

## Chapter 5: Large Random Samples

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Fall 2023



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# 1. Introduction: What Statistics Does

- **Statistics** aims to draw conclusions about a population using information from a sample.
- An **experiment** is a process that generates data —that is, a sample.
- A **sample** is a set of observed values of one or more random variables, obtained from one or several repetitions of the experiment.
- To make valid **inferences** about the population, we need to know how likely it is to observe a given sample.
- This, in turn, requires understanding the **probability distributions of the random variables** that generated the data.



# 1. Introduction: Probability and Approximations

- In practice, knowing the exact probability distribution is often difficult:
  - The calculations may be complex.
  - The number of variables involved may be large.
- When we have **large random samples**, we can use powerful **approximation results** to simplify inference:
  - Law of Large Numbers
  - Central Limit Theorem
- These approximations make statistical inference feasible, even when exact probability functions are unknown.



# 1. Introduction

## Proportion of Heads

If you flip a fair coin, you know that the probability of heads is  $\frac{1}{2}$ . However, in a small number of flips, the observed proportion of heads will rarely be exactly  $\frac{1}{2}$ . For instance, if you flip the coin 10 times, it is unlikely to get exactly 5 heads; and if you flip it 100 times, it is even less likely to get exactly 50. We can compute these probabilities using the binomial distribution with parameters  $n$  and  $p = \frac{1}{2}$ . If  $X$  denotes the number of heads in 10 independent flips, then:

$$P(X = 5) = \binom{10}{5} \left(\frac{1}{2}\right)^5 \left(1 - \frac{1}{2}\right)^5 = 0.2461.$$



# 1. Introduction

## Proportion of Heads

If  $Y$  is the number of heads in 100 independent flips, we have

$$\begin{aligned} Pr(Y = 50) &= \binom{100}{50} \left(\frac{1}{2}\right)^{50} \left(1 - \frac{1}{2}\right)^{50} \\ &= 0.0796. \end{aligned}$$

Even though the probability of exactly  $\frac{n}{2}$  heads in  $n$  flips is quite small, especially for large  $n$ , you still expect the proportion of heads to be close to  $\frac{1}{2}$  if  $n$  is large.



# 1. Introduction

## Proportion of Heads

For example, if  $n = 10$ , the proportion of heads is  $\frac{X}{10}$ . In this case, the probability that the proportion is within 0.1 of  $\frac{1}{2}$  is

$$\begin{aligned} P\left(0.4 \leq \frac{X}{10} \leq 0.6\right) &= P(4 \leq X \leq 6) \\ &= \sum_{i=4}^6 \binom{10}{i} \left(\frac{1}{2}\right)^i \left(1 - \frac{1}{2}\right)^{10-i} \end{aligned}$$





# 1. Introduction

## Proportion of Heads

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$$P\left(0.4 \leq \frac{X}{10} \leq 0.6\right) = P(4 \leq X \leq 6)$$

$$= \sum_{i=4}^6 \binom{10}{i} \left(\frac{1}{2}\right)^i \left(1 - \frac{1}{2}\right)^{10-i}$$

$$= B(10, 0.5, x = 6) - B(10, 0.5, x = 3) = 0.828 - 0.172 = 0.6563$$

For  $n = 100$ , the **probability that the sample proportion is within 0.1 of  $\frac{1}{2}$**  increases to approximately **0.965**.



# 1. Introduction

## Queuing Time

Consider a queue serving customers, where the waiting time of the  $i$ -th customer is a random variable  $X_i$ . Suppose  $X_1, X_2, \dots$  are i.i.d. random variables following a uniform distribution on the interval  $[0, 1]$ . The mean waiting time is therefore 0.5.

Intuitively, the average waiting time over many customers should be close to this mean. However, for any finite sample size  $n > 1$ , the exact distribution of the sample average  $\bar{X}$  is quite complicated. It is generally difficult to compute precisely the probability that the sample average is close to 0.5 when  $n$  is large.

So, even when each  $X_i$  has a simple uniform distribution, finding the exact distribution of  $\bar{X}_n$  quickly becomes computationally cumbersome.



# 1. Introduction

In these cases, when sample is large we will be able to use:

- The **law of large numbers**: to show that the average of a large sample of i.i.d. random variables should be close to their mean.
- The **central limit theorem**: to approximate the probability distribution function of large random samples.



# A Random Variable vs. An Observation

## Definition

A **random variable (RV)** is a rule that assigns a numerical value to each possible outcome of a random experiment. It is a *mathematical object*.

- It represents uncertainty: it represents the outcome of a random process *before* we actually observe it.
- It has a **probability distribution** describing how likely each outcome is.

## Definition

An **observation** (or **realization**) is the actual value taken by a random variable once the experiment has been performed.

- It is the **observed outcome** of a random variable.
- It is a fixed number, not random.



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## 2. The Law of Large Numbers

- The average of a random sample of i.i.d. random variables is called the **sample mean**.
- The sample mean provides a summary of the sample, just as the **expected value** summarizes the information contained in a probability distribution of a random variable.
- In this section, we explore the connection between the sample mean and the expected value of the underlying random variables that make up the sample.



## 2. The Law of Large Numbers

- The **Law of Large Numbers (LLN)** states that:  
*“If an experiment is repeated independently many times and the results are averaged, the average will be close to the expected value.”*
- There are two main versions of the LLN:
  - The *Weak Law of Large Numbers (WLLN)*
  - The *Strong Law of Large Numbers (SLLN)*
- The distinction between the two is mostly theoretical —both express the idea that sample averages converge to expected values as the sample size grows.
- In what follows, we focus on the **Weak Law of Large Numbers**.
- But before that, let us define the *sample mean*.



## 2. The Law of Large Numbers

**Definition.** For i.i.d. random variables  $X_1, X_2, \dots, X_n$ , the **sample mean**, denoted by  $\bar{X}$ , is defined as

$$\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}.$$

Since each  $X_i$  is a random variable, the sample mean  $\bar{X}$  is also a random variable. In particular, its expected value is:

$$\begin{aligned} E[\bar{X}] &= \frac{E[X_1] + E[X_2] + \dots + E[X_n]}{n} \\ &= \frac{nE[X]}{n} && \text{(since } E[X_i] = E[X] \text{ for all } i) \\ &= E[X]. \end{aligned}$$

*Note:* All random variables  $X_1, \dots, X_n$  are drawn from the same probability distribution and share the same expected value.





## 2. The Law of Large Numbers

Let us now look at the **variance** of the sample mean  $\bar{X}$ .

Recall that for any random variable  $X$ ,

$$\text{Var}(X) = E[(X - E[X])^2].$$

Before we compute the variance of  $\bar{X}$ , let us first review how variance behaves under scaling. Specifically, we will find the variance of  $aX$ , where  $a$  is a constant.



## 2. The Law of Large Numbers

To find how variance behaves when a random variable is multiplied by a constant  $a$ :

$$\begin{aligned}\text{Var}(aX) &= E[(aX - E[aX])^2] \\ &= E[(aX - aE[X])^2] \\ &= E[a^2(X - E[X])^2] \\ &= a^2 E[(X - E[X])^2] \\ &= a^2 \text{Var}(X).\end{aligned}$$

**Conclusion:** Multiplying a random variable by a constant  $a$  scales its variance by  $a^2$ .



## 2. The Law of Large Numbers

The variance of the sample mean  $\bar{X}$  is given by:

$$\text{Var}(\bar{X}) = \text{Var}\left(\frac{X_1 + X_2 + \cdots + X_n}{n}\right)$$

$$\text{Var}(\bar{X}) = \frac{\text{Var}(X_1 + X_2 + \cdots + X_n)}{n^2} \quad (\text{since } \text{Var}(aX) = a^2 \text{Var}(X))$$

$$= \frac{\text{Var}(X_1) + \text{Var}(X_2) + \cdots + \text{Var}(X_n)}{n^2}$$

(independent  $X_i$ : covariances are zero)

$$= \frac{n \text{Var}(X)}{n^2}$$

(since all  $X_i$  have the same

$$= \frac{\text{Var}(X)}{n}.$$

*Note:* The variance of the sample mean decreases with the sample size —as  $n$  grows,  $\bar{X}$  becomes less variable and concentrates around  $E[X]$ .



## 2. The Law of Large Numbers

Let us now state and prove the **Weak Law of Large Numbers (WLLN)**.

**Theorem (WLLN).** Let  $X_1, X_2, \dots, X_n$  be i.i.d. random variables with finite expected value  $E[X_i] = \mu < \infty$ . Then, for any  $\epsilon > 0$ ,

$$\lim_{n \rightarrow \infty} P(|\bar{X} - \mu| \geq \epsilon) = 0.$$

**Interpretation:** As the sample size  $n$  increases, the probability that the sample mean  $\bar{X}$  deviates from the population mean  $\mu$  by more than  $\epsilon$  approaches zero.

*In other words, the sample mean converges in probability to the expected value.*



## 2. The Law of Large Numbers

### Proof of the Weak Law of Large Numbers

To prove the WLLN, we will use two key results:

- *Markov's Inequality*
- *Tchebysheff's Inequality*

**Markov's Inequality.** Let  $X$  be a random variable that takes only nonnegative values. Then, for any  $a > 0$ ,

$$P(X \geq a) \leq \frac{E[X]}{a}.$$

**Interpretation:** Markov's inequality bounds the probability that a nonnegative random variable exceeds a multiple of its mean. For example:

$$P(X \geq 2E[X]) \leq \frac{1}{2}, \quad P(X \geq kE[X]) \leq \frac{1}{k}.$$



## 2. The Law of Large Numbers

### Markov Example

Suppose the average grade on an upcoming Probability exam is  $E[X] = 12$ . What is the maximum possible proportion of students who can score at least 15?

$$P(X \geq 15) \leq \frac{E[X]}{15} = \frac{12}{15} = \frac{4}{5}.$$

So, at most 80% of students could possibly score this high. However, to achieve this average, we would need a very extreme distribution:

$\frac{4}{5}$  of the class scoring exactly 15, and  $\frac{1}{5}$  scoring 0.

*This illustrates how Markov's bound can be very loose in realistic cases.*



## 2. The Law of Large Numbers

### Markov Example 2

Consider a random variable  $X$  that takes the value 0 with probability  $\frac{24}{25}$  and the value 5 with probability  $\frac{1}{25}$ .

$$E[X] = \frac{24}{25} \cdot 0 + \frac{1}{25} \cdot 5 = \frac{1}{5}.$$

Let's use Markov's inequality to find an upper bound on the probability that  $X \geq 5$ :

$$P(X \geq 5) \leq \frac{E[X]}{5} = \frac{1/5}{5} = \frac{1}{25}.$$

But this is exactly the true probability that  $X = 5$ ! *In this case, Markov's inequality is exact —we say it is **tight**.*



## 2. The Law of Large Numbers

**From Markov to Tchebysheff:** Let  $Y = (X - E[X])^2$ . Then  $E[Y] = \text{Var}(X)$ . By Markov's inequality, we have:

$$P(Y \geq a^2) \leq \frac{E[Y]}{a^2} = \frac{\text{Var}(X)}{a^2}.$$

Notice that the event  $Y = (X - E[X])^2 \geq a^2$  is equivalent to  $|X - E[X]| \geq a$ . Hence, we obtain:

$$P(|X - E[X]| \geq a) \leq \frac{\text{Var}(X)}{a^2}.$$

*This result is known as 'Tchebysheff's inequality'.*





## 2. The Law of Large Numbers

### Tchebysheff's inequality:

$$P(|X - E[X]| \geq a) \leq \frac{\text{Var}(X)}{a^2}.$$

Tchebysheff's inequality provides a bound on the probability that a random variable deviates from its expected value. If we set  $a = k\sigma$ , where  $\sigma$  is the standard deviation, then:

$$P(|X - \mu| \geq k\sigma) \leq \frac{\text{Var}(X)}{k^2\sigma^2} = \frac{1}{k^2}.$$

*Example:* At least 75% of the probability mass lies within  $2\sigma$  of the mean, and at least 89% within  $3\sigma$ . [▶ 'Ch2'](#)



## 2. The Law of Large Numbers

**Proof of the Weak Law of Large Numbers (WLLN).** Apply Tchebysheff's inequality to the sample mean  $\bar{X}$ , using  $a = \epsilon$ :

$$P(|\bar{X} - \mu| \geq \epsilon) \leq \frac{\text{Var}(\bar{X})}{\epsilon^2}.$$

Since  $\text{Var}(\bar{X}) = \frac{\text{Var}(X)}{n}$ , we obtain:

$$P(|\bar{X} - \mu| \geq \epsilon) \leq \frac{\text{Var}(X)}{n\epsilon^2},$$

which clearly goes to zero as  $n \rightarrow \infty$ . **Conclusion:**

$$\lim_{n \rightarrow \infty} P(|\bar{X} - \mu| \geq \epsilon) = 0.$$

*In words:* as the sample size increases, the sample mean becomes arbitrarily close to the population mean.



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### 3. The Central Limit Theorem

Many phenomena observed in the real world can be described reasonably well by a normal probability distribution. In such cases, it is often assumed that the observable random variables in a random sample  $X_1, X_2, \dots, X_n$  are independent and identically distributed, each following a **normal distribution**. We have seen that the sample mean

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

has expectation  $E[\bar{X}] = \mu$  and variance  $\text{Var}(\bar{X}) = \frac{\text{Var}(X)}{n}$ . Next, let us determine the **distribution** of  $\bar{X}$ .



### 3. The Central Limit Theorem

If  $X_1, X_2, \dots, X_n$  are drawn from a normal distribution with mean  $\mu$  and variance  $\sigma^2$ , then their sample mean

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

is also normally distributed, with

$$E[\bar{X}] = \mu, \quad \text{Var}(\bar{X}) = \frac{\sigma^2}{n}.$$

This property holds for *any* sample size  $n$ : when the underlying distribution is normal, the mean of the sample is itself normal.



### 3. The Central Limit Theorem

We can standardize the sample mean to express it in standard normal units:

$$Z = \frac{\bar{X} - \mu_{\bar{X}}}{\sigma_{\bar{X}}} = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \sqrt{n} \frac{\bar{X} - \mu}{\sigma}.$$

The random variable  $Z$  follows a **standard normal distribution**.

*Interpretation:* when sampling from a normal population, we can measure how far the sample mean deviates from the population mean in units of its standard deviation.



### 3. The Central Limit Theorem

#### Bottling Machine Example

A bottling machine fills bottles with an average of  $\mu$  cl of liquid. The amount dispensed per bottle follows a normal distribution with standard deviation  $\sigma = 1$  cl.

A random sample of  $n = 9$  filled bottles is selected from the production line. What is the probability that the sample mean will be within 0.3 cl of the true mean  $\mu$ ?



### 3. The Central Limit Theorem

#### Bottling Machine Example (continued)

Let  $X_1, X_2, \dots, X_9$  denote the amounts filled in each bottle. Since each  $X_i \sim \mathcal{N}(\mu, \sigma^2)$  with  $\sigma = 1$ , we know that the sample mean

$$\bar{X} = \frac{1}{9} \sum_{i=1}^9 X_i$$

is also normally distributed, with

$$E[\bar{X}] = \mu, \quad \text{Var}(\bar{X}) = \frac{\sigma^2}{n} = \frac{1}{9}.$$

We are interested in:

$$P(|\bar{X} - \mu| \leq 0.3) = P[-0.3 \leq (\bar{X} - \mu) \leq 0.3].$$





### 3. The Central Limit Theorem

#### Bottling Machine Example (continued)

Standardizing with

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}},$$

we obtain:

$$\begin{aligned} P(|\bar{X} - \mu| \leq 0.3) &= P\left(\frac{-0.3}{1/\sqrt{9}} \leq Z \leq \frac{0.3}{1/\sqrt{9}}\right) \\ &= P(-0.9 \leq Z \leq 0.9). \end{aligned}$$

From the standard normal table:

$$P(-0.9 \leq Z \leq 0.9) = 1 - 2P(Z > 0.9) = 1 - 2(0.1841) = 0.6318.$$

*Interpretation:* There is about a 63% chance that the sample mean falls within 0.3 cl of the true mean.



### 3. The Central Limit Theorem

#### Bottling Machine Example

How many observations should be included in the sample if we want the sample mean  $\bar{X}$  to be within 0.3 cl of the true mean  $\mu$  with probability 0.95?

We want

$$P(|\bar{X} - \mu| \leq 0.3) = P[-0.3 \leq (\bar{X} - \mu) \leq 0.3] = 0.95.$$



### 3. The Central Limit Theorem

#### Bottling Machine Example (continued)

$$P(|\bar{X} - \mu| \leq 0.3) = P[-0.3 \leq (\bar{X} - \mu) \leq 0.3] = 0.95.$$

Dividing by the standard deviation of the sample mean,  $\sigma_{\bar{X}} = \frac{\sigma}{\sqrt{n}}$  (with  $\sigma = 1$ ), we obtain:

$$P(-0.3\sqrt{n} \leq Z \leq 0.3\sqrt{n}) = 0.95,$$

where  $Z$  follows a standard normal distribution. We know that

$$P(-z_0 \leq Z \leq z_0) = 0.95 \quad \Rightarrow \quad z_0 = 1.96.$$

$$\text{Therefore, } 0.3\sqrt{n} = 1.96 \quad \Rightarrow \quad n = \left(\frac{1.96}{0.3}\right)^2 = 42.68.$$

**Hence,**  $n \approx 43$  observations are required.



### 3. The Central Limit Theorem

- When we sample from a normal population, the sample mean  $\bar{X}$  follows a normal sampling distribution.
- But what happens if the underlying variables  $X_i$  are *not* normally distributed?
- Fortunately, even in that case, the distribution of  $\bar{X}$  becomes approximately normal as the sample size  $n$  increases.
- This remarkable result is known as the **Central Limit Theorem (CLT)**.



### 3. The Central Limit Theorem

**The Central Limit Theorem.** Let  $X_1, X_2, \dots, X_n$  be independent and identically distributed random variables with  $E[X_i] = \mu$  and  $\text{Var}(X_i) = \sigma^2 < \infty$ . Define the standardized sample mean as:

$$Z_n = \sqrt{n} \frac{\bar{X} - \mu}{\sigma}, \quad \text{where} \quad \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i.$$

Then, as  $n \rightarrow \infty$ , the distribution of  $Z_n$  converges to the standard normal distribution:

$$\lim_{n \rightarrow \infty} P(Z_n \leq x) = \Phi(x) \quad \text{for all } x \in \mathbb{R},$$

where  $\Phi(x)$  is the cumulative distribution function (CDF) of the standard normal variable.



### 3. The Central Limit Theorem

#### Test Scores Example

Achievement test scores of all high school seniors in a state have a mean of  $\mu = 60$  and a variance of  $\sigma^2 = 64$ . A random sample of  $n = 100$  students from one large high school has a mean score of  $\bar{X} = 58$ . Is there evidence to suggest that this high school is performing below the state average? *Compute the probability that the sample mean is at most 58 when  $n = 100$ .*



### 3. The Central Limit Theorem

#### Test Scores Example

Let  $\bar{X}$  denote the sample mean of  $n = 100$  test scores from a population with  $\mu = 60$  and  $\sigma^2 = 64$  (so  $\sigma = 8$ ). We want to approximate:  $P(\bar{X} \leq 58)$ . Since

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$$

is approximately standard normal, we have:

$$P(\bar{X} \leq 58) = P\left(Z \leq \frac{58 - 60}{8/\sqrt{100}}\right) = P(Z \leq -2.5).$$

From the standard normal table,  $P(Z \leq -2.5) = 0.0062$ .

**Interpretation:** The probability is very small, providing strong evidence that this high school's average score is below the state average.



*The End*





# Tchebysheff's Theorem

- In certain scenarios, empirical rule may not provide useful approximations.
- Tchebysheff's theorem offers a lower bound for the probability of  $Y$  being within an interval  $\mu \pm k\sigma$ .
- **Tchebysheff's theorem**

$$P(|Y - \mu| < k\sigma) \geq 1 - \frac{1}{k^2}$$

$$P(|Y - \mu| \geq k\sigma) \leq \frac{1}{k^2}$$

- The theorem:
  - Is valid for any probability distribution.
  - Provides conservative estimates.
  - Doesn't contradict empirical rule (verify!).