Quantitative Methods for Finance

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QMF2_3_Modelling_Regression

Estimating a Linear Regression Model with OLS

This activity introduces ordinary least squares (OLS) linear regression in MATLAB using a tiny, didactic house-pricing dataset. We'll keep the workflow practical.

(A short theoretical background is provided at the end of the Live Script.)

Workflow (the basic one):

- Load a small table of house prices and basic features
- Fit a linear model with fitlm
- Inspect fit quality with the built-in diagnostic plot
- Make a point prediction for a new house profile using predict

Dataset

HousePrices is a table file (XLSX) with a numeric response variable Price and a couple of predictors (Size and a simple indicator feature).

This example shows the typical workflow for linear regression analysis using fitlm. The general workflow should include preparing a data set, fitting a linear regression model, evaluating and improving the fitted model, and predicting response values for new predictor data.

In the present example we perform only the very basic steps

```
% Import data from spreadsheet
HousePrices2 = readtable("C:\Users\mdolf\OneDrive\Documents\Didattica
2026\London\Job\GitHub Materials\HousePrices.xlsx");

% Display results
HousePrices2
```

HousePrices2 = 47×3 table

| | Sfeet | Bedrooms | Price |
|---|-------|----------|--------|
| 1 | 2104 | 3 | 399900 |

| | Sfeet | Bedrooms | Price |
|----|-------|----------|--------|
| 2 | 1600 | 3 | 329900 |
| 3 | 2400 | 3 | 369000 |
| 4 | 1416 | 2 | 232000 |
| 5 | 3000 | 4 | 539900 |
| 6 | 1985 | 4 | 299900 |
| 7 | 1534 | 3 | 314900 |
| 8 | 1427 | 3 | 198999 |
| 9 | 1380 | 3 | 212000 |
| 10 | 1494 | 3 | 242500 |
| 11 | 1940 | 4 | 239999 |
| 12 | 2000 | 3 | 347000 |
| 13 | 1890 | 3 | 329999 |
| 14 | 4478 | 5 | 699900 |

:

Fit an OLS linear model using the table; by default the **last column** is treated as the response (set 'ResponseVar', 'Price' to be explicit).

model=fitlm(HousePrices)

model =
Linear regression model:
 Price ~ 1 + Sfeet + Bedrooms

Estimated Coefficients:

| | Estimate | SE | tStat | pValue |
|-------------|------------|--------|---------|------------|
| | | | | |
| (Intercept) | 1.0293e+05 | 52587 | 1.9573 | 0.056669 |
| Sfeet | 110.01 | 17.521 | 6.2784 | 1.3143e-07 |
| Bedrooms | 6948.8 | 19233 | 0.36129 | 0.71961 |

Number of observations: 47, Error degrees of freedom: 44

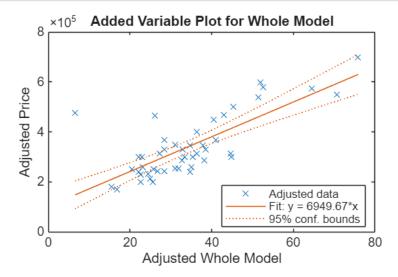
Root Mean Squared Error: 8.33e+04

R-squared: 0.576, Adjusted R-Squared: 0.556

F-statistic vs. constant model: 29.9, p-value = 6.44e-09

Show the built-in diagnostic plot: with **one predictor it's data + fitted line; with multiple predictors it's an added-variable** (partial regression) plot.

plot(model)



Predict for a new case; when you pass a **matrix/row**, **columns must match the model's predictor order**. (Using a table with named variables avoids order mistakes.)

```
estimatedPrice= predict(model,[5000,1]);
fprintf('Estimated price: $%.0f\n', estimatedPrice);
```

Estimated price: \$659908

Theoretical background (example of unconstrained quadratic optimization).

Model and assumptions

Let $y \in \mathbb{R}^n$ be the response (or regressand), $X \in \mathbb{R}^{n \times p}$ the matrix of regressors, $\beta \in \mathbb{R}^p$ the coefficients, and $\varepsilon \in \mathbb{R}^n$ the errors.

$$y = X\beta + \varepsilon$$
, $\mathbb{E}[\varepsilon] = 0$, $\operatorname{Var}(\varepsilon) = \sigma^2 I_n$.

OLS minimizes the residual sum of squares (RSS):

$$RSS(\beta) = \|y - X\beta\|_2^2 = (y - X\beta)^{\mathsf{T}}(y - X\beta) = y^{\mathsf{T}}y - 2\beta^{\mathsf{T}}X^{\mathsf{T}}y + \beta^{\mathsf{T}}X^{\mathsf{T}}X\beta.$$

Since $X^TX \ge 0$, $RSS(\beta)$ is a convex quadratic in β . If rank(X) = p, the minimizer is unique.

Setting the gradient to zero yields the equations:

$$X^{\mathsf{T}}X\,\hat{\beta} = X^{\mathsf{T}}y \quad \Rightarrow \quad \hat{\beta} = (X^{\mathsf{T}}X)^{-1}X^{\mathsf{T}}y \quad (X^{\mathsf{T}}X \text{ is invertible under Gauss-Markov assumptions}).$$

Under Gauss-Markov assumptions:

$$\widehat{\sigma}^2 = \frac{\mathrm{RSS}(\widehat{\beta})}{n-p}, \qquad \mathrm{Var}(\widehat{\beta}) = \sigma^2(X^\top X)^{-1} \approx \widehat{\sigma}^2(X^\top X)^{-1}.$$

"Linear" refers to linearity in the coefficients, not necessarily in the raw predictors. A degree-*d* polynomial regression augments predictors:

$$\phi(x) = [1, x, x^2, \dots, x^d],$$

and fits $y = \phi(x)\beta + \varepsilon$ by OLS. The model is linear in β even though it is nonlinear in x. (In practice, center/scale predictors before adding higher-order terms to mitigate multicollinearity.)

RESET (linearity/specification)

To probe functional-form misspecification, augment the baseline model with powers of the fitted values, e.g. \hat{y}^2 , \hat{y}^3 , and test their joint significance. For nested models M_0 (baseline) and M_1 (augmented) with $p_0 < p_1$ RSS₀, RSS₁, the usual nested F-statistic is parameters and residual sums of squares

$$F = \frac{(RSS_0 - RSS_1)/(p_1 - p_0)}{RSS_1/(n - p_1)}.$$

A small *p*-value suggests misspecification/nonlinearity in the baseline specification.