

# Water fraction uncertainty, part 2

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Previously I showed that the posterior distribution for water fraction,  $\alpha$ , is Inv-Gamma distributed, having expectation  $\frac{\frac{N_L}{N_L-1}p - \mu_l}{\mu_w - \mu_l}$  and variance  $\frac{N_L^2 p^2}{(N_L-1)^2(N_L-2)(\mu_w - \mu_l)^2}$ . Here I'll use those equations to propagate uncertainty from the  $\mu_l$  and  $\mu_w$  terms.

Using the law of total variance we have

$$\begin{aligned} Var(\alpha) &= Var(E[\alpha|\mu_w, \mu_l]) + E[Var(\alpha|\mu_w, \mu_l)] \\ &= Var\left(\frac{\frac{N_L}{N_L-1}p - \mu_l}{\mu_w - \mu_l}\right) + \frac{N_L^2 p^2}{(N_L-1)^2(N_L-2)} E\left[\frac{1}{(\mu_w - \mu_l)^2}\right] \end{aligned}$$

Next, defining  $W = \mu_w - \mu_l$ , continue with another total-variance decomposition of the first term.

$$\begin{aligned} Var\left(\frac{\frac{N_L}{N_L-1}p - \mu_l}{W}\right) &= Var\left(E\left[\frac{\frac{N_L}{N_L-1}p - \mu_l}{W} \middle| W\right]\right) + E\left[Var\left(\frac{\frac{N_L}{N_L-1}p - \mu_l}{W} \middle| W\right)\right] \\ &= Var\left(\frac{\frac{N_L}{N_L-1}p - E[\mu_l]}{W}\right) + E\left[\frac{1}{W^2} Var(\mu_l)\right] \\ &= \left(\frac{N_L}{N_L-1}p - E[\mu_l]\right)^2 Var\left(\frac{1}{W}\right) + Var(\mu_l) E\left[\frac{1}{W^2}\right] \end{aligned}$$

Combining with the first result, this gives a formula for  $Var(\alpha)$  that is a function of the mean and variance of two variables,  $\mu_l$  and  $\frac{1}{W}$  (since  $E[\frac{1}{W^2}] = Var(\frac{1}{W}) + (E[\frac{1}{W}])^2$ ).

To go further, we need to express the moments of  $\frac{1}{W}$  in terms of the moments of  $\mu_w$  and  $\mu_l$ . The only way I can think to do that is to approximate  $W$  as a gamma-distributed random variable, so that  $\frac{1}{W}$  is inverse-gamma distributed. I don't know exactly how well this approximation will work, but generally it will be better if  $SD(\mu_w - \mu_l)$  is small compared to  $E(\mu_w - \mu_l)$ .

Define  $\mu_* = E[\mu_w] - E[\mu_l]$  and  $\sigma_*^2 = Var(\mu_w) + Var(\mu_l)$ . Then, method of moments gives the  $k, \theta$  parameters of  $W$ 's approximated Gamma distribution as  $\frac{\mu_*^2}{\sigma_*^2}$  and  $\frac{\sigma_*^2}{\mu_*}$ , respectively. Thus by the inverse gamma distribution we arrive at

$$E\left[\frac{1}{W}\right] = \frac{\mu_*}{\mu_*^2 - \sigma_*^2}; Var\left(\frac{1}{W}\right) = \frac{\mu_*^2 \sigma_*^2}{(\mu_*^2 - \sigma_*^2)^2 (\mu_*^2 - 2\sigma_*^2)}$$

## Conclusion

The final reassembled result for variance of water fraction is as follows:

$$\begin{aligned} Var(\alpha) &= \left(\frac{N_L}{N_L-1}p - E[\mu_l]\right)^2 Var\left(\frac{1}{W}\right) \\ &\quad + Var(\mu_l) E\left[\frac{1}{W^2}\right] \\ &\quad + \frac{N_L^2 p^2}{(N_L-1)^2(N_L-2)} E\left[\frac{1}{W^2}\right] \end{aligned}$$

where  $Var\frac{1}{W}$  is computed as above and  $E\left[\frac{1}{W^2}\right] = Var\left(\frac{1}{W}\right) + (E[\frac{1}{W}])^2 = \frac{\mu_*^2}{(\mu_*^2 - \sigma_*^2)(\mu_*^2 - 2\sigma_*^2)}$ .