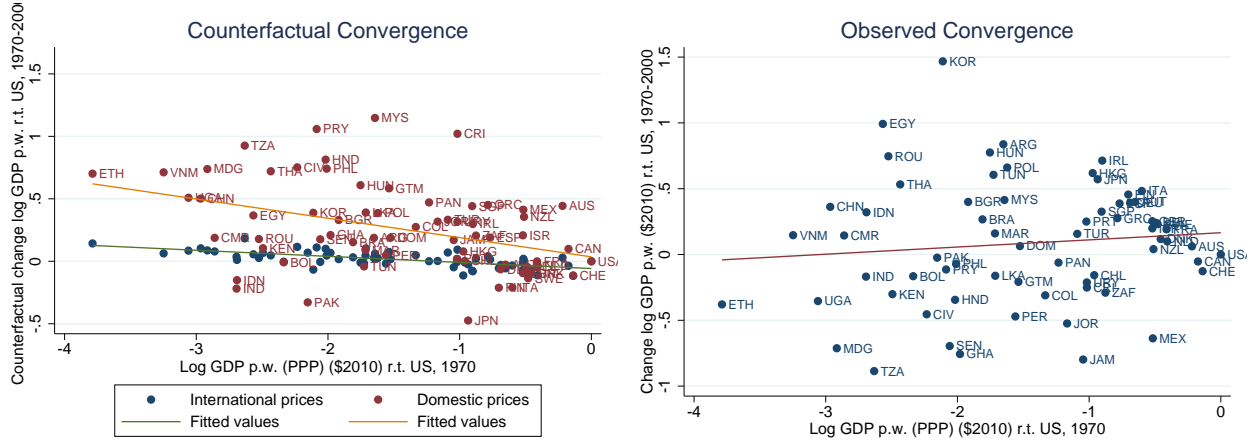
	(1)	(2)	(3)
	International prices	Domestic prices	Actual $\ln y_{n,2000} - \ln y_{n,1970}$
$\ln y_{n,1970}$	-0.0488*** (0.00619)	-0.154*** (0.0390)	0.0544 (0.0573)
Constant	-0.0578*** (0.0113)	0.0356 (0.0600)	0.166** (0.0800)
Observations	67	67	67
Adjusted $R^2$	0.451	0.145	-0.004

Robust standard errors in parentheses

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ 

**Notes:** This table reports convergence regressions for counterfactual changes in real GDP per worker between 1970 and 2000 on real GDP per worker (PPP) in 1970. Column (1) uses the counterfactual change in welfare based on international prices (equation 50) as a measure of changes in GDP per worker whereas column (2) uses the change in welfare based on prices of a domestically produced bundle of goods (equation 51). Column (3) reports a convergence regression with actual changes real GDP per worker as the dependent variable. In all cases, real GDP per worker data (PPP at constant \$2010) are from Penn World Tables 9.0. The sample includes all countries in the World Trade Flow sample but excludes oil exporters (Ecuador, Norway, Algeria, Venezuela, Iran, Kuwait, Libya, Nigeria, Oman, Saudi Arabia, Syria and Trinidad & Tobago).

Figure 5: Counterfactual and Observed Convergence of real GDP per worker (PPP)



**Notes:** The left figure presents a scatterplot with associated linear fit of counterfactual changes in welfare between 1970 and 2000 against actual real GDP per worker in 1970. Blue dots and green fit line refer to counterfactual welfare changes using an international price bundle (equation 50). Red dots and orange fit line refer to counterfactual welfare changes using a price bundle of domestically produced goods (equation 51). All variables are expressed relative to the United States. For both figures, real GDP per worker (PPP, constant \$2010) are from Penn World Tables 9.0.

Figure 6(a) plots these two types of gains against each other for the given sample of countries. Countries with larger static gains (mainly fast-growing developing countries) tend to experience slightly larger dynamic gains, although this relationship is weak. In most countries, dynamic gains of trade are substantial and equal 5 percent on average, which is roughly 1/3 of the average static gains. There exists considerable heterogeneity across countries in terms of the dynamic gains from trade. Several countries have negative estimated dynamic gains, indicating that they shift labor into sectors with lower spillovers when trade barriers decrease during the period studied.

To what extent are these differences explained by initial underlying comparative advantage? Figure 6(b)-(d) plots dynamic gains against countries' revealed comparative advantage in 1970, for each of the three clusters defined in section 2. Countries with a comparative advantage in agriculture, such as Argentina, United States, and Brazil, tend to have lower dynamic gains, which is not surprising given that estimated spillovers are low in this cluster. At the same time, estimated gains are generally higher in countries with a comparative advantage in low-skilled manufacturing, while they are roughly flat



with respect to comparative advantage in high-skilled manufacturing.

There are at least two reasons why countries with initial comparative advantage in low-skilled manufacturing have higher estimated dynamic gains than those with a comparative advantage in high-skilled manufacturing. First, the reduction in trade barriers since the 1970s has been smaller in advanced economies that tend to specialize in high skill intensive exports relative to developing economies. Second, due to non-homothetic preferences, trade barriers tend to restrict the reallocation of labor towards manufacturing in developing economies (Tombe, 2015), while richer countries tend to consume and export skill intensive products (Caron et al., 2014). Both of these mechanisms amplify the trade-induced reallocation of labor towards sectors with higher spillovers more in developing countries with an initial comparative advantage in low-skilled manufacturing.

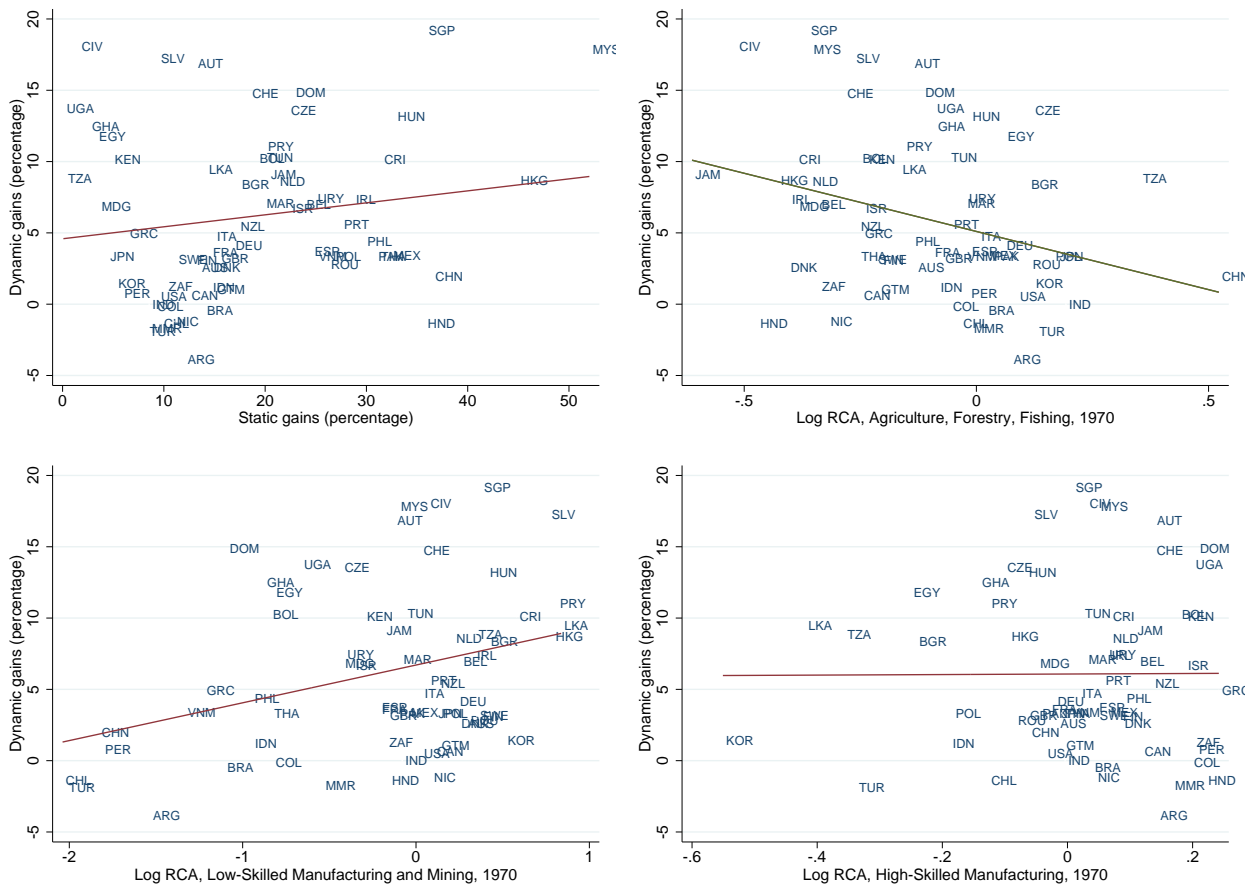
Finally, in the Appendix, I show that while incorporating I-O linkages into the model increases the *static* gains from trade (Caliendo and Parro, 2015), this does not change the *dynamic* gains from trade. Under both models, the dynamic gains from trade equal roughly 1/3 of the static gains of the standard model, which corresponds to the exercises summarized above.

## 8. Conclusions

Developing economies that catch up to the global economic frontier tend to produce more skill-intensive products as they grow. This paper aims to quantify the role of inter-industry productivity spillovers in this catch-up process. Through the lens of a general equilibrium, multi-sector trade model featuring occupation-specific dynamic scale economies, heterogeneous inter-industry spillovers are important for understanding two stylized facts in the growth literature.

First, the model implies that heterogeneous spillovers can account for the lack of cross-country convergence in aggregate productivity (Rodrik, 2012) if dynamic scale economies are stronger for high-skilled occupation, leading to stronger productivity spillovers for countries that produce relatively more in high-skilled intensive sectors. Indeed, the estimates in this paper suggest that dynamic scale economies are substantial in high-skill intensive production but negligible in low-skill intensive production.

Figure 6: Dynamic Gains from Trade and Initial Comparative Advantage



**Notes:** All four figures plot a country's counterfactual dynamic gains from trade (second term in equation 52) for each country, in a counterfactual dynamic equilibrium (section 3.2.2) in which the country's inferred trade costs remain at their 1970 levels. The upper left figure plots countries' dynamic gains against the static gains (first term in equation 52). The other three figures plot countries' dynamic gains against the revealed comparative advantage in the three clusters defined in section 2: agriculture (upper left), low-skilled manufacturing and mining (bottom left), and high-skilled manufacturing (bottom right).

As a consequence, countries farther away from the global economic frontier would experience faster catch-up growth in the absence of inter-industry productivity spillovers.

Second, the model suggest heterogeneous inter-industry spillovers are important for understanding why some countries (mostly East Asian and Eastern European) have been able to catch up to the frontier in the last five decades. Through the lens of the model, exporting relatively more complex, skill-intensive products has a positive effect on subsequent economic growth ([Hausmann et al., 2007](#); [Hidalgo and Hausmann, 2009](#)) because it leads to higher future productivity in sectors in which a country is initially noncompetitive. In this sense, the findings in this paper mirror those of [Hanson \(2017\)](#), who documents that labor-abundant East Asian countries with an initial comparative advantage in low-skilled manufacturing tend to climb a ladder of complex industries as they become more productive, whereas these patterns are not present in resource abundant Latin American countries.

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# Catch-Up Growth and Inter-Industry Productivity Spillovers

## Online Appendix

Marijn Bolhuis

### A. Motivating Facts: Additional Material

This section presents additional material on Fact 1 in section 2 in the form of three different sets of figures. The first set of figures plots a binned scatterplot countries' share of tradable employment in a given sector against its (log) GDP p.w. (const. \$2010) for all countries in the IPUMS International sample over the period 1970-2012. The second set of figures plots the United States share of tradable employment in a given sector against its (log) GDP p.c. (const. \$2011) in the IPUMS USA sample over the period 1850-2010. The third set of figures plots a binned scatterplot of countries' (log) revealed comparative advantage (RCA) in a given sector against its (log) GDP p.w. (const \$2010) in the World Trade Flows sample over the period 1962-2000.

Figure A1: (1) Binned Scatterplots Employment vs. GDP p.w. (IPUMS International, 1970-2012)

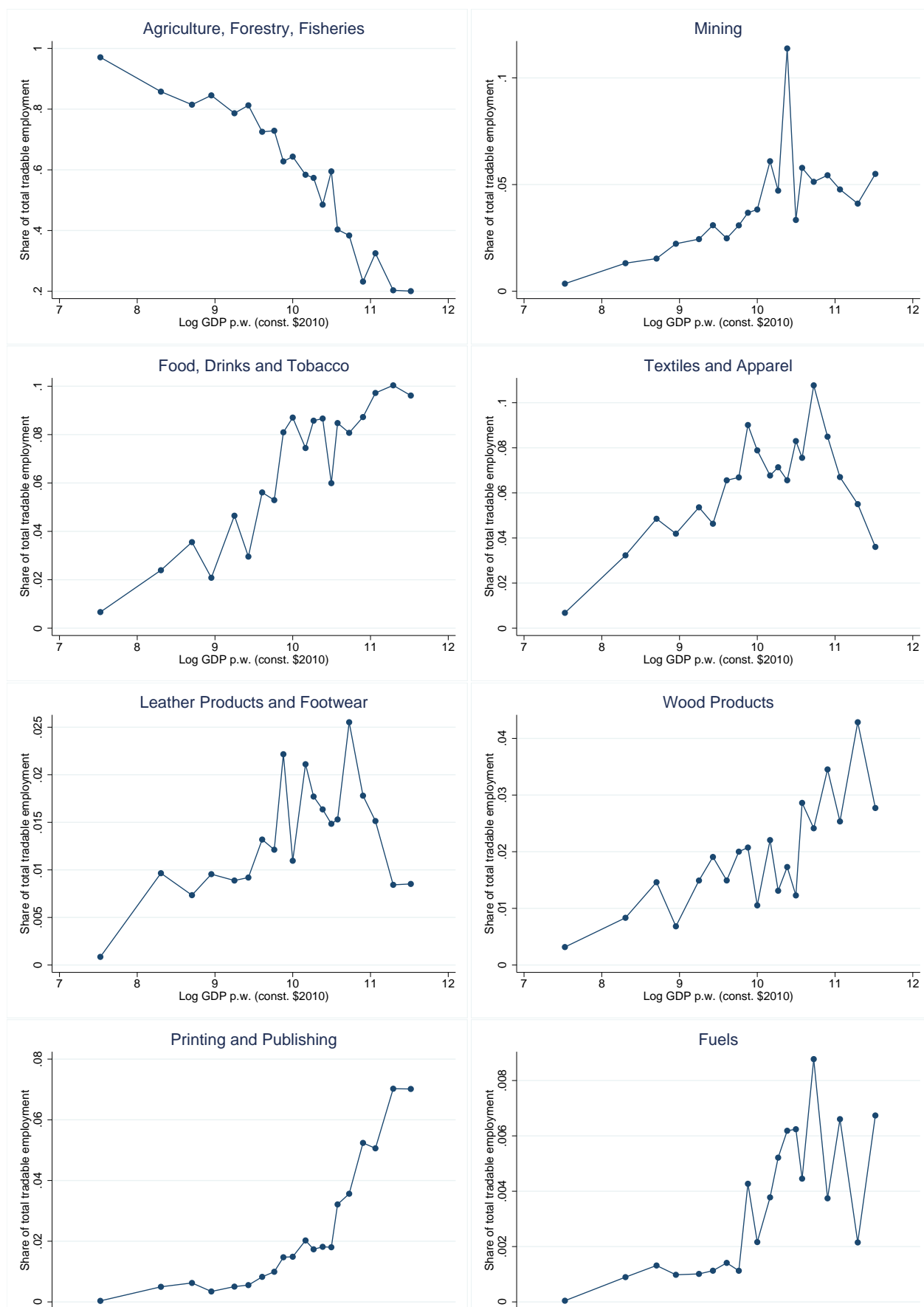


Figure A2: (2) Binned Scatterplots Employment vs. GDP p.w. (IPUMS International, 1970-2012)

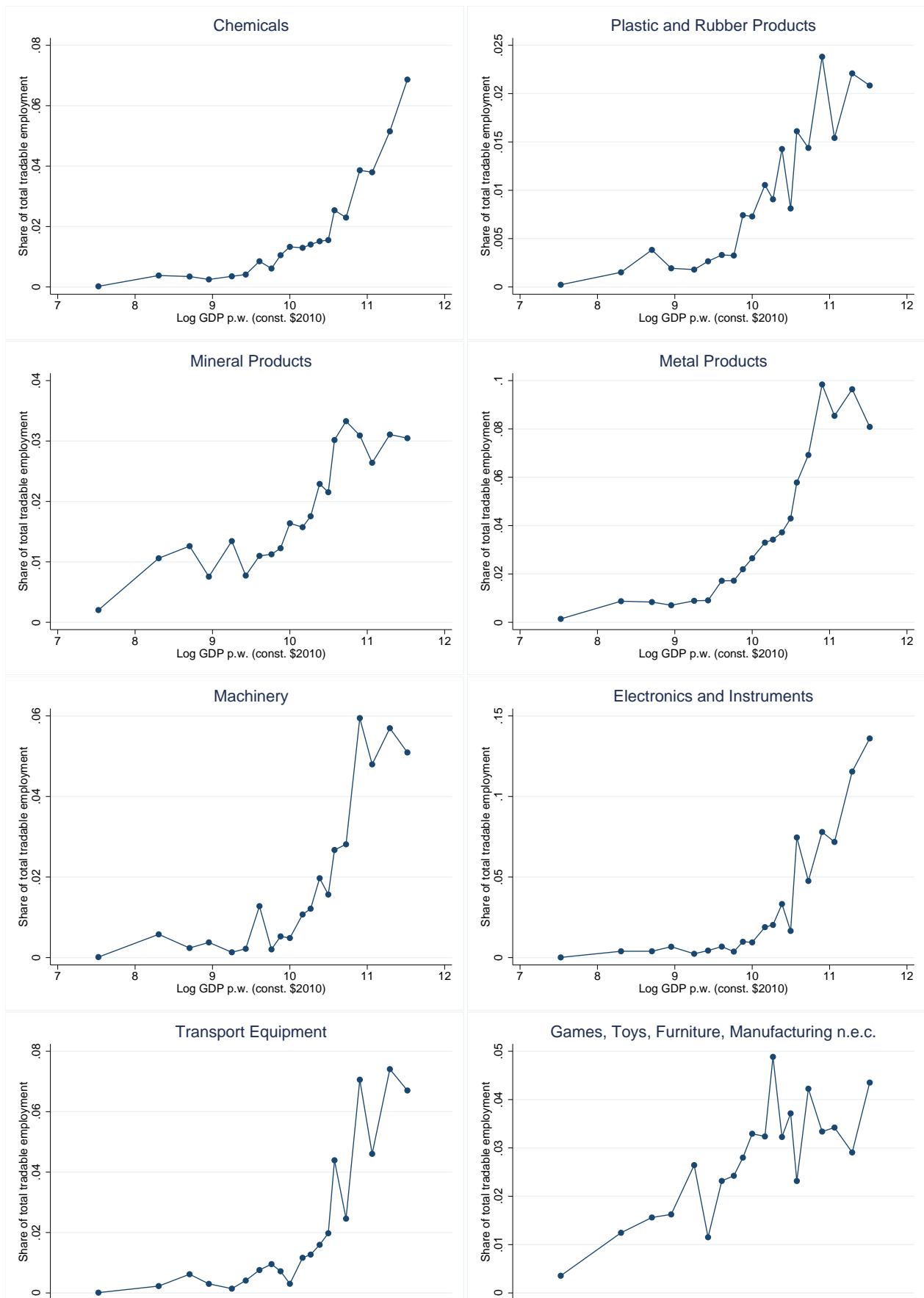


Figure A3: (1) Binned Scatterplots Employment vs. GDP p.w. (IPUMS USA), 1850-2010)

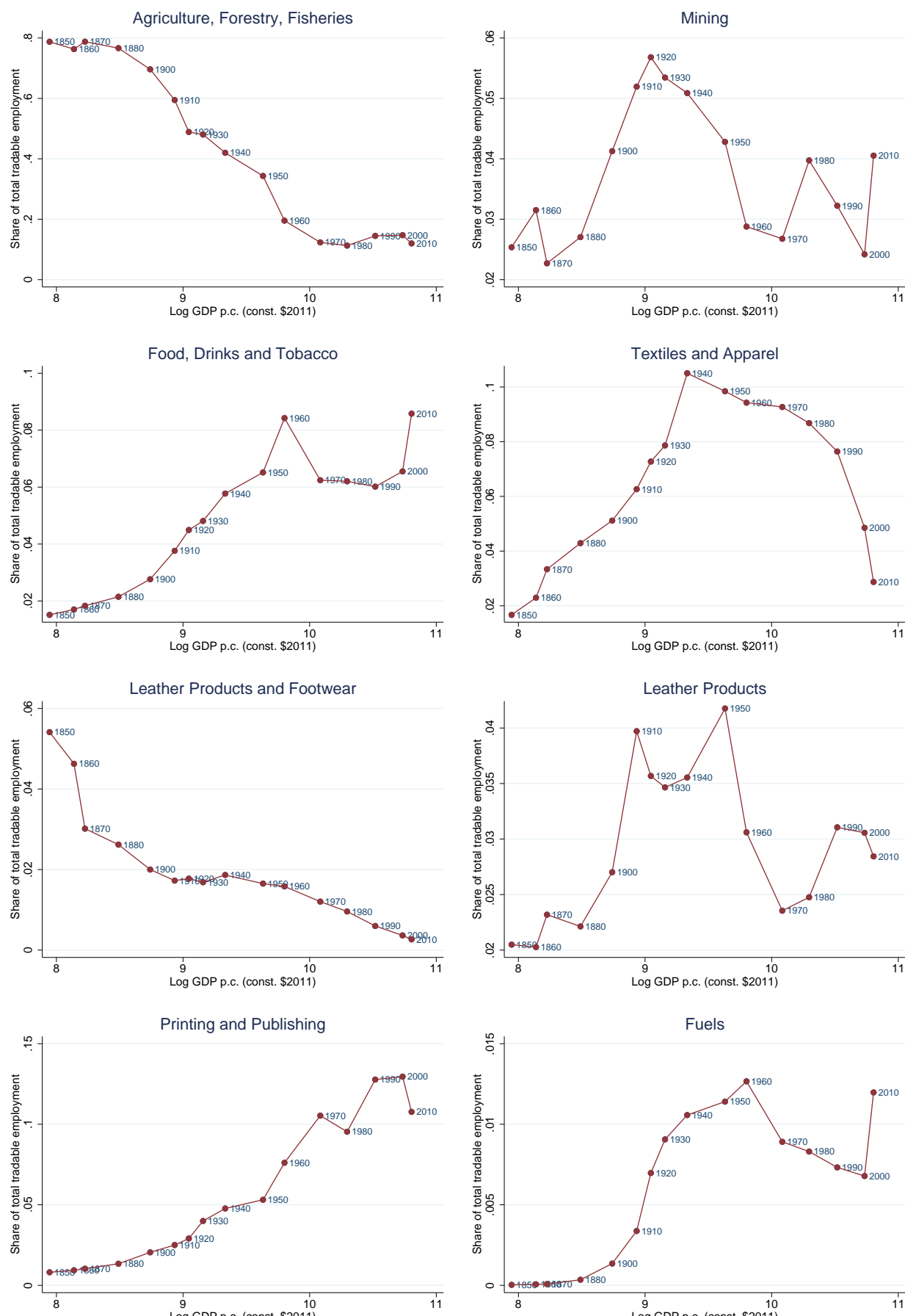


Figure A4: (2) Binned Scatterplots Employment vs. GDP p.w. (IPUMS USA), 1850-2010

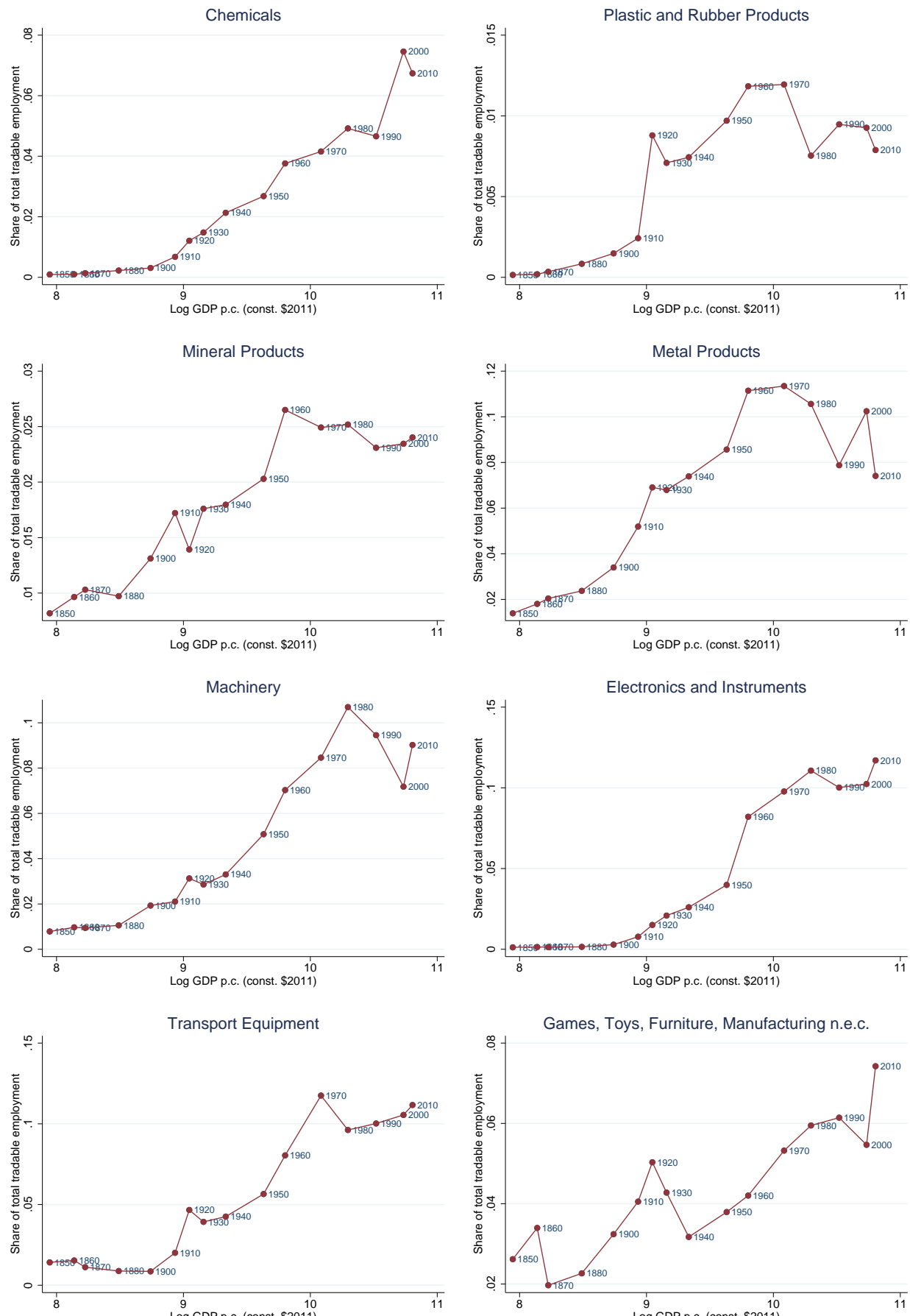




Figure A5: (1) Binned Scatterplots RCA vs. GDP p.w. (World Trade Flows), 1962-2000)

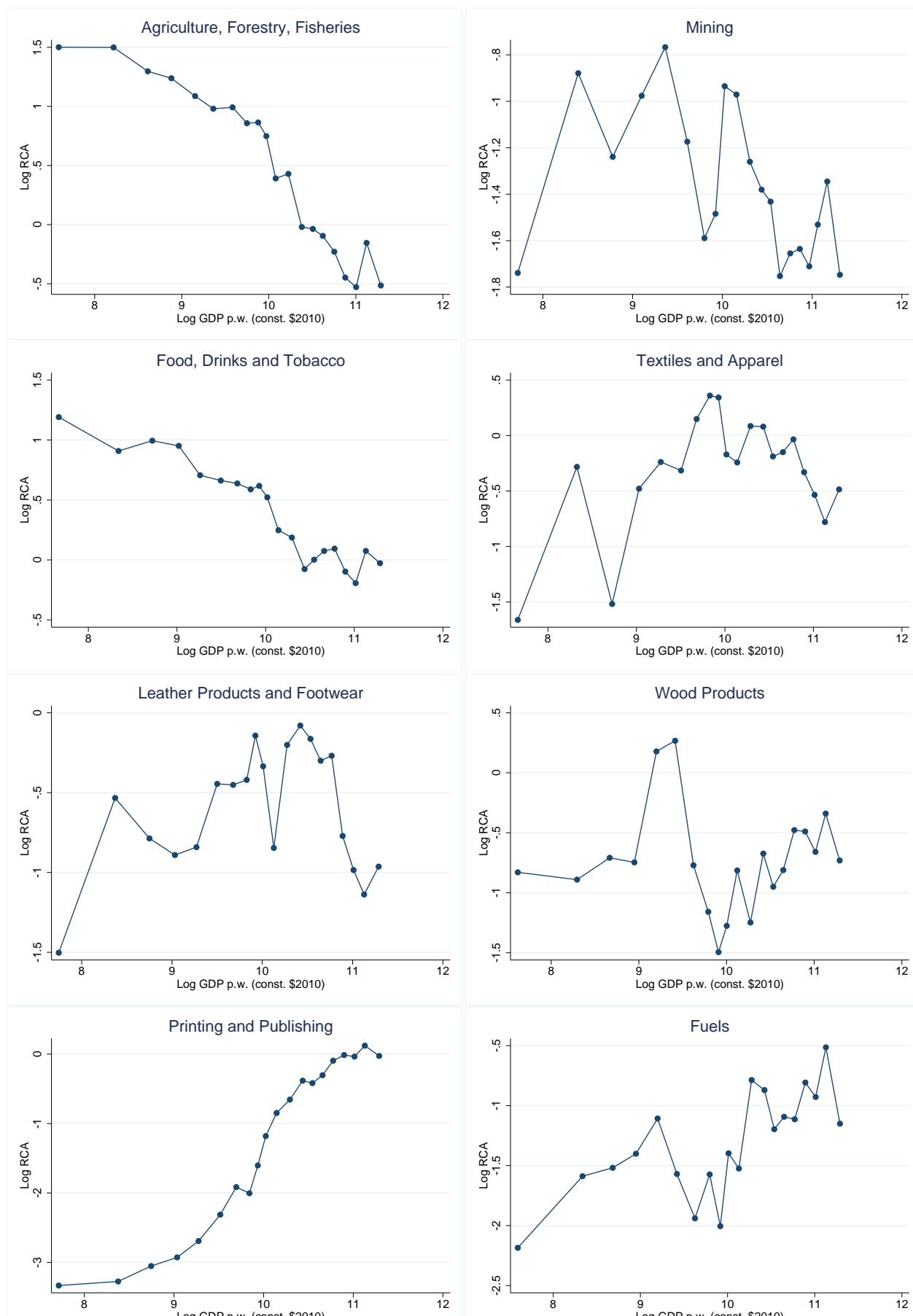
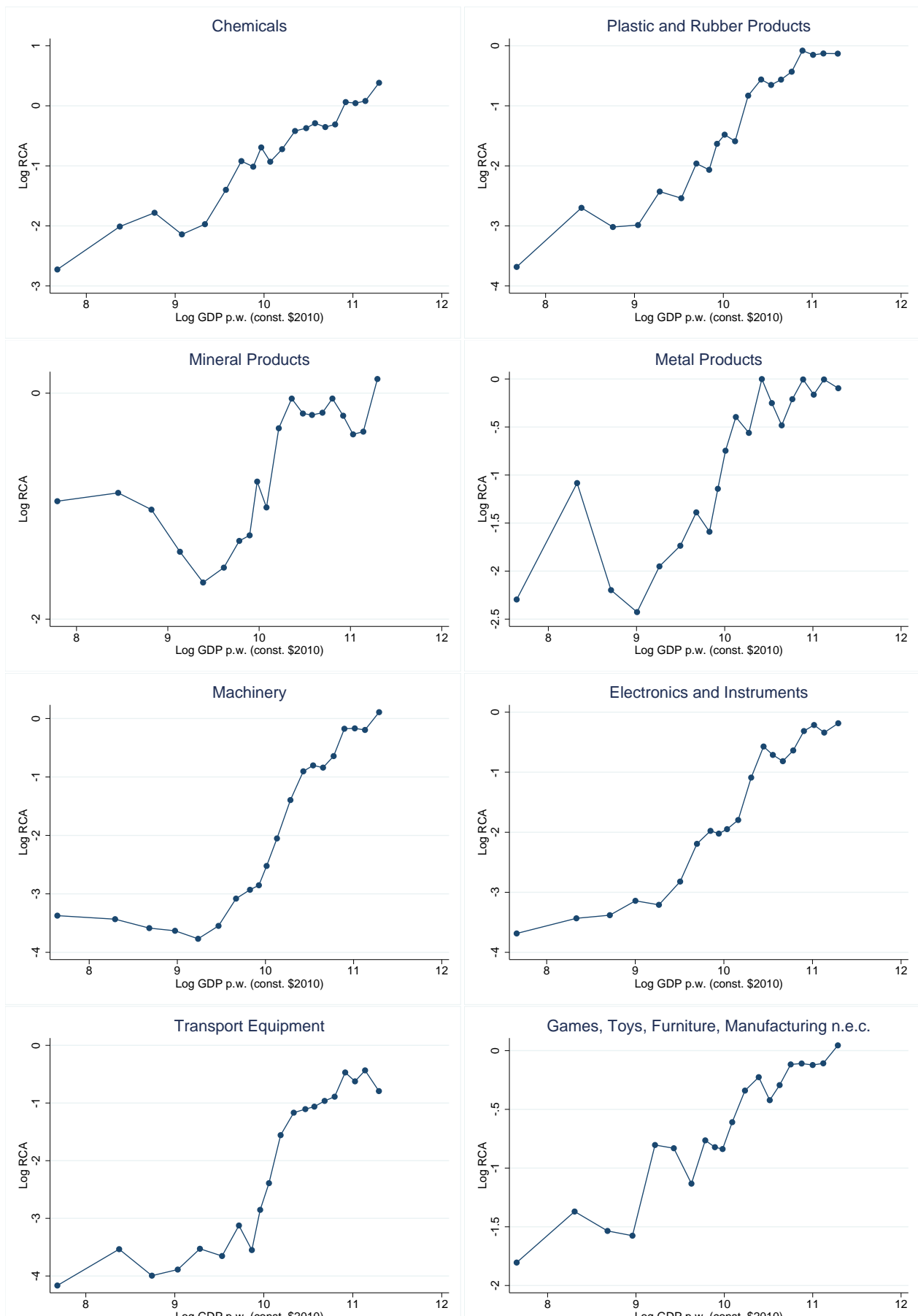


Figure A6: (2) Binned Scatterplots RCA vs. GDP p.w. (World Trade Flows), 1962-2000)



## B. Motivating Facts: Relationship Between RCA and Theoretical Framework

- **Claim 1:** Differences in RCA reflect differences in sector-specific effective unit costs

- Assume no trade costs ( $\tau_{nn',t}^k = 1 \forall n, n', t, k$ ), and constant sector-specific Cobb-Douglas preferences ( $X_{n,t}^k / (I_{n,t} L_{n,t}) = \alpha_k \forall n, t$ ).
- Using

$$Y_{n,t}^k = \alpha_k \sum_{n=1}^N I_{n,t} L_{n,t} \sum_{n'=1}^N \pi_{n'n,t}^k \text{ and } \pi_{n'n,t}^k = T_{n,t}^k (c_{n,t}^k P_{n,t}^k / \Lambda_k)^{-\theta}$$

express relative RCA as a function of relative sector-specific effective unit costs:

$$\frac{RCA_{n,t}^k / RCA_{n,t}^{k'}}{RCA_{m,t}^k / RCA_{m,t}^{k'}} = \frac{(T_{n,t}^k / T_{n,t}^{k'}) (c_{n,t}^k / c_{n,t}^{k'})^{-\theta}}{(T_{m,t}^k / T_{m,t}^{k'}) (c_{m,t}^k / c_{m,t}^{k'})^{-\theta}}$$

Given that exporter fixed effects from the gravity equation equal

$$\delta_{m,t}^k = -\theta \ln c_{m,t}^k / c_{US,t}^k + \ln T_{n,t}^k / T_{US,t}^k$$

relative RCA also correspond to relative exporter fixed effects (in logs):

$$\ln \frac{RCA_{n,t}^k / RCA_{n,t}^{k'}}{RCA_{US,t}^k / RCA_{US,t}^{k'}} = \frac{\delta_{n,t}^k / \delta_{n,t}^{k'}}{\delta_{m,t}^k / \delta_{m,t}^{k'}}.$$

- **Claim 2:** Estimating equation 5 in section 2 corresponds to the sector-level version of dynamic scale economies in equation 20 (section 3.2)

- In addition, to assuming no trade costs and constant sector-specific Cobb-Douglas preferences, assume sector-specific TFP terms are constant over time (e.g.  $T_{n,t}^k = 1 \forall n, t, k$ ) and there are no deficits ( $D_{n,t} = 0$ ). Express log changes in RCA as

$$* \Delta \ln RCA_{n,t}^k = \Delta \ln X_{n,t}^k - \Delta \ln X_{n,t} - \Delta \ln X_t^k + \Delta \ln X_t$$

where the last three terms will be captured by country-time and sector-time fixed effects. Inserting the previous expression for  $X_{n,t}^k$

$$* \Delta \ln RCA_{n,t}^k = -\theta \Delta \ln c_{n,t}^k - \Delta \ln X_{n,t} - \Delta \ln X_t^k + 2\Delta \ln X_t - \theta \Delta \ln P_t^k$$

where the sectoral price index does not have a country-specific subscript as it is the same across countries. Using the definition of unit costs:

$$* \Delta \ln RCA_{n,t}^k = \theta \sum_{a=1}^A \zeta_a^k \Delta \ln T_{n,t}^a - \theta \Delta \ln w_{n,t} - \Delta \ln X_{n,t} - \Delta \ln X_t^k + 2\Delta \ln X_t - \theta \Delta \ln P_t^k$$

Using endogenous task productivity growth ( $\Delta \ln T_{n,t}^a = \beta_0 + \tilde{\eta}_a \ln L_{n,t-1}^a + (\beta - 1) \ln T_{n,t-1}^a + \epsilon_{n,t}^a$  with  $\theta_a \approx 1$ ):

$$* \Delta \ln RCA_{n,t}^k = \theta [\beta_0 + (1 - \beta) \ln c_{n,t-1}^k + \sum_{a=1}^A \zeta_a^k \tilde{\eta}_a \ln L_{n,t-1}^a] + \theta [(\beta - 1) \ln w_{n,t-1} + \sum_{a=1}^A \zeta_a^k \eta_{n,t}^a] - \theta \Delta \ln w_{n,t} - \Delta \ln X_{n,t} - \Delta \ln X_t^k + 2\Delta \ln X_t - \theta \Delta \ln P_t^k$$

Inserting relationship between unit costs and revealed comparative advantage used earlier gives:

$$* \Delta \ln RCA_{n,t}^k = \theta \beta_0 - (1 - \beta) \ln RCA_{n,t-1}^k + \theta \sum_{a=1}^A \zeta_a^k \tilde{\eta}_a \ln L_{n,t-1}^a + \theta[(\beta - 1) \ln w_{n,t-1} + \sum_{a=1}^A \zeta_a^k \eta_{n,t}^a] - \theta \Delta \ln w_{n,t} - \Delta \ln X_{n,t} - \Delta \ln X_t^k + 2\Delta \ln X_t - \theta \Delta \ln P_t^k + \frac{(1-\beta)}{\theta} [\ln \alpha_k + \ln X_{t-1} + \theta \ln \Lambda_k - \theta \ln P_{t-1}^k]$$

- If the diffusion parameter  $\tilde{\eta}_a$  is constant across tasks, substituting in the definition of related RCA  $RR_{n,t}^k$  yields

$$* \Delta \ln RCA_{n,t}^k = \theta \beta_0 - (1 - \beta) \ln RCA_{n,t-1}^k + \theta \tilde{\eta} \ln RR_{n,t-1}^k + \theta[(\beta - 1) \ln w_{n,t-1} + \sum_{a=1}^A \zeta_a^k \eta_{n,t}^a] - \theta \Delta \ln w_{n,t} - \Delta \ln X_{n,t} - \Delta \ln X_t^k + 2\Delta \ln X_t - \theta \Delta \ln P_t^k + \frac{(1-\beta)}{\theta} [\ln \alpha_k + \ln X_{t-1} + \theta \ln \Lambda_k - \theta \ln P_{t-1}^k] + \theta \sum_{a=1}^A \frac{\zeta_a^k X_t^k}{X_t}$$

- Finally, rearranging yields the estimating equation 5 in section 2

$$* \Delta \ln RCA_{n,t}^k = \frac{\theta}{\theta_a} \ln \Gamma(1 - \beta) - (1 - \beta) \ln RCA_{n,t-1}^k + \theta \tilde{\eta} \ln RR_{n,t-1}^k + \delta_{n,t} + \delta_n^k + \delta_t^k + \epsilon_{n,t}^k$$

- where the country-time fixed effect  $\delta_{n,t}$  captures  $\theta(\beta - 1) \ln w_{n,t-1} - \theta \Delta \ln w_{n,t} - \Delta \ln X_{n,t}$ , the country-sector fixed effect  $\delta_n^k$  captures  $\frac{(1-\beta)}{\theta} \ln \alpha_k + (1 - \beta) \ln \Lambda_k$ , and the sector-time fixed effect  $\delta_t^k$  captures  $-\Delta \ln X_t^k + 2\Delta \ln X_t - \theta \Delta \ln P_t^k \frac{(1-\beta)}{\theta} \ln X_{t-1} - (1 - \beta) \ln P_{t-1}^k + \theta \sum_{a=1}^A \frac{\zeta_a^k X_t^k}{X_t}$ .

## C. Theoretical Framework: Mapping to Literature on External Economies of Scale

In this section, I present a mapping between the dynamic scale economies in this paper and the external economies of scale in the recent papers by [Bartelme et al. \(2019a\)](#) and [Kucheryavy et al. \(2020\)](#). In particular, I show that in a steady state, dynamic scale economies manifest themselves as external economies of scale.

Let's lay the groundwork first. Consider a firm producing variety  $\omega_k$  in sector  $k$  in country  $n$ , hiring different types of labor at a common wage rate  $w$ . Normalize the total labor endowment to one. We can express the firm's production function as

$$q_{n,t}^k(\omega_k) = z_{n,t}^k(\omega_k) T_{n,t}^k L_{n,t}^k(\omega_k)$$

where  $T_{n,t}^k$  is a Hicks-neutral term capturing effective labor productivity, and  $L_{n,t}^k$  is a composite input of labor. Define  $L_{n,t}^k \equiv \int L_{n,t}^k(\omega)$  as the total amount of labor employed in sector  $k$ . In [Bartelme et al. \(2019a\)](#), the labor productivity term takes the form:

$$T_{n,t}^k = E_n^k(L_{n,t}^k)$$

where  $E_n^k$  is a structural object, estimated by [Bartelme et al. \(2019a\)](#), that determines external economies of scale across sectors. As in the literature on agglomeration economies ([Duranton and Puga, 2004](#)), [Kucheryavy et al. \(2020\)](#) assume that  $E_n^k$  takes a log-linear parametric form such that

$$T_{n,t}^k = (L_{n,t}^k)^{\phi^k}$$

with  $\phi^k \geq 0$  a sector-specific parameter that determines the degree of economies of scale. In my model, on the other hand, effective labor productivity can be expressed as:

$$T_{n,t}^k = \prod_{a=1}^A (\zeta_a^k T_{n,t}^a)^{\zeta_a^k}$$

I now show that, in a steady state, dynamic scale economies manifest themselves as external economies of scale. As explained in section 3.2.3, a balanced growth path exists when (i) countries are in autarky and preferences and sector-specific productivities are constant, or (ii) countries are the same in terms of fundamentals. We consider the first case. Let  $\gamma_n^a$  denote the constant growth rate of task productivity, which we can express as

$$\Delta \ln T_{n,t}^a = \gamma_n^a = \beta_a + \tilde{\eta}_a \ln L_n^a + \kappa_n^a$$

where  $\kappa_n^a$  is a constant term that reflects the balanced forces of the fishing-out effect and exponential growth of ideas (see section 3.2.3). Now consider two countries,  $i$  and  $n$ , with the same growth of ideas but with different allocations of labor across tasks (e.g. due to differences in preferences). For comparison, normalize initial task productivity levels to one. In this case, taking a snapshot of relative sectoral productivity levels gives:

$$\frac{T_{n,t}^k}{T_{i,t}^k} = \prod_{a=1}^A \left( \frac{L_n^a}{L_i^a} \right) \tilde{\eta}_a \varsigma_a^k$$

which is a general case of the productivity specification in [Kucheryavyy et al. \(2020\)](#). If tasks are specific to sectors, this collapses to:

$$\frac{T_{n,t}^k}{T_{i,t}^k} = \prod_{a=1}^A \left( \frac{L_n^k}{L_i^k} \right) \tilde{\eta}_k$$

in which case dynamic scale economies manifest as external economies of scale in [Kucheryavyy et al. \(2020\)](#), i.e.

$$\phi^k = \tilde{\eta}^k$$

## D. Theoretical Framework: Extensions

This section presents three extensions of the theoretical framework outlined in the main text.

### D.1 Capital

As is common in the literature on job polarization (Burstein et al., 2015; Goos et al., 2014), I model capital as an additional input in the production of tasks. Now a typical task producer solves the problem:

$$\max_{L_{n,t}^a \geq 0, K_{n,t}^a \geq 0} p_{n,t}^a (K_{n,t}^a)^\kappa (T_{n,t}^a L_{n,t}^a)^{1-\kappa} - r_{n,t} K_{n,t}^a - w_{n,t}^a T_{n,t}^a L_{n,t}^a$$

where  $\kappa \in [0, 1]$ , calibrated to 0.24 in the quantitative exercises (Burstein et al., 2015).<sup>45</sup> I do not micro-found household savings behavior, instead, each period households save an exogenous fraction  $s_{n,t}$  of their current income and invest this in the next period's capital.<sup>46</sup> If capital is freely mobile at global rate  $r_t$ , the static counterfactual equilibrium changes in three ways:

- (task price)  $\hat{p}_{n,t}^a = (\hat{r}_t)^\kappa (\hat{w}_{n,t}^a)^{1-\kappa}$
- (trade balance)  $\sum_{k=1}^K \sum_{i=1}^N \alpha_{n,t}^k ((1-s_{n,t})(w_{n,t} L_{n,t} \hat{w}_{n,t} + r_t \hat{r}_t K_{n,t}) + D_{n,t}) \pi_{ni,t}^k = \sum_{k=1}^K \sum_{i=1}^N \alpha_{i,t}^k ((1-s_{i,t})(w_{i,t} L_{i,t} \hat{w}_{i,t} + r_t \hat{r}_t K_{i,t}) + D_{i,t}) \pi_{in,t}^k + D_{n,t}$
- (additional)  $r_t \hat{r}_t \sum_{n=1}^N K'_{n,t} = \kappa \sum_{n=1}^N (w_{n,t} L_{n,t} \hat{w}_{n,t} + D_{n,t} + r_t \hat{r}_t K'_{n,t})$

where the last condition is global capital market clearing.

The additional condition for the dynamic counterfactual equilibrium becomes:

- Between two periods  $t = \tau$  and  $t = \tau - 1$ , the law of motions for counterfactual capital levels satisfy:

$$K'_{n,t} = K'_{n,t-1} + s_{n,t-1} (w'_{n,t} L_{n,t} + D_{n,t} + r'_t K'_{n,t-1})$$

Finally, note that including capital in the model changes the interpretation of the estimates of the spillover parameters  $\eta_a$  in section 6.1.2. These need to be scaled by a factor  $(1 - \kappa)$  to get proper estimates  $\hat{\eta}_a$ .

### D.2 Input-Output Linkages

I model the production and trade of intermediate goods as roundabout production (e.g. Caliendo and Parro, 2015). Now the variety's production function is

$$q_{n,t}^k(\omega_k) = z_{n,t}^k(\omega_k) \left[ \prod_{a=1}^A (t_{n,t}^a(\omega_k))^{\zeta_a^k} \right]^{\lambda_{n,t}^k} (M_{n,t}^k(\omega_k))^{1-\lambda_{n,t}^k} \quad (53)$$

where  $M_{n,t}^k(\omega_k)$  is the composite intermediate good for sector  $k$ . A typical variety producer solves the problem

<sup>45</sup>The assumption of a common output elasticity of capital across tasks -common in the literature- is due to a lack of data on task-specific income shares of capital.

<sup>46</sup>Exogenous saving rates are necessary to match the evolution of the capital stock. Any variation in  $s_{n,t}$  thus captures the effects of economic growth, structural transformation, investment and savings frictions, government policy, cultural differences, as well as demographic developments.

$\max_{\{t_{n,t}^a(\omega_k) \geq 0\}_{a=1}^A, M_{n,t}^k(\omega_k) \geq 0} p_{n,t}^k(\omega_k) q_{n,t}^k(\omega_k) - \sum_{a=1}^A p_{n,t}^a t_{n,t}^a(\omega_k) - p_{n,t}^{k,M} M_{n,t}^k(\omega_k)$   
 which gives rise to the cost of an input bundle:

$$c_{n,t}^k = \Gamma_k \left[ \prod_{a=1}^A (p_{n,t}^a)^{\zeta_a^k} \right]^{\lambda_{n,t}^k} [p_{n,t}^{k,M}]^{1-\lambda_{n,t}^k} \quad (54)$$

where  $p_{n,t}^{k,M}$  is the price of a unit of the composite intermediate good for sector  $k$ . Intermediate goods producers for sector  $k$  in country  $n$  combine aggregates from different sectors to produce according to

$$Q_{n,t}^{k,M} = \prod_{k'=1}^K (Q_{n,t}^{kk'})^{\iota_{n,t}^{kk'}} \quad (55)$$

where  $Q_{n,t}^{kk'}$  is the quantity of sector aggregates from sector  $k'$  in country  $'$  used as intermediates in the production of materials for sector  $k$ .  $\iota_{n,t}^{kk'}$  is an input weight that captures the relative importance of different intermediates.

A typical material producer solves the problem:

$$\max_{\{Q_{n,t}^{kk'}\}_{k'=1}^K} p_{n,t}^{k,M} Q_{n,t}^{k,M} - \sum_{k'=1}^K P_{n,t}^{k'} Q_{n,t}^{kk'}$$

The solution to this problem gives rise to the input price  $p_{n,t}^{k,M}$ :

$$p_{n,t}^{k,M} = \Gamma^{k,M} \prod_{k'=1}^K (P_{n,t}^{k'})^{\iota_{n,t}^{kk'}} \quad (56)$$

where  $\Gamma^{k,M}$  is a sector-specific constant. Now the only changes to the static counterfactual equilibrium are:

- (unit costs)  $\hat{c}_{n,t}^k = [\prod_{a=1}^A (\hat{p}_{n,t}^a)^{\zeta_a^k}]^{\lambda_{n,t}^k} [\prod_{k'=1}^K (\hat{P}_{n,t}^{k'})^{\iota_{n,t}^{kk'}}]^{1-\lambda_{n,t}^k}$
- (trade balance)  $\sum_{k=1}^K \sum_{i=1}^N (\Psi_{n,t}^k)' \pi_{ni,t}^k \hat{\pi}_{ni,t}^k = \sum_{k=1}^K \sum_{i=1}^N (\Psi_{i,t}^k)' \pi_{in,t}^k \hat{\pi}_{in,t}^k + D_{n,t}$

where  $\Psi_{n,t}^k = \sum_{j=1}^K \iota_{n,t}^{jk} (1 - \lambda_{n,t}^k) \sum_{n=1}^N \pi_{in,t}^j \Psi_{i,t}^j + \alpha_{n,t}^k (w_{n,t} L_{n,t} + D_{n,t})$  is the total expenditure on sector  $k$  in country  $n$ , which is the sum of intermediate and final expenditure.

Changes in real wages (welfare) now become:

$$\ln \hat{w}_{n,t} / \hat{P}_{n,t} = - \sum_{k=1}^K \frac{\alpha_{n,t}^k}{\theta} \ln \hat{\pi}_{nn,t}^k - \sum_{k=1}^K \frac{\alpha_{n,t}^k}{\theta} \frac{1 - \lambda_{n,t}^k}{\lambda_{n,t}^k} \ln \hat{\pi}_{nn,t}^k - \sum_{k=1}^K \frac{\alpha_{n,t}^k}{\lambda_{n,t}^k} \ln \prod_{j=1}^K (\hat{P}_{n,t}^j / \hat{P}_{n,t}^k)^{\iota_{n,t}^{jk} (1 - \lambda_{n,t}^k)} - \sum_{k=1}^K \alpha_{n,t}^k \sum_{a=1}^A \zeta_a^k \tilde{\eta}_a \ln \hat{L}_{n,t}^a \quad (57)$$

where the first three terms summarize the static gains (as in (Caliendo and Parro, 2015)), and the last term captures the dynamic gains.

### D.3 Worker Assignment

I now show that the household behavior presented in the main text is a limiting case of a common worker assignment (Ricardo-Roy) model (Lee, 2015; Galle et al., 2017).

Households are heterogeneous with respect to their endowments of task efficiency units, and sort into the production of a specific task. More specifically, each household draws a vector of independent



idiosyncratic productivity shifters  $\{z_{n,t}^a\}_{a=1}^A$  from the same Frechet distribution with the cumulative distribution function  $F(z) = \exp(z^{-\theta_x})$ .<sup>47</sup>

At time  $t$ , each household in country  $n$  has the same baseline productivity  $T_{n,t}^a$  that differs by task, country and time period. A worker with task-specific productivity  $z_{n,t}^a T_{n,t}^a$  receives an income equal to  $z_{n,t}^a w_{n,t}^a$  when producing task  $a$ .

Within one period, a typical household faces a two-stage problem. First, given task wages  $\{w_{n,t}^a\}_{a=1}^A$  it maximizes its income by sorting into a specific task  $a$ . Second, given its income and prices the household maximizes its utility by picking a consumption vector. The first problem can be summarized as:

$$\max_{\{\tilde{a}_a\}_{a=1}^A} \sum_{a=1}^A \tilde{a}_a z_{n,t}^a \sum_{a=1}^A w_{n,t}^a T_{n,t}^a \text{ s.t. } \sum_{a=1}^A \tilde{a}_a = 1, \{\tilde{a}_a\}_{a=1}^A \in \{0, 1\}^A$$

The share of workers choosing task  $a$  is then

$$\pi_{n,t}^a = \frac{(w_{n,t}^a T_{n,t}^a)^{\theta_x}}{\sum_{a=1}^A (w_{n,t}^a T_{n,t}^a)^{\theta_x}} \quad (58)$$

The household behavior in the main text is thus a limiting case of this assignment model with  $\theta_x \rightarrow \infty$ , in which case the distribution of worker-specific productivity draws becomes degenerate.

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<sup>47</sup>Using a Frechet distribution is not just convenient here. The distribution's implied wage distribution approximates the actual conditional wage distribution in the U.S. quite well (Burstein et al., 2015; Saez, 2001).

## E. Estimation Strategy: Algorithms

### Algorithm for Computing Static Counterfactual Equilibrium

Given  $\hat{\tau}_{ni,t}^k, \hat{T}_{n,t}^k$ ,

1. Guess  $\hat{w}_{n,t}$ 
  - compute (in this order)  $\hat{w}_{n,t}^a, \hat{p}_{n,t}^a, \hat{c}_{n,t}^k, \hat{P}_{n,t}^k, \hat{\pi}_{ni,t}^k$
  - compute imports and exports
  - if counterfactual trade balance differs from deficit, go back to (1) and adjust guess of  $\hat{w}_{n,t}$
2. Iterate until trade balance has converged to deficit for all countries

## **F. Data: Details**

Table A1: Work Activities O\*NET and their assigned task categories.

<b>Task description</b>	<b>Category</b>
Getting Information	Information
Identifying Objects, Actions, and Events	Information
Estimating the Quantifiable Characteristics of Products, Events, or Information	Information
Judging the Qualities of Things, Services, or People	Information
Processing Information	Information
Evaluating Information to Determine Compliance with Standards	Information
Analyzing Data or Information	Information
Making Decisions and Solving Problems	Information
Thinking Creatively	Information
Updating and Using Relevant Knowledge	Information
Interacting With Computers	Information
Drafting, Laying Out, and Specifying Technical Devices, Parts, and Equipment	Information
Documenting/Recording Information	Information
Interpreting the Meaning of Information for Others	Information
Provide Consultation and Advice to Others	Information
Developing Objectives and Strategies	Planning
Scheduling Work and Activities	Planning
Organizing, Planning, and Prioritizing Work	Planning
Performing Administrative Activities	Planning
Performing General Physical Activities	Physical
Handling and Moving Objects	Physical
Operating Vehicles, Mechanized Devices, or Equipment	Physical
Controlling Machines and Processes	Equipment
Repairing and Maintaining Mechanical Equipment	Equipment
Repairing and Maintaining Electronic Equipment	Equipment
Monitor Processes, Materials, or Surroundings	Equipment
Inspecting Equipment, Structures, or Material	Equipment
Communicating with Supervisors, Peers, or Subordinates	Contact
Communicating with Persons Outside Organization	Contact
Establishing and Maintaining Interpersonal Relationships	Contact
Assisting and Caring for Others	Contact
Selling or Influencing Others	Contact
Resolving Conflicts and Negotiating with Others	Contact

Table A2: Top 5 occupations and sectors, by task. Occupations are ranked by their normalized task score. Sectors are ranked by the share of wages paid to occupations assigned to a task.

<b>Task</b>	<b>Top 5 occupations (1970 census)</b>	<b>Top 5 sectors (1970 census)</b>
Information	Physicists Aeronautical engineers Chemical engineers Mining engineers Draftsmen	Office and store machines and devices Aircraft and parts Drugs and medicines Photographic equipment and supplies Forestry
Planning	Authors Electrotypers and stereotypers Hucksters and peddlers Newsboys Advertising agents and salesmen	Miscellaneous food industries Drugs and medicines Printing, publishing and allied industries Paints, varnishes and related products Professional equipment and supplies
Equipment	Furnacemen, smeltermen and pourers Heaters (metal) Motion picture projectionists Oilers and greasers Pressmen and plate printers	Knitting mills Apparel and accessories Footwear Coal mining Miscellaneous fabricated textile products
Mechanical	Fishermen and oystermen Glaziers Mail carriers Plumbers and pipe fitters Welders and flame cutters	Agriculture Fisheries Logging Dairy products Bakery products

Table A3: List of tradable sectors in US Census 1950 classification.

Sector	Sector (2)
Agriculture	Photographic equipment and supplies
Own farm	Watches, clocks, and clockwork-operated devices
Forestry	Meat products
Fisheries	Dairy products
Hunting	Canning and preserving fruits, vegetables, and seafoods
Metal mining	Grain-mill products
Coal mining	Bakery products
Crude petroleum and natural gas extraction	Confectionery and related products
Nonmetallic mining and quarrying, except fuel	Beverage industries
Logging	Miscellaneous food preparations and kindred products
Sawmills, planing mills, and mill work	Tobacco manufactures
Miscellaneous wood products	Knitting mills
Furniture and fixtures	Dyeing and finishing textiles, except knit goods
Glass and glass products	Carpets, rugs, and other floor coverings
Cement, concrete, gypsum and plaster products	Yarn, thread, and fabric mills
Structural clay products	Miscellaneous textile mill products
Pottery and related products	Apparel and accessories
Miscellaneous nonmetallic mineral and stone products	Miscellaneous fabricated textile products
Blast furnaces, steel works, and rolling mills	Pulp, paper, and paperboard mills
Other primary iron and steel industries	Paperboard containers and boxes
Primary nonferrous industries	Miscellaneous paper and pulp products
Fabricated steel products	Printing, publishing, and allied industries
Fabricated nonferrous metal products	Synthetic fibers
Not specified metal industries	Drugs and medicines
Agricultural machinery and tractors	Paints, varnishes, and related products
Office and store machines and devices	Miscellaneous chemicals and allied products
Miscellaneous machinery	Petroleum refining
Electrical machinery, equipment, and supplies	Miscellaneous petroleum and coal products
Motor vehicles and motor vehicle equipment	Rubber products
Aircraft and parts	Leather: tanned, curried, and finished
Ship and boat building and repairing	Footwear, except rubber
Railroad and miscellaneous transportation equipment	Leather products, except footwear
Professional equipment and supplies	

Table A4: List of tradable sectors in WIOD classification.

ISIC rev.3 code	Industry name
A-B	Agriculture, hunting, forestry and fishing
C/E	Mining and quarrying / Electricity, gas and water supply
D15-16	Food, beverages and tobacco
D17-18	Textiles and textile products
D19	Leather, leather products and footwear
D20	Wood and products of wood and cork
D21-22	Pulp, paper, printing and publishing
D23	Coke, refined petroleum and nuclear fuel
D24	Chemicals and chemical products
D25	Rubber and plastics
D26	Other non-metallic minerals
D27-28	Basic metals and fabricated metals
D29	Machinery, not elsewhere classified
D30-33	Electrical and optical equipment
D34-35	Transport equipment
D36-37	Manufacturing, not elsewhere classified; recycling

## F.1 Bilateral Trade Flows

Data on bilateral sectoral trade flows are from the World Trade Flows (WTF) database developed by Feenstra et al. (2005). These cover bilateral trade between country pairs at the disaggregated four digit SITC2 level for the years 1962-2000. I aggregate goods to the level of industries from the 1950 Census Bureau industrial classification system. Table A3 shows the corresponding list of sectors. In the rest of the procedure, I follow Hanson et al. (2018). I create a balanced panel of countries by maintaining as single units countries that split up or unite (Czech Republic, Russia, Yugoslavia, Germany, Yemen) and restrict the analysis to countries that form a connected set to be able to identify importer and exporter fixed effects (Abowd et al., 2002). This leaves me with 87 countries and 59 (tradable) sectors.

The WTF database does not contain information on the consumption of domestically produced goods, 'self trade'  $X_{nn,t}^k$ . However, to be able to perform counterfactuals and calibrate expenditure shares, estimates of self-trade are necessary. As a result, I need to infer self trade at the country-industry level to compute industry-level expenditure. Hanson et al. (2018) show that, if a country's log trade costs have a common additively separable component and there are no internal trade costs,  $\tau_{nn,t}^k = 1$ , a country's self trade in sector  $k$  as a share of total self trade is given by

$$\frac{X_{nn,t}^k}{\sum_{k'=1}^K X_{nn,t}^{k'}} = \frac{\exp(\kappa_{n,t}^k + \tilde{\mu}_{n,t}^k)}{\sum_{k'=1}^K \exp(\kappa_{n,t}^{k'} + \tilde{\mu}_{n,t}^{k'})} \quad (59)$$

where  $\kappa_{n,t}^k$  and  $\tilde{\mu}_{n,t}^k$  are a country's fixed effects estimates from the gravity equation estimation in section 4.

One can then use use production data in tradable sectors to infer aggregate self trade as the difference between aggregate production  $Y_{n,t}$  and exports:

$$\sum_{k=1}^K X_{nn,t}^k = Y_{n,t} - \sum_{k=1}^K \sum_{n' \neq n} X_{n'n,t}^k \quad (60)$$

As some trade costs differ systematically across sectors, using aggregate production data to infer self trade using this method leads to substantial measurement error in the self trade estimates. I therefore deviate from Hanson et al. (2018) by assuming trade costs have a common additively separable *sector-level* component, and use estimates of sector-level production data to estimate self trade at the *industry* level.

I classify sectors according to the World Input Output Database (Table A4). I take value added production data for primary sectors and manufacturing from UN National Accounts. This leaves me with obtaining estimates of value added *shares* of manufacturing subsectors, for which I use production data from UNIDO INDSTAT 2.0. Along the way, to obtain estimates for country-sector-year cells with missing data, I extrapolate from non-missing observations by projecting variables onto log GDP per capita (Penn World Tables 9.0) and a time trend.

One cannot simply combine the WTF and production data from national accounts because the former are in terms of gross output and the latter in terms of value added. As gross output production data are not widely available, I convert any estimates of value added production data into gross output using yearly sector-level estimates from Korea KLEMS.

Finally, I estimate industry-level self trade as the product of equations 59 and 60. Doing so requires importer and exporter fixed effects for all industry-country-year cells, however. As not all countries import and/or export in all industries, I estimate synthetic fixed effects by extrapolating from non-missing



observations by projecting fixed effects onto log GDP per capita and a time trend.

## F.2 Occupations (O\*NET)

I use detailed occupation-level information to assign occupations to subgroups (task groups). The Occupational Information Network (O\*NET) is my primary source for information on the standardized work characteristics of occupations and sectors. Its O\*NET database contains hundreds of standardized occupation-specific descriptors on almost 1,000 occupations that cover the entire U.S. economy. In particular, the database provides information on "Work Activities", which "(...) *summarize the kinds of tasks that may be performed across multiple occupations.*" As such, it provides a standardized set of tasks that are comparable across occupations and sectors.

Each descriptor in O\*NET is associated with at least one scale, which are standardized to a score ranging from 0 to 100. The values of these scores are the average response of survey participants that work in a specific occupation. The database contains two scales for Work Activities: Importance and Level. The Importance scale "(...) *indicates the degree of importance a particular descriptor is to the occupation.*" The Level scale "(...) *indicates the degree, or point along a continuum, to which a particular descriptor is required or needed to perform the occupation..*" I choose to work with the Level scale. In total, O\*NET contains data on 41 standardized work activities, which are summarized in Table [A1](#).

## G. Results

### G.1 Spillovers at the Sector-Level

To illustrate how differences in diffusion parameters at the task level translate into differences at the sector-level, I construct a measure that captures the effect of moving labor into a certain sector on global aggregate GDP relative to other sectors. Denote the effect of moving all labor into sector  $k$  on productivity in sector  $\tilde{k}$  by  $\Delta \ln \tilde{y}_{n,t}^{\tilde{k}}$ . The effect on aggregate GDP is then given by

$$\sum_{\tilde{k}=1}^K \theta_t^{\tilde{k}} \Delta \ln \tilde{y}_{n,t}^{\tilde{k}} \quad (61)$$

where  $\theta_t^{\tilde{k}}$  is the share of global exports accounted for by sector  $\tilde{k}$ . Substituting in dynamic scale economies when all labor is in sector  $k$ :

$$\sum_{\tilde{k}=1}^K \theta_t^{\tilde{k}} \sum_{a=1}^A \zeta_a^{\tilde{k}} \tilde{\eta}_a \ln L_{n,t}^{a,k} \quad (62)$$

Lastly, I express this measure of spillovers relative to a benchmark sector, agriculture (denoted by AB), which has the lowest extent of spillovers due to its intensive use mechanical occupations. The final measure of spillovers is thus

$$\sum_{\tilde{k}=1}^K \theta_t^{\tilde{k}} \sum_{a=1}^A \zeta_a^{\tilde{k}} \tilde{\eta}_a (\ln L_{n,t}^{a,k} - L_{n,t}^{a,AB}) \quad (63)$$

The estimated magnitude of this measure of sector-specific spillovers are summarized in Table A5. Spillovers tend to be highest as a result of allocating labor to sectors that use high-skilled occupations intensively, such as electrical and optical equipment, chemicals and fuels.

Table A5: Spillovers by Tradable Sector r.t. Agriculture (WIOD classification)

Subsector	Spillover
Agriculture, hunting, forestry and fishing (A & B)	0
Mining and quarrying (C)	0.84
Food, beverages and tobacco (D15-16)	0.90
Textiles and textile products (D17-18)	0.88
Leather, leather products and footwear (D19)	0.88
Wood and products of wood and cork (D20)	0.71
Pulp, paper, printing and publishing (D21-22)	0.97
Coke, refined petroleum and nuclear fuel (D23)	0.95
Chemicals and chemical products (D24)	0.98
Rubber and plastics (D25)	0.92
Other non-metallic minerals (D26)	0.87
Basic metals and fabricated metals (D27-28)	0.91
Machinery, not elsewhere classified (D29)	0.93
Electrical and optical equipment (D30-33)	0.96
Transport equipment (D34-35)	0.86
Manufacturing, not elsewhere classified; recycling (D36-37)	0.94

**Notes:** This table reports the magnitude of spillovers by tradable sector relative to the agricultural sector, taken as the unweighed mean over the years 1970-2000 in the World Trade Flows sample. Relative spillovers are defined as in equation 63. Table uses WIOD classification of sectors, which roughly corresponds to a 2-digit ISIC classification.

## H. Quantitative Implications

### H.1 Exercises with and without Input-Output Linkages

In this section, I present results of the exercises outlined in section 7 of the model with and without input-output linkages (see section D.2).

Note that there is no reliable cross-country data on sector-to-sector trade flows for the time period studied in the main text. I use the 1995-2010 data from the World Input Output Database (WIOD). These data allow me to exactly calibrate to parameters  $\lambda_{n,t}^k$ ,  $\alpha_{n,t}^k$  and  $\ell_{n,t}^{kk'}$  so that this version of the model can be solved using hat algebra.

Table A6: Static and dynamic gains from trade with and without I-O linkages

	Static			Dynamic		
	Mean	Median	Correlation	Mean	Median	Correlation
(A) Without I-O linkages	1.067	1.052	0.98	1.021	1.014	0.84
(B) With I-O linkages	1.132	1.106		1.02	1.012	
Relative (B)/(A)	1.06	1.049		1.00	1.00	

**Notes:** This table reports the static and dynamic gains from trade for the quantitative exercises with and without input-output linkages, on the WIOD sample from 1995 to 2010, with Head and Ries trade costs set at their 1995 levels. Numbers in columns (1)-(2) and (4)-(5) summarize the relative real wage in the data relative to the counterfactual. For example, the unweighed mean static gains from trade in the exercise without I-O linkages equal 6.7 %. Columns (3) and (6) display the correlation between the country's static and dynamic gains from trade under the two models.

Table A6 plots the corresponding static and dynamic gains from trade for the model with and without I-O linkages. For the 40 countries in the WIOD, static gains from trade (holding trade costs at 1995 levels) average 6.7 % in the standard model. These gains roughly double to 13.2 % in the model with I-O linkages.<sup>48</sup>

While incorporating I-O linkages into the model increases the *static* gains from trade (Caliendo and Parro, 2015), this does not change the *dynamic* gains from trade. Under both models, the dynamic gains from trade equal roughly 1/3 of the static gains of the standard model, which corresponds to the exercises summarized in the main text.

<sup>48</sup>This doubling corresponds to a back of the envelope calculation that can be done by rewriting the decomposition of the static gains from trade in 57:  $\ln \hat{w}_{n,t} / \hat{P}_{n,t} = - \sum_{k=1}^K \frac{\alpha_{n,t}^k}{\theta} \frac{1}{\lambda_{n,t}^k} \ln \hat{\pi}_{nn,t}^k - \sum_{k=1}^K \frac{\alpha_{n,t}^k}{\lambda_{n,t}^k} \ln \Pi_{j=1}^K (\hat{P}_{n,t}^j / \hat{P}_{n,t}^k)^{\ell_{n,t}^{jk} (1 - \lambda_{n,t}^k)}$ . As the third term equals zero on average, the change in gains of the model with I-O linkages is determined by the sectoral multiplier  $\frac{1}{\lambda_{n,t}^k}$ , which equals the sector-specific share of value added in sales. In tradable sectors, this multiplier is roughly 2 as value added accounts for roughly 50 % of sales.