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# The Stochastic Container Relocation Problem

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**Abstract.** The container relocation problem (CRP) is concerned with finding a sequence of moves of containers that minimizes the number of relocations needed to retrieve all containers, while respecting a given order of retrieval. However, the assumption of knowing the full retrieval order of containers is particularly unrealistic in real operations. This paper studies the stochastic CRP, which relaxes this assumption. A new multistage stochastic model, called the *batch model*, is introduced, motivated, and compared with an existing model (the *online model*). The two main contributions are an optimal algorithm called *Pruning-Best-First-Search (PBFS)* and a randomized approximate algorithm called *PBFS-Approximate* with a bounded average error. Both algorithms, applicable in the batch and online models, are based on a new family of lower bounds for which we show some theoretical properties. Moreover, we introduce two new heuristics outperforming the best existing heuristics. Algorithms, bounds, and heuristics are tested in an extensive computational section. Finally, based on strong computational evidence, we conjecture the optimality of the “leveling” heuristic in a special “no information” case, where, at any retrieval stage, any of the remaining containers is equally likely to be retrieved next.

**Supplemental Material:** The online appendix is available at <https://doi.org/10.1287/trsc.2018.0828>.

**Keywords:** container relocation problem • block relocation problem • combinatorial optimization • multistage stochastic models • decision trees

## 1. Introduction

With the growth in international container shipping in maritime ports, there has been an increasing interest in improving operations in container terminals, both on the sea side and land side. The operations on the sea side involve the assignment of quay cranes to ships, the loading of export containers on vessels, and the discharging of import containers from vessels onto internal trucks. Import containers are then transferred to the land side and are stacked in the storage area. Operations on the land side (also called yard operations) include the routing of internal trucks within the yard, the stacking of containers for storage, and the delivery of import containers to external trucks for delivery to another location. This work focuses on the latter problem.

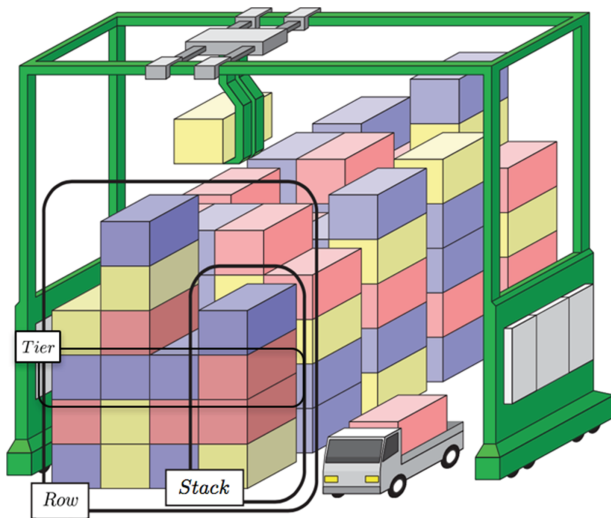
Because of limited space in the storage area, containers are stacked on top of each other. The resulting stacks create rows of containers as shown in Figure 1. If a container that needs to be retrieved (*target* container) is not located on a topmost tier and is covered by other containers, the blocking containers must be relocated to another stack. As a result, during the retrieval process, the yard cranes perform one or more *relocation* moves. Such relocations (also called reshuffles) are costly for the port operators and result in delays in the

retrieval process. Thus, reducing the number of relocations is one of the main goals of port operators. The *container relocation problem (CRP)* (also known as the *block relocation problem*) addresses this challenge by minimizing the number of relocations. As this problem is the main subject of this paper, we provide a formal definition and an extensive literature review.

### 1.1. The Container Relocation Problem

First, it is commonly known that the time to relocate a container within a row is insignificant compared to the time to relocate a container between two distinct rows. Therefore, in most cases, port operators tend to avoid relocations between rows. The CRP assumes that only relocations within rows are allowed, and problems for different rows should be considered independently. Furthermore, a row usually stores containers of the same type for the sake of stability and simplicity.

Using these facts, the CRP models one row using a two-dimensional array of size  $(T, S)$ , where  $S$  is the number of stacks and  $T$  is the maximum height, that is, the maximum number of containers in a stack as limited by the height of the crane. Each element of this array represents a potential slot for a container and contains a number only if a container is currently stored in this slot. Stacks are numbered from left (1) to right ( $S$ ) and tiers from bottom (1) to top ( $T$ ). We refer to this

**Figure 1.** (Color online) Illustration of Stacks of Containers in a Storage Yard

Source: Tanaka and Takii (2016).

array as a *configuration*. The common assumptions of the CRP are the following:

$A_1$ : The initial configuration has  $T$  tiers,  $S$  stacks, and  $C$  containers. For the problem to always be feasible, we suppose that the triplet  $(T, S, C)$  satisfies  $0 \leq C \leq ST - (T - 1)$  (see Caserta, Schwarze, and Voß 2012).

$A_2$ : A container can be retrieved/relocated only if it is in the topmost tier of its stack, that is, no other container is blocking it.

$A_3$ : A container can be relocated only if it is blocking the target container. This assumption was suggested by Caserta, Schwarze, and Voß (2012), and the problem under this assumption is commonly referred to as the *restricted* CRP. Most studies focus on this restricted version, because it is the current practice in many yards, and it helps decrease the dimensionality of the problem, while not losing much optimality (see Petering and Hussein 2013). As is common practice, we will not mention the term “restricted” in the rest of the paper even though we always assume  $A_3$ .

$A_4$ : The cost of relocating a container from a stack does not depend on the stack to which the container is relocated. This allows us to consider the stacks of a configuration as interchangeable. In addition, it motivates the objective of minimizing the number of relocations, since the cost of each relocation can be normalized to 1. Note that this assumption is not required for all of the results stated below, hence our approaches could be

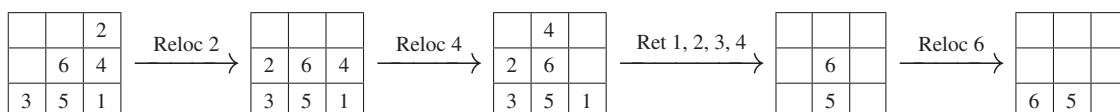
easily extended to the case when Assumption  $A_4$  does not hold.

$A_5$ : The retrieval order of containers is known, so that each container can be labeled from 1 to  $C$ , representing the departure order, that is, container 1 is the first one to be retrieved, and  $C$  is the last one.

The CRP involves finding a sequence of moves to retrieve containers  $1, 2, \dots, C$  (respecting the order) with a minimum number of relocations. Figure 2 provides a simple example of the CRP. The CRP with the above classical assumptions is referred to as *static and full information*: “static” because no new containers arrive during the retrieval process (see Assumption  $A_1$ ) and “full information” because we know the full retrieval order at the beginning of the retrieval process (see Assumption  $A_5$ ). This problem was first formulated by Kim and Hong (2006) in a dynamic programming model.

Researchers have tackled the static CRP with full information from two points of view. The first approach aims to find the optimal solution. Primarily, researchers have used integer programming (IP) to address this problem. For example, Caserta, Schwarze, and Voß (2012) propose an intuitive formulation of the problem. Petering and Hussein (2013) develop a more tractable formulation that is, however, unable to solve real-sized instances efficiently. Zehendner et al. (2015) fix the formulation from Caserta, Schwarze, and Voß (2012) and improve it by removing some variables, tightening some constraints, introducing a new upper bound, and applying a preprocessing step to fix several variables. In all these IP formulations, because of the combinatorial nature of the problem, the number of variables and constraints dramatically increases as the size of the problem grows, and the IP cannot be solved for large instances. To bypass this problem, a recent trend has been to look at more efficient ways to explore the branch-and-bound tree, or even decrease its size using the structural properties of the problem. Ünlüyurt and Aydın (2012) and Expósito-Izquierdo, Melián-Batista, and Moreno-Vega (2015) suggest two branch-and-bound approaches with several heuristics based on this idea. Another solution using the  $A^*$  algorithm is explored by Zhu et al. (2012), and built on by Tanaka and Takii (2016) and Borjian et al. (2015b). Another solution using branch-and-price is presented by Zehendner and Feillet (2014b).

As the problem is  $\mathcal{NP}$ -hard (Caserta, Schwarze, and Voß 2012), an alternative approach is to use quick and

**Figure 2.** Configuration for the CRP with Three Tiers, Three Stacks, and Six Containers

Note. The optimal solution performs three relocations.

efficient heuristics providing suboptimal solutions. For the sake of conciseness, we mention only some of them that are relevant to this paper. Caserta, Schwarze, and Voß (2012) introduce a “MinMax” policy that is defined and generalized in Section 3. Wu and Ting (2010) propose a beam search heuristic, and Wu and Ting (2012) develop the group assignment heuristic (GAH) also generalized later in Section 3. Finally, in Kim and Hong (2006) and Zhu et al. (2012), lower bounds for the CRP are introduced: Kim and Hong (2006) count the number of blocking containers as a straightforward lower bound, and Zhu et al. (2012) refine this idea by taking into account additional unavoidable relocations using a family of lower bounds.

There are many problems related to the CRP. The stacking problem is concerned with how to store incoming containers in a configuration given an arrival order of containers. The pre-marshalling problem deals with rearranging the containers prior to the retrieval process to minimize future relocations, but no container is removed in this process. For both problems, general review and classification surveys of the existing literature on the CRP can be found in Stahlbock and Voß (2008); Steenken, Voß, and Stahlbock (2004); and Lehnfeld and Knust (2014). In addition, Assumption  $A_1$  can be relaxed, which leads to the dynamic CRP where stacking and retrieving are done simultaneously as new containers arrive. For this problem, see Borjani et al. (2015a) and Hakan Akyüz and Lee (2014).

Finally, the main focus of this paper is an extension of the CRP where the full information assumption ( $A_5$ ) is relaxed. Indeed, Assumption  $A_5$  is unrealistic given that arrival times of external trucks at the terminal are generally unpredictable because of uncertain conditions. Nevertheless, new technology advancements such as truck appointment systems (TASs) and GPS tracking can help predict relative truck arrival times. Thus, although the exact retrieval order might not be known, some information on trucks’ arrival times might be available. This leads us to introduce a stochastic version of the CRP.

## 1.2. The Stochastic CRP (SCRCP)

A common assumption is that, for each container, there is a time window during which a truck driver will arrive to retrieve it. We refer to a *batch of containers* as the set of containers stacked in the same row and with the same arrival time window. This information can be either inferred using machine learning algorithms, rarely discussed in the literature, or obtained by using the appointment time windows in a TAS, which has gained attention over the last decade. The first TAS was implemented by Hongkong International Terminals (HIT) in 1997. It uses 30-minute time slots, where trucks can register (Murty et al. 2005). Another TAS was introduced in New Zealand in 2007. Two other

studies, Giuliano and O’Brien (2007) and Morais and Lord (2006), evaluate the benefits of TAS, in reducing truck idling time by increasing the on-time ratio. More recent information can be found in Phillips (2015) and Bonney (2015).

On the modeling side, Zehendner and Feillet (2014a) formulate an IP to get the optimal number of slots a TAS should offer for each batch. Very few studies have tackled the SCRCP, also referred to as a CRP with time windows. This problem was first modeled by Zhao and Goodchild (2010). In their original model, each container is assigned to a batch. Batches of containers are ordered such that all containers in a batch must be retrieved before any containers from a later batch are retrieved. Furthermore, the relative retrieval order of containers within a given batch is assumed to be a random permutation. Throughout the paper, we will refer to the model of Zhao and Goodchild (2010) as the *online model*. In Section 2, we discuss in more detail how this model assumes information is revealed. For the online model, Zhao and Goodchild (2010) develop a myopic heuristic (called RDH) and study, in different settings with two or more groups, the value of information using RDH. They conclude that a small improvement in the information system reduces the number of relocations significantly. Van Asperen, Borgman, and Dekker (2013) use a simulation tool to evaluate the effect of a TAS on many statistics including the ratio of relocations to retrievals. Their decision rules are based on several heuristics including the “leveling,” Random, or “traveling distance” heuristics. More recently, Ku and Arthanari (2016) also use the online model. They formulate the SCRCP under the online model as a finite horizon dynamic programming problem and suggest a decision tree scheme to solve it optimally. They also introduce a new heuristic called expected reshuffling index (ERI), which outperforms RDH, and they perform computational experiments based on available test instances. We will refer to this work frequently and use some of their techniques as well as their available test instances to evaluate our algorithms.

In another recent study related to the SCRCP, Zehendner, Feillet, and Jaillet (2017) study the *online* container relocation problem, which corresponds to an adversarial model. They prove that the number of relocations performed by the leveling policy can be upper-bounded by a linear function of the number of blocking containers and provide a tight competitive ratio for this policy. Moreover, Galle et al. (2016) show that the ratio of the expected number of relocations to the expected blocking lower bound converges to one. Finally, Tierney and Voß (2016) study the robust pre-marshalling problem which also considers uncertainty in the retrieval times of containers. For a general review of techniques on finite horizon dynamic programming, we refer the reader to Bertsekas (2005) and Sennott (2009).

### 1.3. Contributions of the Paper

The contributions of this paper are the following:

1. A new stochastic model, referred to as the *batch model*. This new model uses the same probability distribution as the online model. However, the two models differ in the way each reveals new information on the retrieval order. The batch model is motivated, described, and compared with the online model.

2. Lower and upper bounds for the SCRP. We derive a *new family of lower bounds* for which we show theoretical properties. Furthermore, we develop *two new fast and efficient heuristics*.

3. A *novel optimal algorithm scheme based on decision trees and pruning strategies referred to as pruning best first search (PBFS)*, which takes advantage of the properties of the aforementioned lower bounds. The algorithm is explained with pseudocode in Algorithm 2.

4. A *second novel algorithm tuned for the case of larger batches referred to as PBFS-approximate (PBFSA)*. We build on PBFS and derive a sampling strategy resulting in an approximate algorithm with an expected error that we bound theoretically. The pseudocode of the second algorithm is presented in Algorithm 3.

5. *Extensive computational experiments using an existing set of instances*. The first experiment exhibits the efficiency of PBFS, our lower bounds and two new heuristics for the batch model based on existing instances, presented by Ku and Arthanari (2016), where batches of containers are small (two containers per batch, on average). The second experiment illustrates the advantage of using PBFSA when batches of containers are larger, based on instances obtained by modifying the existing set. In addition, most of our techniques including lower bounds, heuristics, and the PBFS algorithm also apply to the online model. The third experiment shows that, in this model, PBFS outperforms the

algorithm introduced in Ku and Arthanari (2016) in the sense that it is faster for instances that Ku and Arthanari (2016) could solve, and it can solve problems of larger size. Furthermore, our two new heuristics also outperform the best existing heuristic (ERI) for the online model. Finally, the last experiment is used to test the conjecture about the optimality of the leveling heuristic in the special case of the online model with a unique batch of containers.

The rest of the paper is structured as follows: Section 2 thoroughly describes the batch model, its assumptions and objective, the difference with the online model, and the general theory of decision trees applied to the SCRP. Section 3 gives a good intuition into the problem and defines heuristics and a class of lower bounds for the SCRP used in subsequent sections. Then Sections 4 and 5 introduce, respectively, the PBFS and PBFSA algorithms. Computational experiments for both batch and online models are carried through Section 6. We conclude the paper in Section 7 with a discussion of future directions for the SCRP.

## 2. SCRP and Decision Trees

### 2.1. Motivation

Before stating the general assumptions of the batch model, let us motivate our problem using a typical example. We consider a port with a TAS offering 30-minute time windows during which truck drivers who want to retrieve a container can register to arrive at the port. For the sake of the example, we consider the time window between 9:00–9:30 A.M. Multiple trucks can be registered in this time window; in this example, presented in Figure 3, three trucks (designated  $i_1$ ,  $i_4$ , and  $i_6$ ) are registered for this time window.

**Figure 3.** (Color online) Timeline of Events for the Batch Model with Three Trucks

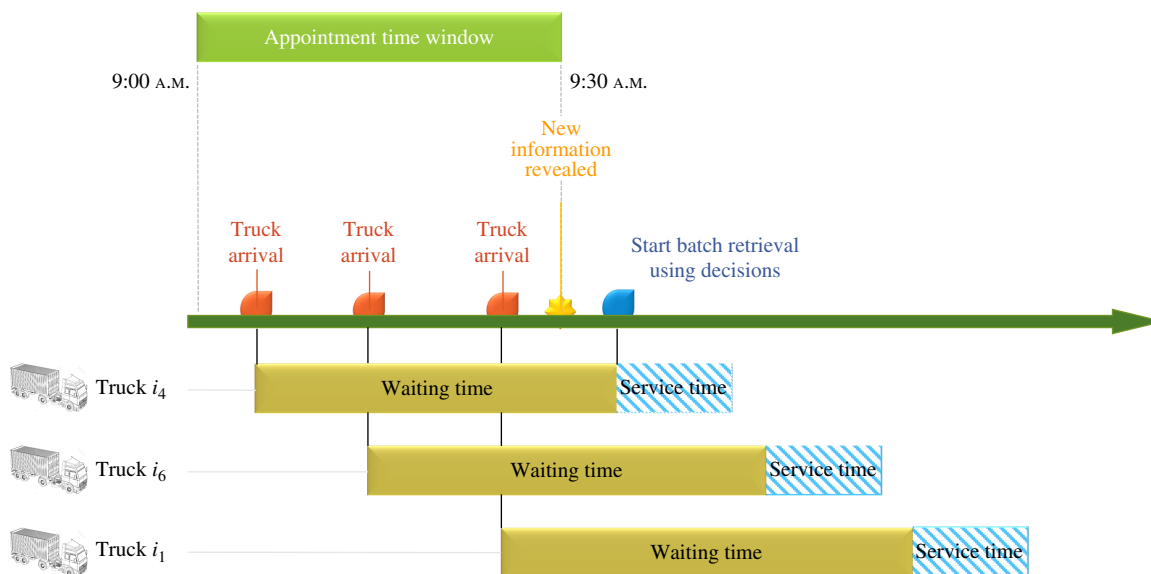
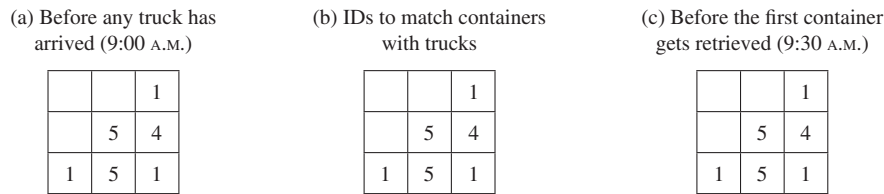


Figure 4. SCRP Example



Notes. The configuration in panel (a) is the input to our problem. The configuration in panel (b) denotes each container with an ID  $i_l$ , where  $l = 1, \dots, 6$ . The configuration in panel (c) denotes the order of the first batch after it is revealed.

We assume that all three trucks arrive on time (between 9:00–9:30 A.M.) and that their containers (similarly designated  $i_1$ ,  $i_4$ , and  $i_6$ ) form a batch to be retrieved. We display the configuration of interest in Figure 4 (3 tiers, 3 stacks, and 6 containers).

We assume that trucks arrive randomly within the time window, so each truck arrival order is equally likely to happen. In this example, there are six potential arrival orders, each with  $1/6$  chance of occurring.

At 9:00 A.M., none of the three trucks have arrived, and their relative retrieval order is unknown. Consequently, these three containers are all labeled 1 in Figure 4(a). In Figure 4(b), the IDs of all containers and their locations are depicted.

Between 9:00–9:30 A.M., trucks arrive in a particular order (e.g., Truck  $i_4$  first, then  $i_6$ , and  $i_1$  last). In busy terminals, trucks typically queue up as they wait to be served. Their place in line is based on their arrival order, so the port operator generally retrieves containers based on the arrival order. Processing in this way, on a first-come, first-served basis, avoids issues with truck unions and maintains fairness among drivers. Consequently, we take the retrieval order to be exogenously determined and we do not consider it a potential decision for port operators.

To provide a specified level of service to the truck drivers, the terminal operator often sets a target average waiting time. If the appointment time window is about the same as or shorter than the target average waiting time, the operator has information about the retrieval order of containers in the batch before the retrieval of those containers must begin to meet the target waiting time. Given these conditions, we make the simplifying assumption in this work that the retrieval of a batch begins at the end of the appointment time window associated with the batch, and the retrieval order of all containers in the batch is known at the moment the retrieval of the batch commences. In our example, the target average waiting time is 30 minutes. At 9:30 A.M., the retrieval order of the batch ( $i_4$ ,  $i_6$ ,  $i_1$ ) is known and the retrieval of the batch commences soon after. The updated information is depicted in the configuration of Figure 4(c).

The assumption that containers to be retrieved are revealed on a batch basis models the reality that port

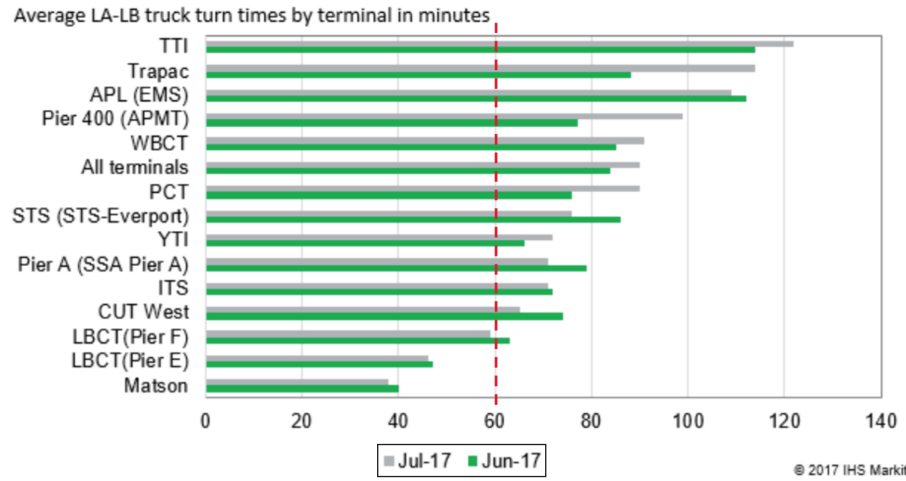
operators typically know information about all of the containers in the same batch before starting to retrieve them. This is especially true for busy ports that have a TAS. Moreover, we assume that no information about future batches is available when making decisions for the current batch. Similar modeling assumptions have been made in previous work (see Zhao and Goodchild 2010 and Ku and Arthanari 2016).

The general assumptions we apply to our model are formally stated in Section 2.2 ( $A_5^*$  and  $A_6^*$ ) and result in the *batch model*, the main focus of this paper. The goal of the SCRP is to find a sequence of moves minimizing the expected number of relocations needed to empty the initial configuration.

Labels in Figures 4(a) and 4(c) are defined such that two containers have the same label only if they are in the same batch and their relative order is yet to be revealed. In our example, since Container  $i_5$  is the only container in the second batch and is retrieved after the first three containers, it is necessarily the fourth container to be retrieved (thus labeled 4). Containers  $i_2$  and  $i_3$  are labeled 5, and when their relative order is revealed, one will be labeled 5 and the other 6.

We mention that these assumptions apply to an increasing number of port terminals. For example, Figure 5 presents the average truck turn time in minutes for all 15 terminals in the Los Angeles–Long Beach port (the largest U.S. port for containers) in June and July 2017. In many of these terminals, a truck appointment system with 60-minutes time windows (see Davies 2009) has been implemented, and Figure 5 shows that most terminals have an average truck turn time longer than 60 minutes.

**The online model.** The main distinguishing difference between the batch model and the online model introduced by Zhao and Goodchild (2010) is the information revelation process. The online model disregards any within-batch information available when it plans the moves to retrieve a container. Hence new information is revealed one container at a time. The online model especially applies to less busy ports where the waiting time is significantly shorter than the appointment time windows. In this case, because there is a limited number of trucks waiting, limited information about the future is known. We show through Lemma 1 that ignoring information (if available) results in a

**Figure 5.** (Color online) Average Truck Turn Times in Minutes by Terminal at Los Angeles–Long Beach Port in June and July 2017

Note. The dashed line shows the length of a time window (60 minutes) in the truck appointment system.  
Source. JOC.com; Harbor Trucking Association.

potential loss of operational efficiency as measured by the expected number of relocations. In addition, most of our batch-based approaches also apply to the online model, and we provide numerical results based on it in Section 6.

## 2.2. Assumptions, Notations, and Formulation

To define our problem as a multistage stochastic optimization problem (the number of “stages” is the number of batches), we need to define a probabilistic model of the container retrieval order, a scheme for revealing new information about this order, and an objective function.

**Batch model.** Let us state the assumptions and objective of the SCRP under the batch model. Assumptions  $A_1^*$ ,  $A_2^*$ ,  $A_3^*$ , and  $A_4^*$  are, respectively, identical to Assumptions  $A_1$ ,  $A_2$ ,  $A_3$ , and  $A_4$ .

$A_5^*$  (Probabilistic model). Given an ordering of batches, the probability distribution of the retrieval order is such that (1) all of the containers in a given batch are retrieved before any containers in a later batch, and (2) within each batch of containers, the order of the containers is drawn from a uniform random permutation. This paper focuses mainly on this latter assumption, but our model can be extended to the more general case of any probability distribution on permutations (not necessarily uniform) that respects the order of batches.

$A_6^*$  (Revelation of new information). For each batch  $w$ , the full relative order of containers from the  $w$ th batch (i.e., the specific random permutation drawn) is revealed after all containers in batch 1 through  $w - 1$  have been retrieved.

Given these assumptions, we want to find the minimum expected number of relocations to retrieve all containers from a given initial configuration. We refer

to this problem as the “stochastic CRP.” Let us introduce some notations:

—The problem size is given by  $(T, S, C)$ , respectively, the number of tiers, stacks, and containers in the initial configuration.

—The number of batches of containers in the initial configuration is denoted by  $W$ . We consider that the batches are ordered from 1 to  $W$ .

—For each batch  $w \in \{1, \dots, W\}$ , let  $C_w$  be the number of containers in  $w$ . By definition  $\sum_{w=1}^W C_w = C$ .

—Each container has two attributes:

- The first attribute, denoted by  $(c_l)_{l \in \{1, \dots, C\}}$ , is the label and is defined as follows: initially, containers in batch  $w$  are labeled  $K_w$  such that  $K_w = 1 + \sum_{u=1}^{w-1} C_u$ , where the sum is empty for  $w = 1$ . Then, for  $k \in \{1, \dots, C\}$ , if a container is revealed to be the  $k$ th container to be retrieved, its label changes to  $k$ . Using this labeling, at any point in the retrieval process, two containers have the same label only if they are in the same batch and their relative order is yet to be revealed.

- The second attribute is a unique ID denoted by  $(i_l)_{l \in \{1, \dots, C\}}$ . This ID is used only to uniquely identify containers in the initial configuration (see Figure 4(b)) and for the sake of clarity of the following probabilistic model. Note that for  $l \in \{1, \dots, C\}$ , if container  $i_l$  is in batch  $w$ , then  $c_l = K_w$  (until the actual retrieval order of  $i_l$  is revealed).

—For  $k \in \{1, \dots, C\}$ , let  $\zeta_k$  be a random variable taking values in  $(i_l)_{l \in \{1, \dots, C\}}$ , such that  $\{\zeta_k = i_l\}$  is the event that container  $i_l$  is the  $k$ th container to be retrieved. According to Assumption  $A_5^*$ , the distribution of  $(\zeta_k)_{k \in \{1, \dots, C\}}$  is given by

$$\mathbb{P}[\zeta_k = i_l] = \begin{cases} \frac{1}{C_w}, & \text{if } w = \min\{u \in \{1, \dots, W\} \mid K_u \geq k\} \\ & \text{and } c_l = K_w, \\ 0, & \text{otherwise.} \end{cases}$$

In this case, there are a total of  $\prod_{w=1}^W (C_w!)$  orders with equal probabilities. More generally, we consider the case where the probability of each order within each batch is not necessarily equally likely. However, we still assume that the batches are ordered, thus  $\mathbb{P}[\zeta_k = i_l]$  can be positive only if  $w = \min\{u \mid K_u \geq k\}$  and  $c_l = K_w$ . In the practical case where probabilities are not considered to be uniform, a list of potential retrieval orders and their associated probability is given for each batch of containers (based on historical data), thus  $\mathbb{P}[\zeta_k = i_l]$  can easily be inferred from these probabilities.

—An action corresponds to a sequence of relocations to retrieve one container. For  $k \in \{1, \dots, C\}$ , we denote the action for the  $k$ th retrieval by  $a_k$ , and the feasible set of actions is defined according to Assumptions  $A_2^*$  and  $A_3^*$ .

—For a given batch  $w \in \{1, \dots, W\}$ ,

1. Let  $y_w$  denote the configuration before the retrieval order of containers in batch  $w$  is revealed, that is, before  $\zeta_{K_w}, \dots, \zeta_{K_w+C_w-1}$  are revealed. Note that  $y_1$  corresponds to the initial configuration. We denote  $x_{K_w}$  the configuration after the retrieval order of containers in batch  $w$  is revealed, and before action  $a_{K_w}$  is taken. If  $\xrightarrow{\zeta_{K_w}, \dots, \zeta_{K_w+C_w-1}}$  represents the revelation of the random variables  $\zeta_{K_w}, \dots, \zeta_{K_w+C_w-1}$ , we can write  $y_w \xrightarrow{\zeta_{K_w}, \dots, \zeta_{K_w+C_w-1}} x_{K_w}$ .

2. After the retrieval order of batch  $w$  has been revealed, actions to retrieve the revealed containers must be made. If  $C_w > 1$ , then  $\{K_w, \dots, K_w + C_w - 2\} \neq \emptyset$ . In this case, for  $k \in \{K_w, \dots, K_w + C_w - 2\}$ , let  $x_{k+1}$  be the configuration after applying action  $a_k$  to state  $x_k$  and before action  $a_{k+1}$ . Therefore, if  $\xrightarrow{a_k}$  represents the application of action  $a_k$ , we have  $x_k \xrightarrow{a_k} x_{k+1}$ .

3. The last container to be retrieved in batch  $w$  is the  $(K_w + C_w - 1)$ th container, thus, according to the previous point,  $x_{K_w+C_w-1}$  corresponds to the configuration before  $a_{K_w+C_w-1}$  is taken. After this retrieval, the order of the next batch (batch  $w + 1$ ) has to be revealed, and according to the first point, the configuration is  $y_{w+1}$ . The configuration after retrieving batch  $W$  will be empty, thus we define  $y_{W+1}$  to be the empty configuration. So if  $\xrightarrow{a_{K_w+C_w-1}}$  represents the application of action  $a_{K_w+C_w-1}$ , then we have  $x_{K_w+C_w-1} \xrightarrow{a_{K_w+C_w-1}} y_{w+1}$ .

In summary, we have

$$\forall w \in \{1, \dots, W\}, \begin{cases} y_w \xrightarrow{\zeta_{K_w}, \dots, \zeta_{K_w+C_w-1}} x_{K_w}, \\ x_k \xrightarrow{a_k} x_{k+1}, & \text{if } C_w > 1, \forall k \in \{K_w, \dots, K_w + C_w - 2\}, \\ x_{K_w+C_w-1} \xrightarrow{a_{K_w+C_w-1}} y_{w+1}. \end{cases}$$

—Let the function  $r(\cdot)$  be such that  $r(x)$  is the number of relocations to retrieve the target container in configuration  $x$ . It is also equal to the number of containers blocking the target container. This function is defined only for configurations in which the target container is

identified. Specifically, it is defined for  $(x_k)_{k=1, \dots, C}$  (but not for  $(y_w)_{w=1, \dots, W}$ ). For  $k \in \{1, \dots, C\}$ , we refer to  $r(x_k)$  as the *immediate cost* for the  $k$ th retrieval.

—Let the function  $f(\cdot)$  be such that  $f(x)$  is the minimum expected number of relocations required to retrieve all containers from configuration  $x$ . In addition,  $f(x)$  is commonly referred to as the *cost-to-go* function of configuration  $x$ . Note that it is well defined for both  $(x_k)_{k=1, \dots, C}$  and  $(y_w)_{w=1, \dots, W}$ .

By definition, we have

$$\forall w \in \{1, \dots, W\}, \begin{cases} f(y_w) = \mathbb{E}_{\zeta_{K_w}, \dots, \zeta_{K_w+C_w-1}} [f(x_{K_w})], & \text{where } k = K_w, \\ f(x_k) = r(x_k) + \min_{a_k} \{f(x_{k+1})\}, & \text{if } C_w > 1 \text{ and} \\ & \forall k \in \{K_w, \dots, K_w + C_w - 2\}, \\ f(x_k) = r(x_k) + \min_{a_k} \{f(y_{w+1})\}, & \text{where } k = K_w + C_w - 1, \end{cases}$$

which can be written as

$$\forall w \in \{1, \dots, W\}, \begin{cases} f(y_w) = \mathbb{E}_{\zeta_{K_w}, \dots, \zeta_{K_w+C_w-1}} [f(x_{K_w})], \\ f(x_{K_w}) = \min_{a_{K_w}, \dots, a_{K_w+C_w-1}} \left\{ \left( \sum_{k=K_w}^{K_w+C_w-1} r(x_k) \right) + f(y_{w+1}) \right\}, \end{cases} \quad (1)$$

and  $f(y_{W+1}) = 0$ . Therefore, the SCRPP is concerned with finding  $f(y_1)$ , where by induction

$$f(y_1) = \mathbb{E}_{\zeta_{K_1}, \dots, \zeta_{K_1+C_1-1}} \left[ \min_{a_{K_1}, \dots, a_{K_1+C_1-1}} \left\{ \mathbb{E}_{\zeta_{K_2}, \dots, \zeta_{K_2+C_2-1}} \left[ \dots \right. \right. \right. \\ \left. \left. \left. \dots \mathbb{E}_{\zeta_{K_W}, \dots, \zeta_{K_W+C_W-1}} \left[ \min_{a_{K_W}, \dots, a_{K_W+C_W-1}} \left\{ \sum_{k=1}^C r(x_k) \right\} \right] \dots \right] \right\} \right] \quad (2)$$

As a final remark, we note that batches should be as small as possible if only information is at stake. Indeed, smaller batches correspond to an efficient information system since more information is known about the retrieval order. Yet the size of the batches is restricted by two intrinsic constraints:

1. *Batches should be larger than a certain size.* Indeed, a terminal offers time slots for trucks to register, and these slots cannot be brief, as trucks would most certainly not arrive during their appointed slot because of traffic or other uncertainties. Therefore, given the minimum time of a slot, the terminal will allow at least a certain number of trucks to register for each slot, that is, the minimum batch size.

2. *Batches cannot be too large* for the batch model to be applicable, since in this model, the appointment time windows are supposed to be the same as or shorter than the target waiting time. As the target waiting time is limited, only a limited number of containers can be retrieved in a certain batch.

**The online model.** Using our notations, we briefly present the SCRP under the online model. Instead of Assumption A<sub>6</sub><sup>\*</sup>, the online model assumes that for each retrieval  $k \in \{1, \dots, C\}$ , only the next target container is revealed (i.e.,  $\zeta_k$ ). Therefore, we consider the states  $(y_k^o, x_k^o)_{k=1, \dots, C}$  defined such that  $k \in \{1, \dots, C\}$ ,  $y_k^o \xrightarrow{\zeta_k} x_k^o \xrightarrow{a_k} y_{k+1}^o$ , where  $y_{C+1}^o$  is the empty configuration. In this case, if  $f^o$  denotes the cost-to-go function, then by definition we have  $f^o(y_k^o) = \mathbb{E}_{\zeta_k} [f^o(x_k^o)]$  (with  $f^o(y_{C+1}^o) = 0$ ), and  $\forall k \in \{1, \dots, C\}$ ,  $f^o(x_k^o) = \min_{a_k} \{r(x_k^o) + f^o(y_{k+1}^o)\}$ . By induction, the SCRP under the online model is hence concerned with finding

$$f^o(y_1^o) = \mathbb{E}_{\zeta_1} \left[ \min_{a_1} \left\{ \mathbb{E}_{\zeta_2} \left[ \dots \mathbb{E}_{\zeta_C} \left[ \min_{a_C} \left\{ \sum_{k=1}^C r(x_k^o) \right\} \right] \dots \right] \right\} \right].$$

The next lemma compares the batch and the online models theoretically (the proof can be found in Online Appendix A). It states that it is beneficial in terms of the expected number of relocations to use the batch model rather than the online model, if the first one is applicable. Practically, this suggests that the operator should always use available information.

**Lemma 1.** Let  $y$  be a given initial configuration, then we have

$$f(y) \leq f^o(y).$$

### 2.3. Decision Trees

Multistage stochastic optimization problems can be solved using decision trees in which *chance nodes* and *decision nodes* typically alternate. A *chance node* embodies the stochasticity of the model, whereas a *decision node* models the possible actions of the algorithm. In a decision tree for the SCRP, a node represents a configuration. The root node (denoted by 0) is the initial configuration, and the leaf nodes are configurations for which we can compute the cost-to-go function.

In our case, we slightly modify the structure of a typical decision tree, in the sense that chance nodes and decision nodes do not necessarily alternate. A chance node is a configuration for which the target container is not yet known, and information needs to be revealed (note that this occurs only at the beginning of each batch). A decision node is a configuration for which the target container is known. Using our notations, chance nodes correspond to  $(y_w)_{w=1, \dots, W}$  and decision nodes correspond to  $(x_k)_{k=1, \dots, C}$ .

Let  $n$  be a node corresponding to a configuration in the decision tree. Thus  $f(n)$ , the cost-to-go function of  $n$ , is defined for all nodes  $n$ , and  $r(n)$ , the immediate cost function of  $n$ , is well defined when  $n$  is a decision node. We denote by  $\lambda_n$  the level of  $n$  in the decision tree, and define it as the number of containers remaining in the configuration. We denote the lowest level of the tree by  $\lambda^* = \min_{n \in \text{Tree}} \{\lambda_n\}$ . It corresponds to the level such that if  $\lambda_n = \lambda^*$ ,  $f(n)$  can be computed

efficiently (the empty configuration being an obvious candidate with a cost-to-go of 0). Moreover, we have the following:

- If  $n$  is a chance node, then there exists a unique  $w \in \{1, \dots, W\}$  such that  $\lambda_n = C - K_w + 1$ . We denote by  $\Omega_n$  the set of offspring of  $n$ , each offspring being a decision node corresponding to a realization of the random variables  $\zeta_{K_w}, \dots, \zeta_{K_w+C_w-1}$ , that is, the full retrieval order of containers in batch  $w$ . Note that  $n$  has a priori  $|\Omega_n| = C_w!$  offspring.

- If  $n$  is a decision node, then  $r(n)$  is well defined and is equal to the number of containers blocking the target container in configuration  $n$  (i.e., the  $(C - \lambda_n + 1)$ th container to be retrieved). We denote by  $\Delta_n$  the set of offspring of  $n$ , which can either be chance nodes if there exists  $w \in \{1, \dots, W\}$  such that  $\lambda_n + 1 = C - K_w + 1$ , or decision nodes otherwise. For brevity, we compute  $\Delta_n$  greedily by considering all feasible combinations of relocations of the  $r(n)$  containers blocking the target container in  $n$ . Note that  $|\Delta_n|$  is of the order of  $(S - 1)^{r(n)}$ , where  $S$  is the number of stacks.

Equation (1) provides the relation to compute the cost-to-go by backtracking. For all  $n$  in the decision tree, we have

$$f(n) = \begin{cases} \frac{1}{|\Omega_n|} \sum_{n_i \in \Omega_n} f(n_i), & \text{if } n \text{ is a chance node,} \\ r(n) + \min_{n_i \in \Delta_n} \{f(n_i)\}, & \text{if } n \text{ is a decision node.} \end{cases} \quad (3)$$

In the case in which the probability of each permutation (in each batch) is not uniform, we mentioned that in practice, operators provide the probability of potential orders for each batch. Given a chance node  $n$  and one of its offspring  $n_i \in \Omega_n$ , this input probability is exactly the probability that the actual retrieval order is the one revealed in  $n_i$ . For a given node  $n$ , we denote these probabilities by  $(p_{n_i})_{n_i \in \Omega_n}$ . In this case, Equation (3) is replaced by

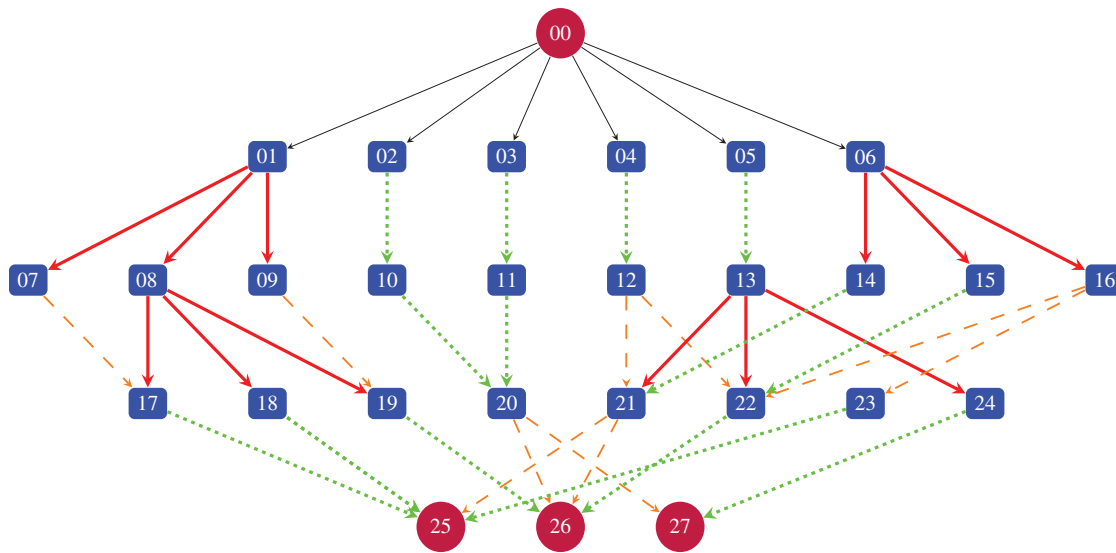
$$f(n) = \begin{cases} \sum_{n_i \in \Omega_n} p_{n_i} f(n_i), & \text{if } n \text{ is a chance node,} \\ r(n) + \min_{n_i \in \Delta_n} \{f(n_i)\}, & \text{if } n \text{ is a decision node,} \end{cases}$$

for all  $n$  in the decision tree.

Figures 6 and 7 provide the description of the decision tree corresponding to the example in Figure 4, using chance/decision nodes and configurations, respectively. A chance node is depicted by a circle, and a decision node is depicted by a square.

To illustrate how to use Equation (3), we derive the calculations using the example in Figure 6. Suppose that  $f(25) = f(26) = f(27) = 0.5$  are known, then we get  $f(17) = f(18) = f(19) = 0.5$ ,  $f(07) = f(09) = 1.5$ , and  $f(08) = 2.5$ , leading to  $f(01) = 3.5$ . Similarly, by backtracking, we can compute  $f(02) = f(03) = f(04) = 1.5$  and  $f(05) = f(06) = 2.5$ , giving us  $f(00) = 13/6$ .

Figure 6. (Color online) Decision Tree Represented with Nodes



Note. The colored arrows represent different values of immediate cost, that is, the number of containers blocking the target container (dotted green: 0, dashed orange: 1, thick solid red: 2).

As the example shows, considering the full decision tree can become intractable even for small examples, hence making this approach impractical for larger problems. As previously mentioned, the number of decision offspring of a chance node scales with  $C_w!$ , and the number of offspring of a decision node is of the order of  $(S - 1)^{r(n)}$ . So the size of the tree grows exponentially with the size of the problem. However, there exist general and specific techniques to reduce substantially the size of this tree, as we discuss below.

First, there are suboptimal approaches. One way to approximate  $f(n)$ , when  $n$  is a chance node, is to sample from its offspring. When  $\Omega_n$  is large (which can happen with large batches), one might sample a certain number of realizations, resulting in a set of offspring  $\Psi_n \subset \Omega_n$ . By sampling “enough,” we show in Section 5, we can ensure guarantees on expectation for such an algorithm. Another popular suboptimal approach is to use techniques from approximate dynamic programming. These techniques can provide good empirical results, but no guarantee on how far from optimality can be obtained. This direction is not discussed in this paper but can be a direction for future work. Finally, another approach is to consider heuristics such as the ones described in Section 3, which select a subset of the offspring of decision nodes, and lead to upper bounds on the optimal value  $f(0)$ .

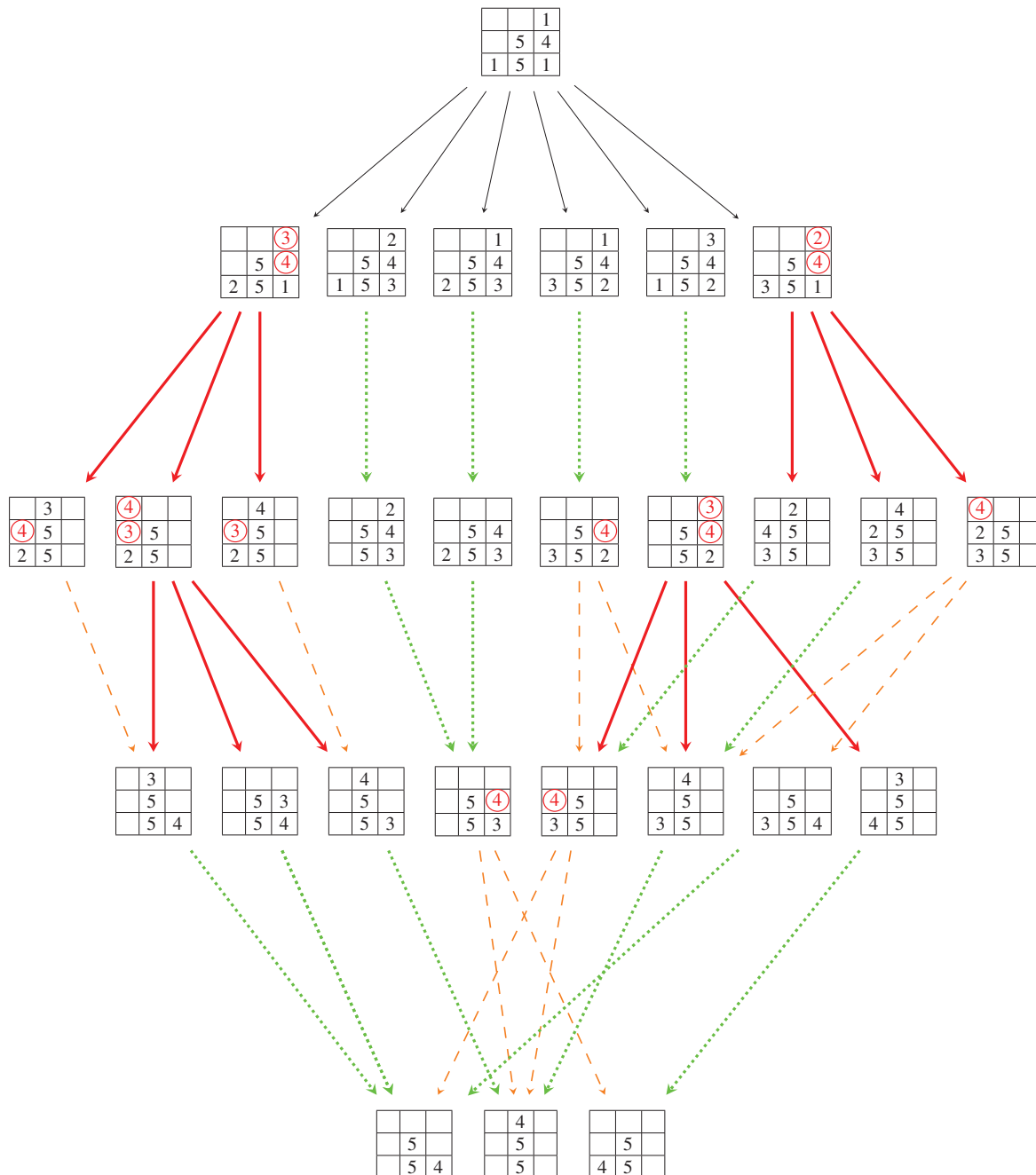
Second, there exist ways to decrease the size of the tree while ensuring optimality. One is to reduce the number of nodes using the problem structure of the SCRP. In the online setting, Ku and Arthanari (2016) propose an “abstraction” technique that significantly shrinks the size of the tree. Thanks to Assumption  $A_4^*$ , we can consider that the stacks of a configuration are interchangeable, making many configurations equivalent in

terms of number of relocations. For instance, in Figure 6, nodes 20 and 21 are identical in terms of number of relocations.

We describe the abstraction step with an example in Figure 8. The five configurations at the top are all equivalent to the configuration at the bottom. The abstraction transformation first ranks the stacks by increasing height. For stacks with equal height, it breaks ties by ranking them lexicographically going from the bottom to the top. Stacks are rearranged to have the first ranked on the left and the last on the right. Ku and Arthanari (2016) use a slightly different rule and add the extra step of removing empty stacks. Using the abstraction procedure, the proposed algorithm should not generate a node twice with identical abstracted versions. Even though Ku and Arthanari (2016) introduce this rule for the online model, this abstraction step also applies in the batch setting. Throughout the remainder of the paper, we will refer in pseudocode to the function  $\text{ABSTRACT}(n)$ , when applying this method to a given node  $n$ . We mention that Ku and Arthanari (2016) also suggest caching strategies that could be added on the top of this simplification step, including caching part of the tree or using a transportation table.

Finally, the performance of a decision-tree-based algorithm depends on the exploration strategy of the tree. For the online model, Ku and Arthanari (2016) use depth-first-search (DFS), and we propose to explore the best-first-search (BFS) approach. Note that BFS requires some kind of measure that we define in Section 4.

In Sections 4 and 5, we explore two other ways to decrease the size of the tree while still ensuring optimality. The first one is specific to the SCRP, and uses

**Figure 7.** (Color online) Decision Tree Represented with Configurations

*Notes.* The colored arrows represent different values of immediate cost, that is, the number of containers blocking the target container (dotted green: 0, dashed orange: 1, thick solid red: 2). Red circled numbers highlight containers blocking the target container.

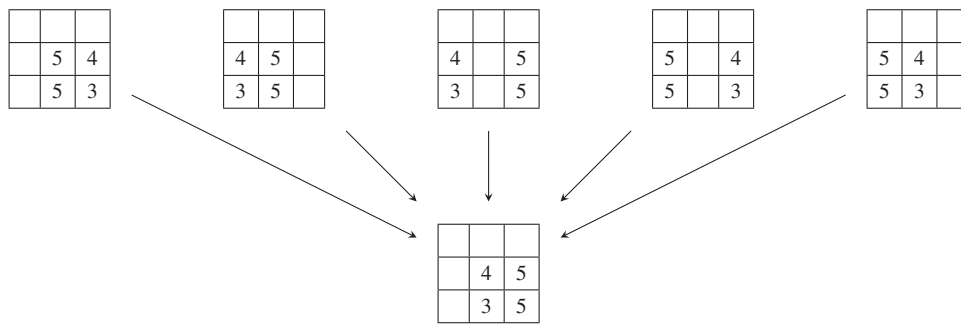
properties of the problem to increase  $\lambda^*$ . Recall that  $\lambda^*$  is the minimum level of the tree at which we can find the optimal expected number of relocations, without further branching. We show that we can set  $\lambda^*$  to  $\max\{S, C_W\}$ , where  $S$  is the number of stacks, and  $C_W$  is the number of containers in the last batch. Comparatively, Ku and Arthanari (2016) branch until  $\lambda^* = 0$ . The second optimal pruning strategy uses lower bounds in a BFS scheme.

After we introduce the batch model for the SCRP (as well as the online model) and some preliminary concepts about decision trees, Sections 3–5 develop approaches to solve the SCRP.

### 3. Heuristics and Lower Bounds

Before introducing the two main algorithms, we describe in this section heuristics and lower bounds for the SCRP. Indeed, PBFS and PBFSA build on some of

Figure 8. “Abstraction” Procedure



these bounds. In addition, these algorithms provide good intuition on how to solve the SCRP.

Let  $n$  be a given configuration with  $S$  stacks and  $T$  tiers. We say that a container  $c$  is a *blocking container* in  $n$  if it is stacked above at least one container that is to be removed before  $c$ . Note that all of the bounds mentioned below apply both in the batch and online models.

### 3.1. Heuristics

For the sake of completeness, we first describe three existing heuristics used in the proofs and/or our computational experiments before describing two novel heuristics.

#### 3.1.1. Existing Heuristics.

**Random.** For every relocation of a blocking container  $c$  from stack  $s$ , the Random heuristic picks any stack  $s' \neq s$  uniformly at random among stacks that are not “full,” that is, stacks containing strictly less than  $T$  containers.

**Leveling (L).** For every relocation of a blocking container  $c$  from stack  $s$ , L chooses the stack  $s' \neq s$  currently containing the fewest containers, breaking ties arbitrarily by selecting the leftmost stack.

Heuristic L is interesting for several reasons. Most important, it is an intuitive and commonly used heuristic in real operations in that it uses no more than the height of each stack. In addition, it does not require any information about batches or departure times, which means it is robust with respect to the inaccuracy of information.

**Expected reshuffling index (ERI).** This index-based heuristic was introduced by Ku and Arthanari (2016) for the online model. For every relocation of a blocking container  $c$  from stack  $s$ , ERI computes a score called the expected reshuffling index for each stack  $s' \neq s$  that is not full, denoted by  $\text{ERI}(s', c)$ . ERI chooses the stack  $s' \neq s$  with the lowest  $\text{ERI}(s', c)$ . In the case of a tie, the policy breaks it by selecting the highest stack among the ones minimizing  $\text{ERI}(s', c)$ . Further ties are arbitrarily broken by selecting the leftmost stack verifying

the two previous conditions. The value  $\text{ERI}(s', c)$  corresponds to the expected number of containers in stack  $s'$  that depart earlier than  $c$ . Let  $H_{s'}$  be the current number of containers in  $s'$ . If  $H_{s'} = 0$ , then  $\text{ERI}(s', c) = 0$ . Otherwise, let  $(c_1, \dots, c_{H_{s'}})$  be the containers in  $s'$ , then  $\text{ERI}(s', c) = \sum_{i=1}^{H_{s'}} \mathbb{1}\{c_i < c\} + \mathbb{1}\{c_i = c\}/2$ .

#### 3.1.2. First New Heuristic: Expected Minmax (EM).

EM considers an idea similar to that of Caserta, Schwarze, and Voß (2012) for the CRP. Let  $\min(s)$  be the smallest label of a container in  $s$  ( $\min(s) = C + 1$ , if stack  $s$  is empty). For every relocation of a blocking container  $c$  from stack  $s$ , we select the stack to which we relocate  $c$  using the following rules:

**Rule 1.** If there exists  $s' \neq s$  such that  $\min(s') > c$ , let

$$M = \min_{s' \in \{1, \dots, S\} \setminus s} \{\min(s') : \min(s') > c\}.$$

Select a stack such that  $\min(s') = M$ , breaking ties by choosing from the highest ones, finally taking the leftmost stack if any ties remain.

**Rule 2.** If for all stacks  $s' \neq s$ ,  $\min(s') \leq c$ , let

$$M = \max_{s' \in \{1, \dots, S\} \setminus s} \{\min(s')\}.$$

Select a stack such that  $\min(s') = M$ . If there are several such stacks, select those with the minimum number of containers labeled  $M$ . Further ties are broken by taking the highest ones, and finally choosing the leftmost one arbitrarily.

Rule 1 says, if there is a stack where  $\min(s)$  is greater than  $c$  ( $c$  can almost surely avoid being relocated again), then choose such a stack where  $\min(s)$  is minimized, since stacks with larger minimums can be useful for larger blocking containers.

If there is no stack satisfying  $\min(s) > c$  (Rule 2), then we have two cases following the same rule. On one hand, if  $M = c$ , then  $M$  is the maximum of the minimum labels of each stack, and  $c$  can potentially avoid being relocated again. If there are several stacks that maximize the minimum label, then by selecting the one with the least number of containers labeled  $M$ , EM minimizes the probability of  $c$  being relocated again.

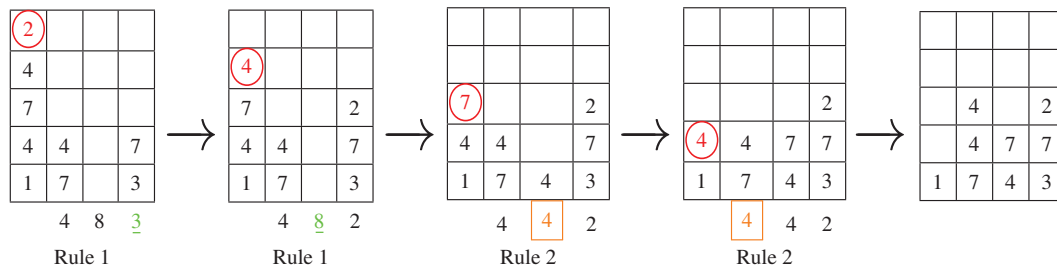
On the other hand, if  $M < c$ ,  $c$  will almost surely be relocated again, and then EM chooses the stack where  $\min(s)$  is maximized to delay the next unavoidable relocation of  $c$  as much as possible. We show how EM makes a decision on a simple example in Figure 9.

**3.1.3. Second New Heuristic: Expected Group Assignment (EG).** EM is quite intuitive because it tries to minimize the number of blocking containers after each retrieval. EG aims for the same goal, but uses some more sophisticated rules. EG is inspired by a heuristic designed by Wu and Ting (2012) for the complete information case, and we generalize this idea to the SCRP. It is different from ERI and EM because it considers

a group of blocking containers together, whereas ERI and EM consider them one at a time. EG can be decomposed into two main phases for each retrieval. The decisions made by EG on the same example as shown in Figure 9 are given in Figure 10.

The first phase assigns the blocking containers for which there exists  $s' \neq s$  such that  $\min(s') > c$ . If this is not the case, the assignments of these containers will be ignored during the first phase. The acceptable containers are assigned in descending order of labels, that is, the container with the highest label is assigned first (breaking ties for the highest one first). To assign these acceptable containers, the first phase applies the first

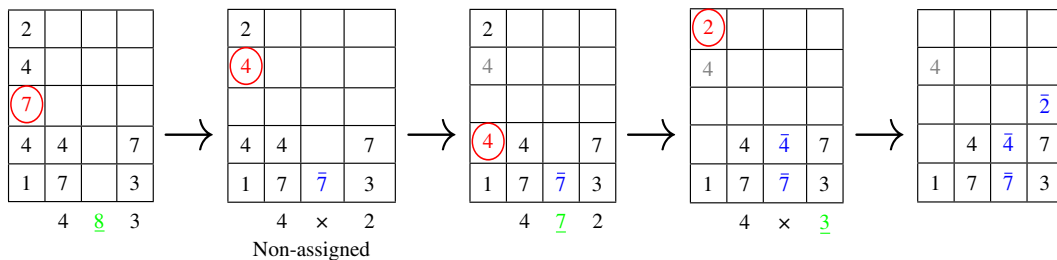
**Figure 9.** (Color online) Decisions of the EM Heuristic on an Example with Five Tiers, Four Stacks, and Nine Containers (Three per Batch)



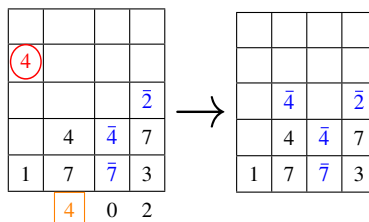
*Notes.* Under the batch model, the first batch has been revealed, and we present the decisions to retrieve the first container made by EM. The container with the circled red label is the current blocking container. Numbers under the configuration correspond to the stack indices  $\min(s)$ . The underlined green (respectively, squared orange) indices correspond to the selected stack with the corresponding  $M$  when Rule 1 (respectively, Rule 2) applies.

**Figure 10.** (Color online) Decisions of the EG Heuristic in an Example with Five Tiers, Four Stacks, and Nine Containers (Three in Each Batch)

(a) First phase: EG assigns acceptable containers in descending order. The container with the circled red label is the next acceptable container that EG tries to assign to a stack. Containers with overlined blue labels are assigned, and with gray labels are unassigned. Below, we show the indices  $\min(s)$  to apply the first rule of EM ( $\times$  means that a container below the considered container has already been assigned to a stack)



(b) Second phase, EG assigns all nonassigned containers using the index  $G \min(s)$



*Note.* Under the batch model, the first batch has been revealed, and we present the two-phase decisions to retrieve the first container made by EG.

of the EM rules. Finally, an acceptable container cannot be assigned to a stack if there is a container below it that was previously assigned to this stack.

The assignment in the second phase for the blocking containers not yet assigned might lead to additional relocations. These containers are assigned to other stacks in an ascending order of labels. The second phase computes a modified  $\min(s')$  index for each stack denoted by  $G\min(s')$ , which is defined as follows: Let  $H_{s'}$  be the height of stack  $s'$  and  $B(s')$  be the subset of containers assigned in the first phase to stack  $s'$

$$G\min(s') = \begin{cases} -1, & \text{if } |B(s')| + H_{s'} = T, \\ \min(s'), & \text{if } |B(s')| = 0, \\ B(s'), & \text{if } |B(s')| = 1, \\ 0, & \text{otherwise.} \end{cases}$$

If a stack is full after we assign the containers during the first phase, then it cannot be selected. If no container was assigned, the index remains as the min. If one container was assigned, it is “artificially” the new minimum of the stack. Finally, if more than one container was assigned, the index becomes very unattractive by being as low as possible (0). The second phase is similar to the EM heuristic, but it considers  $G\min$  instead of the  $\min$  index, breaking ties identically. Note that, after each assignment in the second phase, we update  $G\min$  accordingly for the remaining containers to be assigned. For more details in the complete information case, we refer the reader to Wu and Ting (2012).

### 3.2. Lower Bounds

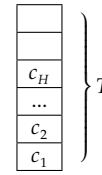
After defining heuristics (upper bounds), we are now concerned with defining valid lower bounds for the SCRP. More specifically, we care about computing lower bounds for decision nodes in the decision tree previously defined. Note that the computation of lower bounds easily extends to chance nodes.

**3.2.1. Blocking Lower Bound.** Suppose that the departure order is known, as in the CRP. The following lower bound was introduced by Kim and Hong (2006) and is based on the following simple observation. If a container is blocking in  $n$ , then it must be relocated at least once. Thus, the optimal number of relocations is lower bounded by the number of blocking containers.

In the SCRP, the retrieval order is a random variable, so the fact that a container is blocking is also random. Let us denote the expected number of blocking containers in  $n$  by  $b(n)$ . Therefore, by taking the expectation on the retrieval order of the previous fact, which holds for every retrieval order, we have the following observation.

**Observation 1.** For all configurations  $n$ ,  $f(n)$  is the minimum expected number of relocations to empty  $n$ ,

**Figure 11.** Example of a Single Stack Configuration



and  $b(n)$  is the expected number of blocking containers, then

$$f(n) \geq b(n).$$

Lemma 2 shows one way to compute the expected number of blocking containers for one stack, and  $b(n)$  is the sum of the expected number of blocking containers of each stack of  $n$ . Mathematically, let  $b_s(n)$  be the expected number of blocking containers in stack  $s$  of  $n$ ; we have  $b(n) = \sum_{s=1}^S b_s(n)$ .

**Lemma 2.** Let  $n$  be a single stack configuration with  $T$  tiers, and  $H \geq 0$  containers ( $H \leq T$ ). If  $H = 0$ , we have

$$b(n) = 0.$$

If  $H \geq 1$ , we denote the label of containers by  $(c_i)_{i=1,\dots,H}$ , where  $c_1$  is the container at the bottom and  $c_H$  is the container at the top (see Figure 11), then we have

$$b(n) = H - \sum_{h=1}^H \frac{\mathbb{I}\{c_h = \min_{i=1,\dots,h}\{c_i\}\}}{\sum_{i=1}^h \mathbb{I}\{c_h = c_i\}},$$

where  $\mathbb{I}\{A\}$  is the indicator function of  $A$ .

**Proof.** Clearly, if  $H = 0$ , then  $b(n) = 0$ . If  $H \geq 1$ , then, by definition, we have

$$\begin{aligned} b(n) &= \mathbb{E} \left[ \sum_{h=1}^H \mathbb{I}\{c_h \text{ is a blocking container}\} \right] \\ &= \sum_{h=1}^H \mathbb{P}[c_h \text{ is a blocking container}]. \end{aligned}$$

Let us fix  $h \in \{1, \dots, H\}$ , and compute the probability that  $c_h$  is blocking. We consider two cases:

- If  $c_h > \min_{i=1,\dots,h}\{c_i\}$ , then  $c_h$  is almost surely blocking.
- Otherwise it is  $c_h = \min_{i=1,\dots,h}\{c_i\}$ , and there are  $\sum_{i=1}^h \mathbb{I}\{c_h = c_i\} - 1$  containers below  $c_h$  with the same label (or batch). Since each departure sequence between containers of the same batch is equally likely, the probability that  $c_h$  is blocking is equal to  $\frac{\sum_{i=1}^h \mathbb{I}\{c_h = c_i\}}{\sum_{i=1}^h \mathbb{I}\{c_h = c_i\}} - 1 / \sum_{i=1}^h \mathbb{I}\{c_h = c_i\} = 1 - 1 / \sum_{i=1}^h \mathbb{I}\{c_h = c_i\}$ .

Consequently, we get

$$\begin{aligned} \mathbb{P}[c_h \text{ is a blocking container}] &= 1 \times \mathbb{I}\left\{c_h > \min_{i=1,\dots,h}\{c_i\}\right\} + \left(1 - \frac{1}{\sum_{i=1}^h \mathbb{I}\{c_h = c_i\}}\right) \\ &\quad \times \mathbb{I}\left\{c_h = \min_{i=1,\dots,h}\{c_i\}\right\} \\ &= 1 - \frac{\mathbb{I}\{c_h = \min_{i=1,\dots,h}\{c_i\}\}}{\sum_{i=1}^h \mathbb{I}\{c_h = c_i\}}. \end{aligned}$$

We sum the above expression for  $h = 1, \dots, H$  to conclude the proof.

Therefore, one can compute the blocking lower bound as follows: let  $H^s$  be the number of containers in stack  $s$ , and  $(c_1^s, \dots, c_{H^s}^s)$  be the containers in stack  $s$  listed from the bottom to the top, then

$$b(n) = \sum_{\substack{s=1, \dots, S \\ H^s \geq 1}} \left( H^s - \sum_{h=1}^{H^s} \frac{\mathbb{I}\{c_h^s = \min_{i=1, \dots, h} \{c_i^s\}\}}{\sum_{i=1}^h \mathbb{I}\{c_h^s = c_i^s\}} \right). \quad (4)$$

**Nonuniform case.** In the case where probabilities are not uniform across retrieval orders, we still consider a similar lower bound. For each container  $c_h^s$ , let  $q_{c_h^s}$  be the probability that  $c_h^s$  is the first container to be retrieved among the ones with the same batch, and positioned below in its stack. Equation (4) extends to give

$$b(n) = \sum_{\substack{s=1, \dots, S \\ H^s \geq 1}} \left( H^s - \sum_{h=1}^{H^s} q_{c_h^s} \mathbb{I}\left\{c_h^s = \min_{i=1, \dots, h} \{c_i^s\}\right\} \right).$$

**3.2.2. Look-Ahead Lower Bounds.** Note that the blocking lower bound  $b$  is only taking into account the current configuration. However, some relocations lead necessarily to an additional relocation. We refer to such relocations as “bad.” Formally, let  $s$  be a stack of a configuration, and  $\min(s)$  be the smallest label of a container in  $s$ . Recall that, if  $s$  is empty, we set  $\min(s) = C + 1$ . We say that the relocation of container  $c$  from stack  $s$  is bad if  $c > \max_{s'=1, \dots, S, s' \neq s} \{\min(s')\}$ . We propose to construct a lower bound that anticipates bad relocations.

The basic idea is based on a similar one used by Zhu et al. (2012) for the CRP. We consider the first look-ahead lower bound denoted by  $b_1(n)$ . By definition, we take  $b_1(n) = b(n) + d_1(n)$ , where  $b(n)$  is the blocking lower bound, and  $d_1(n)$  is the expected number of unavoidable bad relocations while performing the first removal. We compute the term  $d_1(n)$  by considering all realizations of the first target container. For each realization, we compute the number of unavoidable bad relocations and average them. Formally, for a given configuration  $n$ , consider  $U_n$  the set of potential next target container in  $n$  (which can be a singleton if it is already known), that is,  $U_n = \{c \mid c = \min_{s=1, \dots, S} (\min(s))\}$ . Based on the definition of a bad relocation, we compute the number of unavoidable bad moves for each  $u \in U_n$  denoted by  $\beta(n, u)$ , and we take

$$d_1(n) = \frac{1}{|U_n|} \sum_{u \in U_n} \beta(n, u),$$

or  $d_1(n) = \sum_{u \in U_n} p_{n,u} \beta(n, u)$ , where  $p_{n,u}$  is the probability that  $u$  is the next target container in  $n$  if the probabilities considered are not uniform (which

**Figure 12.** Example for Look-Ahead Lower Bounds

		4
		3
1	3	1

can be computed using  $(p_{n_i})_{n_i \in \Omega_n}$  if  $n$  is a chance node).

For example, in Figure 12, the presented configuration denoted by  $n$  is such that  $b(n) = 2$ . Now consider a container  $u \in U_n$ : if  $u$  is the container labeled 1 in stack 1, then there is no blocking container, so  $\beta(n, u) = 0$ ; if  $u$  is the other container labeled 1, the relocation of the container labeled 4 from stack 3 is necessarily a bad relocation, since  $\min(1) = 1 < 4$  and  $\min(2) = 3 < 4$ , but it is not the case for the blocking container labeled 3, hence  $\beta(n, u) = 1$ . Therefore,  $d_1(1) = 0.5(0 + 1) = 0.5$ , and  $b_1(n) = 2 + 0.5 = 2.5$ , hence giving a lower bound closer to the optimal solution than  $b(n)$ . Note that, if  $n$  has an empty stack, then  $\beta(n, u) = 0$  for all  $u \in U_n$ , and hence  $d_1(n) = 0$ .

We can refine this idea by trying to find unavoidable bad relocations for the second removal. In this case, the configuration depends on the first removal and the decisions that have been made accordingly. For clarity, consider that the first target container has been revealed, and denote it  $u_1$ . After retrieving  $u_1$ , only containers blocking  $u_1$  have changed from their initial position. It can be very challenging to detect future unavoidable bad moves for these containers. To bypass this issue, we consider that all containers blocking  $u_1$  are also removed, resulting in a configuration without  $u_1$  and its blocking containers. Given this new configuration denoted by  $n(u_1)$ , we can compute the expected number of unavoidable bad moves  $d_1(n(u_1))$ . Since  $u_1$  is actually random, we have to consider each scenario with its associated probability and compute a new configuration where blocking containers are retrieved with the target container. We denote the result  $d_2(n)$ , and it is a lower bound on the expected number of unavoidable bad relocations for the first two removals starting at  $n$ . Finally, our second look-ahead lower bound is given by  $b_2(n) = b(n) + d_2(n)$ .

**Algorithm 1** (Lower bound on the number of unavoidable bad relocations for the  $k$  first removals)

- 1: **procedure**  $[d_k(n)] = \text{UNAVOIDABLEBADRELOC}(n, k)$
- 2:   **if**  $k = 0$  or  $n$  has an empty stack or  $n$  is empty  
      **then**  $d_k(n) = 0$
- 3:   **else** let  $U_n = \{\text{containers with minimum label in } n\}$
- 4:     **if**  $k = 1$  **then**  $d_k(n) = \frac{1}{|U_n|} \sum_{u \in U_n} \beta(n, u)$
- 5:   **else**
- 6:     **for**  $u \in U_n$  **do**
- 7:       Let  $n(u)$  be the configuration  $n$  without  $u$   
          and all containers blocking  $u$

- 8:       Compute recursively  $d_{k-1}(n(u)) =$   
           UnavoidableBadReloc( $n(u), k-1$ )  
 9:       Compute  $d_k(n) = \frac{1}{|U_n|} \sum_{u \in U_n} \beta(n, u) + d_{k-1}(n(u))$ .

This idea can easily be generalized for  $k \geq 2$  by induction with  $b_k(n) = b(n) + d_k(n)$ . Here  $k$  is the number of removals that the lower bound considers to compute the expected number of unavoidable bad relocations (see pseudocode of Algorithm 1).

#### 4. PBFS, A New Optimal Algorithm for the SCR

Building on lower bounds introduced in Section 3, this section introduces, studies, and proves the optimality of one of the main contributions of this paper, the PBFS Algorithm.

##### 4.1. PBFS Algorithm

We start by giving the pseudocode of our algorithm, and we derive its optimality in Lemmas 3 and 4. PBFS takes two inputs, the configuration  $n$  for which we aim to compute  $f(n)$ , and a valid lower bound  $l$ . This algorithm uses a combination of four features to return  $f(n)$ . The first one is the BFS exploration of the tree based on a given lower bound  $l$ . We first compute  $f$  for the “most promising nodes,” because nodes with small lower bounds are more likely to result in small  $f$ . The second technique is stopping to compute  $f$  recursively after level  $\lambda^* = \max\{S, C_W\}$ , by calculating it either using  $b$  or the  $A^*$  algorithm defined later. In Algorithm 2,  $A^*(n)$  denotes the optimal number of relocations for node  $n$  obtained using the  $A^*$  algorithm. The third feature is pruning with a lower bound, revealing the suboptimality of some nodes without actually computing  $f$ . As its fourth feature, the algorithm uses the abstraction technique described previously.

##### Algorithm 2 (PBFS algorithm)

- 1: **procedure**  $[f(n)] = \text{PBFS}(n, l)$
- 2:   **if**  $\lambda_n \leq S$  ( $n$  has less than  $S$  containers)  
       **then**  $f(n) = b(n)$
- 3:   **else**
- 4:     **if**  $n$  is a *chance node* **then** start with  $\Psi_n^{\text{PBFS}} = \{\}$
- 5:     **for**  $n_i \in \Omega_n$  **do**  $n_i \leftarrow \text{ABSTRACT}(n_i)$
- 6:       **if** there exists  $m = n_i$  already in  $\Psi_n^{\text{PBFS}}$   
           **then**  $p_m^n \leftarrow p_m^n + 1/|\Omega_n|$
- 7:       **else if** there exists  $m = n_i$  already in the  
           decision tree **then** add  $m$  to  $\Psi_n^{\text{PBFS}}$  and  
            $p_m^n = 1/|\Omega_n|$
- 8:       **else** add  $n_i$  to  $\Psi_n^{\text{PBFS}}$ ,  $p_{n_i}^n = 1/|\Omega_n|$  and  
           compute  $f(n_i) = \text{PBFS}(n_i, l)$
- 9:     Compute  $f(n) = \sum_{n_i \in \Psi_n^{\text{PBFS}}} p_{n_i}^n f(n_i)$
- 10:   **else**  $n$  is a *decision node*
- 11:    **if**  $\lambda_n = C_W$  (the full retrieval order is  
       known) **then**  $f(n) = A^*(n)$

- 12:   **else** construct  $\Delta_n$  by considering all feasible  
       sets of decisions to deliver the target  
       container
- 13:    Compute  $l(n_i)$  for each  $n_i \in \Delta_n$
- 14:    Sort  $(n_{(1)}, n_{(2)}, \dots, n_{(|\Delta_n|)})$  in nondecreasing  
       order of  $l(\cdot)$
- 15:    Compute  $f(n_{(1)}) = \text{PBFS}(n_{(1)}, l)$
- 16:    Start with  $\Gamma_n^{\text{PBFS}} = \{n_{(1)}\}$  and  $k = 2$
- 17:    **while**  $k \leq |\Delta_n|$  and  $l(n_{(k)}) <$   
        $\min_{j=1, \dots, k-1} \{f(n_{(j)})\}$  **do**  $n_{(k)} \leftarrow$   
        $\text{ABSTRACT}(n_{(k)})$
- 18:      **if** there exists  $m = n_{(k)}$  already in the  
       decision tree **then** add  $m$  to  $\Gamma_n^{\text{PBFS}}$
- 19:      **else** add  $n_{(k)}$  to  $\Gamma_n^{\text{PBFS}}$  and compute  
        $f(n_{(k)}) = \text{PBFS}(n_{(k)}, l)$
- 20:      Update  $k = k + 1$
- 21:     $f(n) = r(n) + \min_{n_i \in \Gamma_n^{\text{PBFS}}} \{f(n_i)\}$ .

##### 4.1.1. Decreasing the Size of Decision Tree by Increasing $\lambda^*$ to $\max\{S, C_W\}$ .

**If  $C_W \leq S$ , then compute  $f(n)$  using  $b(n)$ .** Recall that, for every relocation, heuristic L chooses the stack with the fewest containers, breaking ties arbitrarily by choosing the leftmost stack. Note that L always provides a valid upper bound for the SCR. If we denote the resulting expected number of relocations to empty configuration  $n$  using L by  $f_L(n)$ , then we have  $f_L(n) \geq f(n)$ .

**Lemma 3.** Let  $n$  be a configuration with  $S$  stacks,  $T$  tiers, and  $C$  containers such that  $C \leq S$ , then we have

$$f_L(n) = b(n) = f(n).$$

**Proof.** Consider a retrieval order of containers from  $n$  that has a nonzero probability of occurring. If there are no blocking containers, then the lemma clearly holds. Otherwise, let  $c$  be one of the blocking containers for this retrieval order, and consider the first removal for which  $c$  has to be relocated. For this removal, there are at most  $S$  containers in the configuration, hence there exists at least one empty stack to relocate  $c$ . Since heuristic L always chooses empty stacks if one exists, L would move  $c$  to one of the existing empty stacks. Note that in this case,  $c$  will never be blocking again, and hence will never be relocated again. This observation holds for any blocking containers, thus L relocates each blocking container at most once.

Since this fact holds for any retrieval order with nonzero probability, by taking expectation on the retrieval order, we have  $f_L(n) \leq b(n)$ , thus  $f_L(n) \leq b(n) \leq f(n) \leq f_L(n)$ , which concludes the proof.

Lemma 3 states that for configurations with  $S$  containers or fewer, heuristic L is optimal for the SCR.

This implies that, for nodes at level  $S$ , we have access to the cost-to-go function using  $b(n)$ , as well as an optimal solution (provided by heuristic L). Hence PBFS can stop branching at  $\lambda^* = S$  (line 2 of Algorithm 2).

**If  $C_W > S$ , then compute  $f(n)$  using the  $A^*$  algorithm.** If  $n$  is a decision node at level  $C_W$ , the full order of retrieval is known, and computing  $f(n)$  reduces to solving a classical CRP, so we can leverage the existence of efficient solutions to the classical CRP such as the  $A^*$  algorithm, and take  $\lambda^* = C_W$ . Throughout the remainder of the paper,  $A^*$  refers to the improved version of this algorithm presented in Borjani et al. (2015b).

Combined with the two previous observations, we can take  $\lambda^* = \max\{S, C_W\}$ .

**4.1.2. Decreasing the Size of Decision Tree by Pruning Using Lower Bounds.** We would also like to reduce the size of the tree before level  $\lambda^*$ . For a decision node  $n$ , PBFS considers only a subset  $\Gamma_n^{\text{PBFS}}$  of all of the offspring  $\Delta_n$  (line 21 of Algorithm 2). Our goal is to set  $\Gamma_n^{\text{PBFS}}$  to still guarantee optimality.

First, PBFS generates all nodes  $n_i \in \Delta_n$  by considering all feasible sets of decisions to deliver the target container in  $n$  (line 12 of Algorithm 2), and for each of them, compute a lower bound  $l(n_i)$ , where  $l$  is the input lower bound (line 13 of Algorithm 2). Let  $(n_{(1)}, n_{(2)}, \dots, n_{(|\Delta_n|)})$  be the list of offspring of  $n$  sorted by nondecreasing lower bound (line 14 of Algorithm 2). The algorithm first considers  $n_{(1)}$ , adds it to  $\Gamma_n^{\text{PBFS}}$  and computes  $f(n_{(1)})$  recursively (lines 15–16 of Algorithm 2). For  $k \geq 2$ , we consider  $n_{(k)}$ 's sequentially, and check if  $l(n_{(k)}) < \min_{j=1, \dots, k-1} \{f(n_{(j)})\}$  (line 17 of Algorithm 2). If so, add  $n_{(k)}$  to  $\Gamma_n^{\text{PBFS}}$  and compute  $f(n_{(k)})$  recursively. If not, we stop branching on all nodes  $n_{(k)}, \dots, n_{(|\Delta_n|)}$ . An illustration of the pruning rule is shown in Figure 13 and the next lemma shows the optimality of this rule.

**Lemma 4.** Let  $n$  be a decision node in the decision tree, and  $\Gamma_n^{\text{PBFS}}$  be the subset of nodes considered for this node

in Algorithm 2, and constructed as aforementioned, then we have

$$\min_{m_i \in \Gamma_n^{\text{PBFS}}} \{f(m_i)\} = \min_{n_i \in \Delta_n} \{f(n_i)\}.$$

**Proof.** Let  $(n_{(1)}, n_{(2)}, \dots, n_{(|\Delta_n|)})$  be the list of offspring of  $n$ , sorted by nondecreasing lower bounds. We consider two cases:

- If  $\Gamma_n^{\text{PBFS}} = \Delta_n$ , the statement clearly holds.
- Otherwise, there exists  $k \leq |\Delta_n|$  such that  $l(n_{(k)}) \geq \min_{j=1, \dots, k-1} \{f(n_{(j)})\}$ , and  $\Gamma_n^{\text{PBFS}} = \{n_{(1)}, \dots, n_{(k-1)}\}$ . Note that we have  $\forall k' \geq k, f(n_{(k')}) \geq l(n_{(k')}) \geq l(n_{(k)}) \geq \min_{j=1, \dots, k-1} \{f(n_{(j)})\}$ . Hence

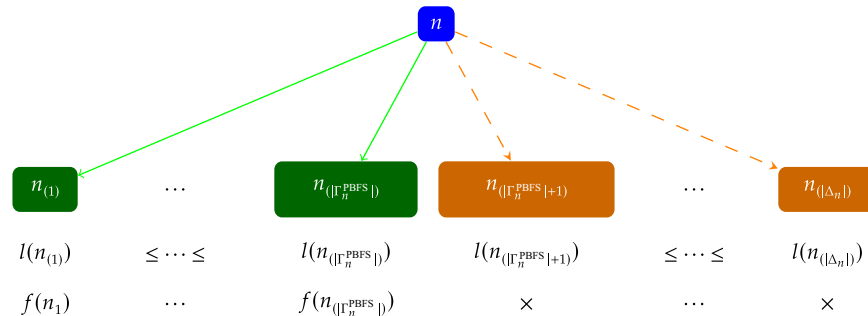
$$\min_{n_i \in \Delta_n} \{f(n_i)\} = \min_{j=1, \dots, k-1} \{f(n_{(j)})\} = \min_{m_i \in \Gamma_n^{\text{PBFS}}} \{f(m_i)\}.$$

We claim that increasing  $\lambda^*$  to  $\max\{S, C_W\}$  together with pruning in a Best-First-Search scheme, dramatically helps in the efficiency of PBFS while guaranteeing optimality. However, this algorithm faces the issue that  $|\Delta_n| = C_w!$  if  $n$  is a chance node. So if  $C_w$  is large, typically  $C_w \geq 4$ , the number of nodes to consider gets too large. We tackle this issue by considering a near-optimal algorithm in the Section 5. This leads us to consider an alternative to PBFS in the case of larger batches.

## 5. PBFSA, Near-Optimal Algorithm with Guarantees for Large Batches

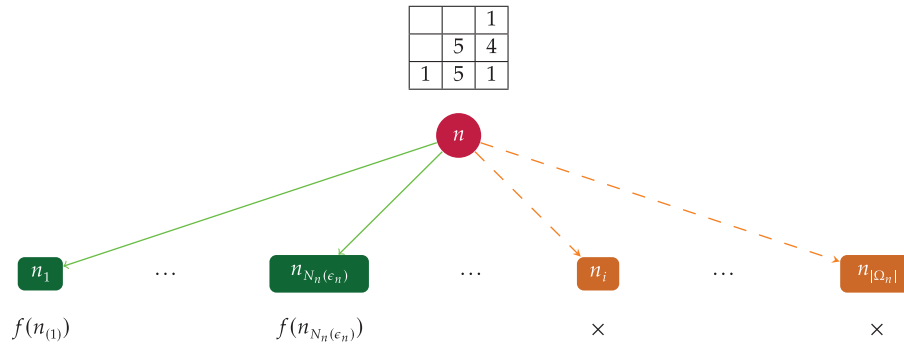
Building on PBFS introduced in Section 4, this section describes the randomized algorithm PBFSA and shows some theoretical guarantees on expectation. This new algorithm is identical to PBFS except when computing the value function of a chance node (lines 4 to 19 of Algorithm 3). To decrease the number of decision offspring to consider for each chance node, we sample a certain number of i.i.d. permutations and consider only the decision nodes associated with these permutations as illustrated in Figure 14. In particular, PBFSA uses  $f_{\min}(\cdot)$  and  $f_{\max}(\cdot)$ , lower and upper bound functions

**Figure 13.** (Color online) Illustration of the Pruning Rule



**Notes.** First, offspring are ordered by nondecreasing lower bounds. Then, we start computing the objective function starting at  $n_{(1)}$ . We stop computing the objective functions once the pruning rule is reached. Green nodes linked with full green arrows are nodes in  $\Gamma_n^{\text{PBFS}}$ , that is,  $f(\cdot)$  has been computed. Orange nodes linked with dashed orange arrows are nodes in  $\Delta_n \setminus \Gamma_n^{\text{PBFS}}$ , that is,  $f(\cdot)$  does not need to be computed, which is represented here by  $\times$ .

Figure 14. (Color online) Illustration of the Sampling Rule



Notes. In this figure, the smallest batch is batch 1; therefore  $w_{\min} = 1$ , and there are six containers, thus  $\lambda_n = 6$ . We have  $\lambda^* = 3$  so  $\delta_n = 1$  and thus  $\epsilon_n = \epsilon$ . These values allow us to compute the number of samples required  $N_n(\epsilon_n)$ . If  $N_n(\epsilon_n)$  is less than the total number of offspring  $|\Omega_n| = C_{w_{\min}}! = 3!$ , then we compute only  $f(\cdot)$  for sampled nodes. Let  $\Psi_n^{\text{PBFSa}}$  represent the subset of sampled nodes colored green and linked with full green arrows, and for which  $f(\cdot)$  needs to be computed. Note that  $|\Psi_n^{\text{PBFSa}}| = N_n(\epsilon_n)$ . Orange nodes linked with dashed orange arrows are nodes in  $\Omega_n \setminus \Psi_n^{\text{PBFSa}}$ , that is, they were not sampled and  $f(\cdot)$  does not need to be computed, which is represented here by  $\times$ . Finally, the approximate value of  $f(n)$  is the average of the objective values over all sampled nodes.

on  $f(\cdot)$  for offspring of chance nodes. In this paper, we use certain  $f_{\min}(\cdot)$  and  $f_{\max}(\cdot)$  defined in Equations (9)–(10), although others could be used. Combined with the fact that our problem has a finite number of sampling stages, this allows us to independently sample nodes to approximate the objective function. Using concentration inequalities, we can choose the number of samples needed to control the approximation error.

Since we perform a sampling at each chance node, PBFSa incurs an approximation error at each chance node. Lemma 5 proves that the total approximation error at the root node is, on average, the sum of all approximation errors from the root node to any leaf node. Therefore, consider a node  $n$ , then  $\delta_n$  is the number of chance nodes between  $n$  and any leaf node in the decision tree. If the target error is  $\epsilon$  at node  $n$ , PBFSa “allocates” evenly the remaining error to the next chance nodes, giving  $\epsilon_n = \epsilon/\delta_n$  error at each remaining chance node where sampling occurs.

Formally, when  $n$  is a given chance node, recall that PBFSa computes  $f(n) = (1/|\Omega_n|) \sum_{n_i \in \Omega_n} f(n_i)$ , where each  $n_i \in \Omega_n$  represents one retrieval order (a random permutation) of batch  $w$  (if  $\lambda_n = C - K_w + 1$ ). Instead, PBFSa computes the number of chance nodes between  $n$  and a leaf node denoted by  $\delta_n$ . If  $\epsilon$  is the target error at node  $n$ , the algorithm allocates  $\epsilon_n = \epsilon/\delta_n$  error for sampling at node  $n$  and the remaining (i.e.,  $\epsilon - \epsilon_n$ ) to its offspring (lines 13 and 18 of Algorithm 3). Using  $\epsilon_n$ , we compute  $N_n(\epsilon_n)$  and obtain a subset  $\Psi_n \subset \Omega_n$  from  $N_n(\epsilon_n)$  offspring drawn i.i.d. uniformly from  $\Omega_n$ . The important part is to define  $N_n(\epsilon_n)$ , such that  $\tilde{f}(n) = (1/|\Psi_n|) \sum_{m \in \Psi_n} f(m)$  is a “good” approximation of  $f(n)$ , that is,  $|\tilde{f}(n) - f(n)|$  is bounded by  $\epsilon$  on average. Note that if  $N_n(\epsilon_n) > C_{w_{\min}}!$ , we would need to sample more elements than the total number of offspring of  $n$ , and thus we do not sample (lines 15–18 of Algorithm 3).

PBFSa takes three input arguments, the configuration  $n$  for which we want to evaluate  $f$ , a valid lower bound  $l$ , and an upper bound  $\epsilon > 0$  on the total expected “error” ensured by the algorithm. It outputs  $\tilde{f}(n)$ , which is a randomized approximation of  $f(n)$ . Because of the samplings performed in line 10 in Algorithm 3, the output of PBFSa is random. The average error incurred by the algorithm is  $\mathbb{E}[|\tilde{f}(n) - f(n)|]$ , where the expectation is taken over the aforementioned samplings. Our main result (Lemma 5) states that PBFSa ensures  $\mathbb{E}[|\tilde{f}(n) - f(n)|] \leq \epsilon$ , in other words, PBFSa guarantees an average error of at most  $\epsilon$ . The proof of Lemma 5 can be found in Online Appendix B.

**Lemma 5.** Let  $n$  be a configuration with  $\lambda_n \geq 0$  containers,  $l$  be a valid lower bound function, and  $\epsilon > 0$ . If  $\tilde{f}(n) = \text{PBFSa}(n, l, \epsilon)$ , then

$$\mathbb{E}[|\tilde{f}(n) - f(n)|] \leq \epsilon.$$

**Sketch of the proof.** Lemma 5 is proven by backtracking from leaf nodes to the root node, that is, consider a node  $n$  at level  $\lambda_n$ , then the proof is done by induction on  $\lambda_n$ . To show that the expected absolute value of the error at node  $n$  (i.e.,  $\tilde{f}(n) - f(n)$ ) is bounded by  $\epsilon$ , we actually show that the expected positive and negative parts of the error are both bounded by  $\epsilon/2$ , which implies our result.

First, we consider the case where  $n$  is a decision node and we show that there is no additional error incurred by PBFSa at such node, that is, if all of the offspring of node  $n$  (in  $\Gamma_n^{\text{PBFSa}}$  and at level  $\lambda_n - 1$ ) have the expected positive and negative parts of their error bounded by  $\epsilon/2$  (the induction hypothesis), then the expected positive and negative parts of the error at node  $n$  are also bounded by  $\epsilon/2$ .

Second, we consider the case where  $n$  is a chance node, and we show that an additional error is incurred because of sampling. Using the lemmas proven below,

**Algorithm 3** (PBFSFA algorithm)

```

1: procedure  $[\tilde{f}(n)] = \text{PBFSFA}(n, l, \epsilon)$ 
2:   if  $\lambda_n \leq S$  then  $\tilde{f}(n) = b(n)$ 
3:   else
4:     if  $n$  is a chance node then start with  $\Psi_n^{\text{PBFSFA}} = \{\}$ . Let  $w_{\min}$  be such that  $\lambda_n = C - K_{w_{\min}} + 1$ 
5:     Compute  $\delta_n = \min\{w \in \{w_{\min}, \dots, W\} \mid \sum_{u=w_{\min}}^w C_u \geq \lambda_n - \lambda^*\}$  to get  $\epsilon_n = \epsilon / \delta_n$ 
6:     Compute  $f_{\max}(n)$  and  $f_{\min}(n)$  from Equations (9)–(10)
7:     Get  $N_n(\epsilon_n) = \pi(f_{\max}(n) - f_{\min}(n))^2 / (2\epsilon_n^2)$ 
8:     if  $N_n(\epsilon_n) \leq C_{w_{\min}}!$  then
9:       for  $i = 1, \dots, N_n(\epsilon_n)$  do
10:        Sample a random permutation, get corresponding  $n_i \in \Omega_n$  and  $n_i \leftarrow \text{ABSTRACT}(n_i)$ 
11:        if there is  $m = n_i$  already in  $\Psi_n^{\text{PBFSFA}}$  then  $p_m^n \leftarrow p_m^n + 1/N_n(\epsilon_n)$ 
12:        else if there is  $m = n_i$  already in the decision tree then add  $m$  to  $\Psi_n^{\text{PBFSFA}}$ ,  $p_m^n = 1/N_n(\epsilon_n)$ 
13:        else add  $n_i$  to  $\Psi_n^{\text{PBFSFA}}$ ,  $p_{n_i}^n = 1/N_n(\epsilon_n)$  and compute  $\tilde{f}(n_i) = \text{PBFSFA}(n_i, l, \epsilon - \epsilon_n)$ 
14:      else
15:        for  $n_i \in \Omega_n$  do  $n_i \leftarrow \text{ABSTRACT}(n_i)$ 
16:        if there exists  $m = n_i$  already in  $\Psi_n^{\text{PBFSFA}}$  then  $p_m^n \leftarrow p_m^n + 1/|\Omega_n|$ 
17:        else if there exists  $m = n_i$  already in decision tree then add  $m$  to  $\Psi_n^{\text{PBFSFA}}$ ,  $p_m^n = 1/|\Omega_n|$ 
18:        else add  $n_i$  to  $\Psi_n^{\text{PBFSFA}}$ ,  $p_{n_i}^n = 1/|\Omega_n|$  and compute  $\tilde{f}(n_i) = \text{PBFSFA}(n_i, l, \epsilon - \epsilon_n)$ 
19:      Compute  $\tilde{f}(n) = \sum_{n_i \in \Psi_n^{\text{PBFSFA}}} p_{n_i}^n \tilde{f}(n_i)$ 
20:    else  $n$  is a decision node
21:    if  $\lambda_n \leq C_W$  then  $\tilde{f}(n) = A^*(n)$ 
22:    else Construct  $\Delta_n$  by considering all feasible sets of decisions to deliver the target container
23:    Compute  $l(n_i)$  for each  $n_i \in \Delta_n$ 
24:    Sort  $(n_{(1)}, n_{(2)}, \dots, n_{(|\Delta_n|)})$  in nondecreasing order of  $l(\cdot)$ 
25:    Compute  $\tilde{f}(n_{(1)}) = \text{PBFSFA}(n_{(1)}, l, \epsilon)$ 
26:    Start with  $\Gamma_n^{\text{PBFSFA}} = \{n_{(1)}\}$  and  $k = 2$ 
27:    while  $k \leq |\Delta_n|$  and  $l(n_{(k)}) < \min_{j=1, \dots, k-1} \{\tilde{f}(n_{(j)})\}$  do  $n_{(k)} \leftarrow \text{ABSTRACT}(n_{(k)})$ 
28:    if there exists  $m = n_{(k)}$  already in the decision tree then add  $m$  to  $\Gamma_n^{\text{PBFSFA}}$ 
29:    else add  $n_{(k)}$  to  $\Gamma_n^{\text{PBFSFA}}$  and compute  $\tilde{f}(n_{(k)}) = \text{PBFSFA}(n_{(k)}, l, \epsilon)$ 
30:    Update  $k = k + 1$ 
31:     $\tilde{f}(n) = r(n) + \min_{n_i \in \Gamma_n^{\text{PBFSFA}}} \{\tilde{f}(n_i)\}$ .

```

the positive and negative parts of this additional error are bounded by  $\epsilon_n/2$ . Since all of the offspring of node  $n$  (in  $\Psi_n^{\text{PBFSFA}}$ ) are decision nodes at level  $\lambda_n$  and PBFSFA sets a target error of  $(\epsilon - \epsilon_n)/2$  for these nodes, then the first part of the proof shows that the positive and negative parts of the error of each offspring of node  $n$  are bounded by  $(\epsilon - \epsilon_n)/2$ . Combining both observations leads to  $n$  having the positive and negative parts of its error bounded by  $\epsilon/2$ , which proves the lemma.

**5.1. Hoeffding's Inequality Applied to the SCRP**

To prove this result, we use Hoeffding's inequality to compute the number of samples to ensure probabilistic guarantees. We first state the well-known inequality and a direct corollary. The proofs of Corollary 1, and Lemmas 6 and 7 can be found in Online Appendix C.

**Theorem 1** (Hoeffding's Inequality). *Let  $X \in [x_{\min}, x_{\max}]$  be a real-valued bounded random variable with mean value  $\mathbb{E}[X]$ . Let  $N \in \mathbb{N}$  and  $(X_1, \dots, X_N)$  be  $N$  i.i.d. samples of  $X$ . If  $\bar{X} = (1/N) \sum_{i=1}^N X_i$ , then we have*

$$\forall \delta > 0, \quad \mathbb{P}(\bar{X} - \mathbb{E}[X] > \delta) \leq \exp\left(\frac{-2N\delta^2}{(x_{\max} - x_{\min})^2}\right), \quad (5)$$

and

$$\forall \delta > 0, \quad \mathbb{P}(\bar{X} - \mathbb{E}[X] < -\delta) \leq \exp\left(\frac{-2N\delta^2}{(x_{\max} - x_{\min})^2}\right). \quad (6)$$

**Corollary 1.** *Let  $X \in [x_{\min}, x_{\max}]$  be a real-valued bounded random variable with mean value  $\mathbb{E}[X]$ . Let  $N \in \mathbb{N}$  and  $(X_1, \dots, X_N)$  be  $N$  i.i.d. samples of  $X$ . If  $\bar{X} = (1/N) \sum_{i=1}^N X_i$ , then  $\forall \epsilon > 0$  such that  $N \geq \pi(x_{\max} - x_{\min})^2 / (2\epsilon^2)$ , we have*

$$\mathbb{E}[(\bar{X} - \mathbb{E}[X])^+] \leq \frac{\epsilon}{2}, \quad (7)$$

$$\mathbb{E}[(\bar{X} - \mathbb{E}[X])^-] \leq \frac{\epsilon}{2}, \quad (8)$$

where  $x^+ = \max\{x, 0\}$  (respectively,  $x^- = -\min\{x, 0\}$ ) is the positive (respectively, negative) part of  $x$ .

To use Hoeffding's inequality, we need to define lower ( $f_{\min}$ ) and upper ( $f_{\max}$ ) bound functions, such that for each chance node  $n$ ,  $f_{\min}(n) \leq \min_{n_i \in \Omega_n} \{f(n_i)\}$  and  $f_{\max}(n) \geq \max_{n_i \in \Omega_n} \{f(n_i)\}$ .

**Lemma 6.** *Let  $n$  be a chance node, if*

$$f_{\min}(n) = \min_{n_i \in \Omega_n} \{b(n_i)\}, \quad (9)$$

and

$$f_{\max}(n) = \min\left\{((\lambda_n - S)(T - 1))^+ + (\min\{S, \lambda_n\} - 1), \right. \\ \left. (2\lceil \lambda_n / S \rceil - 1) \max_{n_i \in \Omega_n} \{b(n_i)\}\right\}, \quad (10)$$

then

$$f_{\min}(n) \leq \min_{n_i \in \Omega_n} \{f(n_i)\} \quad \text{and} \quad f_{\max}(n) \geq \max_{n_i \in \Omega_n} \{f(n_i)\}.$$

Note that the previous lemma involves computing  $\min_{n_i \in \Omega_n} \{b(n_i)\}$  and  $\max_{n_i \in \Omega_n} \{b(n_i)\}$ . The following lemma provides an efficient way to compute these values.

**Lemma 7.** Let  $n$  be a chance node, and  $w_{\min} \in \{1, \dots, W\}$  be such that  $\lambda_n = C - K_{w_{\min}} + 1$  (i.e., the minimum batch in  $n$ ). For each stack  $s$  of  $n$  with  $H^s \geq 1$  containers, let  $(c_h^s)_{h=1, \dots, H^s}$  be the containers in  $s$ , where  $c_1^s$  is the container at the bottom and  $c_{H^s}^s$  is the container at the top (see Figure 11, for the case  $H = H^s$ ). Finally, consider  $C_{w_{\min}}^s = |\{c_h^s = K_{w_{\min}}, h = 1, \dots, H^s\}|$ . Then we have

$$\begin{aligned} \min_{n_i \in \Omega_n} \{b(n_i)\} &= \sum_{\substack{s=1, \dots, S \\ H^s \geq 1}} \left( H^s - C_{w_{\min}}^s - \sum_{\substack{h=1, \dots, H^s \\ c_h^s \neq K_{w_{\min}}}} \frac{\mathbb{I}\{c_h^s = \min_{i=1, \dots, h} \{c_i^s\}\}}{\sum_{i=1}^h \mathbb{I}\{c_h^s = c_i^s\}} \right), \end{aligned} \quad (11)$$

and

$$\max_{n_i \in \Omega_n} \{b(n_i)\} = \sum_{\substack{s=1, \dots, S \\ H^s \geq 1}} \left( H^s - \sum_{\substack{h=1, \dots, H^s \\ c_h^s \neq K_{w_{\min}}}} \frac{\mathbb{I}\{c_h^s = \min_{i=1, \dots, h} \{c_i^s\}\}}{\sum_{i=1}^h \mathbb{I}\{c_h^s = c_i^s\}} \right). \quad (12)$$

**Nonuniform case.** Similar to the blocking lower bound, we can extend Lemma 7 to the case where probabilities are not uniform across retrieval orders. Recall that  $q_{c_h^s}$  denotes the probability that  $c_h^s$  is the first one to be retrieved among the ones positioned below in its stack and with the same label. Then we have

$$\begin{aligned} \min_{n_i \in \Omega_n} \{b(n_i)\} &= \sum_{\substack{s=1, \dots, S \\ H^s \geq 1}} \left( H^s - C_{w_{\min}}^s - \sum_{\substack{h=1, \dots, H^s \\ c_h^s \neq K_{w_{\min}}}} q_{c_h^s} \mathbb{I}\left\{c_h^s = \min_{i=1, \dots, h} \{c_i^s\}\right\} \right), \end{aligned}$$

and

$$\max_{n_i \in \Omega_n} \{b(n_i)\} = \sum_{\substack{s=1, \dots, S \\ H^s \geq 1}} \left( H^s - \sum_{\substack{h=1, \dots, H^s \\ c_h^s \neq K_{w_{\min}}}} q_{c_h^s} \mathbb{I}\left\{c_h^s = \min_{i=1, \dots, h} \{c_i^s\}\right\} \right).$$

## 6. Computational Experiments

Having introduced lower and upper bounds, PBFS, PBFSA, and theoretical guarantees in previous sections, we present several experimental results in this section to understand the effectiveness of our algorithms for the SCRP. For clarity, we refer to the set of instances from Ku and Arthanari (2016) as the *existing data set*. We present four sets of experiments:

1. Based on instances from the existing data set, which have relatively small batches, we test the PBFS algorithm, as well as the two new heuristics and our lower bounds.

2. We slightly modify the existing data set to obtain the *modified data set*, to obtain instances with relatively larger batches. We test the efficiency of PBFSA on this modified data set.

3. Based on the existing data set, we show that PBFS improves on the algorithm proposed in Ku and Arthanari (2016) for the online model. Moreover, the two new heuristics (EM and EG) outperform the ERI algorithm on expectation for the majority of the instances in the data set.

4. We change the existing data set by considering that all containers belong to a unique batch. We show strong computational evidence to support Conjecture 1, which states that the leveling policy is optimal for the SCRP under the online model with a unique batch.

All experiments are performed on a MacBook Pro with 2.2 GHz Intel Core i7 processor and 8 GB RAM, and the programming language is MATLAB 2016a. Finally, all results and instances used in this section are available at <https://github.com/vgalle/StochasticCRP>.

### Implementation of heuristics.

1. Computing the number of relocations using  $b$  when there are  $S$  containers or less: In the retrieval process, when there are  $S$  containers or fewer remaining in the configuration, the expected number of relocations performed by ERI, EM, EG, and L is computed using  $b$ . This is motivated by the following observation: ERI, EM, EG, and L are optimal when there are  $S$  containers or fewer remaining in the configuration, and Lemma 3 shows that the optimal expected number of relocations in this case is equal to  $b$ . Therefore, for all heuristics (except Random), instead of running simulations until there are no containers left, we stop when there are  $S$  containers left and compute the expected number of relocations using  $b$  instead.

2. Estimating the expected number of relocations using sampling: To estimate the exact objective value for a given heuristic, one would have to consider all possible retrieval scenarios. Instead, for each heuristic unless otherwise specified, we report the average over 5,000 samples (of retrieval orders) for each instance, where samples are uniformly drawn at random.

**Existing data set description.** The full description of the data set can be found in Ku and Arthanari (2016), and the original data set is available at <http://crp-timewindow.blogspot.com>. Note the following:

- Configuration sizes vary from  $T = 3, \dots, 6$  tiers, and  $S = 5, \dots, 10$  stacks.
- Two occupancy rates are considered, 50% and 67%. The occupancy rate ( $\mu \in [0, 1]$ ) is defined such that the initial number of containers is  $C = \text{round}(\mu \times S \times T)$ , where  $\text{round}(x)$  rounds  $x$  to the closest integer. Therefore, a given triplet  $(T, S, \mu)$  is equivalent to a given triplet  $(T, S, C)$ , and note that if  $C = \text{round}(0.67 \times S \times T)$ , the condition  $0 \leq C \leq ST - (T - 1)$  is satisfied.
- Given a configuration size ( $T$  and  $S$ ) and an occupancy rate ( $\mu$ ) resulting in a given initial number of

containers (C), the data set includes 30 different initial configurations.

- For all 1,440 instances, the ratio between the number of batches and C is taken to be around half, that is, there are, on average, two containers per batch, which is the smallest size for a batch.

In all our experiments, the time limit is set to an hour, and the first look-ahead lower bound  $b_1$  is used as input for both PBFS and PBFSA. All instances are solved by heuristics and lower bounds within seconds or less.

### 6.1. Experiment 1: Batch Model with Small Batches

Table 1 gives a summary of the results as follows: ✓ indicates that all 30 instances are solved optimally by PBFS. In this case, the average solution time to solve these instances is given in seconds. Otherwise, the number of instances solved optimally is provided in red and in the form  $x/30$ . This table shows the efficiency of PBFS as it can solve all instances except two, for  $T = 3$  and  $T = 4$ . Most important, the average time to solve these instances is under 10 seconds for these problem sizes. Since many ports today have a maximum tier requirement of four and need fast solutions, PBFS could be used in practice in the case of small batches. However, for  $T = 5$  and  $T = 6$ , PBFS cannot solve all instances optimally in a timely manner. This suggests that, as the problem grows slightly, given instances become very hard to solve, which should not be a surprise, given the  $\mathcal{NP}$ -hardness of the problem. To avoid such situations in real operations, heuristics can be used to provide a “good” suboptimal solution (good in the sense of not being too far from optimality). Therefore, we want to evaluate the performance of

these heuristics to know which one should be used in real operations.

We measure the performance of heuristics and the tightness of lower bounds in Tables 5 and 6. Concerning lower bounds,  $b$  encompasses a significant number of relocations. Adding unavoidable bad relocations in  $b_1$  and  $b_2$  improves the lower bound slightly. Yet experiments seem to confirm that  $b_2(n) - b_1(n) \leq b_1(n) - b(n)$  holds, supporting our intuition that the relative increase of lower bounds  $b_k(n) - b_{k-1}(n)$  decreases with  $k$ .

Concerning heuristics, EG and EM clearly outperform ERI as they result in lower expected numbers of relocations. When we have access to the optimal solutions, both heuristics are, on average, at most 2% more than the optimal solution. We expect this behavior to be similar for larger instances, however, we only have access to lower bounds to evaluate their performances. In this case, heuristics are, on average, at most 11% more than  $b_2$ , hence at most 11% from the optimal solution (even though we believe that this number is very conservative, as our lower bounds are not “tight”). Therefore, both EG and EM appear to be good solutions for the SCRP under the batch model with small batches. In this case, we recommend using EM for its simplicity of implementation and its understandability.

### 6.2. Experiment 2: Batch Model with Larger Batches

**6.2.1. Modifying Existing Instances.** For the sake of reproducibility, we use the existing set of instances, but modify it slightly to consider larger batches. To create these instances, for each original instance  $n$ , consider  $n'$  with the same containers in the same configuration.

**Table 1.** Instances Solved by PBFS in the Batch Model with Small Batches

S	T	3		4		5		6	
		50%	67%	50%	67%	50%	67%	50%	67%
5	Fill rate								
	C	8	10	10	13	13	17	15	20
	Solved	✓	✓	✓	✓	✓	<b>28/30</b>	✓	<b>15/30</b>
6	Time (s)	0.01	0.02	0.03	0.12	0.17		5.17	
	C	9	12	12	16	15	20	18	24
	Solved	✓	✓	✓	✓	✓	<b>25/30</b>	✓	<b>14/30</b>
7	Time (s)	0.01	0.03	0.04	0.86	2.90		15.94	
	C	11	14	14	19	18	23	21	28
	Solved	✓	✓	✓	✓	✓	<b>24/30</b>	<b>23/30</b>	<b>5/30</b>
8	Time (s)	0.02	0.04	0.04	0.83	1.37			
	C	12	16	16	21	20	27	24	32
	Solved	✓	✓	✓	✓	✓	<b>20/30</b>	<b>22/30</b>	<b>5/30</b>
9	Time (s)	0.01	0.06	0.16	10.04	6.84			
	C	14	18	18	24	23	30	27	36
	Solved	✓	✓	✓	✓	<b>29/30</b>	<b>10/30</b>	<b>19/30</b>	<b>2/30</b>
10	Time (s)	0.03	0.10	0.37	8.84				
	C	15	20	20	27	25	34	30	40
	Solved	✓	✓	✓	<b>28/30</b>	<b>29/30</b>	<b>12/30</b>	<b>22/30</b>	<b>2/30</b>
	Time (s)	0.03	0.10	0.54					

**Table 2.** Instances Solved by PBFSA in the Batch Model with Larger Batches

S	T	3		4		5		6	
		50%	67%	50%	67%	50%	67%	50%	67%
5	C	8	10	10	13	13	17	15	20
	Solved	✓	✓	✓	✓	✓	21/30	✓	3/30
	Time (s)	0.08	0.29	0.14	4.55	3.20		72.70	
6	C	9	12	12	16	15	20	18	24
	Solved	✓	✓	✓	✓	✓	18/30	27/30	1/30
	Time (s)	0.08	0.47	0.25	126.37	14.74			
7	C	11	14	14	19	18	23	21	28
	Solved	✓	✓	✓	29/30	✓	9/30	14/30	0/30
	Time (s)	0.13	0.71	0.58		17.74			
8	C	12	16	16	21	20	27	24	32
	Solved	✓	✓	✓	28/30	29/30	6/30	17/30	1/30
	Time (s)	0.08	1.67	1.26					
9	C	14	18	18	24	23	30	27	36
	Solved	✓	✓	✓	26/30	25/30	5/30	15/30	0/30
	Time (s)	0.13	1.49	1.47					
10	C	15	20	20	27	25	34	30	40
	Solved	✓	✓	✓	22/30	29/30	7/30	14/30	0/30
	Time (s)	0.17	0.79	3.10					

Yet, if  $w$  is the batch of a container  $c$  in  $n$ , then we take the batch of  $c$  in  $n'$  to be  $w' = \lceil w/\gamma \rceil$ , where  $\gamma > 1$ , that is, we merge  $\gamma$  batches together. In these experiments, we take  $\gamma = 2$ , which implies that batches have an average size of four.

**6.2.2. Target Error  $\epsilon$ .** To set our target error, we consider the following. Let  $n_0$  be a given instance, and set  $\epsilon = b(n_0)/2$ . In this case, we know that  $\epsilon \leq f(n_0)/2$ , which implies that we are making an error of at most 50%. Note that this error is very conservative because of two things: first,  $b(n_0)$  is not necessarily representative of  $f(n_0)$ , especially if  $n_0$  has many containers. Second, the number of samples given by Hoeffding's inequality is also very conservative, probably making our approximation substantially more accurate than what we can theoretically prove.

**6.2.3. Results.** The results are summarized in Table 2. Similar to Table 1, a ✓ indicates that all 30 instances are solved approximately by PBFSA within the given expected error. In this case, the average solution time to solve these instances is given in seconds. Otherwise, the number of instances solved is provided in red and in the form  $x/30$ . This table shows that PBFSA presents several advantages. First, it solves most of the instances with  $T = 4$  and  $S \leq 9$  approximately within two minutes, while we note that PBFS was not able to solve most of these. Moreover, as can be seen in Tables 7 and 8, PBFSA still outperforms the best heuristics despite the fact that we only set the theoretical guarantee to 50% of optimality. Together, these two advantages show the practicality of PBFSA for problem sizes typically encountered in real ports. Moreover, we note that increasing the batch size appears to make

the problem significantly more complicated to solve as we can solve optimally larger instances in Experiment 1 (see Table 1). Finally, we remark that similar conclusions of Experiment 1 can be drawn for lower and upper bounds (see Tables 7 and 8).

### 6.3. Experiment 3: Online Model and Comparison with Ku and Arthanari (2016)

Table 3 gives a summary similar to the two previous experiments. In addition, we report the results of Ku and Arthanari (2016) who set a time limit of eight hours for each instance. In this table, ✓ (✓) indicates that all 30 instances are solved to optimality by both PBFS and Ku and Arthanari (2016). In this case, the average solution time in seconds to solve these instances is given for PBFS and for Ku and Arthanari (2016) in parentheses. The sign ✓ indicates that all 30 instances are solved optimally only by PBFS but not by Ku and Arthanari (2016). In this case, only the average solution time to solve these instances with PBFS is given in seconds. Otherwise, the number of instances solved by PBFS is provided in red and in the form  $x/30$ .

The results show strong evidence that our solution significantly improves the best existing results for the SCRPP under the online model, given that we solve many larger instances optimally. Furthermore, it also outperforms the most recent algorithm in solution time for the problem sizes it can solve. It appears that for problems for which we can solve all (or almost all) instances, most instances are “easy” to solve as the algorithm finds a solution within seconds. However, as in the batch model, there exist some instances for

**Table 3.** Instances Solved by PBFS and Ku and Arthanari (2016) in the Online Model with Small Batch

S	T	3		4		5		6	
		50%	67%	50%	67%	50%	67%	50%	67%
5	C	8	10	10	13	13	17	15	20
	Solved	✓ (✓)	✓	✓ (✓)	✓	✓ (✓)	28/30	✓	18/30
	Time (s)	0.01 (0.02)	0.02	0.01 (2.51)	0.09	0.16 (2,483.30)		3.74	
6	C	9	12	12	16	15	20	18	24
	Solved	✓ (✓)	✓	✓ (✓)	✓	✓	25/30	✓	15/30
	Time	0.01 (0.01)	0.04	0.06 (139.08)	0.81	1.85		14.92	
7	C	11	14	14	19	18	23	21	28
	Solved	✓ (✓)	✓	✓ (✓)	✓	✓	24/30	23/30	5/30
	Time (s)	0.02 (0.33)	0.04	0.04 (207.62)	0.67	1.38			
8	C	12	16	16	21	20	27	24	32
	Solved	✓ (✓)	✓	✓	✓	✓	20/30	22/30	5/30
	Time (s)	0.01 (0.33)	0.05	0.10	8.29	5.85			
9	C	14	18	18	24	23	30	27	36
	Solved	✓ (✓)	✓	✓	✓	29/30	10/30	19/30	2/30
	Time (s)	0.02 (32.24)	0.09	0.38	7.26				
10	C	15	20	20	27	25	34	30	40
	Solved	✓ (✓)	✓	✓	28/30	29/30	12/30	16/30	2/30
	Time (s)	0.03 (58.85)	0.08	0.52					

which the optimal solution still requires an exponential number of nodes, which makes our algorithm not tractable.

In Tables 9 and 10, we also report in parentheses the averages for ERI and Random found by Ku and Arthanari (2016). The results for Random are consistent. However, we find significantly better results for our implementation of ERI. This is unexpected since the only difference between the two implementations is the use of lower bound  $b$ , when the configuration has fewer than  $S$  containers remaining. Nevertheless, ERI should also be optimal in this case, as it reduces to heuristic L. So this should not affect the expected number of relocations, and we cannot explain this difference. Finally, we point out that the results are quite similar to those in Experiment 1. Indeed, the existing data set has relatively small batches (on average, two containers), which inherently makes the two models, batch and online, very close to each other.

#### 6.4. Experiment 4: Online Model with a Unique Batch

In this experiment, we consider the existing data set, but assign all containers into a unique batch ( $W = 1$ ). We consider the SCRP under the online model, where containers are revealed one at a time. Note that, in this case, each container is equally likely to be retrieved, and it is equivalent to know no-information about containers relative retrieval order. For each instance, we solve it twice: first using PBFS, and then using heuristic L, for which we sample 10,000 scenarios (this is different from the 5,000 samples considered in previous experiments). We report the results in Table 4. In this table,

for each problem size, we report the expected optimal number of relocations averaged over 30 instances. The sign “—” means that all 30 instances could not be solved optimally with PBFS within the given time limit of an hour. Note that the expected number of relocations using heuristic L reported in this experiment might be less than the one of PBFS; this is only because of the fact that we are sampling. Intuitively, L should be the optimal solution in this setting, and this experiment shows strong evidence that the next conjecture holds.

**Conjecture 1.** Consider a configuration  $n$  with a unique batch. Let  $f^o(n)$  be the minimum expected number of relocations to empty  $n$  under the online model, and let  $f^{o,L}(n)$  be the expected number of relocations performed by the leveling heuristic under the online model, then

$$f^o(n) = f^{o,L}(n). \quad (13)$$

This conjecture could also be made in the dynamic case, when containers arrive to be stacked. These results would have important ramifications for port operations, namely, the optimal policy to minimize relocations, when no information is given in advance, is leveling configurations.

## 7. Discussion

Managing relocation moves is one of the main challenges in the storage yard of container terminals, because it directly affects the costs and efficiency of yard operations. The container relocation problem, notorious for its computational intractability, addresses this issue. In this paper, we extend the CRP to the more practical case in which the retrieval order of

**Table 4.** Instances Solved with PBFS and Heuristic L in the Online Model with a Unique Batch

S	T	3		4		5		6	
		50%	67%	50%	67%	50%	67%	50%	67%
5	C	8	10	10	13	13	17	15	20
	PBFS	2.08	3.33	3.54	6.53	6.56	12.05	9.13	17.28
	L	2.08	3.33	3.54	6.52	6.57	12.06	9.14	17.29
6	C	9	12	12	16	15	20	18	24
	PBFS	2.10	4.04	4.23	8.02	7.01	13.48	10.73	20.13
	L	2.10	4.04	4.23	8.02	7.01	13.48	10.73	20.13
7	C	11	14	14	19	18	23	21	28
	PBFS	2.69	4.60	4.82	9.55	8.58	14.75	12.22	23.14
	L	2.69	4.61	4.82	9.55	8.58	14.74	12.22	23.14
8	C	12	16	16	21	20	27	24	32
	PBFS	2.61	5.19	5.51	9.96	9.12	17.75	13.83	—
	L	2.62	5.19	5.51	9.95	9.12	17.75	13.83	26.04
9	C	14	18	18	24	23	30	27	36
	PBFS	3.31	5.72	6.10	11.58	10.89	19.15	—	—
	L	3.31	5.72	6.10	11.58	10.89	19.14	15.40	28.84
10	C	15	20	20	27	25	34	30	40
	PBFS	3.36	6.38	6.68	12.98	11.13	22.07	—	—
	L	3.36	6.38	6.67	12.99	11.13	22.06	16.87	31.77

containers is not known far in advance. First, we introduce a new stochastic model, called the batch model, show the applicability of this model, and compare it theoretically with the existing model of Zhao and Goodchild (2010). Then, we derive lower bounds and fast and efficient heuristics for the SCRP. Subsequently, we develop two novel algorithms (PBFS and PBFSa) to solve the stochastic CRP in different settings. Efficiencies of all algorithms are supported through computational experiments, for which all results are made available online at <https://github.com/vgalle/StochasticCRP>. Finally, using our solution methods and based on extensive experiments, we conjecture the optimality of the simple leveling heuristic in the online stochastic setting. More generally, the methods developed in this paper apply to multistage stochastic optimization problems, where the number of stages is finite, the set of feasible actions at each stage is finite, the objective function is bounded, and bounds on the objective function can be easily computed.

Future work could include the proof of Conjecture 1. Important future work can also be done on the optimal design of time windows for a TAS. On one hand, small time windows imply more information on the retrieval sequence, hence higher operational efficiency of port operators. On the other hand, large time windows insure higher flexibility for truck drivers and a high rate of on-time arrivals. To balance this trade-off, one would need to quantify two important metrics with respect to the expected number of relocations: the “value of information” and the assignment of containers to “wrong” batches. Finally, in the grand scheme of port operations, the study of stacking and retrieving simultaneously, as well as the extension in the row

dimension of blocks, is important for future studies of operations to take into account.

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