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- Many types of data can take any value in some interval (sometimes all) of the real numbers.
- Here, the probability density function for discrete random variables is not enough because
 - the number of possible outcomes is uncountable, so we can't just add up all probabilities
 - the probability of any particular value on the continuum typically has to be zero.
- We have to deal with this type of random variables separately from the discrete case.





Definition 1 A random variable Y has a continuous distribution if Y can take on any values in some interval -bounded or unbounded of the real line.

- We can "discretize" the distribution by putting the possible values the random variable can take into "bins"
- i.e. instead of looking at the probabilities P(Y = y), we'll look at probabilities for intervals, i.e. $P(y_1 \le Y \le y_2)$.
- Then, we can plot the bins into a histogram

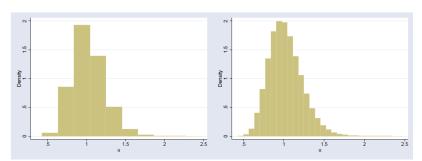




1. Introduction - Definition

Introduction

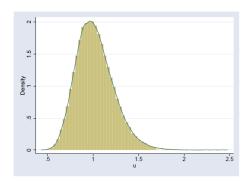
Histograms of the same Distribution for 10 and 30 Bins, respectively







Histogram with 60 Bins and Continuous Density







We can compute

Introduction

$$P(y_j \le Y \le y_k) = \sum_{i=j+1}^k P(y_{i-1} \le Y \le y_i)$$

- We can make the intervals of the histogram finer and finer until we get to the integral of a function
- In the end, we need to compute the area below a function in an interval [a,b] (integral)

$$F(y) = P(a \le y \le b) = \int_a^b f(y)dx$$





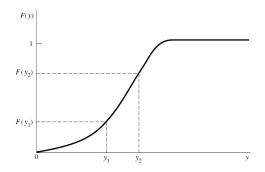
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Definition 2 A random variable Y with distribution function (CDF) F(y) is said to be continuous if F(y) is continuous, for $-\infty < y < \infty$.







What does it mean to have P(Y = y) = 0 ?

• If this were not true and $P(Y=y_0)=p_0>0$, then F(y) would have a discontinuity (jump), violating the continuity assumption.

Rainfall

Consider the example of measuring daily rainfall. What is the probability that we will see a daily rainfall measurement of exactly 2.193 cm? It is quite likely that we would never observe that exact value even if we took rainfall measurements for a lifetime, although we might see many days with measurements between 2 and 3 cm.



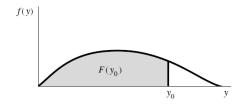


The probability density function PDF is the derivative of F(y):

$$f(y) = \frac{dF(y)}{dy} = F'(y)$$

It then follows that

$$F(y) = \int_{-\infty}^{y} f(t)dt$$







Properties:

The pdf must satisfy that:

Positive probability

$$f(y) \ge 0 \qquad \forall y \in \mathbb{R}$$

Add up to 1

$$\int_{-\infty}^{\infty} f(y)dy = 1$$

Note that for any $Y \in \mathbb{R}$, P(Y = y) = 0





Properties of a CDF:

The CDF must satisfy that:

- $P(\infty) \equiv \lim_{y \to \infty} F(y) = 1.$
- **③** F(y) is a nondecreasing function of y. [If y_1 and y_2 are any values such that $y_1 < y_2$, then $F(y_1) ≤ F(y_2)$.]

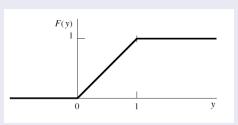




Numerical example 1

Find the PDF of

$$F(y) = \begin{cases} 0 & \text{if } y < 0 \\ y & \text{if } 0 \le y \le 1 \\ 1 & \text{if } y > 1 \end{cases}$$



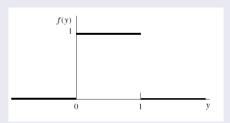


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Numerical example 1

We need to derivate F(y)

$$f(y) = \begin{cases} 0 & \text{if } y < 0 \\ 1 & \text{if } 0 \le y \le 1 \\ 0 & \text{if } y > 1 \end{cases}$$

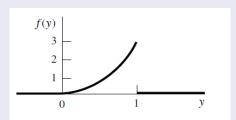




Numerical example 2

Find F(y)

$$f(y) = \begin{cases} 3y^2 & \text{if } 0 \le y \le 1\\ 0 & \text{otherwise} \end{cases}$$





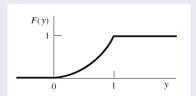


Numerical example 2

We need to integrate f(y) Now to integrate

$$F(y) = \int_0^y 3t^2 dt = t^3]_0^y = y^3$$

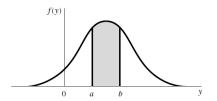
$$F(y) = \begin{cases} 0 & \text{if } y < 0\\ y^3 & \text{if } 0 \le y \le 1\\ 1 & \text{if } y > 1 \end{cases}$$





Here is how we can work with Continuous RV: If we want to know the proba that Y falls in a given interval [a,b], we can compute

$$P(Y \in [a, b]) = P(a \le Y \le b) = \int_a^b f(y)dy$$



Here the equality sign does not matter as much as in the discrete case.



Find c

Given

$$f(y) = \begin{cases} cy^2, & \text{if } 0 \le y \le 2\\ 0, & \text{elsewhere} \end{cases}$$

Find the value of c for which f(y) is a valid density function.





Find c

We require a value for c such that

$$F(\infty) = \int_{-\infty}^{\infty} f(y) \, dy = 1$$

Given the function f(y), this can be written as:

$$\int_0^2 cy^2 \, dy = \frac{cy^3}{3} \Big|_0^2 = \frac{8c}{3}.$$

Thus, $\frac{8}{3}c = 1$, and we find that $c = \frac{3}{8}$.





Find c

Find $P(1 \leq Y \leq 2)$ for the previous example. Also find P(1 < Y < 2).





Find c

Find $P(1 \leq Y \leq 2)$ for the previous example. Also find P(1 < Y < 2).

We have:

$$P(1 \le Y \le 2) = \int_{1}^{2} f(y) \, dy = \frac{3}{8} \int_{1}^{2} y^{2} \, dy = \frac{3}{8} \left[\frac{y^{3}}{3} \right]_{1}^{2} = \frac{7}{8}.$$

Because Y has a continuous distribution, it follows that:

$$P(Y = 1) = P(Y = 2) = 0$$

and, therefore, that:

$$P(1 < Y < 2) = P(1 \le Y \le 2) = \frac{7}{8}.$$



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Sometimes, it is difficult to find the PDF of a continuous RV. We can then use its moments:

Definition 3 The expected value of a continuous RV Y is

$$E(Y) = \int_{-\infty}^{\infty} y f(y) dy \tag{1}$$

- ullet f(y)dy corresponds to p(y) for the discrete case
- integration corresponds to summation
- Hence, E(Y) is also a *mean*





As in the discrete case...

ullet We can compute the expected value of a function g(Y)

$$E(g(Y)) = \int_{-\infty}^{\infty} g(y)f(y)dy$$
 (2)

- \bullet E(c) = c
- $\bullet \ E[cg(Y)] = cE[g(Y)]$
- $E[g_1(Y) + g_2(Y) + \dots + g_k(Y)] = E[g_1(Y)] + E[g_2(Y)] + \dots + E[g_k(Y)]$





If, Y has density function

$$f(y) = \begin{cases} \frac{1}{2}(2-y), & 0 \le y \le 2, \\ 0, & \text{elsewhere,} \end{cases}$$

find the mean and variance of Y.





Mean of Y:

$$E(Y) = \int_{-\infty}^{\infty} y f(y) dy$$

For the given range:

$$\int_0^2 y \left(\frac{1}{2}(2-y)\right) dy = \frac{1}{2} \int_0^2 (2y - y^2) dy = \frac{1}{2} \left[y^2 - \frac{1}{3}y^3\right]_0^2$$
$$= \frac{1}{2} \left[4 - \frac{8}{3}\right] = \frac{1}{2} \left[\frac{4}{3}\right] = \frac{2}{3}$$





The variance is given by (slide 30, CH2):

$$\sigma^2 = E(Y^2) - [E(Y)]^2$$

First, compute $E(Y^2)$:

$$E(Y^2) = \int_0^2 y^2 \left(\frac{1}{2}(2-y)\right) dy = \frac{1}{2} \int_0^2 (2y^2 - y^3) dy$$

$$E(Y^2) = \frac{1}{2} \left[\frac{2}{3} y^3 - \frac{1}{4} y^4 \right]_0^2 = \frac{1}{2} \left[\frac{16}{3} - 4 \right] = \frac{1}{2} \left[\frac{4}{3} \right] = \frac{2}{3}$$





The variance is given by (slide 30, CH2):

$$\sigma^2 = E(Y^2) - [E(Y)]^2$$

Now, using the formulas:

$$Var(Y) = E[Y^2] - (E[Y])^2$$

$$Var(Y) = \frac{2}{3} - \left(\frac{2}{3}\right)^2$$

$$Var(Y) = \frac{2}{3} - \frac{4}{9}$$

$$Var(Y) = \frac{2}{9}$$



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Let a < b be integers. Suppose that the value of a random variable Y is equally likely to be each of the integers a,...,b. Then we say that Y has the uniform distribution on the integers a,...,b. Definition 4 A random variable Y is **uniformly** distributed on the interval [a,b],a < b, if it has the probability density function

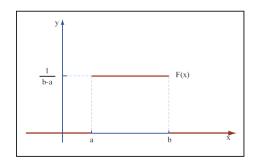
$$f(y) = \begin{cases} \frac{1}{b-a} & \text{if } a \le y \le b\\ 0 & \text{otherwise} \end{cases}$$

We write $Y \sim U(a, b)$



4. The Uniform Distribution

p.d.f for a Uniform Random Variable, $Y \sim U(a,b)$







c.d.f. of a uniform distribution

If $Y \sim U[0,1]$, then the c.d.f. is

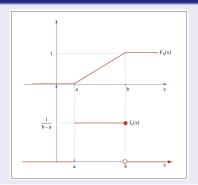
$$F(y) = \begin{cases} 0 & \text{if } y < 0 \\ y & \text{if } 0 \le y \le 1 \\ 1 & \text{if } y \ge 1 \end{cases}$$





4. The Uniform Distribution

c.d.f. of a uniform distribution







4. The Uniform Distribution

Uniform distribution

For example, if $Y \sim U[0, 10]$, can you find $P(3 \le Y \le 4)$?





Uniform distribution

For example, if $Y \sim U[0, 10]$, then, its p.d.f. is

$$f(y) = \frac{1}{b-a} = \frac{1}{10-0} = \frac{1}{10}$$

Then we can find

$$P(3 \le Y \le 4) = \int_3^4 \frac{1}{10} dy = \left[\frac{y}{10}\right]_3^4 = \frac{4}{10} - \frac{3}{10} = \frac{1}{10}$$





4. The Uniform Distribution

Checkout counter

It is known that, during a given 30-minute period, one customer arrived at a checkout counter. Find the probability that the customer arrived during the last 5 minutes of the 30-minute period. The actual time of arrival follows a uniform distribution over the interval of (0, 30).

First, what is the pdf?





Checkout counter

It is known that, during a given 30-minute period, one customer arrived at a checkout counter. Find the probability that the customer arrived during the last 5 minutes of the 30-minute period. The actual time of arrival follows a uniform distribution over the interval of (0, 30). If Y denotes the arrival time, then

$$P(25 \le Y \le 30) = \int_{25}^{30} \frac{1}{30} dy = \frac{30 - 25}{30} = \frac{5}{30} = \frac{1}{6}$$





4. The Uniform Distribution

Expected value of a Uniform distribution

$$\mu = E(Y) = \frac{b+a}{2}$$

Note that the mean is simply the mid-value between the two parameters.

Variance of a Uniform distribution

$$\sigma^2 = V(Y) = \frac{(a-b)^2}{12}$$





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Many measurements are closely approximated by a normal distribution (or bell-shaped).

Definition 5 A random variable Y is normally distributed if the density function of Y is

$$f(y) = \frac{e^{(y-\mu)^2/2\sigma^2}}{\sigma\sqrt{2\pi}} \tag{3}$$

It contains 2 parameters μ and σ such that

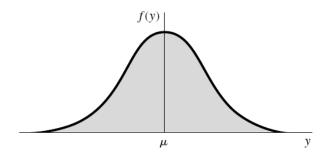
$$E(Y) = \mu$$
 and $V(Y) = \sigma^2$

We write $Y \sim N(\mu, \sigma)$





The parameter μ is located at the center of the distribution and σ measures its spread. It is symmetric with respect to μ .





But DON'T WORRY, we will not integrate the complicated expression of f(y) to obtain F(Y). We will use an approximation presented in next slide's Table.

We use the standardized normal distribution Z, having $Z \sim N(0,1)$.

Next Table show all F(Y) values for each z point in the random variable Z.





Table 4 Normal Curve Areas Standard normal probability in right-hand tail (for negative values of z, areas are found by symmetry)



	Second decimal place of z									
z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
0.0	.5000	.4960	.4920	.4880	.4840	.4801	.4761	.4721	.4681	.4641
0.1	.4602	.4562	.4522	.4483	.4443	.4404	.4364	.4325	.4286	.4247
0.2	.4207	.4168	.4129	.4090	.4052	.4013	.3974	.3936	.3897	.3859
0.3	.3821	.3783	.3745	.3707	.3669	.3632	.3594	.3557	.3520	.3483
0.4	.3446	.3409	.3372	.3336	.3300	.3264	.3228	.3192	.3156	.3121
0.5	.3085	.3050	.3015	.2981	.2946	.2912	.2877	.2843	.2810	.2776
0.6	.2743	.2709	.2676	.2643	.2611	.2578	.2546	.2514	.2483	.2451
0.7	.2420	.2389	.2358	.2327	.2296	.2266	.2236	.2206	.2177	.2148
0.8	.2119	.2090	.2061	.2033	.2005	.1977	.1949	.1922	.1894	.1867
0.9	.1841	.1814	.1788	.1762	.1736	.1711	.1685	.1660	.1635	.1611
1.0	.1587	.1562	.1539	.1515	.1492	.1469	.1446	.1423	.1401	.1379
1.1	.1357	.1335	.1314	.1292	.1271	.1251	.1230	.1210	.1190	.1170
1.2	.1151	.1131	.1112	.1093	.1075	.1056	.1038	.1020	.1003	.0985
1.3	.0968	.0951	.0934	.0918	.0901	.0885	.0869	.0853	.0838	.0823
1.4	.0808	.0793	.0778	.0764	.0749	.0735	.0722	.0708	.0694	.0681
1.5	.0668	.0655	.0643	.0630	.0618	.0606	.0594	.0582	.0571	.0559
1.6	.0548	.0537	.0526	.0516	.0505	.0495	.0485	.0475	.0465	.0455
1.7	.0446	.0436	.0427	.0418	.0409	.0401	.0392	.0384	.0375	.0367
1.8	.0359	.0352	.0344	.0336	.0329	.0322	.0314	.0307	.0301	.0294
1.9	.0287	.0281	.0274	.0268	.0262	.0256	.0250	.0244	.0239	.0233





A Normal example

Let Z denote a normal random variable with mean 0 and standard deviation 1.

- Find P(Z > 2).
- ② Find $P(-2 \le Z \le 2)$.
- **3** Find $P(0 \le Z \le 1.73)$.





A Normal example

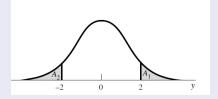
• Find P(Z>2). Since $\mu=0$ and $\sigma=1$, the value 2 is actually z=2. Proceed down the first (z) column in Table 4, and read the area opposite z=2.0. This area, denoted by the symbol A(z), is A(2.0)=.0228. Thus, P(Z>2)=.0228.





A Normal example

2 Find $P(-2 \le Z \le 2)$.



In part (1) we determined that $A_1=A(2.0)=.0228$. Because the density function is symmetric about the mean, it follows that $A_2=A_1=.0228$ and hence that

$$P(-2 \le Z \le 2) = 1 - A1 - A2 = 1 - 2(.0228) = .9544$$





A Normal example

③ Find $P(0 \le Z \le 1.73)$. Because P(Z > 0) = A(0) = .5, we obtain that $P(0 \le Z \le 1.73) = .5 - A(1.73)$, where A(1.73) is obtained by proceeding down the z column in Table 4, to the entry 1.7 and then across the top of the table to the column labeled .03 to read A(1.73) = .0418. Thus,

$$P(0 \le Z \le 1.73) = .5 - .0418 = .4582.$$





We can always transform a normal random variable Y to a standard normal random variable Z by using the relationship

$$Z = \frac{Y - \mu}{\sigma}$$

So we go from $Y \sim N(\mu, \sigma)$ to $Z \sim N(0, 1)$





Test scores

The achievement scores for a college entrance examination are normally distributed with mean 75 and standard deviation 10. What fraction of the scores lies between 80 and 90?





Test scores

Recall that z is the distance from the mean of a normal distribution expressed in units of standard deviation. Thus,

$$z = \frac{y - \mu}{\sigma}$$

Then the desired fraction of the population is given by the area between $z_1 = \frac{80-75}{10} = 0.5$ and $z_2 = \frac{90-75}{10} = 1.5$.

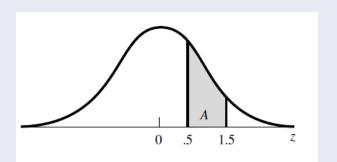
$$A = A(0.5) - A(1.5) = 0.3085 - 0.0668 = 0.2417.$$





Test scores

$$A = A(0.5) - A(1.5) = 0.3085 - 0.0668 = 0.2417.$$







How to integrate

The integral of any polynomial is the sum of the integrals of its terms. A general term of a polynomial can be written as

$$ax^n$$

and the indefinite integral of that term is

$$\int ax^n \, dx = \frac{a}{n+1}x^{n+1} + C$$

where a and C are constants. The expression applies for both positive and negative values of n except for the special case of n=-1. In general, C is set equal to zero.



How to integrate

If definite limits are set for the integration, it is called a definite integral.

The definite integral of any polynomial is the sum of the integrals of its terms. A general term of a polynomial can be written as

$$ax^n$$

and the definite integral of that term is

$$\int_{b}^{c} ax^{n} dx = \left[\frac{a}{n+1} x^{n+1} \right]_{b}^{c} = \frac{a}{n+1} c^{n+1} - \frac{a}{n+1} b^{n+1}$$

where b and c are constants, called the limits of the integral. The procedure is basically the same as in the indefinite integral except for the evaluation at the two limits.