# Stochastic Processes: Homework 0

### Martín Prado

November 2023 Universidad de los Andes — Bogotá Colombia

### Exercise 1

Consider a sequence of i.i.d. random variables  $(X_i)_{i\in\mathbb{N}}$  with  $\mathbf{E}\,X_i=0$  and  $\mathbf{Var}\,X_i=1$  for every  $i\in\mathbb{N}$ .

1. Show with th Law of Large Numbers that,

$$\lim_{n\to\infty} \|X_1,\dots,X_n\|_2 - \sqrt{n} \to 0$$

- (a) in  $\mathbb{P}$ ,
- (b) a.e.,
- (c) in distribution,
- (d) Show that if  $X_i \in L^p$  for some p > 1, then it converges in  $L^q$  for every  $q \in [1 \le p)$ .
- 2. Infer from the previous results that for

$$\text{Law}(X_1, \dots, X_n) \approx \text{UNI}(\sqrt{n}\mathbb{S}^{n-1})$$

### **Solution Part 1**

**Theorem 1** (Laws of Large Numbers). Let  $(X_i)_{i \in \mathbb{N}}$  be a sequence of i.i.d. random variables such that  $\mathbf{E} X_i = \mu$  for every  $i \in \mathbb{N}$ , and let  $\overline{X_n} = \frac{1}{n} \sum_{i=1}^n X_i$ . Then,

$$\lim_{n\to\infty} \mathbf{P}\{\|\overline{X_n} - \mu\| > \varepsilon\} = 0, \ \forall \varepsilon > 0.$$
 (Weak Law of Large Numbers)

$$\mathbf{P}\{\lim_{n\to\infty}\overline{X_n}\neq\mu\}=0.$$

(Strong Law of Large Numbers)

**Definition 1.1** (Convergence in probability). Let  $(X_n)_{n\in\mathbb{N}}$  be a sequence of random variables. We say that  $X_n$  converges to X in probability i.e.  $X_n \stackrel{p}{\to} X$  when

$$\lim_{n \to \infty} \mathbf{P}\{\|X_n - X\| > \varepsilon\} = 0, \quad \forall \varepsilon > 0.$$

**Definition 1.2** (Convergence almost everywhere). Let  $(X_n)_{n\in\mathbb{N}}$  be a sequence of random variables. We say that  $X_n$  converges to X almost everywhere (or almost surely) i.e.  $X_n \stackrel{a.e.}{\to} X$  when

$$\mathbf{P}\{\lim_{n\to\infty} X_n \neq X\} = 0$$

According to theorem 1 and the previous definitions, since  $(X_i^2)_{n\in\mathbb{N}}$  is still a sequence of i.i.d. random variables,

(a) 
$$\frac{1}{n} \|X_1, \dots, X_n\|_2^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 \xrightarrow{p} \mathbf{E} X^2 = \mathbf{Var} X - \mathbf{E} X = 1$$

(b) 
$$\frac{1}{n} \|X_1, \dots, X_n\|_2^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 \xrightarrow{a.e.} \mathbf{E} X^2 = \mathbf{Var} X - \mathbf{E} X = 1$$

Therefore,

$$||X_1, \dots, X_n||_2 - \sqrt{n} \xrightarrow{p} 0, \qquad ||X_1, \dots, X_n||_2 - \sqrt{n} \xrightarrow{a.e.} 0.$$

**Definition 1.3** (Convergence of distribution). Let  $(X_n)_{n\in\mathbb{N}}$  be a sequence of random variables with probability distributions  $P_n$ . Let X a random variable with a probability distribution P. We say that  $X_n$  converges to X in distribution i.e.  $X_n \stackrel{d}{\to} X$  if

$$\lim_{n \to \infty} \mathbf{E}\left[f(X_n)\right] = \mathbf{E}\left[f(X)\right]$$

for every bounded and continuous function  $f: \mathcal{X} \to \mathbb{R}$ .

**Theorem 2.** A direct consequence of Fatou's Lemma and Dominated Convergence is that,

$$X_n \xrightarrow{a.e.} X \implies X_n \xrightarrow{p} X \implies X_n \xrightarrow{d} X.$$

Thus, this proves that

(c) 
$$||X_1, \dots, X_n||_2 - \sqrt{n} \stackrel{d}{\longrightarrow} 0.$$

**Definition 2.1** (Convergence in  $L_p$ ). Let  $(X_n)_{n\in\mathbb{N}}$  be a sequence of random variables. For some  $p\in[1,\infty)$ , we say that  $X_n$  converges to X in  $L_p$  norm i.e.  $X_n\stackrel{L_p}{\to} X$  if  $\mathbf{E}|X_n|^p$  and  $\mathbf{E}|X|^p$  exist, and

$$\lim_{n \to \infty} \mathbf{E} |X_n - X|^p = 0.$$

for every bounded and continuous function  $f: \mathcal{X} \to \mathbb{R}$ .

**Theorem 3** (Exercise 1.d.). Convergence in  $L_p$  implies convergence in  $L_q$  for every  $q \in [1, p)$ .

*Proof.* Hölder inequality states that for  $a, b \in [1, \infty]$  such that  $a^{-1} + b^{-1} = 1$  and random variables A, B,

$$||AB||_1 = \mathbf{E} |AB| \le (\mathbf{E} |A|^a)^{a^{-1}} (\mathbf{E} |B|^b)^{b^{-1}} = ||A||_a ||B||_b.$$

By letting

$$A = |X_n - X|^q, \quad B = 1,$$
  
$$a = \frac{p}{q}, \qquad b = \frac{p}{p - q},$$

we obtain,

$$\mathbf{E} |X_n - X|^q = \mathbf{E} ||X_n - X|^q|$$

$$\leq (\mathbf{E} |X_n - X|^{q \cdot p/q})^{q/p} \cdot (\mathbf{E} |1|^{p/(p-q)})^{(p-q)/p}$$

$$= (\mathbf{E} |X_n - X|^p)^{q/p}.$$

The hypothesis of  $q \in [1, p)$  is used on the fact that, if  $q \ge p$ , then  $p - q \le 0$ , and if q < 1, then the  $L_q$  norm wouldn't be defined. Finally, since

$$\lim_{n \to \infty} \mathbf{E} |X_n - X|^p = 0,$$

it follows that

$$\lim_{n\to\infty} \mathbf{E} |X_n - X|^q \le \lim_{n\to\infty} (\mathbf{E} |X_n - X|^p)^{q/p} = 0.$$

#### **Solution Part 2**

In the first place note that the function  $\|\cdot\|_2$  is invariant under rotations because for any orthogonal (rotation) matrix  $O \in \{M : M^T M = I\}$ ,

$$||OX||_2 = \langle OX, OX \rangle^{1/2}$$

$$= \langle X, O^T OX \rangle^{1/2}$$

$$= \langle X, X \rangle^{1/2}$$

$$= ||X||_2$$

Therefore, for any vector  $Z_n \sim \mathcal{N}(0, I_n)$  the density formula, which only depends on  $||Z_n||_2$ , is invariant under rotations. Thus, for any  $n \times n$  rotation matrix O and measurable set  $A \subset \mathbb{R}^2$ ,

$$P(Z_n \in A) = P(OZ_n \in A).$$

Then, by normalizing everything, we would obtain that for every measurable subset of the (n-1)-sphere  $A \subset \mathbb{S}^{n-1}$ ,

$$P(Z_n/\|Z_n\|_2 \in A) = P(OZ_n/\|Z_n\|_2 \in A) = P(OZ_n/\|OZ_n\|_2 \in A).$$

So it follows that  $Z_n/\|Z_n\|_2$  is uniformly distributed over the (n-1)-sphere.

On the other hand, according to the Central Limit Theorem,

$$\overline{Z_n} = \frac{1}{\sqrt{n}} \sum_{i=1}^n (X_i - \mathbf{E} X_i) / (\sqrt{\mathbf{Var} X_i}) \stackrel{d}{\longrightarrow} Z_n.$$

According to the exercise's statement,  $\mathbf{E} X_i = 0$  and  $\mathbf{Var} X_i = 1$ . Thus,

$$\overline{Z_n} = \frac{\sqrt{n}}{n} \sum_{i=1}^n X_i = \sqrt{n} \cdot \overline{X_n}.$$

Also, using the a.e. convergence from the part 1, we can make the following approximation,

$$||Z_n||_2 \approx \sqrt{n}$$
.

Finally, it follows that

$$\overline{X_n} = \frac{\overline{Z_n}}{\sqrt{n}} \approx \frac{\overline{Z_n}}{\|Z_n\|_2} \xrightarrow{d} \frac{Z_n}{\|Z_n\|_2},$$

which will let us conclude that the distribution of  $\overline{X_n}$  which is Law $(X_1, \ldots, X_n)$  might be similar to a uniform distribution on a (n-1)-sphere.

## Exercise 2

Show that for every random variable  $X \in L^2$ ,

$$\mathbf{E} [|X - \mathbf{E} X|^2] \le \mathbf{E} [|X|^2].$$

#### Solution

Since  $|x|^2 = x^2$ , it follows that,

$$\mathbf{E} [ |X - \mathbf{E} X|^2 ] = \mathbf{E} (X - \mathbf{E} X)^2$$

$$= \mathbf{E} [X^2] - E[X]^2$$

$$= \mathbf{E} [ |X|^2 ] - E[X]^2$$

$$\leq \mathbf{E} [ |X|^2 ].$$

### Exercise 3

Let  $f:[0,1]\to\mathbb{R}$  be a continuous function. Show with the Law of Large Numbers that there exists a sequence of polynomial  $(P_n)_{n\in\mathbb{N}}$  such that  $\deg(P_n)=n$  and,

$$\lim_{n \to \infty} \sup_{x \in [0,1]} |f(x) - P_n(x)| = 0.$$

#### Solution

The goal of this solution is to prove the Stone-Weierstrass approximation theorem using the Law of Large Numbers. I will use Hoeffding's inequality in a similar fashion its used in the Glivenko-Cantelli theorem's proof.

In the first place, let  $(X_n)_{n\in\mathbb{N}}$  be a sequence of random i.i.d. Bernoulli variables, each with probability  $x\in[0,1]$ . Define  $\overline{X_n}=\sum_{i=1}^n X_i$ , and note that  $\overline{X_n}\sim \mathrm{Bi}(n,x)$ . The polynomial we are going to use in this proof is the Bernstein polynomial:

$$P_n(x) = \sum_{k=0}^n \binom{n}{k} x^k (1-x)^{n-k} f\left(\frac{k}{n}\right) = \mathbf{E}\left[f\left(\frac{\overline{X_n}}{n}\right)\right].$$

Now, in conjunction with the Heine-Cantor theorem, we use the fact that f is a continuous on a compact set to say that f is uniformly continuous. Therefore, for  $\varepsilon/2 > 0$  there exists  $\delta > 0$  such that

$$|f(x) - f(y)| < \varepsilon/2$$
, whenever  $|x - y| \le \delta$ ,  $\forall x, y \in [0, 1]$ .

We rewrite our expression as follows,

$$|f(x) - P_n(x)| = \left| f(x) - \mathbf{E} \left[ f\left(\frac{\overline{X_n}}{n}\right) \right] \right|$$
$$= \left| \mathbf{E} \left[ f(x) - f\left(\frac{\overline{X_n}}{n}\right) \right] \right|.$$

Then, we separate the expression in two cases.

$$1 = \mathbb{1}\left\{|x - \frac{\overline{X_n}}{n}| \le \delta\right\}(x) + \mathbb{1}\left\{|x - \frac{\overline{X_n}}{n}| > \delta\right\}(x).$$

For the case  $|x - \frac{\overline{X_n}}{n}| \le \delta$ , use uniform continuity:

$$\forall n \in \mathbb{N}: \mathbf{E}\left[\left(f(x) - f\left(\frac{\overline{X_n}}{n}\right)\right) \cdot \mathbb{1}\left\{|x - \frac{\overline{X_n}}{n}| \le \delta\right\}\right] < \frac{\varepsilon}{2}.$$
 (1)

For the case  $|x - \frac{\overline{X_n}}{n}| > \delta$ , define  $M = ||f||_{\infty} = \max_{x \in [0,1]} |f(x)|$ ,

$$\begin{aligned} \forall n \in \mathbb{N} : \mathbf{E} \left[ \left( f(x) - f\left( \frac{\overline{X_n}}{n} \right) \right) \cdot \mathbb{1} \left\{ |x - \frac{\overline{X_n}}{n}| > \delta \right\} \right] \\ & \leq 2M \cdot \mathbf{E} \left[ \mathbb{1} \left\{ |x - \frac{\overline{X_n}}{n}| > \delta \right\} \right] \\ & = 2M \cdot \mathbf{P} \left\{ |x - \frac{\overline{X_n}}{n}| > \delta \right\}. \end{aligned}$$

Since  $x = \mathbf{E} X_n$  for every  $n \in \mathbb{N}$ , we can use Hoeffding inequality to assert that

$$\forall x \in [0,1]: \mathbf{P}\left\{|x - \frac{\overline{X_n}}{n}| > \delta\right\} \le 2e^{-2n\delta^2}.$$
 (2)

Then, choose  $n_0$  big enough to satisfy  $2e^{-2n\delta^2} \leq \frac{\varepsilon}{4M}$  for every  $n > n_0$ . Finally, combine (1), (2) to obtain,

$$\forall \varepsilon > 0, \ \exists n_0: \ |f(x) - P_n(x)| < \frac{\varepsilon}{2} + \frac{2M\varepsilon}{4M} = \varepsilon, \ \forall x \in [0, 1], \ \forall n > n_0.$$

Thus,

$$\sup_{x \in [0,1]} |f(x) - P_n(x)| < \varepsilon, \ \forall n > n_0.$$

This implies that

$$\lim_{n \to \infty} \sup_{x \in [0,1]} |f(x) - P_n(x)| = 0.$$