

# Statistical Inference

PCHN62121  
Image Analysis

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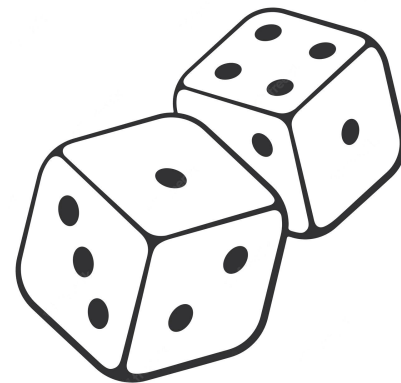
- Statistics lie **at the heart** of everything we will be doing during this module
- Our ability to reach conclusions about our fMRI and M/EEG data depends entirely upon **statistical modelling** and **statistical inference**
- Poldrack, Mumford & Nichols (2011) name *Probability and Statistics* as their **number 1** prerequisite for fMRI data analysis

1. *Probability and statistics*. There is probably no more important foundation for fMRI analysis than a solid background in basic probability and statistics. Without this, nearly all of the concepts that are central to fMRI analysis will be foreign.

- In this session we will review the **fundamentals** of statistical inference to prepare you for the content that is to come on this module

# Probability

- Probability is the foundation of everything in statistics.
- Statistics is the **science of uncertainty** - of reaching conclusions based on **noisy** or **incomplete** information
- Probability is the **language of uncertainty**
- Statistics uses probability to **describe the nature of data** and how we can reach **general conclusions** about a phenomena by examining **a small part of it**
- Probability provides a mechanism for **inductive reasoning** - going from the **specific** to the **general**
  - **Induction** is a big philosophical problem that is not fully resolved - hence why we cannot *prove* anything in science



## Kolmogorov axioms

- Mathematically, for a number to be called a probability it must adhere to some **rules** – known as the Kolmogorov axioms
- Imagine rolling a six-sided die:
  - There are 6 **mutually exclusive** events – the numbers 1 to 6
  - Each event needs to be assigned a number that is  $\geq 0$  (Axiom 1)
  - They cannot all be 0 (Axiom 2)
  - The sum of all the probabilities must be equal to 1 (Axiom 3)
- Therefore, if we believe all the outcomes to be equally probable we can define



$$P(E) = \frac{\text{number of favourable events}}{\text{total number of events}}$$

## Kolmogorov axioms

- The probability of rolling a 5 would be

$$P(5) = \frac{1}{6}$$

- The probability of rolling an even number would be

$$P(2 \cup 4 \cup 6) = \frac{3}{6} = \frac{1}{2}$$

- All these examples satisfy the Kolmogorov axioms and thus can be called probabilities
- Notice that any notion of **what probability means** is completely absent – Kolmogorov tells us how to calculate the numbers, but he does not tell us **what they mean**

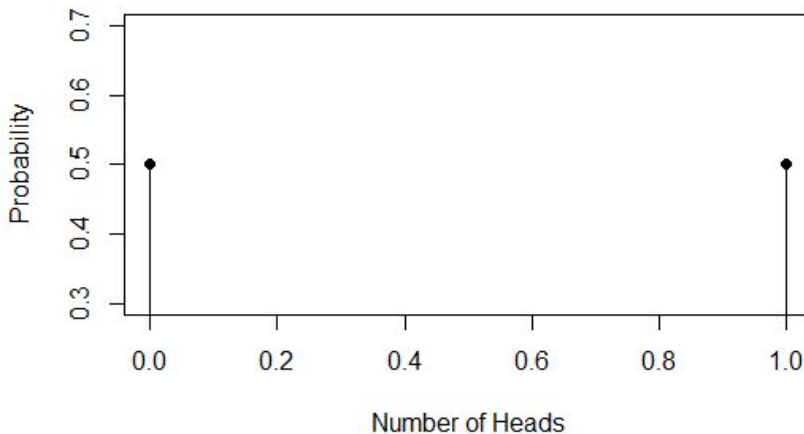
## Interpreting Probability

- One of the greatest divides in modern science between the **Frequentist** and **Bayesian** approaches to statistics
- For the Frequentist, probabilities represent **physical phenomena** that can be **counted**
  - A probability is the **long-run frequency** of an event
- For the Bayesian, probabilities represent **degrees of belief**
  - A probability indicates, based on the available evidence, how likely an event is to occur
- A Bayesian can apply probability to events that cannot be counted (e.g. the probability of rain tomorrow)
- The Bayesian view leads to a **much more flexible analysis framework** – the notion of **degree of belief** has been criticised as **too subjective**
- The development during the 20th century of inferential statistics by **Ronald Fisher** was motivated by his **deep disdain** for the Bayesian perspective on probability

# Random variables

- Irrespective of your philosophical views on interpretation, one of the most important concepts from probability for statistics is the **random variable**
- A random variable is
  - A variable whose value is dependent upon the outcome of some random processes
  - A variable where we will measure a different value every time we observe it
  - A variable where each possible values can be associated with a probability
- A basic example would be the outcome of flipping a coin

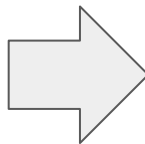
Outcome	Probability
H	$1/2$
T	$1/2$



# Random variables

- Another example would be counting the number of **heads** after 3 flips of a coin

Outcome	Number of Heads
HHH	3
HHT	2
HTH	2
THH	2
HTT	1
THT	1
TTH	1
TTT	0



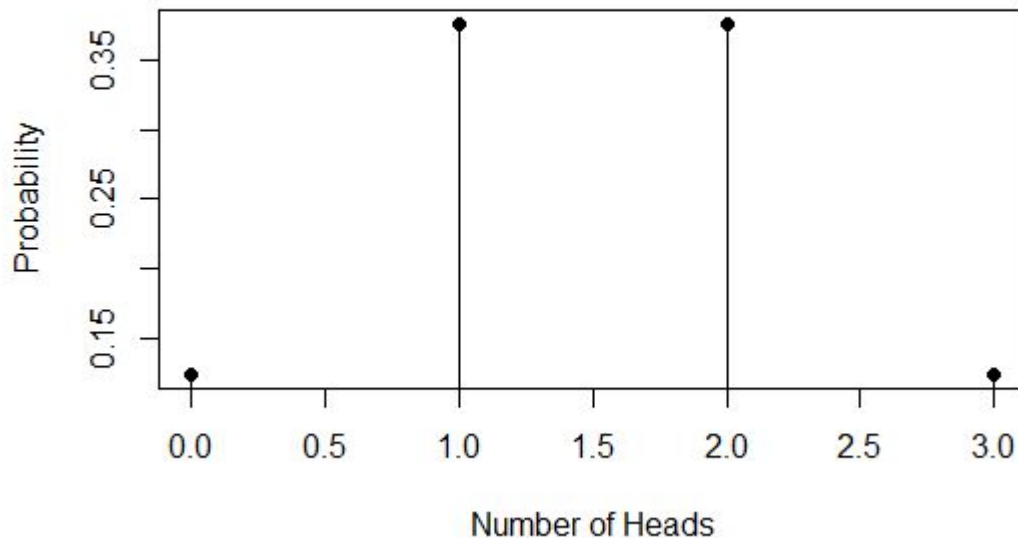
Number of heads	Probability
3	1/8
2	3/8
1	3/8
0	1/8



# Random variables

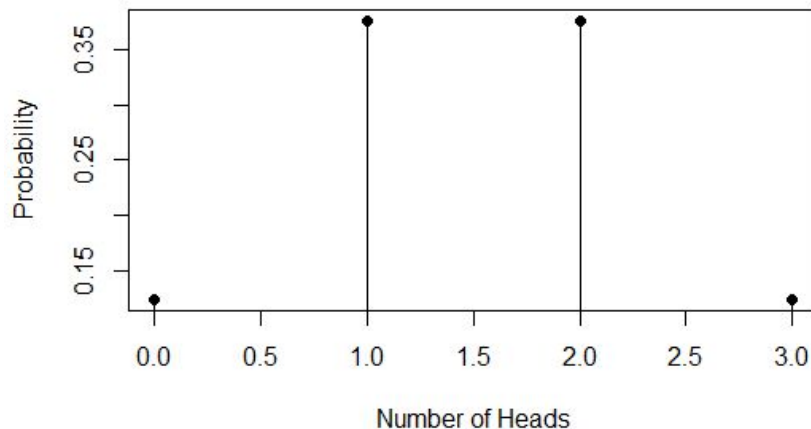
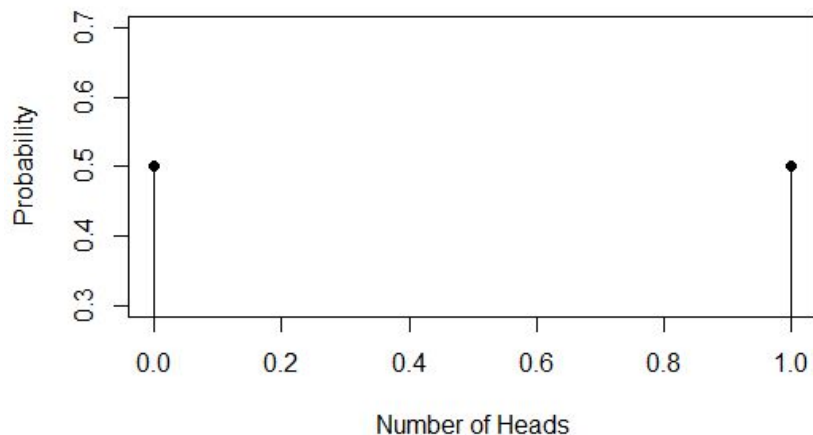
- Another example would be counting the number of **heads** after 3 flips of a coin

Number of heads	Probability
3	$1/8$
2	$3/8$
1	$3/8$
0	$1/8$



# Probability distributions

- These shapes are known as **probability distributions** - they tell us the probability of **all possible values** of the random variable

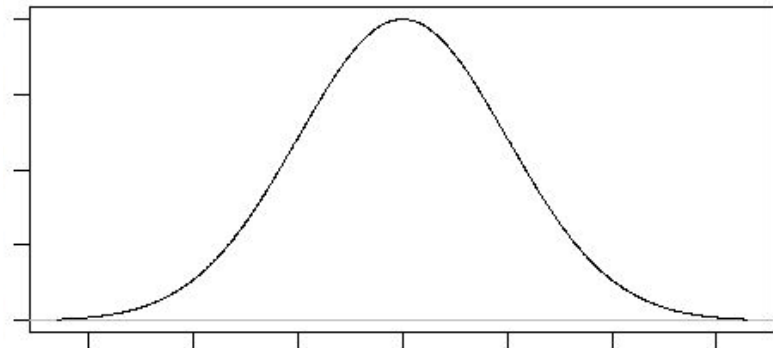


- These shapes are examples of the *binomial distribution*  $y \sim \text{Binomial}(n, p)$
- Each probability distribution is controlled by **parameters** that describe the **shape**
  - $n$  = the number of trials,  $p$  = the probability of success on a single trial

# Probability distributions

## The Normal Distribution

- The binomial distribution is an example of a **discrete** probability distribution because the measurements are **whole numbers**
- In the real world we often deal with random variables that can take on an **infinite** number of possible values
  - Time, height, weight, reaction time, BOLD signal etc.
- In these cases we have to use a **continuous** probability distribution
- Although there are many continuous distributions available, the most commonly used is the **normal distribution**
- Also known as the **Gaussian** distribution



# Probability distributions

## The Normal Distribution

- The normal distribution is fully described by the formula

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

- The important point is that this is parameterised by **two** values:
  - The **mean** ( $\mu$ ) - the centre of the distribution
  - The **standard deviation** ( $\sigma$ ) - the width of the distribution

$$y \sim \mathcal{N}(\mu, \sigma)$$

- If we assume our random variable of interest comes from a normal distribution, our aim is to estimate the **mean** and **standard deviation** and how these **change** under **different experimental conditions**

# Random samples

- Imagine we have an interest in the **weight** of **males** who are suffering from **major depressive disorder** in the **UK**
- **Weight** is a **continuous random variable** with some distribution - if we assume this is a **normal distribution** then

$$\text{weight} \sim \mathcal{N}(\mu, \sigma)$$

- This distribution represents the **entire population** under study
- We want to know the **parameter values** of this distribution – we would need to weight **every male who has major depression in the UK**
- Instead we take a **sample** and use this to **infer** something about the population
- Using a **sample** to say something about the **population distribution** lies at the heart of **parametric** statistical methods

# Random samples

- A **random sample** of size  $n$  from a **population** can be conceptualised as a sequence of  $n$  independent random variables  $(y_1, y_2, y_3, \dots, y_n)$ , where each random variable is drawn from the **same distribution** ( $i = 1, \dots, n$ )

$$y_i \sim \mathcal{N}(\mu, \sigma)$$

- These are known as **independent and identically distributed** (i.i.d.) random variables
- This random sampling model describes an experimental situation where **repeated observations** are made on the **same variable**  $y$
- For our current example, each observation represents the **weight** of a **different subject** in our experiment

# Estimating population parameters

- Our aim is now to use our **random sample** to **estimate** values for the **population mean** and the **population standard deviation**
- It would seem that what we want to calculate is

$$P(\mu, \sigma | y)$$

- This cannot be calculated without using Bayesian methods (which Fisher hated) so classical statistical methods instead use something called the **likelihood**

$$\mathcal{L}(\mu, \sigma | y) = P(y | \mu, \sigma)$$

- An **optimisation algorithm** is used to search through different values of the parameters to find those that **maximise the likelihood**
- In some cases, optimisation is not needed because there are **closed form** solutions to finding the estimates that maximise the likelihood

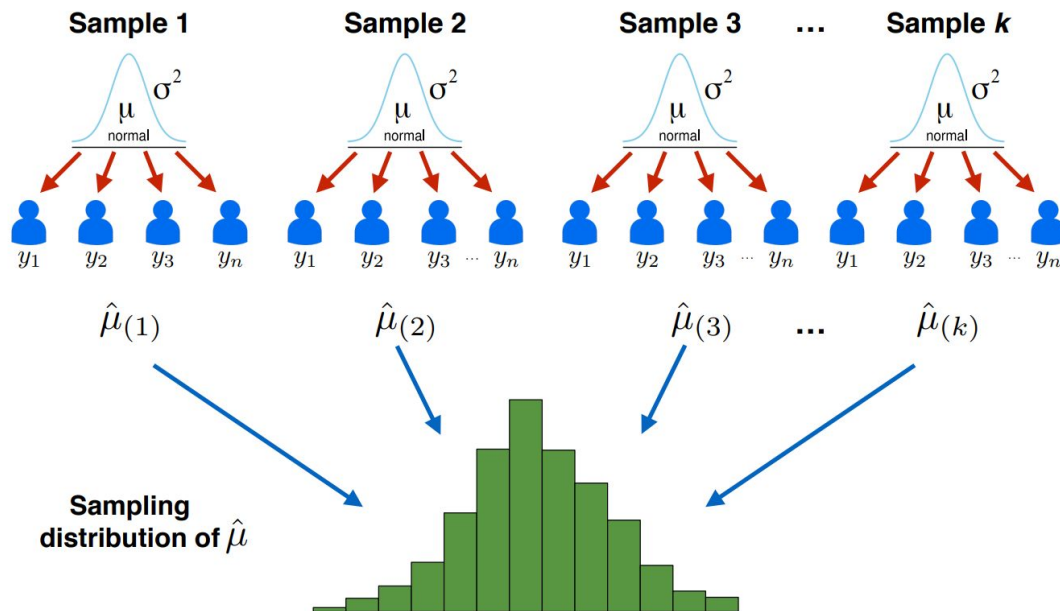
# Uncertainty in the parameter estimates

- So let us take a step back:
  - We have a phenomena of interest characterised as a **random variable** from a **normal population distribution**
  - We want to know the **population parameter values**, but **cannot measure the whole population**
  - We **take a sample** and use the method of **maximum likelihood** to **estimate** the population parameters
  - For a normal distribution, this involves calculating the **sample mean** and **sample standard deviation**
- There is a problem with doing this:
  - What happens if we take a **different sample**? Will we get the same estimates?
  - No! Because a **different sample** will contain **different data** - so which estimates do we use?
  - We need some way of characterising the **uncertainty** in our estimates.



# Uncertainty in the parameter estimates

- The key insight is to recognise that with **each new sample** we will get **different parameter estimates** - our estimates are **also random variables**
- This means they have an associated **probability distribution** – the **sampling distribution**



# Uncertainty in the parameter estimates

- The key insight is to recognise that with **each new sample** we will get **different parameter estimates** - our estimates are **also random variables**
- This means they have an associated **probability distribution** – the **sampling distribution**
- For a **normal population distribution** the sampling distribution of the mean is **also normal**

$$\hat{\mu} \sim \mathcal{N} \left( \mu, \frac{\sigma}{\sqrt{n}} \right)$$

- The **mean of this distribution** is the **true population mean** – on average, we should estimate this correctly across samples
- The **standard deviation** of this distribution depends upon the **sample size** - the more data the more accurate we will be – this is known as the **standard error**

# Hypothesis testing

- By this point we have successfully managed to:
  - Characterise our phenomena of interest as a **random variable** with a **distribution**
  - Use formulas derived from **maximum likelihood** to estimate the **parameters** of this distribution based on a single sample
  - Calculate the **standard error** of these estimates as a means of characterising their **uncertainty**
- So we now have **estimates** and **standard errors** – how do we use these to reach conclusions about the population under study?
- This is where the process of **null hypothesis significance testing** comes in

## Test Statistics

- Trying to draw conclusions based on the **parameter estimates** has two issues:
  - The estimates are on the **same scale** as the data (e.g. weight) so depend upon our **domain knowledge** to interpret
  - The estimates alone do **not** take the **uncertainty** into account
- Both of these issues can be solved by **dividing** the estimate by the standard error

$$t = \frac{\text{estimate}}{\text{standard error}} = \frac{\hat{\mu}}{\sigma \{\hat{\mu}\}}$$

- The quantity  $t$  is now a **standardised** variable – same units irrespective of the data
- The quantity  $t$  contains both the **estimate** and **uncertainty** – the value will increase as the uncertainty decreases

## Test Statistics

- Using  $t$  for **hypothesis testing** involves comparing our estimate with some **hypothesised value** for the population parameter

$$t = \frac{\hat{\mu} - \mu^{H_1}}{\sigma \{\hat{\mu}\}}$$

- The **larger** the value of  $t$ , the greater the **discrepancy** between our estimate and our hypothesised value
- So **big values** of  $t$  suggest that our **hypothesised population** value is **incorrect**
- In this example, the hypothesised value of the mean would depend upon domain knowledge (e.g. average weight of males in the UK)

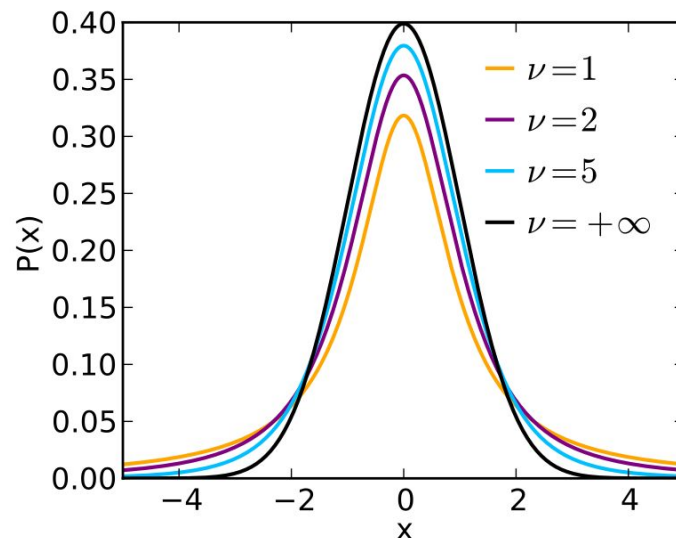
## Null Hypothesis Significance Testing

- The insight that Ronald Fisher provided was that our test should form a **null hypothesis**
- In this instance, it would be there the **difference** between the **true mean** and the **hypothesised mean** is **0** in the population
- To see why this is useful, consider that  $t$  is **also a random variable** because it is calculated from two other random variables
- This means that  $t$  has a **distribution** that can be derived from knowing the sampling distribution of the estimates
- If we assume that the null hypothesis is **true** then the  $t$ -distribution will be centered on **0** with a width that depends upon the sample size

# Hypothesis testing

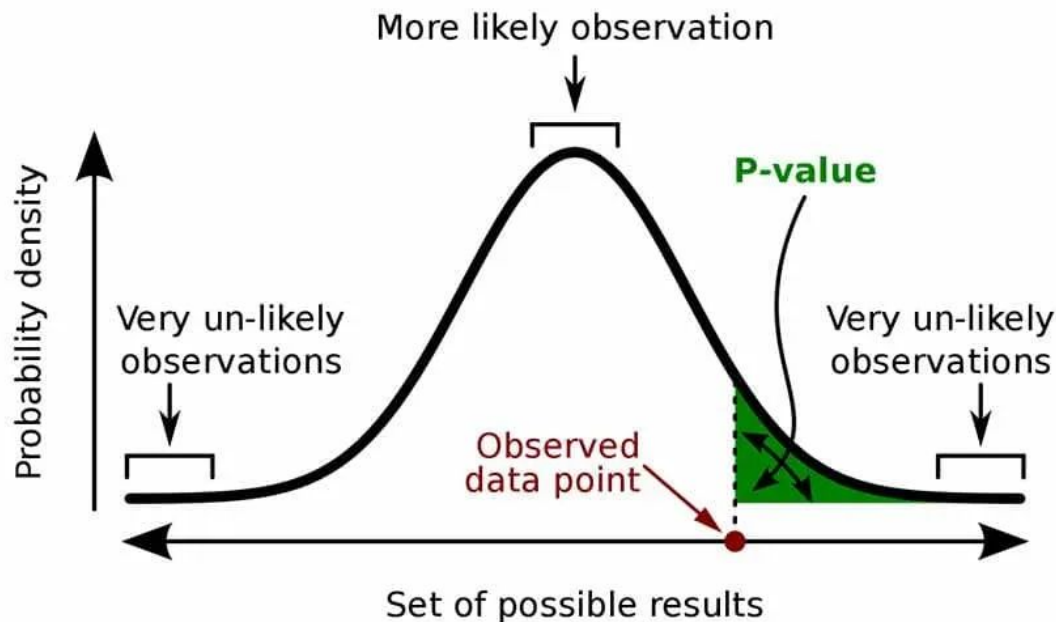
## Null Hypothesis Significance Testing

- This distribution tells us the various values of  $t$  we would expect to calculate **if the null hypothesis were true**
- So what we can do is use this distribution to calculate the **probability** of obtaining our particular value of  $t$
- This gives the  $p$ -value
  - The probability of obtaining a test statistic as larger, or larger, assuming the null hypothesis is true
- A **small**  $p$ -value suggests that our calculated test statistic is **unlikely**, if the null were true – either we observed a **rare event** or the **null hypothesis is not accurate**



# Hypothesis testing

## **P-values**



A **p-value** (shaded green area) is the probability of an observed (or more extreme) result assuming that the null hypothesis is true.



## *P*-values

- How do we use this information?
  - Fisher's recommendation was to count any  $p < 0.05$  as **evidence against the null**
  - In our example, the null was that the population mean is the same as the hypothesised mean (their difference was **0**)
  - If  $p < 0.05$ 
    - We would call this a **significant** result and **reject** the null hypothesis – it is unlikely that the population mean is the same as the hypothesised mean
  - If  $p > 0.05$ 
    - We would call this a **non-significant** result and **fail to reject** the null hypothesis – it is possible that the population mean is the same as the hypothesis mean
- The  $p$ -value is a way of reaching binary conclusions from our results

# Application to multiple groups

## Two sample tests

- To see how this method applies to more complex experiments, consider comparing the weights of depressed individuals taking **two different drugs**

$$y_{ij} \sim \mathcal{N}(\mu_j, \sigma_j)$$

- We now have **two** population distributions

$$y_i^{(\text{Drug A})} \sim \mathcal{N}(\mu^{(\text{Drug A})}, \sigma^{(\text{Drug A})})$$

$$y_i^{(\text{Drug B})} \sim \mathcal{N}(\mu^{(\text{Drug B})}, \sigma^{(\text{Drug B})})$$

- Our aim is still to **estimate** the **parameters** of these distributions – we want to **compare** the means to see whether average weight changes due to the drug

# Application to multiple groups

## Two sample tests

- We use the same procedure as before to estimate the **means** and **standard deviations** of these populations, as well as the **standard errors** of the estimates
- The  $t$ -statistic then involves comparing the **mean difference** to a **hypothesised mean difference** – typically taken to be **0**

$$t = \frac{(\hat{\mu}_1 - \hat{\mu}_2) - D^{H_1}}{\sqrt{\sigma\{\hat{\mu}_1\} + \sigma\{\hat{\mu}_2\}}} = \frac{\hat{\mu}_1 - \hat{\mu}_2}{\sqrt{\sigma\{\hat{\mu}_1\} + \sigma\{\hat{\mu}_2\}}}$$

- We can then use the **same** null  $t$ -distribution to calculate a  $p$ -value to provide evidence for or against the null hypothesis of the population distributions having the **same mean**
- The process is **the same** – assume a population distribution, estimate the parameters from a sample, form a hypothesis test about the parameters, calculate a  $p$ -value

# Application to continuous variables

## Regression models

- We can also use the same framework to reach conclusions about the relationship between our random variable of interest and other continuous measures
- Imagine that we are interested in how the **weight** of our **depressed males** relates to the **severity of their symptoms**
- In this situation, we might start with the normal distribution model

$$\text{weight} \sim \mathcal{N}(\mu, \sigma)$$

- But then specify a more complex form for the mean

$$\mu = \beta_0 + \beta_1 \text{severity}$$

- So the value of the mean depends upon the severity of symptoms

# Application to continuous variables

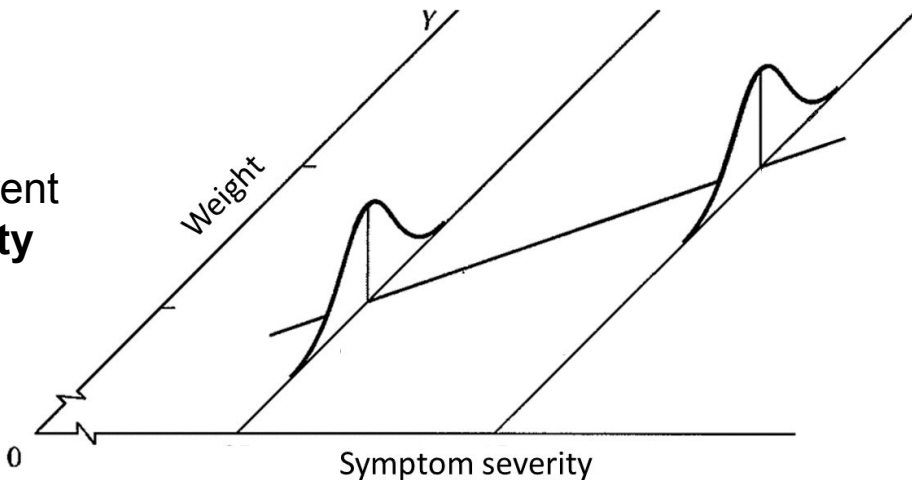
## Regression models

- Assuming a **mean function** of

$$\mu = \beta_0 + \beta_1 \text{severity}$$

is an example of a **linear regression model**

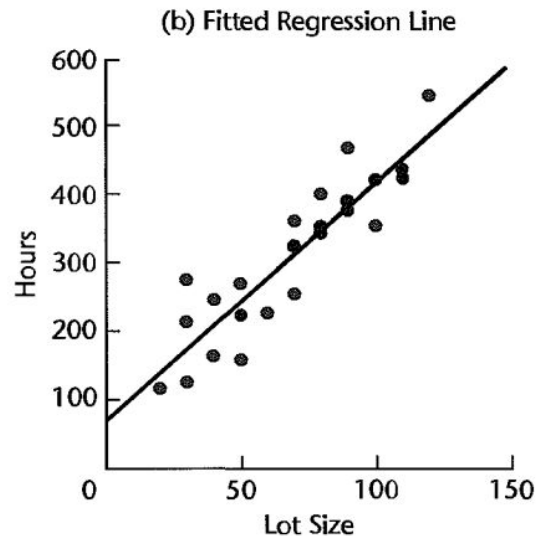
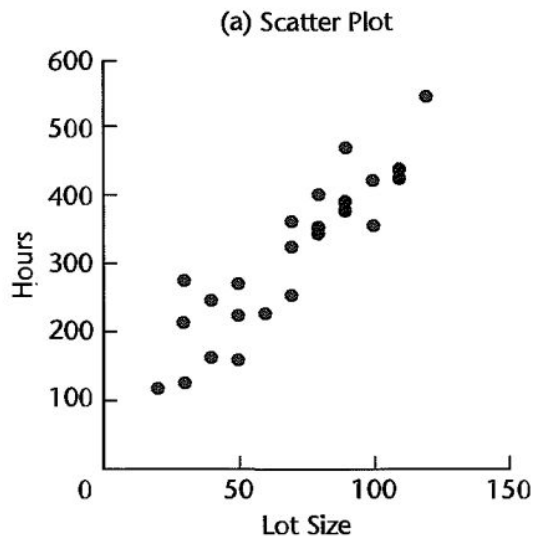
- This assumes that the relationship between **weight** and **symptom severity** is a **straight-line**
  - $\beta_0$  is the **intercept**
  - $\beta_1$  is the **slope**
- The probability model is that there is a different normal distribution for **each value** of **severity**
- The standard deviations are the **same** and the **means** sit along a **straight line** defined by the two parameters



# Application to continuous variables

## Regression models

- In order to estimate the **mean** of our **population distribution** we need to estimate the values of the **intercept** and the **slope** – in this example the mean depends upon two further parameters
- **Maximum likelihood** can do this for us



# Application to continuous variables

## Test Statistics

- We can again calculate a  $t$ -statistic, but this time on the *intercept* and the *slope*

$$t = \frac{\hat{\beta}_1 - \beta_1^{H_0}}{\sigma\{\hat{\beta}_1\}}$$

- The hypothesised value for the slope is usually taken as **0** – **no relationship** between **weight** and **severity**

$$t = \frac{\hat{\beta}_1 - 0}{\sigma\{\hat{\beta}_1\}} = \frac{\hat{\beta}_1}{\sigma\{\hat{\beta}_1\}}$$

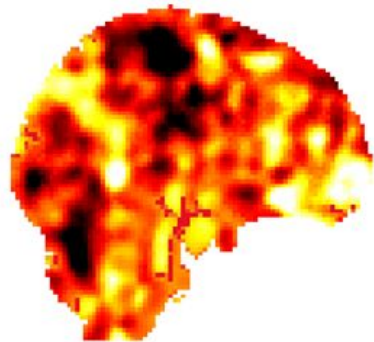
- So this is the same approach as **before** – the only difference is that the **mean function** is more complex – this is the difference between different **statistical models**

- We have now seen the process of **statistical inference**, from **first principles** about probability, all the way up to  **$p$ -values** and **hypothesis testing**
- This is a somewhat complex process:
  - Our data are conceptualised as random variables drawn from a distribution
  - This distribution has parameters that characterise the whole population
  - We want to know these parameters but cannot use the whole population
  - Instead, we take a sample and estimate the population parameters
  - These estimates are random variables with an associated sampling distribution
  - The standard deviation of the sampling distribution is known as the standard error
  - Dividing the estimates by the standard error produces a test statistic
  - This test statistic is also a random variable with a distribution
  - We can calculate the shape of this distribution under the null hypothesis of no effect
  - We can then calculate a  $p$ -value to tell us how likely it would have been to obtain our test statistic if the null hypothesis were true
  - $p < 0.05$  is evidence against the null



# Summary

- This may take some time to sink in if it is new to you – it is the fundamental process used to reach conclusions about fMRI and M/EEG data
- We will see all of this in action as we learn how statistical modelling and inference works inside of SPM



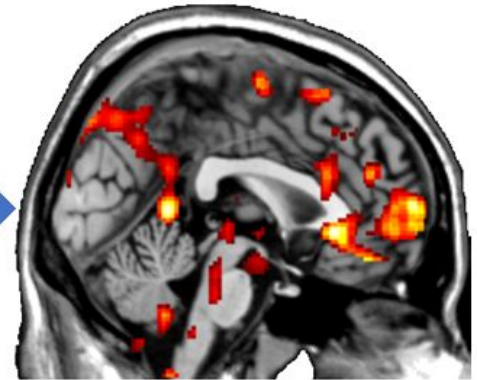
SPM{ $t$ }

Threshold



$t$ -values with  
 $p < 0.05$

Overlay



Final results