

## Lecture 8

### Reducing the $L_2$ Norm of a Distribution via Flattening

A useful technique in distribution testing is reducing the  $L_2$  norm of a distribution. This can be achieved through *flattening*, a process that transforms the original distribution  $p$  into a new distribution  $p'$ . The core idea is to distribute the probability mass of elements with high probability in  $p$  among multiple elements in  $p'$ , effectively “flattening” the distribution and reducing its  $L_2$  norm. This process is illustrated in Figure 1.

The transformation from  $p$  to  $p'$  involves determining, for each element  $i$  in the domain  $[n]$ , the number of elements in  $p'$  that will correspond to  $i$ . Let  $b_i$  denote this number for element  $i$ . We will discuss how to choose  $b_i$  shortly; for now, assume  $b_i$  is given. For each element  $i$  in the domain of  $p$ , we associate  $b_i$  elements in the domain of  $p'$  with  $i$ . We refer to these associated elements as “buckets.” The probability mass of  $i$  in  $p$  is then distributed equally among its  $b_i$  buckets in  $p'$ .

Formally, we define:

$$\text{New domain of } p' : \quad D' := \{(i, j) \mid i \in [n], j \in [b_i]\} \quad (1)$$

$$\text{New domain size:} \quad |D'| = \sum_{i=1}^n b_i \quad (2)$$

$$\text{Probability of a domain element in } p' : \quad p'_{(i,j)} = \frac{p_i}{b_i} \quad (3)$$

It is straightforward to verify from this definition that the probabilities in  $p'$  sum to one.

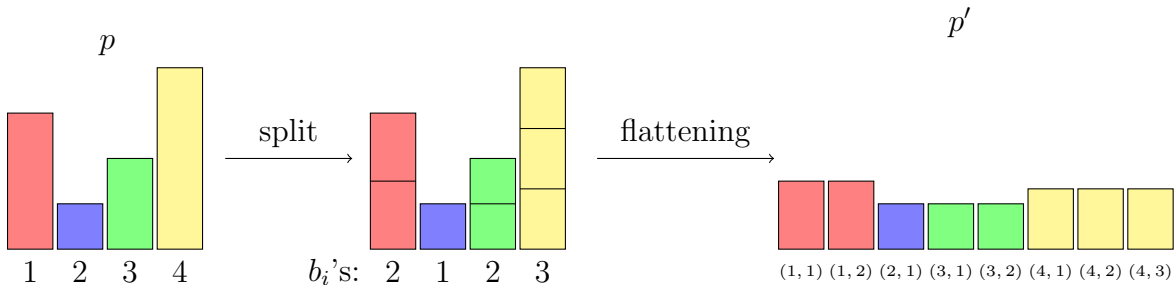


Figure 1: Flattening of a distribution  $p$  to  $p'$

This transformation ideally has the following properties:

1. **Access:** For known  $b_i$ 's, we have the same access to  $p'$  as to  $p$ . That is, if we know the probabilities of every element in  $p$ , we also know the probabilities of every element in  $p'$ . Similarly, if we have sample access to  $p$ , we have sample access to  $p'$ .
2. **Preservation of  $L_1$  distance:** For a fixed set of  $b_i$ 's, the transformation preserves the  $L_1$  distance between distributions.
3. **Reduced  $L_2$  norm:** We can choose  $b_i$ 's such that the  $L_2$  norm of  $p'$  is low.

The first two properties are straightforward to establish for any choice of  $b_i$ 's. Regarding access, if  $p$  is known, the probability of each bucket  $(i, j)$  is given by Equation (3). To sample from  $p'$ , we can use a sample from  $p$ . Specifically, if  $i$  is a sample drawn from  $p$ , we can then draw a uniform random sample  $j$  from  $[b_i]$ . The pair  $(i, j)$  then constitutes a sample from  $p'$ .

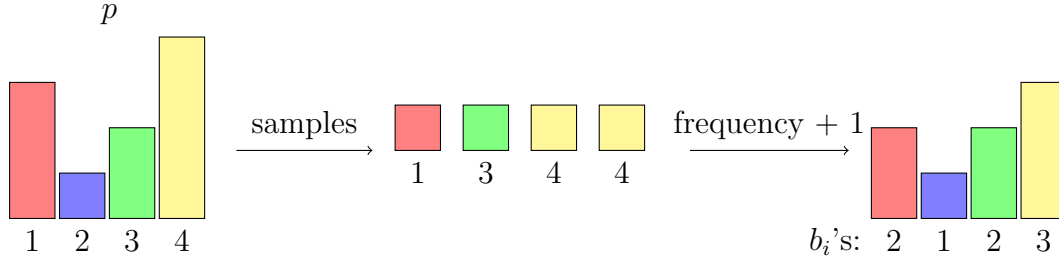
Next, we show that if we flatten  $p$  and  $q$  using the same set of  $b_i$ 's, the resulting distributions  $p'$  and  $q'$  have the same  $L_1$  distance as  $p$  and  $q$ . Formally,

$$\begin{aligned} \|p' - q'\|_1 &= \sum_{i \in [n]} \sum_{j \in [b_i]} |p'_{(i,j)} - q'_{(i,j)}| = \sum_{i \in [n]} \sum_{j \in [b_i]} \left| \frac{p_i}{b_i} - \frac{q_i}{b_i} \right| \\ &= \sum_{i \in [n]} |p_i - q_i| \sum_{j \in [b_i]} \frac{1}{b_i} = \sum_{i \in [n]} |p_i - q_i| = \|p - q\|_1. \end{aligned}$$

## Determining the Number of Buckets

For the third property (reduced  $L_2$  norm), we need to carefully choose the number of buckets,  $b_i$ , for each element  $i$ . Ideally, we want to decompose elements of  $p$  with high probability into smaller pieces, effectively "flattening" the distribution and making it more uniform. Thus, we aim for  $b_i$  to be proportional to  $p_i$ . Various methods exist for determining the  $b_i$  values. Here, we focus on the approach proposed in [DK16]. This approach is illustrated in Figure 2. Given a parameter  $k$ , the method proceeds as follows:

1. Draw  $k'$  from a Poisson distribution with mean  $k$  (i.e.,  $k' \leftarrow \text{Poi}(k)$ ).
2. Draw a set  $F$  of  $k'$  independent samples from  $p$ .
3. For each  $i \in [n]$ , let  $f_i$  denote the frequency of element  $i$  in  $F$ .
4. For each  $i \in [n]$ , set  $b_i = f_i + 1$ .

Figure 2: Calculating  $b_i$ 's from samples of  $p$ 

Let's analyze how this flattening affects the  $L_2^2$  norm of  $p'$ . We focus on the expected value of the  $L_2^2$  norm, where the expectation is taken over the randomness of the sample set  $F$ .

$$\begin{aligned} \mathbf{E}_F \left[ \|p'\|_2^2 \right] &= \mathbf{E}_F \left[ \sum_{i=1}^n \sum_{j=1}^{b_i} (p'_{(i,j)})^2 \right] = \mathbf{E}_F \left[ \sum_{i=1}^n \sum_{j=1}^{b_i} \frac{p_i^2}{b_i} \right] \\ &= \sum_{i=1}^n p_i^2 \cdot \mathbf{E}_F \left[ \frac{1}{b_i} \right] = \sum_{i=1}^n p_i^2 \cdot \mathbf{E}_F \left[ \frac{1}{f_i + 1} \right]. \end{aligned}$$

Recall that we defined  $b_i = f_i + 1$ , where  $f_i$  is the frequency of element  $i$  in the sample set  $F$ . From the discussion on Poissonization, we know that  $f_i$  is a random variable drawn from  $\text{Poi}(p_i k)$ . Let's focus on the expected value of  $1/(f_i + 1)$ :

$$\begin{aligned} \mathbf{E}_{f_i \sim \text{Poi}(p_i \cdot k)} \left[ \frac{1}{f_i + 1} \right] &= \mathbf{E} \left[ \int_0^1 x^{f_i} dx \right] = \int_0^1 \mathbf{E} [x^{f_i}] dx \quad (\text{via linearity of expectation}) \\ &= \int_0^1 \left( \sum_{t=0}^{\infty} x^t \cdot \frac{e^{-p_i k} (p_i k)^t}{t!} \right) dx \quad (\text{via definition of Poisson dist.}) \\ &= \int_0^1 e^{-p_i k + p_i k x} \left( \sum_{t=0}^{\infty} \frac{e^{-p_i k x} (p_i k x)^t}{t!} \right) dx \end{aligned}$$

Note that the terms in the sum are probabilities of  $Z = t$ , where  $Z$  is drawn from  $\text{Poi}(p_i k x)$ . Thus, the sum of those terms is equal to one. Therefore, we have:

$$\mathbf{E}_{f_i \sim \text{Poi}(p_i \cdot k)} \left[ \frac{1}{f_i + 1} \right] = \int_0^1 e^{p_i k (x-1)} dx = \frac{1}{p_i k} \cdot e^{p_i k (x-1)} \Big|_{x=0}^1 \leq \frac{1}{p_i k}$$

Now, returning to the bound for the  $L_2^2$  norm of  $p'$ , we have:

$$\mathbf{E}_F \left[ \|p'\|_2^2 \right] = \sum_{i=1}^n p_i^2 \cdot \mathbf{E}_F \left[ \frac{1}{f_i + 1} \right] \leq \frac{1}{k} \sum_{i=1}^n p_i = \frac{1}{k}$$

**Costs of flattening:** While flattening significantly reduces the  $L_2^2$ -norm, it introduces some costs. First, the process of determining the  $b_i$ 's requires drawing samples from  $p$ , thus increasing the overall sample complexity. Second, flattening inflates the domain size, which can indirectly increase the sample complexity of any subsequent algorithms that operate on the flattened distribution:

$$\text{new domain size: } |D'| = \sum_{i=1}^n b_i = \sum_{i=1}^n f_i + 1 = \text{Poi}(k) + n$$

**Other flattening schemes:** There are various methods for determining the  $b_i$  values, allowing us to choose a flattening strategy tailored to the specific problem structure. For instance, some flattening techniques are designed for testing the independence of random variables. Here, we focused on a scheme suitable for an *unknown* distribution  $p$ . If  $p$  were *known* (i.e., all probabilities  $p_i$  were available), we could set  $b_i = \lfloor np_i \rfloor + 1$ . As an exercise, the reader can verify that this approach reduces the  $L_2^2$  norm of  $p'$  to  $O(1/n)$ .

## Back to Closeness Tester

Recall from our previous lecture that there exists an algorithm that, for two distributions over  $[n]$ , distinguishes whether  $p = q$  or they are  $\epsilon$ -far with a probability of at least 0.9 using the following number of samples:

$$s = O \left( \frac{n \cdot \max(\|p\|_2, \|q\|_2)}{\epsilon^2} \right).$$

Given the flattening technique introduced in this lecture, we can efficiently test closeness between  $p$  and  $q$ . For some  $k$  (to be determined), we draw  $\text{Poi}(k)$  samples from  $p$  and  $q$  and use them to create flattened distributions  $p'$  and  $q'$ , respectively. We then apply the closeness tester to  $p'$  and  $q'$  to determine if they are equal or  $\epsilon$ -far. This process is depicted in Figure 3. As shown previously, flattening preserves the  $L_1$  distance between distributions. Thus,  $p = q$  if and only if  $p' = q'$ , and if  $p$  is  $\epsilon$ -far from  $q$ , then  $p'$  is  $\epsilon$ -far from  $q'$ .

To reduce the  $L_2$  norm of both  $p$  and  $q$ , we combine the flattening steps. Drawing two sets of  $\text{Poi}(k)$  samples (one from  $p$  and one from  $q$ ), we set  $b_i = f_i^{(p)} + f_i^{(q)}$ , where  $f_i^{(p)}$  and  $f_i^{(q)}$  are the frequencies of element  $i$  in the respective sample sets. As long as  $b_i \geq f_i^{(p)} + 1$ , the  $L_2^2$  norm reduction is guaranteed, as shown earlier.

To ensure that the testing step has a reduced sample complexity, we must show that the

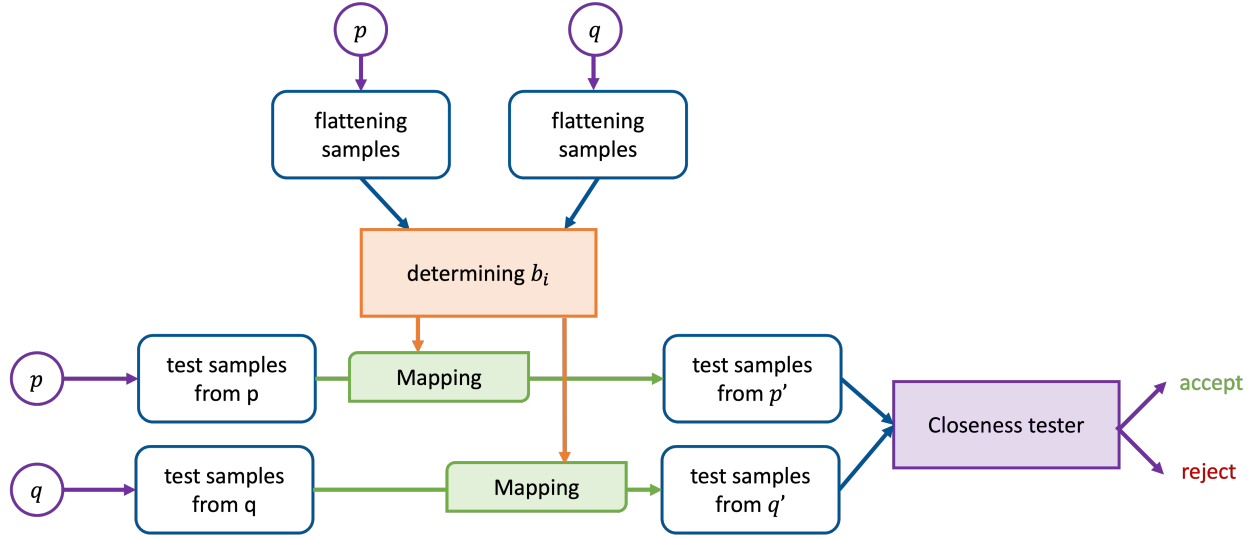


Figure 3: Diagram of the flattening and testing process

$L_2$  norm is reduced with high probability. Since the expected  $L_2$  norm of  $p'$  is at most  $1/k$ , Markov's inequality implies that the  $L_2$  norms of  $p'$  and  $q'$  are each at most  $100/k$  with probability at least 0.99.

The optimal choice of  $k$  balances the sample complexity of the flattening and testing steps. The total sample complexity is

$$O(k + s) = O\left(k + \frac{n'}{\epsilon^2 \sqrt{k}}\right) = O\left(k + \frac{n + k}{\epsilon^2 \sqrt{k}}\right),$$

where  $n' = O(n + k)$  is the size of the new domain.

If  $k \geq n$ , the sample complexity increases with  $k$ , so we need not consider  $k > n$ . When  $k \leq n$ , we have two competing terms:  $O(k)$  (increasing) and  $O\left(\frac{n}{\sqrt{k}\epsilon^2}\right)$  (decreasing). Minimizing their sum yields  $k = n^{2/3}/\epsilon^{4/3}$ . Since  $k \leq n$ , the optimal choice is  $k = \min(n, n^{2/3}/\epsilon^{4/3})$ . Substituting this value into the total sample complexity gives a final sample complexity of

$$O\left(\frac{n^{2/3}}{\epsilon^{4/3}} + \frac{\sqrt{n}}{\epsilon^2}\right).$$

This sample complexity is known to be optimal for this problem.

**Bibliographic Note:** The content of this lecture is based on [CDVV14, DK16]. Further applications of the flattening technique can be found in that work.

## References

- [CDVV14] Siu-on Chan, Ilias Diakonikolas, Paul Valiant, and Gregory Valiant. Optimal algorithms for testing closeness of discrete distributions. In *SODA*, pages 1193–1203, 2014.
- [DK16] Ilias Diakonikolas and Daniel M. Kane. A new approach for testing properties of discrete distributions. In *IEEE 57th Annual Symposium on Foundations of Computer Science, FOCS 2016, 9-11 October 2016, Hyatt Regency, New Brunswick, New Jersey, USA*, pages 685–694, 2016.