



230Q: Problem Set #3

Professor Johan Walden

Group 20

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Problem 1**Solution****Part A**

$$X(t) = e^{at}$$

$$dX = \alpha e^{\alpha t} dt$$

Part B

$$X(t) = \int_0^t g(s) dW(s)$$

$$dX = g(t) dW(t)$$

Part C

$$X(t) = e^{\alpha W(t)}$$

$$dX = \alpha e^{\alpha W(t)} dW + \frac{1}{2} \alpha^2 e^{\alpha W(t)} dt$$

Part D

$$X(t) = e^{\alpha Y(t)}$$

$$dX = \alpha e^{\alpha Y(t)} dY + \frac{1}{2} \alpha^2 e^{\alpha Y(t)} (dY)^2$$

$$= \alpha e^{\alpha Y(t)} (\mu dt + \sigma dW) + \frac{1}{2} \alpha^2 e^{\alpha Y(t)} \cdot \sigma^2 dt$$

$$= \alpha e^{\alpha Y(t)} \left(\mu + \frac{1}{2} \alpha \sigma^2 \right) dt + \alpha e^{\alpha Y(t)} \sigma dW$$

Part E

$$x(t) = y^2(t)$$

$$dx = 2y dy + (dy)^2$$

$$= 2y(\alpha y dt + \sigma y dW) + \sigma^2 y^2 dt$$

$$= y^2 (2\alpha + \sigma^2) dt + 2\sigma y^2 dW$$

Part F

$$x(t) = \frac{1}{y(t)}$$

$$dx = -\frac{1}{y^2} dy + 2\frac{1}{y^3} (dy)^2$$

$$= -\frac{1}{y^2} (\alpha y dt + \sigma y dW) + \frac{2}{y^3} \cdot \sigma^2 y^2 dt$$

$$= \frac{1}{y} (-\alpha + 2\sigma^2) dt - \frac{\sigma}{y} dW$$

Problem 2

Solution

First, we calculate the derivatives of X as follows

$$X(t) = \int_0^t \sigma_s dw_s \Rightarrow \begin{cases} dx = \sigma_t dw \\ (dx)^2 = \sigma_t^2 dt \end{cases}$$

Then, Let $Y(t) = e^{iuX(t)}$. We know $Y(0) = 1$.

Applying Ito lemma, we get

$$\begin{aligned} dY &= iue^{iuX(t)} dx + \frac{1}{2}(iu)^2 e^{iuX(t)} (dx)^2 \\ &= iue^{iuX(t)} \sigma_t dw - \frac{1}{2}u^2 e^{iuX(t)} \sigma_t^2 dt \end{aligned}$$

Take an integral on both sides

$$Y(t) - 1 = \int_0^t iue^{iuX(t)} \sigma_t dW - \frac{1}{2} \int_0^t u^2 e^{iuX(t)} \sigma_t^2 dt$$

Take the expectation of the formula. By the martingale properties, we know $\int_0^t iue^{iuX(t)} \sigma_t dw = 0$. So, we have

$$E[Y(t)] = 1 - \frac{1}{2}u^2 \int_0^t \sigma^2(s) E[e^{iuX(t)}] dt$$

Let $m_t = E[Y(t)] = E[e^{iuX(t)}]$. We know that $m(0) = 1$

Then, solve the ODE question as below

$$\begin{aligned} m_t &= 1 - \frac{1}{2}u^2 \int_0^t \sigma^2(s) m_s ds \\ m'_t &= -\frac{1}{2}u^2 \sigma^2(t) m_t \\ m_t &= e^{-\frac{1}{2}u^2 \int_0^t \sigma^2(s) ds} \end{aligned}$$

So, we have

$$E[e^{iuX(t)}] = e^{-\frac{u^2}{2} \int_0^t \sigma^2(s) ds}$$

Problem 3

Solution

$$\begin{aligned}
 dX &= \alpha X dt + \sigma_t dW \\
 \int_0^t dX &= \int_0^t \alpha X dt + \int_0^t \sigma_t dw \\
 E[X_t - X_0] &= E\left[\int_0^t \alpha X_s ds\right] + E\left[\int_0^t \sigma_s dW_s\right] \\
 E[X_t] &= E[X_0] + E\left[\int_0^t \alpha X_s ds\right] \\
 E[X_t] &= E[X_0] + \int_0^t \alpha E[X_s] ds
 \end{aligned}$$

Let $E[x_t] = m_t$, $m_0 = E[x_0]$. We solve the following ODE problem

$$\begin{aligned}
 m_t &= E[X_0] + \int_0^t \alpha m_s ds \\
 m'_t &= \alpha m_t \Rightarrow m_t = E[X_0] \cdot e^{\alpha t}
 \end{aligned}$$

To conclude, $E[X_t] = E[X_0] \cdot e^{\alpha t}$

Problem 4

Solution

By applying Taylor expansion, we have

$$\begin{aligned}dR &= 2XdX + 2YdY + \frac{1}{2} [2(dX)^2 + 2(dY)^2 + 0(dX)(dY)] \\&= 2XdX + 2YdY + (dX)^2 + (dY)^2 \\&= 2X(\alpha Xdt - YdW) + 2Y(\alpha Ydt + XdW) + Y^2dt + X^2dt \\&= (2\alpha X^2 + 2\alpha Y^2 + X^2 + Y^2) dt + (2XY - 2XY)dw \\&= (X^2 + Y^2) (2\alpha + 1)dt\end{aligned}$$

To conclude, we have $dR = (2\alpha + 1)Rdt$

Problem 5

Solution

Part A

First, calculate the derivatives of Y

$$Y = e^{-\alpha t} X_t, Y_t = -\alpha Y, Y_x = e^{-\alpha t}, Y_{xx} = 0$$

Then, apply Ito lemma

$$\begin{aligned} dY &= -\alpha Y dt + e^{-\alpha t} dX \\ &= -\alpha Y dt + e^{-\alpha t} (\alpha x dt + \sigma dw) \\ &= (-\alpha Y + \alpha X e^{-\alpha t}) dt + \sigma e^{-\alpha t} dW \\ &= \sigma e^{-\alpha t} dW \end{aligned}$$

Part B

$$\begin{aligned} \int_0^t dY &= \int_0^t \sigma e^{-\alpha s} dW_s \\ Y_t - Y_0 &= \int_0^t \sigma e^{-\alpha s} dW_s \\ e^{-\alpha t} X_t - X_0 &= \int_0^t \sigma e^{-\alpha s} dW_s \\ e^{-\alpha t} X_t &= X_0 + \int_0^t \sigma e^{-\alpha s} dW_s \\ X_t &= e^{\alpha t} x_0 + \sigma \int_0^t e^{\alpha(t-s)} dW_s \end{aligned}$$

Part C

$X(t) \mid \tilde{F}_0$ is normally distributed since each increment is independent and normally distributed. The increment comprises $\alpha X dt$, which is a constant. The increment also comprises σdW , which is independent normally distributed. So, $X(t) \mid \tilde{F}_0$ is normally distributed.

For the expected value,

$$\begin{aligned} E[x(T) \mid x(0)] &= E[e^{\alpha T} x_0] + E\left[\sigma \int_0^T e^{\alpha(t-s)} dW_s\right] \\ &= e^{\alpha T} x_0 + \sigma e^{\alpha T} E\left[\int_0^T e^{-\alpha s} dW_s\right] \\ &= e^{\alpha T} X_0 \end{aligned}$$

For the variance,

$$\begin{aligned}\text{Var}[X(T) \mid X(0)] &= \text{Var}\left(e^{\alpha T} X_0 + \sigma \int_0^T e^{\alpha(T-s)} dW_s\right) \\&= \sigma^2 \text{Var}\left(\int_0^T e^{\alpha(T-s)} dW_s\right) \\&= \sigma^2 \left(E \left[\left(\int_0^T e^{\alpha(T-s)} dW_s \right)^2 \right] - E \left[\int_0^T e^{\alpha(T-s)} dW_s \right]^2 \right) \\&= \sigma^2 \int_0^T E \left[e^{2\alpha(T-s)} \right] ds \\&= \sigma^2 \left(-\frac{1}{2\alpha} \right) \cdot e^{2\alpha(T-s)} \Big|_0^T \\&= -\frac{\sigma^2}{2\alpha} (1 - e^{2\alpha T}) \\&= \frac{\sigma^2}{2\alpha} (e^{2\alpha T} - 1)\end{aligned}$$