

Parameter Uncertainty for Credit Risk in the Vasicek Model

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The Vasicek model for credit risk is inferred from the hypothesis that a firm would defaults if its firm value  $V_i(T) < d$ , where

$$V_i = \sqrt{\rho}Z + \sqrt{1 - \rho}\epsilon_i$$

where

- $lacksquare Z \sim \mathcal{N}(0,1)$  is the common risk factor
- ullet  $\epsilon_i \sim \mathcal{N}(0,1)$  i.i.d are idiosyncratic risk factors
- lacksquare  $\rho$  is the correlation between firms values

### HP case

### LHP case

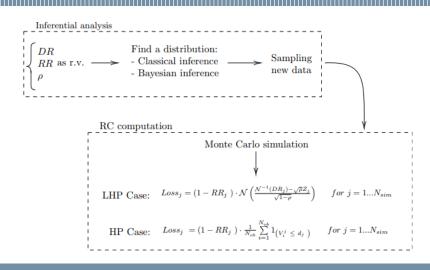
Regulatory Capital:

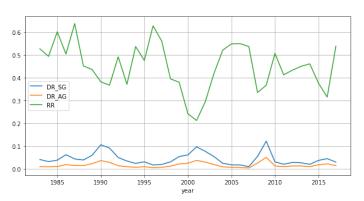
$$\alpha RC := \alpha VaR - \mathbb{E}[Loss]$$

## Remark (Deterministic parameters)

 $DR, RR, \rho$  are parameters obtained from the average of the corresponding observed quantities

### Altenative model





Correlation  $\rho$  is inferred from publicly traded firms in the US from asset returns in the month of July 2006.

Frequentistic inference parameters distribution assumption:

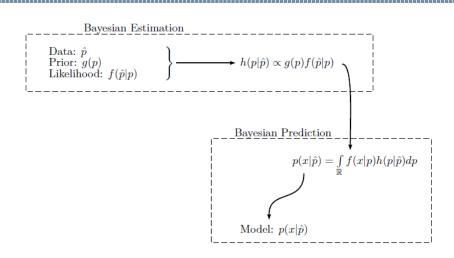
- $lacksquare d = \mathcal{N}^{-1}(DR) \sim \mathcal{N}(\hat{d}, \sigma_d)$
- $\blacksquare$   $RR \sim \mathcal{N}(\overline{RR}, \sigma_{RR})$
- $ho \sim Beta(lpha_
  ho,eta_
  ho)$  vs  $ho_{Basel}(DR,k)|_{k=50}$

Gaussianity of d and RR verified with Shapiro-Wilk test

#### Remark (Default barrier)

We have two ways of choosing the mean of the distribution of d:

- sample mean of  $\{\mathcal{N}^{-1}(DR_i)\}_i \ (\hat{d} = -1.7733)$
- $\hat{d}$  such that  $\mathbb{E}[\mathcal{N}(d)] = \frac{1}{N} \sum_{i=1}^{N} DR_i \ (d = -1.7731)$



#### Correlation

- Prior distribution:  $g(\rho) = \mathbb{1}_{[0,1]}(\rho)$
- Likelihood  $f(\hat{\rho}|\rho)$  is a beta distribution with  $\alpha(\rho)$  and  $\beta(\rho)$  s.t.

$$\rho = \frac{\alpha}{\alpha + \beta} \qquad \sigma^2 = \frac{\alpha\beta}{(1 + \alpha + \beta)(\alpha + \beta)^2}$$

Hence the posterior can be written as

$$h(
ho|\hat{
ho}) \propto \mathbb{1}_{[0,1]}(
ho) \textit{Beta}(\hat{
ho}|lpha(
ho),eta(
ho))$$

We recover the class conditional density as

$$p(x|\hat{
ho}) = \int\limits_{\mathbb{R}} \mathsf{Beta}(x|
ho, \sigma_{
ho}) h(
ho|\hat{
ho}) d
ho$$

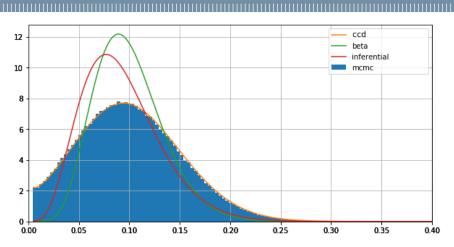


Figure: Distribution of the parameter  $\rho$  according to different approaches.

- lacksquare Prior distribution:  $d \sim \mathcal{N}(0,1)$  as  $DR \sim \mathcal{U}[0,1]$
- Likelihood:  $f(\{d_i\}_i|d,\sigma_d) = \prod_{i=1}^N \phi(d_i|d,\sigma_d)$  where  $\sigma_d$  is the sample standard deviation of  $\{d_i\}_i = \{\mathcal{N}^{-1}(PD_i)\}_i$ .

Since the prior and the likelihood are conjugated, we know that also the posterior  $h(d|\{d_i\}_i) \sim \mathcal{N}\left(\frac{\sum d_i}{N+\sigma_d^2}, \frac{\sigma_d^2}{N+\sigma_d^2}\right)$ 

#### Remark

In [Baviera et al.] was used just  $\hat{d}$  and the variance of the mean of  $\sigma_d^2/N$  for a normal i.i.d. sample. Therefore the posterior  $h \sim \mathcal{N}\left(\frac{\hat{d}}{1+\sigma_c^2/N}, \frac{\sigma_d^2}{N+\sigma_c^2}\right)$ 

# Cramer Rao lower bound (HP case)

We might also want to use the Cramer-Rao lower bound for variance of unbiased estimators for the parameter, this would be a function of the unknown parameter p. It is defined as:

$$\sigma_{CR}^2(p) = -\frac{1}{T} \left( \mathbb{E}_y \left[ \frac{\partial^2 LL(y|p)}{\partial p^2} \right] \right)^{-1}$$

where  $LL(\cdot|p)$  is the log-likelihood function for the observed data and T the length of time series.

# Cramer Rao lower bound (HP case)

■ For  $\rho$  there is a closed formula:

$$\sigma_{CR,\rho}^2(\rho) = \frac{2(1-\rho)^2(1+(N_{ob}-1)\rho)^2}{N_{ob}(N_{ob}-1)T}$$

 $\blacksquare$  For d, it is computed numerically using as likelihood:

$$P(X=m) = \int_{-\infty}^{\infty} {N_{ob} \choose m} p(y)^m (1-p(y))^{N_{ob}-m} \phi(y) dy$$

where m is the number of defaults and  $p(y) = \mathcal{N}\left(\frac{d-\sqrt{\rho}y}{\sqrt{1-\rho}}\right)$ .

# Cramer Rao lower bound (HP case)

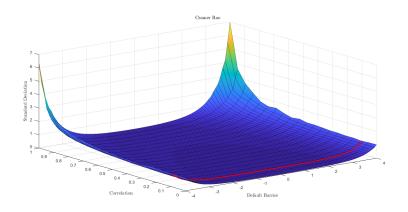


Figure: Surface for the standard deviation (Cramer Rao)

Add-on	99%	99,90%	
d	47,33%	54,28%	
d <sup>[closed]</sup>	47,25%	53,20%	
RR	8,41%	12,04%	
ρ	4,16%	16,60%	
<b>d</b> , ρ	49,11%	62,54%	
All	57,30%	75,08%	

Table: Add-on for LHP with frequentist inference (Speculative Grade)

Add-on	99%	99,90%
d	29,33%	36,97%
RR	7,93%	10,58%
ρ	0,01%	9,23%
d, ρ	29,34%	46,22%
All	39,30%	58,81%

Table: Add-on for HP with frequentist inference (Speculative Grade)

Add-on	99%	99,90%
d	49,52%	56,00%
ρ	14,65%	31,22%
All	59,34%	76,99%

Table: Add-on for LHP with Bayesian inference (Speculative Grade)

Add-on	99%	99,90%	
d	28,97%	36,76%	
ρ	14,66%	18,47%	
All	43,66%	55,25%	

Add-on	99%	99,90%	
d	28,66%	36,52%	
ρ	14,63%	18,47%	
All	43,29%	54,96%	

Table: Add-on for HP with Bayesian inference fixing variance equal to empiric one (table on the left) and Cramer Rao lower bound (table on the right)

In the 1980, Global Sensitivity Analysis (GSA) has been developed, which aims to identify what inputs are the most influential in the model.

$$f(X_1,\cdots,X_n)=Y$$

Importance of fixing  $X_i$  is high when  $Var[Y|X_i=x_i^*]$  is low. But the "true" value  $x_i^*$  is unknown  $\implies$  use  $\mathbb{E}[Var[Y|X_i]]$ .

$$Var[Y] = Var[\mathbb{E}[Y|X_i]] + \mathbb{E}[Var[Y|X_i]]$$
  
 $\implies S_1^{(i)} = \frac{Var[\mathbb{E}[Y|X_i]]}{V[Y]}$ (Sobol index)

	DR	RR	$\rho$
$S_1$	36.84	7.99	23.16
$NS_1$	53.76	12.25	33.99

Table: Sobol indices for LHP (%)

	DR	RR	$\rho$
$S_1$	51.24	12.08	30.26
$NS_1$	54.75	12.91	32.33

Table: Sobol indices for LHP (%). Independent parameters

Conclusions 17/

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- It would be unfair to say that banks are for sure under-capitalized, since our add-on was computed with respect to the mean of a many years worth of data.

Conclusions

Default Rate is the most sensitive parameter in the model and provides the most variance to the output

- Frequentist and Bayesian analysis have produced very similar results, but with very different degrees of complexity.
- It would be unfair to say that banks are for sure under-capitalized, since our add-on was computed with respect to the mean of a many years worth of data.
- Most important parameter in the model is the Default Rate! more capital/larger team/most defined model is needed to estimate it.