

# On the Estimation of Security Price Volatilities from Historical Data

Mark B. Garman and Michael J. Klass

The Journal of Business, 1980

Homework 4  
Matteo Zucca  
Martina Chiesa

# Topic:

Estimating capital asset price conditional volatility through public data

Data considered:

- opening prices (O)
- closing prices (C)
- high prices (H)
- low prices (L)

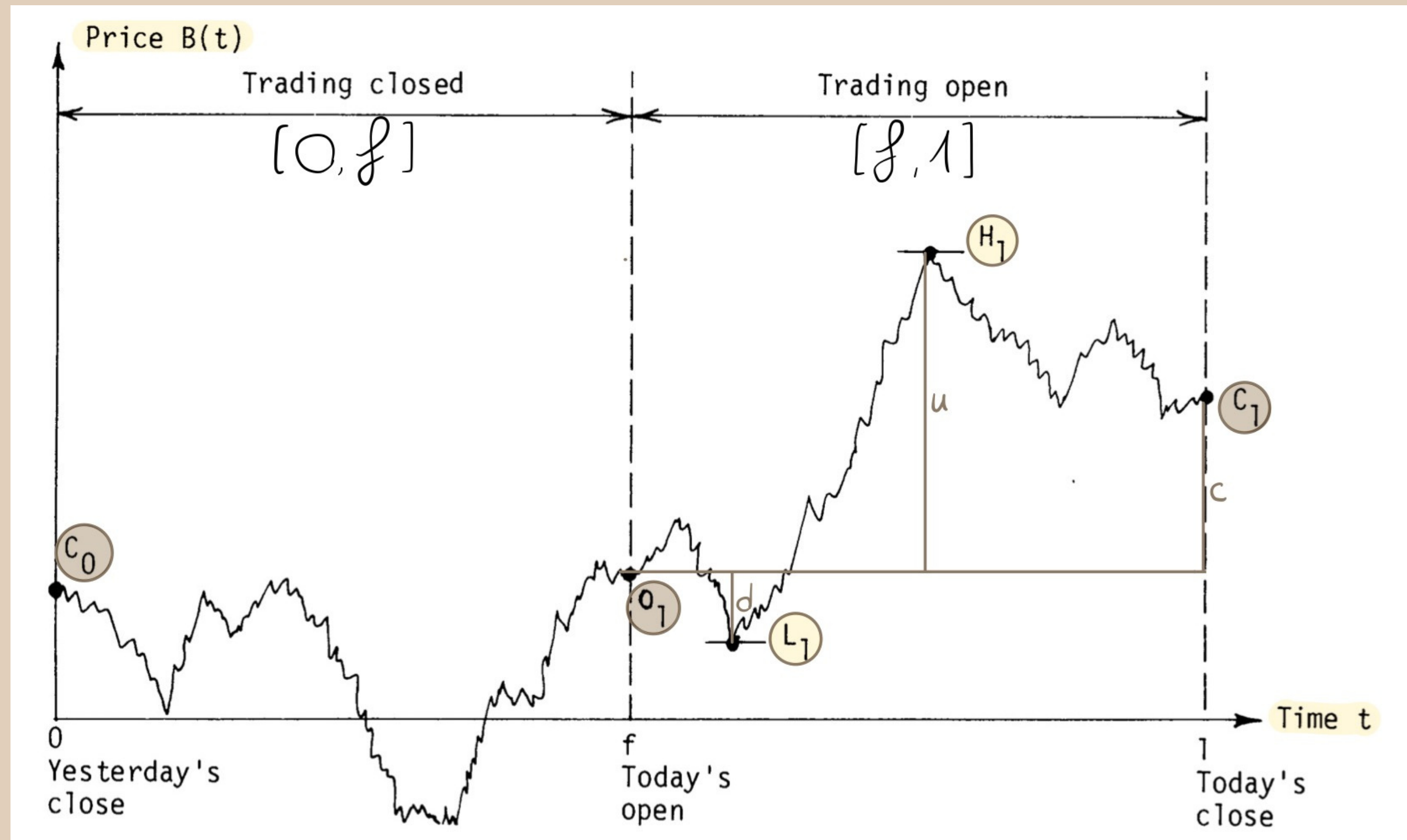
## Model assumption

Security prices are governed by a diffusion process with differential representation

$$P(t) = \Phi(B(t)) \quad dB = \sigma dz$$

Transformed price series  $B = \Phi^{-1}(P)$

# One trading day graph



Price versus Time

# Classical estimator

## Benchmark role

$$\hat{\sigma}_0^2 = (C_1 - C_0)^2 \quad f=0$$

### Advantages:

- simplicity of use
- unbiased estimation

### Disadvantages:

- ignores other information

### Relative Efficiency measure:

Consider the classical estimator as benchmark for other estimators

$$Eff(\hat{y}) = \frac{var(\hat{\sigma}_0^2)}{var(\hat{y})}$$

# Alternative estimators

## Opening & closing prices estimator

$$\hat{\sigma}_1^2 = \frac{(O_1 - C_0)^2}{2f} + \frac{(C_1 - O_1)^2}{2(1-f)} \quad 0 < f < 1$$

### Disadvantages

- It considers only opening and closing prices, which are snapshot, whereas low and high prices can provide more information on the volatility

# Alternative estimators

## High & Low prices estimators

### Disadvantages

- they ignore joint effects between  $u$ ,  $d$ ,  $c$  (normalized quantities of  $H$ ,  $L$  and  $C$  with respect to  $O$ )

$$\hat{\sigma}_2^2 = \frac{(H_1 - L_1)^2}{4 \ln 2} = \frac{(u-d)^2}{4 \ln 2} \quad f=0$$

$$\hat{\sigma}_3^2 = a \frac{(O_1 - C_0)^2}{f} + (1-a) \frac{(u-d)^2}{(1-f) 4 \ln 2} \quad 0 < f < 1$$

with  $a = 0.17$  minimum variance of the estimator

# Alternative estimators

## Unbiased analytic scale-invariant estimators

Aim: to find  $\hat{\sigma}^2 = D(u, d, c)$  minimum squared error estimator which is

- unbiased

$$E(D(u, d, c)) = \sigma^2$$

- analytic

$$D(u, d, c) = a_{200} u^2 + a_{020} d^2 + a_{002} c^2 + a_{110} ud + a_{101} uc + a_{011} dc$$

- scale invariant

$$D(\lambda u, \lambda d, \lambda c) = \lambda^2 D(u, d, c) \quad \lambda > 0$$



Suppose  $f=0$

$$D(u, d, c) = a_{200}u^2 + a_{020}d^2 + a_{002}c^2 + a_{110}ud + a_{101}uc + a_{011}dc =$$

$$\left. \begin{array}{l} a_{200} = a_{020} \\ a_{101} = a_{011} \end{array} \right] \text{ by symmetry property } B(t) \text{ and } -B(t) \text{ have same distribution}$$

$$2a_{200} + a_{110} + 2a_{101} = 0 \text{ ] by invariant property } B(t) \text{ and } B(1-t) - B(1) \text{ have same distribution}$$

$$= a_{200}(u^2 + d^2) + a_{002}c^2 - 2(a_{200} + a_{101})ud + a_{101}c(u + d) =$$

$$E(u^2) = E(d^2) = E(c^2) = E(c(u+d)) = \sigma^2$$

$$E(ud) = (1 - 2\ln 2)\sigma^2$$

$$= a_1(u-d)^2 + a_2(c(u+d) - 2ud) + (1 - (a_1 + a_2)4\ln 2 + a_2)c^2$$



$$V(a_1, a_2) = E \left[ \left( a_1 \underbrace{(u-d)}_X - (4 \ln 2) c^2 \right) + a_2 \underbrace{(c(u+d) - 2ud)}_Y + \underbrace{c^2}_Z \right)^2 \right]$$

is minimum for  $a_1, a_2$  which satisfy

$$E[(a_1 X + a_2 Y + Z)X] = E[(a_1 X + a_2 Y + Z)Y] = 0$$

$$a_1^* = \frac{E(XY)E(YZ) - E(Y^2)E(XZ)}{E(X^2)E(Y^2) - (E(XY))^2} = 0.511$$

$$a_2^* = \frac{E(XY)E(XZ) - E(Y^2)E(YZ)}{E(X^2)E(Y^2) - (E(XY))^2} = -0.019$$

Substituting  $a_1^*, a_2^*$  in  $D(u, d, c)$

$$D(u, d, c) = 0.511(u-d)^2 - 0.019(c(u+d) - 2ud) - 0.383c^2 = \hat{\sigma}_4^2$$

# Alternative estimators

## Unbiased analytic scale-invariant estimators

Eliminating cross product terms:

$$\hat{\sigma}_5^2 = 0.5(u-d)^2 - (2\ln 2 - 1)c^2$$

Suppose  $0 < f < 1$

$$\hat{\sigma}_6^2 = a \frac{(O_1 - C_0)^2}{f} + (1-a) \frac{\hat{\sigma}_4^2}{1-f}$$

with  $a = 0.12$  minimum variance of the estimator

# Summary estimators

estimator of $\sigma^2$	relative efficiency
$\hat{\sigma}_1^2 = \frac{(O_1 - C_0)^2}{2f} + \frac{(C_1 - O_1)^2}{2(1-f)}$	$\text{Eff}(\hat{\sigma}_1^2) = 2$
$\hat{\sigma}_2^2 = \frac{(H_1 - L_1)^2}{4 \ln 2} = \frac{(u-d)^2}{4 \ln 2}$	$\text{Eff}(\hat{\sigma}_2^2) \approx 5.2$
$\hat{\sigma}_3^2 = a \frac{(O_1 - C_0)^2}{f} + (1-a) \frac{(u-d)^2}{(1-f) 4 \ln 2}$	$\text{Eff}(\hat{\sigma}_3^2) \approx 6.2$
$\hat{\sigma}_4^2 = 0.511(u-d)^2 - 0.019(c(u+d) - 2ud) - 0.383c^2$	$\text{Eff}(\hat{\sigma}_4^2) \approx 7.4$
$\hat{\sigma}_5^2 = 0.5(u-d)^2 - (2 \ln 2 - 1)c^2$	$\text{Eff}(\hat{\sigma}_5^2) \approx 7.4$
$\hat{\sigma}_6^2 = a \frac{(O_1 - C_0)^2}{f} + (1-a) \frac{\hat{\sigma}_4^2}{1-f}$	$\text{Eff}(\hat{\sigma}_6^2) \approx 8.4$

# Conclusions:

Satisfying improvements in terms of relative efficiency

The new **unbiased analytic scale-invariant estimators** are more efficient than the classical ones: about 8 times better!

They exploit all information captured by the four data "available on newspaper's pages".

Together scale invariance and analyticity provides a *space reduction*: from an infinite dimensional problem to a six parameter formula for  $D(u,d,c)$  which is further simplified.



# Limitations

Unpredictable factors:

Exogenous shocks (e.g. oil crisis) influence stock prices

Predictable factors:

Using a continuous process on discrete time points provides biased statistics (finite transaction volume)

# Further developments

1

Consider more than one security (multivariate models, Sharpe 1970)

2

Seek for better models (since the one used is assumed to be good for price volatility)

3

Estimate additional quantities despite conditional volatility



# Possible application

