

LATENT VARIABLE MODELING

Session 5: Multiple group analyses

Multigroup SEM

Multi-group SEM allows for assessing parameter differences between groups:

- ▣ Measurement parameters
 - A.k.a. measurement invariance, measurement equivalence, differential item functioning
- ▣ Structural parameters
 - Means and variances of LVs
 - Regression relationships between OV's and/or LV's, e.g.,
 - differential prediction (e.g., intelligence test predicts functioning in job differently among males/females, majority/minority, ...)
 - genetically informative design
 - ...

Contents

- ▣ Measurement invariance: Differences in measurement models between groups
- ▣ Structural invariance: Differences in means, variances or regressions between groups
- ▣ Model fit comparisons

Measurement invariance (MI)

- ▣ Are measurement parameters equal across
 - ▣ two or more groups (between-group MI), or
 - ▣ two or more measurement occasions (longitudinal MI)
- ▣ If not, there is measurement bias (lack of measurement invariance)
 - ▣ Observed score differences do not only reflect true differences in the construct of interest, but also reflect group membership

Multigroup SEM

Categorical variables can be included within a SEM as:

- ▣ Endogenous variables
 - Discussed in last session
- ▣ Exogenous variable
 - Model main effect only:
 - Include (several) 0-1 coded variable as a variable
 - Model possible interaction effects:
 - Use multi-group SEM: Fit same model in each group and compare parameter estimates
 - or: Create interaction(s) before analysis and include as variables in model (not possible for every research question)

(lack of) measurement invariance

- ▣ The observed score on observed variable i (i.e., item or subscale) of person j is given by: $Y_{ij} = \tau_i + \lambda_i \eta_j + \epsilon_{ij}$
- ▣ Therefore: $E(Y_{ij} | \eta_j) = \tau_i + \lambda_i \eta_j$
- ▣ If intercepts or loadings differ between groups

$$\tau_{ig} \neq \tau_{ig'} \text{ or } \lambda_{ig} \neq \lambda_{ig'}$$

and

$$E(Y_{ijg} | \eta_j) \neq E(Y_{ijg'} | \eta_j)$$
- ▣ Thus, given the same latent trait value, we would expect a different item score for a person in group g , than a person in group g'
- ▣ We say: X_i is a biased indicator of η with respect to group
 - ▣ In other words: differences in item scores reflect group membership, in addition to true score differences

(lack of) measurement invariance

- The observed score on item or subscale i , of person j is given by $Y_{ij} = \tau_i + \lambda_i \eta_j + \epsilon_{ij}$
- By definition, ϵ_i follows a normal distribution with mean 0 and variance $\sigma_{\epsilon_i}^2$
- When variance of measurement error differs over groups
 - ▣ No systematic bias in observed scores, but
 - ▣ Construct is not measured with same precision across groups (i.e., different reliability across groups)

Mean structure

- Graphically: mean structure is represented by one or more triangles, which
 - ▣ Denote a constant with a value of 1
 - ▣ Have outgoing, single-headed arrow(s), of which the corresponding coefficient is the value of the intercept
- Algebraically: mean structure is represented by two vectors in lavaan:
 - ▣ \mathbf{v} (nu; contains intercepts of observed variables)
 - ▣ $\mathbf{\alpha}$ (contains intercepts of latent or variables)

(lack of) measurement invariance

- With CFA, we can statistically test whether the parameters of the measurement model are equal across groups
- We subsequently test for equality across groups of:
 1. Pattern of zero and non-zero loadings
 - ▣ 'configural' invariance
 2. In addition to 1: loadings (λ_i 's)
 - ▣ 'metric' or 'weak' invariance
 3. In addition to 2: intercepts (τ_i 's)
 - ▣ 'scalar' or 'strong' invariance
 4. In addition to 3: error variances ($\sigma_{\epsilon_i}^2$'s)
 - ▣ 'uniqueness' or 'strict' invariance

Mean structure

- As discussed earlier, the model-implied covariance matrix is given by

$$\hat{\Sigma} = \Lambda(\mathbf{I} - \beta)^{-1} \Psi [(\mathbf{I} - \beta)^{-1}]^T \Lambda^T + \Theta$$
- The model-implied mean vector is given by

$$\hat{\mu} = \Lambda(\mathbf{\alpha} + \beta\mathbf{\alpha}) + \mathbf{v}$$
- In CFA, there are no structural regression parameters, and the equations simplify to

$$\hat{\Sigma} = \Lambda\Psi\Lambda^T + \Theta$$

$$\hat{\mu} = \Lambda\mathbf{\alpha} + \mathbf{v}$$

(lack of) structural invariance

- We can also test whether structural coefficients are equal across groups:
 - ▣ equality of β (structural or latent regressions)
 - ▣ equality of Ψ (structural or latent (co)variances)
 - ▣ equality of $\mathbf{\alpha}$ (structural or latent means)

Mean structure: Identification

- For identification of the mean structure of a latent variable, we take a similar approach as for identification of the covariance structure:
- ▣ Standardized latent variable: Set intercept of LV to 0 (in addition to setting variance of LV to 1)
 - ▣ Marker variable: Set intercept of an indicator variable to 0 (in addition to setting loading of indicator to 1)

Testing invariance

To test whether a set of parameters (loadings, intercepts, residual variances, latent (co)variances, or latent means) are equal across groups, we fit two models:

1. Model with parameters of interest **estimated freely** in both groups
2. Model with parameters of interest **restricted to equality** across groups
3. Assess difference in fit between models 1 and 2
 - $\chi^2(df)$, CFI, AIC, BIC and/or SSABIC

More restricted model will (almost always) have worse fit, but is it significantly or substantially worse?

Testing invariance

- $\Delta\chi^2$ test often significant with larger sample sizes. Alternatives:
 - Use AIC, BIC or RMSEA (lower value is better model)
 - Use difference in CFI values:
 - Cheung and Rensvold (2000): $\Delta CFI > .01$ indicates that null hypothesis of invariance should be rejected
 - Meade et al. (2008): $\Delta CFI > .002$ indicates that null hypothesis of invariance should be rejected

Testing invariance

- Like model fit, tenability of MI is not an all-or-nothing question, researcher should make informed decision
- Rules-of-thumb offer a good starting point:
 - For evaluating configural invariance (as usual):
Non-significant χ^2 -value; CFI > .95; RMSEA < .06; SRMS < .08 (for good fit; can also use more lenient 'acceptable-fit' criteria)
 - For evaluating metric, scalar and uniqueness invariance:
Non-significant $\Delta\chi^2(df)$, ΔCFI
Model with lowest AIC or BIC value fits best

Examples and exercises

- Example 4.4
- Exercise 4.1
 - Assessing measurement and structural invariance of the WAIS (continuous indicators)

Chi-square difference test

- Statistical significance of difference in fit between two nested models can be assessed using $\Delta\chi^2(\Delta df)$ test

$$\Delta\chi^2 = \chi^2_{\text{model2}} - \chi^2_{\text{model1}}$$

$$\Delta df = df_{\text{model2}} - df_{\text{model1}}$$
- Nested model: all free parameters in less complex model are also free in more complex model
 - More complex model will have χ^2 equal to or lower than that of less complex model, by definition
 - More complex model also has lower df
- $\Delta\chi^2$ tests whether more complex model fits significantly better than less complex model
 - If so, retain more complex model
 - If not, retain less complex model

Reporting your results

- When reporting on your SEM model, you should provide at least two tables:
 - Table with indices of overall model fit indices
 - When doing model comparisons, also report differences in fit between models ($\Delta\chi^2$, Δdf , ΔCFI)
- Example table for presenting the results
- | | χ^2 | df | p | CFI | TLI | RMSEA | BIC | AIC |
|---------|----------|------|-----|-----|-----|-------|-----|-----|
| Model 1 | | | | | | | | |
| Model 2 | | | | | | | | |
| Model 3 | | | | | | | | |
| Model 4 | | | | | | | | |
- Table with parameter estimates (from your final, best-fitting model)
 - Useful examples: see Van de Schoot, Lugtig and Hox (2012); Vandenberg and Lance (2000)

invariance testing with ordered-categorical indicators

- We subsequently test for equality across groups of:
 1. Pattern of zero and non-zero loadings
 - 'configural' invariance
 2. In addition to 1: loadings (λ_i 's)
 - 'metric' or 'weak' invariance
 3. In addition to 2: threshold (τ_i 's)
 - threshold invariance
- Equality of residual variances is not tested
 - Delta parameterization: $\sigma_{\epsilon_i} = 1 - \lambda_i^2$. Thus, test of equal loadings also tests equality of residual variances.
 - Theta parameterization: $\sigma_{\epsilon_i} = 1$. All equal by default.
- Structural invariance tests as in continuous-indicator case:
 - equality of β (structural or latent regressions)
 - equality of Ψ (structural or latent (co)variances)
 - equality of α (structural or latent means)

References

- Cheung, G. W., & Rensvold, R. B. (2002). Evaluating goodness-of-fit indexes for testing measurement invariance. *Structural equation modeling*, 9(2), 233-255.
- Meade, A. W., Johnson, E. C., & Braddy, P. W. (2008). Power and sensitivity of alternative fit indices in tests of measurement invariance. *Journal of Applied Psychology*, 93(3), 568.
- Van de Schoot, R., Lugtig, P., & Hox, J. (2012). A checklist for testing measurement invariance. *European Journal of Developmental Psychology*, 9(4), 486-492.
- Vandenberg, R. J., & Lance, C. E. (2000). A review and synthesis of the measurement invariance literature: Suggestions, practices, and recommendations for organizational research. *Organizational research methods*, 3(1), 4-70.

Including non-linear effects in SEMs

- Lavaan exclusively fits linear SEM models
- If you want to include non-linear effects, you should calculate non-linear transformations of your variables first, then specify these as linear predictors in your SEM
- For example:

```
data <- data.frame(x1, x2)
## Calculate quadratic effect of x1:
data$x1_x1 <- data$x1^2
## Calculate interaction between x1 and x2:
data$x1_x2 <- data$x1*data$x2
## Specify and fit "linear" sem model:
mod <- "y ~ x1 + x1_x1 + x2 + x1_x2"
fit <- sem(mod, data = data)
```

Examples and exercises

- Additional Exercise: Invariance of the HADS
 - ordered categorical indicators
 - adding regressions and interactions to an LVM