



Figure 1: Plot of  $F(\alpha, t)$  in q3 as a function of  $\alpha$  (for  $t = 1$  fixed).

## Homework 5

1. Plot the value of a European call option as a function of the initial stock price at  $t = 0$  for two different maturities  $T_1$  and  $T_2$  (assume that  $0 < T_1 < T_2$  and the interest rate  $r = 0$ ). Derive the asymptotic behaviour of the call price as  $S_0 \rightarrow \infty$ .

**Solution.** The unique no-arbitrage call option price  $C(S_0, 0)$  at  $t = 0$  satisfies

$$\frac{C(S_0, 0)}{S_0} = \frac{1}{S_0} \mathbb{E}^{\mathbb{Q}}((S_T - K)^+) = \frac{1}{S_0} \mathbb{E}^{\mathbb{Q}}((S_0 e^{\sigma W_T - \frac{1}{2} \sigma^2 T} - K)^+) = \mathbb{E}^{\mathbb{Q}}((e^{\sigma W_T - \frac{1}{2} \sigma^2 T} - \frac{K}{S_0})^+).$$

This tends to 1 as  $S_0 \rightarrow \infty$  by the dominated convergence theorem, since the expression inside the expectation is less than or equal to  $\mathcal{E} = e^{\sigma W_T - \frac{1}{2} \sigma^2 T}$  (and tends monotonically to  $\mathcal{E}$  as  $S_0 \rightarrow \infty$  a.s. since  $\frac{K}{S_0} \rightarrow 0$ ), and  $\mathbb{E}^{\mathbb{Q}}(\mathcal{E}) = 1 < \infty$ . Hence we have shown that  $C(S_0, 0) \sim S_0$  as  $S_0 \rightarrow \infty$ .

2. Consider a portfolio of buying  $\frac{1}{\Delta}$  call options with strike  $K$ , and selling  $\frac{1}{\Delta}$  call options of strike  $K + \Delta$  for  $\Delta > 0$ . Plot the terminal payoff function  $\tilde{f}(S)$  of this portfolio as a function of  $S$ . What does this payoff tend to as  $\Delta \rightarrow 0$ ?

**Solution.** The payoff function  $f(S) = 0$  for  $S \leq K$ ,  $f(S) = \frac{1}{\Delta}(S - K - (S - K - \Delta)) = 1$  for  $S \geq K + \Delta$ , and  $f(S) = \frac{1}{\Delta}(S - K)$  for  $S \in [K, K + \Delta]$ , and in particular we see that  $f(\cdot)$  is continuous at  $S = K$  and  $K + \Delta$  (This is known as a **call spread** strategy). From a picture, we see that this tends to a digital call payoff  $1_{S_T \geq K}$  as  $\Delta \rightarrow 0$ .

3. A **symmetric  $\alpha$ -stable process**  $X$  with parameters  $\alpha \in (0, 2]$ ,  $\sigma > 0$  is a generalization of Brownian motion, which has independent stationary increments like Brownian motion but now  $\mathbb{E}(e^{iu(X_t - X_s)} | X_s) = e^{-(t-s)\sigma^\alpha |u|^\alpha}$  for  $u \in \mathbb{R}$  and  $0 \leq s \leq t$ , so  $X$  is only a (multiple of) BM if  $\alpha = 2$ , but for  $\alpha < 2$  the increments of  $X$  are not normally distributed. If  $\bar{X}_t = \max_{0 \leq s \leq t} X_s$  and  $\underline{X}_t = \min_{0 \leq s \leq t} X_s$ , it can be shown that

$$\mathbb{E}(\bar{X}_t - \underline{X}_t) = F(\alpha, t) := \frac{2\alpha\Gamma(1 - \frac{1}{\alpha})}{\pi} t^{\frac{1}{\alpha}}$$

for  $\alpha \in (1, 2]$ . Use this identity to define a statistical estimator  $\hat{\alpha}$  for  $\alpha$  (assuming  $\sigma = 1$ ) from observed values for  $\bar{X}_t$  and  $\underline{X}_t$ . Is  $\hat{\alpha}$  biased? (you may use that  $\frac{\partial^2}{\partial \alpha^2} F(\alpha, t) > 0$ ).

**Solution.** We just solve  $\bar{X}_t - \underline{X}_t = F(\hat{\alpha}, t)$  for  $\hat{\alpha}$  to get  $\hat{\alpha}$  (see plot of  $F(\alpha, t)$  above for  $t = 1$ ). For the 2nd part, we see that

$$F(\alpha, t) = \mathbb{E}(\bar{X}_t - \underline{X}_t) = \mathbb{E}(F(\hat{\alpha}, t)) \geq F(\mathbb{E}(\hat{\alpha}), t)$$

where the final inequality follows from **Jensen's inequality** from last lecture. Assuming the  $\geq$  is actually a  $>$  here, we can apply  $F^{-1}$  to both sides and (since  $F^{-1}$  is decreasing), we see that  $\alpha < \mathbb{E}(\hat{\alpha})$ , so  $\hat{\alpha}$  is biased.

4. Let  $M_t^{(n)} = \max_{0 \leq k \leq n} W_{kt/n}$  denote the **discretely sampled** maximum of  $W$ . Is  $M_t^{(n)}$  more or less than  $M_t = \max_{0 \leq s \leq t} W_s$ ? It is known that

$$\mathbb{E}(M_t^{(n)}) = \sqrt{\frac{t}{2\pi n}} H_n^{(\frac{1}{2})}$$

where  $H_n^{(\frac{1}{2})} = \sum_{k=1}^n k^{-\frac{1}{2}}$  is known as the  $n$ th Harmonic number of order  $\frac{1}{2}$ . Using that  $H_n^{(\frac{1}{2})} \sim 2\sqrt{n}$  as  $n \rightarrow \infty$ , show this formula is consistent with the formula  $\mathbb{E}(R_t) = 2\sqrt{\frac{2t}{\pi}}$  in Hwk4, q5, where  $R_t$  is the range of  $W$ .

**Solution.**  $M_t^{(n)} \leq M_t$ . Using the given asymptotic relation, we see that

$$\mathbb{E}(M_t^{(n)}) \sim \sqrt{\frac{t}{2\pi n}} 2\sqrt{n} = \sqrt{\frac{2t}{\pi}}$$

and we see this is half  $\mathbb{E}(R_t)$  as expected, since  $\mathbb{E}(R_t) = \mathbb{E}(M_t) - \mathbb{E}(m_t) = 2\mathbb{E}(M_t)$  since  $m_t \sim -M_t$ , where  $m_t = \min_{0 \leq s \leq t} W_s$ .

5. Consider the famous **GARCH(1,1)** discrete-time stochastic volatility model where

$$\begin{aligned} r_t &= \sqrt{V_t} \varepsilon_t \\ V_t &= \omega + \alpha r_{t-1}^2 + \beta V_{t-1} \end{aligned}$$

for  $t \in \mathbb{Z}$  and  $\omega, \alpha, \beta > 0$ , where  $r_t = (S_t - S_{t-1})/S_{t-1}$  is the return on a stock price process  $S$  at time  $t$ , and  $\varepsilon_t$  is an i.i.d. sequence of random variables with  $\mathbb{E}(\varepsilon_t) = 0$  and  $\text{Var}(\varepsilon_t) = 1$ . Express  $V_t$  in terms of the  $\varepsilon_t$  sequence only, and compute  $\mathbb{E}(V_t)$  and give a condition on the model parameters for the formula to make sense.

**Solution.** We first note that

$$V_t = \omega + \alpha r_{t-1}^2 + \beta V_{t-1} = \omega + \alpha V_{t-1} \varepsilon_{t-1}^2 + \beta V_{t-1} = \omega + A_t V_{t-1} \quad (1)$$

where  $A_t = \alpha \varepsilon_{t-1}^2 + \beta$ . “Unrolling” this expression, we see that

$$\begin{aligned} V_t &= \omega + A_t(\omega + A_{t-1}V_{t-2}) = \omega + A_t(\omega + A_{t-1}(\omega + A_{t-2}V_{t-3})) \\ &= \omega + A_t\omega + A_tA_{t-1}\omega + A_tA_{t-1}A_{t-2}V_{t-3} \\ &= \omega(1 + A_t + A_tA_{t-1} + A_tA_{t-1}A_{t-2} + \dots). \end{aligned}$$

We also note that  $\mathbb{E}(A_t) = \alpha + \beta$  for all  $t$ , so

$$\mathbb{E}(V_t) = \omega(1 + (\alpha + \beta) + (\alpha + \beta)^2 + \dots) = \frac{\omega}{1 - \alpha - \beta}$$

(using the formula for the sum of a **Geometric series**), so the model only makes sense when  $\alpha + \beta < 1$ , which is known as the **stationarity condition** which ensures that  $V_t$  has the same distribution for all  $t$  (see notes on my webpage for more details on the model, many students have looked at this model in their projects and we often use the student  $t$ -distribution for the  $\varepsilon_t$ 's).