```
P(A1B) = P(A1B)/P(B)
             P(A/B) = P(B) P(A/B)
             If A. .. An partition SZ, P(B) = & P(Ai) P(BIAi) = E[P(BIA)]
            P(A|B) = P(B|A)P(A) Independence: P(AB) = P(A)P(B) (orditional Indep: P(AB)C) = P(A|C)P(B|C)
P(AB) = P(AB) = P(A)
            n objects: ni perms, n!/(n-k)! | 12-perms, \binom{n}{k} = \frac{n!}{k!(n-k)!} combos, n_i!ne!...n_r! sized partitions
            var(X)= E[(X-E(X))2] = E[X2] - (E[X])2 = 0x
            E[g(x)] = \underbrace{\xi}_{g(x)} \rho_{x}(x) \text{ or } \int g(x) f_{x}(x) dx \quad Y = aX + b \Rightarrow E[Y] = aE[X] + b, \quad \text{var}(Y) = a^{2} \text{var}(X)
            px(x) = & px, x(xy) px, x(x,y) = px(y) px1x(x/y) px(x) = & px(y) px1x(x/y)
             E[X] = E[E[XIY]] = & PY/y) E[XIY=y]
            Independence: \rho_{X,Y}(x,y) = \rho_{X}(x) \rho_{Y}(y) \forall x,y; E[XY] = E[X]E[Y]; var(x+Y) = var(x) + var(x)

Discrete uniform [a,b]: \rho_{X}(|x|) = \frac{\lambda(x)}{2} \rho_{X}(x) = \frac{\lambda(x)}{2} \rho_{X}(x) = \frac{\lambda(x)}{2} \rho_{X}(x) = \frac{\lambda(x)}{2} \rho_{X}(x)
            Bernoulli: P(1) = p, p(0) = 1-p, E[X]=p, var(x)=p(1-p)
             Binomial: Px(k)= (n)pk(1-p)n-k, E(X)=np, var(X)=np(1-p)
             Geometric: px(k) = (1-p)k-1p, E(X)=p, var(x)= p2
            Poisson: px(x)= e-2 } , E[x]=1, var(x)=2. appreximates binomial w/ n>p, 2=np
(ontinuous E[X]= Jxfx(x)dx, var(X)= S(x-E[X])2fx(x)dx = E[X²]-(E[X])2 Y=aX+b ⇒ same as 1
            CDF: Fx(x)= P(x < x) Fx(k)= Si=- px(i) or Fx(x)= S= fx(t)dt
             p_{X}(k) = F_{X}(k) - F_{X}(k-1), f_{X}(x) = \frac{\partial}{\partial x} F_{X}(x) Y = aX + b is normal if X is normal 1 + Y \sim N(0, 1), X \sim N(u, \sigma^{2}), Y = \frac{X - u}{\sigma}. F_{X}(x) = \Phi(\frac{X - u}{\sigma})
             If Af real line, P(X \in A) > 0, f_{X}(A(X)) = \frac{f_{X}(X)}{P(X \in A)} if X \in A, O else
            E[XIA] = SxfxIa(x)dx fx(x) = Si P(Ai) fxIAi(x) E[X] = Si P(Ai) E[XIAi] if Ai's partition of
            fx, y(x,y) = fy(y) fx 1 y(x/y) fx(x) = Sfy(y) fx 1 y(x/y) dy E[X] = Sfy(y) E[X/Y=y] dy
             f_{X|Y}(x|y) = f_{X}(x)f_{Y|X}(y|x) \quad \text{Independence:} \quad f_{X,Y}(x,y) = f_{X}(x)f_{X}(y) \quad \forall x,y
f_{Y}(y) = f_{X}(x)f_{Y|X}(y|x) \quad \text{Independence:} \quad f_{X,Y}(x,y) = f_{X}(x)f_{X}(y) \quad \forall x,y
f_{Y}(y) = f_{X}(x)f_{Y|X}(y|x) \quad \text{Independence:} \quad f_{X,Y}(x,y) = f_{X}(x)f_{X}(y) \quad \forall x,y
f_{Y}(y) = f_{X}(x)f_{Y|X}(y|x) \quad \text{Independence:} \quad f_{X,Y}(x,y) = f_{X}(x)f_{X}(y) \quad \forall x,y
            derived distributions: get Fy(y)= P(g(x) = y), fy(y)= ay Fy(y)
             Transforms! Mx(s) = E[esx] = E[x1] = do Mx(s) |s=0
More
            Y=aX+b \Rightarrow M_X(s)=e^{sb}M_X(as) X, Y indep \Rightarrow M_{X+X}(s)=M_X(s)M_Y(s)
             PX+X(X+y)= PX(X) * PX(y), fx+x (x+y)= fx(x) * fx(y) Heated E: E[E[XIY]] = E[X]
             conditional var: var(x)= E[var(x/Y)] + var(E[X/Y])
             Sums of Random #s of RVs: Y= &Xi, Xi's have same (u, o2): E(Y] = u E[N]
             var(Y) = 02 E[N] + 12 var(N) My(s) = MN(s) (es = Mx(s) P= (OU(X, Y)
            (ou(X,Y) = E[(X-E(X)(Y-E(Y))]=E(XY)-E(X) E(Y) independence = 0 cov
                                                                                                                Nvar(X) var(Y)
            Least Squares; E[(X-c)2) minimized by C=E(X), E[(X-C)2/Y=q]by C=E[X/Y=q]
            E((X-q(Y))3] by q(Y) = E[XIV] Least squares linear estimator of X based on Y:
            Cont. uniform [a,b]: f_{X}(x) = \frac{1}{5-a} (x) + \frac{(ov(X,Y))}{5^{2}}(Y - E(Y)) w/ MSE (1-p^{2})var(X)
            Exponential! f_X(x) = \lambda e^{-\lambda x} for x \ge 0, f_X(x) = 1 - e^{-\lambda x} for x \ge 0, f(x) = \frac{1}{2}
            Normal: 0027 e (x21)202=fx(x), E(X)=4 var(x)=02
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