



Differential expression with the limma package

- Admin: exercise due dates, journal club sign-up, project proposals
- Review from last week: *biological versus technical replication*
- In class exercise: how do we know methods work well in practice?
- Preliminaries – really quick comments on pre-processing microarrays: background correction, normalization, exploratory analysis
- The philosophy of limma
- limma – INTRO TO linear models for microarray data



Admin

Exercise due dates: from now on, always the following Wednesday by 8pm

Evaluation of exercise: 30% of total mark (we will accumulate the marks of the best 7 exercises, worth 4% each and then add 2%)

Journal club sign-up: presentations will start 26.10.2015; specify your paper + your date by 19.10.2015

Project proposals: 1 paragraph by 02.11.2015, due 08.01.2016 by 8pm.



University of
Zurich^{UZH}

Institute of Molecular Life Sciences

From the feed

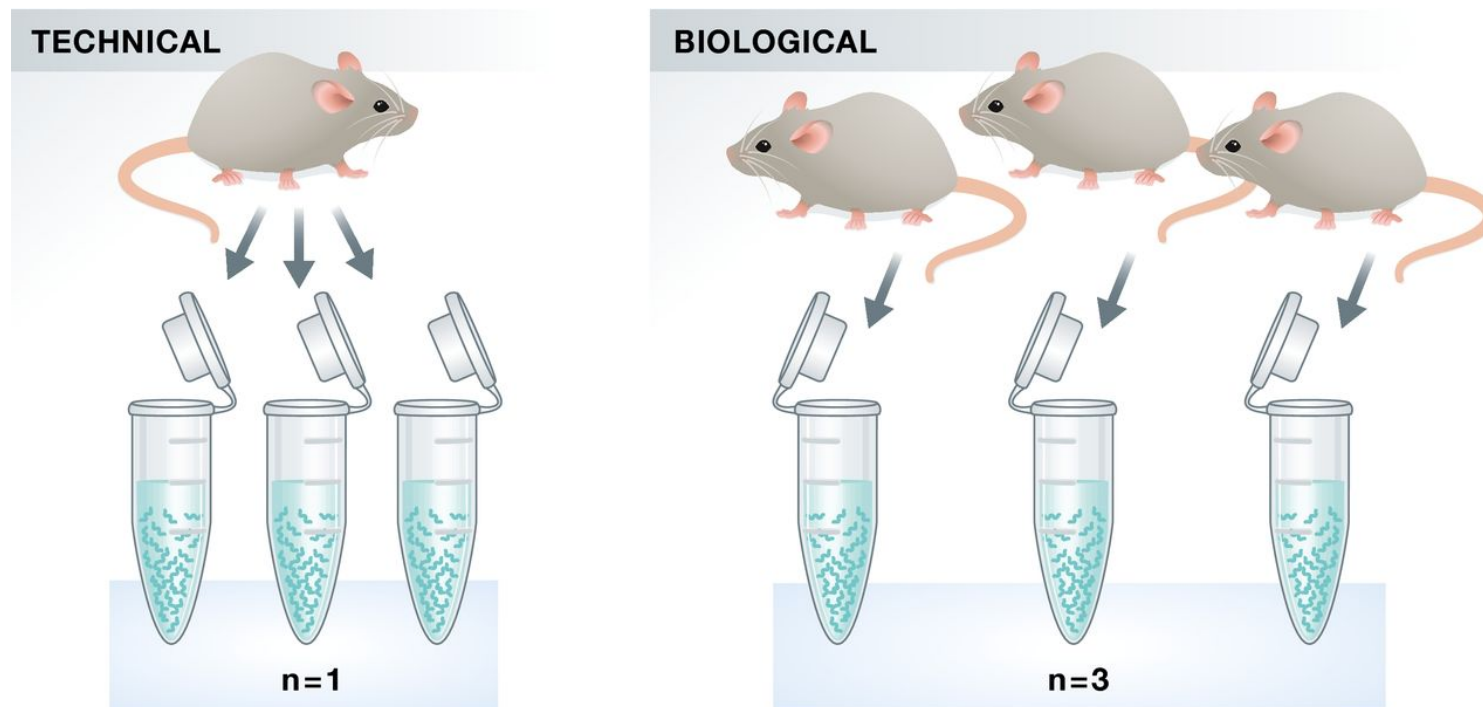
The unreasonable effectiveness of GitHub browsability →
may have relevance to handing in assignments

Mission Bay (science publishing) manifesto

The Battle Over Genome Editing Gets Science All Wrong



Review: one thing we didn't get back to last week



What do we want? Why?



From the feed: Terry's IMS Bulletin + “Over-optimism”

We will see a lot of methods in this course – **how do we evaluate what works well in practice ?**

<http://bulletin.imstat.org/2012/11/terences-stuff-does-it-work-in-practice/>

Gene expression

Advance Access publication June 26, 2010

Over-optimism in bioinformatics: an illustration

Monika Jelizarow¹, Vincent Guillemot^{1,2}, Arthur Tenenhaus², Korbinian Strimmer³ and Anne-Laure Boulesteix^{1,*}

¹Department of Medical Informatics, Biometry and Epidemiology, University of Munich, Marchioninstr. 15, 81377 Munich, Germany, ²SUPELEC Sciences des Systèmes (E3S)-Department of Signal Processing and Electronics Systems - 3, rue Joliot Curie, Plateau de Moulon, 91192 Gif-sur-Yvette Cedex, France and ³Department of Medical Informatics, Statistics and Epidemiology, University of Leipzig, Härtelstr. 16-18, 04107 Leipzig, Germany

Associate Editor: John Quackenbush

“if the improvement of a quantitative criterion such as the error rate is the main contribution of a paper, the superiority of new algorithms should always be demonstrated on independent validation data.”



In class exercise + discussion

- (5 minutes) Read the excerpt from “Terence’s Stuff” column
- (5-10 minutes; discuss with your neighbour) Answer the following 3 questions:
 - How do we tell what works in practice?
 - What problems arise using simulated (synthetic) data?
 - What problems arise using real data?
 - What are positive/negative controls?
- Discuss
- What metrics could/would/should we use?



Differential expression, small sample inference

- Table of data (e.g., microarray gene expression data with replicates of each of condition A, condition B)
 - *rows* = features (e.g., genes), *columns* = experimental units (samples)
- Most common problem in statistical bioinformatics: want to infer whether there is a **change in the response** → a statistical test for each row of the table.

What test might you use? Why is this hard? What issues arise? How much statistical power is there [1] ?

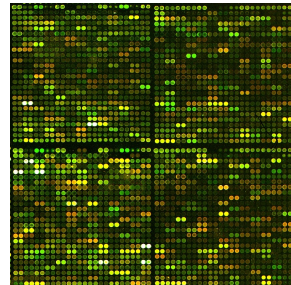
```
> head(y)
      group0      group0      group0      group1      group1      group1
gene1 -0.1874854  0.2584037 -0.05550717 -0.4617966 -0.3563024 -0.03271432
gene2 -3.5418798 -2.4540999  0.11750996 -4.3270442 -5.3462622 -5.54049106
gene3 -0.1226303  0.9354707 -1.10537767 -0.1037990  0.5221678 -1.72360854
gene4 -2.3394536 -0.3495697 -3.47742610 -3.2287093  6.1376670 -2.23871974
gene5 -3.7978820  1.4545702 -7.14796503 -4.0500796  4.7235714 10.00033769
gene6  1.4627078 -0.3096070 -0.26230124 -0.7903434  0.8398769 -0.96822312
```

[1] <http://www.stat.ubc.ca/~rollin/stats/ssize/n2.html>



Microarray expression measures

Two-colour



$$y_{ga} = \log_2(R/G)$$

array

probe or gene

Affymetrix



$$y_{ga} =$$

log-intensity
(summarized over
probes)

Illumina



$$y_{ga} =$$

log-intensity
(summarized over
beads)

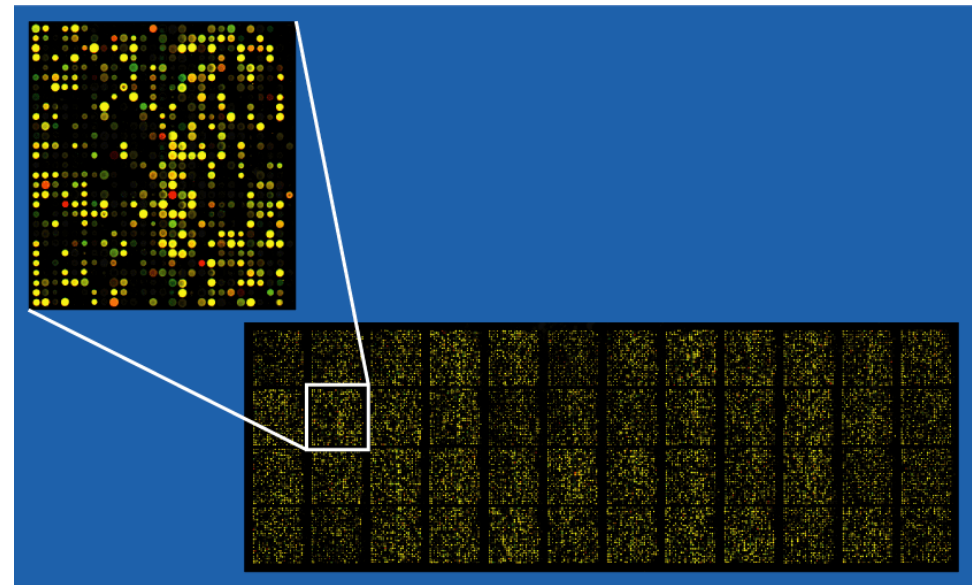
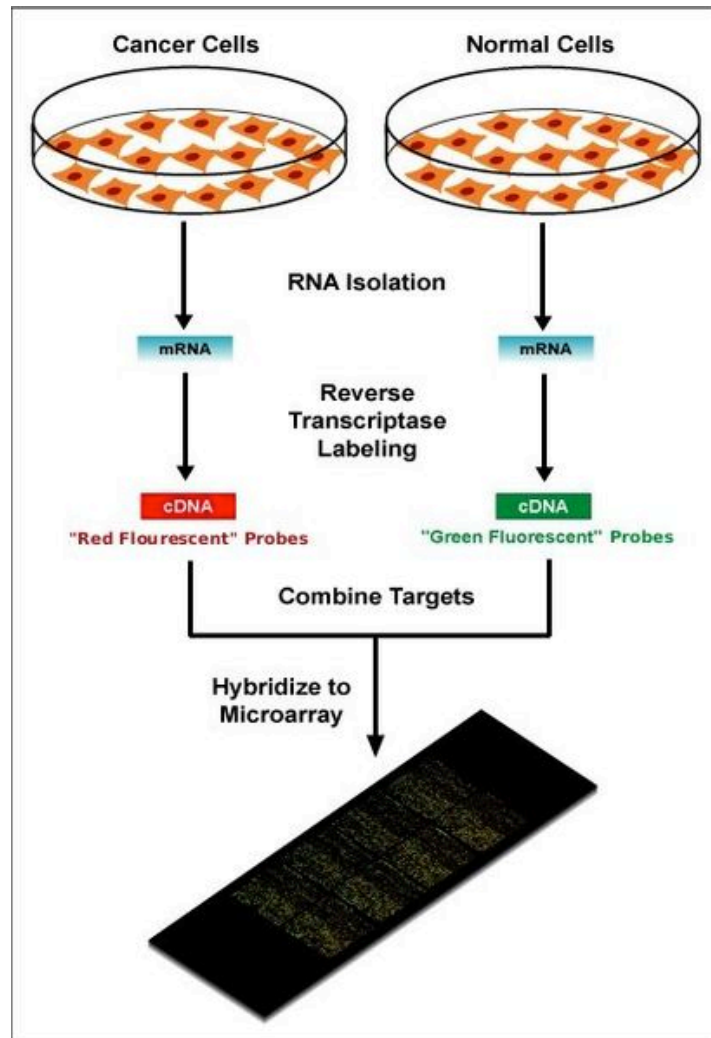


Questions of Interest

- What genes have changed in expression? (e.g. between disease/normal, affected by treatment) **Gene discovery, differential expression**
- Is a specified group of genes all up-regulated in a particular condition?
Gene set differential expression
- Can the expression profile predict outcome?
Class prediction, classification
- Are there tumour sub-types not previously identified?
Do my genes group into previously undiscovered pathways?
Class discovery, clustering



Two colour microarrays





Preprocessing: additive + multiplicative error model

Observe intensity for one probe on one array

Intensity = background + signal

$$I = B + S$$

additive
errors

multiplicative errors

This idea underlies variance stabilizing transformations vsn (two colour data) and vst (for Illumina data)



University of
Zurich^{UZH}

Institute of Molecular Life Sciences

Microarray background correction: maximum likelihood estimation for the normal-exponential convolution

JEREMY D. SILVER

Bioinformatics Division, Walter and Eliza Hall Institute, Parkville 3050, Victoria, Australia and
Department of Biostatistics, University of Copenhagen, Øster Farimagsgade 5, Entrance B,
PO Box 2099, DK-1014 Copenhagen K, Denmark
j.silver@biostat.ku.dk

MATTHEW E. RITCHIE

Department of Oncology, University of Cambridge, Cambridge CB2 0RE, UK

GORDON K. SMYTH*

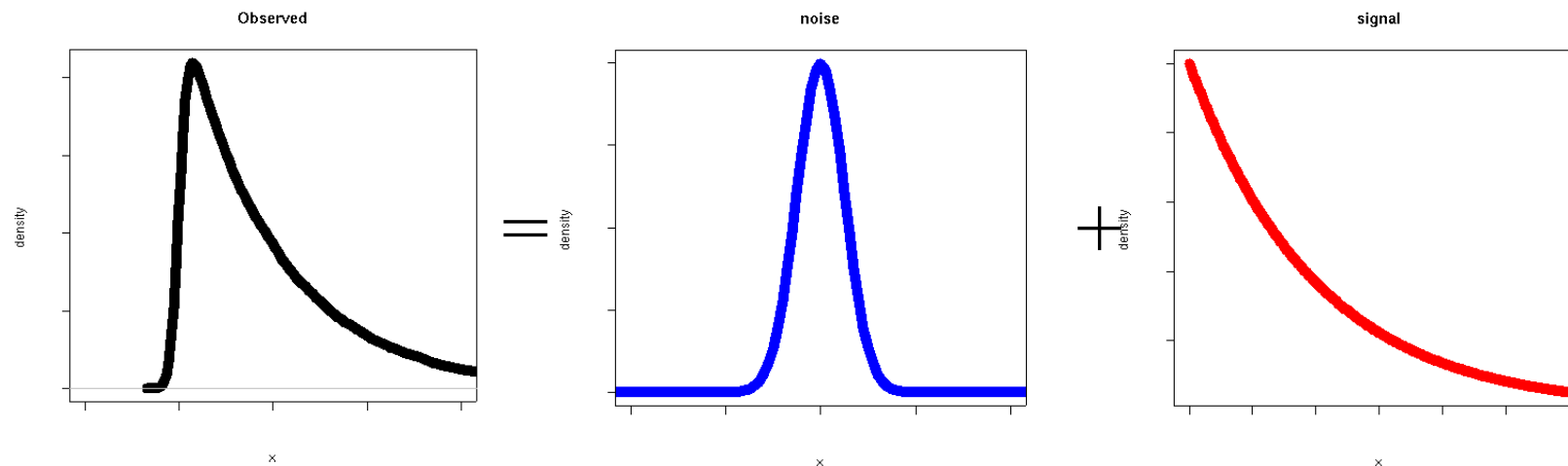
Bioinformatics Division, Walter and Eliza Hall Institute, Parkville 3050, Victoria, Australia
smyth@wehi.edu.au

normexp convolution model

$$\text{Intensity} = \text{Background} + \text{Signal}$$

$N(\mu, \sigma^2)$

Exponential(α)

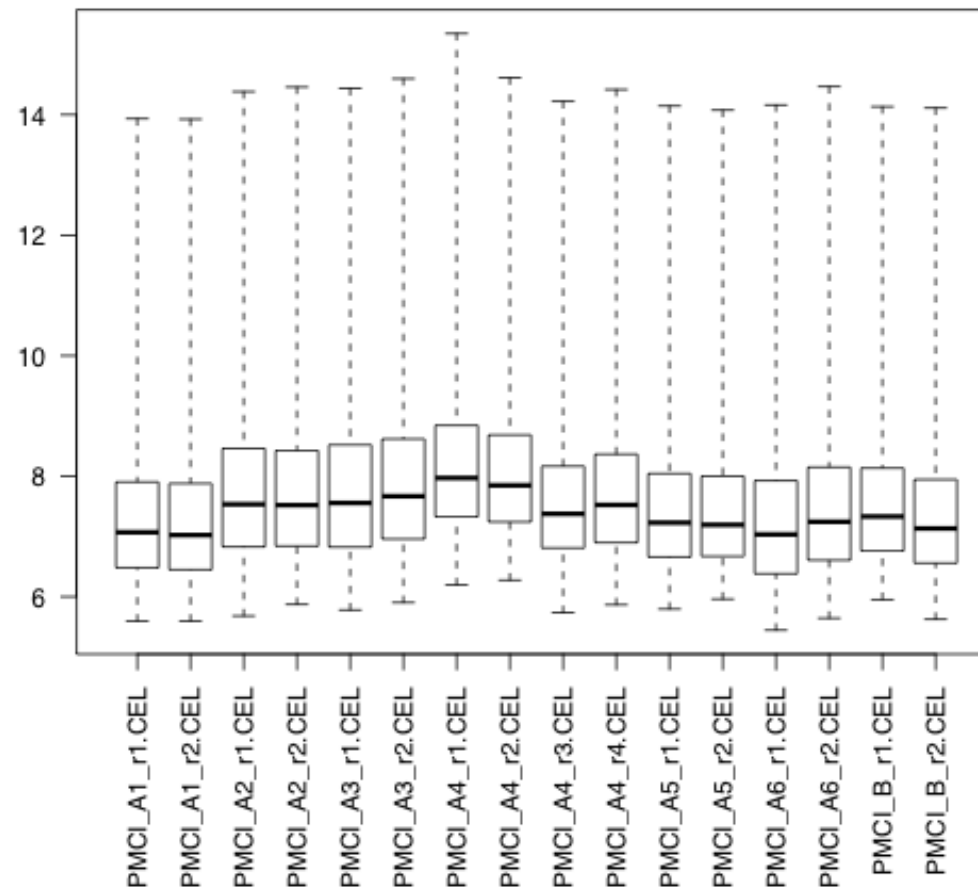




Normalization: one-colour

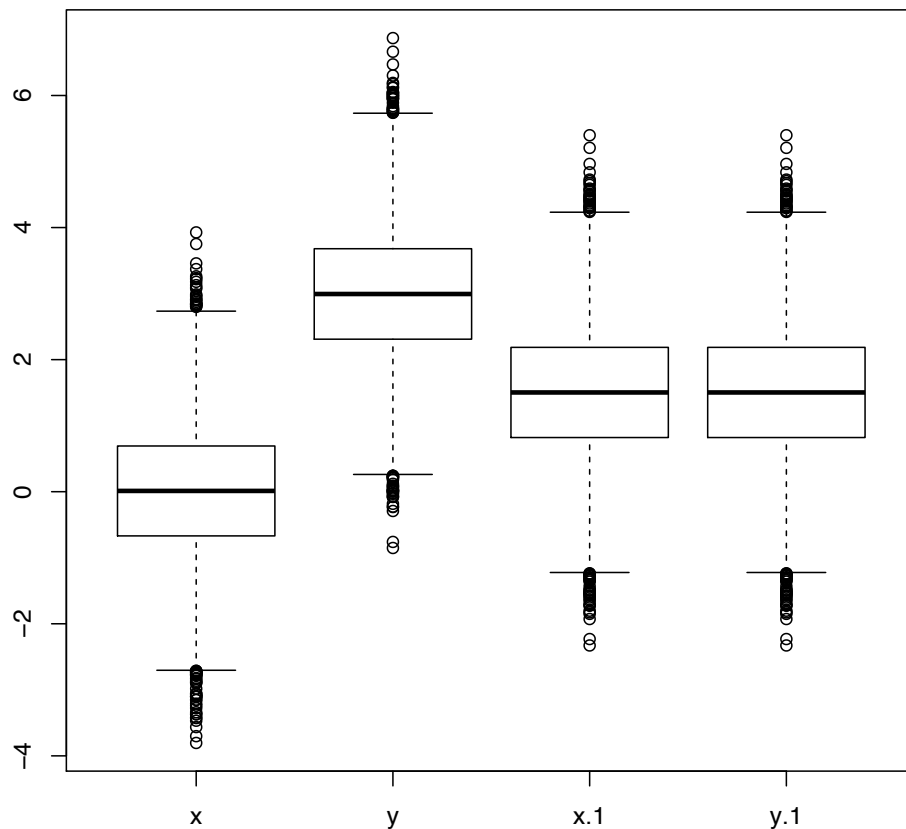


Similarly for single channel data, adjustments need to be made for all samples to be comparable.





Quantile normalization



```
x <- rnorm(10000, mean=0, sd=1)
y <- rnorm(10000, mean=3)
z <- cbind(x,y)
```

```
# create "reference" distribution
s <- apply(z,2,sort)
sm <- rowMeans(s)
```

```
# impose ref. distribution by ranks
r <- apply(z,2,rank)
n <- apply(r,2,function(u) sm[u])
```

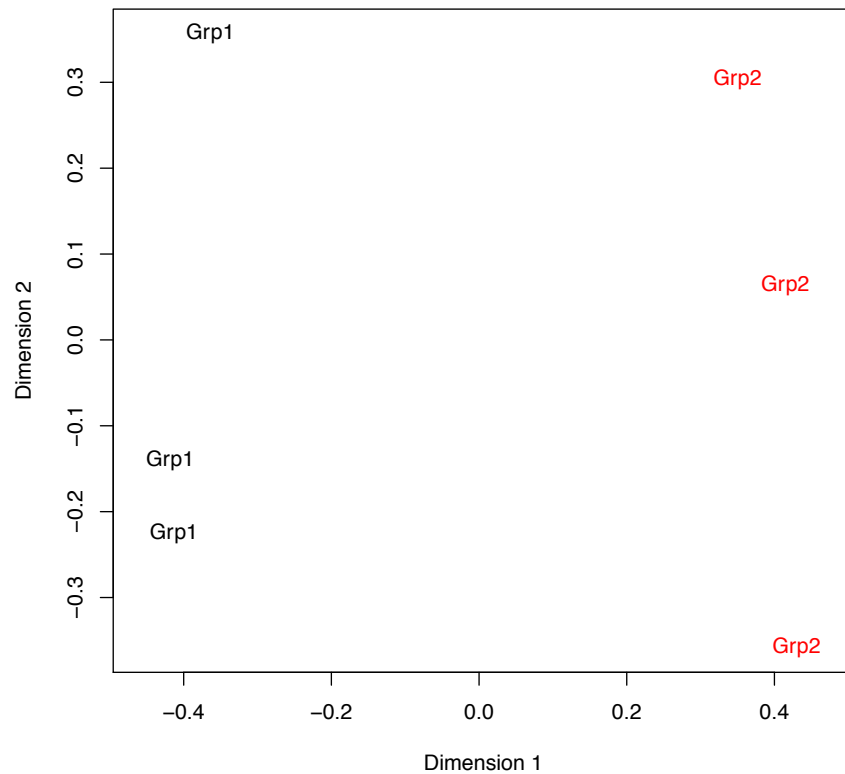
```
boxplot( data.frame(x=x,y=y,n) )
```

```
#> library(limma)
#> zn <- normalizeQuantiles(z)
#> all(zn==n)
#[1] TRUE
```



Quality assessments

Multidimensional scaling plot



```
sd <- 0.3*sqrt(4/rchisq(1000,df=4))  
x <- matrix(rnorm(1000*6,sd=sd),1000,6)  
x[1:50,4:6] <- x[1:50,4:6] + 2
```

```
mds <- plotMDS(x)
```

```
> round(mds$distance.matrix,3)
```

| | [,1] | [,2] | [,3] | [,4] | [,5] | [,6] |
|------|-------|-------|-------|-------|------|------|
| [1,] | 0.000 | 0.000 | 0.000 | 0.000 | 0.00 | 0 |
| [2,] | 0.835 | 0.000 | 0.000 | 0.000 | 0.00 | 0 |
| [3,] | 0.850 | 0.793 | 0.000 | 0.000 | 0.00 | 0 |
| [4,] | 1.089 | 1.068 | 1.058 | 0.000 | 0.00 | 0 |
| [5,] | 1.050 | 1.058 | 1.072 | 0.863 | 0.00 | 0 |
| [6,] | 0.991 | 1.047 | 1.046 | 0.865 | 0.85 | 0 |



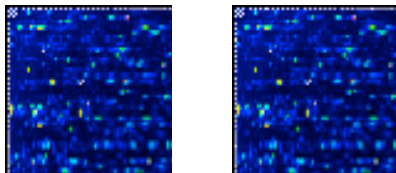
Limma concept: borrowing information across genes

- **Small data sets**: few samples, generally under-powered for 1 gene
- **Curse of dimensionality**: many tests, need to adjust for multiple testing (= loss of power)
- **Benefit of parallelism**: same model is fit for every gene. Can *borrow information* from one gene to another
 - **Hard**: assume parameters are constant across genes
 - **Soft**: smooth genewise parameters towards a common value in a graduated way, e.g., Bayes, empirical Bayes, Stein shrinkage ...

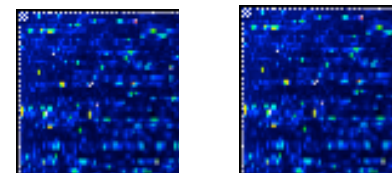


A very common experiment (1-colour)

Mutant x 2



WT x 2



Gene X



Which genes are differentially expressed?

$n_1 = n_2 = 2$ Affymetrix arrays

~30,000 probe-sets



Ordinary t-tests (1-colour)

$$t_g = \frac{\bar{y}_{\text{mu}} - \bar{y}_{\text{wt}}}{s_g c}$$

give very high false discovery rates

$$c = \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

Residual df = 2



t-tests with common variance

$$t_{g,\text{pooled}} = \frac{\bar{y}_{\text{mu}} - \bar{y}_{\text{wt}}}{s_0 c}$$

with residual standard deviation
across genes

s_0

pooled

More stable, but ignores gene-specific variability

$$c = \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$



A better compromise

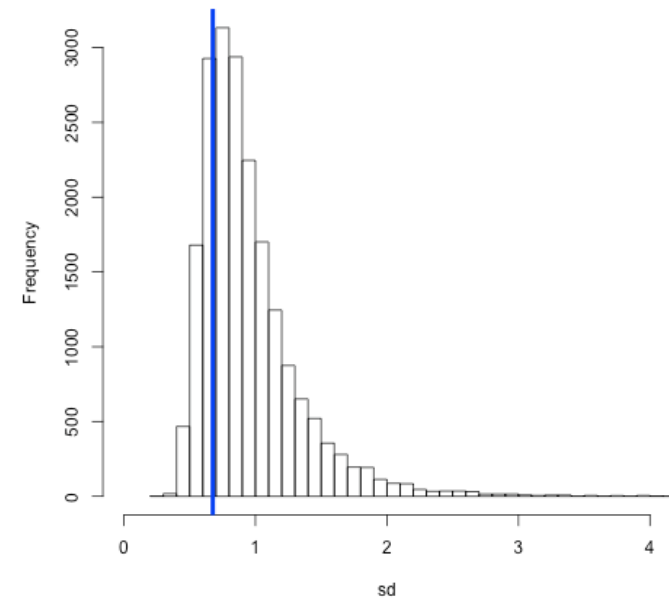
Shrink standard deviations towards common value

$$\tilde{s}_g^2 = \frac{d_0 s_0^2 + d_g s_g^2}{d_0 + d_g}$$

d = degrees
of freedom

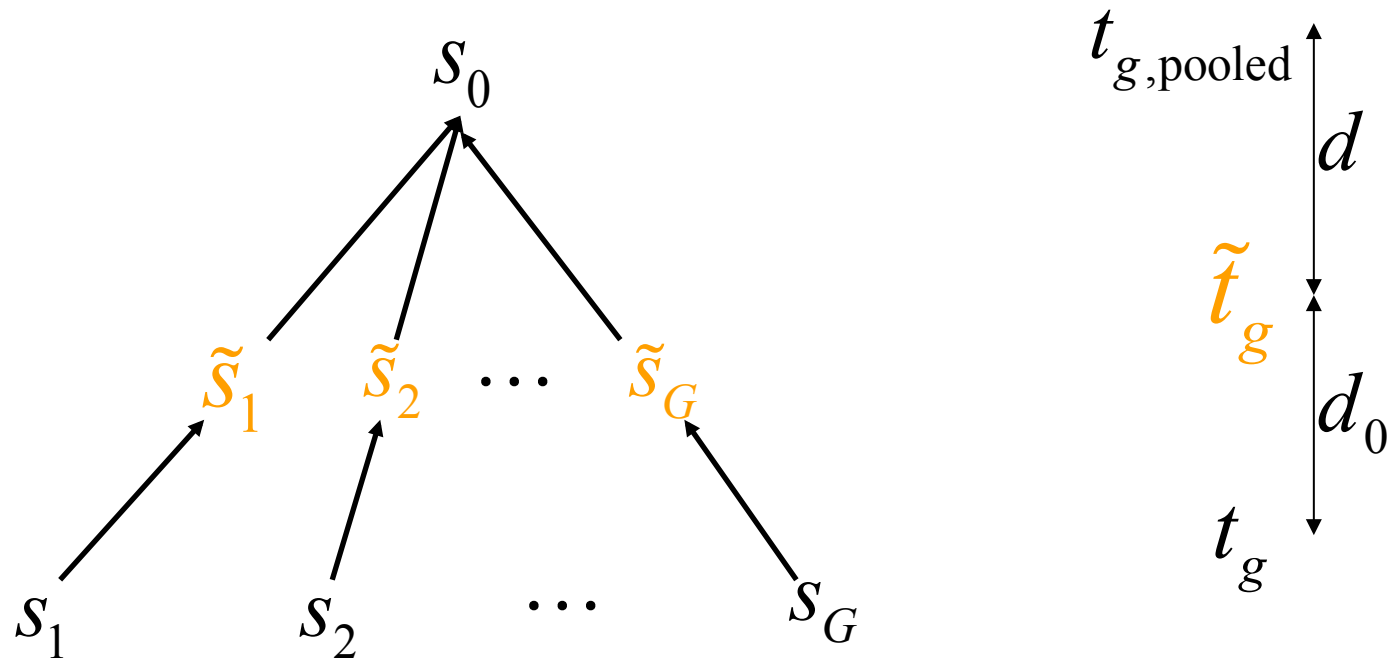
Moderated t-statistics

$$\tilde{t}_g = \frac{\bar{y}_{\text{mu}} - \bar{y}_{\text{wt}}}{\tilde{s}_g u}$$





Shrinkage of standard deviations



The **data decides** whether \tilde{t}_g should be closer to $t_{g,\text{pooled}}$ or to t_g



Why does it work?

- We learn what is the **typical** variability level by looking at all genes, but allow some **flexibility** from this for individual genes
- Adaptive – data (through hyperparameter estimates, d_0 and s_0) suggests how much to “squeeze” toward common value



Hierarchical model for variances

Data

$$s_g^2 \sim \sigma_g^2 \frac{\chi_{d_g}^2}{d_g}$$

Prior

$$\frac{1}{\sigma_g^2} \sim s_0^2 \frac{\chi_{d_0}^2}{d_0}$$

Posterior

$$E\left(\frac{1}{\sigma_g^2} \mid s_g^2\right) = \frac{d_0 + d_g}{s_0^2 d_0 + s_g^2 d_g}$$



Posterior Statistics

Posterior variance estimators

$$\tilde{s}_g^2 = \frac{s_0^2 d_0 + s_g^2 d_g}{d_0 + d_g}$$

Moderated t-statistics

$$\tilde{t}_{gj} = \frac{\hat{\beta}_{gj}}{\tilde{s}_g \sqrt{c_{gj}}}$$

Baldi & Long 2001, Wright & Simon 2003, Smyth 2004



Exact distribution for moderated t

An **unexpected piece of mathematics** shows that, under the null hypothesis,

$$\tilde{t}_g \sim t_{d_0 + d_g}$$

The degrees of freedom add!

The Bayes prior in effect adds d_0 extra arrays for estimating the variance.

Wright and Simon 2003, Smyth 2004



Aside: Marginal Distributions to calculate

Under usual likelihood model, s_g is independent of the estimated coefficients.

Under the hierarchical model, s_g is independent of the moderated t-statistics instead

$$s_g^2 \sim s_0^2 F_{d, d_0}$$



Multiple testing and adjusted p-values

- Each statistical test has an associated false error rate
- Traditional method in statistics is to control family wise error rate, e.g., by Bonferroni.
- Controlling the false discovery rate (FDR) is more **appropriate** in microarray studies
- Benjamini and Hochberg method controls expected FDR for independent or weakly dependent test statistics. Simulation studies support use for genomic data.
- All methods can be implemented in terms of adjusted p-values.



Linear Models

- In general, need to specify:
 - Dependent variable
 - Explanatory variables (experimental design, covariates, etc.)
- More generally:

$$y = X\beta + \epsilon$$

vector of
observed
data

design
matrix

Vector of
parameters to
estimate



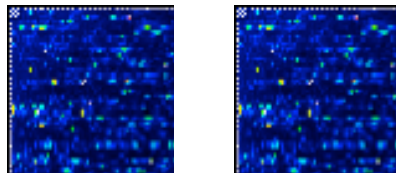
Linear Models for microarrays

- Analyse all arrays together combining information in optimal way
- Combined estimation of precision
- Extensible to arbitrarily complicated experiments
- **Design matrix**: specifies RNA targets used on arrays
- **Contrast matrix**: specifies which comparisons are of interest

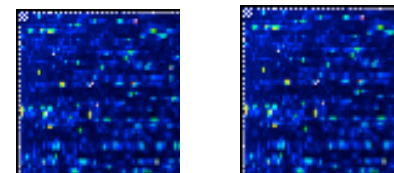


Design → Linear models

WT x 2



Mutant x 2



$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 1 & 0 \\ 1 & 1 \\ 1 & 1 \end{bmatrix} \begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix}$$

β_1 = wt log-expression

β_2 = mutant - wt

$$E[y_1] = E[y_2] = \beta_1$$

$$E[y_3] = E[y_4] = \beta_1 + \beta_2$$