
STATISTICAL ANALYSIS PLAN

ADVANCE TRAUMA

**Effects of Advanced Trauma Life Support®
Training Compared to Standard Care on
Adult Trauma Patient Outcomes: A
Cluster Randomised Trial**

Version 0.3.0, 2024-09-18

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1 Administrative information

1.1 Study identifiers

- Protocol version 1.2.0 dated 2024-08-26
- ClinicalTrials.gov ID NCT06321419
- Clinical Trial Registry - India

1.2 Changelog

Once version 1.0.0 is finalised, this section will be updated with a changelog.

1.3 Contributors

2 Trial synopsis

Title Effects of Advanced Trauma Life Support® Training Compared to Standard Care on Adult Trauma Patient Outcomes: A Cluster Randomised Trial

Rationale Trauma is a massive global health issue. Many training programmes have been developed to help physicians in the initial management of trauma patients. Among these programmes, Advanced Trauma Life Support® (ATLS®) is the most popular, having trained over one million physicians worldwide. Despite its widespread use, there are no controlled trials showing that ATLS® improves patient outcomes. Multiple systematic reviews emphasise the need for such trials.

Aim To compare the effects of ATLS® training with standard care on outcomes in adult trauma patients.

Primary Outcome In-hospital mortality within 30 days of arrival at the emergency department.

Trial Design Batched stepped-wedge cluster randomised trial in India.

Trial Population Adult trauma patients presenting to the emergency department of a participating hospital.

Sample Size 30 clusters and 4320 patients.

Eligibility Criteria

Hospitals are secondary or tertiary hospitals in India that admit or refer/transfer for admission at least 400 patients with trauma per year.

Clusters are one or more units of physicians providing initial trauma care in the emergency department of tertiary hospitals in India.

Patients participants are adult trauma patients who presents to the emergency department of participating hospitals and are admitted or transferred for admission.

Intervention The intervention will be ATLS® training, a proprietary 2.5 day course teaching a standardised approach to trauma patient care using the concepts of a primary and secondary survey. Physicians will be trained in an accredited ATLS® training facility in India.

Ethical Considerations We will use an opt-out consent approach for collection of routinely recorded data. We will obtain informed consent for collection of non-routinely recorded data, such as quality of life and disability outcomes. Patients who are unconscious or lack a legally authorized representative will be included under a waiver of informed consent. Note that consent here refers to consent to data collection.

Trial Period October 1, 2024, to September 30, 2029

3 Special considerations

3.1 Funding

This trial is not yet fully funded. The Trial Management Group has decided to proceed with the trial with the expectation that additional funding will be secured. The Trial Steering Committee will be informed of the funding status at each meeting. If funding is not secured, the trial will be stopped. This will likely result in an underpowered trial. The justification for this decision is that the intervention is considered standard of care in many countries and the data collection is considered minimal risk. There is therefore a very small risk of harm to patient participants, but a potential direct benefit to those patient participants who receive the intervention. The benefit-risk ratio is therefore considered to be favourable, even in the case of an underpowered trial.

3.2 Potential amendments

There are ongoing discussions about re-framing the trial as a hybrid effectiveness-implementation trial and include a cost-effectiveness analysis. This would involve adding additional data collection to assess the implementation and costs of the intervention. This would involve additional funding and amended ethical approvals.

4 Statistical analysis

5 Aim and objectives

5.1 Aim

To compare the effects of ATLS® training with standard care on outcomes in adult trauma patients.

5.2 Objectives

- To estimate the effects of ATLS® training on 30-day in-hospital mortality compared to standard care.
- To assess the robustness of the main analysis results to different model specifications through sensitivity analyses.
- To estimate the effects of ATLS® training on secondary outcomes compared to standard care.
- To explore the effects of ATLS® training on subgroups of patients compared to standard care.

5.3 Design

This is a batched stepped-wedge cluster randomised trial, including a total of 30 (see [trial-design?](#)). The trial will be composed of 6 batches of identical 12-period 5-sequence design, with one cluster being assigned to each sequence of each batch¹. Each period is one month, and each cluster will be in the trial for a total of 13 months. The intervention will be implemented during a one-month transition period, which will be excluded from the analysis. There will be an overlap of 6 months between successive batches.

5.4 Randomisation of clusters

We will assign clusters to batches as they are found to be eligible and receive ethical approval. Batches will include clusters from hospitals in different regions to optimize trial logistics. We will randomise the clusters allotted to each batch to the different intervention implementation sequences within that batch¹. We will balance the randomisation within each batch on cluster size, defined as monthly volume of eligible patient participants, using covariate constrained randomisation. The cluster sizes are expected to

¹**Question:** Karla and James, can you please provide more details on how the randomisation will be done?

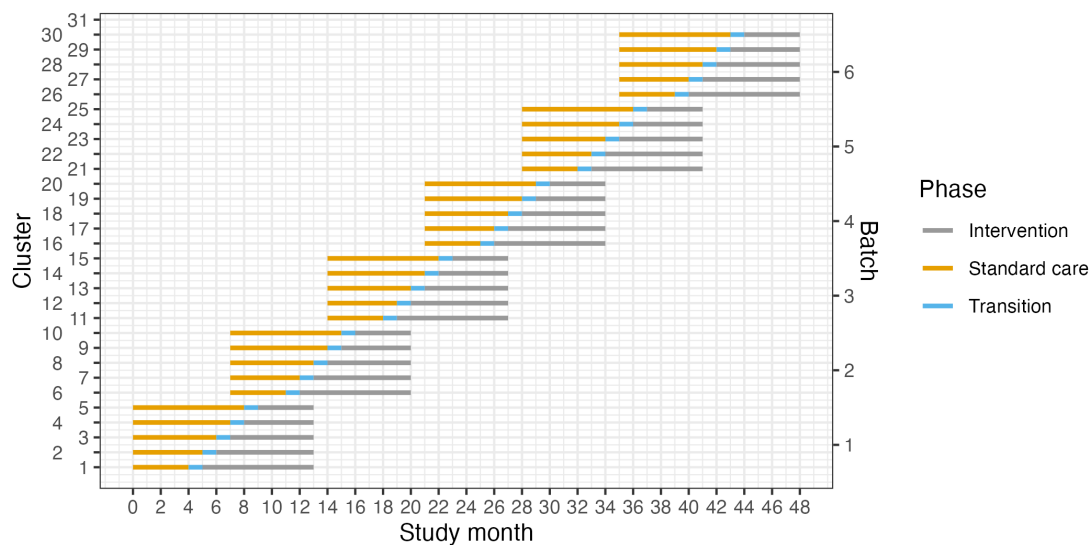


Figure 1: Trial design. Lines represent the duration of patient enrolment across clusters and phases. Clusters will be sequentially allocated to a batch based on when they enter the study. Within each batch clusters will then be randomised to an intervention implementation sequence.

vary between 12 and 20 patients per month, based on our previous experiences. We will conceal the randomisation order for as long as it is logistically possible, considering that arrangements for sending physicians to ATLS® training need to be made in advance.

5.5 Outcomes

5.5.1 Primary outcome

The primary outcome will be in-hospital mortality within 30 days of arrival at the emergency department. Clinical research coordinators will extract information on death from patient hospital records. If the patient has been transferred to another hospital, the clinical research coordinators will collect data on this outcome by calling the patient or a patient representative, or by contacting the hospital to which the patient was transferred. Data on this outcome will be collected continuously during the trial.

5.5.2 Secondary outcomes

- All cause mortality within 24 hours, 30 days and three months of arrival at the emergency department. Data on this outcome will be collected in the same way as for the primary outcome.

- Quality of life within seven days of discharge, and at 30 days and three months of arrival at the emergency department, measured by the official and validated translations of the EQ5D3L. Data on this outcome will be collected in person if the patient is still in hospital, or by phone if the patient has been discharged. We will collect this data from a stratified random sample (site and period) of patient participants. The sampling will be designed so that it maximises statistical efficiency.
- Disability within seven days of discharge, and at 30 days and three months of arrival at the emergency department, assessed using the WHO Disability Assessment Schedule 2.0 (WHODAS 2.0). Data on this outcome will be collected in person if the patient is still in hospital, or by phone if the patient has been discharged. This data will also be collected from a stratified random sample of participants.
- Return to work at 30 days and three months after arrival at the emergency department. Data on this outcome will be collected in person if the patient is still in hospital, or by phone if the patient has been discharged.
- Length of emergency department stay. Data on this outcome will be collected from patient hospital records.
- Length of hospital stay. Data on this outcome will be collected from patient hospital records.
- Intensive care unit admission. Data on this outcome will be collected from patient hospital records.
- Length of intensive care unit stay. Data on this outcome will be collected from patient hospital records.
- Adherence to ATLS® principles during initial patient resuscitation, up to one hour after the physician has first seen the patient. This assessment will be done using a 14 item checklist covering the key steps of the ATLS® primary survey, which was modelled based on previous work on ATLS® adherence ^{aukstakalnis_impact_2024?}. We will consider completion of all 14 steps as 100% adherence. We will collect this data by observing the care of a random sample of patients. The sampling will be designed so that it maximises statistical efficiency. The clinical research coordinators collecting the data will be trained by participating in an ATLS® course as observers, prior to the start of the trial.

5.6 Statistical hypotheses

Our primary statistical hypotheses are:

- **Null hypothesis:** There is no difference in the primary outcome of 30-day in-hospital mortality between those randomised to ATLS® and standard care, meaning that the odds ratio (OR) for ATLS® vs standard care would be 1.
- **Alternative hypothesis:** There is an absolute difference in the primary outcome of 30-day in-hospital mortality between those randomised to ATLS® and standard care of at least 5% units, meaning that the OR for ATLS® vs standard care would

be different from 1. Our expectation, based on our pilot study and review of the literature, is that the OR will be less than 1, indicating lower odds of 30-day in-hospital mortality among those randomised to ATLS[®] group compared to those randomised to the standard care group.

5.7 Sample size calculations

With 30 clusters across 6 batches and a total sample size of 4320 our study has ~90% power across different combinations of cluster autocorrelations (CAC) and intra-cluster correlations (ICC) to detect a reduction in the primary outcome of in-hospital mortality within 30 days from 20% under standard care to 15% after ATLS[®] training (see Figure 2). This effect is a conservative estimate and the reduction equals a risk ratio of 0.75, which would be clinically important while also being consistent with our pilot study and updated systematic review. We allowed for the clustered design and assumed an ICC of 0.02, but considered sensitivity across the range 0.01-0.05^{2,3}, and a CAC of 0.9 but considered sensitivity across the range 0.8-1.0, based on our pilot study and current guidance⁴⁻⁶. We included the CAC to allow for variation in clustering over time. We assume that each cluster will contribute approximately 12 observations per month to the analysis, based on our previous work.

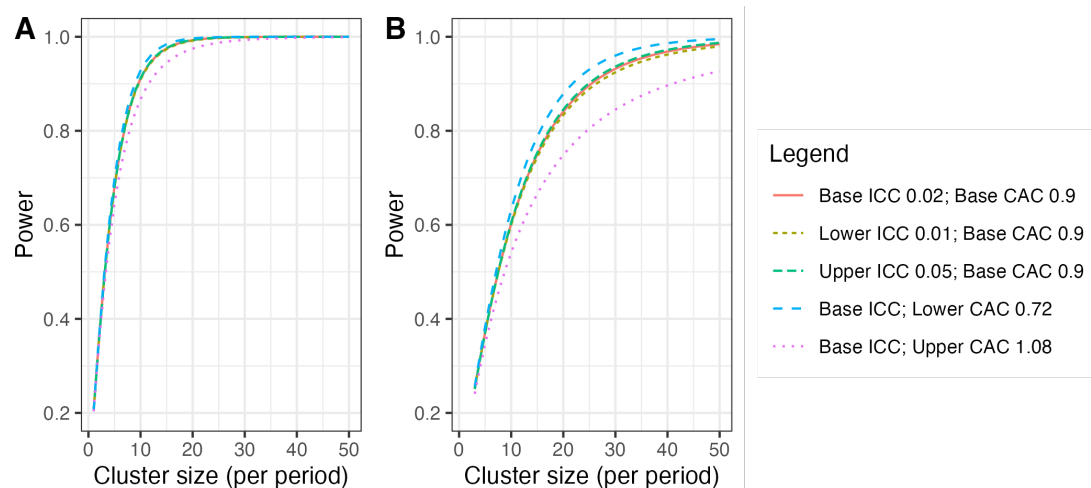


Figure 2: Power curves for different combinations of cluster autocorrelations (CAC) and intra-cluster correlations (ICC). **A)** Shows power curves assuming a reduction in the primary outcome of in-hospital mortality within 30 days from 20% under standard care to 15% after ATLS[®] training. **B)** Shows power curves assuming a reduction in the primary outcome from 10% under standard care to 7.5% after ATLS[®] training. Under this scenario, we would need to increase the sample size per month to around 30 observations to achieve 90% power under most combinations of CAC and ICC.

5.8 Statistical principles

5.8.1 Statistical software

We will use the R Statistical Software for all analyses⁷.

5.8.2 Levels of statistical significance and confidence

We will not perform any formal hypothesis testing as part of our planned interim analyses. We will use a two-sided significance level of 0.05 for all analyses, and we will report 95% confidence intervals (CI) for all estimates. We will not adjust for multiple testing because no secondary outcome is regarded as singularly more important.

5.9 Analysis populations

The unit of randomisation is the hospital, because all units trained in the same hospital will be treated as one cluster, but the unit of analysis is the individual patient. The group allocation for a patient depends on the period in which the patient was admitted to the hospital, and patients will be considered exposed to the intervention if they were admitted to the hospital at any time point following the transition period. We will use an intention-to-treat approach for all analyses, according to the planned date of transition from standard care. We might consider a per protocol analysis if cluster deviate considerably from the planned date of transition, using the actual date of transition as the cut-off point. We will report the results of the per protocol analysis as a sensitivity analysis. We will use a CONSORT diagram to display the flow of hospitals, clusters and patients through the trial. We will present cluster level summaries of the intervention effect. We will report the study according to the CONSORT guidelines for stepped-wedge randomised trials⁸.

5.10 Baseline analyses

5.10.1 Cluster characteristics

We will describe cluster characteristics including location and size using frequencies and percentages for discrete variables and means, standard deviations, medians and interquartile ranges (Q1-Q3) for continuous variables.

5.10.2 Patient characteristics

We will describe patient characteristics at baseline, meaning all pre-training periods, per treatment group and overall using frequencies and percentages for discrete variables and means, standard deviations, medians and interquartile ranges (Q1-Q3) for continuous variables. We will not adjust for clustering when presenting baseline characteristics.

5.11 Interim analysis

There will be one interim analyses after half of the batches have completed the trial. The interim analyses will be assessed by the joint Trial Steering and Data Monitoring Committee. The interim analysis will include the baseline analyses including the frequencies and proportions of the primary outcome in the intervention and control arms. We will recalculate the sample size based on the observed period sizes and the observed effect size, using the Shiny CRT Calculator⁴. In the revised sample size calculation, we will assume an ICC of 0.02, but consider sensitivity across the range 0.01-0.05^{2,3}, and a CAC of 0.9, but consider sensitivity across the range 0.8-1.0⁴⁻⁶. These assumptions are the same as in the original sample size calculation. We will not perform any formal hypothesis testing as part of our planned interim analyses.

5.12 Analysis of the primary outcome

The primary outcomes is in-hospital mortality within 30 days of arrival at the emergency department and will be analysed as a dichotomous variable. We will estimate the primary intervention effect as the OR of death between the ATLS[®] and standard care arms, with an OR < 1 indicating lower odds of death in the ATLS[®] arm compared to the standard care arm and vice versa.

5.12.1 Main analysis: mixed effects binomial model with logit link

We will use a mixed effects binomial model with a logit link to estimate the OR. We will include fixed effects for period as a categorical variable and a fixed effect for intervention exposure². The primary analysis will allow for clustering as a random cluster and random cluster by period effect³, both assumed to follow a normal distribution. The full model is specified in Equation 1. To correct the potential inflation of the type I error rate due to small number of clusters, the Kenward and Roger small sample correction will be

²**Question:** Should we adjust for calendar time because the last batch will be two years later than the first batch?

³**Question:** Just a thought: keep in mind the complexity of the random effects structure; if it leads to convergence problems or overfitting, should we consider simplifying the random effects?

used [9]⁴. This model will be fitted using residual pseudo-likelihood estimation based on linearization with subject-specific expansion (RSPL).

$$\text{logit}(\Pr(Y_{bkti} = 1)) = \mu + \beta_{bt} + \theta X_{bkt} + \alpha_{bk} + \gamma_{bkt} \quad (1)$$

Where:

- $\Pr(Y_{bkti} = 1)$: the probability of death for patient $i = 1, \dots, m$ in cluster $k = 1, \dots, 30$ during period $t = 1, \dots, 12$ in batch $b = 1, \dots, 6$.
- μ : the intercept, representing the baseline log-odds of the outcome when all predictors are set to 1.
- β_{bt} : the fixed effect of period t in batch b , accounting for a separate period effect for each batch, so that there is a total of 72 period effects.⁵
- θ : the fixed effect of intervention exposure, representing the effect of ATLS® exposure on the probability of death.
- X_{bkt} : the treatment arm indicator for patient i in cluster k during period t in batch b , with $X_{bkt} = 1$ for ATLS® and $X_{bkt} = 0$ for standard care.
- α_{bk} : the random effect of cluster k in batch b , representing the variability in outcomes between different clusters (hospitals) across batches.
- γ_{bkt} : the random effect of cluster k in period t in batch b , accounting for the random variation within clusters across periods and batches, including potential interaction effects between clusters and time.

We will present the effect of ATLS® exposure as an odds ratio (OR) for mortality with an associated 95% CI, using the standard care arm as the reference. Additionally, we will present the risk difference with a 95% CI. The randomization within each batch will be balanced based on cluster size, defined as the expected monthly volume of eligible patient participants. Therefore, no further adjustment for cluster size will be made in the main analysis.

5.12.2 Sensitivity analyses

The sensitivity analyses will be conducted to assess the robustness of the main analysis results to different model specifications. We will first model the primary outcome using an identity link function to estimate the risk difference instead of the OR. Henceforth, each additional sensitivity analyses will be operationalised using two separate models, one with the logit link and one with the identity link. We will first explore more complex

⁴**Questions about Kenward and Roger small sample correction:** 1. What clusters do we consider here? Doctors? Patients? Hospitals? 2. Why use this method and not eg. bootstrapping? I thought Kenward and Roger was primarily used for linear mixed models and not for binary outcomes?

⁵**Question:** The model will estimate some 78 parameters. Is this a problem considering the low number of participants per period?

correlation structures. We will then model time using a spline function. Finally, we will conduct a fully adjusted covariate analysis.

Model with identity link

We will use an identity link used to estimate the risk difference, meaning that the coefficient will be interpreted as the difference in the probability of death between the ATLS® and standard care arms. We will present the risk difference with a 95% CI. This model is specified in Equation 2 and will also be fitted using RSPL. If the binomial model with the identity link does not converge then only a odds ratio will be reported.

$$\Pr(Y_{bkti} = 1) = \mu + \beta_{bt} + \theta X_{bkt} + \alpha_{bk} + \gamma_{bkt} \quad (2)$$

Where:

- μ : the intercept, representing the baseline probability of the outcome when all predictors are zero.

Models with different correlation structure

We will explore if models with more complicated correlation structures are a better fit to the data. These models are not being used as our primary analysis models as there is limited understanding as to when such models will converge and how to choose between the various different correlation structures which might be plausible. First, we will include a discrete time decay correlation structure including a random cluster effect with auto-regressive structure (AR(1)), described in Equation 3.

$$\alpha_{bk,t} = \rho\alpha_{bk,t-1} + \epsilon_t, \quad \epsilon_t \sim N(0, \sigma_\alpha^2) \quad (3)$$

Where:

- ρ : the correlation between the random effects of two consecutive periods, the period t and the period $t - 1$, reflecting how outcomes in one period are correlated with outcomes in the previous period within the same cluster.
- $\alpha_{bk,t}$: the random effect of cluster k in period t in batch b . This represents the variation in the outcome that is specific to each cluster and period, accounting for time-related effects.
- ϵ_t is the error term for period t , which is assumed to be normally distributed with mean 0 and variance σ_α^2 .

To allow for the randomisation by batches, we will also include a different secular trend for each batch as a random effect interaction term between batch and period. The full model is specified in Equation 4.

$$g(\Pr(Y_{bkti} = 1)) = \mu + \beta_{bt} + \theta X_{bkt} + \alpha_{bk,t} + \gamma_{bkt} + \delta_{bt} \quad (4)$$

Where:

- $g(\cdot)$: the link function.
- $\alpha_{bk,t}$: the updated random effect of cluster k in batch b during period t with the AR(1) correlation structure. This accounts for correlation between adjacent periods in the same cluster.
- δ_{bt} : the random effect of batch b during period t , which captures any variability across different batches over time.

Models with random cluster by intervention effects

Models will also be extended to include random cluster by intervention effects (with a non-zero covariance term) to examine if results are sensitive to the assumption of no intervention by cluster interaction. The model is specified in Equation 5.

$$g(\Pr(Y_{bkti} = 1)) = \mu + \beta_{bt} + \theta X_{bkt} + \alpha_{bk} + \gamma_{bkt} + u_{bk} \times X_{bkt} \quad (5)$$

Where:

- u_{bk} : the random effect of cluster k by intervention interaction.
- X_{bkt} : the treatment arm indicator for patient i in cluster k during period t .
- $u_{bk} \times X_{bkt}$: the random interaction effect between the intervention and cluster k .

Models with time modelled with a spline function

We will further explore the potential for a time-varying treatment effect¹⁰. To explore if the fixed period effect is both parsimonious and adequate to represent the extent of any underlying secular trend, we will model the time effect using natural cubic splines with knots at the equally spaced time points 3, 6 and 9. This will result in five spline basis functions, because the natural cubic splines are modelled with three degrees of freedom but are constrained to be linear before the first and after the last knot. The model is specified in Equation 6.

$$g(\Pr(Y_{bkti} = 1)) = \mu + \sum_{j=1}^5 \beta_j S_j(t, \{3, 6, 9\}) + \theta X_{bkt} + \alpha_{bk} + \gamma_{bkt} \quad (6)$$

Where:

- $S_j(t, \{3, 6, 9\})$: the natural cubic spline basis functions with knots placed at times 3, 6 and 9.
- β_j : the coefficient for the j -th spline basis function.

Models exploring lag and weaning effects

Models will also be extended to include an interaction between treatment and number of periods since first treated, to examine if there is any indication of a relationship between duration of exposure to the intervention and outcomes. This will allow us to model different lag effects (whereby it takes time for the intervention to become embedded within the culture before its impact can properly start to be realised); as well as weaning effects (whereby the effect of the intervention starts to decrease – or fade). This type of analysis attempts to disentangle how some clusters end up having a long exposure to the intervention and others have a much shorter exposure time. The model is specified in Equation 7.

$$g(\Pr(Y_{bkti} = 1)) = \mu + \beta_{bt} + \theta X_{bkt} + \theta_{\text{int}} X_{bkt} \times T_{bkt} + \alpha_{bk} + \gamma_{bkt} \quad (7)$$

Where:

- θ_{int} is the coefficient for the interaction between treatment and time since first treated.
- T_{bkt} is the number of periods since first treated.

5.12.3 Adjusted analyses

Fully adjusted covariate analysis will additionally adjust for⁶ (expected % missing data in parentheses):

- Age (<5%)
- Sex (<5%)
- Systolic blood pressure (25%)
- Glasgow Coma Scale (25%)
- Injury Severity Score (10%)
- Mechanism of injury (10%)

⁶**Questions:** 1. Should we adjust for calendar time because the last batch will be two years later than the first batch? 2. Should we add it as covariate in the list above? 3. Should it be continuous (e.g., months since the study started) or categorical calendar time (if you expect non-linear trends or stepwise changes over years)? 4. Should we analyse Time Since Admission? Depending on the nature of the trauma, the time since patient admission might also influence outcomes, especially if the timing of interventions or care decisions plays a critical role. Eg. time from admission to surgery.

These are known individual-level prognostic factors for the primary outcome. These covariates will be included in the models specified in Equation 1 and Equation 2 as fixed effects. We will model the continuous covariates assuming linear effects.

5.12.4 Subgroup analyses

We will perform the following subgroup analyses⁷:

- geographical region, defined using the state in which the participating hospital is located. Demonstrating the consistency of any effect across multiple regions will enhance the generalisability of the results;
- age groups, defined as older adolescents (15-19 years), young adults (20-24 years), adults (25-59 years), and older adults (60 years and older) [11];
- sex, using the levels male and female;
- clinical cohorts, defined as blunt multisystem trauma, penetrating trauma, and severe isolated traumatic brain injury, with modification to avoid overlap between the cohorts; and
- cluster size (number of patients per hospital)

These subgroup analyses will be conducted by adding the subgroup variable and the interaction between the subgroup variable and the intervention exposure variable as fixed effects to the models specified in Equation 1 and Equation 2.

5.12.5 Treatment of missing data

We will present the frequency and percentage of missing data for all variables. If the percentage of missing data for the primary outcome is less than 10%, we will perform a complete case analysis. If the percentage of missing data for the primary outcome is 10% or more, we will handle missing data depending on the missing data mechanism. If the data are missing at random (MAR), we will perform multiple imputation using multiple imputation by chained equations (MICE), imputing data for the primary outcome as well as all covariates included in the fully adjusted model. The number of imputations will be determined by the percentage of missing data, with a minimum of 20 imputations. If there is evidence that the data are missing not at random (MNAR), we will explore the impact of this assumption using a sensitivity analysis (e.g., pattern mixture models or selection models) to assess how robust our findings are to different assumptions about the missing data mechanism. Additionally, we will perform diagnostic checks after multiple imputation to ensure the quality of the imputation process, including comparing distributions of observed and imputed data and checking convergence.

⁷**Question:** Are we exploring too many subgroups?

5.13 Analysis of secondary outcomes

5.13.1 All cause mortality within 24 hours, 30 days and three months of arrival at the emergency department

We will use the model with the logit link specified in Equation 1 and the model with the identity link specified in Equation 2 to estimate the OR and risk difference for these mortality outcomes.

5.13.2 Quality of life within seven days of discharge, and at 30 days and three months of arrival at the emergency department

Quality of life will be measured by the official and validated translations of the EQ5D5L. This tool assesses five dimensions of health-related quality of life: mobility, self-care, usual activities, pain/discomfort, and anxiety/depression. Each dimension is rated on a likert scale from 1 to 5. There is also a visual analogue scale (VAS) for self-rated quality of life, ranging from 0 to 100. For each of the five dimensions we will use a mixed effects ordinal model as specified in Equation 8.

$$\text{logit}(\Pr(Y_{bkti} \leq j)) = \mu_j + \beta_{bt} + \theta X_{bkt} + \alpha_{bk} + \gamma_{bkt} \quad (8)$$

Where:

- μ_j : the intercept for the j -th category of the EQ5D dimension ($j = 1, 2, 3, 4, 5$).

The VAS will be analysed using a linear mixed effects model as specified in Equation 9.

$$\text{VAS}_{bkti} = \mu + \beta_{bt} + \theta X_{bkt} + \alpha_{bk} + \gamma_{bkt} + \epsilon_{bkti}, \quad \epsilon_{bkti} \sim N(0, \sigma^2) \quad (9)$$

Where:

- ϵ_{bkti} : the error term for patient i in cluster k in period t in batch b , assumed to be normally distributed with mean 0 and variance σ^2 .

5.13.3 Disability within seven days of discharge, and at 30 days and three months of arrival at the emergency department

We will measure disability using the WHO Disability Assessment Schedule 2.0 (WHODAS 2.0)¹². This tool assesses six domains of functioning: cognition, mobility, self-care, getting along, life activities, and participation. Each domain is rated on a likert scale from 1 to 5, with 1 indicating no difficulties and 5 indicating extreme difficulties. We will analyse each domain separately using a mixed effects ordinal model as specified in Equation 8. We will also calculate a WHODAS 2.0 summary score using the method referred to as the “complex scoring” method. This method involves summing the item scores within each of the six domains, then summing the scores of all domains, and finally transforming the total score to a 0-100 scale. We will analyse the summary score using a linear mixed effects model as specified in Equation 9.

5.13.4 Return to work at 30 days and three months after arrival at the emergency department

We will analyse return to work as a dichotomous variable using a mixed effects binomial model with a logit link as specified in Equation 1.

5.13.5 Length of emergency department stay

We will analyse length of emergency department stay as a continuous variable using a linear mixed effects model as specified in Equation 9.

5.13.6 Length of hospital stay

We will analyse length of hospital stay as a continuous variable using a linear mixed effects model as specified in Equation 9.

5.13.7 Intensive care unit admission

We will analyse intensive care unit admission as a dichotomous variable using a mixed effects binomial model with a logit link as specified in Equation 1.

5.13.8 Length of intensive care unit stay

We will analyse length of intensive care unit stay as a continuous variable using a linear mixed effects model as specified in Equation 9.

5.13.9 Adherence to ATLS® principles during initial patient resuscitation

We will assess adherence to ATLS® principles using a 14-item checklist that covers the key steps of the ATLS® primary survey. Adherence will be measured as the proportion of steps completed, with 100% adherence defined as completing all 14 steps. We will analyse adherence using a mixed-effects beta regression model with a logit link as specified in Equation 10.

$$\text{logit}(E(Y_{bkti})) = \mu + \beta_{bt} + \theta X_{bkt} + \alpha_{bk} + \gamma_{bkt} \quad (10)$$

Where:

- $E(Y_{bkti})$: The expected proportion of completed ATLS steps for patient i in cluster k during period t in batch b .
- μ : The intercept, representing the baseline log-odds of completing ATLS steps when all predictors are set to 1.
- β_{bt} : The fixed effect of period t in batch b , allowing for different period effects for each batch.
- θ : The fixed effect of ATLS exposure, representing the effect of the intervention on adherence to ATLS principles.
- X_{bkt} : The treatment arm indicator for patient i in cluster k during period t in batch b (1 for ATLS, 0 for standard care).
- α_{bk} : The random effect of cluster k in batch b , capturing variability in adherence across clusters (hospitals).
- γ_{bkt} : The random effect of cluster k in period t within batch b , accounting for within-cluster variability across time periods and batches.

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