Estimation of Possibly Non-Stationary Long Memory Processes via Adaptive Overdifferencing

Maryclare Griffin

University of Massachusetts Amherst Department of Mathematics and Statistics

May 20, 2023

Hello and congratulations, Ruev!

Journal of Forecasting, Vol. 13, 109-131 (1994).

Biometrika (1997), 84, 4, np. 791-802 Printed in Great Britain

Some Advances in Non-linear and Adaptive Modelling in Time-series

GEORGE C. TIAO AND RUEY S. TSAY University of Chicago, IL, U.S.A.

Bandwidth selection for kernel regression with long-range dependent errors

BY BONNIE K RAY Center for Applied Mathemat mology, Newark, New Jersey e-mail: borayx@m.njit.edu

Iterative Bandwidth Estimation for Nonparametric Regression with Long-range Dependent Errors Bonnie K. Rav

Ruev S. Tsav

AND RUFY S TSAY

ABSTRACT. We discuss the problem of bandwidth selection for a kernel or when the errors are long-range dependent. The BAYESIAN METHODS FOR CHANGE-POINT DETECTION 3th selection method is investigated and modified ory in the errors. We compare the mean average-IN LONG-RANGE DEPENDENT PROCESSES

By BONNIE K. RAY AND RUEY S. TSAY

and estimates using the bandwidth obtained from ve apply the modified method to estimate a trend

True or Spurious Long Memory? A New Test

Arek OHANISSIAN, Jeffrey R. RUSSELL, and Ruey S. TSAY Graduate School of Business, University of Chicago, Chicago, IL 60637 (ieffrey.russell@chicagogsb.edu)

> It is well known that long memory characteristics observed in data can be generated by nonstationary structural-break or slow regime switching models. We propose a statistical test to distinguish between true long memory and spurious long memory based on invariance of the long memory parameter for temporal aggregates of the process under the null of true long memory. Goweke Porter-Hudak estimates of the long memory parameter obtained from different temporal aggregates of the underlying time series are shown to be asymptotically jointly normal, leading to a test statistic that is constructed as the quadratic form of a demeaned vector of the estimates. The result is a test statistic that is very simple to implement, Simulations show the test to have good size and power properties for the classic alternatives to true long memory that have been suggested in the literature. The asymptotic distribution of the test statistic is also valid for a stochastic volatility with Gaussian long memory model. The test is applied to foreign exchange rate data. Based on all the models considered in this article, we conclude that the long memory property in exchange rate volatility is generated by a true long memory process.

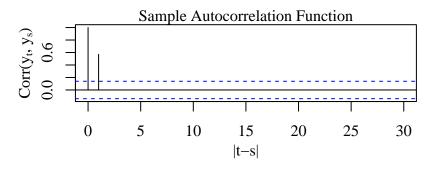
KEY WORDS: Regime switching: Structural change: Temporal appregation.

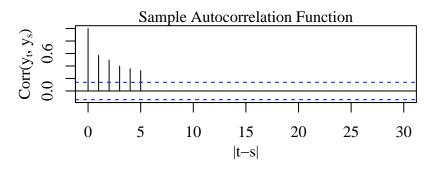
Long-range Dependence in Daily Stock Volatilities

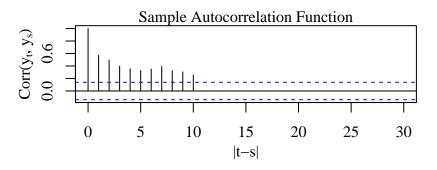
Department of Mathematical Sciences and Center for Applied Math and Statistics. New Jersey Institute of Technology, Newark, NJ 07102 (borayx@m.n)it.edu)

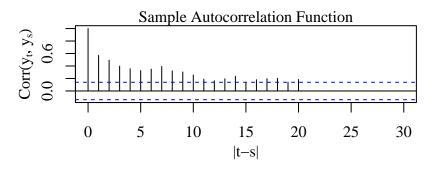
Ruev S. TSAY

Graduate School of Business, University of Chicago, Chicago, IL. 60637 (ruev travial net hicago, edit)









$$(y_t - \mu_t) = \sum_{j=1}^{p} \phi_j (y_{t-j} - \mu_{t-j}) + z_t$$
$$z_t \stackrel{i.i.d.}{\sim} \text{normal} (0, \sigma^2)$$

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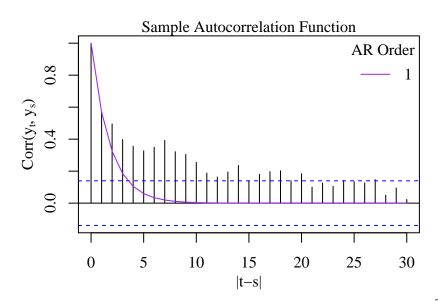
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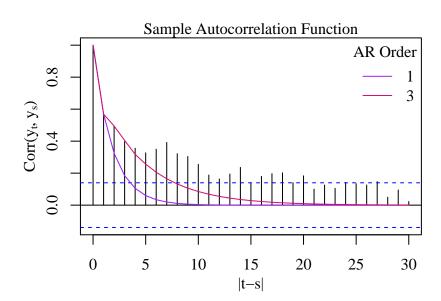
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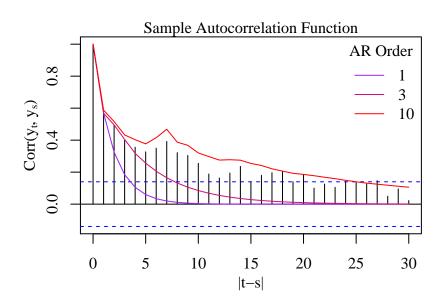
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In practice, long memory requires $p \to \infty$ as $n \to \infty$







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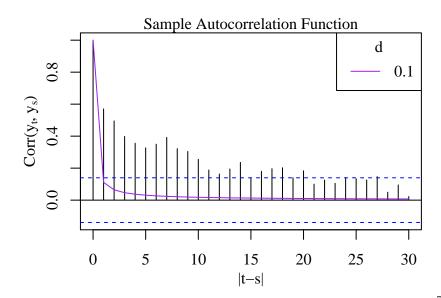
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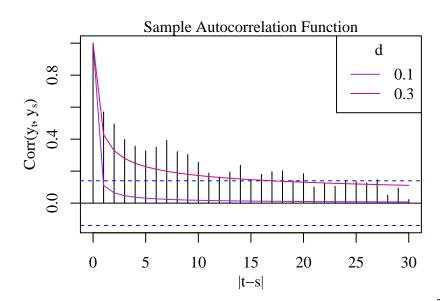
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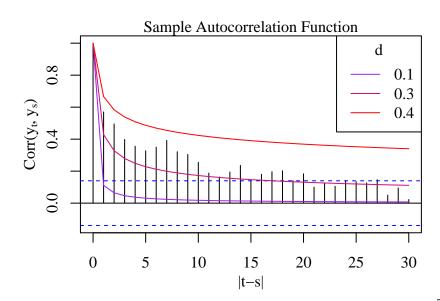
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Unique values of $d \leq -0.5$ give unique likelihood

Granger and Joyeux (1980); Hosking (1981)







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(Content applies to ARFIMA(p, d, q) as well)

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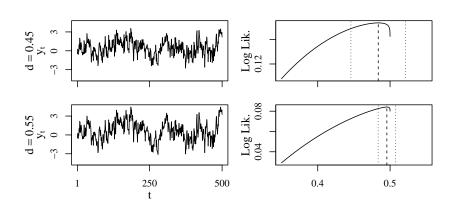
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Why do we care about assumptions on d?



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$$(1-B)^{2} (y_{t} - \mu_{t}) = y_{t} - 2y_{t-1} + y_{t-2} - (\mu_{t} - 2\mu_{t-1} + \mu_{t-2})$$

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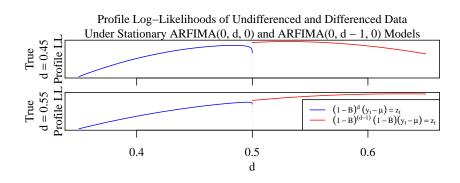
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What's wrong with this procedure?



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• **Problem:** Data changes as d changes

• Solution: Fix data in advance to get stationarity for $d < \bar{d}$

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$$(1-B)^m (y_t - \mu_t)$$
 stationary, constrained ARFIMA $(0, d-m+r, r)$

Odaki (1993); Doornik and Ooms (2003)

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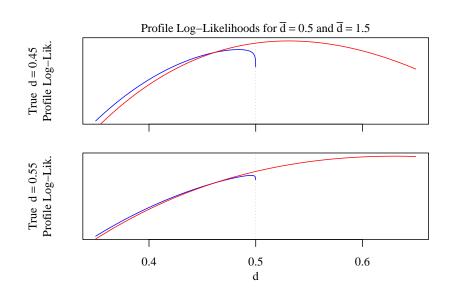
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This gives a <u>continuous</u> likelihood in d!



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 $\Pr(\hat{d}_{\bar{d}} > \bar{d}) \approx \epsilon$, small ϵ facilitates:

- Construction of CI's for d with nominal coverage
- Implementation of tests of d = c with nominal level

When $\epsilon = 0.5$ we'll call estimator **BND**

Need a principled/optimal way to choose \bar{d}

Choose smallest \bar{d} for which:

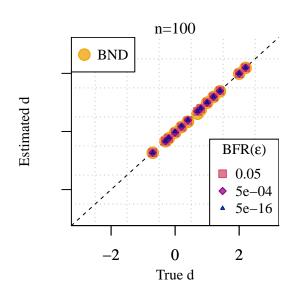
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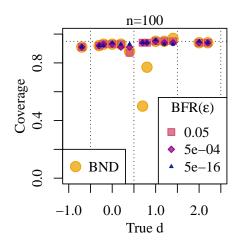
- Construction of CI's for d with nominal coverage
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When $\epsilon = 0.5$ we'll call estimator **BND** When $\epsilon < 0.5$, we'll call estimator **BFR**(ϵ)

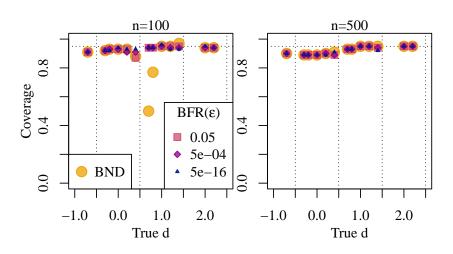
Adaptive Procedure Simulations



Adaptive Procedure Simulations

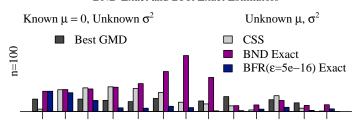


Adaptive Procedure Simulations



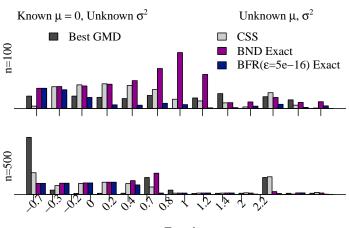
Adaptive Procedure vs. Competition

Mean Absolute Bias $|\hat{d} - d|$ of CSS and Best GMD Estimators vs. BND Exact and BFR Exact Estimators



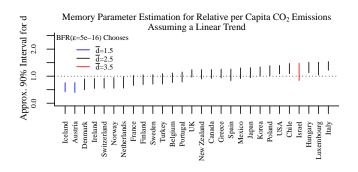
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Application to CO² Emission Convergence

$$(1-B)^{d_c} (y_{tc} - \mu_c - \beta_c t) = z_{tc}, \quad z_{tc} \stackrel{i.i.d.}{\sim} \text{normal} (0, \sigma_c^2),$$



Barassi et al. (2018)

We took a difficult problem and turned it into an easy problem!

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Maximizing over possibly non-stationary FI models

≈ Maximizing a sequence of stationary ARFIMA models

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Maximizing over possibly non-stationary FI models \approx Maximizing a sequence of stationary ARFIMA models

Everything we have done:

• Extends to ARFIMA(p, d, q)

$$\phi(B)(1-B)^d(y_t-\mu_t) = \theta(B)z_t, z_t \stackrel{i.i.d.}{\sim} \text{normal}(0,\sigma^2)$$

- * Optimization is harder
- * Likelihood may be multimodal

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Thank you to:

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