

# Stochastic Processes: Homework 2

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## Exercise 1

Let  $f(x) = \frac{1}{2} \sin(x) \mathbb{1}_{[0, \pi]}(x)$ .

1. For  $n = 4$  **calculate** the first moment of the order statistic  $X_{i:n}$  for every  $i \in 1, \dots, 4$ .
2. **Sketch a telling drawing** with the original density, the densities  $f_{i:4}$  and  $\mathbf{E}[X_{i:4}]$

## Solution Part 1

In the first place, for  $x \in [0, \pi]$

$$\begin{aligned} F(x) &= \frac{1}{2} \int_{-\infty}^x \sin(t) \mathbb{1}(t) dt = \frac{1}{2} \int_0^x \sin(t) dt \\ &= \cos(x) \Big|_0^x = \frac{1}{2} (1 - \cos(x)) \cdot \mathbb{1}_{[0, \pi]}(x) \end{aligned}$$

Thus,

$$\overline{F}(x) = \frac{1}{2} (1 + \cos(x)) \cdot \mathbb{1}_{[0, \pi]}(x),$$

and,

$$\begin{aligned} f_{i:n}(x) &= i \cdot \binom{n}{i} f(x) \cdot F^{i-1}(x) \cdot \overline{F}^{n-i}(x) \\ &= i \cdot \binom{n}{i} \frac{1}{2^n} \cdot \sin(x) \cdot (1 - \cos(x))^{i-1} \cdot (1 + \cos(x))^{n-i}. \end{aligned}$$

For simplicity, define  $C_{i:n} = i \binom{n}{i} 2^{-n}$ . Now let  $n = 4$ . In order to calculate the exact expected value formula, we must use the following angle identities,

$$\sin(x) \cos(x) = \frac{1}{2} \sin(2x), \quad \cos^2(x) = \frac{1}{2} (1 + \cos(2x)),$$

and the result of these integrals

$$\int x \sin(kx) dx = \frac{\sin(kx) - kx \cos(kx)}{k^2}.$$

From now on I'm going to simplify the trigonometric expressions using the TR8 algorithm provided by the [package sympy](#). Also, use the package to integrate and evaluate the final expression. As always, the code is included with this document.

•  $i = 1$  :

$$C_{1:4}^{-1} \cdot f_{1:4} = \frac{7 \sin(x)}{4} + \frac{7 \sin(2x)}{4} + \frac{3 \sin(3x)}{4} + \frac{\sin(4x)}{8},$$

$$\begin{aligned} C_{1:4}^{-1} \cdot \mathbf{E}[X_{1:4}] &= C_{1:4}^{-1} \cdot \int_0^\pi x f_{1:4} dx. \\ &= \left. -\frac{7x \cos(x)}{4} - \frac{7x \cos(2x)}{8} - \frac{x \cos(3x)}{4} - \frac{x \cos(4x)}{32} \right|_0^\pi \\ &\quad + \left. \frac{7 \sin(x)}{4} + \frac{7 \sin(2x)}{16} + \frac{\sin(3x)}{12} + \frac{\sin(4x)}{128} \right|_0^\pi = \frac{35\pi}{32}. \end{aligned}$$

•  $i = 2$  :

$$C_{2:4}^{-1} \cdot f_{2:4} = \frac{3 \sin(x)}{4} + \frac{\sin(2x)}{4} - \frac{\sin(3x)}{4} - \frac{\sin(4x)}{8},$$

$$\begin{aligned} C_{2:4}^{-1} \cdot \mathbf{E}[X_{2:4}] &= C_{2:4}^{-1} \cdot \int_0^\pi x f_{2:4} dx. \\ &= \left. -\frac{3x \cos(x)}{4} - \frac{x \cos(2x)}{8} + \frac{x \cos(3x)}{12} + \frac{x \cos(4x)}{32} \right|_0^\pi \\ &\quad + \left. \frac{3 \sin(x)}{4} + \frac{\sin(2x)}{16} - \frac{\sin(3x)}{36} - \frac{\sin(4x)}{128} \right|_0^\pi = \frac{55\pi}{96}. \end{aligned}$$

•  $i = 3$  :

$$C_{3:4}^{-1} \cdot f_{3:4} = \frac{3 \sin(x)}{4} - \frac{\sin(2x)}{4} - \frac{\sin(3x)}{4} + \frac{\sin(4x)}{8},$$

$$\begin{aligned} C_{3:4}^{-1} \cdot \mathbf{E}[X_{3:4}] &= C_{3:4}^{-1} \cdot \int_0^\pi x f_{3:4} dx. \\ &= \left. -\frac{3x \cos(x)}{4} + \frac{x \cos(2x)}{8} + \frac{x \cos(3x)}{12} - \frac{x \cos(4x)}{32} \right|_0^\pi \\ &\quad + \left. \frac{3 \sin(x)}{4} - \frac{\sin(2x)}{16} - \frac{\sin(3x)}{36} + \frac{\sin(4x)}{128} \right|_0^\pi = \frac{73\pi}{96}. \end{aligned}$$

•  $i = 4$  :

$$C_{4:4}^{-1} \cdot f_{4:4} = \frac{7 \sin(x)}{4} - \frac{7 \sin(2x)}{4} + \frac{3 \sin(3x)}{4} - \frac{\sin(4x)}{8},$$

$$\begin{aligned} C_{4:4}^{-1} \cdot \mathbf{E}[X_{4:4}] &= C_{4:4}^{-1} \cdot \int_0^\pi x f_{4:4} dx. \\ &= \left. \frac{7x \cos(x)}{4} + \frac{7x \cos(2x)}{8} - \frac{x \cos(3x)}{4} + \frac{x \cos(4x)}{32} \right|_0^\pi \\ &\quad - \left. \frac{7 \sin(x)}{4} - \frac{7 \sin(2x)}{16} + \frac{\sin(3x)}{12} - \frac{\sin(4x)}{128} \right|_0^\pi = \frac{93\pi}{32}. \end{aligned}$$

Therefore,

$$\begin{aligned} \mathbf{E}[X_{1:4}] &= \frac{35\pi}{128}, & \mathbf{E}[X_{2:4}] &= \frac{55\pi}{128}, \\ \mathbf{E}[X_{3:4}] &= \frac{73\pi}{128}, & \mathbf{E}[X_{4:4}] &= \frac{93\pi}{128}. \end{aligned}$$

## Solution Part 2

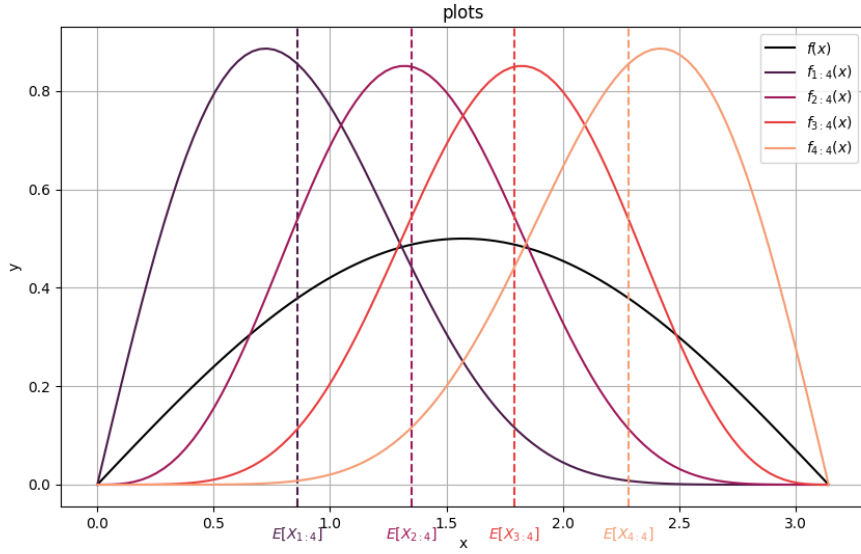


Figure 1: This is a floating figure with an image.

## Exercise 2

1. **Formulate** the criterion that characterizes the existence of an extreme distribution (in terms of  $\overline{F}$ ).
2. **Determine and justify** by this criterion if the following distribution has an extreme distribution. The cumulative distribution function is given by

$$F(n) := 1 - \frac{C}{(n+1)^{\ln(n+1)}}, \quad n \in \mathbb{N}.$$

3. In the case that it has a limit distribution, **argue** about which should be the limit distribution.

## Solution Part 2

$$\overline{F}(n) = \frac{C}{(n+1)^{\ln(n+1)}}.$$

## Exercise 3

Let  $(X_i)_{i \in \mathbb{N}}$  be a sequence of i.i.d. random variables with  $X_1 \sim \mu$ ,

$$M_n := \max\{X_1, \dots, X_n\}, \quad \text{and} \quad N_n := \min\{X_1, \dots, X_n\}.$$

1. For  $\mu = \text{Gamma}(\alpha, \beta)$ ,  $\alpha, \beta > 0$ , **determine and justify** the extreme distribution of  $(M_n)$  and  $(N_n)$ .
2. For  $\mu = \text{Beta}(\alpha, \beta)$ ,  $\alpha, \beta > 0$ , that is  $f(x) = C_{\alpha, \beta} x^{\alpha-1} (1-x)^{\beta-1}$ ,  $\alpha, \beta > 0$ , **determine and justify** the extreme distribution of  $(M_n)$  and  $(N_n)$ .
3. For  $\mu$  such that for  $\alpha > 0$ ,

$$F(x) = \begin{cases} 0, & \text{for } x < 1, \\ \frac{\ln(x)}{x^\alpha}, & \text{for } x \geq 1. \end{cases}$$

**determine and justify** the extreme distribution of  $(M_n)$  and  $(N_n)$

## Solution Part 1

The gamma distribution with parameters  $\alpha, \beta > 0$  has the following density function:

$$f(x) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x}, \quad x \in [0, \infty).$$

It's clear that the Gamma distribution has extreme distribution convergence since it's absolutely continuous. The extreme distribution cannot be Weibull because the population maximum  $x_F = \infty$ . Now, we are going to prove that the tails are not regular, and thus, the extremes do not converge to Fréchet (using L'Hôpital rule):

$$\begin{aligned}
\lim_{x \rightarrow \infty} \frac{\overline{F}(\lambda \cdot x)}{\overline{F}(x)} &= \lim_{x \rightarrow \infty} \frac{\lambda \cdot f(\lambda x)}{f(x)} \\
&= \lim_{x \rightarrow \infty} \frac{\lambda^\alpha x^{\alpha-1} e^{-\beta \lambda^2 x^2}}{x^{\alpha-1} e^{-\beta x^2}} \\
&= \lim_{x \rightarrow \infty} \lambda^\alpha \exp(-\beta x^2(1 - \lambda^2)) \\
&= 0, \forall \lambda > 0.
\end{aligned}$$

Therefore, the only option left for  $M_n$  is to be in the Gumbell's domain of attraction.

Now, for  $N_n$ , we are going to use the von Mises condition to prove that it is in the domain of attraction of the Weibull distribution. In the first place, note that the population minimum  $x_F$  equals 0 and we are approaching to it from the right. Thus, after reflecting the whole distribution over the  $y$  axis, we obtain

$$\begin{aligned}
f'(x) &= (\alpha - 1)x^{\alpha-2}e^{-\beta x^2} - x^{\alpha-1} \cdot 2\beta x e^{-\beta x^2} \\
&= x^{\alpha-2}e^{-\beta x^2}((\alpha - 1) - 2\beta x^2)
\end{aligned}$$

$$\begin{aligned}
\lim_{x \rightarrow 0^-} \frac{x f(x)}{F(x)} &= \lim_{x \rightarrow 0^-} \frac{x f'(x) + f(x)}{f(x)} \\
&= \lim_{x \rightarrow 0^-} \frac{x^{\alpha-1}e^{-\beta x^2}((\alpha - 1) - 2\beta x^2)}{x^{\alpha-1}e^{-\beta x^2}} + \frac{f(x)}{f(x)} \\
&= \lim_{x \rightarrow 0^-} (\alpha - 1) - 2\beta x^2 + 1 = \alpha.
\end{aligned}$$

Thus,  $c_n N_n + d_n$  converges to a Weibull distribution with parameter  $\alpha$ .

## Solution Part 2

Both  $M_n$  and  $N_n$  will be in the domain of attraction of Weibull. For  $M_n$  the von Mises condition is satisfied on the following limit:

$$f'(x) = (\alpha - 1)x^{\alpha-2}(1-x)^{\beta-1} - (\beta - 1)x^{\alpha-1}(1-x)^{\beta-2}$$

$$= x^{\alpha-2}(1-x)^{\beta-2}((\alpha - 1)(1-x) - (\beta - 1)x)$$

$$\begin{aligned} \lim_{x \rightarrow 1^+} \frac{(1-x)f(x)}{\bar{F}(x)} &= \lim_{x \rightarrow 1^+} \frac{(1-x)f'(x) - f(x)}{-f(x)} \\ &= \lim_{x \rightarrow 1^+} \frac{(1-x) \cdot x^{\alpha-2}(1-x)^{\beta-2}((\alpha - 1)(1-x) - (\beta - 1)x)}{-x^{\alpha-1}(1-x)^{\beta-1}} + 1 \\ &= \lim_{x \rightarrow 1^+} -(\alpha - 1) \underbrace{\frac{1-x}{x}}_{\rightarrow 0} + (\beta - 1) + 1 \\ &= \beta - 1 + 1 = \beta. \end{aligned}$$

Now, for  $N_n$ , after reflecting the whole distribution, we obtain:

$$\begin{aligned} \lim_{x \rightarrow 0^-} \frac{xf(x)}{F(x)} &= \lim_{x \rightarrow 0^-} \frac{xf'(x) + f(x)}{f(x)} \\ &= \lim_{x \rightarrow 0^-} \frac{x \cdot x^{\alpha-2}(1-x)^{\beta-2}((\alpha - 1)(1-x) - (\beta - 1)x)}{x^{\alpha-1}(1-x)^{\beta-1}} + 1 \\ &= \lim_{x \rightarrow 0^-} (\alpha - 1) - (\beta - 1) \underbrace{\frac{x}{1-x}}_{\rightarrow 0} + 1 \\ &= \alpha - 1 + 1 = \alpha. \end{aligned}$$

For  $M_n$  is Weibull with parameter  $\beta$  and for  $N_n$  is Weibull with parameter  $\alpha$ .

### Solution Part 3