

# Testing and Estimation for Sparsity-Inducing Power Penalties

Maryclare Griffin

Department of Mathematics and Statistics  
University of Massachusetts Amherst

November 15, 2023

# Introduction



# Introduction



# Introduction



# Introduction



# Introduction



# Introduction



# Sparsity Inducing Power Penalties

Power penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_q^q$



# Sparsity Inducing Power Penalties

Power penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_q^q$

Posterior mode of  $\boldsymbol{\beta}$  under:

# Sparsity Inducing Power Penalties

Power penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_q^q$

Posterior mode of  $\boldsymbol{\beta}$  under:

$$\begin{aligned}\mathbf{y} &= \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, & \mathbf{z} &\sim \text{normal}(\mathbf{0}, \mathbf{I}_n) \\ p(\beta_j) &\propto \exp\{-\lambda|\beta_j|^q\}\end{aligned}$$

# Sparsity Inducing Power Penalties

Power penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_q^q$

Posterior mode of  $\boldsymbol{\beta}$  under:

$$\begin{aligned}\mathbf{y} &= \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, & \mathbf{z} &\sim \text{normal}(\mathbf{0}, \mathbf{I}_n) \\ p(\beta_j) &\propto \exp\{-\lambda|\beta_j|^q\}\end{aligned}$$

- $p(\beta_j)$  is density of the [exponential power distribution](#)

# Sparsity Inducing Power Penalties

Power penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_q^q$

Posterior mode of  $\boldsymbol{\beta}$  under:

$$\begin{aligned}\mathbf{y} &= \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, & \mathbf{z} &\sim \text{normal}(\mathbf{0}, \mathbf{I}_n) \\ p(\beta_j) &\propto \exp\{-\lambda|\beta_j|^q\}\end{aligned}$$

- $p(\beta_j)$  is density of the [exponential power distribution](#)
- Special cases include:

# Sparsity Inducing Power Penalties

Power penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_q^q$

Posterior mode of  $\boldsymbol{\beta}$  under:

$$\begin{aligned}\mathbf{y} &= \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, & \mathbf{z} &\sim \text{normal}(\mathbf{0}, \mathbf{I}_n) \\ p(\beta_j) &\propto \exp\{-\lambda|\beta_j|^q\}\end{aligned}$$

- $p(\beta_j)$  is density of the **exponential power distribution**
- Special cases include:
  - $q = 1$ : Laplace model/lasso penalty

# Sparsity Inducing Power Penalties

Power penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_q^q$

Posterior mode of  $\boldsymbol{\beta}$  under:

$$\begin{aligned}\mathbf{y} &= \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, & \mathbf{z} &\sim \text{normal}(\mathbf{0}, \mathbf{I}_n) \\ p(\beta_j) &\propto \exp\{-\lambda|\beta_j|^q\}\end{aligned}$$

- $p(\beta_j)$  is density of the **exponential power distribution**
- Special cases include:
  - $q = 1$ : Laplace model/lasso penalty
  - $q = 2$ : Normal distribution/ridge penalty

# Sparsity Inducing Power Penalties

Power penalized regression solves  $\min_{\beta} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \|\beta\|_q^q$

Posterior mode of  $\beta$  under:

$$\begin{aligned}\mathbf{y} &= \mathbf{X}\beta + \sigma\mathbf{z}, & \mathbf{z} &\sim \text{normal}(\mathbf{0}, \mathbf{I}_n) \\ p(\beta_j) &\propto \exp\{-\lambda|\beta_j|^q\}\end{aligned}$$

- $p(\beta_j)$  is density of the **exponential power distribution**
- Special cases include:
  - $q = 1$ : Laplace model/lasso penalty
  - $q = 2$ : Normal distribution/ridge penalty
  - $q \leq 1$ : Bridge penalty

What is this talk going to be about?

“Testing Sparsity-Inducing Penalties”



# What is this talk going to be about?

“Testing Sparsity-Inducing Penalties”

Testing when  $q = 1$  is a “good” choice, picking a “better”  $q$

# What is this talk going to be about?

“Testing Sparsity-Inducing Penalties”

Testing when  $q = 1$  is a “good” choice, picking a “better”  $q$

“Improved Pathwise Coordinate Descent for Power Penalties”

# What is this talk going to be about?

## “Testing Sparsity-Inducing Penalties”

Testing when  $q = 1$  is a “good” choice, picking a “better”  $q$

## “Improved Pathwise Coordinate Descent for Power Penalties”

Solving  $\min_{\boldsymbol{\beta}} ||\mathbf{y} - \mathbf{X}\boldsymbol{\beta}||_2^2 + \lambda ||\boldsymbol{\beta}||_q^q$  for a range of  $q$ ,  $\lambda$

# “Testing Sparsity-Inducing Penalties”

## Problem

# “Testing Sparsity-Inducing Penalties”

## Problem

Lasso penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_1$

# “Testing Sparsity-Inducing Penalties”

## Problem

Lasso penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_1$

☺ Computationally simple

# “Testing Sparsity-Inducing Penalties”

## Problem

Lasso penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_1$

- ☺ Computationally simple
- ☺ Yields possibly sparse estimates of  $\boldsymbol{\beta}$

# “Testing Sparsity-Inducing Penalties”

## Problem

Lasso penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_1$

- ☺ Computationally simple
- ☺ Yields possibly sparse estimates of  $\boldsymbol{\beta}$
- ☺ Same as finding posterior mode of  $\boldsymbol{\beta}$  under

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$
$$p(\beta_j) \propto \exp\{-\lambda |\beta_j|\}$$



# “Testing Sparsity-Inducing Penalties”

## Problem

Lasso penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_1$

- ☺ Computationally simple
- ☺ Yields possibly sparse estimates of  $\boldsymbol{\beta}$
- ☺ Same as finding posterior mode of  $\boldsymbol{\beta}$  under

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$
$$p(\beta_j) \propto \exp\{-\lambda |\beta_j|\}$$

- ☺ Can perform poorly if  $\boldsymbol{\beta}$  is very sparse or not sparse at all

# “Testing Sparsity-Inducing Penalties”

## Problem

Lasso penalized regression solves  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_1$

- ☺ Computationally simple
- ☺ Yields possibly sparse estimates of  $\boldsymbol{\beta}$
- ☺ Same as finding posterior mode of  $\boldsymbol{\beta}$  under

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$
$$p(\beta_j) \propto \exp\{-\lambda |\beta_j|\}$$

- ☺ Can perform poorly if  $\boldsymbol{\beta}$  is very sparse or not sparse at all

## Objectives

# “Testing Sparsity-Inducing Penalties”

## Problem

Lasso penalized regression solves  $\min_{\beta} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \|\beta\|_1$

- ☺ Computationally simple
- ☺ Yields possibly sparse estimates of  $\beta$
- ☺ Same as finding posterior mode of  $\beta$  under

$$\mathbf{y} = \mathbf{X}\beta + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$
$$p(\beta_j) \propto \exp\{-\lambda |\beta_j|\}$$

- ☺ Can perform poorly if  $\beta$  is very sparse or not sparse at all

## Objectives

- Test Laplace prior appropriateness

# “Testing Sparsity-Inducing Penalties”

## Problem

Lasso penalized regression solves  $\min_{\beta} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \|\beta\|_1$

- ☺ Computationally simple
- ☺ Yields possibly sparse estimates of  $\beta$
- ☺ Same as finding posterior mode of  $\beta$  under

$$\mathbf{y} = \mathbf{X}\beta + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$
$$p(\beta_j) \propto \exp\{-\lambda |\beta_j|\}$$

- ☺ Can perform poorly if  $\beta$  is very sparse or not sparse at all

## Objectives

- Test Laplace prior appropriateness
- Adaptively choose a better prior

Why bother treating the lasso penalty as a prior?

## Why bother treating the lasso penalty as a prior?

- Tuning parameters become variance components

## Why bother treating the lasso penalty as a prior?

- Tuning parameters become variance components,  $\lambda = 2/\tau^2$

# Why bother treating the lasso penalty as a prior?

- Tuning parameters become variance components,  $\lambda = 2/\tau^2$ 
  - Variance components are interpretable



# Why bother treating the lasso penalty as a prior?

- Tuning parameters become variance components,  $\lambda = 2/\tau^2$ 
  - Variance components are interpretable
  - Variance component estimation has a long history

# Why bother treating the lasso penalty as a prior?

- Tuning parameters become variance components,  $\lambda = 2/\tau^2$ 
  - Variance components are interpretable
  - Variance component estimation has a long history
- Posterior summaries have decision theoretic justifications

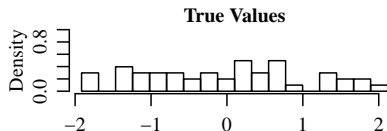
# Why bother treating the lasso penalty as a prior?

- Tuning parameters become variance components,  $\lambda = 2/\tau^2$ 
  - Variance components are [interpretable](#)
  - Variance component estimation has a long history
- Posterior summaries have decision theoretic justifications

$$\overbrace{\mathbb{E} [\boldsymbol{\beta} | \mathbf{X}, \mathbf{y}, \sigma^2, \lambda]}^{\text{Posterior Mean}} = \underset{\mathbf{b}}{\operatorname{argmin}} \underbrace{\int \|\boldsymbol{\beta} - \mathbf{b}\|_2^2 p(\mathbf{y} | \mathbf{X}, \boldsymbol{\beta}, \sigma^2) p(\boldsymbol{\beta} | \lambda) d\boldsymbol{\beta}}_{\text{Weighted average squared error loss}}$$

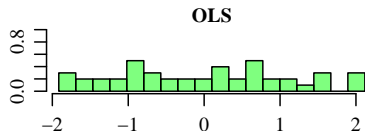
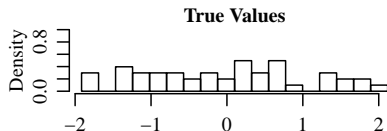
What if  $\beta$  is not sparse at all?

Simulated example with  $n = 50$ ,  $p = 45$ ,  $\sigma^2 = \tau^2 = 1$



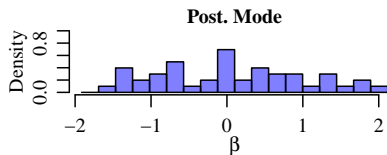
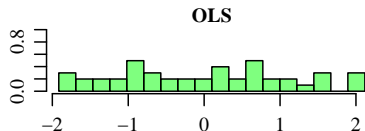
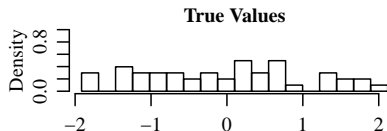
What if  $\beta$  is not sparse at all?

Simulated example with  $n = 50$ ,  $p = 45$ ,  $\sigma^2 = \tau^2 = 1$



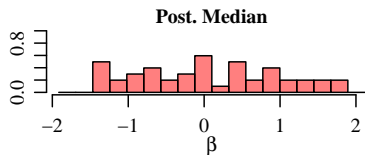
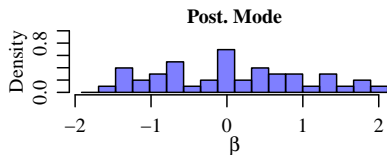
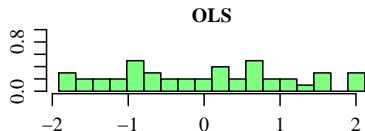
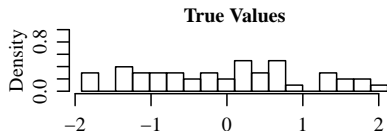
# What if $\beta$ is not sparse at all?

Simulated example with  $n = 50$ ,  $p = 45$ ,  $\sigma^2 = \tau^2 = 1$



# What if $\beta$ is not sparse at all?

Simulated example with  $n = 50$ ,  $p = 45$ ,  $\sigma^2 = \tau^2 = 1$



What if  $\beta$  is sparse?



What if  $\beta$  is sparse?

- Sparsity rate of mode under Laplace depends on  $\sigma^2/\tau^2$

What if  $\beta$  is sparse?

- Sparsity rate of mode under Laplace depends on  $\sigma^2/\tau^2$
- $\beta$  could be spike-and-slab distributed

## What if $\beta$ is sparse?

- Sparsity rate of mode under Laplace depends on  $\sigma^2/\tau^2$
- $\beta$  could be spike-and-slab distributed
  - $\beta_j$  is exactly equal to zero with probability  $1 - \pi$

## What if $\beta$ is sparse?

- Sparsity rate of mode under Laplace depends on  $\sigma^2/\tau^2$
- $\beta$  could be spike-and-slab distributed
  - $\beta_j$  is exactly equal to zero with probability  $1 - \pi$
  - $\beta_j \sim \text{normal}(0, \tau^2/\pi)$  distribution otherwise

## What if $\beta$ is sparse?

- Sparsity rate of mode under Laplace depends on  $\sigma^2/\tau^2$
- $\beta$  could be spike-and-slab distributed
  - $\beta_j$  is exactly equal to zero with probability  $1 - \pi$
  - $\beta_j \sim \text{normal}(0, \tau^2/\pi)$  distribution otherwise
- Sparsity rate of mode independent of sparsity rate of  $\beta$ !

# Existing Literature on Testing Specification of $p(\beta)$

Largely appropriateness of normal random effects models

# Existing Literature on Testing Specification of $p(\beta)$

Largely appropriateness of normal random effects models

Tricky because:

# Existing Literature on Testing Specification of $p(\beta)$

Largely appropriateness of normal random effects models

Tricky because:

- $\beta$  is latent



# Existing Literature on Testing Specification of $p(\beta)$

Largely appropriateness of normal random effects models

Tricky because:

- $\beta$  is latent
- Estimates of  $\beta$  often reflect assumed model

# Existing Literature on Testing Specification of $p(\beta)$

Largely appropriateness of normal random effects models

Tricky because:

- $\beta$  is latent
- Estimates of  $\beta$  often reflect assumed model

Existing tests we could consider may:

# Existing Literature on Testing Specification of $p(\beta)$

Largely appropriateness of normal random effects models

Tricky because:

- $\beta$  is latent
- Estimates of  $\beta$  often reflect assumed model

Existing tests we could consider may:

- Depend on unknown parameters,  $\sigma^2$ ,  $\tau^2$

# Existing Literature on Testing Specification of $p(\beta)$

Largely appropriateness of normal random effects models

Tricky because:

- $\beta$  is latent
- Estimates of  $\beta$  often reflect assumed model

Existing tests we could consider may:

- Depend on unknown parameters,  $\sigma^2$ ,  $\tau^2$
- Not offer guidance given evidence of misspecification

## A More General Model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$

$$p(\beta_j) \propto \exp \left\{ - \left( \frac{\Gamma(1/q)}{\Gamma(3/q)} \right)^{-q/2} \left| \frac{\beta_j}{\tau} \right|^q \right\}$$

## A More General Model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$

$$p(\beta_j) \propto \exp \left\{ - \left( \frac{\Gamma(1/q)}{\Gamma(3/q)} \right)^{-q/2} \left| \frac{\beta_j}{\tau} \right|^q \right\}$$

- $p(\beta_j)$  is density of the exponential power distribution

## A More General Model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$

$$p(\beta_j) \propto \exp \left\{ - \left( \frac{\Gamma(1/q)}{\Gamma(3/q)} \right)^{-q/2} \left| \frac{\beta_j}{\tau} \right|^q \right\}$$

- $p(\beta_j)$  is density of the exponential power distribution
- Special cases include:

## A More General Model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$

$$p(\beta_j) \propto \exp \left\{ - \left( \frac{\Gamma(1/q)}{\Gamma(3/q)} \right)^{-q/2} \left| \frac{\beta_j}{\tau} \right|^q \right\}$$

- $p(\beta_j)$  is density of the exponential power distribution
- Special cases include:

$q = 1$ : Laplace model/lasso penalty



## A More General Model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$

$$p(\beta_j) \propto \exp \left\{ - \left( \frac{\Gamma(1/q)}{\Gamma(3/q)} \right)^{-q/2} \left| \frac{\beta_j}{\tau} \right|^q \right\}$$

- $p(\beta_j)$  is density of the **exponential power distribution**
- Special cases include:
  - $q = 1$ : Laplace model/lasso penalty
  - $q = 2$ : Normal distribution/ridge penalty

## A More General Model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$

$$p(\beta_j) \propto \exp \left\{ - \left( \frac{\Gamma(1/q)}{\Gamma(3/q)} \right)^{-q/2} \left| \frac{\beta_j}{\tau} \right|^q \right\}$$

- $p(\beta_j)$  is density of the **exponential power distribution**
- Special cases include:
  - $q = 1$ : Laplace model/lasso penalty
  - $q = 2$ : Normal distribution/ridge penalty
  - $q \leq 1$ : Bridge penalty

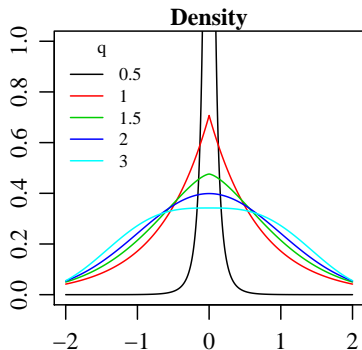
## A More General Model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sigma\mathbf{z}, \quad \mathbf{z} \sim \text{normal}(\mathbf{0}, \mathbf{I}_n)$$

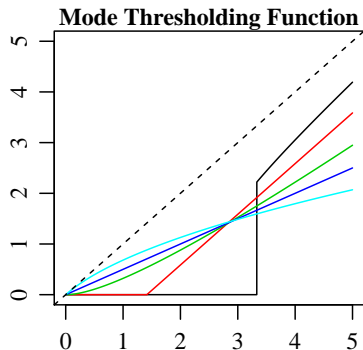
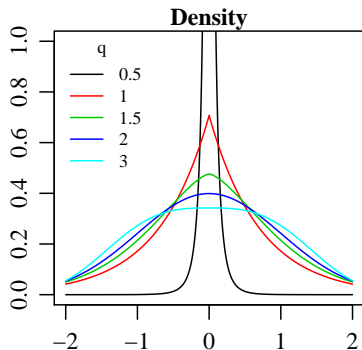
$$p(\beta_j) \propto \exp \left\{ - \left( \frac{\Gamma(1/q)}{\Gamma(3/q)} \right)^{-q/2} \left| \frac{\beta_j}{\tau} \right|^q \right\}$$

- $p(\beta_j)$  is density of the **exponential power distribution**
- Special cases include:
  - $q = 1$ : Laplace model/lasso penalty
  - $q = 2$ : Normal distribution/ridge penalty
  - $q \leq 1$ : Bridge penalty
  - $q \rightarrow \infty$ : Uniform distribution on  $(-\sqrt{3}\tau, \sqrt{3}\tau)$

# Flexibility of the Exponential Power Distribution



# Flexibility of the Exponential Power Distribution



# How does the exponential power distribution help?

Relates **uninterpretable**  $q$  to **interpretable**

# How does the exponential power distribution help?

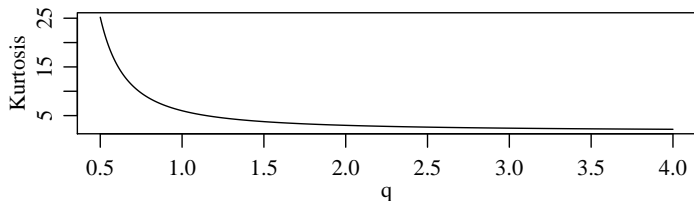
Relates **uninterpretable**  $q$  to **interpretable**

$$\underbrace{\kappa + 3 = \mathbb{E}[\beta_j^4] / \mathbb{E}[\beta_j^2]^2}_{\text{kurtosis}}$$

# How does the exponential power distribution help?

Relates **uninterpretable**  $q$  to **interpretable**

$$\underbrace{\kappa + 3 = \mathbb{E}[\beta_j^4] / \mathbb{E}[\beta_j^2]^2}_{\text{kurtosis}}$$

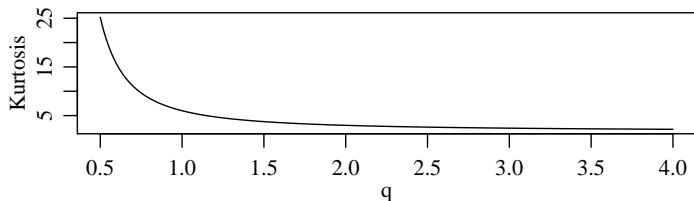




# How does the exponential power distribution help?

Relates **uninterpretable**  $q$  to **interpretable**

$$\underbrace{\kappa + 3 = \mathbb{E}[\beta_j^4] / \mathbb{E}[\beta_j^2]^2}_{\text{kurtosis}}$$



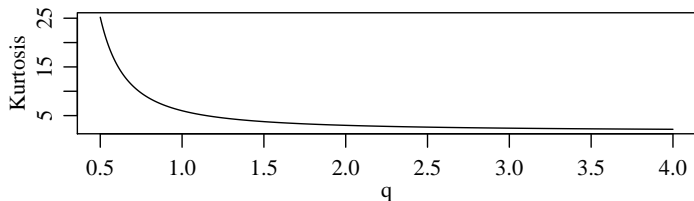
Facilitates:

- Testing  $q = 1$

# How does the exponential power distribution help?

Relates **uninterpretable**  $q$  to **interpretable**

$$\underbrace{\kappa + 3 = \mathbb{E}[\beta_j^4] / \mathbb{E}[\beta_j^2]^2}_{\text{kurtosis}}$$



Facilitates:

- Testing  $q = 1$
- Specifying a “better” model if evidence suggests  $q \neq 1$

Is it easy to test the null hypothesis that  $q = 1$ ?

Is it easy to test the null hypothesis that  $q = 1$ ?

## Likelihood Ratio Test

Is it easy to test the null hypothesis that  $q = 1$ ?

### Likelihood Ratio Test

- Needs maximum marginal likelihood estimation of  $\sigma^2$ ,  $\tau^2$ ,  $q$

Is it easy to test the null hypothesis that  $q = 1$ ?

### Likelihood Ratio Test

- Needs maximum marginal likelihood estimation of  $\sigma^2$ ,  $\tau^2$ ,  $q$
- Elements of  $\mathbf{y}$  are not marginally independent

Is it easy to test the null hypothesis that  $q = 1$ ?

### Likelihood Ratio Test

- Needs maximum marginal likelihood estimation of  $\sigma^2, \tau^2, q$
- Elements of  $\mathbf{y}$  are not marginally independent
- Null distribution of test statistic may depend on  $\sigma^2, \tau^2$

Is it easy to test the null hypothesis that  $q = 1$ ?

### Likelihood Ratio Test

- Needs maximum marginal likelihood estimation of  $\sigma^2, \tau^2, q$
- Elements of  $\mathbf{y}$  are not marginally independent
- Null distribution of test statistic may depend on  $\sigma^2, \tau^2$

### Score Test



Is it easy to test the null hypothesis that  $q = 1$ ?

### Likelihood Ratio Test

- Needs maximum marginal likelihood estimation of  $\sigma^2$ ,  $\tau^2$ ,  $q$
- Elements of  $\mathbf{y}$  are not marginally independent
- Null distribution of test statistic may depend on  $\sigma^2$ ,  $\tau^2$

### Score Test

- Needs estimation of  $\sigma^2$  and  $\tau^2$  under Laplace prior

Is it easy to test the null hypothesis that  $q = 1$ ?

## Likelihood Ratio Test

- Needs maximum marginal likelihood estimation of  $\sigma^2$ ,  $\tau^2$ ,  $q$
- Elements of  $\mathbf{y}$  are not marginally independent
- Null distribution of test statistic may depend on  $\sigma^2$ ,  $\tau^2$

## Score Test

- Needs estimation of  $\sigma^2$  and  $\tau^2$  under Laplace prior
- Elements of  $\mathbf{y}$  are not marginally independent

Is it easy to test the null hypothesis that  $q = 1$ ?

### Likelihood Ratio Test

- Needs maximum marginal likelihood estimation of  $\sigma^2$ ,  $\tau^2$ ,  $q$
- Elements of  $\mathbf{y}$  are not marginally independent
- Null distribution of test statistic may depend on  $\sigma^2$ ,  $\tau^2$

### Score Test

- Needs estimation of  $\sigma^2$  and  $\tau^2$  under Laplace prior
- Elements of  $\mathbf{y}$  are not marginally independent
- Null distribution of test statistic may depend on  $\sigma^2$ ,  $\tau^2$

Is it easy to test the null hypothesis that  $q = 1$ ?

## Likelihood Ratio Test

- Needs maximum marginal likelihood estimation of  $\sigma^2$ ,  $\tau^2$ ,  $q$
- Elements of  $\mathbf{y}$  are not marginally independent
- Null distribution of test statistic may depend on  $\sigma^2$ ,  $\tau^2$

## Score Test

- Needs estimation of  $\sigma^2$  and  $\tau^2$  under Laplace prior
- Elements of  $\mathbf{y}$  are not marginally independent
- Null distribution of test statistic may depend on  $\sigma^2$ ,  $\tau^2$

Relationship between kurtosis and  $q \rightarrow$

Is it easy to test the null hypothesis that  $q = 1$ ?

## Likelihood Ratio Test

- Needs maximum marginal likelihood estimation of  $\sigma^2$ ,  $\tau^2$ ,  $q$
- Elements of  $\mathbf{y}$  are not marginally independent
- Null distribution of test statistic may depend on  $\sigma^2$ ,  $\tau^2$

## Score Test

- Needs estimation of  $\sigma^2$  and  $\tau^2$  under Laplace prior
- Elements of  $\mathbf{y}$  are not marginally independent
- Null distribution of test statistic may depend on  $\sigma^2$ ,  $\tau^2$

Relationship between kurtosis and  $q \rightarrow$  [Moment-Based Test](#)

## An Approximate Level- $\alpha$ Test for $q = 1$ , Full Rank $\mathbf{X}$

**Oracle Test:** Given  $\beta$ , use  $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  as test statistic

## An Approximate Level- $\alpha$ Test for $q = 1$ , Full Rank $\mathbf{X}$

**Oracle Test:** Given  $\beta$ , use  $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  as test statistic

- Obtain quantiles  $s_{\alpha/2}$ ,  $s_{1-\alpha/2}$  of  $\psi$  under null

# An Approximate Level- $\alpha$ Test for $q = 1$ , Full Rank $\mathbf{X}$

**Oracle Test:** Given  $\beta$ , use  $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  as test statistic

- Obtain quantiles  $s_{\alpha/2}$ ,  $s_{1-\alpha/2}$  of  $\psi$  under null
  - Distribution of  $\psi$  is independent of  $\tau^2$ , can be simulated



## An Approximate Level- $\alpha$ Test for $q = 1$ , Full Rank $\mathbf{X}$

**Oracle Test:** Given  $\beta$ , use  $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  as test statistic

- Obtain quantiles  $s_{\alpha/2}$ ,  $s_{1-\alpha/2}$  of  $\psi$  under null
  - Distribution of  $\psi$  is independent of  $\tau^2$ , can be simulated

# An Approximate Level- $\alpha$ Test for $q = 1$ , Full Rank $\mathbf{X}$

**Oracle Test:** Given  $\beta$ , use  $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  as test statistic

- Obtain quantiles  $s_{\alpha/2}$ ,  $s_{1-\alpha/2}$  of  $\psi$  under null
  - Distribution of  $\psi$  is independent of  $\tau^2$ , can be simulated
- Reject when  $\psi \notin (s_{\alpha/2}, s_{1-\alpha/2})$

# An Approximate Level- $\alpha$ Test for $q = 1$ , Full Rank $\mathbf{X}$

**Oracle Test:** Given  $\beta$ , use  $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  as test statistic

- Obtain quantiles  $s_{\alpha/2}$ ,  $s_{1-\alpha/2}$  of  $\psi$  under null
  - Distribution of  $\psi$  is independent of  $\tau^2$ , can be simulated
- Reject when  $\psi \notin (s_{\alpha/2}, s_{1-\alpha/2})$

**Approximate Test:** Construct  $\mathbf{b} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{y}$ ,  $\hat{\psi} = \frac{\bar{b}^{(4)}}{(\bar{b}^{(2)})^2}$

# An Approximate Level- $\alpha$ Test for $q = 1$ , Full Rank $\mathbf{X}$

**Oracle Test:** Given  $\beta$ , use  $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  as test statistic

- Obtain quantiles  $s_{\alpha/2}$ ,  $s_{1-\alpha/2}$  of  $\psi$  under null
  - Distribution of  $\psi$  is independent of  $\tau^2$ , can be simulated
- Reject when  $\psi \notin (s_{\alpha/2}, s_{1-\alpha/2})$

**Approximate Test:** Construct  $\mathbf{b} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{y}$ ,  $\hat{\psi} = \frac{\bar{b}^{(4)}}{(\bar{b}^{(2)})^2}$

- Reject when  $\hat{\psi} \notin (s_{\alpha/2}, s_{1-\alpha/2})$

# An Approximate Level- $\alpha$ Test for $q = 1$ , Full Rank $\mathbf{X}$

**Oracle Test:** Given  $\beta$ , use  $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  as test statistic

- Obtain quantiles  $s_{\alpha/2}$ ,  $s_{1-\alpha/2}$  of  $\psi$  under null
  - Distribution of  $\psi$  is independent of  $\tau^2$ , can be simulated
- Reject when  $\psi \notin (s_{\alpha/2}, s_{1-\alpha/2})$

**Approximate Test:** Construct  $\mathbf{b} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{y}$ ,  $\hat{\psi} = \frac{\bar{b}^{(4)}}{(\bar{b}^{(2)})^2}$

- Reject when  $\hat{\psi} \notin (s_{\alpha/2}, s_{1-\alpha/2})$
- Valid when  $\text{tr} \left( (\mathbf{X}'\mathbf{X})^{-1} \right) \rightarrow 0$

# An Approximate Level- $\alpha$ Test for $q = 1$ , Full Rank $\mathbf{X}$

**Oracle Test:** Given  $\beta$ , use  $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  as test statistic

- Obtain quantiles  $s_{\alpha/2}$ ,  $s_{1-\alpha/2}$  of  $\psi$  under null
  - Distribution of  $\psi$  is independent of  $\tau^2$ , can be simulated
- Reject when  $\psi \notin (s_{\alpha/2}, s_{1-\alpha/2})$

**Approximate Test:** Construct  $\mathbf{b} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{y}$ ,  $\hat{\psi} = \frac{\bar{b}^{(4)}}{(\bar{b}^{(2)})^2}$

- Reject when  $\hat{\psi} \notin (s_{\alpha/2}, s_{1-\alpha/2})$
- Valid when  $\text{tr} \left( (\mathbf{X}'\mathbf{X})^{-1} \right) \rightarrow 0$ 
  - $\text{tr} \left( (\mathbf{X}'\mathbf{X})^{-1} \right) = \sum_{j=1}^p \frac{1}{\eta_j}$ , i.e.  $\mathbf{X}'\mathbf{X}$  not ill-conditioned

**Extension:**  $n < p$ , Ill-Conditioned  $\mathbf{X}'\mathbf{X}$

## Extension: $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

Let  $\mathbf{V} = \sqrt{\text{diag}\{\mathbf{X}'\mathbf{X}\}}$  and  $\mathbf{C} = \mathbf{V}^{-1}\mathbf{X}'\mathbf{X}\mathbf{V}^{-1}$



## Extension: $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

Let  $\mathbf{V} = \sqrt{\text{diag}\{\mathbf{X}'\mathbf{X}\}}$  and  $\mathbf{C} = \mathbf{V}^{-1}\mathbf{X}'\mathbf{X}\mathbf{V}^{-1}$

$$\bullet \mathbf{b}_{\delta, \mathbf{X}} = \overbrace{\mathbf{V}^{-1}(\mathbf{C} + \delta^2 \mathbf{I}_p)^{-1}\mathbf{V}^{-1}}^D \mathbf{X}'\mathbf{y}$$

## Extension: $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

Let  $\mathbf{V} = \sqrt{\text{diag}\{\mathbf{X}'\mathbf{X}\}}$  and  $\mathbf{C} = \mathbf{V}^{-1}\mathbf{X}'\mathbf{X}\mathbf{V}^{-1}$

- $\mathbf{b}_{\delta,\mathbf{X}} = \overbrace{\mathbf{V}^{-1}(\mathbf{C} + \delta^2\mathbf{I}_p)^{-1}\mathbf{V}^{-1}}^{\mathbf{D}} \mathbf{X}'\mathbf{y}$
- $\mathbf{b}_{\delta,\mathbf{X}} - \underbrace{\mathbf{D}\mathbf{X}'\mathbf{X}\boldsymbol{\beta}}_{\boldsymbol{\beta}_{\delta,\mathbf{X}}} \sim \text{normal}(\mathbf{0}, \sigma^2\mathbf{D}\mathbf{X}'\mathbf{X}\mathbf{D})$

## Extension: $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

Let  $\mathbf{V} = \sqrt{\text{diag}\{\mathbf{X}'\mathbf{X}\}}$  and  $\mathbf{C} = \mathbf{V}^{-1}\mathbf{X}'\mathbf{X}\mathbf{V}^{-1}$

- $\mathbf{b}_{\delta,\mathbf{X}} = \overbrace{\mathbf{V}^{-1}(\mathbf{C} + \delta^2\mathbf{I}_p)^{-1}\mathbf{V}^{-1}}^{\mathbf{D}} \mathbf{X}'\mathbf{y}$
- $\mathbf{b}_{\delta,\mathbf{X}} - \underbrace{\mathbf{D}\mathbf{X}'\mathbf{X}\boldsymbol{\beta}}_{\boldsymbol{\beta}_{\delta,\mathbf{X}}} \sim \text{normal}(\mathbf{0}, \sigma^2\mathbf{D}\mathbf{X}'\mathbf{X}\mathbf{D})$

**Approximate Test:** Define  $\psi_{\delta,\mathbf{X}} = \frac{\bar{\beta}_{\delta,\mathbf{X}}^{(4)}}{(\bar{\beta}_{\delta,\mathbf{X}}^{(2)})^2}$

## Extension: $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

Let  $\mathbf{V} = \sqrt{\text{diag}\{\mathbf{X}'\mathbf{X}\}}$  and  $\mathbf{C} = \mathbf{V}^{-1}\mathbf{X}'\mathbf{X}\mathbf{V}^{-1}$

- $b_{\delta,\mathbf{X}} = \overbrace{\mathbf{V}^{-1}(\mathbf{C} + \delta^2\mathbf{I}_p)^{-1}\mathbf{V}^{-1}}^{\mathbf{D}} \mathbf{X}'\mathbf{y}$
- $b_{\delta,\mathbf{X}} - \underbrace{\mathbf{D}\mathbf{X}'\mathbf{X}\boldsymbol{\beta}}_{\boldsymbol{\beta}_{\delta,\mathbf{X}}} \sim \text{normal}(\mathbf{0}, \sigma^2 \mathbf{D}\mathbf{X}'\mathbf{X}\mathbf{D})$

**Approximate Test:** Define  $\psi_{\delta,\mathbf{X}} = \frac{\bar{\beta}_{\delta,\mathbf{X}}^{(4)}}{(\bar{\beta}_{\delta,\mathbf{X}}^{(2)})^2}$

- Obtain quantiles of  $\psi_{\delta,\mathbf{X}}$  under null

## Extension: $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

Let  $\mathbf{V} = \sqrt{\text{diag}\{\mathbf{X}'\mathbf{X}\}}$  and  $\mathbf{C} = \mathbf{V}^{-1}\mathbf{X}'\mathbf{X}\mathbf{V}^{-1}$

- $b_{\delta,\mathbf{X}} = \overbrace{\mathbf{V}^{-1}(\mathbf{C} + \delta^2\mathbf{I}_p)^{-1}\mathbf{V}^{-1}}^{\mathbf{D}} \mathbf{X}'\mathbf{y}$
- $b_{\delta,\mathbf{X}} - \underbrace{\mathbf{D}\mathbf{X}'\mathbf{X}\boldsymbol{\beta}}_{\boldsymbol{\beta}_{\delta,\mathbf{X}}} \sim \text{normal}(\mathbf{0}, \sigma^2 \mathbf{D}\mathbf{X}'\mathbf{X}\mathbf{D})$

**Approximate Test:** Define  $\psi_{\delta,\mathbf{X}} = \frac{\bar{\beta}_{\delta,\mathbf{X}}^{(4)}}{(\bar{\beta}_{\delta,\mathbf{X}}^{(2)})^2}$

- Obtain quantiles of  $\psi_{\delta,\mathbf{X}}$  under null
  - Distribution of  $\psi_{\delta,\mathbf{X}}$  still independent of  $\tau^2$ , can simulate

## Extension: $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

Let  $\mathbf{V} = \sqrt{\text{diag}\{\mathbf{X}'\mathbf{X}\}}$  and  $\mathbf{C} = \mathbf{V}^{-1}\mathbf{X}'\mathbf{X}\mathbf{V}^{-1}$

- $\mathbf{b}_{\delta,\mathbf{X}} = \overbrace{\mathbf{V}^{-1}(\mathbf{C} + \delta^2\mathbf{I}_p)^{-1}\mathbf{V}^{-1}}^{\mathbf{D}} \mathbf{X}'\mathbf{y}$
- $\mathbf{b}_{\delta,\mathbf{X}} - \underbrace{\mathbf{D}\mathbf{X}'\mathbf{X}\boldsymbol{\beta}}_{\boldsymbol{\beta}_{\delta,\mathbf{X}}} \sim \text{normal}(\mathbf{0}, \sigma^2 \mathbf{D}\mathbf{X}'\mathbf{X}\mathbf{D})$

**Approximate Test:** Define  $\psi_{\delta,\mathbf{X}} = \frac{\bar{\beta}_{\delta,\mathbf{X}}^{(4)}}{(\bar{\beta}_{\delta,\mathbf{X}}^{(2)})^2}$

- Obtain quantiles of  $\psi_{\delta,\mathbf{X}}$  under null
  - Distribution of  $\psi_{\delta,\mathbf{X}}$  still independent of  $\tau^2$ , can simulate
- Compute  $\hat{\psi}_{\delta,\mathbf{X}} = \frac{\bar{b}_{\delta,\mathbf{X}}^{(4)}}{(\bar{b}_{\delta,\mathbf{X}}^{(2)})^2}$ , compare to quantiles of  $\psi_{\delta,\mathbf{X}}$

Choice of  $\delta^2$  for  $n < p$ , Ill-Conditioned  $\mathbf{X}'\mathbf{X}$

Choice of  $\delta^2$  for  $n < p$ , Ill-Conditioned  $\mathbf{X}'\mathbf{X}$

As  $\delta \rightarrow \infty$



## Choice of $\delta^2$ for $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

As  $\delta \rightarrow \infty$

☺  $\hat{\psi}_{\delta, \mathbf{X}} \rightarrow \psi_{\delta, \mathbf{X}}$ , nominal level more likely to be achieved

## Choice of $\delta^2$ for $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

As  $\delta \rightarrow \infty$

- ☺  $\hat{\psi}_{\delta, \mathbf{X}} \rightarrow \psi_{\delta, \mathbf{X}}$ , nominal level more likely to be achieved
- ☹  $\mathbf{b}_{\delta, \mathbf{X}} \rightarrow \mathbf{0}$ , power decreases

## Choice of $\delta^2$ for $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

As  $\delta \rightarrow \infty$

- ☺  $\hat{\psi}_{\delta, \mathbf{X}} \rightarrow \psi_{\delta, \mathbf{X}}$ , nominal level more likely to be achieved
- ☹  $\mathbf{b}_{\delta, \mathbf{X}} \rightarrow \mathbf{0}$ , power decreases

**Our Recommendation:**

## Choice of $\delta^2$ for $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

As  $\delta \rightarrow \infty$

- ☺  $\hat{\psi}_{\delta, \mathbf{X}} \rightarrow \psi_{\delta, \mathbf{X}}$ , nominal level more likely to be achieved
- ☹  $\mathbf{b}_{\delta, \mathbf{X}} \rightarrow \mathbf{0}$ , power decreases

### Our Recommendation:

Letting  $\eta_1, \dots, \eta_p$  be eigenvalues of  $\mathbf{C}$

## Choice of $\delta^2$ for $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

As  $\delta \rightarrow \infty$

- ☺  $\hat{\psi}_{\delta, \mathbf{X}} \rightarrow \psi_{\delta, \mathbf{X}}$ , nominal level more likely to be achieved
- ☹  $\mathbf{b}_{\delta, \mathbf{X}} \rightarrow \mathbf{0}$ , power decreases

### Our Recommendation:

Letting  $\eta_1, \dots, \eta_p$  be eigenvalues of  $\mathbf{C}$

$$\delta^2 = (1 - \min_j \eta_j)_+$$

## Choice of $\delta^2$ for $n < p$ , Ill-Conditioned $\mathbf{X}'\mathbf{X}$

As  $\delta \rightarrow \infty$

- ☺  $\hat{\psi}_{\delta, \mathbf{X}} \rightarrow \psi_{\delta, \mathbf{X}}$ , nominal level more likely to be achieved
- ☹  $\mathbf{b}_{\delta, \mathbf{X}} \rightarrow \mathbf{0}$ , power decreases

### Our Recommendation:

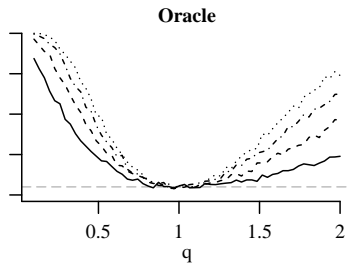
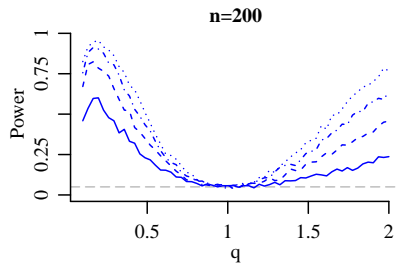
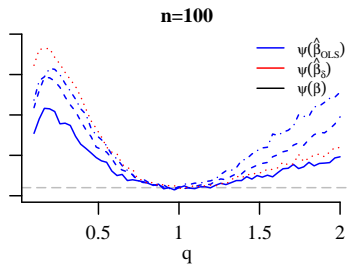
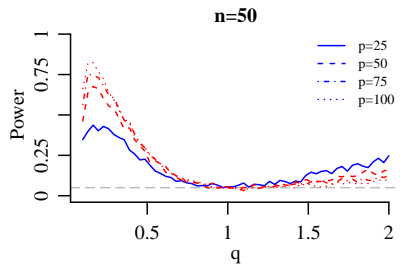
Letting  $\eta_1, \dots, \eta_p$  be eigenvalues of  $\mathbf{C}$

$$\delta^2 = (1 - \min_j \eta_j)_+$$

If columns of  $\mathbf{X}$  standardized to have norm  $n$ ,

$$\text{tr} \left( \underbrace{\mathbf{D}\mathbf{X}'\mathbf{X}\mathbf{D}}_{\text{Variance of } \boldsymbol{\beta}_{\delta, \mathbf{X}}} \right) \leq \frac{1}{n} \sum_{j=1}^p \frac{1}{\delta^2 + \eta_j}$$

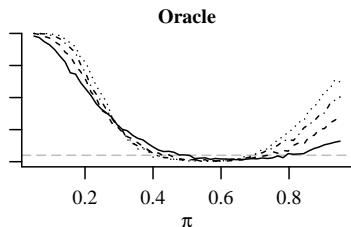
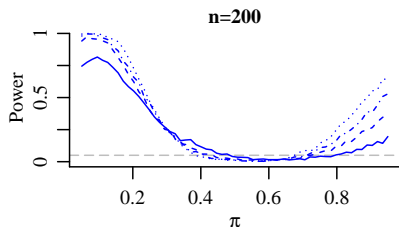
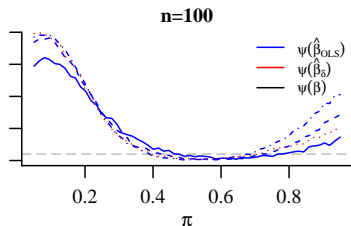
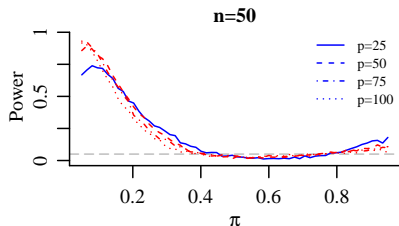
# Test Performance: Exponential Power $\beta$



# Test Performance: Spike-and-Slab $\beta$

Laplace prior often used when  $\beta$  may be **sparse**

$$p(\beta_j \neq 0) = \pi, \beta_j | \beta_j \neq 0 \sim \text{normal}(0, \tau^2/\pi), \kappa = 3 \left( \frac{1-\pi}{\pi} \right)$$





# Specifying a “Better” Model After Rejection

- 1 Compute  $\hat{q}$

# Specifying a “Better” Model After Rejection

- 1 Compute  $\hat{q}$
- 2 Compute unbiased  $\hat{\tau}^2$ ,  $\hat{\sigma}^2$  from  $\mathbf{b}_{\delta, \mathbf{X}}$  and residuals

# Specifying a “Better” Model After Rejection

- 1 Compute  $\hat{q}$
- 2 Compute unbiased  $\hat{\tau}^2$ ,  $\hat{\sigma}^2$  from  $\mathbf{b}_{\delta, \mathbf{X}}$  and residuals
- 3 Compute  $\hat{\boldsymbol{\beta}}$  by:

# Specifying a “Better” Model After Rejection

- ① Compute  $\hat{q}$
- ② Compute unbiased  $\hat{\tau}^2$ ,  $\hat{\sigma}^2$  from  $\mathbf{b}_{\delta, \mathbf{X}}$  and residuals
- ③ Compute  $\hat{\boldsymbol{\beta}}$  by:
  - Solving for posterior mode

# Specifying a “Better” Model After Rejection

- 1 Compute  $\hat{q}$
- 2 Compute unbiased  $\hat{\tau}^2$ ,  $\hat{\sigma}^2$  from  $\mathbf{b}_{\delta, \mathbf{X}}$  and residuals
- 3 Compute  $\hat{\boldsymbol{\beta}}$  by:
  - Solving for posterior mode (**nonconvex** when  $q < 1$ )

# Specifying a “Better” Model After Rejection

- ① Compute  $\hat{q}$
- ② Compute unbiased  $\hat{\tau}^2$ ,  $\hat{\sigma}^2$  from  $\mathbf{b}_{\delta, \mathbf{X}}$  and residuals
- ③ Compute  $\hat{\boldsymbol{\beta}}$  by:
  - Solving for posterior mode (**nonconvex** when  $q < 1$ )
  - Obtain posterior mean/median via Gibbs sampling

## Specifying a “Better” Model After Rejection

**Idea:** Use  $\hat{\psi}_{\delta, \mathbf{X}}$  to estimate  $\kappa + 3$

## Specifying a “Better” Model After Rejection

**Idea:** Use  $\hat{\psi}_{\delta, \mathbf{X}}$  to estimate  $\kappa + 3$

- $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  is a method of moments estimator of  $\kappa + 3$



## Specifying a “Better” Model After Rejection

**Idea:** Use  $\hat{\psi}_{\delta, \mathbf{X}}$  to estimate  $\kappa + 3$

- $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  is a method of moments estimator of  $\kappa + 3$

→ Use  $\hat{\psi} = \widehat{\kappa + 3}$

# Specifying a “Better” Model After Rejection

**Idea:** Use  $\hat{\psi}_{\delta, \mathbf{X}}$  to estimate  $\kappa + 3$

- $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  is a method of moments estimator of  $\kappa + 3$

→ Use  $\hat{\psi} = \widehat{\kappa + 3}$

When we use  $\psi_{\delta, \mathbf{X}} = \frac{\bar{\beta}_{\delta, \mathbf{X}}^{(4)}}{(\bar{\beta}_{\delta, \mathbf{X}}^{(2)})^2}$ , need to correct for bias of  $\beta_{\delta, \mathbf{X}}$ :

# Specifying a “Better” Model After Rejection

**Idea:** Use  $\hat{\psi}_{\delta, \mathbf{X}}$  to estimate  $\kappa + 3$

- $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  is a method of moments estimator of  $\kappa + 3$

→ Use  $\hat{\psi} = \widehat{\kappa + 3}$

When we use  $\psi_{\delta, \mathbf{X}} = \frac{\bar{\beta}_{\delta, \mathbf{X}}^{(4)}}{(\bar{\beta}_{\delta, \mathbf{X}}^{(2)})^2}$ , need to correct for bias of  $\beta_{\delta, \mathbf{X}}$ :

$$\widehat{\kappa + 3} = \left( \frac{\alpha_{\delta, \mathbf{X}}^2}{\gamma_{\delta, \mathbf{X}}} \right) \left( \hat{\psi}_{\delta, \mathbf{X}} - \frac{\omega_{\delta, \mathbf{X}}}{\alpha_{\delta, \mathbf{X}}^2} \right)$$

# Specifying a “Better” Model After Rejection

**Idea:** Use  $\hat{\psi}_{\delta, \mathbf{X}}$  to estimate  $\kappa + 3$

- $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  is a method of moments estimator of  $\kappa + 3$

→ Use  $\hat{\psi} = \widehat{\kappa + 3}$

When we use  $\psi_{\delta, \mathbf{X}} = \frac{\bar{\beta}_{\delta, \mathbf{X}}^{(4)}}{(\bar{\beta}_{\delta, \mathbf{X}}^{(2)})^2}$ , need to correct for bias of  $\beta_{\delta, \mathbf{X}}$ :

$$\widehat{\kappa + 3} = \left( \frac{\alpha_{\delta, \mathbf{X}}^2}{\gamma_{\delta, \mathbf{X}}} \right) \left( \hat{\psi}_{\delta, \mathbf{X}} - \frac{\omega_{\delta, \mathbf{X}}}{\alpha_{\delta, \mathbf{X}}^2} \right)$$

- Reduces to  $\hat{\psi}$  when  $\delta = 0$

# Specifying a “Better” Model After Rejection

**Idea:** Use  $\hat{\psi}_{\delta, \mathbf{X}}$  to estimate  $\kappa + 3$

- $\psi = \frac{\bar{\beta}^{(4)}}{(\bar{\beta}^{(2)})^2}$  is a method of moments estimator of  $\kappa + 3$

→ Use  $\hat{\psi} = \widehat{\kappa + 3}$

When we use  $\psi_{\delta, \mathbf{X}} = \frac{\bar{\beta}_{\delta, \mathbf{X}}^{(4)}}{(\bar{\beta}_{\delta, \mathbf{X}}^{(2)})^2}$ , need to correct for bias of  $\beta_{\delta, \mathbf{X}}$ :

$$\widehat{\kappa + 3} = \left( \frac{\alpha_{\delta, \mathbf{X}}^2}{\gamma_{\delta, \mathbf{X}}} \right) \left( \hat{\psi}_{\delta, \mathbf{X}} - \frac{\omega_{\delta, \mathbf{X}}}{\alpha_{\delta, \mathbf{X}}^2} \right)$$

- Reduces to  $\hat{\psi}$  when  $\delta = 0$
- Can obtain  $\hat{q}$  via  $\widehat{\kappa + 3} = \Gamma(\frac{1}{\hat{q}})\Gamma(\frac{5}{\hat{q}})/\Gamma(\frac{3}{\hat{q}})^2$

# Estimating Variance Components

# Estimating Variance Components

- Gibbs-within-EM needed for maximum marginal likelihood

# Estimating Variance Components

- Gibbs-within-EM needed for maximum marginal likelihood
  - Can be slow to converge



# Estimating Variance Components

- Gibbs-within-EM needed for maximum marginal likelihood
  - Can be slow to converge
  - Scales slowly with  $p$

# Estimating Variance Components

- Gibbs-within-EM needed for maximum marginal likelihood
  - Can be slow to converge
  - Scales slowly with  $p$
  - Estimates of  $\sigma^2$ ,  $\tau^2$  will depend on  $q$

# Estimating Variance Components

- Gibbs-within-EM needed for maximum marginal likelihood
  - Can be slow to converge
  - Scales slowly with  $p$
  - Estimates of  $\sigma^2$ ,  $\tau^2$  will depend on  $q$
- Moment-based estimators can be negative

# Estimating Variance Components

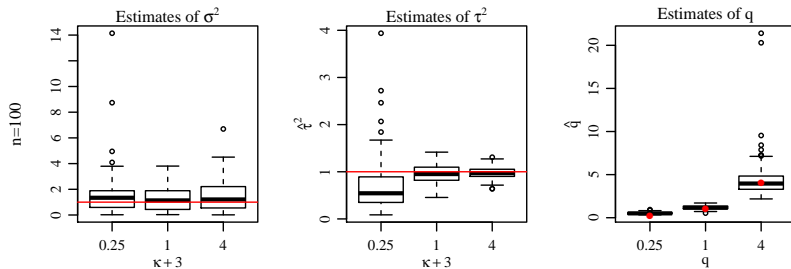
- Gibbs-within-EM needed for maximum marginal likelihood
  - Can be slow to converge
  - Scales slowly with  $p$
  - Estimates of  $\sigma^2$ ,  $\tau^2$  will depend on  $q$
- Moment-based estimators can be negative
- Approximate  $\sigma^2$ ,  $\tau^2$  obtained via MML for normal  $\beta$

# Estimating Variance Components

- Gibbs-within-EM needed for maximum marginal likelihood
  - Can be slow to converge
  - Scales slowly with  $p$
  - Estimates of  $\sigma^2$ ,  $\tau^2$  will depend on  $q$
- Moment-based estimators can be negative
- Approximate  $\sigma^2$ ,  $\tau^2$  obtained via MML for normal  $\beta$

$$\min_{\tau^2, \sigma^2} \log \left( \left| \mathbf{X} \mathbf{X}^\top \tau^2 + \mathbf{I}_n \sigma^2 \right| \right) + \text{tr} \left( \mathbf{y} \mathbf{y}^\top \left( \mathbf{X} \mathbf{X}^\top \tau^2 + \mathbf{I}_n \sigma^2 \right)^{-1} \right)$$

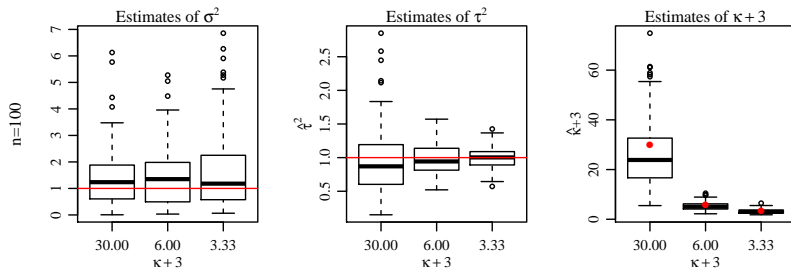
# Adaptive Simulations: Exponential Power $\beta$



# Adaptive Simulations: Exponential Power $\beta$

$p$	100		
$q$	1/4	1	4
Null Hypothesis Rejected	83	5	69
MSE APM <sub>e</sub> $\leq$ MSE LPM <sub>e</sub>	<b>97</b>	<b>97</b>	<b>95</b>
MSE APM <sub>e</sub> $\leq$ MSE EPM <sub>e</sub>	<b>85</b>	<b>85</b>	50
MSE APM <sub>e</sub> $\leq$ MSE BB	<b>100</b>	<b>100</b>	49
MSE APM <sub>e</sub> $\leq$ MSE CV	<b>52</b>	<b>52</b>	<b>63</b>

# Adaptive Simulations: Spike-and-slab $\beta$





# Adaptive Simulations: Spike-and-slab $\beta$

$p$	100		
$\pi$	0.1	0.5	0.9
Pow. ADMo $\geq$ Pow. LPMo	<b>68</b>	<b>98</b>	<b>100</b>
Pow. ADMo $\geq$ Pow. EPMo	<b>100</b>	42	31
Pow. ADMo $\geq$ Pow. CV	<b>84</b>	<b>97</b>	<b>95</b>
FDR ADMo $\leq$ FDR LPMo	<b>100</b>	<b>100</b>	<b>75</b>
FDR ADMo $\leq$ FDR EPMo	<b>93</b>	<b>70</b>	<b>97</b>
FDR ADMo $\leq$ FDR CV	<b>70</b>	9	24
CS ADMo $\geq$ CS LPMo	<b>100</b>	<b>100</b>	<b>98</b>
CS ADMo $\geq$ CS EPMo	<b>93</b>	<b>70</b>	31
CS ADMo $\geq$ CS CV	<b>70</b>	16	<b>94</b>

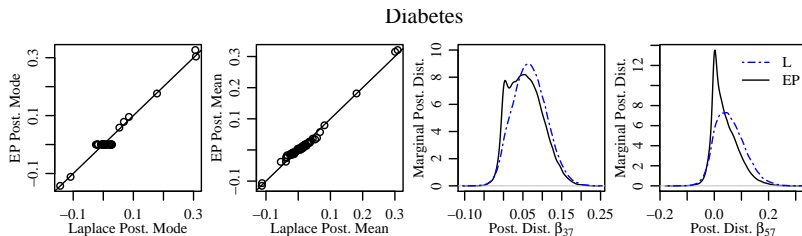
# Applications: Canonical Examples

Dataset	$n$	$p$	$\psi_{\delta,0.025}$	$\psi_{\delta,0.975}$	$\psi(\hat{\beta}_{\delta})$
Diabetes	422	64	2.31	7.68	10.36
Boston Housing	506	104	1.97	7.61	6.57
Motif	287	195	2.87	10.35	5.77
Glucose	68	72	2.31	7.06	7.99

# Application: Diabetes Data with Interactions

- $y$  is a measure of diabetes progression for  $n = 442$
- $p = 64$  main effects and interactions for 10 covariates

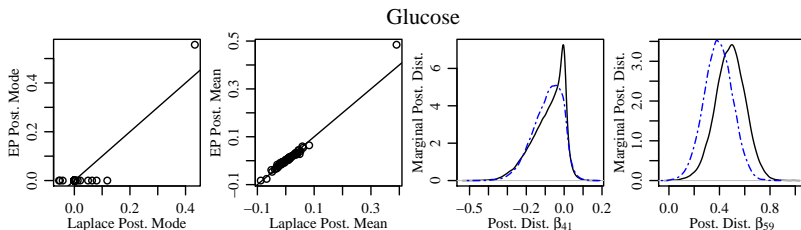
Dataset	Par. Ests.			Mode Sparsity	
	$\hat{\sigma}^2$	$\hat{\tau}^2$	$\hat{q}$	$L$	$EP$
Diabetes	0.4720	0.0071	0.5505	50.0%	87.5%



# Application: Glucose Data

- $y$  is blood glucose concentration  $y$  for  $n = 68$  subjects
- $p = 72$  health indicators, e.g. metabolite measurements

Dataset	Par. Ests.			Mode Sparsity	
	$\hat{\sigma}^2$	$\hat{\tau}^2$	$\hat{q}$	$L$	$EP$
Glucose	0.4754	0.0070	0.5939	83.3%	98.6%



## Recap: Testing Laplace Model/Lasso Penalty

We've shown:

- We can construct a test of the Laplace model/lasso penalty

## Recap: Testing Laplace Model/Lasso Penalty

We've shown:

- We can construct a test of the Laplace model/lasso penalty
- Test statistic can guide model selection after rejection

## Recap: Testing Laplace Model/Lasso Penalty

We've shown:

- We can construct a test of the Laplace model/lasso penalty
- Test statistic can guide model selection after rejection

Our approach has natural extensions to:

## Recap: Testing Laplace Model/Lasso Penalty

We've shown:

- We can construct a test of the Laplace model/lasso penalty
- Test statistic can guide model selection after rejection

Our approach has natural extensions to:

- Testing appropriateness of different distributions for  $\beta$



## Recap: Testing Laplace Model/Lasso Penalty

We've shown:

- We can construct a test of the Laplace model/lasso penalty
- Test statistic can guide model selection after rejection

Our approach has natural extensions to:

- Testing appropriateness of different distributions for  $\beta$
- Generalized linear models for  $\mathbf{y}$

## Recap: Testing Laplace Model/Lasso Penalty

We've shown:

- We can construct a test of the Laplace model/lasso penalty
- Test statistic can guide model selection after rejection

Our approach has natural extensions to:

- Testing appropriateness of different distributions for  $\beta$
- Generalized linear models for  $\mathbf{y}$
- Nonnormal  $\mathbf{z}$

# “Improved Pathwise Coordinate Descent for Power...”

## Problem

# “Improved Pathwise Coordinate Descent for Power...”

## Problem

Solving  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_q^q$  is hard

# “Improved Pathwise Coordinate Descent for Power...”

## Problem

Solving  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_q^q$  is hard

☹ No closed form solution for arbitrary  $q$ , even when  $\boldsymbol{\beta} = \beta$

# “Improved Pathwise Coordinate Descent for Power...”

## Problem

Solving  $\min_{\boldsymbol{\beta}} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_q^q$  is hard

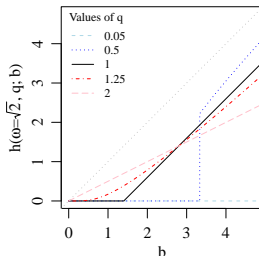
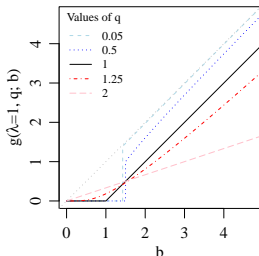
- ☹ No closed form solution for arbitrary  $q$ , even when  $\boldsymbol{\beta} = \beta$
- ☹ Solutions for arbitrary  $q$  aren't nested

# “Improved Pathwise Coordinate Descent for Power...”

## Problem

Solving  $\min_{\beta} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \|\beta\|_q^q$  is hard

- ☹ No closed form solution for arbitrary  $q$ , even when  $\beta = \beta$
- ☹ Solutions for arbitrary  $q$  aren't nested



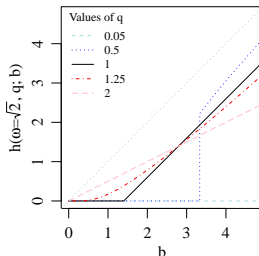
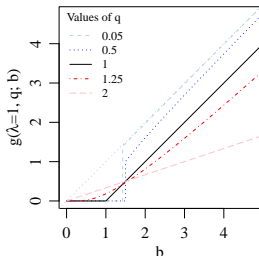
## Solution

# “Improved Pathwise Coordinate Descent for Power...”

## Problem

Solving  $\min_{\beta} \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \|\beta\|_q^q$  is hard

- ☹ No closed form solution for arbitrary  $q$ , even when  $\beta = \beta$
- ☹ Solutions for arbitrary  $q$  aren't nested



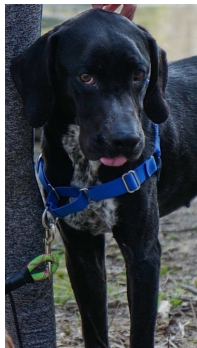
## Solution

- ☺ Reparametrization to  $\frac{1}{2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \left(\frac{\omega^{2-q}}{q}\right) \|\beta\|_q^q$



# Acknowledgements

- My advisor Peter Hoff
- NSF Grants DGE-1256082, DMS-1505136, & DMS-2113079
- My canine assistants



- Box, G. E. P. (1953). A Note on Regions for Tests of Kurtosis. *Biometrika* 40(3/4), 465–468.
- Box, G. E. P. and G. C. Tiao (1973). Bayesian Assessment of Assumptions. In *Bayesian Inference in Statistical Analysis*, Chapter 3, pp. 149–202. Reading, MA: Addison-Wesley Pub. Co.
- Bühlmann, P. and S. van de Geer (2011). *Statistics for High-Dimensional Data: Methods, Theory and Applications*. Heidelberg: Springer.
- Drikvandi, R., G. Verbeke, and G. Molenberghs (2017). Diagnosing misspecification of the random-effects distribution in mixed models. *Biometrics* (March), 63–71.
- Efron, B., T. Hastie, I. Johnstone, and R. Tibshirani (2004). Least Angle Regression. *The Annals of Statistics* 32(2), 407–499.
- Frank, I. E. and J. H. Friedman (1993). A Statistical View of Some Chemometrics Regression Tools. *Technometrics* 35(2), 109.
- Friedman, J., T. Hastie, and R. Tibshirani (2010). Regularization Paths for Generalized Linear Models via Coordinate Descent. *Journal of Statistical Software* 33(1).
- Marjanovic, G. and V. Solo (2014). lq Sparsity Penalized Linear Regression with Cyclic Descent. *IEEE Transactions on Signal Processing* 62(6), 1464–1475.
- Mazumder, R., J. H. Friedman, and T. Hastie (2011). Sparsenet: Coordinate descent with nonconvex penalties. *Journal of the American Statistical Association* 106, 1125–1138.
- Park, T. and G. Casella (2008). The Bayesian Lasso. *Journal of the American Statistical Association* 103(482), 681–686.
- Polson, N. G., J. G. Scott, and J. Windle (2014). The Bayesian bridge. *Journal of the Royal Statistical Society. Series B: Statistical Methodology* 76(4), 713–733.
- Priami, C. and M. J. Morine (2015). *Analysis of Biological Systems*. London: Imperial College Press.
- Subbotin, M. T. (1923). On the Law of Frequency of Error. *Matematicheskii Sbornik* 31(2), 296–301.
- Verbeke, G. and E. Lesaffre (1996). A Linear Mixed-Effects Model With Heterogeneity in the Random-Effects Population. *Journal of the American Statistical Association* 91(433), 217–221.

## Details for Estimating $\hat{q}$

$$\left. \begin{aligned} \mathbb{E}[\bar{\beta}_{\delta, \mathbf{X}}^{(2)}] &= \alpha_{\delta, \mathbf{X}} \tau^2 \\ \mathbb{E}[\bar{\beta}_{\delta, \mathbf{X}}^{(4)}] &= \gamma_{\delta, \mathbf{X}} (\kappa + 3) \tau^4 + \omega_{\delta, \mathbf{X}} \tau^4 \end{aligned} \right\} \rightarrow \widehat{\kappa + 3} = \left( \frac{\alpha_{\delta, \mathbf{X}}^2}{\gamma_{\delta, \mathbf{X}}} \right) \left( \hat{\psi}_{\delta, \mathbf{X}} - \frac{\omega_{\delta, \mathbf{X}}}{\alpha_{\delta, \mathbf{X}}^2} \right)$$

$$\alpha_{\delta, \mathbf{X}} = \text{tr}(\mathbf{X}' \mathbf{X} \mathbf{D} \mathbf{D} \mathbf{X}' \mathbf{X}) / p$$

$$\gamma_{\delta, \mathbf{X}} = \sum_{j=1}^p (\mathbf{D} \mathbf{X}' \mathbf{X})_{jj}^4 / p$$

$$\omega_{\delta, \mathbf{X}} = 3 \left( \sum_{j=1}^p (\mathbf{X}' \mathbf{X} \mathbf{D} \mathbf{D} \mathbf{X}' \mathbf{X})_{jj}^2 / p - \gamma_{\delta, \mathbf{X}} \right)$$