

# ORF526 - Problem Set 7

Bachir EL KHADIR

November 19, 2015

## Question 1

1. It is easy to check that  $C$  is symmetric and that:

$$C(s, t) = \begin{cases} \min(|s|, |t|) & \text{if } ts \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

Let  $(t_i)_1^n \in \mathbb{R}$ , and  $f_i = 1_{(0, t_i)}$  where  $(0, t) = (t, 0)$  if  $t < 0$ , then:  $C(t_i, t_j) = \int_{\mathbb{R}} f_i f_j$  which is a scalar product in  $L_2$ .  $C$  is definite positive semi-definite as a conclusion.

2.  $C(t, s) = \min(t, s)$  when  $t, s \geq 0$ .
3.
  - $Var(B_0) = 0$  so  $B_0 = 0$  as.
  - $B_t - B_s$  is normal because  $B_t$  is a gaussian process.  $E[B_t - B_s] = 0$ , and  $Var(B_t - B_s) = Var(B_t) + Var(B_s) + 2Cov(B_t, B_s) = |t| + |s| + 2C(t, s) = |t - s|$
  - $Cov(B_t - B_s, B_u - B_v) = C(t, u) + C(s, v) - C(s, u) - C(t, v) = \frac{1}{2}(|t| + |u| - |t - u| + |s| + |v| - |s - v| - |s| - |u| + |s - u| - |t| - |v| + |t - v|) = \frac{1}{2}(u - t + v - s + u - s + v - t) = 0$ , and since the 2d process  $B_t - B_s, B_u - B_v$  is gaussian, its componenets are independent.

## Question 2

Let's call  $C_1$  the function  $C$  defined on question 1.

1.  $C(u, v) = C_1(u_1, v_1)C_1(u_2, v_2) = \int_{\mathbb{R}^2} 1_{(0, u_1)}(x)1_{(0, v_1)}(x)1_{(0, u_2)}(y)1_{(0, v_2)}(y)dxdy = \int_{\mathbb{R}^2} 1_{(0, u_1) \times (0, u_2)} 1_{(0, v_1) \times (0, v_2)} = < 1_{(0, u_1) \times (0, u_2)}, 1_{(0, v_1) \times (0, v_2)} >$

So  $C$  is positive semi-definite.

2.  $C(u, v) = \min(u_1, v_1) \min(u_2, v_2)$  when  $u, v \geq 0$
3. if one componenet of  $u$  is 0, then  $Var(X_u) = C(u_2, u_2)C(u_1, u_1) = 0$ , ie  $X_u = 0$  as.
4.  $B_t = X_{(t, 1)}$  is a gaussian process.  $E[B_t] = 0$  and  $Cov(B_t, B_s) = Cov(X_{(t, 1)}, X_{(s, 1)}) = C_1(t, s)$ , so  $B_t$  is a two sided browninan motion
5.  $Var(X_{(t, t)}) = C_1(t, t)^2 = |t|^2$

## Question 3

Let's first show the following lemma: For every  $X \in L_1$ , there exist a sequence of simple function  $Z_n$  bounded by  $|X|$  and converging to  $X$ . To show that, we write  $X = X^+ - X^-$ , and let  $Z_n^+$  (resp.  $Z_n^-$ ) a sequence of positive simple functions converging to  $X^+$  (resp.  $X^-$ ) from below (resp. above). And set  $Z_n = Z_n^+ - Z_n^-$ , which verifies the lemma.

1.  $X$  is  $\mathcal{G}$  measurable, and trivially verifies the definition of conditional probability, so  $E[X|\mathcal{G}] = X$
2.  $aE[X|\mathcal{G}] + bE[Y|\mathcal{G}]$  is  $\mathcal{G}$ -measurable as sum of two functions that are  $\mathcal{G}$ -measurable, and if  $A \in \mathcal{G}$ :

$$\begin{aligned} E[(aE[X|\mathcal{G}] + bE[Y|\mathcal{G}])1_A] &= aE[E[X|\mathcal{G}]1_A] + bE[E[Y|\mathcal{G}]1_A] \\ &= aE[X1_A] + bE[Y1_A] && \text{because } A \text{ is } \mathcal{G}\text{-measurable} \\ &= E[(aX + bY)1_A] \end{aligned}$$

so  $E[aX + bY|\mathcal{G}] = aE[X|\mathcal{G}] + bE[Y|\mathcal{G}]$ .

3.  $E[X|\mathcal{G}] - E[Y|\mathcal{G}] = E[X - Y|\mathcal{G}]$  Let  $H := E[X - Y|\mathcal{G}]$ , and  $A := \{H \leq 0\}$ .  $A$  is  $\mathcal{G}$ -measurable and by positivity of the expectation:  $0 \geq E[H1_A] = E[(X - Y)1_A] \geq 0$ .

Since  $-H1_A \leq 0$  a.s and its expectation is 0,  $H1_A = H^- = 0$  as, and therefore  $H \geq 0$  as.

4. For  $A \in \mathcal{H} \subseteq \mathcal{G}$ ,  $E[E[X|\mathcal{G}]|H]$  is  $H$ -measurable and :

$$\begin{aligned} E[1_A E[E[X|\mathcal{G}]|H]] &= E[1_A E[X|\mathcal{G}]] \\ &= E[1_A X] \end{aligned}$$

5. Let  $A \in \mathcal{G}$ , and prove that  $E[1_A Y E[X|\mathcal{G}]] = E[1_A X Y]$  If we denote  $Z := 1_A Y$ , this is equivalent to  $E[Z E[X|\mathcal{G}]] = E[Z X]$ .

$Z$  is  $\mathcal{G}$ -measurable and  $|ZX| \leq |YX| \in L_1$

- If  $Z$  is a simple function  $\sum_{i=0..n} \alpha_i 1_{A_i}$ , where  $A_i \in \mathcal{G}$  for  $i = 0..n$ , then by linearity of the expectation:

$$E[Z E[X|\mathcal{G}]] = \sum_i \alpha_i E[1_{A_i} E[X|\mathcal{G}]] = \sum_i \alpha_i E[1_{A_i} X] = E[Z X]$$

- If  $X$  and  $Y$  are non-negative, Let  $Z_n$  be a sequence of non-negative simple  $\mathcal{G}$ -measurable functions s.t.  $Z_n \uparrow Z$  and therefore  $|Z_n X| \leq |ZX| \in L_1$ . By monotnous convergence theorem:

$$E[Z E[X|\mathcal{G}]] = \lim E[Z_n E[X|\mathcal{G}]] = \lim E[Z_n X] = E[Z X]$$

- $X$  now can be in  $L_1$ .

We use  $h$ ), to show that  $|E[X|\mathcal{G}]| < E[|X||\mathcal{G}]$ . (take  $\phi : x \rightarrow |x|$ )

Let  $Z_n$  a sequence of simple functions converging to  $Z$  and bounded by  $|Z|$ . Then  $|Z_n X| \leq |ZX| \in L_1$  and  $|Z_n E[X|\mathcal{G}]]| = |E[Z_n X|\mathcal{G}]]| \leq E[|XZ||\mathcal{G}] \in L_1$  because  $EE[|XZ||\mathcal{G}] = E[|XZ|] < \infty$ .

By dominated convergence theorem:

$$E[Z E[X|\mathcal{G}]] = \lim E[Z_n E[X|\mathcal{G}]] = \lim E[Z_n X] = E[Z X]$$

- If  $Y \in L_1$ ,  $Z = Z^+ - Z^-$ , and by linearity

$$E[Z E[X|\mathcal{G}]] = E[Z^+ E[X|\mathcal{G}]] - E[Z^- E[X|\mathcal{G}]] = EE[XZ^+|\mathcal{G}] - EE[XZ^-|\mathcal{G}] = E[XZ^+] - E[XZ^-] = E[XZ]$$

6. Let's first prove that if  $A \in \mathcal{G}$ ,  $E[X1_A] = E[X]E[1_A]$ .

- (a) If  $X$  is an indicator function, then it follows from the definition of independence
- (b) If  $X$  is a simple function it follows from the linearity of the expectation.

- (c) If  $Z_n$  a sequence of simple functions converging to  $X$  and uniformly bounded by an  $|X|$ , then by CVD:

$$E[X1_A] = \lim E[Z_n 1_A] = \lim E[Z_n]E[1_A] = \lim E[X]E[1_A]$$

So now we have:

$$E[1_A X] = E[1_A]E[X] = E[1_A E[X]]$$

$E[X]$  is a constant, so  $\mathcal{G}$ -measurable.

7.

$$E[X1_\emptyset] = 0 = E[X]E[1_\emptyset]$$

$$E[X1_\Omega] = E[X] = E[X]E[1_\Omega]$$

so  $X$  is independent of  $\mathcal{G}$ , and therefore  $E[X|\mathcal{G}] = E[X]$ .

8. If  $\varphi$  is affine  $= ax + b$ , then by linearity  $E[\varphi(X)|\mathcal{G}] = \varphi(E[X|\mathcal{G}])$

If  $\varphi$  is convex not linear, we can write  $\varphi = \sup \varphi_n = \sup_n a_n x + b_n$  where  $a_n, b_n \in R$ , then  $\forall n E[\varphi(X)|\mathcal{G}] \geq E[\varphi_n(X)|\mathcal{G}] \geq \varphi_n(E[X|\mathcal{G}])$  as. Let  $\Omega_n$  the set where this equality holds, so on  $\Omega' := \cap_n \Omega_n$  we have that:

$$E[\varphi(X)|\mathcal{G}] \geq \sup_n \varphi_n(E[X|\mathcal{G}]) = \varphi(E[X|\mathcal{G}]) \text{ on } \Omega'$$

and  $P(\Omega') = 1 - P(\cup_n \Omega_n^c) \geq 1 - \sum_n P(\Omega_n^c) \geq 1$

#### Question 4

- $E[X_n|Y]$  is non-decreasing, let's call  $L := \lim E[X_n|Y]$ , and prove that  $L = E[X|\mathcal{G}]$ .

Since  $Y \leq X_n \uparrow X$ ,  $Y \wedge n \leq X \wedge n \uparrow X$  and  $E[Y|\mathcal{G}] \leq E[X_n|\mathcal{G}] \uparrow L$ , by monotonous convergence theorem, for all  $A \in \mathcal{G}$ :

$$\begin{aligned} E[1_A L] &= \lim_n E[1_A E[X_n|\mathcal{G}]] \\ &= \lim_n E[1_A X_n] \\ &= E[1_A X] \\ &= \lim E[1_A (X \wedge k)] \\ &= \lim E[1_A E[X \wedge k]] \\ &= E[1_A E[X|\mathcal{G}]] \end{aligned}$$

Let  $B \in \mathcal{B}(R)$ , and for  $a > 0$ ,  $A = B \cap \{|L| < a, |E[X|\mathcal{G}]| < a\}$ . And now we have:

$$0 = E[1_A (L - E[X|\mathcal{G}])] = E[1_B (L - E[X|\mathcal{G}]) 1_{|L| < a, |E[X|\mathcal{G}]| < a}]$$

By taking  $B$  to be the set where  $L - E[X|\mathcal{G}] > 0$  and then  $L - E[X|\mathcal{G}] < 0$ , we have that  $(L - E[X|\mathcal{G}]) 1_{|L| < a, |E[X|\mathcal{G}]| < a} = 0$ , and by taking  $a$  to  $\infty$ , we have that  $L = E[X|\mathcal{G}]$ .

- Let's define  $L_k := \inf_{n \geq k} X_n \leq X_k$ , so that

$$E[L_k|\mathcal{G}] \leq E[X_k|\mathcal{G}] \tag{1}$$

But  $Y \leq L_k \uparrow \liminf_n X_n$ , by a)  $E[L_k|\mathcal{G}] \uparrow_k E[\liminf_n X_n|\mathcal{G}]$ , and by taking the  $\liminf$  in the inequality 1 we have the result.

- $X_n$  and  $-X_n$  verify the conditions of the last question, so:

$$\liminf E[-X_n|\mathcal{G}] \geq E[\liminf -X_n|\mathcal{G}] \Rightarrow \limsup E[X_n|\mathcal{G}] \leq E[\limsup X_n|\mathcal{G}]$$

$$\liminf E[X_n|\mathcal{G}] \geq E[\liminf X_n|\mathcal{G}]$$

as a result

$$E[\limsup X_n|\mathcal{G}] \geq \limsup E[X_n|\mathcal{G}] \geq \liminf E[X_n|\mathcal{G}] \geq E[\liminf X_n|\mathcal{G}]$$

Since  $\limsup X_n = \liminf X_n = X$ , we have the result.

### Question 5

- $y \rightarrow p(y, A)$  is measurable because:

1.  $(x, y) \rightarrow f(x, y)$  is measurable since it is a density
2.  $N(y)$  is measurable by Fubini
3.  $y \rightarrow \int \frac{f(x, y)}{N(y)} 1_{0 < N(y) < \infty} + (1 - 1_{0 < N(y) < \infty})\varphi(x)$  is also measurable by Fubini

$p(Y, A)$  is then  $\sigma(Y)$ -measurable.

- We have that:

$$N(y)f_y(x) = \begin{cases} N(y)\phi(x) = 0 & \text{if } N(y) = 0 \\ N(y)\phi(x) & \text{if } N(y) = \infty \\ f(x, y) & \text{otherwise} \end{cases}$$

But since  $N \in L_1$  (By Fubini), the set  $\{N = \infty\}$  is of measure 0, so  $N(y)f_y(x) = 1_{N(y) \neq 0}f(x, y)$  a.s.

Let  $B \in \mathcal{B}(R)$ , all function integrated below are non negative, so:

$$\begin{aligned} E[p(Y, A)1_{Y \in B}] &= \int_{R^2} p(y, A)1_{y \in B}f(x, y)dxdy \\ &= \int p(y, A)1_{y \in B}N(y)dy && \text{By Tonelli} \\ &= \int_y 1_{y \in B} \int_x 1_{x \in A}N(y)f_y(x)dxdy \\ &= \int 1_{y \in B}1_{x \in A}1_{N(y) \neq 0}f(x, y)dxdy \\ &= \int 1_{y \in B}1_{x \in A}f(x, y)dxdy && \text{because if } N(y) = 0 \text{ then } \int_A f(x, y) = 0 \\ &= E[1_{Y \in B}1_{X \in A}] \\ &= E[1_{Y \in B}P(X \in A|Y)] \end{aligned}$$

Which prove the result.