#### Class 5 Outline

- 1. Time to event (survival) data, censoring survival times and the survivor function
- 2. The survivor function and the hazard rate
  - Uncensored data, Censored data
- 3. Kaplan-Meier estimates of the survivor function, S(t), for ungrouped survival data
- 4. Example using AML Data
- 5. Summary

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### 0. Learning Objectives

- Describe ungrouped survival data in which the exact time to event or censoring is known
- Describe the survival function and hazard function
- Describe how to estimate the survivor function using the Kaplan-Meier survival curve and confidence intervals for ungrouped survival data

Key words - censoring, survival function, hazard function, Kaplan-Meier estimates

#### 1. Time to Event Data

- The outcome (event) of interest is dichotomous
- The study design may not be able to assure that the outcome is know for all individuals at the endpoint of the study.
- Uncensored data: The event has occurred
- Censored data: The event has yet to occur
  - Event-free at the current follow-up time
  - A competing event that is not an endpoint stops follow-up
  - Death (if not part of the endpoint)
  - Clinical event that requires treatment, etc.

3

#### 1.1 Survival Times

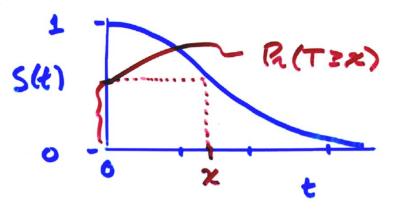
- Distribution of times to event often called "survival times," even when the "event" is not "death"
- Survival times follow a continuous distribution with times ranging from zero to infinity
- The probability distribution of the survival times can be described by:
  - cumulative distribution function
  - density function
  - survivor function = 1 cumulative distribution function
  - hazard function

## 2. Survivor Function, S(t)

- The survival function, denoted S(t), is a useful way to represent the probability distribution of the survival time T, when some of the observed times are censored – only know that T>t, rather than T=t
- S(t) = Pr(T > t)  $0 < t < \infty$
- *S*(*t*) is the probability of surviving beyond *t*

5





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### 2.2a Handling Uncensored Survival Data

- · Use the "usual" methods
  - t-tests
  - regression
  - ANOVA
  - transforms
- The survival curve as an important and complete summary of a population

$$S(t) = \frac{\text{("alive"}^* \text{ at followup time t)}}{\text{(alive at time 0)}}$$

$$0 = \begin{cases} \text{Date of randomization} \\ \text{Birth} \end{cases}$$

\* Not had the event

7

### 2.2b Handling Uncensored Survival Data

- The survival curve starts at 1.0 and decreases over time
- Estimating these curves and comparing them among groups constitutes a "survival analysis"
- Need to decide on what summary measure is important
  - Mean survival time
  - Median survival time
  - Value at a specific time: S(12)
  - Difference of curves: S1(12) S2(12)
  - Maximal difference

### 2.3 Example with Uncensored Data

• Data: 2, 3, 3, 5, 6, 9, 9, 10, 13, 16

	6
Time	S(t)
0	1.00
1	1.00
2	0.90
3	0.70
4	0.70
5	0.60
6	0.50
7	0.50
8	0.50
9	0.30
10	0.20
11	0.20
12	0.20
13	0.10
14	0.10
15	0.10
16	0.00

9

### 2.4 Ways of Handling Censoring

- 1. Ignore the incomplete cases; drop them (never!)
  - Produces bias in the estimated curve
  - Unbalanced censoring produces biased comparisons
- 2. Impute a missing event time
  - Depends on a detailed probability model
- 3. Calculate the overall event rate
- 4. Use the available information on each participant
  - Important issue: If no events are reported in the interval from last follow-up to "now", we need to choose between:
    - No news is good news?
    - · No news is no news?

#### 2.5 Overall Event Rate

Overall event rate:

Event rate =  $\frac{\text{# events}}{\text{total observation time}}$ 

Example: 5 events in 600 person months

= 1 event per 120 months

= 1 event per 10 years

= 0.1 events per year

= 10 events per 100 person-years

- Gives an <u>average event rate</u> over the follow-up period; actual event rate may vary over time
- For a finer time resolution, use small intervals,

# 3.0 Kaplan-Meier Estimates of the Survivor Function, S(t)

Biostatistics Trivia: Professor Paul Meier was an assistant professor in the JHU Department of Biostatistics from 1952 to 1957. He teamed with E.L. Kaplan to write their seminal paper "Non-parametric Estimation from Incomplete Observations," which appeared in the *Journal of the American Statistical Association* in 1958. This paper was to lay the groundwork for modern survival analysis.

#### 3.1a The Hazard Function

- Basic idea: Estimate the hazard of death at each event time t using available data and then use them to produce the survival curve by multiplying (1 - hazard) terms
- The hazard =
   Pr(event "now" | no event yet) )/unit time
   where "now" means in the current unit interval
- Thus, (1-hazard) = Pr(no event "now" | no event yet)

13

#### 3.1b The Hazard Function

- The hazards for time intervals i=1, 2, 3, ... are h<sub>1</sub>, h<sub>2</sub>, h<sub>3</sub>,...
- Example:  $S(3) = (1 h_1) \times (1 h_2) \times (1 h_3)$
- If the hazard is large, the survival curve decreases rapidly
- Estimate  $\mathbf{h}_i = \frac{\text{\# events at t=i}}{\text{\# at risk at t=i}} = \frac{\mathbf{Y}_i}{\mathbf{n}_i}$
- Kaplan-Meier estimate:  $\hat{S}(t) = \prod_{i \text{ for } t_i < t} (1-h_i)$

# 3.2 Relationship between the Survivor and Hazard Functions

#### Example:

S(3) = Pr(survive for 3 months)
= Pr(survive 1st month) ×
Pr(survive 2nd month | survive 1st month) ×
Pr(survive 3rd month | survive 2nd month)

• Thus,  $S(3) = S(1) \cdot \frac{S(2)}{S(1)} \cdot \frac{S(3)}{S(2)}$   $S(3) = S(1) \cdot \frac{S(2)}{S(1)} \cdot \frac{S(3)}{S(2)}$   $S(3) = (1-h_1) \cdot (1-h_2) \cdot (1-h_3)$ <sup>15</sup>

# 3.3a Calculating the Hazard – Uncensored Data

• Data: 2, 3, 3, 5, 6, 9, 9, 10, 13, 16 (uncensored data)

$$\begin{array}{lll} h_1 = 0 & (0/10) \\ h_2 = 0.10 & (1/10) \\ h_3 = 0.22 & (2/9) \\ h_4 = 0 & (0/7) \\ h_5 = 0.14 & (1/7) \\ h_6 = 0.17 & (1/6) \\ h_7 = h_8 = 0 & (0/5) \\ h_9 = 0.40 & (2/5) \\ h_{10} = 0.33 & (1/3) \\ h_{11} = h_{12} = 0 & (0/2) \\ h_{13} = 0.50 & (1/2) \\ h_{14} = h_{15} = 0 & (0/1) \\ h_{16} = 1.00 & (1/1) \end{array}$$

### 3.3b Calculating the Survivor Function -**Uncensored Data**

Uncensored (complete) data example

$$S(3) = (1 - h_1)(1 - h_2)(1 - h_3)$$

· Estimate by

$$= (1 - y_1/n_1)(1 - y_2/n_2)(1 - y_3/n_3)$$

$$= (1 - 0/10)(1 - 1/10)(1 - 2/9)$$

$$= 1 \times 9/10 \times 7/9 = 0.70$$

where y<sub>i</sub> is the number of events at time i and n<sub>i</sub> is the number at risk at time i

#### 3.4 Hazard Function and the Survival Curve

· If we know the hazard function, we know the survival curve:

$$S(t_i) = \prod_{i \le t} (1 - h_i) = (1 - h_i)S(t_{i-1})$$

· If we know the survival curve, we know the hazard function:

cumulative hazard = -log[S(t)]

hazard = increments of cumulative hazard

# 3.5a Calculating the Hazard – Censored Data

Data: 2, 3, 3\*, 5, 6\*, 9, 9\*, 10, 13, 16 (\* = censored at end of the interval)

```
h1 = 0
                     (0/10)
h2 = 0.10
                     (1/10)
h3 = 0.11
                     (1/9)
h4 = 0
                     (0/7)
h5 = 0.14
                     (1/7)
h6 = 0
                     (0/6)
h7 = h8 = 0
                     (0/5)
h9 = 0.20
                     (1/5)
h10 = 0.33
                     (1/3)
h11 = h12 = 0
                     (0/2)
h13 = 0.50
                     (1/2)
h14 = h15 = 0
                     (0/1)
h16 = 1.00
                     (1/1)
```

# 3.5b Calculating the Survivor Function Censored Data

· Censored data example

$$S(3) = (1 - h_1)(1 - h_2)(1 - h_3)$$

Estimate by

$$= (1-0)(1-0.10)(1-0.11)$$

$$= 1 \times .90 \times 8/9$$

$$= .80$$

20

### 4. Example- AML Data

- Example: Acute Myelogenous Leukemia (AML)
  - Y time from start of treatment to cancer relapse
  - X Indicates chemotherapy group (=0 for not maintained on chemotherapy group and

=1 for maintained on chemotherapy)

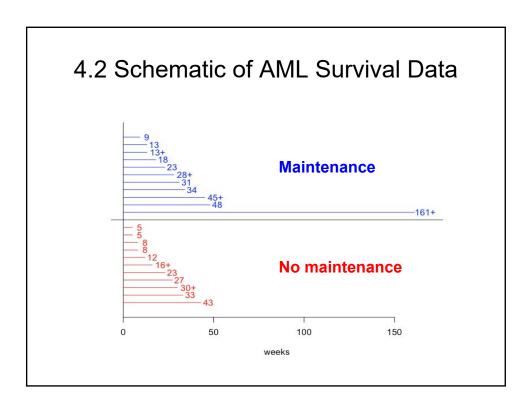
21

### 4.1 Survival Data - AML Example

 Consider a clinical trial in patients with acute myelogenous leukemia (AML) comparing two groups of patients: no maintenance treatment with chemotherapy (X=0) -vs- maintenance chemotherapy treatment (X=1)

Group	Weeks in remission ie, time to relapse
Maintenance chemo (X=1)	9, 13, 13+, 18, 23, 28+, 31, 34, 45+, 48, 161+
No maintenance chemo (X=0)	5, 5, 8, 8, 12, 16+, 23, 27, 30+, 33, 43, 45

 + indicates a censored time to relapse; e.g., 13+ = more than 13 weeks to relapse



#### 4.3a Stata Commands for Survival Data

- There are many Stata commands for input, management, and analysis of survival data, most of which are found in the manual in the st section – all survival data commands start with st
- st can be used to analyze individual level data (Kaplan-Meier, Cox regression, etc) or to group the individual level data for grouped analysis (SMRs, output for Poisson regression, etc.)
- Stata 15 Reference manual

#### 4.3b Stata Commands for Survival Data

- With ungrouped survival data on individuals:
  - 1. Use the ordinary **Stata** input commands to input and/or generate the following variables:
  - X variables
  - Person-time (denominator) variable (if applicable)
  - Time variable containing follow-up time
  - Censoring variable indicating status at the end of follow-up either "failed" or "censored"
  - 2. Then, use the *st* commands, as illustrated, to process and analyze the data

25

#### 4.3c Stata Commands for Survival Data

 Define survival data: stset command

Used to define the time variable, the status variable with the codes for "failures," and an "id" variable the uniquely identifies each individual observation

stset t , failure(failed==1) id(id)

Descriptive statistics for survival data:

stdes, stsum command

# 4.4 Listing of AML Data

. list id Chemo time failed

•	id	Chemo	time	failed
1.	j 1	1	9	1
2.	2	1	13	1
3.	3	1	13	0
4.	4	1	18	1
•				
•				
•				
•				
•				
19.	19	0	27	1
20.	20	0	30	0
21.	21	0	33	1 j
22.	22	0	43	1
23.	23	0	45	1
	<b>.</b>			+

27

## 4.5 Defining Survival Data

.stset time, failure(failed==1) id(id)

id: id
failure event: failed == 1
obs. time interval: (time[\_n-1], time]
exit on or before: failure

23 total obs.

0 exclusions

22 obs. maniping managementing

23 obs. remaining, representing

23 subjects

17 failures in single failure-per-subject data

678 total analysis time at risk, at risk from t = 0
earliest observed entry t = 0
last observed exit t = 161

# 4.6 Description of Survival Data

. stdes if Chemo==0

failure \_d: failed == 1
analysis time \_t: time
 id: id

			per sub	er subject	
Category	total	mean	min	median	max
no. of subjects	12				
no. of records	12	1	1	1	1
(first) entry time		0	0	0	O
(final) exit time		21.25	5	19.5	45
subjects with gap	0				
time on gap if gap	0	•	•	•	
time at risk	255	21.25	5	19.5	45
failures	10	.8333333	0	1	1
					 29

# 4.7b Summary of Survival Data

.stsum

failure \_d: failed == 1 analysis time \_t: time id: id

		incidence	no. of	:	Survival ti	me
	time at risk	rate	subjects	25%	50%	75%
total	678	.0250737	23	12	27	43

# 4.7b Description of Survival Data

. stdes if Chemo==1

failure \_d: failed == 1
analysis time \_t: time
 id: id

Category	total	mean	per subj min	ect median	max
no. of subjects	11				
no. of records	11	1	1	1	1
(first) entry time		0	0	0	0
(final) exit time		38.45455	9	28	161
subjects with gap	0				
time on gap if gap	0	•		•	
time at risk	423	38.45455	9	28	161
failures	7	.6363636	0	1	1

31

# 4.8 Overall Incidence Rates by Group

note: Exposed <-> Chemo==1 and Unexposed <-> Chemo==0

	Chemo   Exposed Unexposed	   Total	
Failure Time	7 10 423 255	17   678	
Incidence Rate	   .0165485 .0392157	.0250737	
	Point estimate	95% Conf. Int	erval]
Inc. rate diff.	  0226672	+  0498895 .	004555
Inc. rate ratio	.4219858	.1363296 1.	228119 (exact)
Prev. frac. ex.	.5780142	2281186 .8	636704 (exact)
Prev. frac. pop	.3606195	İ	
-	+		
	(midp) Pr(k <= 7) =		0.0418 (exact)
	(midp) 2*Pr(k<=7) =		0.0836 (exact) 32

# 4.9a Kaplan-Meier Estimates of the Survivor Function – Not Maintained Group

4.9b Kaplan-Meier Estimates of the Survivor Function – Not Maintained Group

Time	Number at Risk	Events	$S(t_i) = (1 - h_i)S(t_{i-1})$
0	12	0	1.0
5	12	2	1.0(1- 2/12) = 0.833
8	10	2	0.833(1- 2/10)=0.666
12	8	1	0.666(1-1/8) = 0.583
23	6	1	0.583(1- 1/6) =0.486
27	5	1	0.486(1- 1/5) =0.389
33	3	1	0.389(1- 1/3) = 0.259
43	2	1	0.259(1 -1/2) = 0.130
45	1	1	0.130(1-1/1)=0

# 4.10a Kaplan-Meier Estimates of the Survivor Function – Maintained Group

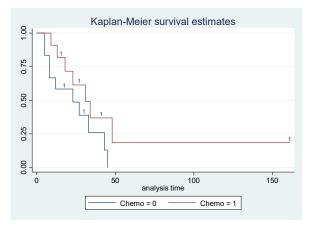
analy	failure sis time	_	ailed == : d	1			
	Beg.		Net	Survivor	Std.		
Time	Total	Fail	Lost	Function	Error	[95% Con	f. Int.
 9	11	1	 0	0.9091	0.0867	0.5081	0.986
13	10	1	1		0.1163		
18	8	1	0	0.7159	0.1397	0.3502	0.899
23	7	1	0	0.6136	0.1526	0.2658	0.835
28	6	0	1	0.6136	0.1526	0.2658	0.835
31	5	1	0	0.4909	0.1642	0.1673	0.753
34	4	1	0	0.3682	0.1627	0.0928	0.657
45	3	0	1	0.3682	0.1627	0.0928	0.657
48	2	1	0	0.1841	0.1535	0.0117	0.525
161	1	0	1	0.1841	0.1535	0.0117	0.525

# 4.10b Kaplan-Meier Estimates of the Survivor Function – Maintained Group

Time	Number at Risk	Events	$S(t_i) = (1-h_i)S(t_{i-1})$
0	11	0	1.0
9	11	1	1.0(1- 1/11) = 0.909
13	10	1	0.909(1- 1/10) = 0.818
18	8	1	0.818(1- 1/8) = 0.716
23	7	1	0.716(1- 1/7) = 0.614
31	5	1	0.614(1- 1/5) = 0.491
34	4	1	0.491(1- 1/4) = 0.368
48	2	1	0.368(1- 1/2) = 0.184

## 4.11 Graph of Kaplan Meier Survival Curves

- Estimates of the survival function S(t) versus time -separate curves for each group
- .sts graph, by(Chemo)lost or .sts graph, by(Chemo)risktable



37

### 5.a Summary

- There are statistical techniques for describing and making inferences for time to event data (survival times) in the presence of censoring:
  - Overall incidence rate, survivor function
- The survivor function S(t) represents the probability distribution of the survival times; S(t) is the probability of surviving beyond t
- The hazard function, h<sub>i</sub> =
   Pr(event "now" | no event yet)/unit time
   where "now" means in the current unit interval i

## 5.b Summary

• There is a relationship between the survivor function and hazard function in discrete time

$$S(t_i) = \prod_{i < t} (1 - h_i) = (1 - h_i)S(t_{i-1})$$

- The survivor function is a product of (1-hazard) terms
- Kaplan-Meier estimates of the survival curve for ungrouped data are calculated only at times that events occur