

# Instrumental Variables

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# Readings

- Angrist, J. and Krueger, A. (2001) *Instrumental Variables and the Search for Identification*, Journal of Economic Perspectives, Fall .
- Angrist, J and Pischke, J.S. (2009) *Mostly Harmless econometrics*, Princeton University Press.
- See also Steve Pischke's lectures on labor economics for research students – these notes are based on them...

# The problem

- The basic problem is to estimate a model with an endogenous variable. The classic example is earnings ( $y$ ) and education ( $S$ ) where ability ( $A$ ) is not observed
- Long form model:  $y_i = \alpha + \rho S + \gamma A + \epsilon_i$
- Estimate structural model:  $y_i = \alpha + \rho S + e_i$ , and so  $e_i = \gamma A + \epsilon_i$
- Note:  $\text{cov}(S, \epsilon) \neq 0$ . Estimation leads to OVB.
- IV solution is to find a variable  $Z_i$  such that  $\text{cov}(Z, \epsilon) = 0$  &  $\text{cov}(Z, S) \neq 0$

# Instrument validity

Conditions for a good instrument

1. Random assignment of  $Z_i$  (at least as good as)
2. Exclusion restriction:  $cov(Z, \epsilon) = 0$ . The instrument is not correlated with the disturbance in the long form model
3. Instrument validity:  $cov(Z, S) \neq 0$  The instrument is correlated with the endogenous variable.

Numbers 1 and 2 cannot be tested but argued from institutional context. 3 can be tested.

# Causal models

Three causal models:

1. The effect of  $Z_i$  on  $S_i$
2. The effect of  $Z_i$  on  $y_i$
3. The effect of  $S_i$  on  $y_i$

Ultimately, we're interested in model 3

Model estimation:

1. Effect of  $Z_i$  on  $S_i \Rightarrow$  randomization & instrument validity
2. Effect of  $Z_i$  on  $y_i \Rightarrow$  randomization & instrument validity
3. Effect of  $S_i$  on  $y_i \Rightarrow$  randomization, exclusion restriction & instrument validity

Again, model 3 is what we are interested in.

The following causal models can be estimated:

1. First stage:  $S_i = \pi_{10} + \pi_{11}Z_i + \psi_{1i}$
2. Reduced form:  $y_i = \pi_{20} + \pi_{21}Z_i + \psi_{1i}$
3. Structural equation:  $y_i = \alpha + \rho S_i + \epsilon_{1i}$

Interested in identifying the causal effect  $\rho$ .

# Implementation

Substitution of first stage into the structural model:

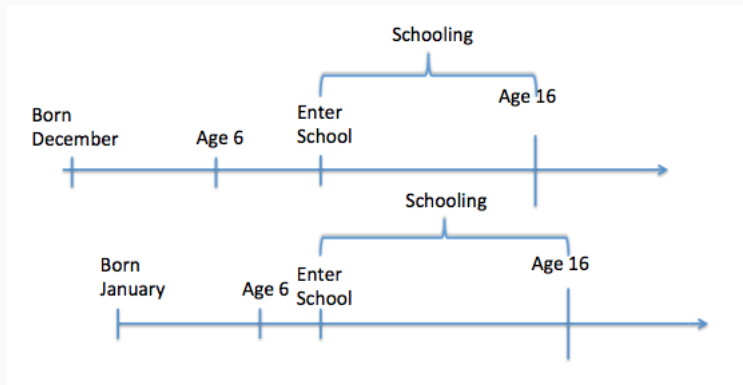
- $y_i = \alpha + \rho S_i + \epsilon_{1i}$
- $y_i = \alpha + \rho[\pi_{10} + \pi_{11}Z_i + \psi_{1i}] + \epsilon_{1i}$
- $y_i = \alpha + \rho\pi_{10} + \rho\pi_{11}Z_i + \rho\psi_{1i} + \epsilon_{1i}$
- $y_i = (\alpha + \rho\pi_{10}) + \rho\pi_{11}Z_i + (\rho\psi_{1i} + \epsilon_{1i})$
- $y_i = \pi_{20} + \pi_{21}Z_i + \epsilon_{2i}$
- where  $\pi_{20} = (\alpha + \rho\pi_{10})$  and  $\pi_{21} = \rho\pi_{11}$
- $\Rightarrow \frac{\pi_{21}}{\pi_{11}} = \rho$
- $\Rightarrow$  causal effect in the structural eq. ( $\rho$ ) is the ratio of the coefficient on the reduced form model divided by the coefficient from the first stage



## Example

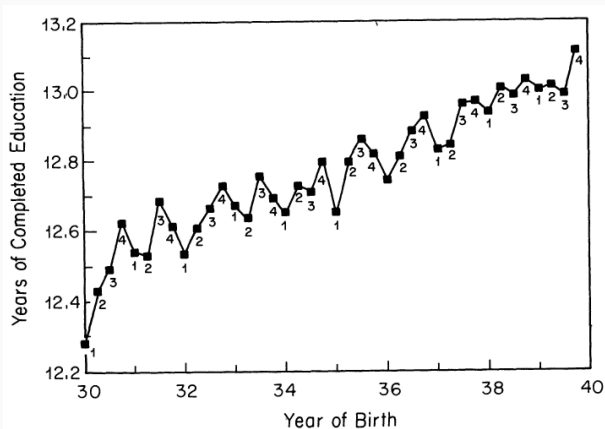
- Angrist & Krueger (1991) Does compulsory school attendance affect schooling and earnings. QJE.
- How does this work? School districts require children to have turned 6 years of age in the year that they enter the school system. Legal dropout rate is exactly 16.
- So, people born earlier in the year will be older when they enter school.
- They will also have fewer years of formal education
- Birth date therefore can be used as a legitimate instrument for years of schooling which is uncorrelated to earning since it is random....

# Schooling and birth-quarter



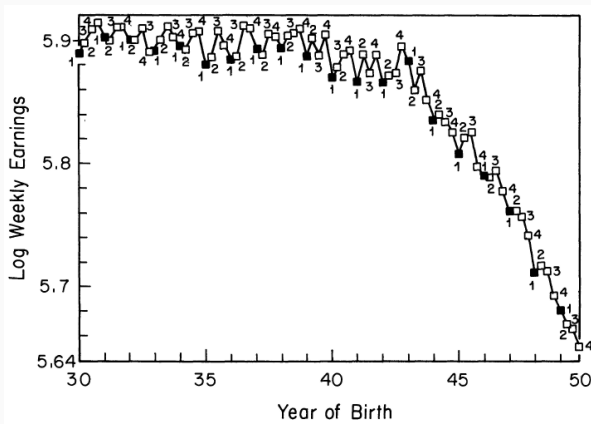
*Source: from J. Angrist data*

## Completed schooling and birth-quarter



Source: from J. Angrist data

# Earnings and birth-quarter



Source: from J. Angrist data

- Angrist & Krueger (1991) use 1980 data from the US census.
- The data set contains information on 329,509 men born between 1930 and 1939.
- The men are in the 40s when the sample is taken
- The data set contains information on 1979 earnings (logged), birth quarter, years of schooling
- Data available from Josh Angrist's data archive....

## Birth quarter statistics

| birth_quarter | Freq.   | Percent | Cum.   |
|---------------|---------|---------|--------|
| 1             | 81,671  | 24.79   | 24.79  |
| 2             | 80,138  | 24.32   | 49.11  |
| 3             | 86,856  | 26.36   | 75.47  |
| 4             | 80,844  | 24.53   | 100.00 |
| Total         | 329,509 | 100.00  |        |

*Source: from J. Angrist data*

## Schooling and birth quarter

| Variables    | (1)<br>school        | (2)<br>school        |
|--------------|----------------------|----------------------|
| q2           | 0.057***<br>(0.016)  |                      |
| q3           | 0.117***<br>(0.016)  |                      |
| q4           | 0.151***<br>(0.016)  | 0.092***<br>(0.013)  |
| Constant     | 12.688***<br>(0.011) | 12.747***<br>(0.007) |
| Observations | 329,509              | 329,509              |
| R-squared    | 0.000                | 0.000                |

Standard errors in parentheses

## Earnings and schooling

| Variables    | (1)<br>lwage (OLS)  | (2)<br>lwage (IV)   | (3)<br>lwage (IV)   |
|--------------|---------------------|---------------------|---------------------|
| school       | 0.071***<br>(0.000) | 0.074***<br>(0.028) | 0.103***<br>(0.020) |
| Constant     | 4.995***<br>(0.004) | 4.955***<br>(0.358) | 4.590***<br>(0.249) |
| Observations | 329,509             | 329,509             | 329,509             |
| R-squared    | 0.117               | 0.117               | 0.094               |

Standard errors in parentheses

\*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Instruments: Col 1: OLS; Col 2: IV using Q4 birth; Col 3: IV using Q3 to Q4



## Wald estimator

Consider the model:  $y = \beta_0 + \beta_1 x + u$  and let  $z$  be a binary ( $z = 0$ , or  $z = 1$ ) instrumental variable for  $x$ . Then the IV estimator is:

$$\hat{\beta}_1 = \hat{\beta}_{1IV} = \frac{\sum_{i=1}^N (z_i - \bar{z})(y_i - \bar{y})}{\sum_{i=1}^N (z_i - \bar{z})(x_i - \bar{x})},$$

The IV estimator  $\hat{\beta}_1$  can also be written as:

$$\hat{\beta}_1 = (\bar{y}_1 - \bar{y}_0)/(\bar{x}_1 - \bar{x}_0),$$

where  $\bar{y}_0$  and  $\bar{x}_0$  are the sample averages of  $y_i$  over the part of the sample with  $z_i = 0$ , and the terms  $\bar{y}_1$  and  $\bar{x}_1$  are the sample averages of  $y_i$  over the part of the sample with  $z_i = 1$ . This is known as the Wald (1940) estimator.

## Wald estimator: example

We can use the Angrist data to calculate the Wald estimator. Let  $y_i$  be log wages, and  $S_i$  years of schooling, and Q4 quarter 4 birth

$$\beta_{Wald} = \frac{Cov(y_i, Q4)}{Cov(S_i, Q4)}$$

$$\beta_{Wald} = \frac{E[y_i|Q4 = 1] - E[y_i|Q4 = 0]}{E[S_i|Q4 = 1] - E[S_i|Q4 = 0]}$$

## Wald estimates

Calculate the Wald estimator from the first stage and reduced form regressions.

$$S_i = \alpha_{10} + \alpha_{14}Q4_i + \theta_{1i} \Rightarrow \text{First Stage regression}$$

$$y_i = \alpha_{20} + \alpha_{24}Q4_i + \theta_{2i} \Rightarrow \text{Reduced form regression}$$

$$E[y_i|Q4 = 1] = \alpha_{20} + \alpha_{24} \text{ and}$$

$$E[y_i|Q4 = 0] = \alpha_{20}$$

$$\Rightarrow E[y_i|Q4 = 1] - E[y_i|Q4 = 0] = \alpha_{24} \text{ and}$$

## Wald estimates

- The Wald estimate is the ratio of these two:

$$\beta_{Wald} = \frac{\alpha_{24}}{\alpha_{14}}$$

- In the case of earnings: the Wald estimator is the difference in average earnings across the two groups divided by the difference in average schooling across the two groups

## Reduced form

```
. reg lwage q4 // regress wage on Q4
```

| Source   | SS               | df | MS         |
|----------|------------------|----|------------|
| Model    | 2.83199415       | 1  | 2.83199415 |
| Residual | 151835.039329507 |    | .460794578 |
| Total    | 151837.871329508 |    | .460801774 |

Number of obs = 329509

F( 1,329507) = 6.15

Prob > F = 0.0132

R-squared = 0.0000

Adj R-squared = 0.0000

Root MSE = .67882

| lwage | Coef.    | Std. Err. | t       | P> t  | [95% Conf. Interval] |          |
|-------|----------|-----------|---------|-------|----------------------|----------|
| q4    | .0068132 | .0027482  | 2.48    | 0.013 | .0014267             | .0121996 |
| _cons | 5.898272 | .0013613  | 4332.90 | 0.000 | 5.895604             | 5.90094  |

*Source: from J. Angrist data*

## First stage

```
. reg school q4 // regress school on Q4
```

| Source   | SS               | df         | MS       | Number of obs = 329509 |  |  |
|----------|------------------|------------|----------|------------------------|--|--|
| Model    | 517.7393         | 1          | 517.7393 | F( 1,329507) = 48.09   |  |  |
| Residual | 3547149.92329507 | 10.7650215 |          | Prob > F = 0.0000      |  |  |
| Total    | 3547667.66329508 | 10.76656   |          | R-squared = 0.0001     |  |  |
|          |                  |            |          | Adj R-squared = 0.0001 |  |  |
|          |                  |            |          | Root MSE = 3.281       |  |  |

| school | Coef.    | Std. Err. | t       | P> t  | [95% Conf. Interval] |          |
|--------|----------|-----------|---------|-------|----------------------|----------|
| q4     | .0921209 | .0132834  | 6.94    | 0.000 | .0660857             | .118156  |
| _cons  | 12.74731 | .0065796  | 1937.40 | 0.000 | 12.73441             | 12.76021 |

*Source: from J. Angrist data*

## Wald Estimate

- Difference in mean earnings = 0.0068
- Difference in mean schooling = 0.0921
- Ratio of the difference =  $0.0068 / 0.0921 = 0.074$
- Compare this quantity to the IV estimate arising from the 2SLS method – they are the same...

## IV estimates

```
. ivregress 2sls lwage (school= q4)
```

Instrumental variables (2SLS) regression

Number of obs = 329509

Wald chi2(1) = 6.96

Prob > chi2 = 0.0083

R-squared = 0.1171

Root MSE = .63785

| lwage  | Coef.    | Std. Err. | z     | P> z  | [95% Conf. Interval] |          |
|--------|----------|-----------|-------|-------|----------------------|----------|
| school | .0739589 | .0280328  | 2.64  | 0.008 | .0190157             | .1289021 |
| _cons  | 4.955495 | .3579775  | 13.84 | 0.000 | 4.253872             | 5.657118 |

Instrumented: school

Instruments: q4

*Source: from J. Angrist data*