

08.03.2016

Definition (Random variable (RV)). $X : (\Omega, \mathcal{A}, \mathbb{P}) \rightarrow (\mathbb{R}, \mathcal{B}_{\mathbb{R}})$ where $\mathcal{B}_{\mathbb{R}}$ are Borel sets (think of intervals).

$$X^{-1}(I) \in \mathcal{A}$$

for all interval I (Borel set I).

Definition (Distribution of X). P_X on $(\mathbb{R}, \mathcal{B})$:

$$P_X(I) = \mathbb{P}[X \in I] = \mathbb{P}(X^{-1}(I))$$

Definition (Discrete/continuous RVs (and more!)). 1. discrete density:

$$p_X(x) = \mathbb{P}[X = x] > 0$$

only in the (finite or countable) value set of X .

$$\mathbb{P}[X \in I] = \sum_{x \in I} p_X(x)$$

2. density:

$f_X : \mathbb{R} \rightarrow [0, \infty)$ such that

$$\mathbb{P}[X \in I] = \int_I f_X(x) dx$$

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Definition. The expected value of a RV X is

$$\mathbb{E}(X) = \int_{\Omega} X d\mathbb{P}$$

if the integral exists. In particular, if X is discrete, then

$$\mathbb{E}(X) = \sum_{x \in \mathbb{R}} xp_X(x)$$

If X is continuous, then

$$\mathbb{E}(X) = \int_{\mathbb{R}} xf_X(x) dx$$

We are interested in the situation where

$$\mathbb{E}(|X|) < \infty,$$

that is:

$$\sum |x|p_X(x) < \infty \text{ (discrete case)}$$

$$\int_{\mathbb{R}} f_X(x) dx < \infty \text{ (continuous case)}$$

Remark (for Math). $X = X^+ - X^-$: at least one of $\int X^{\pm}$ is finite

Example. Let $p(n) = \frac{6}{\pi^2} \frac{1}{n^2}$ for $n \in \mathbb{N}$.

$$\sum np(n) = \frac{6}{\pi^2} \sum \frac{1}{n} = \infty$$

0.1 Convergence of sequences of RVs

Definition. Let (X_n) be a sequence of random variables and X another random variable. Then

1. $X_n \rightarrow X$ (X_n converges to X) almost surely

$$\begin{aligned}\mathbb{P}\left[\lim_{n \rightarrow \infty} X_n = X\right] &= \mathbb{P}(\{\omega \in \Omega : (X_n(\omega))_{n \in \mathbb{N}} \text{ converges and the limit is } X(\omega)\}) \\ &= 1\end{aligned}$$

2. $X_n \rightarrow X$ in probability
 $\forall \varepsilon > 0$

$$\mathbb{P}[|X_n - X| \geq \varepsilon] = \mathbb{P}(\{\omega : |X_n(\omega) - X(\omega)| \geq \varepsilon\}) \rightarrow 0$$

Question: Is the statement $\mathbb{E}(X_n) \xrightarrow{?} \mathbb{E}(X)$ true?

Theorem 0.1 (Monotone convergence). If $0 \leq X_n \leq X_{n+1}$ and $X = \lim_{n \rightarrow \infty} X_n$ almost surely, then

$$\mathbb{E}(X) = \lim_{n \rightarrow \infty} \mathbb{E}(X_n)$$

Theorem 0.2 (Lemma of Fatou). If $X_n \geq 0$ for all n , then

$$\mathbb{E}(\liminf X_n) \leq \liminf \mathbb{E}(X_n)$$

Theorem 0.3. If $X_n \rightarrow X$ almost surely and there is a random variable Y such that $|X_n| \leq Y$ a.s. for all n and $\mathbb{E}(Y) < \infty$, then

$$\lim \mathbb{E}(X_n) = \mathbb{E}(X).$$

Example. A counter example for the case “If $X_n \rightarrow X$ a.s., then $\lim \mathbb{E}(X_n) = \mathbb{E}(X)$.”:

$\Omega = [0, 1]$, $\mathcal{A} = \mathcal{B}_{[0,1]}$, $\mathbb{P} = \text{lebesgue measure}$. Choose the curve such that the triangle has a small base and a high height. Then $\mathbb{E}(X_n) = 1$, $\lim X_n(\omega) = 0 = X(\omega)$, $\mathbb{E}(X) = 0$.

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Remark. Let f map from $X_n \rightarrow X$ a.s. on $X_n \rightarrow X$ in probability, then

$$\mathbb{P}[X \neq X'] = 0 : X = X' \text{ a.s.}$$

Proof.

$$\begin{aligned}|X - X'| &\leq |X - X_n| + |X_n - X'| \\ \implies [|X - X'| \geq \varepsilon] &\subseteq [|X - X_n| \geq \frac{\varepsilon}{2}] \cup [|X_n - X'| \geq \frac{\varepsilon}{2}] \\ \mathbb{P}[|X - X'| \geq \varepsilon] &\leq \mathbb{P}[|X_n - X| \geq \frac{\varepsilon}{2}] + \mathbb{P}[|X_n - X'| \geq \frac{\varepsilon}{2}] \\ &\xrightarrow{(n \rightarrow \infty)} 0\end{aligned}$$

$$\text{so } \mathbb{P}[|X - X'| \geq \frac{1}{r}] = 0 \text{ for all } r \in \mathbb{N}$$

$$\implies \mathbb{P}\left(\bigcup_{r=1}^{\infty} [|X - X'| \geq \frac{1}{r}]\right) = \mathbb{P}[|X - X'| > 0] = 0$$

□

Theorem 0.4. $X_n \rightarrow X$ a.s. if and only if $U_k \rightarrow 0$ in probability where

$$U_k = \sup_{n \geq k} |X_n - X|$$

In particular, $X_n \rightarrow X$ in prob. (because $U_k \geq |X_k - X|$)

Proof. Let $\varepsilon = \frac{1}{r}$.

$$\begin{aligned} \mathbb{P}[\forall r \exists k \forall n \geq k : |X_n - X| < \frac{1}{r}] &= 1 \\ &= \mathbb{P}(\underbrace{\bigcap_r \bigcup_k \bigcap_{n \geq k} [|X_n - X| < \frac{1}{r}]}_{A_{r+1} \subset A_r \in \mathcal{A}}) \\ \implies \mathbb{P}(A_r) &= 1 \quad \forall r \in \mathbb{N} \end{aligned}$$

Since A_r are decreasing and $\mathbb{P}(A_r) = 1 \quad \forall r \in \mathbb{N}$, the other direction holds as well.

$$\bigcap_{n \geq k} [|X_n - X| < \frac{1}{r}] \subset [U_k \leq \frac{1}{r}]$$

$$\text{Conversely, } [U_k \leq \frac{1}{r}] \subset \bigcap_{n \geq k} [|X_n - X| < \frac{1}{r-1}]$$

$$A_r \subset \bigcup_k [U_k \leq \frac{1}{r}]$$

$$\text{So, } 1 = \mathbb{P}(A_r) \leq \mathbb{P}(\bigcup_k [U_k \leq \frac{1}{r}]) \leq 1$$

Since $U_k \geq U_{k+1}$, the events $U_k \leq \frac{1}{r}$ are increasing in k .

$$\begin{aligned} \iff \forall r \quad \lim_{k \rightarrow \infty} \mathbb{P}[U_k \leq \frac{1}{r}] &= 1 \\ \iff \forall r \quad \lim_{k \rightarrow \infty} \mathbb{P}[U_k > \frac{1}{r}] &= 0 \end{aligned}$$

This means, $U_k \rightarrow 0$ in prob. This concludes the entire proof. \square