

Lecture 8

Reducing the L₂ Norm of a Distribution via Flattening

A useful technique in distribution testing is reducing the L₂ norm of a distribution. This can be achieved through *flattening*, a process that transforms the original distribution p into a new distribution p' . The core idea is to distribute the probability mass of elements with high probability in p among multiple elements in p' , effectively “flattening” the distribution and reducing its L₂ norm. This process is illustrated in Figure 1.

The transformation from p to p' involves determining, for each element i in the domain $[n]$, the number of elements in p' that will correspond to i . Let b_i denote this number for element i . We will discuss how to choose b_i shortly; for now, assume b_i is given. For each element i in the domain of p , we associate b_i elements in the domain of p' with i . We refer to these associated elements as “buckets.” The probability mass of i in p is then distributed equally among its b_i buckets in p' .

Formally, we define:

$$\text{New domain of } p' : \quad D' := \{(i, j) \mid i \in [n], j \in [b_i]\} \quad (1)$$

$$\text{New domain size:} \quad |D'| = \sum_{i=1}^n b_i \quad (2)$$

$$\text{Probability of a domain element in } p' : \quad p'_{(i,j)} = \frac{p_i}{b_i} \quad (3)$$

It is straightforward to verify from this definition that the probabilities in p' sum to one.

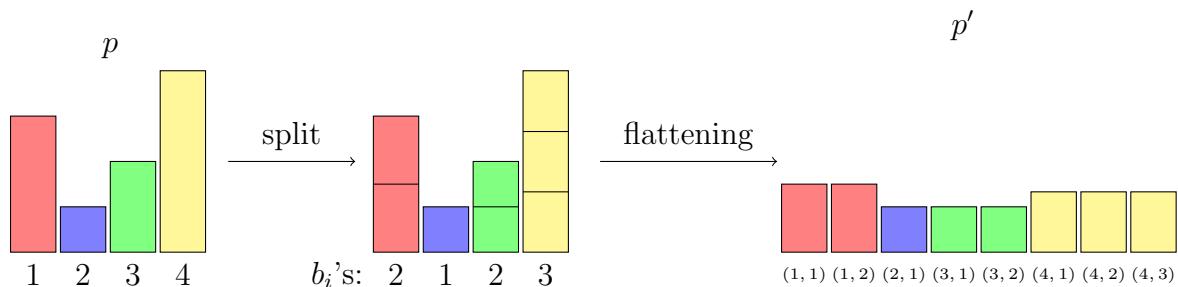


Figure 1: Flattening of a distribution p to p'

This transformation ideally has the following properties:

1. **Access:** For known b_i 's, we have the same access to p' as to p . That is, if we know the probabilities of every element in p , we also know the probabilities of every element in p' . Similarly, if we have sample access to p , we have sample access to p' .
2. **Preservation of L₁ distance:** For a fixed set of b_i 's, the transformation preserves the L₁ distance between distributions.
3. **Reduced L₂ norm:** We can choose b_i 's such that the L₂ norm of p' is low.

The first two properties are straightforward to establish for any choice of b_i 's. Regarding access, if p is known, the probability of each bucket (i, j) is given by Equation (3). To sample from p' , we can use a sample from p . Specifically, if i is a sample drawn from p , we can then draw a uniform random sample j from $[b_i]$. The pair (i, j) then constitutes a sample from p' .

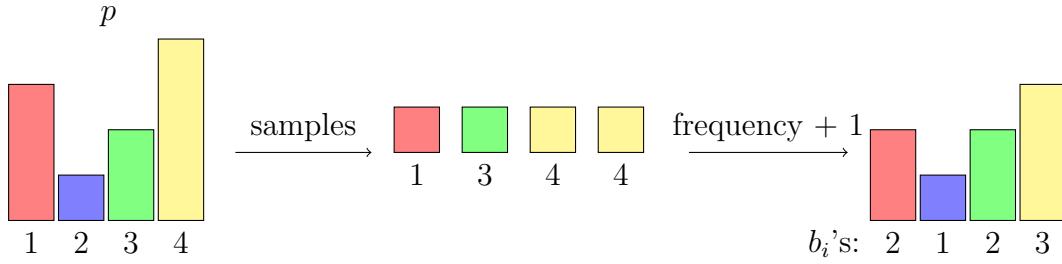
Next, we show that if we flatten p and q using the same set of b_i 's, the resulting distributions p' and q' have the same L₁ distance as p and q . Formally,

$$\begin{aligned} \|p' - q'\|_1 &= \sum_{i \in [n]} \sum_{j \in [b_i]} |p'_{(i,j)} - q'_{(i,j)}| = \sum_{i \in [n]} \sum_{j \in [b_i]} \left| \frac{p_i}{b_i} - \frac{q_i}{b_i} \right| \\ &= \sum_{i \in [n]} |p_i - q_i| \sum_{j \in [b_i]} \frac{1}{b_i} \sum_{i \in [n]} |p_i - q_i| = \|p - q\|_1. \end{aligned}$$

Determining the Number of Buckets

For the third property (reduced L₂ norm), we need to carefully choose the number of buckets, b_i , for each element i . Ideally, we want to decompose elements of p with high probability into smaller pieces, effectively "flattening" the distribution and making it more uniform. Thus, we aim for b_i to be proportional to p_i . Various methods exist for determining the b_i values. Here, we focus on the approach proposed in [DK16]. This approach is illustrated in Figure 2. Given a parameter k , the method proceeds as follows:

1. Draw k' from a Poisson distribution with mean k (i.e., $k' \leftarrow \text{Poi}(k)$).
2. Draw a set F of k' independent samples from p .
3. For each $i \in [n]$, let f_i denote the frequency of element i in F .
4. For each $i \in [n]$, set $b_i = f_i + 1$.

Figure 2: Calculating b_i 's from samples of p

Let's analyze how this flattening affects the L_2^2 norm of p' . We focus on the expected value of the L_2^2 norm, where the expectation is taken over the randomness of the sample set F .

$$\begin{aligned} \mathbf{E}_F \left[\|p'\|_2^2 \right] &= \mathbf{E}_F \left[\sum_{i=1}^n \sum_{j=1}^{b_i} (p'_{(i,j)})^2 \right] = \mathbf{E}_F \left[\sum_{i=1}^n \sum_{j=1}^{b_i} \frac{p_i^2}{b_i^2} \right] \\ &= \sum_{i=1}^n p_i^2 \cdot \mathbf{E}_F \left[\frac{1}{b_i} \right] = \sum_{i=1}^n p_i^2 \cdot \mathbf{E}_F \left[\frac{1}{f_i + 1} \right]. \end{aligned}$$

Recall that we defined $b_i = f_i + 1$, where f_i is the frequency of element i in the sample set F . From the discussion on Poissonization, we know that f_i is a random variable drawn from $\text{Poi}(p_i k)$. Let's focus on the expected value of $1/(f_i + 1)$:

$$\begin{aligned} \mathbf{E}_{f_i \sim \text{Poi}(p_i k)} \left[\frac{1}{f_i + 1} \right] &= \mathbf{E} \left[\int_0^1 x^{f_i} dx \right] = \int_0^1 \mathbf{E}[x^{f_i}] dx \quad (\text{via linearity of expectation}) \\ &= \int_0^1 \left(\sum_{t=0}^{\infty} x^t \cdot \frac{e^{-p_i k} (p_i k)^t}{t!} \right) dx \quad (\text{via definition of Poisson dist.}) \\ &= \int_0^1 e^{-p_i k + p_i k x} \left(\sum_{t=0}^{\infty} \frac{e^{-p_i k x} (p_i k x)^t}{t!} \right) dx \end{aligned}$$

Note that the terms in the sum are probabilities of $Z = t$, where Z is drawn from $\text{Poi}(p_i k x)$. Thus, the sum of those terms is equal to one. Therefore, we have:

$$\mathbf{E}_{f_i \sim \text{Poi}(p_i k)} \left[\frac{1}{f_i + 1} \right] = \int_0^1 e^{p_i k (x-1)} dx = \frac{1}{p_i k} \cdot e^{p_i k (x-1)} \Big|_{x=0}^1 \leq \frac{1}{p_i k}$$

Now, returning to the bound for the L_2^2 norm of p' , we have:

$$\mathbf{E}_F \left[\|p'\|_2^2 \right] = \sum_{i=1}^n p_i^2 \cdot \mathbf{E}_F \left[\frac{1}{f_i + 1} \right] \leq \frac{1}{k} \sum_{i=1}^n p_i = \frac{1}{k}$$

Costs of flattening: While flattening significantly reduces the L_2^2 -norm, it introduces some costs. First, the process of determining the b_i 's requires drawing samples from p , thus increasing the overall sample complexity. Second, flattening inflates the domain size, which can indirectly increase the sample complexity of any subsequent algorithms that operate on the flattened distribution:

$$\text{new domain size: } |D'| = \sum_{i=1}^n b_i = \sum_{i=1}^n f_i + 1 = \text{Poi}(k) + n$$

Other flattening schemes: There are various methods for determining the b_i values, allowing us to choose a flattening strategy tailored to the specific problem structure. For instance, some flattening techniques are designed for testing the independence of random variables. Here, we focused on a scheme suitable for an *unknown* distribution p . If p were *known* (i.e., all probabilities p_i were available), we could set $b_i = \lfloor np_i \rfloor + 1$. As an exercise, the reader can verify that this approach reduces the L_2^2 norm of p' to $O(1/n)$.

Back to Closeness Tester

Recall from our previous lecture that there exists an algorithm that, for two distributions over $[n]$, distinguishes whether $p = q$ or they are ϵ -far with a probability of at least 0.9 using the following number of samples:

$$s = O \left(\frac{n \cdot \max(\|p\|_2, \|q\|_2)}{\epsilon^2} \right).$$

Given the flattening technique introduced in this lecture, we can efficiently test closeness between p and q . For some k (to be determined), we draw $\text{Poi}(k)$ samples from p and q and use them to create flattened distributions p' and q' , respectively. We then apply the closeness tester to p' and q' to determine if they are equal or ϵ -far. This process is depicted in Figure 3. As shown previously, flattening preserves the L_1 distance between distributions. Thus, $p = q$ if and only if $p' = q'$, and if p is ϵ -far from q , then p' is ϵ -far from q' .

To reduce the L_2 norm of both p and q , we combine the flattening steps. Drawing two sets of $\text{Poi}(k)$ samples (one from p and one from q), we set $b_i = f_i^{(p)} + f_i^{(q)}$, where $f_i^{(p)}$ and $f_i^{(q)}$ are the frequencies of element i in the respective sample sets. As long as $b_i \geq f_i^{(p)} + 1$, the L_2^2 norm reduction is guaranteed, as shown earlier.

To ensure that the testing step has a reduced sample complexity, we must show that the

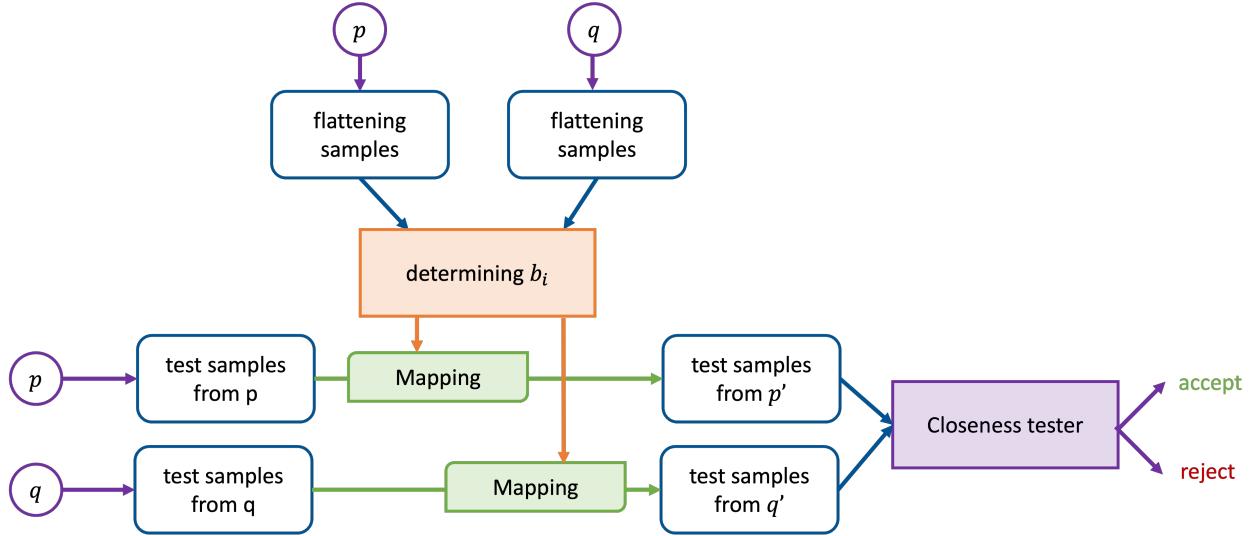


Figure 3: Diagram of the flattening and testing process

L_2 norm is reduced with high probability. Since the expected L_2 norm of p' is at most $1/k$, Markov's inequality implies that the L_2 norms of p' and q' are each at most $100/k$ with probability at least 0.99.

The optimal choice of k balances the sample complexity of the flattening and testing steps. The total sample complexity is

$$O(k + s) = O\left(k + \frac{n'}{\epsilon^2 \sqrt{k}}\right) = O\left(k + \frac{n + k}{\epsilon^2 \sqrt{k}}\right),$$

where $n' = O(n + k)$ is the size of the new domain.

If $k \geq n$, the sample complexity increases with k , so we need not consider $k > n$. When $k \leq n$, we have two competing terms: $O(k)$ (increasing) and $O\left(\frac{n}{\sqrt{k}\epsilon^2}\right)$ (decreasing). Minimizing their sum yields $k = n^{2/3}/\epsilon^{4/3}$. Since $k \leq n$, the optimal choice is $k = \min(n, n^{2/3}/\epsilon^{4/3})$. Substituting this value into the total sample complexity gives a final sample complexity of

$$O\left(\frac{n^{2/3}}{\epsilon^{4/3}} + \frac{\sqrt{n}}{\epsilon^2}\right).$$

This sample complexity is known to be optimal for this problem.

Bibliographic Note: The content of this lecture is based on [CDVV14, DK16]. Further applications of the flattening technique can be found in that work.

References

- [CDVV14] Siu-on Chan, Ilias Diakonikolas, Paul Valiant, and Gregory Valiant. Optimal algorithms for testing closeness of discrete distributions. In *SODA*, pages 1193–1203, 2014.
- [DK16] Ilias Diakonikolas and Daniel M. Kane. A new approach for testing properties of discrete distributions. In *IEEE 57th Annual Symposium on Foundations of Computer Science, FOCS 2016, 9-11 October 2016, Hyatt Regency, New Brunswick, New Jersey, USA*, pages 685–694, 2016.