

Computational Methods Assignment 6 - Monte Carlo Integration

Martik Aghajanian

1 Question 1 - Discussion of the Method

In this program, the *Mcint* class takes a specific, potentially multivalued function of N_d variables to integrate between upper and lower limits, each individually specified by N_d -dimensional vectors. To integrate the function between the limits using importance sampling the user must supply the function to be integrated and the probability distribution to weight the samples. The user must also supply either (a) a function to generate the distribution of random numbers from a uniform deviate using the transformation method or (b) a data structure which contains a large array of N_d -component vectors distributed by an arbitrary N_d -variable probability function. The former applies if the probability function is a function of one variable and has an easily invertible cumulative distribution function. The uniform deviate generator used for this was the same as for assignment 3.

In this method, the function $f(\vec{x})$ (or $f(\vec{x})/p(\vec{x})$ if importance sampling) is sampled over volume V at M_{init} values, to calculate the average $V\langle f \rangle$, or $\langle f/p \rangle$ for importance sampling (no factor of V). At each step, an additional amount of samples, equal to the number currently included in the average, M , are taken which doubles the number of samples. The old evaluation of the integral is multiplied by M and then the further M samples are added to it, to later divide by $2M = M_{\text{new}}$. As a measure of error, the standard error is used as a statistical metric to determine the point to stop iterating the method. At each step, as well as calculating $\langle f \rangle$ (or $\langle f/p \rangle$), the value of $\langle f^2 \rangle$ (or $\langle f^2/p^2 \rangle$) is also calculated to obtain the standard error according to:

$$\begin{aligned}\sigma_m^{(1)} &= \sqrt{\frac{\langle f^2 \rangle - \langle f \rangle^2}{M}} \\ \sigma_m^{(2)} &= 2\sqrt{\frac{\langle f^2 \rangle - \langle f \rangle^2}{M}} \\ \sigma_m^{(2.58)} &= 2.58\sqrt{\frac{\langle f^2 \rangle - \langle f \rangle^2}{M}}\end{aligned}$$

The first and second values of the standard errors give the bounds around the estimated values of the integration for the 68% and 95.5% confidence intervals of the estimate of the integral (\hat{I}) respectively. Using either of these for the estimate of the error would be equivalent to saying that with 68% or 95.5% confidence, the value calculated for this integral lies within these bounds. The error estimate that was used for this investigation is the 99% confidence interval. Using the value of this, it can be stated with 99% confidence that the true value of the integral lies between $\hat{I} - \sigma_m^{(2.58)}$ and $\hat{I} + \sigma_m^{(2.58)}$. A 99% certainty that the value calculated lies in this range is used because the evaluation can stop as soon as this is reached, without having to over compensate for more significant uncertainty of the value. This gives a good criterion for convergence of the integral evaluation such that the addition of extra samples to the estimate will cease when $\sigma_m^{(2.58)} < \epsilon = 10^{-6}$. This can be used due to the large number of samples taken, which means that from an arbitrary distribution, the central limit theorem states that as the number of samples tends to infinity, the probability distribution function tends to that of a normal distribution. This expressions means that the error, as expected, should decrease as $M^{-1/2}$ as the number of samples M is increased. This means that by doubling the samples it is expected that the error should decrease by a factor of $\sqrt{2}$ each time.

2 Question 2 - Uniform vs Importance sampling

The error function was supplied as the integrand and firstly the evaluation of $\text{erf}(2)$ was performed using the uniform sampling of points in an integral, later to be calculated using a distribution $\text{pdf}_1(y) = 0.98 - 0.48y$ with a corresponding cumulative distribution of:

$$\text{pdf}_1(y) = -0.48y + 0.98 = x \implies y = \text{cdf}_1^{-1}(x) = \frac{0.98 - \sqrt{0.98^2 - 0.96x}}{0.48}$$

The cumulative distribution function for this is a quadratic, so there are two roots. However, only the '+' answer of the quadratic equation gives a viable transformation, so this is used. Also to test the effect of different probability distributions on the convergence of the method, a second probability distribution function was used which bared a similar functional form to the integrand.

$$\text{pdf}_2(y) = \frac{3e^{-1.5y}}{2(1 - e^{-3})} \implies y(x) = \text{cdf}_2^{-1}(x) = -\frac{2}{3}\ln(1 - (1 - e^{-3})x)$$

The function $f(\vec{x})$ was passed to the generalised integrator class along with the probability distribution function and cumulative distribution. A step constitutes the doubling of the number of samples and calculation of the average and the average square of the integrand. After 22 steps, it became too computationally cumbersome to iterate the method.

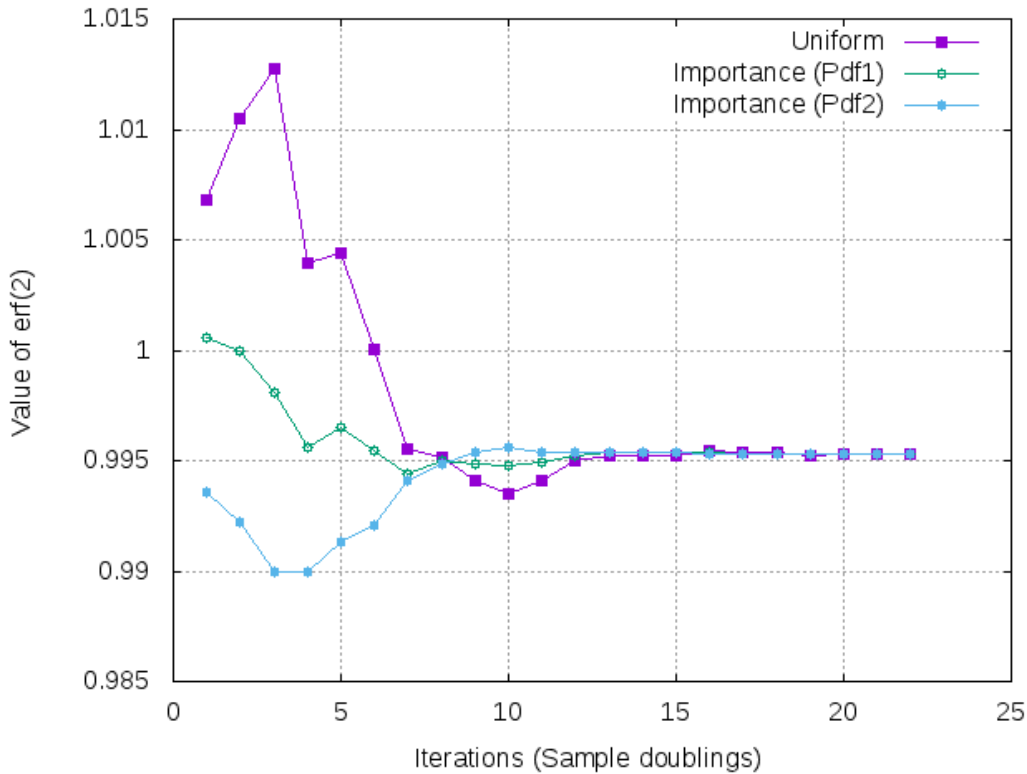


Figure 1: Convergence of values for $\text{erf}(2)$ with repeated doubling of the sample size for uniform sampling and importance sampling with both $\text{pdf}_1(x)$ and $\text{pdf}_2(x)$.

The value of $\text{erf}(2)$ after 22 steps, the value of $\sigma_m^{(2.58)}$ after 22 steps, and the predicted iterations required to obtain an accuracy of $\epsilon = 10^{-6}$ are shown in Table 1. This shows that implementing importance sampling results in a higher convergence of the integral, with a smaller error $\sigma_m^{(2.58)}$ than the value obtained by uniform sampling, after 22 steps. The estimated value of the integral and the associated $\sigma_m^{(2.58)}$ is shown in Fig. 1 and Fig. 2 respectively, against the number of steps/iterations in which the number of samples required is

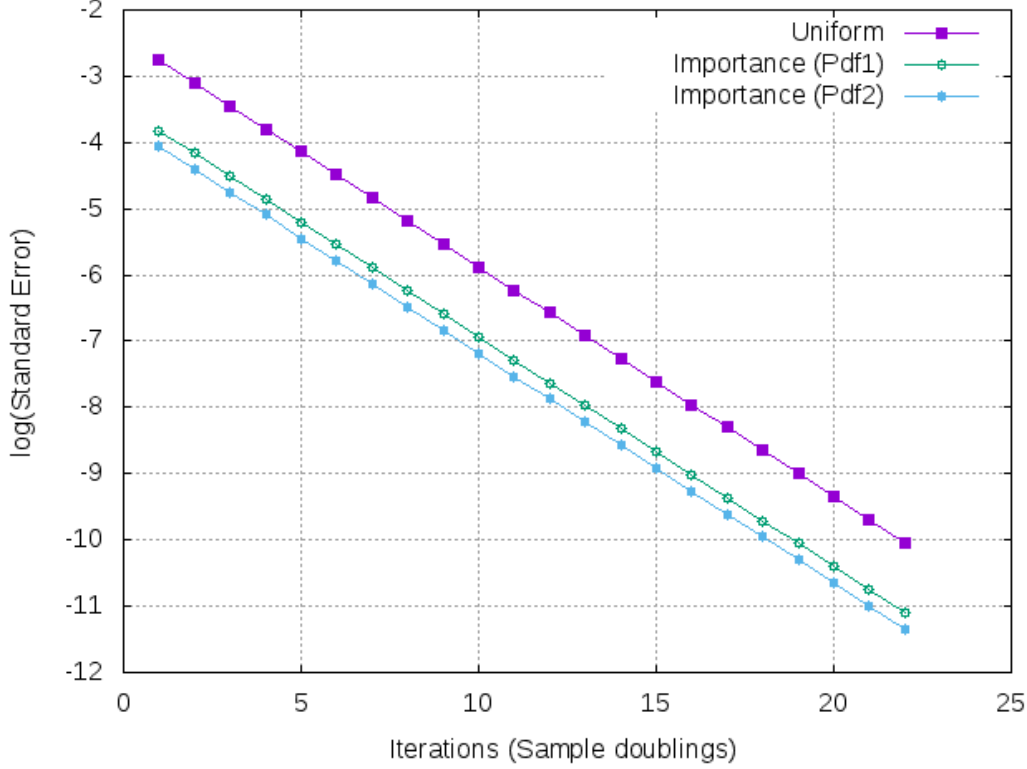


Figure 2: The logarithm of the error $\sigma_m^{(2.58)}$ of the evaluation of $\text{erf}(2)$ using Monte Carlo integration with increasing number of steps in which the sample size doubles each time.

doubled, from an initial number of $M_{\text{init}} = 500$. Firstly, Fig. 1 shows that this method of integration does converge to a value that is in the vicinity of the true value 0.995322 to $\epsilon = 10^{-6}$. It also can be seen that as the number of steps increased, the factor that the 99% standard error $\sigma_m^{(2.58)}$ decreased by, converged to $\sqrt{2}$, and this can be seen from the data tables outputted by the program. This meant that, assuming that the error decreases by a factor of $\sqrt{2}$ each step, the number of steps required to reach the convergence criterion of $\epsilon = 10^{-6}$ could be predicted. This was found by taking the current $\sigma_m^{(2.58)}$ value at 22 steps, finding how many divisions by $\sqrt{2}$ were needed to obtain $\sigma_m^{(2.58)} < \epsilon$, and adding that to the number of steps already carried out (22). The number of samples this corresponds to is $M_{\text{start}} \times 2^{n_r}$ where $M_{\text{start}} = 500$ is the initial number of iterations carried out in the first step, and n_r is the number of required iterations for convergence. However, this number of samples required is not necessarily the minimum number, since this method overshoots due to the fact that the number of samples doubles each step. If the error after including N samples is ϵ_N , then the constant of proportionality c (assuming that $\epsilon_N = c \times N^{0.5}$ for large N) is $\epsilon_N \sqrt{N}$. To calculate this, the highest number of iterations (22) is used to get c for each method and for this, by converting the number of steps M into number of samples N . Then this c is used to estimate the number of samples N_{th} needed for each sampling method to acquire an accuracy of $\epsilon = 10^{-6}$, by calculating it as $N_{th} = (c \times 10^6)^2$.

Table 1 shows that the number of samples required for the same accuracy is around 8.6 times higher for the uniform sampling than the importance sampling for the probability distribution function $0.98 - 0.48x$, which is close to 2^3 corresponding to three extra doublings of the sample size. For the second probability distribution, the sample size is 13.5 times lower than the uniform sampling, and around 1.6 times less than for the importance sampled method with the first probability distribution. In comparison to the other methods of integration in previous assignments, including Runge-Kutta 45 Integration (with adaptive step size), the trapezium rule and Simpson's rule, this Monte Carlo method requires a higher number of samples to be taken. These values are compared in Table 2, and shows that the number of samples required for this

	Uniform	Importance (pdf ₁)	Importance (pdf ₂)
$\text{erf}(x = 2)$	0.995311	0.995312	0.995317
$\sigma_m^{(2.58)}$ (99% conf. int.)	4.4×10^{-5}	1.5×10^{-5}	1.2×10^{-5}
Predicted steps for $\epsilon = 10^{-6}$	33	30	30
Corresponding samples	4.295×10^{12}	5.369×10^{11}	5.369×10^{11}
Value for c	2.015	0.687	0.550
Predicted min. samples for $\epsilon = 10^{-6}$	4.06×10^{12}	4.72×10^{11}	3.03×10^{11}

Table 1: Results from calculation of $\text{erf}(2)$ using both uniform sampling and importance sampling, for to different probability distribution functions.

	Samples N	Scaling of ϵ with N
Monte Carlo (Importance sampling pdf_1)	3.025×10^{11}	$\epsilon \sim N^{-1/2}$
Monte Carlo (Importance sampling pdf_2)	1.771×10^{11}	$\epsilon \sim N^{-1/2}$
Runge-Kutta 45 (Adaptive step size)	12	$\epsilon \sim N^{-6}$
Simpson's Rule (Fixed step)	65	$\epsilon \sim N^{-4}$
Trapezium Rule	513	$\epsilon \sim N^{-2}$

Table 2: Comparison of the number of samples required to obtain an accuracy of $\epsilon = 10^{-6}$ for the evaluation of $\text{erf}(2)$, additionally showing the expected scaling of ϵ with sample size.

accuracy is much higher for Monte Carlo than any others.

The values in Table 2 can be understood considering the scaling of the error with number of steps for the different methods. If the formula for error goes as $\epsilon = cN^{-n}$ where n is some number indicating scaling whilst c is a constant, and if it is assumed that $c \approx 0.5$ similar to the Monte-Carlo method, then:

$$\begin{aligned}
\text{Runge-Kutta 45 (Adaptive): } \epsilon &= cN^{-6} \implies N = (c/\epsilon)^{1/6} = (0.5 \times 10^6)^{1/6} = 8.91 \sim 10 \\
\text{Simpson's Rule (Fixed step): } \epsilon &= cN^{-4} \implies N = (c/\epsilon)^{1/4} = (0.5 \times 10^6)^{1/4} = 26.59 \sim 30 \\
\text{Trapezium: } \epsilon &= cN^{-2} \implies N = (c/\epsilon)^{1/2} = (0.5 \times 10^6)^{1/2} = 707.1 \sim 700
\end{aligned}$$

This shows that the order of the number of steps required for evaluating $\text{erf}(2)$ to an accuracy of $\epsilon = 10^{-6}$ calculated from the known scaling of the error with sample size, is of the same order as what has been found in the programs written for the corresponding assignments (Since the evaluation is very close to one, the error for the RK45 method is approximately equal to the relative error).

3 Question 3

The interpolating function used in assignment 2 for the data given, produces the image shown in Fig. 3. If the MC integration program was used to integrate this, then the a sampling probability distribution which may be useful for this is of a Gaussian distribution form. This would be centred on 0 and have a full-width-at-half-maximum of approximately 1.6 (to match the width of the function). This is also shown in Fig. 3, indicating the functional similarity and hence suitability to this calculation. There are several issues regarding this however, the first being that this Gaussian would now need to be normalised between the maximum and minimum independent variable values, requiring the integration of the error function. However, this can easily be approximated using confidence tables and used to renormalise the Gaussian. Another problem is that unless the user supplied a data set of random Gaussian distributed numbers, the user would have to supply the inverse cumulative distribution function to the method. Luckily, methods are available for both generating normally distributed random numbers and transforming uniform deviates into such. Primarily, the user can utilize a generator called the Box-Muller Transform in which, given 2 uniform

deviates x_1 and x_2 , two independent normally distributed random numbers z_1 and z_2 can be determined through the equations (Ref. Wolfram Mathworld: Box-Muller Transformation):

$$\begin{aligned} z_1 &= \sqrt{-2\ln x_1} \cos(2\pi x_2) & x_1 &= e^{(-z_1^2 - z_2^2)/2} \\ z_2 &= \sqrt{-2\ln x_2} \sin(2\pi x_1) & x_2 &= \frac{1}{2\pi} \tan^{-1} \left(\frac{z_2}{z_1} \right) \end{aligned}$$

The Jacobian used to transform between these two coordinates (which for random variables gives the pdf quotient), is the product of two *independent* random normally distributed variables. This would then allow for the interpolant function to be integrated for a faster convergence than if uniform deviates were used. Another method is to use an approximated form of the cumulative distribution function of the Gaussian distribution which is the error function. It is difficult to obtain the inverse of this, however a possibility is for the user to supply this approximation to the program as the inverse error function (Ref. S. Winitzki "A handy approximation of the error function", 2008):

$$\operatorname{erf}^{-1}(x) = \operatorname{sign}(x) \sqrt{\sqrt{\left(\frac{2}{a\pi} + \frac{\ln(1-x^2)}{2}\right)^2 - \frac{\ln(1-x^2)}{a}} - \left(\frac{2}{a\pi} + \frac{\ln(1-x^2)}{2}\right)}$$

where $a \approx 0.147$, and the approximation has a small error of 0.00012. Using the Box-Muller method to generate an array of vectors which contain Gaussian distributed numbers and supplying it to the *Mcint* class along with the renormalised probability distribution (turning off the transformation method), this cubic spline interpolant could be integrated using Monte-Carlo methods.

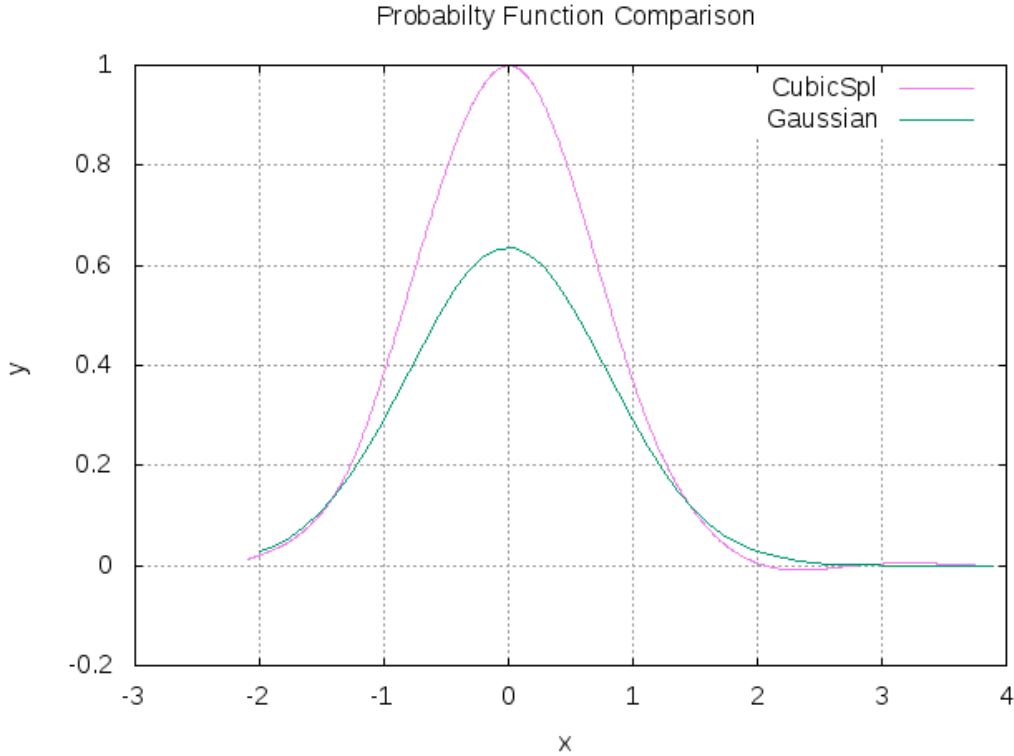


Figure 3: Comparison of the cubicspline interpolant function (red) and a Gaussian probability distribution function (green) with mean $\mu = 0$ and standard deviation $\sigma = 0.8$.