

Exercises ordered categorical indicator variables

Exercise 6.1

Bartholomew, Steele, Galbraith, and Moustaki (2008) analyzed four items from the British Social Attitudes Survey concerning abortion. The item responses from 379 respondents are available in the Abortion data from the ltm package. For each item, respondents were to indicate yes (1) or no (0) on whether abortion should be allowed. We will rename the items I1-I4.

```
library(ltm)
names(Abortion) <- c(paste0("I", 1:4))
```

Hint: use 'ordered = paste0("I", 1:4)' to declare the items as ordered categorical in using the cfa() function.

a) Find the proportion who endorsed each item (i.e., the mean score).

```
colMeans(Abortion)
```

```
##          I1          I2          I3          I4
## 0.4379947 0.5936675 0.6358839 0.6174142
```

b) Fit a CFA for binary responses using the CFA function, assuming a single latent variable underlies the item responses.

```
library(lavaan)
model <- '
  Theta =~ I1 + I2 + I3 + I4
'
write.table(Abortion, file = "Abortion.txt")
fit.abo <- cfa(model, data = Abortion, std.lv = TRUE,
               ordered = paste0("I", 1:4))
summary(fit.abo, standardized = TRUE)
```

```
## lavaan (0.6-1) converged normally after 13 iterations
##
##   Number of observations              379
##
##   Estimator                        DWLS      Robust
##   Model Fit Test Statistic          7.291    12.647
##   Degrees of freedom                 2         2
##   P-value (Chi-square)              0.026     0.002
##   Scaling correction factor          0.587
##   Shift parameter                   0.234
##   for simple second-order correction (Mplus variant)
##
## Parameter Estimates:
##
##   Information                      Expected
##   Information saturated (h1) model  Unstructured
##   Standard Errors                  Robust.sem
##
## Latent Variables:
##
##           Estimate  Std.Err  z-value  P(>|z|)  Std.lv  Std.all
##   Theta =~
##     I1           0.921    0.022   42.552   0.000    0.921    0.921
##     I2           0.940    0.021   44.737   0.000    0.940    0.940
##     I3           0.964    0.019   50.568   0.000    0.964    0.964
```

```

##      I4              0.905      0.025      35.507      0.000      0.905      0.905
##
## Intercepts:
##      Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##      .I1      0.000
##      .I2      0.000
##      .I3      0.000
##      .I4      0.000
##      Theta    0.000
##
## Thresholds:
##      Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##      I1|t1     0.156      0.065      2.410      0.016      0.156      0.156
##      I2|t1    -0.237      0.065     -3.639      0.000     -0.237     -0.237
##      I3|t1    -0.347      0.066     -5.273      0.000     -0.347     -0.347
##      I4|t1    -0.299      0.066     -4.559      0.000     -0.299     -0.299
##
## Variances:
##      Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##      .I1      0.151
##      .I2      0.117
##      .I3      0.071
##      .I4      0.182
##      Theta    1.000
##
## Scales y*:
##      Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##      I1      1.000
##      I2      1.000
##      I3      1.000
##      I4      1.000

```

c) Evaluate overall model fit.

```
fitmeasures(fit.abo)
```

```

##      npar      fmin
##      8.000      0.010
##      chisq      df
##      7.291      2.000
##      pvalue      chisq.scaled
##      0.026      12.647
##      df.scaled      pvalue.scaled
##      2.000      0.002
##      chisq.scaling.factor      baseline.chisq
##      0.587      4919.479
##      baseline.df      baseline.pvalue
##      6.000      0.000
##      baseline.chisq.scaled      baseline.df.scaled
##      3905.848      6.000
##      baseline.pvalue.scaled      baseline.chisq.scaling.factor
##      0.000      1.260
##      cfi      tli
##      0.999      0.997
##      nnfi      rfi

```

##	0.997	0.996
##	nfi	pnfi
##	0.999	0.333
##	ifi	rni
##	0.999	0.999
##	cfi.scaled	tli.scaled
##	0.997	0.992
##	cfi.robust	tli.robust
##	NA	NA
##	nnfi.scaled	nnfi.robust
##	0.992	NA
##	rfi.scaled	nfi.scaled
##	0.990	0.997
##	ifi.scaled	rni.scaled
##	0.997	0.997
##	rni.robust	rmsea
##	NA	0.084
##	rmsea.ci.lower	rmsea.ci.upper
##	0.025	0.153
##	rmsea.pvalue	rmsea.scaled
##	0.145	0.119
##	rmsea.ci.lower.scaled	rmsea.ci.upper.scaled
##	0.062	0.185
##	rmsea.pvalue.scaled	rmsea.robust
##	0.025	NA
##	rmsea.ci.lower.robust	rmsea.ci.upper.robust
##	NA	NA
##	rmsea.pvalue.robust	rmr
##	NA	0.025
##	rmr_nomean	srmr
##	0.029	0.029
##	srmr_bentler	srmr_bentler_nomean
##	0.025	0.029
##	srmr_bollen	srmr_bollen_nomean
##	0.025	0.029
##	srmr_mplus	srmr_mplus_nomean
##	0.025	0.029
##	cn_05	cn_01
##	311.626	478.508
##	gfi	agfi
##	0.999	0.993
##	pgfi	mfi
##	0.200	0.993

Inspect the estimated thresholds and loadings to answer the following questions:

d) If you would have to create a 1-item abortion attitude test, which item would you select?

The item with the highest discrimination parameter: Item 3.

e) If the 1-item test has to be used to find persons with extremely liberal views on abortion, which item would you select?

The item with the highest threshold (difficulty): Item 1

f) Looking at the discrimination parameters (loadings) and their standard errors, would you expect the Rasch or 2pl model to fit better?

The loadings are quite similar, they differ about 1 SE amongst each other, so the differences do not seem statistically significant. Therefore, the assumption that all loadings are equal seems tenable, and the Rasch model is probably more appropriate.

g) Statistically test whether the Rasch or 2pl model fits better.

```
model.rasch <- '
  Theta =~ 1*I1 + 1*I2 + 1*I3 + 1*I4
'

fit.rasch <- cfa(model.rasch, data = Abortion, ordered = paste0("I", 1:4))
fitinds <- c("cfi.scaled", "rmsea.scaled", "srmr")
fitMeasures(fit.abo, fitinds)
```

```
##   cfi.scaled rmsea.scaled      srmr
##         0.997         0.119      0.029
```

```
fitMeasures(fit.rasch, fitinds)
```

```
##   cfi.scaled rmsea.scaled      srmr
##         0.998         0.067      0.040
```

```
lavTestLRT(fit.rasch, fit.abo)
```

```
## Scaled Chi Square Difference Test (method = "satorra.2000")
##
##           Df AIC BIC   Chisq Chisq diff Df diff Pr(>Chisq)
## fit.abo    2           7.291
## fit.rasch  5          10.171      3.3525      3      0.3404
```

The chi-square difference test indicates no significant difference between the fit of the two models. Then we prefer the most parsimonious model, in this case the Rasch model.

Exercise 6.2

Beaujean and Sheng (2010) conducted an IRT analysis of the ten-item vocabulary test from the General Social Survey. Data from the respondents with responses to all 10 items ($n = 2943$) from the 2000 decade group are available as a space delimited file (gss2000.dat), and the items are named word.a-word.j. Get the file gss2000.dat from the github repository. To load it in R, type:

```
gssdat <- read.table("gss2000.dat", header = TRUE)
```

Hint: use following code in cfa() function: ordered = paste0("word.", letters[1:3])

a) Conduct an item-level confirmatory factor analysis with one latent variable. Analyze only the first four items, as analyzing all 10 will involve a lot of typing.

```
gssmod <- '
  vocab =~ word.a + word.b + word.c + word.d
'

gssfit <- cfa(gssmod, ordered = paste0("word.", letters[1:4]), data = gssdat)
summary(gssfit, standardized = TRUE, fit.measures = TRUE)
```

```
## lavaan (0.6-1) converged normally after 23 iterations
##
##   Number of observations              2943
##
##   Estimator                        DWLS      Robust
##   Model Fit Test Statistic          4.014      5.149
##   Degrees of freedom                  2          2
```

```

##      P-value (Chi-square)                0.134      0.076
##      Scaling correction factor            0.786
##      Shift parameter                      0.044
##      for simple second-order correction (Mplus variant)
##
## Model test baseline model:
##
##      Minimum Function Test Statistic      673.679    625.493
##      Degrees of freedom                   6          6
##      P-value                             0.000      0.000
##
## User model versus baseline model:
##
##      Comparative Fit Index (CFI)          0.997      0.995
##      Tucker-Lewis Index (TLI)            0.991      0.985
##
##      Robust Comparative Fit Index (CFI)    NA
##      Robust Tucker-Lewis Index (TLI)      NA
##
## Root Mean Square Error of Approximation:
##
##      RMSEA                               0.019      0.023
##      90 Percent Confidence Interval        0.000 0.045    0.000 0.049
##      P-value RMSEA <= 0.05               0.978      0.959
##
##      Robust RMSEA                         NA
##      90 Percent Confidence Interval        0.000      NA
##
## Standardized Root Mean Square Residual:
##
##      SRMR                               0.027      0.027
##
## Parameter Estimates:
##
##      Information                        Expected
##      Information saturated (h1) model    Unstructured
##      Standard Errors                    Robust.sem
##
## Latent Variables:
##
##      Estimate  Std.Err  z-value  P(>|z|)  Std.lv  Std.all
##      vocab =~
##      word.a      1.000
##      word.b      1.921    0.222    8.640    0.000    0.466    0.466
##      word.c      0.678    0.117    5.814    0.000    0.896    0.896
##      word.d      0.678    0.117    5.814    0.000    0.316    0.316
##      word.d      1.705    0.174    9.794    0.000    0.796    0.796
##
## Intercepts:
##
##      Estimate  Std.Err  z-value  P(>|z|)  Std.lv  Std.all
##      .word.a    0.000
##      .word.b    0.000
##      .word.c    0.000
##      .word.d    0.000
##      vocab      0.000
##      .word.a    0.000
##      .word.b    0.000
##      .word.c    0.000
##      .word.d    0.000
##      vocab      0.000
##

```

```
## Thresholds:
##           Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## word.a|t1    -1.061   0.029  -37.203   0.000  -1.061  -1.061
## word.b|t1    -1.473   0.035  -42.114   0.000  -1.473  -1.473
## word.c|t1     0.614   0.025   24.813   0.000   0.614   0.614
## word.d|t1    -1.649   0.039  -42.209   0.000  -1.649  -1.649
##
## Variances:
##           Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## .word.a       0.782           0.782   0.782
## .word.b       0.197           0.197   0.197
## .word.c       0.900           0.900   0.900
## .word.d       0.367           0.367   0.367
## vocab         0.218   0.040   5.451   0.000   1.000   1.000
##
## Scales y*:
##           Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## word.a        1.000           1.000   1.000
## word.b        1.000           1.000   1.000
## word.c        1.000           1.000   1.000
## word.d        1.000           1.000   1.000
```

Model fit is perfect (by definition, there are only three indicators), standardized loadings are substantial.

b) What are the easiest and most difficult items?

Easiest item is word.d, most difficult item is word.c.

c) What are the best and worst indicators of the latent trait?

Best indicator is word.b, worst indicator is word.c.

d) Does the Rasch, or the 2pl model fit the 3 vocabulary items better?

```
gssmod.rasch <- '
  vocab =~ a*word.a + a*word.b + a*word.c + a*word.d
'
gssfit.rasch <- cfa(gssmod.rasch, ordered = paste0("word.", letters[1:4]),
  data = gssdat)
fitMeasures(gssfit, fitinds)
```

```
## cfi.scaled rmsea.scaled      srmr
##      0.995      0.023      0.027
```

```
fitMeasures(gssfit.rasch, fitinds)
```

```
## cfi.scaled rmsea.scaled      srmr
##      0.767      0.099      0.142
```

```
lavTestLRT(gssfit.rasch, gssfit)
```

```
## Scaled Chi Square Difference Test (method = "satorra.2000")
```

```
##
##           Df AIC BIC      Chisq Chisq diff Df diff Pr(>Chisq)
## gssfit         2      4.0139
## gssfit.rasch    5     137.0874      135.42      3 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The fit of the 2pl model is significantly better than that of the Rasch model. So for this vocabulary test, maybe the 2pl model should be preferred. However, the sample size is very large, so we have a lot of power to detect differences which may in fact be small.

Additional exercise: HADS

```
library(foreign)
HADS <- read.spss("HADS.sav", use.value.labels = TRUE, to.data.frame = TRUE)
summary(HADS)
```

```
## Respondentnummer    leeftijd          geslacht          HADS1
## Min.    :500002    Min.    :18.00    een man   :217    bijna nooit : 43
## 1st Qu.:500162    1st Qu.:35.00    een vrouw :285    soms       :160
## Median :500333    Median :43.00                                vaak        :202
## Mean    :500335    Mean    :42.84                                bijna altijd: 97
## 3rd Qu.:500512    3rd Qu.:51.00
## Max.    :500689    Max.    :80.00
##
##          HADS2          HADS3          HADS4
## bijna nooit :214    bijna nooit : 75    bijna altijd: 31
## soms        :151    soms        :175    vaak         : 81
## vaak        :103    vaak        :180    soms        :219
## bijna altijd: 34    bijna altijd: 72    bijna nooit  :171
##
##
##          HADS5          HADS6          HADS7
## bijna nooit :179    bijna nooit : 67    bijna nooit :199
## soms        :170    soms        :204    soms        :187
## vaak        :116    vaak        :167    vaak        :101
## bijna altijd: 37    bijna altijd: 64    bijna altijd: 15
##
##
```

a) Fit a graded response model to the data:

```
HADS.GRM.mod <- '
  anx =~ HADS1 + HADS2 + HADS3 + HADS4 + HADS5 + HADS6 + HADS7
'
HADS.GRM.fit <- cfa(HADS.GRM.mod, data = HADS, ordered = paste("HADS", 1:7, sep=""))
```

The warning about empty bivariate cell tables is common, do not worry about this. The bivariate tables are frequency tables, with response categories of one item in the rows and those of another item in the columns. These bivariate tables are used to calculate the tetrachoric correlation matrix. With 7 items with 4 response categories, there are $7 \times (7 - 1)$ bivariate tables, with 4×4 cells in each table. If twelve of them are empty, that is certainly not a lot.

```
summary(HADS.GRM.fit, standardized = TRUE)
```

```
## lavaan (0.6-1) converged normally after 18 iterations
##
## Number of observations                    502
##
## Estimator                                DWLS            Robust
## Model Fit Test Statistic                 94.652          171.090
## Degrees of freedom                       14             14
## P-value (Chi-square)                     0.000            0.000
```

```

## Scaling correction factor                                0.559
## Shift parameter                                         1.733
##   for simple second-order correction (Mplus variant)
##
## Parameter Estimates:
##
## Information                               Expected
## Information saturated (h1) model          Unstructured
## Standard Errors                           Robust.sem
##
## Latent Variables:
##      Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##   anx =~
##   HADS1      1.000
##   HADS2      0.961    0.033   29.081    0.000    0.799    0.799
##   HADS3      0.962    0.033   29.428    0.000    0.800    0.800
##   HADS4      0.756    0.042   18.089    0.000    0.629    0.629
##   HADS5      0.737    0.041   18.153    0.000    0.613    0.613
##   HADS6      0.878    0.034   25.691    0.000    0.730    0.730
##   HADS7      0.912    0.034   27.160    0.000    0.759    0.759
##
## Intercepts:
##      Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##   .HADS1      0.000
##   .HADS2      0.000
##   .HADS3      0.000
##   .HADS4      0.000
##   .HADS5      0.000
##   .HADS6      0.000
##   .HADS7      0.000
##   anx      0.000
##
## Thresholds:
##      Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##   HADS1|t1    -1.368    0.080  -17.124    0.000   -1.368   -1.368
##   HADS1|t2    -0.242    0.057   -4.276    0.000   -0.242   -0.242
##   HADS1|t3     0.866    0.064   13.462    0.000    0.866    0.866
##   HADS2|t1    -0.186    0.056   -3.298    0.001   -0.186   -0.186
##   HADS2|t2     0.604    0.060   10.089    0.000    0.604    0.604
##   HADS2|t3     1.493    0.086   17.407    0.000    1.493    1.493
##   HADS3|t1    -1.039    0.068  -15.170    0.000   -1.039   -1.039
##   HADS3|t2    -0.005    0.056   -0.089    0.929   -0.005   -0.005
##   HADS3|t3     1.065    0.069   15.388    0.000    1.065    1.065
##   HADS4|t1    -1.540    0.088  -17.450    0.000   -1.540   -1.540
##   HADS4|t2    -0.762    0.062  -12.224    0.000   -0.762   -0.762
##   HADS4|t3     0.411    0.058    7.113    0.000    0.411    0.411
##   HADS5|t1    -0.368    0.057   -6.406    0.000   -0.368   -0.368
##   HADS5|t2     0.511    0.059    8.696    0.000    0.511    0.511
##   HADS5|t3     1.449    0.084   17.336    0.000    1.449    1.449
##   HADS6|t1    -1.110    0.071  -15.740    0.000   -1.110   -1.110
##   HADS6|t2     0.100    0.056    1.783    0.075    0.100    0.100
##   HADS6|t3     1.138    0.071   15.944    0.000    1.138    1.138
##   HADS7|t1    -0.263    0.057   -4.632    0.000   -0.263   -0.263
##   HADS7|t2     0.735    0.062   11.887    0.000    0.735    0.735

```



```
##      HADS7|t3          1.883    0.112   16.784    0.000    1.883    1.883
##
## Variances:
##           Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##      .HADS1         0.308             0.308    0.308
##      .HADS2         0.361             0.361    0.361
##      .HADS3         0.360             0.360    0.360
##      .HADS4         0.604             0.604    0.604
##      .HADS5         0.624             0.624    0.624
##      .HADS6         0.467             0.467    0.467
##      .HADS7         0.424             0.424    0.424
##      anx           0.692    0.033   21.088    0.000    1.000    1.000
##
## Scales y*:
##           Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##      HADS1         1.000             1.000    1.000
##      HADS2         1.000             1.000    1.000
##      HADS3         1.000             1.000    1.000
##      HADS4         1.000             1.000    1.000
##      HADS5         1.000             1.000    1.000
##      HADS6         1.000             1.000    1.000
##      HADS7         1.000             1.000    1.000
```

```
fitMeasures(HADS.GRM.fit, fitinds)
```

```
##      cfi.scaled rmsea.scaled      srmr
##      0.954      0.150      0.066
```

b) Which category from which item is the 'easiest'?

HADS4, it has the lowest thresholds for all categories. Note that you can also see this from the histograms printed earlier.

c) What do we mean by 'easiest' in this case?

For for this item, lower latent trait (Anxiety) values are needed to endorse higher response categories.

d) Are all category thresholds ordered similarly across items?

Yes, they go from low to high.

e) Fit a partial credit model to the data:

```
HADS.PCM.mod <- '
  anx =~ 1*HADS1 + 1*HADS2 + 1*HADS3 + 1*HADS4 + 1*HADS5 + 1*HADS6 + 1*HADS7
'
HADS.PCM.fit <- cfa(HADS.PCM.mod, data = HADS, ordered = paste("HADS", 1:7, sep=""))
summary(HADS.PCM.fit, standardized = TRUE)
```

```
## lavaan (0.6-1) converged normally after 3 iterations
##
##      Number of observations          502
##
##      Estimator              DWLS      Robust
##      Model Fit Test Statistic    192.056    206.433
##      Degrees of freedom           20         20
##      P-value (Chi-square)         0.000      0.000
##      Scaling correction factor          0.950
##      Shift parameter              4.277
```

```

##      for simple second-order correction (Mplus variant)
##
## Parameter Estimates:
##
##      Information                                Expected
##      Information saturated (h1) model          Unstructured
##      Standard Errors                          Robust.sem
##
## Latent Variables:
##      Estimate  Std.Err  z-value  P(>|z|)  Std.lv  Std.all
##      anx =~
##      HADS1      (1)    1.000                0.750    0.750
##      HADS2      (1)    1.000                0.750    0.750
##      HADS3      (1)    1.000                0.750    0.750
##      HADS4      (1)    1.000                0.750    0.750
##      HADS5      (1)    1.000                0.750    0.750
##      HADS6      (1)    1.000                0.750    0.750
##      HADS7      (1)    1.000                0.750    0.750
##
## Intercepts:
##      Estimate  Std.Err  z-value  P(>|z|)  Std.lv  Std.all
##      .HADS1      0.000                0.000    0.000
##      .HADS2      0.000                0.000    0.000
##      .HADS3      0.000                0.000    0.000
##      .HADS4      0.000                0.000    0.000
##      .HADS5      0.000                0.000    0.000
##      .HADS6      0.000                0.000    0.000
##      .HADS7      0.000                0.000    0.000
##      anx        0.000                0.000    0.000
##
## Thresholds:
##      Estimate  Std.Err  z-value  P(>|z|)  Std.lv  Std.all
##      HADS1|t1    -1.368    0.080   -17.124    0.000   -1.368   -1.368
##      HADS1|t2    -0.242    0.057    -4.276    0.000   -0.242   -0.242
##      HADS1|t3     0.866    0.064   13.462    0.000    0.866    0.866
##      HADS2|t1    -0.186    0.056    -3.298    0.001   -0.186   -0.186
##      HADS2|t2     0.604    0.060   10.089    0.000    0.604    0.604
##      HADS2|t3     1.493    0.086   17.407    0.000    1.493    1.493
##      HADS3|t1    -1.039    0.068   -15.170    0.000   -1.039   -1.039
##      HADS3|t2    -0.005    0.056    -0.089    0.929   -0.005   -0.005
##      HADS3|t3     1.065    0.069   15.388    0.000    1.065    1.065
##      HADS4|t1    -1.540    0.088   -17.450    0.000   -1.540   -1.540
##      HADS4|t2    -0.762    0.062   -12.224    0.000   -0.762   -0.762
##      HADS4|t3     0.411    0.058    7.113    0.000    0.411    0.411
##      HADS5|t1    -0.368    0.057    -6.406    0.000   -0.368   -0.368
##      HADS5|t2     0.511    0.059    8.696    0.000    0.511    0.511
##      HADS5|t3     1.449    0.084   17.336    0.000    1.449    1.449
##      HADS6|t1    -1.110    0.071   -15.740    0.000   -1.110   -1.110
##      HADS6|t2     0.100    0.056    1.783    0.075    0.100    0.100
##      HADS6|t3     1.138    0.071   15.944    0.000    1.138    1.138
##      HADS7|t1    -0.263    0.057    -4.632    0.000   -0.263   -0.263
##      HADS7|t2     0.735    0.062   11.887    0.000    0.735    0.735
##      HADS7|t3     1.883    0.112   16.784    0.000    1.883    1.883
##

```

```
## Variances:
##           Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##   .HADS1         0.438                0.438  0.438
##   .HADS2         0.438                0.438  0.438
##   .HADS3         0.438                0.438  0.438
##   .HADS4         0.438                0.438  0.438
##   .HADS5         0.438                0.438  0.438
##   .HADS6         0.438                0.438  0.438
##   .HADS7         0.438                0.438  0.438
##   anx           0.562      0.019   30.186   0.000   1.000   1.000
##
## Scales y*:
##           Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##   HADS1         1.000                1.000  1.000
##   HADS2         1.000                1.000  1.000
##   HADS3         1.000                1.000  1.000
##   HADS4         1.000                1.000  1.000
##   HADS5         1.000                1.000  1.000
##   HADS6         1.000                1.000  1.000
##   HADS7         1.000                1.000  1.000
```

Note that again we see Item 4 is the easiest item, with the lowest thresholds.

f) Test whether the GRM or PCM fits better:

```
fitMeasures(HADS.GRM.fit, fitinds)
```

```
##   cfi.scaled rmsea.scaled      srmr
##       0.954      0.150      0.066
```

```
fitMeasures(HADS.PCM.fit, fitinds)
```

```
##   cfi.scaled rmsea.scaled      srmr
##       0.946      0.136      0.097
```

```
anova(HADS.PCM.fit, HADS.GRM.fit)
```

```
## Scaled Chi Square Difference Test (method = "satorra.2000")
```

```
##
```

```
##           Df AIC BIC   Chisq Chisq diff Df diff Pr(>Chisq)
```

```
## HADS.GRM.fit 14      94.652
```

```
## HADS.PCM.fit 20     192.056    67.696      6 1.213e-12 ***
```

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

There is a significant difference in fit, so we should prefer the GRM, which is more complex. This is also indicated by the CFI values. However, if we use the RMSEA as the main criterion for model selection, we would prefer the PCM, because it is more parsimonious.