

## Homework 2

1. Let  $X$  be a Lévy process, and  $e_q$  an  $\text{Exp}(q)$  random variable independent of  $X$ , and let  $\rho_t(x)$  denote the density of  $X_t$ . Using that

$$\Phi_q^+(z) := \mathbb{E}(e^{iz\bar{X}_{e_q}}) = \exp\left(\int_0^\infty t^{-1}e^{-qt} \int_0^\infty (e^{izx} - 1)\rho_t(x)dx dt\right) \quad (1)$$

for  $z \in \mathbb{R}$ , compute  $\mathbb{E}(\bar{X}_t)$  (this question requires no knowledge of Lecture 3).

**Solution.**

$$(\Phi_q^+)'(0) = i\mathbb{E}(\bar{X}_{e_q}) = i \int_0^\infty e^{-qt} \frac{\mathbb{E}(X_t^+)}{t} dt.$$

But we also know that  $i\mathbb{E}(\bar{X}_{e_q}) = i \int_0^\infty qe^{-qt}\mathbb{E}(\bar{X}_t)dt$ , so (dividing both expressions by  $iq$ ) we see that

$$\int_0^\infty e^{-qt}\mathbb{E}(\bar{X}_t)dt = \frac{1}{q} \int_0^\infty e^{-qt} \frac{\mathbb{E}(X_t^+)}{t} dt.$$

But by a standard simple property of Laplace transforms,  $\frac{1}{q}\mathcal{L}f = \mathcal{L}F$ , where  $\mathcal{L}$  denotes the Laplace transform operator and  $F(t) = \int_0^t f(s)ds$  (see e.g. [https://en.wikipedia.org/wiki/Laplace\\_transform#Properties\\_and\\_theorems](https://en.wikipedia.org/wiki/Laplace_transform#Properties_and_theorems)). Hence (comparing both sides) we see that

$$\mathbb{E}(\bar{X}_t) = \int_0^t \frac{\mathbb{E}(X_s^+)}{s} ds$$

(we have seen this formula before in FM02 without proof).

2. The general **Wiener-Hopf** formula for a Lévy process states that

$$\mathbb{E}(e^{izX_{e_q}}) = \mathbb{E}(e^{iz\bar{X}_{e_q}}) \mathbb{E}(e^{izX_{e_q}})$$

for  $z \in \mathbb{R}$ . Using (1), show that  $\bar{X}_{e_q}$  has a **Lévy-Khintchine** representation of the form  $\log \Phi_q^+(z) = \int_0^\infty (e^{izx} - 1)\nu(x)dx$  for some non-negative function  $\nu(x)$  (recall we refer to  $\nu(x)$  as the Lévy density).

**Solution.** Interchanging integrals in (1), we see that

$$\log \Phi_q^+(z) := \log \mathbb{E}(e^{iz\bar{X}_{e_q}}) = \int_0^\infty (e^{izx} - 1) \int_0^\infty t^{-1}e^{-qt}\rho_t(x)dt dx = \int_0^\infty (e^{izx} - 1)\nu(x)dx$$

where  $\nu(x) = \int_0^\infty t^{-1}e^{-qt}\rho_t(x)dt$ .

**Remark 0.1** If  $X$  is standard Brownian motion, using the known density  $\rho_t(x) = \frac{1}{\sqrt{2\pi t}}e^{-\frac{1}{2}x^2/t}$  of  $W_t$ , we can easily check that  $\nu(x) = \frac{e^{-x\sqrt{2q}}}{x}1_{x>0}$ , and we have seen in the FM02 Mock that  $\bar{X}_{e_q} \sim \text{Exp}(\sqrt{2q})$ . A Lévy process with this  $\nu(x)$  function is known as a **Gamma process**, which is essentially a **one-sided CGMY process** with  $Y = 0$ .

3. Let  $X$  be a **Cauchy process** (i.e. an  $\alpha$ -stable process with  $\alpha = 1$  so  $\nu(x) = \frac{1}{\pi x^2}$ ) for which the density of  $X_t$  given  $X_s = x$  (for  $0 \leq s \leq t$ ) is a Cauchy distribution with density

$$p(x, y; \tau) = \frac{1}{\pi} \frac{\tau}{\tau + (y - x)^2}$$

where  $\tau = t - s$ . Compute the density of  $\hat{X}_t$  given  $\hat{X}_s = x$  (for  $0 \leq s \leq t < 1$ ), where  $\hat{X}$  is a Cauchy **bridge process**, i.e.  $X$  conditioned to be 0 at time 1.

**Solution.** Using Bayes' formula, the conditional density is

$$\frac{p(x, y; t - s)p(y, 0; 1 - t)}{p(x, 0; 1 - s)}$$

which evaluates to

$$\frac{(1 - t)(t - s)((1 - s)^2 + x^2)}{\pi(1 - s)((s - t)^2 + (x - y)^2)((1 - t)^2 + y^2)}.$$

This can be used to simulate the bridge process using the usual  $F^{-1}(U)$  method.

4. Let  $X_t = \sigma B_t^H$  where  $B^H$  is fBM, and let  $\theta_{m,k} = -2^{\frac{m}{2}} (X_{\frac{2(k+1)}{2^{m+1}}} - 2X_{\frac{2k+1}{2^{m+1}}} + X_{\frac{2k}{2^{m+1}}})$  and  $s_n^2 = \sum_{m=0}^{n-1} \sum_{k=0}^{2^m-1} \theta_{m,k}^2$ . Using that

$$\mathbb{E}(s_n^2) = \sigma^2(4^{n(1-H)} - 1) \sim \sigma^2 4^{n(1-H)} = \sigma^2 2^{2n(1-H)}$$

and  $s_n^2 \sim \sigma^2 2^{2n(1-H)}$  as  $n \rightarrow \infty$  (i.e. the ratio of both sides tends to 1), derive an estimator  $\hat{H}_n$  for  $H$  using a ratio of  $s_n$  terms at different resolutions. Is  $\hat{H}$  scale-invariant? You may assume that  $\mathbb{E}(\theta_{m,k}^2)$  is independent of  $k$ .

**Solution.**  $\frac{s_n^2}{2^{2n(1-H)}} \rightarrow \sigma^2$  as  $n \rightarrow \infty$ , so  $\frac{s_{n-1}^2}{2^{2(n-1)(1-H)}} \rightarrow \sigma^2$  and the ratio of these two terms satisfies

$$\frac{s_n^2}{s_{n-1}^2} \cdot 2^{-2(1-H)} \rightarrow 1.$$

Setting  $\frac{s_n^2}{s_{n-1}^2} = 2^{2(1-\hat{H}_n)}$ , we see that  $\frac{2^{2(1-\hat{H}_n)}}{2^{2(1-H)}} \rightarrow 1$ , so  $2(1-\hat{H}_n) - 2(1-H) = 2(H-\hat{H}_n) \rightarrow 0$ , i.e.  $\hat{H}_n$  is a consistent estimator, and we can write  $\hat{H}_n$  explicitly as

$$\hat{H}_n = 1 - \frac{1}{2} \log_2 \frac{s_n^2}{s_{n-1}^2}.$$

$\hat{H}_n$  is scale-invariant since if we multiply the path  $X$  by  $\lambda$ , the ratio  $\frac{s_n^2}{s_{n-1}^2}$  remains un-changed because the  $\lambda$  terms cancel.