Revision of the paper "Nonparametric comparison: of epidemic time trends: the case of COVID-19"

First of all, we would like to thank the editor, the associate editor and the reviewers for their many comments and suggestions which were very helpful in improving the paper. In the revision, we have addressed all comments and have rewritten the paper accordingly. Please find our point-by-point responses below.

Generalization of the theoretical results. We have extended the theoretical results as suggested by Referee 1:

(i) We have derived the following result for the asymptotic power of our test:

Let the conditions of Theorem A.1 be satisfied and consider two sequences of functions $\lambda_{i,T}$ and $\lambda_{j,T}$ with the following property: There exists $\mathcal{I}_k \in \mathcal{F}$ such that

$$\lambda_{i,T}(w) - \lambda_{j,T}(w) \ge c_T \sqrt{\log T/(Th_k)} \quad \text{for all } w \in \mathcal{I}_k,$$
 (1)

where $\{c_T\}$ is any sequence of positive numbers with $c_T \to \infty$ faster than $\frac{\sqrt{\log T}\sqrt{\log\log T}}{\log\log\log T}$. We denote the set of triplets $(i,j,k) \in \mathcal{M}$ for which (1) holds true as \mathcal{M}_1 . Then

$$\mathbb{P}\Big(\forall (i, j, k) \in \mathcal{M}_1 : |\hat{\psi}_{ijk,T}| > c_{T,\text{Gauss}}(\alpha, h_k)\Big) = 1 - o(1)$$

for any given $\alpha \in (0,1)$.

This result is stated as Corollary A.2 on p.?? of the revised paper.

(ii)

Application. We have improved the application section in the following way:

(i) In order to make the data even more comparable across countries, we take the starting date t=1 to be the first Monday after reaching 100 confirmed case in each country. Such alignment of the data by starting on Monday takes into account possible differences in reporting the numbers on a weekly level. As a robustness check, we perform our analysis on the time series without the alignment, where we take the starting date t=1 to be the first day after reaching 100 confirmed case in each country, and we report the results of the robustness check in Section S.3 in the Supplement.

(ii) We have extended the considered time period from T=139 to T=150 and since now the data are available for longer time period, we perform the robustness check for longer time series with T=200 days and report the results in Section S.4 in the Supplement.

Reply to Referee 1

Thank you very much for the constructive and helpful comments. In our revision, we have addressed all of them. Please see our replies to your comments below.

- (1) Since the difference could be canceled out, why should one consider $\sum_t (X_{it} X_{jt}) \mathbf{1}_{t/T \in I_k}$? Isn't it more appropriate to use $|X_{it} X_{jt}|$ or $|X_{it} X_{jt}|^2$ to capture the distance?
- (2) Now the conclusion will be largely interfered by the choice of interval sets. I am wondering whether we can reach some unified result without the influence of such selection. That is whether we can aggregate the rejected intervals Ik and draw some meaningful conclusion?
- (3) The author mentioned this method can be used to identify locations of changes in the trends. But the detail is not very clear to me. For example consider a very simple case: if the two series i, j differ from time t₁ to t₂ and are the same before and after this interval, where t₁, t₂, t₂-t₁ are all unknown. Can we somehow able to identify this interval [t₁, t₂] using our method and how well can we estimate t₁ and t₂? If one takes difference of each pair (i, j), and then the trends are zero except some unknown intervals. Then the task is to detect those unknown intervals. Such problem can be possibly solved by for example MOSUM. Can author comments about this?

(4) Some theory question

- (i) Since the result in Chernozukov et al's Gaussian approximation(GA) does not require the series to be independent cross sectionally, I wonder does that mean the current result can be extended to data with cross-sectional dependence?
- (ii) It would be better if the author can derive power under certain alternatives, so that one can get a better idea as how different the trends needs to be in order to be detected.
 - It can indeed be proven that our multiscale test has asymptotic power 1 against local alternatives. We have added this result as Corollary A.2 in the Appendix to the paper. The proof is provided in the Supplementary Material.
- (iii) The argument about no need for time dependent data is reasonable, just a short comment: there already exists result extending Chernozukov et al's GA to time dependent case, maybe this paper can be further extended to time dependent data as well.
- (5) The specific allowance of p = |W|, which is essential in high dimensional analysis, is not mentioned until appendix. Please put them forward in the main context to

provide some guidance in application. Also since the convergence speed of Gaussian approximation depends on T, p, it would be better to keep the bound in terms of those parameters, so that we know how large the sample size we need in order to obtain the desired accuracy.

Reply to Referee 2

Thank you very much for the constructive and useful suggestions. In our revision, we have addressed all of them. Here are our point-by-point responses to your comments.

- (1) The assumption of independence across countries may be debatable, but it seems that in the context of the model, this could be tested, so this may be worth mentioning.
- (2) Some arguments may be worth further details in the text. For instance, the equation involving $\widehat{s}_{ijk,T}/\sqrt{Th_k}$ on Page 2 Line 7, or the bound for $|r_{it}|$ on Page 2 Line -5.
- (3) I am unsure why the statistic in (3.2) is introduced, I feel the discussion in Pages 8-9 could be done without referring to it.
 - We have deferred the introduction of this test statistics to the Appendix.
- (4) The "cp." abbreviation is uncommon, I feel it should be replaced by "see" or "e.g." We have replaced the "cp." abbreviation by "e.g." and in some cases by "see" throughout the paper.