Housing Bubbles and Support for Incumbents

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August 21, 2016

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Abstract

Abstract here

1 Introduction

In this article we examine whether local changes in property values play a part in shaping the electoral success of incumbents. Our focus is on Danish Parliamentary elections between 2005 and 2015. A period in which the Danish real-estate market experienced a dramatic boom and bust, the extent of which was largely driven by policies which deregulated the housing market (Dam et al., 2011). Following the literature on economic voting (Healy and Malhotra, 2013; Lewis-Beck and Stegmaier, 2013), we want to examine whether voters held governing parties electorally accountable for how the housing bubble played out in their local context.

We do this using two complementary empirical approaches. First, we link detailed registry data on local housing prices to election results at the precinct-level across five national elections, allowing us to study whether within-district differences in property values are related to changes in support for governing parties. Second, to test the hypothesized causal mechanism, that voters are able to make inferences about government based on the state of their local housing market, we zoom in on individual voters' local contexts. Specifically, we link a two-period panel survey to uniquely detailed data from the Danish administrative registries, which allows for precise measures of how individuals' neighborhoods –measured at very low levels of aggregation– were affected by changes in house prices.

Analyzing these data, we find that changes in housing prices do leads to a change in support for governing parties; a relationship which is consistent across precinct-level and individual-level data. In particular we find that support for governing parties is 3-5 pct. higher in contexts where the price of property has increased 50 pct the last year than it is in contexts where the price of property has decreased 50 pct.

We are not the first to investigate, whether voters might draw inferences about policy outcomes from local economic conditions. A number of studies have examined the extent to which voters draw inferences about national economic conditions from local economic conditions (Books and Prysby, 1999; Reeves and Gimpel, 2012; Anderson and Roy, 2011; Ansolabehere et al., 2014; Bisgaard et al., 2015), and a number of studies have

examined the extent to which voters draw inferences about whether to support incumbent politicians (Hansford and Gomez, 2015; Eisenberg and Ketcham, 2004; Kim et al., 2003; Healy and Lenz, 2014). The results from these studies are somewhat mixed, but on balance they find that voters do make inferences based on local economic conditions; asserting that the national economy declining or that incumbent politicians are doing a bad job when local economic conditions are declining.

Our study adds to these literature in two ways. It does so by examining a new type of local economic condition: property values. Compared to other features of the economy, the quality and status of one's home has received scant attention in extant literature on economic voting. A small literature exists on patrimonial economic voting (Nadeau et al., 2010; Stubager et al., 2013), that is the extent to which owning assets, like real estate, makes it more likely that you will vote for right-wing parties (see Ansell, 2014, for a similar argument). However, very few studies and have focused on whether housing prices, similarly to other economic indicators like unemployment and GDP per capita, influence electoral support for governing parties (e.g. Hopkins and Pettingill, 2015), and no studies have looked at how local differences in property values affect economic conditions.

We also adds to the existing literature by addressing some methodological short-comings with previous studies. First, previous studies have generally relied on rather large geographical units (e.g. US counties) when estimating the effects of local economic conditions. This is potentially problematic, as the local context voters react to might not map on to these (typically large) geographical areas. Further, to the extent that these larger geographical units map unto media markets the effect of local economic conditions may be confounded with the effect of mass media communications about these issues (Bisgaard et al., 2015). Second, the studies do not generally take structural differences between local contexts into account when relating economic conditions to attitudes or voting behavior. This is potentially problematic, since it seems likely that voters will at least take some structural factors into account. In the present case, voters are probably not likely to infer much about the government based on the fact that there

are differences in property values between cities and rural areas. They are more likely to infer based on the fact that properties are selling for more (or less) than they used too in their own area. More broadly, if one does not take structural differences into account, one risks conflating re-distributive concerns, i.e. voters in comparatively less well off areas having different demands from government than those in well of areas, with inferential concerns, i.e. the question of what my local contexts tell me about the national economy or about the quality of the government. Third, measures of local economic conditions are often based on samples which, while large enough to estimate precise national economic conditions, are not sufficiently precise on geographical levels (Healy and Lenz, 2014).

It is important to note that some previous studies do address some of these methodological challenges, however, our study contributes by addressing all of these shortcomings at once. We do so by (1) employing data on a very small geographical level of aggregation, (2) using panel data which removes influence of time invariant structural factors, and (3) by using detailed register data on all real estate transactions in the period under investigation.

2 Empirical Setting: A Policy Driven Boom and Bust

This is an excellent setting because:

- 1. Housing bubble was induced by policy. As such, not strange that voters infer something about government based on local changes in house prices.
- 2. We have kick-ass administrative data all house sale registered, we can link these houses to individuals local contexts (e.g zipcodes or survey respondents).
- 3. We witnessed a really volatile bubble there is gonna be a lot of changes in IV.

2.1 Identification strategy

We use a difference-in-difference approach controlling for other economic factors.

Use two different empirical approaches: Precinct level and Individual level data.

3 Precinct-level Evidence

We begin our exploration of the relationship between local housing prices and incumbent support by looking at precinct-level election returns in the Danish Parliamentary elections between 2005 and 2015. In particular, we match support for governing parties in precincts with changes in the prices of houses sold in the proximity of the precinct.

3.1 Data sources and indicators

The key dependent variable in our study is *percent of votes cast for government parties* in each voting precinct.¹ Each voting precinct corresponds to a single polling place and is this the smallest unit at which voting returns can be observed. We measure this for all precincts in five elections: 2001, 2005, 2007, 2011 and 2015. A number of precincts are redistricted between each election. This is problematic, as we want to use the precincts as part of a panel data set. There are two ways to deal with this. We can drop precincts, as their geographical boundaries get altered. This would mean dropping roughly 15 pct. of the data on the dependent variable. The other option is to fix the precincts geographical boundaries at one reference election (i.e. 2015), and then recalculate vote returns in any changed precincts, so they match up with the precincts in the reference election. Since there are a lot of minor changes in geographical boundaries from election to election, but only a few major changes, we opt for the latter, which allows us to keep these slightly altered districts.².

The key independent variable is *change in local housing prices*. We obtain housing price data from The Danish Mortgage Banks' Federation (*Realkreditforeningen*), which publishes quarterly data on the average price per square meters of all sales at the zip code level.³ For each election, we calculate change in housing prices as the percentage

¹Results are similar for only using the prime minister's party.

²For details of how returns from the redistricted precincts are calculated, see Søren Risbjerg Thomsen's research note at bit.ly/2050lPi. We use 2015 as a reference elections

³Available at statistik.realkreditforeningen.dk.

change in the quarter of the election compared to the same quarter one year before.

Zip codes are a substantively interesting level of aggregation when it comes to the price of housing, as it is the level at which house prices are most often reported in Denmark (cf. the fact they are published by The Danish Mortgage Banks' Federation). However, since the dependent variable is observed at the voting precinct level, merging these observations is not trivial. The easiest solution would be to extract the zip code of the address of each polling place and link the polling place to housing prices in that zip code. Unfortunately, full addresses are not available for all polling places. Instead, we take a three-stage approach to linking polling places to zip codes. First, we extract the street address and higher-level voting district of each polling place and add 'Denmark' at the end of the string (the full resulting string is of the format 'Streetname X, City, Denmark'). Second, we pass this string to the Google Maps API, which geocodes the string and returns latitude-longitude coordinates.⁴ Third and last, we pass these coordinates to the Danish Addresses Web API (DAWA), a public service provided by the Danish Geodata Agency.⁵ The DAWA returns the zip code for each address, allowing us to link the two sources of data. Since voting returns are thus statistically speaking nested inside zip codes, we estimate models using standard errors clustered at the zip code level.

We use changes in housing prices rather than the level of housing prices. This is in part because previous literature on economic voting has focused on prices they have focused on changes (i.e. inflation) rather than levels, and in part because changes in housing prices should be more salient than the level. As such, changes in the house price will mean either very short or very long turnaround time for house sales, as sellers and buyers try to adjust to the new prices, leaving visible traces in the immediate context – such as the amount of for sale signs, and stories from neighbors about the speed at which their house was sold.

In addition to the main independent variable, we measure both the unemployment rate and median income at the zip-code level. We also measure the population density

 $^{^4}$ Available at developers.google.com/maps/documentation/geocoding/intro.

⁵Available at dawa.aws.dk.

 $(log(inhabitants/km^2))$ of the municipality in which the precinct is located. These are all population based measures calculated from registers provided by Statistics Denmark.

3.2 Estimating the average effect of housing prices

In table 1 we report estimates from a set of linear regression of electoral support for governing parties using changes in local housing prices as the primary independent variable.

In model 1 we simply look at the bivariate relationship between electoral support and changes in housing prices. In model 2 we include controls for the state of the economy in the precinct. In model 3 we include precinct fixed effects. In model 4 we include year fixed effects; this gives us the difference-in difference model with structural controls. In model 4 we include year by structural factors.

Table 1: Estimated effects of house prices on electoral support for governing parties.

	(1)	(2)	(3)	(4)
Δ housing price	0.104**	0.054**	0.040**	0.029**
	(0.008)	(0.008)	(0.007)	(0.007)
Unemployment rate		-1.571**	-0.489**	-1.917**
		(0.057)	(0.086)	(0.221)
T (3.5.1)		0.050**	0.000**	0.000**
Log(Median income)		-0.259**	-0.333**	-0.889**
		(0.010)	(0.010)	(0.064)
D EE				
Precinct FE			\checkmark	✓
Year FE				√
				<u> </u>
Observations	4194	4174	4174	4174
RMSE	8.411	6.739	5.503	5.325
Ctan dand amount in reason	(1			

Standard errors in parentheses

As can be seen from table 1, there is a statistically significant and positive effect estimate of one-year changes in housing prices, indicating that a larger fraction of the electorate casts their vote for governing parties in precincts where housing prices increase. In the most demanding specification, model 5, the effect is 0.032. This implies that if the price of housing sold in the municipality in the last quarter before the election is twice that of the housing sold in the same quarter the year before, governing parties will get 3.2 percentage points support.

^{*} p < 0.05, ** p < 0.01

3.3 Robustness

Linje 1 - Kontrol for differenced Linje 2 - Differenced afhængig Linje 3 - Lagget afhængig Linje 4/5 - Positive/Negative

Brief discussion of heterogeneity. We find (1) no difference in effects between positive and negative changes, but (2) larger effects in more densely populated areas.

Figure 3

Figure 4

4 Individual-level evidence

We continue our study of how local housing prices shape incumbent support by tracking individuals intent to vote for governing parties in a two wave panel survey collected between 2002 and 2011. We link these individuals to the prices of houses sold in their residential context using the national Danish registers. The registers contain very detailed information about all individuals legally residing in Denmark, including the exact geographical location of their residence, the price of any real estate they sell, and a range of other characteristics (Thygesen et al., 2011). This makes it possible to calculate the distance between the individuals in the survey and all other individuals in Denmark and, in turn, the distance to any individuals who are selling their home.

Before describing these data in more detail, It is important to highlight what we hope to gain from using an additional data set. While it is always desirable to try and replicate findings using a different methodology and a different set of data, linking individual level data to these detailed registers has some advantages over the the precinct-level data used above. The flexibility and detail of the Danish registers makes it possible to look at multiple levels of aggregation – not just the official levels of aggregation such as zip codes. This makes it possible to eliminate concerns related to the modifiable area unit problem (MAUP), in that we can rule out that the findings are linked to a specific way of measuring 'local housing prices'. This approach also makes it is possible to link house prices to individual level characteristics such as attitudes, home ownership etc.

This is important, because we can use these variables to rule out alternative explanations and to explore moderators of the housing price effects, which might, in turn, make it possible to get at the causal mechanism underlying the effect.

4.1 Data Sources and indicators

Our independent variable is once again year-over-year changes in housing prices in the residential context of the respondent. We measure the change by comparing the price of housing sold in the quarter prior to the data collection and the price of housing sold in the same quarter the year before. Unlike for the data used above we do not have data on prices per square meter. This makes our change measure more sensitive to random variation in the types of houses put up for sales in the two time periods we compare. That is, some of the change from year to year might be due to the fact that larger houses were put up for sale. To take this, as well as other structural differences in the type of houses put up for sale, into account we divide the sales price of each house by its publicly valued price, before.⁶ As such, we calculate a price change adjusted for the public valuation of the housing sold in the particular quarter. ⁷

We estimate these changes in house prices within each survey respondents residential context, measuring residential context in three different ways. First, and similar to what we did for the precinct-data, we use the respondents zip-code, comparing houses sold within the same zip code a year apart. Second, we look at the prices of the 20 or 40 units of housing sold closest to the respondents own home, comparing the prices of houses sold in the immediate proximity of the respondent to that of houses sold one year earlier. Third, we look at the price of houses sold within a fixed radius of 500 or 1500 meters of the respondent.

Both of these latter methods are superior to the zip code level in that they do not....

⁶The Danish government makes a conservative estimate of the price of all houses in Denmark every two years which is used to calculate property taxes. The public evaluation was constant across the time periods we use to estimate house price changes.

⁷We make the following exceptions: (1) Sales of part of a house or apartment (10 pct. of all sales). (2) Sales of commercial real estate (9 pct). (3) Sales of apartments or houses valued at more than DKK 10 million (0.2 pct. of all sales) (4) Sales with what 'Statistics Denmark' calls an irregular price (i.e. if the sales price is more than three times the valuation or less than forty percent of the valuation, 6 pct. of all sales).

bla bla - differ in important ways too. One take sales fixed varies size, the other takes size fixed varies sales.

4.2 Average Effect

Table 2 shows the effect og house prices.

We use a linear probability model of support for governing parties. Independent variable is year on year change in house prices measured in different ways. For the first three columns we look at changes in prices for the closest 10, 20 and 40 house sales. For the next three columns we look at changes in the price of houses sold within a radius of 500 and 1500 meters.

In table ?? we use lag house prices as a placebo test.

Table 2: Linear Regression of Voting for Governing party

	10.01	20.61	40 Cl +	F00 1	1000	1500 (
	10 Closest	20 Closest	40 Closest	500 metres	1000 metres	1500 metres
Δ housing prices	0.051^{+}	0.057	0.052	0.059	0.056	0.074
	(0.027)	(0.038)	(0.045)	(0.050)	(0.044)	(0.048)
Unemployment rate	0.044	0.052	0.059	-0.507	0.038	0.848^{+}
	(0.290)	(0.290)	(0.291)	(0.458)	(0.494)	(0.489)
Average income	-0.004	-0.004	-0.004	-0.004	-0.006	-0.006
	(0.003)	(0.003)	(0.003)	(0.004)	(0.006)	(0.006)
Personal income	-0.000	-0.000	-0.000	-0.001*	-0.000	-0.000
	(0.000)	(0.000)	(0.000)	(0.000)	(0.001)	(0.000)
Unnemployed (household)	-0.033	-0.032	-0.032	-0.083*	-0.051	-0.042
	(0.035)	(0.035)	(0.035)	(0.042)	(0.038)	(0.037)
Round FE	Yes	Yes	Yes	Yes	Yes	Yes
Voter FE	Yes	Yes	Yes	Yes	Yes	Yes
Observations	3479	3479	3479	2851	3178	3318

Standard errors in parentheses

4.3 Comparing effect sizes

These effects are not all significant, vut point in the rigt direction. Consistent. IN figure x we compare with b

 $^{^{+}}$ p < 0.1, * p < 0.05

Here we will have figure comparing effect sizes for precinct level and individual level data.

4.4 The Causal mechanism

Ideology, Self Interest or Inference

- (1) Home-ownership
- (2) Moving
- (3) Ideology

Ansell (2014) and others find that house prices affect ideology. In our study the DV is primarily a right-wing government. Can the effects we find be explained by house prices affecting government support thorugh ideologY?

Tables ?? and ?? examine whether this is the case substituting incumbent support for a measure of voters ideological orientation. For precinct level data this is net support for right wing government, for individual level data this is self placement on ideological left right scale. We find no

Table 3: Estimated effects of house prices on net electoral support for right wing government parties.

	(1)	(2)	(3)	(4)
Δ housing price	0.076**	-0.011**	-0.015**	-0.004
	(0.008)	(0.004)	(0.004)	(0.004)
Unemployment rate		-1.971**	-1.690**	-0.779**
		(0.055)	(0.058)	(0.148)
Median income		-58.158**	-59.509**	-39.723**
Median income				
		(0.791)	(0.822)	(4.993)
Precinct FE			\checkmark	\checkmark
Voor EE				/
Year FE				√
Observations	4191	4171	4171	4171
RMSE	7.742	3.680	3.073	3.014

Standard errors in parentheses

^{*} p < 0.05, ** p < 0.01

Table 4: Linear Regression of placement on Left/Right ideological scale

st 20 Closest 0.039 (0.188)	40 Closest 0.021 (0.258)	500 metres -0.292 (0.231)	0.133 (0.256)	1500 metres 0.045 (0.255)
(0.188)				
, ,	(0.258)	(0.231)	(0.256)	(0.255)
0.114				(0.233)
0.111				
0.114	0.176	-3.582+	-2.983	-2.992
(1.151)	(1.156)	(2.131)	(2.344)	(2.528)
0.003	0.003	-0.000	0.010	0.021
(0.010)	(0.010)	(0.015)	(0.021)	(0.022)
0.001	0.001	0.003*	0.003*	0.003*
(0.001)	(0.001)	(0.001)	(0.001)	(0.001)
	-0.204		-0.200	-0.227
(0.204)	(0.203)	(0.246)	(0.232)	(0.216)
Yes	Yes	Yes	Yes	Yes
37	37	37	3/	37
				Yes
3301	3301	2702	3013	3148
	0.003 (0.010) 0.001 (0.001) -0.208 (0.204) Yes	(1.151) (1.156) 0.003	(1.151) (1.156) (2.131) 0.003	(1.151) (1.156) (2.131) (2.344) 0.003

Standard errors in parentheses $^+$ p < 0.1, * p < 0.05

Conclusion

Conclusion

References

- Anderson, C. D. and Roy, J. (2011). Local economies and national economic evaluations. *Electoral Studies*, 30(4):795–803.
- Ansell, B. (2014). The political economy of ownership: Housing markets and the welfare state. *American Political Science Review*, 108(02):383–402.
- Ansolabehere, S., Meredith, M., and Snowberg, E. (2014). Mecro-economic voting: Local information and micro-perceptions of the macro-economy. *Economics & Politics*, 26(3):380–410.
- Bisgaard, M., Dinesen, P. T., and Sønderskov, K. M. (2015). Reconsidering the neighborhood effect: Does exposure to residential unemployment influence voters' perceptions of the national economy? *Working paper*.
- Books, J. and Prysby, C. (1999). Contextual effects on retrospective economic evaluations the impact of the state and local economy. *Political Behavior*, 21(1):1–16.
- Dam, N. A., Hvolbøl, T. S., Pedersen, E. H., and Birch, P. (2011). The housing bubble that burst: Can house prices be explained? and can their fluctuations be dampened? *Danmarks Nationalbank Monetary Review 1st Quarter Part*.
- Eisenberg, D. and Ketcham, J. (2004). Economic voting in us presidential elections: Who blames whom for what. *Topics in Economic Analysis & Policy*, 4(1).
- Hansford, T. G. and Gomez, B. T. (2015). Reevaluating the sociotropic economic voting hypothesis. *Electoral Studies*, 39:15–25.
- Healy, A. and Lenz, G. S. (2014). Presidential voting and the local economy.
- Healy, A. and Malhotra, N. (2013). Retrospective voting reconsidered. *Annual Review of Political Science*, 16:285–306.
- Hopkins, D. J. and Pettingill, L. M. (2015). Economic voting in big-city us mayoral elections. *Available at SSRN* 2567294.
- Kim, J., Elliott, E., and Wang, D.-M. (2003). A spatial analysis of county-level outcomes in us presidential elections: 1988–2000. *Electoral Studies*, 22(4):741–761.
- Lewis-Beck, M. S. and Stegmaier, M. (2013). The vp-function revisited: a survey of the literature on vote and popularity functions after over 40 years. *Public Choice*, 157(3-4):367–385.
- Nadeau, R., Foucault, M., and Lewis-Beck, M. S. (2010). Patrimonial economic voting: Legislative elections in france. *West European Politics*, 33(6):1261–1277.
- Reeves, A. and Gimpel, J. G. (2012). Ecologies of unease: Geographic context and national economic evaluations. *Political Behavior*, 34(3):507–534.
- Stubager, R., Lewis-Beck, M. S., and Nadeau, R. (2013). Reaching for profit in the welfare state: Patrimonial economic voting in denmark. *Electoral Studies*, 32(3):438–444.

Thygesen, L. C., Daasnes, C., Thaulow, I., and Brønnum-Hansen, H. (2011). Introduction to danish (nationwide) registers on health and social issues: structure, access, legislation, and archiving. *Scandinavian Journal of Public Health*, 39(7 suppl):12–16.

Supplementary materials

S1:Descriptive statistics

Table of descriptive statistics

S2: Common trends in precinct-level data

In table 5 we look at whether housing prices can predict changes in support for governing parties in the last period.

Table 5: Estimated effects of house prices on electoral support for governing parties at t-1.

	(1)	(2)	(3)	(4)
Δ housing price (lag DV)	-0.021**	-0.034**	-0.029**	0.008**
	(0.007)	(0.004)	(0.003)	(0.003)
Precinct FE			✓	✓
Year FE				\checkmark
Observations	3099	3089	3089	3089
RMSE	4.430	2.290	1.705	1.466

Standard errors in parentheses p < 0.05, ** p < 0.01

S3: Alternative estimation in individual-level data

In table 6 we estimate a conditional logit model on the panel data. We find similar effects as in linear model above.

 Table 6: Conditional Logit Model of Voting for Governing party

	10 Closest	20 Closest	40 Closest	500 metres	1000 metres	1500 metres
Δ housing prices	0.687	0.631	1.015	0.557	0.752	0.837
.	(0.431)	(0.559)	(0.766)	(0.627)	(0.871)	(0.962)
Unemployment rate	-1.950	-2.044	-2.109	-9.478	2.381	15.116^{+}
	(3.126)	(3.142)	(3.068)	(6.613)	(6.747)	(7.775)
	0.040	0.005	0.006	2.246	0.055	0.004
Average income	-0.040	-0.035	-0.036	-0.046	-0.055	-0.081
	(0.040)	(0.039)	(0.039)	(0.071)	(0.059)	(0.075)
Danaan al in aanaa	0.001	0.001	0.001	0.000	0.002	0.002
Personal income	-0.001	-0.001	-0.001	-0.009	-0.002	-0.003
	(0.004)	(0.004)	(0.004)	(0.017)	(0.004)	(0.005)
Years of Education	0.042	0.030	0.017	0.128	0.023	0.027
rears or Education	(0.171)	(0.170)	(0.171)	(0.206)	(0.210)	(0.218)
	(0.171)	(0.170)	(0.171)	(0.200)	(0.210)	(0.210)
Round FE	Yes	Yes	Yes	Yes	Yes	Yes
Voter FE	Yes	Yes	Yes	Yes	Yes	Yes
Observations	562	562	562	420	504	528

Standard errors in parentheses $^+$ p < 0.1, * p < 0.05

S3: Heterogeneous effects in precinct-level data

Tables 7 and 8 examines the heterogeneity of the effects in the precinct level data.

Table 7: Estimated effects of house prices on electoral support for governing parties across positive and negative changes.

(1)	(2)	(3)	(4)
-0.081***	-0.072***	-0.057**	-0.030
(0.022)	(0.017)	(0.019)	(0.019)
0.116***	0.045***	0.031**	0.029*
(0.012)	(0.011)	(0.011)	(0.011)
		√	√
			\checkmark
0.25	0.25	0.32	0.96
4194	4174	4174	4174
8.42	6.74	5.50	5.33
	-0.081*** (0.022) 0.116*** (0.012) 0.25 4194	-0.081*** -0.072*** (0.022) (0.017) 0.116*** 0.045*** (0.012) (0.011) 0.25 0.25 4194 4174	$\begin{array}{cccccccccccccccccccccccccccccccccccc$

Standard errors in parentheses

Table 8: Estimated effects of house prices on electoral support for governing parties across volatility.

	(1)	(2)	(3)	(4)
Δ housing price	-0.01	-0.20**	-0.18**	-0.17**
	(0.03)	(0.02)	(0.02)	(0.03)
I and donates)	E 40**	2 (0**	0.00	0.00
Log(density)	-5.49**	-2.69**	0.00	0.00
	(0.37)	(0.41)	(.)	(.)
Δ housing price \times Log(density)	0.05**	0.12**	0.10**	0.10**
△ Housing price × Log(density)				00
	(0.01)	(0.01)	(0.01)	(0.01)
Precinct FE			\checkmark	\checkmark
Year FE				_
Teal FE				· ·
Observations	4191	4171	4171	4171
RMSE	8.43	6.81	5.50	5.40

Standard errors in parentheses

^{*} p < 0.05, ** p < 0.01, *** p < 0.001

^{*} p < 0.05, ** p < 0.01

S4: Heterogeneous treatment effects in individual-level data

Tables 9 and 10 examines the heterogeneity of the effects in the individual level data.

Table 9: Linear Regression of Voting for Governing party

	10 Closest	20 Closest	40 Closest	500 metres	1000 metres	1500 metres
Δ housing prices (positive)	-0.020	-0.037	0.034	-0.053	-0.005	-0.021
	(0.062)	(0.072)	(0.083)	(0.095)	(0.088)	(0.089)
Δ housing prices (negative)	0.070^{+}	0.071	0.137^{+}	0.063	0.101	0.129^{+}
	(0.041)	(0.065)	(0.075)	(0.074)	(0.063)	(0.075)
Unemployment rate	0.040	0.050	0.044	-0.527	0.035	0.847^{+}
	(0.291)	(0.291)	(0.292)	(0.462)	(0.495)	(0.490)
Average income	-0.004	-0.004	-0.004	-0.004	-0.006	-0.006
	(0.003)	(0.003)	(0.003)	(0.004)	(0.006)	(0.006)
D 1.	0.000	2.222	0.000	0.004*	2.222	0.000
Personal income	-0.000	-0.000	-0.000	-0.001*	-0.000	-0.000
	(0.000)	(0.000)	(0.000)	(0.000)	(0.001)	(0.000)
11 1 1/1 1 11	0.022	0.021	0.021	0.001+	0.050	0.041
Unnemployed (household)	-0.032	-0.031	-0.031	-0.081 ⁺	-0.050	-0.041
	(0.035)	(0.035)	(0.035)	(0.041)	(0.038)	(0.037)
Davis J.EE	Vaa	Vaa	Vaa	Van	Vaa	Vac
Round FE	Yes	Yes	Yes	Yes	Yes	Yes
Voter FE	Yes	Yes	Yes	Yes	Yes	Yes
Observations	3473	3473	3473	2846	3173	3313
- CDSCI VALIOID	01/0	01/0	01/0	2010	0170	

Standard errors in parentheses $^+$ p < 0.1, * p < 0.05

 Table 10: Linear Regression of Voting for Governing party

	10 Closest	20 Closest	40 Closest	500 metres	1000 metres	15
Δ housing prices	-0.007	-0.084	-0.166	0.014	-0.127	10
2 housing prices	(0.078)	(0.107)	(0.137)	(0.173)	(0.157)	
	(0.070)	(0.107)	(0.157)	(0.173)	(0.157)	
Log(No. of ppl in context)	-0.010	-0.010	-0.010	0.006	-0.001	
	(0.010)	(0.010)	(0.010)	(0.020)	(0.015)	
Δ housing prices \times Log(No. of ppl in context)	0.008	0.021	0.031^{+}	0.006	0.024	
	(0.010)	(0.014)	(0.019)	(0.023)	(0.019)	
TT 1	0.000	0.100	0.114	0.500	0.025	
Unemployment rate	0.098	0.102	0.114	-0.580	0.035	
	(0.297)	(0.295)	(0.293)	(0.484)	(0.557)	
Average income	-0.004	-0.004	-0.004	-0.004	-0.007	
	(0.003)	(0.003)	(0.003)	(0.004)	(0.006)	
	(01000)	(0.000)	(0.000)	(0.00-)	(0.000)	
Personal income	-0.000	-0.000	-0.000	-0.001*	-0.000	
	(0.000)	(0.000)	(0.000)	(0.000)	(0.001)	
Unnemployed (household)	-0.034	-0.034	-0.034	-0.081^{+}	-0.050	
	(0.035)	(0.035)	(0.035)	(0.042)	(0.038)	
Round FE	Yes	Yes	Yes	Yes	Yes	
ROUNG FE	ies	ies	ies	ies	ies	
Voter FE	Yes	Yes	Yes	Yes	Yes	
Observations	3473	3473	3473	2846	3173	

Standard errors in parentheses $^+$ p < 0.1, * p < 0.05