

ORF527 - Problem Set 2

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Q.1

Let $\varepsilon > 0, x \in [0, 1]$.

- $\sup |f^n - f| \rightarrow_n 0$, let $n \in \mathbb{N} \forall t \in [0, 1] |f^n(t) - f(t)| \leq \frac{\varepsilon}{3}$.
- f^n is continuous, let $\delta > 0$ such that $\forall y \in [0, 1], |x - y| < \delta \Rightarrow |f^n(x) - f^n(y)| < \frac{\varepsilon}{3}$.
- Let $y \in [0, 1]$ such that $|x - y| < \delta$

$$\begin{aligned} |f(x) - f(y)| &\leq |f(x) - f^n(x)| + |f^n(x) - f^n(y)| + |f^n(y) - f(y)| \\ &\leq 3\frac{\varepsilon}{3} \\ &\leq \varepsilon \end{aligned}$$

Which conclude the proof.

Q.2

3.2

let $X \sim \mathcal{N}(0, 1)$, and $\varepsilon \sim \mathcal{B}(-1, 1, \frac{1}{2})$ be two independant rv. And Let $Y = \varepsilon X$

- By symmetry of the distribution of X : $F_Y(y) = P(Y \leq y) = P(\varepsilon X \leq y) = E[P(\varepsilon X \leq y|\varepsilon)] = \frac{1}{2}P(X \leq y) + \frac{1}{2}P(-X \leq y) = P(X \leq x)$ so $Y \sim \mathcal{N}(0, 1)$.
- $cov(X, Y) = E[X^2\varepsilon] = E[X^2]E[\varepsilon] = 0$
- X, Y are not independent. Indeed, Let $\alpha := \mathbb{P}(|X| > 0.5)$.
 - $P(|X| > 0.5, |Y| > 0.5) = P(|X| > 0.5) = \alpha$,
 - $P(|X| > 0.5)P(|Y| > 0.5) = P(|X| > 0.5)^2 = \alpha^2$

But since $\alpha \notin \{0, 1\}$, $\alpha \neq \alpha^2$.

3.3

For a random variable X , let's call Φ_X its characteristic function.

(a) For $t \in \mathbb{R}^n$, $\Phi_{AV}(t) = \mathbb{E}[e^{iV^T A^T t}] = \Phi_V(A^T t) = e^{(A\mu)^T V - \frac{1}{2}(t^T A \Sigma A^T)t}$, so $V \sim \mathcal{N}(A\mu, A\Sigma A^T)$

(b) By symmetry of the gaussian distribution, $-Y$ has the same distribution as Y . So it suffices to show that the result holds for $X + Y$. By independence: $\Phi_{X+Y}(t) = \Phi_X(t)\Phi_Y(t) = \Phi_X(t)^2 = e^{i2\mu t - \frac{1}{2}2t^2}$, so $X + Y \sim \mathcal{N}(0, 2)$.

(c) If $\text{cov}(X, Y) = 0$, the covariance matrix of the gaussian process (X, Y) is the identity, therefore

$$\forall t = (t_1, t_2) \in \mathbb{R}^2 \quad \Phi_{(X,Y)}(t) = E[e^{i\mu_X t_1 + i\mu_Y t_2 - \frac{1}{2}\sigma_X^2 t_1^2 - \frac{1}{2}\sigma_Y^2 t_2^2}] = \Phi_X(t_1)\Phi_Y(t_2)$$

Where $X \sim \mathcal{N}(\mu_X, \sigma_X^2), Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$. Therefore X, Y are independent.

(d) Let $\alpha := \frac{\text{cov}(X,Y)}{\text{var}(X)}$. $(Y - \alpha X, X)$ is gaussian and $\text{cov}(Y - \alpha X, X) = 0$. So $Y - \alpha X \perp X$. Therefore:

$$\mu_{Y|X=x} = E[Y|X=x] = E[Y - \alpha X] + E[\alpha X|X=x] = \mu_Y - \alpha\mu_X + \alpha x = \mu_Y + \frac{\text{cov}(X,Y)}{\text{var } X}(x - \mu_X)$$

$$\sigma_Y^2 = \text{var}(Y|X=x) = \text{var}(Y - \alpha X) + \underbrace{\text{var}(\alpha X|X=x)}_{=0} = \text{var } Y + \alpha^2 \text{var } X - 2\alpha \text{cov}(X, Y) = \sigma_Y^2 - \frac{\text{cov}(X, Y)^2}{\text{var } X}$$

Q.3

(a) Since $\forall n > 0, \Delta_n(1) = 0$, $B_1 = \sum_{n=0}^{\infty} \lambda_n Z_n \Delta_n(1) = \lambda_0 Z_0 \Delta_0(1)$.

(b) If $s < t$, then $\text{cov}(B_s, B_t) = \text{cov}(B_s - B_t, B_s) + \text{var}(B_s) = s$.

$$\begin{aligned} \text{cov}(U_s, U_t) &= \text{cov}(B_s - sB_1, B_t - tB_1) = \text{cov}(B_s, B_t) + ts \text{cov}(B_1, B_1) - t \text{cov}(B_s, B_1) - s \text{cov}(B_1, B_t) \\ &= s + ts - ts - ts = s - ts = s(1 - t). \end{aligned}$$

(c) For $s < t$, $\text{cov}(X_t, X_s) = g(t)g(s)h(s)$ should be equal to $s(1 - t)$. We can take $g : t \rightarrow 1 - t$, $h : t \rightarrow \frac{s}{1-s}$ defined on $[0, 1)$