

PROBABILITY: QUIZ 2

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1. PROBLEM 1

Indicate True or False for each statement.

1.1. **Part A.** X and Y are independent random variables. X is uniformly distributed on $[-2, 2]$. Y is uniformly distributed on $[-1, 5]$. If $Z = X + Y$, then $f_Z(3) = 1/6$.

True.

This PDF of the sum of two random variables is found through convolution, where $f_Z(z) = \int_{-\infty}^{\infty} f_X(x)f_Y(z-x) dx$.

$f_Z(3) = \int_{-\infty}^{\infty} f_X(x)f_Y(3-x) dx$
 $f_X(x)$ is non-zero only for $x \in [-2, 2]$. $f_Y(y)$ is only non-zero for $y \in [-1, 5]$, thus $f_Y(3-x)$ is only non-zero for $x \in [-2, 4]$. The limits of integration is the union of these two sets: $[-2, 2]$.
 $f_Z(3) = \int_{-2}^2 f_X(x)f_Y(3-x) dx = \left(\frac{1}{6}\right) \left(\frac{1}{4}\right) (2+2) = \frac{1}{6}$

1.2. **Part B.** If X is a Gaussian random variable with mean 0 and variance 1, then the density function of $Z = |X|$ is equal to $2f_X(x), z \geq 0$.

True. Since X is symmetric about 0, $Z = |X|$ means the probability of $-x$ adds to the probability of x , causing the $f_Z(z) = 2f_X(z)$ where $z \geq 0$.

1.3. **Part C.** The sum of a random number (N) of independent Gaussian random variables with zero mean and unit variance results in a Gaussian random variable, regardless of the distribution of N .

False

For $Y = X_1 + X_2 + \dots + X_N$, the transform of Y can be found as such.

$M_Y(s) = M_N(s)|_{e^s = M_X(s)}$. Let N be Gaussian.

$M_N(s) = e^{s^2/2}$ and since X is Gaussian, $M_X(s) = e^{s^2/2}$.

Thus, $M_Y(s) = \sqrt{(e^s)^s}|_{e^s = M_X(s)} = \sqrt{(e^{s^2/2})^s}$, which is clearly not Gaussian.

1.4. **Part D.** If X and Y are independent random variables, both exponentially distributed with parameters λ_1 and λ_2 respectively, then the random variable $Z = \min\{X, Y\}$ is also exponentially distributed.

True

The CDF $F_Z(z) = P(Z \leq z) = P(\min\{X, Y\} \leq z)$ is not immediately evident, but $1 - F_Z(z)$ is.

$$\begin{aligned} 1 - F_Z(z) &= P(Z > z) = P(\min\{X, Y\} > z) \\ &= P(X > z, Y > z) = P(X > z) P(Y > z) \\ &= (1 - F_X(z)) (1 - F_Y(z)) = (1 - 1 + e^{-\lambda_1 z}) (1 - 1 + e^{-\lambda_2 z}) \\ 1 - F_Z(z) &= e^{-(\lambda_1 + \lambda_2)z} \\ f_Z(z) &= (\lambda_1 + \lambda_2) e^{-(\lambda_1 + \lambda_2)z} \end{aligned}$$

1.5. **Part E.** Let the transform associated with random variable X be $M_X(s) = \left(\frac{e^s}{1-s}\right)^{15}$. $E[X]$ is equal to 30.

True

$$\begin{aligned} E[X] &= \frac{d}{ds} M_X(s) \Big|_{s=0} = 15 \left(\frac{e^s}{1-s}\right)^{14} \frac{(1-s)e^s + e^s}{(1-s)^2} \Big|_{s=0} \\ E[X] &= 15 \left(\frac{1}{1-0}\right)^{14} \frac{(1)(1) + 1}{(1-0)^2} = 15 (2) = 30 \end{aligned}$$

1.6. **Part F.** X and Y are independent random variables. Y is normal with mean 0 and variance 1, X is uniform on $[0, 1]$. $Z = X + Y$. The conditional density of Z given X , $f_{Z|X}(z|x)$ is normal with mean x and variance 1.

True

Z given some x is the normal random variable Y plus x , $Z|x = Y + x$. Thus, $E[Z|X] = E[Y] + x$ and $\text{Var}(Z|X) = \text{Var}(Y)$.

1.7. **Part G.** $\text{Var}(Z)=2$.

False

Since X and Y are independent, $\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$
So, $\text{Var}(Z) = \frac{1}{12} + 1 \neq 2$.

1.8. **Part H.** $E[X|Z = -1] = -1$

False

X is bounded by $[0, 1]$, so it cannot take the value of -1.

1.9. **Part I.** $\text{Cov}(X, Z) = \text{Var}(X)$

True

$$\text{Cov}(X, Z) = \mathbf{E}[XZ] - \mathbf{E}[X]\mathbf{E}[Z]$$

$$= \mathbf{E}[X(X + Y)] - \mathbf{E}[X]\mathbf{E}[X + Y]$$

$$= \mathbf{E}[X^2] + \mathbf{E}[XY] - \mathbf{E}[X]^2 - \mathbf{E}[X]\mathbf{E}[Y]$$

$$= \mathbf{E}[X^2] - \mathbf{E}[X]^2 + \mathbf{E}[XY] - \mathbf{E}[X]\mathbf{E}[Y]$$

$$= \text{Var}(X) + \text{Cov}(X, Y)$$

And since X and Y are independent, $\text{Cov}(X, Y) = 0$, thus $\text{Cov}(X, Z) = \text{Var}(X)$

1.10. **Part J.** $Z = \mathbf{E}[X|Z] + \mathbf{E}[Y|Z]$

True

$\mathbf{E}[Z|Z] = \mathbf{E}[X + Y|Z]$ and since conditional expectation is linear, $\mathbf{E}[Z|Z] = Z = \mathbf{E}[X|Z] + \mathbf{E}[Y|Z]$

2. PROBLEM 2

2.1. **Part A.** Find the least squares estimate of Y given $X = x$, for all possible values of x .

$$\mathbf{E}[Y|X] = \begin{cases} \frac{1}{2} & 0 \leq X < 1 \\ X - \frac{1}{2} & 1 \leq X \leq 2 \end{cases}$$

2.2. **Part B.** Let $g(x)$ be the estimate from Part A. Find $\mathbf{E}[g(X)]$ and $\text{Var}(g(X))$.

$$\mathbf{E}[g(X)] = \int_0^2 g(x) f_X(x) dx = \frac{1}{2} \left(\int_0^1 \frac{1}{2} dx + \int_1^2 x - \frac{1}{2} dx \right)$$

$$\mathbf{E}[g(X)] = \frac{1}{2} \left(\frac{1}{2} + 1 - 0 \right) = \frac{3}{4}$$

$$\mathbf{E}[g(X)^2] = \int_0^2 g(x)^2 f_X(x) dx = \frac{1}{2} \left(\frac{1}{4} + \int_1^2 (x - \frac{1}{2})^2 dx \right) = \frac{1}{2} \left(\frac{1}{4} + \frac{13}{12} \right) = \frac{2}{3}$$

$$\text{Var}(g(X)) = \mathbf{E}[g(X)^2] - \mathbf{E}[g(X)]^2 = \frac{2}{3} - \frac{9}{16} = \frac{5}{48} \approx 0.1042$$

2.3. **Part C.** Find the mean squared error $\mathbf{E}[(Y - g(X))^2]$. Is it the same as $\mathbf{E}[\text{var}(Y|X)]$?

Since $g(X) = \mathbf{E}[Y|X]$,

$$\mathbf{E}[(Y - g(X))^2] = \mathbf{E}[(Y - \mathbf{E}[Y|X])^2] = \mathbf{E}[\mathbf{E}[(Y - \mathbf{E}[Y|X])^2|X]]$$

$$= \mathbf{E}[\mathbf{E}[(Y|X - \mathbf{E}[Y|X])^2]] = \mathbf{E}[\text{Var}(Y|X)]$$

Since Y is uniform for all X , $\text{Var}(Y|X) = \frac{1}{12}(Y_1 + 1 - Y_1)^2 = \frac{1}{12}$, $0 \leq X \leq 2$
 $\mathbf{E}[\text{Var}(Y|X)] = \frac{1}{2} \int_0^2 \frac{1}{12} dx = \frac{1}{12}$

3. PROBLEM 3

Each year, an editor is sent a random number of books for review. The number of books received can be modeled as a Poisson random variable N with mean μ . Each book contains a random number of typos, modeled by a Poisson random variable with mean λ . Let B_i denote the number of typos in book i . Assume all random variables are independent. The editor finds typos with probability p , independent of all other findings, and other random variables.

There are two different payment options:

Option 1. \$1 for each typo found

Option 2. \$1 for each book where at least 1 typo is found

Let X_i be the amount of money the editor receives for book i and T the total amount of money the editor receives in a year.

3.1. **Part A.** The the PMF of X_i under option 1.

The PMF of X_i under option 1 is the PMF of the number of typos in book i . For each typo in book i , the editor has a p probability of catching it. Thus, the PMF is a binomial distribution with the number of trials the random variable of B_i .

$$\mathbf{P}(X_i = x | B_i = b) = \binom{b}{x} p^x (1-p)^{b-x}$$

Since X_i and B_i are independent,

$$\mathbf{P}(X_i = x) = \sum_{b=x}^{\infty} \mathbf{P}(X_i = x | B_i = b) \mathbf{P}(B_i = b)$$

$$\mathbf{P}(X_i = x) = \sum_{b=x}^{\infty} \binom{b}{x} p^x (1-p)^{b-x} \frac{\lambda e^{-\lambda}}{b!}$$

$$\mathbf{P}(X_i = x) = \frac{(p\lambda)^x}{x!} e^{-\lambda} \sum_{b=0}^{\infty} \frac{(\lambda(1-p))^b}{b!} = \frac{(p\lambda)^x}{x!} e^{-p\lambda}$$

This is a Poisson PMF with mean $p\lambda$.

3.2. **Part B.** Find $M_T(s)$ under option 1.

$T = X_1 + X_2 + \dots + X_N$. Thus, $M_T(s) = M_N(s) |_{e^s = M_X(s)}$

$$M_N(s) = e^{\mu(e^s - 1)} \quad M_X(s) = e^{p\lambda(e^s - 1)}$$

$$M_T(s) = e^{\mu(e^{p\lambda(e^s - 1)} - 1)}$$

3.3. **Part C.** Find $\mathbf{P}(T = 2)$ under option 2.

Since T is a discrete random variable, we can find probabilities through differentiation of $M_T(s)$ as such:

$$\begin{aligned}\mathbf{P}(T = t) &= \frac{1}{t!} \frac{d^t}{d(e^s)^t} M(s)|_{e^s=0} \\ \mathbf{P}(T = 2) &= \frac{1}{2} \frac{d^2}{d(e^s)^2} M_T(s)|_{e^s=0} \\ &= \left[\frac{\mu p \lambda}{2} \exp\{\mu e^{p\lambda(e^s-1)} - \mu + p\lambda e^s - p\lambda\} \left(\mu p \lambda e^{p\lambda(e^s-1)} + p\lambda \right) \right] |_{e^s=0} \\ \mathbf{P}(T = 2) &= \frac{1}{2} \mu (p\lambda)^2 e^{-p\lambda} e^{\mu(e^{-p\lambda}-1)} (\mu e^{-p\lambda} + 1)\end{aligned}$$

3.4. **Part D.** Find $\mathbf{E}[T]$.

$$\begin{aligned}\mathbf{E}[T] &= \frac{d}{ds} M_T(s)|_{s=0} = e^{\mu(e^{p\lambda(e^s-1)}-1)} \mu e^{p\lambda(e^s-1)} p\lambda e^s |_{s=0} \\ \mathbf{E}[T] &= \mu p \lambda\end{aligned}$$

3.5. **Part E.** Find $\text{Var}(T)$.

$$\begin{aligned}\text{Var}(T) &= \mathbf{E}[T^2] - \mathbf{E}[T]^2 \\ \mathbf{E}[T^2] &= \frac{d^2}{ds^2} M_T(s)|_{s=0} = (\mu p \lambda)^2 + \mu (p\lambda)^2 + \mu p \lambda \\ \text{Var}(T) &= (\mu p \lambda)^2 + \mu (p\lambda)^2 + \mu p \lambda - (\mu p \lambda)^2 = \mu p \lambda (p\lambda + 1)\end{aligned}$$

3.6. **Part F.** Find the PMF of X_i under option 2.

Under option 2, $X_i = 1$ if at least 1 typo is found in book i , and 0 otherwise (Binomial distribution). From Part A, we see that the number of found typos is book i is a Poisson random variable with mean $p\lambda$. Thus, the probability that $X_i = 0$ is $e^{-p\lambda}$. The full PMF is:

$$X_i = \begin{cases} e^{-p\lambda} & x = 0 \\ 1 - e^{-p\lambda} & x = 1 \\ 0 & \text{otherwise} \end{cases}$$

3.7. **Part G.** Find $\mathbf{E}[T]$ under option 2.

Each book viewed by the editor, there is a ρ probability that he or she will get a dollar for the book. There are N number of attempts at getting dollars. Thus, T has a binomial distribution with the number of trials being a Poisson random variable. Following a similar procedure as Part A:

$$\mathbf{P}(T = t) = \sum_{n=t}^N \binom{n}{t} \rho^t (1 - \rho)^{n-t} \frac{\mu^n}{n!} e^{-\mu} = \frac{(\mu\rho)^t}{t!} e^{-\mu\rho}$$

This is a Poisson distribution with mean $\mu\rho$. From Part G, we saw that the probability ρ of receiving a dollar for book i is $1 - e^{-p\lambda}$. Therefore, we have:

$$\mathbf{E}[T] = \mu\rho = \mu(1 - e^{-p\lambda})$$

3.8. **Part H.** Which option should the editor choose?

We will answer this by finding the option with the highest expected value for T . Under option 1, $\mathbf{E}[T] = \mu p \lambda$. Under option 2, $\mathbf{E}[T] = \mu(1 - e^{-p\lambda})$.

Both μ and λ are out of the editor's control. But p is determined by the editor's skill in spotting typos. As p increases towards 1, the expected yearly payout under option 1 grows linearly towards $\mu\lambda$. Whereas under option 2 it is bounded above by $\mu(1 - e^{-\lambda})$. Note that $1 - e^{-\lambda}$ grows sub-linearly, so $\mu\lambda$ will always be greater than $\mu(1 - e^{-\lambda})$.

Therefore, option 1 will result in the highest expected value of T .