

Answers to exercises multigroup LVMs

```
library("lavaan")
```

Exercise 4.1: Measurement invariance between the WAIS and WAIS-IV

We read in the data:

```
WAIS.cor <- lav_matrix_lower2full(c(
  1.00,
  0.31, 1.00,
  0.36, 0.40, 1.00,
  0.51, 0.46, 0.45, 1.00,
  0.29, 0.40, 0.33, 0.43, 1.00,
  0.39, 0.29, 0.27, 0.36, 0.33, 1.00,
  0.32, 0.27, 0.29, 0.33, 0.24, 0.28, 1.00,
  0.22, 0.32, 0.15, 0.22, 0.27, 0.12, 0.26, 1.00
))
WAIS.means <- c(7.83, 5.50, 5.67, 21.50, 7.67, 8.00, 6.50, 34.83)
WAIS.sds <- c(2.69, 1.50, 2.36, 6.06, 1.85, 2.18, 5.97, 9.94)
WAIS.cov <- cor2cov(WAIS.cor, sds=WAIS.sds)

WAISIV.cor <- lav_matrix_lower2full(c(
  1.00,
  0.46, 1.00,
  0.58, 0.55, 1.00,
  0.63, 0.43, 0.73, 1.00,
  0.27, 0.51, 0.37, 0.33, 1.00,
  0.45, 0.38, 0.37, 0.43, 0.13, 1.00,
  0.33, 0.52, 0.49, 0.41, 0.29, 0.43, 1.00,
  0.15, 0.27, 0.16, 0.09, 0.12, 0.25, 0.23, 1.00
))
WAISIV.means <- c(15.17, 15.00, 11.83, 21.67, 12.17, 17.83, 18.67, 45.83)
WAISIV.sds <- c(4.93, 4.10, 5.20, 6.54, 2.72, 5.35, 9.36, 10.44)
WAISIV.cov <- cor2cov(WAISIV.cor, sds=WAISIV.sds)

WAIS.names <- c("Compr", "Arith", "Simil", "Vocab", "DigSpan", "PictCompl",
  "BlockDes", "Cod")

names(WAIS.means) <- names(WAIS.sds) <- colnames(WAIS.cov) <-
  rownames(WAIS.cov) <- names(WAISIV.means) <- names(WAISIV.sds) <-
  rownames(WAISIV.cov) <- colnames(WAISIV.cov) <- WAIS.names

WAIS.cov.list <- list(WAIS.cov, WAISIV.cov)
WAIS.mean.list <- list(WAIS.means, WAISIV.means)
WAIS.n.list <- list(WAIS.n = 200, WAISIV.n = 200)
```

```

WAIS.mod <- '
  ## verbal comprehension
  VC =~ Simil + Vocab + Compr

  ## Perceptual reasoning
  PR =~ PictCompl + BlockDes

  ## Working memory
  WM =~ Arith + DigSpan + Cod
,

```

- a) Considering that both WAIS versions were administered to the same respondents, one would expect their latent verbal comprehension, Perceptual Reasoning and Working Memory scores to be the same between tests. Thus, the latent means and variances can be expected to be equal. Therefore, the standardized LV approach seems the most appropriate for these data.

Note that when equality restrictions on parameters across groups are applied in a model identified using the standardized-LV approach in lavaan, the latent mean and variance will be fixed to 1 and 0, respectively, in the first group. In the second group, the loading and intercept of the first item will be fixed to the respective values in the first group, to identify the model. This allows for differences in the latent mean and variance to be present. For the current example, this may not be completely appropriate, but the current design (testing measurement invariance between two different tests in the same individuals) is quite rare.

- b) We fit the configural invariance model to the datasets:

```

fit.indices <- c("chisq", "df", "pvalue", "cfi", "rmsea", "srmr", "aic", "bic")
WAIS.conf.fit <- cfa(WAIS.mod, sample.cov = WAIS.cov.list,
  sample.mean = WAIS.mean.list, std.lv = TRUE,
  sample.nobs = WAIS.n.list, meanstructure = TRUE)
pars <- parameterestimates(WAIS.conf.fit, standardized = TRUE)
pars[order(pars$op, decreasing = TRUE), c(1:3, 5:7, 9, 13)]

```

##	lhs	op	rhs	group	est	se	pvalue	std.all
## 1	VC	==	Simil	1	1.393	0.170	0.000	0.592
## 2	VC	==	Vocab	1	4.659	0.415	0.000	0.771
## 3	VC	==	Compr	1	1.725	0.190	0.000	0.643
## 4	PR	==	PictCompl	1	1.198	0.191	0.000	0.551
## 5	PR	==	BlockDes	1	3.026	0.509	0.000	0.508
## 6	WM	==	Arith	1	1.018	0.113	0.000	0.681
## 7	WM	==	DigSpan	1	1.147	0.140	0.000	0.621
## 8	WM	==	Cod	1	4.123	0.779	0.000	0.416
## 34	VC	==	Simil	2	4.460	0.313	0.000	0.860
## 35	VC	==	Vocab	2	5.499	0.397	0.000	0.843
## 36	VC	==	Compr	2	3.516	0.318	0.000	0.715
## 37	PR	==	PictCompl	2	3.194	0.398	0.000	0.599
## 38	PR	==	BlockDes	2	6.707	0.713	0.000	0.718
## 39	WM	==	Arith	2	3.761	0.308	0.000	0.920
## 40	WM	==	DigSpan	2	1.495	0.198	0.000	0.551
## 41	WM	==	Cod	2	3.085	0.778	0.000	0.296
## 23	Simil	~1		1	5.670	0.166	0.000	2.409
## 24	Vocab	~1		1	21.500	0.427	0.000	3.557
## 25	Compr	~1		1	7.830	0.190	0.000	2.918
## 26	PictCompl	~1		1	8.000	0.154	0.000	3.679
## 27	BlockDes	~1		1	6.500	0.421	0.000	1.092
## 28	Arith	~1		1	5.500	0.106	0.000	3.676
## 29	DigSpan	~1		1	7.670	0.130	0.000	4.156

## 30	Cod ~1	1	34.830	0.701	0.000	3.513	
## 31	VC ~1	1	0.000	0.000	NA	0.000	
## 32	PR ~1	1	0.000	0.000	NA	0.000	
## 33	WM ~1	1	0.000	0.000	NA	0.000	
## 56	Simil ~1	2	11.830	0.367	0.000	2.281	
## 57	Vocab ~1	2	21.670	0.461	0.000	3.322	
## 58	Compr ~1	2	15.170	0.348	0.000	3.085	
## 59	PictCompl ~1	2	17.830	0.377	0.000	3.341	
## 60	BlockDes ~1	2	18.670	0.660	0.000	2.000	
## 61	Arith ~1	2	15.000	0.289	0.000	3.668	
## 62	DigSpan ~1	2	12.170	0.192	0.000	4.485	
## 63	Cod ~1	2	45.830	0.736	0.000	4.401	
## 64	VC ~1	2	0.000	0.000	NA	0.000	
## 65	PR ~1	2	0.000	0.000	NA	0.000	
## 66	WM ~1	2	0.000	0.000	NA	0.000	
## 9	Simil ~~	Simil	1	3.602	0.419	0.000	0.650
## 10	Vocab ~~	Vocab	1	14.834	2.454	0.000	0.406
## 11	Compr ~~	Compr	1	4.226	0.518	0.000	0.587
## 12	PictCompl ~~	PictCompl	1	3.293	0.451	0.000	0.696
## 13	BlockDes ~~	BlockDes	1	26.304	3.282	0.000	0.742
## 14	Arith ~~	Arith	1	1.202	0.179	0.000	0.537
## 15	DigSpan ~~	DigSpan	1	2.090	0.272	0.000	0.614
## 16	Cod ~~	Cod	1	81.310	8.768	0.000	0.827
## 17	VC ~~	VC	1	1.000	0.000	NA	1.000
## 18	PR ~~	PR	1	1.000	0.000	NA	1.000
## 19	WM ~~	WM	1	1.000	0.000	NA	1.000
## 20	VC ~~	PR	1	0.907	0.106	0.000	0.907
## 21	VC ~~	WM	1	0.837	0.067	0.000	0.837
## 22	PR ~~	WM	1	0.820	0.118	0.000	0.820
## 42	Simil ~~	Simil	2	7.014	1.237	0.000	0.261
## 43	Vocab ~~	Vocab	2	12.323	1.991	0.000	0.290
## 44	Compr ~~	Compr	2	11.823	1.394	0.000	0.489
## 45	PictCompl ~~	PictCompl	2	18.276	2.247	0.000	0.642
## 46	BlockDes ~~	BlockDes	2	42.184	7.143	0.000	0.484
## 47	Arith ~~	Arith	2	2.580	1.642	0.116	0.154
## 48	DigSpan ~~	DigSpan	2	5.127	0.578	0.000	0.696
## 49	Cod ~~	Cod	2	98.930	10.059	0.000	0.912
## 50	VC ~~	VC	2	1.000	0.000	NA	1.000
## 51	PR ~~	PR	2	1.000	0.000	NA	1.000
## 52	WM ~~	WM	2	1.000	0.000	NA	1.000
## 53	VC ~~	PR	2	0.758	0.066	0.000	0.758
## 54	VC ~~	WM	2	0.647	0.062	0.000	0.647
## 55	PR ~~	WM	2	0.751	0.076	0.000	0.751

```
fitMeasures(WAIS.conf.fit, fit.indices)
```

##	chisq	df	pvalue	cfi	rmsea	srmr	aic	bic
##	62.187	34.000	0.002	0.968	0.064	0.037	17617.734	17833.273

Model fit is good according to SRMR and CFI. RMSEA indicates adequate fit. All standardized loadings are substantial in both versions, in both versions Coding is the weakest indicator. The strongest indicator is Vocabulary in the WAIS and Arithmetics in the WAIS-IV. The three factors correlate strongly in both groups, but more strongly in the WAIS than in the WAIS-IV group.

We inspect standardized residuals:

```
residuals(WAIS.conf.fit, type = "cor")
```

```
## $`Group 1`
## $`Group 1`$type
## [1] "cor.bollen"
##
## $`Group 1`$cov
##      Simil  Vocab  Compr  PctCmp BlckDs Arith  DigSpn Cod
## Simil      0.000
## Vocab     -0.006  0.000
## Compr     -0.020  0.015  0.000
## PictCompl -0.025 -0.025  0.069  0.000
## BlockDes   0.017 -0.025  0.024  0.000  0.000
## Arith      0.063  0.021 -0.056 -0.017 -0.014  0.000
## DigSpan    0.022  0.029 -0.044  0.049 -0.019 -0.023  0.000
## Cod       -0.056 -0.048 -0.004 -0.068  0.087  0.037  0.012  0.000
##
## $`Group 1`$mean
##      Simil  Vocab  Compr PictCompl  BlockDes  Arith  DigSpan  Cod
##      0      0      0      0      0      0      0      0
##
##
## $`Group 2`
## $`Group 2`$type
## [1] "cor.bollen"
##
## $`Group 2`$cov
##      Simil  Vocab  Compr  PctCmp BlckDs Arith  DigSpn Cod
## Simil      0.000
## Vocab      0.005  0.000
## Compr     -0.035  0.027  0.000
## PictCompl -0.020  0.047  0.125  0.000
## BlockDes   0.021 -0.049 -0.060  0.000  0.000
## Arith      0.038 -0.072  0.035 -0.033  0.024  0.000
## DigSpan    0.064  0.030  0.015 -0.118 -0.007  0.003  0.000
## Cod       -0.005 -0.072  0.013  0.117  0.070 -0.002 -0.043  0.000
##
## $`Group 2`$mean
##      Simil  Vocab  Compr PictCompl  BlockDes  Arith  DigSpan  Cod
##      0      0      0      0      0      0      0      0
```

WAIS-IV has two residual correlations with absolute values larger than .1: the correlations of Picture Completion with Digit Span and Coding.

All in all, configural invariance appears tenable.

- b) Assess whether loadings, intercepts and residual variances are equal between the two WAIS versions.

When we use the standardized LV identification method, and apply an equality restriction on the loadings between groups, **lavaan** will keep the latent variance of the first group fixed at one, but will freely estimate the latent variance in the second group.

```
WAIS.metr.fit <- cfa(WAIS.mod, sample.cov = WAIS.cov.list,
                    sample.mean = WAIS.mean.list,
                    sample.nobs = WAIS.n.list, std.lv = TRUE,
                    meanstructure = TRUE,
```

```

group.equal = "loadings")
pars <- parameterestimates(WAIS.metr.fit, standardized = TRUE)
pars[pars$op == "~" | pars$rhs %in% c("VC", "PR", "WM"), c(1:3, 5, 7, 14)]

```

```

##   lhs op      rhs group   est std.all
## 1  VC =~      Simil    1 1.831  0.715
## 2  VC =~      Vocab    1 2.742  0.502
## 3  VC =~      Compr    1 1.651  0.618
## 4  PR =~ PictCompl    1 1.262  0.574
## 5  PR =~ BlockDes    1 2.813  0.478
## 6  WM =~      Arith    1 1.295  0.842
## 7  WM =~ DigSpan    1 0.685  0.390
## 8  WM =~      Cod     1 1.617  0.168
## 17 VC ~~      VC      1 1.000  1.000
## 18 PR ~~      PR      1 1.000  1.000
## 19 WM ~~      WM      1 1.000  1.000
## 20 VC ~~      PR      1 0.930  0.930
## 21 VC ~~      WM      1 0.750  0.750
## 22 PR ~~      WM      1 0.683  0.683
## 34 VC =~      Simil    2 1.831  0.806
## 35 VC =~      Vocab    2 2.742  0.878
## 36 VC =~      Compr    2 1.651  0.724
## 37 PR =~ PictCompl    2 1.262  0.582
## 38 PR =~ BlockDes    2 2.813  0.728
## 39 WM =~      Arith    2 1.295  0.834
## 40 WM =~ DigSpan    2 0.685  0.623
## 41 WM =~      Cod     2 1.617  0.387
## 50 VC ~~      VC      2 4.688  1.000
## 51 PR ~~      PR      2 5.934  1.000
## 52 WM ~~      WM      2 6.567  1.000
## 53 VC ~~      PR      2 3.940  0.747
## 54 VC ~~      WM      2 3.713  0.669
## 55 PR ~~      WM      2 4.926  0.789

```

Note that now the latent variance are freely estimated for the WAIS-IV, and are much larger than for the WAIS (where they are fixed to 1).

This may indicate that the measurement precision of the WAIS-IV is much higher than that of the WAIS. It cannot be due to latent differences between the two ‘groups’, because the two tests were administered to the same individuals.

```
fitMeasures(WAIS.metr.fit, fit.indices)
```

```

##      chisq      df    pvalue      cfi      rmsea      srmr      aic      bic
##    132.004    39.000     0.000     0.896     0.109     0.093 17677.551 17873.133

```

```
lavTestLRT(WAIS.conf.fit, WAIS.metr.fit)
```

```

## Chi-Squared Difference Test
##
##           Df   AIC   BIC   Chisq Chisq diff Df diff Pr(>Chisq)
## WAIS.conf.fit 34 17618 17833   62.187
## WAIS.metr.fit 39 17678 17873 132.004    69.817      5 1.119e-13 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

All fit indices indicate a lack of metric invariance.

```
lavTestScore(WAIS.metr.fit)$uni
```

```
##
## univariate score tests:
##
##   lhs op   rhs      X2 df p.value
## 1 .p1. == .p34. 22.090  1  0.000
## 2 .p2. == .p35. 27.669  1  0.000
## 3 .p3. == .p36.  0.052  1  0.819
## 4 .p4. == .p37.  0.680  1  0.410
## 5 .p5. == .p38.  0.680  1  0.410
## 6 .p6. == .p39. 27.254  1  0.000
## 7 .p7. == .p40. 12.023  1  0.001
## 8 .p8. == .p41. 11.753  1  0.001
```

Five out of eight loadings seem non-invariant. Which parameters are the most invariance-offending?

```
pars <- parameterestimates(WAIS.metr.fit)
pars[pars$label %in% c(".p1.", ".p2.", ".p6.", ".p7.", ".p8."), c(1:3, 5:6)]
```

```
##   lhs op   rhs group label
## 1  VC =~ Simil    1 .p1.
## 2  VC =~ Vocab    1 .p2.
## 6  WM =~ Arith    1 .p6.
## 7  WM =~ DigSpan  1 .p7.
## 8  WM =~ Cod      1 .p8.
## 34 VC =~ Simil    2 .p1.
## 35 VC =~ Vocab    2 .p2.
## 39 WM =~ Arith    2 .p6.
## 40 WM =~ DigSpan  2 .p7.
## 41 WM =~ Cod      2 .p8.
```

Vocabulary and Arithmetics seem the worst-offending loadings; let's lift those equality restrictions:

```
WAIS.metr.fit2 <- cfa(WAIS.mod, sample.cov = WAIS.cov.list,
  sample.mean = WAIS.mean.list,
  sample.nobs = WAIS.n.list, std.lv = TRUE,
  auto.fix.first = TRUE,
  meanstructure = TRUE,
  group.equal = "loadings",
  group.partial = c("VC =~ Vocab", "WM =~ Arith"))
fitMeasures(WAIS.metr.fit2, fit.indices)
```

```
##   chisq      df    pvalue      cfi    rmsea      srmr      aic      bic
##   71.647   37.000     0.001    0.961    0.068    0.050 17621.194 17824.759
```

```
lavTestLRT(WAIS.conf.fit, WAIS.metr.fit2)
```

```
## Chi-Squared Difference Test
##
##           Df   AIC   BIC  Chisq Chisq diff Df diff Pr(>Chisq)
## WAIS.conf.fit 34 17618 17833 62.187
## WAIS.metr.fit2 37 17621 17825 71.647     9.4601      3    0.02376 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
lavTestScore(WAIS.metr.fit2)$uni
```

```
##
## univariate score tests:
##
##   lhs op   rhs    X2 df p.value
## 1 .p1. == .p34. 6.010 1   0.014
## 2 .p3. == .p36. 6.010 1   0.014
## 3 .p4. == .p37. 0.605 1   0.437
## 4 .p5. == .p38. 0.605 1   0.437
## 5 .p7. == .p40. 2.423 1   0.120
## 6 .p8. == .p41. 2.423 1   0.120
```

Model fit improved, but not enough.

It may seem odd that we have three identical values for the six remaining modification indices. We lifted two equality restrictions, one for the VC and one for the WM factor. For each latent factor, there are now two equality restrictions left. Releasing any of the two will yield an identical improvement of model fit.

Next, we lift the equality restriction on the loading of Similarity:

```
WAIS.metr.fit3 <- cfa(WAIS.mod, sample.cov = WAIS.cov.list,
  sample.mean = WAIS.mean.list,
  sample.nobs = WAIS.n.list, std.lv = TRUE,
  meanstructure = TRUE,
  group.equal = "loadings",
  group.partial = c("VC =~ Vocab", "WM =~ Arith",
    "VC =~ Simil"))
fitMeasures(WAIS.metr.fit3, fit.indices)
```

##	chisq	df	pvalue	cfi	rmsea	srmr	aic	bic
##	65.316	36.000	0.002	0.967	0.064	0.042	17616.863	17824.419

```
lavTestLRT(WAIS.conf.fit, WAIS.metr.fit3)
```

```
## Chi-Squared Difference Test
##
##           Df   AIC   BIC  Chisq Chisq diff Df diff Pr(>Chisq)
## WAIS.conf.fit  34 17618 17833 62.187
## WAIS.metr.fit3 36 17617 17824 65.316      3.1289      2      0.2092
```

Fit has improved. Next, we restrict item intercepts to equality:

```
WAIS.scal.fit <- cfa(WAIS.mod, sample.cov = WAIS.cov.list,
  sample.mean = WAIS.mean.list, std.lv = TRUE,
  sample.nobs = WAIS.n.list, meanstructure = TRUE,
  group.equal = c("intercepts", "loadings"),
  group.partial = c("VC =~ Vocab", "WM =~ Arith",
    "VC =~ Simil"))
fitMeasures(WAIS.scal.fit, fit.indices)
```

##	chisq	df	pvalue	cfi	rmsea	srmr	aic	bic
##	222.241	41.000	0.000	0.797	0.149	0.117	17763.788	17951.387

```
lavTestLRT(WAIS.metr.fit3, WAIS.scal.fit)
```

```
## Chi-Squared Difference Test
##
##           Df   AIC   BIC  Chisq Chisq diff Df diff Pr(>Chisq)
```

```
## WAIS.metr.fit3 36 17617 17824 65.316
## WAIS.scal.fit 41 17764 17951 222.241 156.93 5 < 2.2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
lavTestScore(WAIS.scal.fit)$uni
```

```
##
## univariate score tests:
##
##      lhs op   rhs      X2 df p.value
## 1   .p3. == .p36. 0.000 1 1.000
## 2   .p4. == .p37. 8.224 1 0.004
## 3   .p5. == .p38. 8.224 1 0.004
## 4   .p7. == .p40. 2.153 1 0.142
## 5   .p8. == .p41. 2.153 1 0.142
## 6  .p23. == .p56. 2.052 1 0.152
## 7  .p24. == .p57. 93.438 1 0.000
## 8  .p25. == .p58. 32.113 1 0.000
## 9  .p26. == .p59. 16.203 1 0.000
## 10 .p27. == .p60. 16.203 1 0.000
## 11 .p28. == .p61. 3.357 1 0.067
## 12 .p29. == .p62. 2.683 1 0.101
## 13 .p30. == .p63. 0.037 1 0.847
```

```
pars <- parameterestimates(WAIS.scal.fit)
pars[pars$label %in% c(".p24.", ".p25.", ".p26.", ".p27."), c(1:3, 5:6)]
```

```
##      lhs op rhs group label
## 24   Vocab ~1      1 .p24.
## 25   Compr ~1      1 .p25.
## 26 PictCompl ~1      1 .p26.
## 27 BlockDes ~1      1 .p27.
## 57   Vocab ~1      2 .p24.
## 58   Compr ~1      2 .p25.
## 59 PictCompl ~1      2 .p26.
## 60 BlockDes ~1      2 .p27.
```

Vocabulary seems to have the worst variance-offending item intercept, we release those restrictions:

```
WAIS.scal.fit2 <- cfa(WAIS.mod, sample.cov = WAIS.cov.list,
  sample.mean = WAIS.mean.list, std.lv = TRUE,
  sample.nobs = WAIS.n.list, meanstructure = TRUE,
  group.equal = c("loadings", "intercepts"),
  group.partial = c("VC =~ Vocab", "WM =~ Arith",
    "VC =~ Simil", "Vocab ~1"))
fitMeasures(WAIS.scal.fit2, fit.indices)
```

```
##      chisq      df    pvalue      cfi      rmsea      srmr      aic      bic
##    99.437    40.000     0.000    0.933    0.086    0.070 17642.984 17834.574
```

```
lavTestLRT(WAIS.scal.fit2, WAIS.metr.fit3)
```

```
## Chi-Squared Difference Test
##
##      Df    AIC    BIC  Chisq Chisq diff Df diff Pr(>Chisq)
## WAIS.metr.fit3 36 17617 17824 65.316
## WAIS.scal.fit2 40 17643 17835 99.436    34.121      4 7.039e-07 ***
```



```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

We should lift additional restrictions:

```
lavTestScore(WAIS.scal.fit2)$uni
```

```
##
## univariate score tests:
##
##      lhs op   rhs      X2 df p.value
## 1   .p3. == .p36.  0.000  1   1.000
## 2   .p4. == .p37.  7.398  1   0.007
## 3   .p5. == .p38.  7.398  1   0.007
## 4   .p7. == .p40.  2.221  1   0.136
## 5   .p8. == .p41.  2.221  1   0.136
## 6  .p23. == .p56. 12.844  1   0.000
## 7  .p25. == .p58. 12.844  1   0.000
## 8  .p26. == .p59. 15.451  1   0.000
## 9  .p27. == .p60. 15.451  1   0.000
## 10 .p28. == .p61.  2.279  1   0.131
## 11 .p29. == .p62.  1.904  1   0.168
## 12 .p30. == .p63.  0.010  1   0.919
```

```
WAIS.scal.fit3 <- cfa(WAIS.mod, sample.cov = WAIS.cov.list,
  sample.mean = WAIS.mean.list, std.lv = TRUE,
  sample.nobs = WAIS.n.list, meanstructure = TRUE,
  group.equal = c("loadings", "intercepts"),
  group.partial = c("VC =~ Vocab", "WM =~ Arith",
    "VC =~ Simil", "Vocab ~1",
    "Compr ~ 1"))
fitMeasures(WAIS.scal.fit3, fit.indices)
```

```
##      chisq      df    pvalue      cfi      rmsea      srmr      aic      bic
##      85.705    39.000      0.000      0.948      0.077      0.064 17631.252 17826.834
```

```
lavTestLRT(WAIS.scal.fit3, WAIS.metr.fit3)
```

```
## Chi-Squared Difference Test
##
##              Df    AIC    BIC   Chisq Chisq diff Df diff Pr(>Chisq)
## WAIS.metr.fit3 36 17617 17824 65.316
## WAIS.scal.fit3 39 17631 17827 85.705      20.389      3  0.000141 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
lavTestScore(WAIS.scal.fit3)$uni
```

```
##
## univariate score tests:
##
##      lhs op   rhs      X2 df p.value
## 1   .p3. == .p36.  0.000  1   1.000
## 2   .p4. == .p37.  7.359  1   0.007
## 3   .p5. == .p38.  7.359  1   0.007
## 4   .p7. == .p40.  2.228  1   0.136
## 5   .p8. == .p41.  2.228  1   0.136
## 6  .p23. == .p56.  0.000  1   1.000
```

```
## 7 .p26. == .p59. 16.705 1 0.000
## 8 .p27. == .p60. 16.705 1 0.000
## 9 .p28. == .p61. 2.243 1 0.134
## 10 .p29. == .p62. 1.875 1 0.171
## 11 .p30. == .p63. 0.011 1 0.915
```

```
WAIS.scal.fit4 <- cfa(WAIS.mod, sample.cov = WAIS.cov.list,
  sample.mean = WAIS.mean.list, std.lv = TRUE,
  sample.nobs = WAIS.n.list, meanstructure = TRUE,
  group.equal = c("loadings", "intercepts"),
  group.partial = c("VC =~ Vocab", "WM =~ Arith",
    "VC =~ Simil", "Vocab ~1",
    "Compr ~ 1", "PictCompl ~ 1"))
fitMeasures(WAIS.scal.fit4, fit.indices)
```

```
##      chisq      df    pvalue    cfi    rmsea    srmr      aic      bic
## 67.775 38.000 0.002 0.967 0.063 0.046 17615.322 17814.895
```

```
lavTestLRT(WAIS.scal.fit4, WAIS.metr.fit3)
```

```
## Chi-Squared Difference Test
```

```
##
##           Df   AIC   BIC  Chisq Chisq diff Df diff Pr(>Chisq)
## WAIS.metr.fit3 36 17617 17824 65.316
## WAIS.scal.fit4 38 17615 17815 67.775      2.4588      2      0.2925
```

Fit is acceptable. We continue with testing equality of measurement error variances:

```
WAIS.uni.fit <- cfa(WAIS.mod, sample.cov = WAIS.cov.list,
  sample.mean = WAIS.mean.list, std.lv = TRUE,
  sample.nobs = WAIS.n.list, meanstructure = TRUE,
  group.equal = c("loadings", "intercepts", "residuals"),
  group.partial = c("VC =~ Vocab", "WM =~ Arith",
    "VC =~ Simil", "Vocab ~1",
    "Compr ~ 1", "PictCompl ~ 1"))
```

```
## Warning in lav_model_estimate(lavmodel = lavmodel, lavpartable = lavpartable, :
## lavaan WARNING: the optimizer warns that a solution has NOT been found!
```

We get a warning about negative LV variances, which indicates that restricting all measurement error variances to equality is not a good idea.

Considering the large number of non-invariant parameters, measurement invariance does not seem tenable. I would rather go back and inspect the parameter estimates of the configural invariant model:

```
pars <- parameterestimates(WAIS.conf.fit, standardized = TRUE)
col_names <- c("lhs", "op", "rhs", "group", "est", "se", "pvalue", "std.all")
pars[pars$op == "=", col_names]
```

```
##    lhs op      rhs group  est   se pvalue std.all
## 1 VC =~ Simil    1 1.393 0.170    0 0.592
## 2 VC =~ Vocab    1 4.659 0.415    0 0.771
## 3 VC =~ Compr    1 1.725 0.190    0 0.643
## 4 PR =~ PictCompl 1 1.198 0.191    0 0.551
## 5 PR =~ BlockDes 1 3.026 0.509    0 0.508
## 6 WM =~ Arith    1 1.018 0.113    0 0.681
## 7 WM =~ DigSpan  1 1.147 0.140    0 0.621
## 8 WM =~ Cod      1 4.123 0.779    0 0.416
```

```
## 34 VC =~ Simil 2 4.460 0.313 0 0.860
## 35 VC =~ Vocab 2 5.499 0.397 0 0.843
## 36 VC =~ Compr 2 3.516 0.318 0 0.715
## 37 PR =~ PictCompl 2 3.194 0.398 0 0.599
## 38 PR =~ BlockDes 2 6.707 0.713 0 0.718
## 39 WM =~ Arith 2 3.761 0.308 0 0.920
## 40 WM =~ DigSpan 2 1.495 0.198 0 0.551
## 41 WM =~ Cod 2 3.085 0.778 0 0.296
```

Standardized loadings are higher for the VC and PR indicators in the WAIS-IV, compared to the WAIS. For the WM indicators, we see that in the WAIS, the indicators have similar standardized loadings, while in the WAIS-IV, Arithmetics has a much higher (standardized) loading than the other two subtests.

```
pars[pars$op == "~" & !pars$lhs %in% c("VC", "WM", "PR"), col_names]
```

```
##      lhs op      rhs group  est    se pvalue std.all
## 9      Simil ~ Simil    1  3.602  0.419  0.000  0.650
## 10     Vocab ~ Vocab    1 14.834  2.454  0.000  0.406
## 11     Compr ~ Compr    1  4.226  0.518  0.000  0.587
## 12 PictCompl ~ PictCompl 1  3.293  0.451  0.000  0.696
## 13 BlockDes ~ BlockDes  1 26.304  3.282  0.000  0.742
## 14     Arith ~ Arith    1  1.202  0.179  0.000  0.537
## 15 DigSpan ~ DigSpan    1  2.090  0.272  0.000  0.614
## 16      Cod ~ Cod      1 81.310  8.768  0.000  0.827
## 42     Simil ~ Simil    2  7.014  1.237  0.000  0.261
## 43     Vocab ~ Vocab    2 12.323  1.991  0.000  0.290
## 44     Compr ~ Compr    2 11.823  1.394  0.000  0.489
## 45 PictCompl ~ PictCompl 2 18.276  2.247  0.000  0.642
## 46 BlockDes ~ BlockDes  2 42.184  7.143  0.000  0.484
## 47     Arith ~ Arith    2  2.580  1.642  0.116  0.154
## 48 DigSpan ~ DigSpan    2  5.127  0.578  0.000  0.696
## 49      Cod ~ Cod      2 98.930 10.059  0.000  0.912
```

We observe the reverse pattern for the residual variances: Standardized residual variances are higher for the VC and PR indicators in the WAIS, compared to the WAIS-IV. For the WM indicators, we see that in the WAIS, the indicators have more similar standardized residual variances, while in the WAIS-IV, Arithmetics has a much lower (standardized) loading than the other two subtests.

```
pars[pars$op == "~1" & !pars$lhs %in% c("VC", "WM", "PR"), col_names]
```

```
##      lhs op rhs group  est    se pvalue std.all
## 23     Simil ~1      1  5.67 0.166    0  2.409
## 24     Vocab ~1      1 21.50 0.427    0  3.557
## 25     Compr ~1      1  7.83 0.190    0  2.918
## 26 PictCompl ~1      1  8.00 0.154    0  3.679
## 27 BlockDes ~1      1  6.50 0.421    0  1.092
## 28     Arith ~1      1  5.50 0.106    0  3.676
## 29 DigSpan ~1      1  7.67 0.130    0  4.156
## 30      Cod ~1      1 34.83 0.701    0  3.513
## 56     Simil ~1      2 11.83 0.367    0  2.281
## 57     Vocab ~1      2 21.67 0.461    0  3.322
## 58     Compr ~1      2 15.17 0.348    0  3.085
## 59 PictCompl ~1      2 17.83 0.377    0  3.341
## 60 BlockDes ~1      2 18.67 0.660    0  2.000
## 61     Arith ~1      2 15.00 0.289    0  3.668
## 62 DigSpan ~1      2 12.17 0.192    0  4.485
```

```
## 63          Cod ~1          2 45.83 0.736          0 4.401
```

The intercepts seem lower for the WAIS than for the WAIS-IV, indicating that the the subtests of the WAIS-IV may be easier. Perhaps the tests have different numbers of items and different scoring rules. But the intercept differences should be interpreted relative to the indicator item.

c)

```
pars[pars$op == "~" & pars$lhs %in% c("VC", "WM", "PR"), col_names]
```

##	lhs	op	rhs	group	est	se	pvalue	std.all
## 17	VC	~	VC	1	1.000	0.000	NA	1.000
## 18	PR	~	PR	1	1.000	0.000	NA	1.000
## 19	WM	~	WM	1	1.000	0.000	NA	1.000
## 20	VC	~	PR	1	0.907	0.106	0	0.907
## 21	VC	~	WM	1	0.837	0.067	0	0.837
## 22	PR	~	WM	1	0.820	0.118	0	0.820
## 50	VC	~	VC	2	1.000	0.000	NA	1.000
## 51	PR	~	PR	2	1.000	0.000	NA	1.000
## 52	WM	~	WM	2	1.000	0.000	NA	1.000
## 53	VC	~	PR	2	0.758	0.066	0	0.758
## 54	VC	~	WM	2	0.647	0.062	0	0.647
## 55	PR	~	WM	2	0.751	0.076	0	0.751

We see stronger correlations between the latent factors in the WAIS, than in the WAIS-IV.

```
pars[pars$op == "~1" & pars$lhs %in% c("VC", "WM", "PR"), col_names]
```

##	lhs	op	rhs	group	est	se	pvalue	std.all
## 31	VC	~1		1	0	0	NA	0
## 32	PR	~1		1	0	0	NA	0
## 33	WM	~1		1	0	0	NA	0
## 64	VC	~1		2	0	0	NA	0
## 65	PR	~1		2	0	0	NA	0
## 66	WM	~1		2	0	0	NA	0

All in all, there does not seem to be measurement invariance between the WISC and the WISC-IV. Thus, observed subtest scores are not comparable between the two test versions. However, the results do indicate that the WAIS-IV is an improvement over the WAIS: It has better measurement precision, and the latent factors seem more separate traits than they were in the WAIS.