

# Multivariate linear Regression

We have  $r$   $n \times 1$  response vectors  $[y_1, \dots, y_r] = Y$   
and  $n \times q$  covariate matrix  $X = [\tilde{x}_1, \dots, \tilde{x}_q]$

we want to find  $B_{q \times r}$  such that  $Y \approx XB$ .

by minimizing  $\min_B \|Y - XB\|_2^2 = \sum_{x=1}^r \sum_{i=1}^n (y_{ix} - x_i' b_x)^2$

$$\|Y\|_2^2 = \sum_{x=1}^r \sum_{i=1}^n y_{ix}^2$$

↑ residual sum of squares

Differentiating with respect to  $B$  yields

$$(X'X)\hat{B} = X'Y \Rightarrow \hat{B} = (X'X)^{-1}X'Y$$

(if  $X$  is full rank)

## Model Formulation of Multivariate Linear Regression

Assume 
$$Y_{n \times r} = X_{n \times q} B_{q \times r} + W_{n \times r}$$

\* If  $E[W] = 0$  then  $\hat{B}$  is unbiased

\* If  $\tilde{w}_i \stackrel{iid}{\sim} N(\mathbf{0}, \Sigma_w)$  where  $\tilde{w}_i$  are rows of  $W$  then:

- $\hat{B}$  is the maximum likelihood estimator of  $B$
- Elements of  $\hat{B}$  are normally distributed

with  $V[\hat{\tilde{b}}_e] = \sigma_{w,ee} (X'X)^{-1}$

$\text{Cov}[\hat{\tilde{b}}_e, \hat{\tilde{b}}_k] = \sigma_{w,ek} (X'X)^{-1}$

(where  $\tilde{b}_e$  is the  $e$ -th column of  $B$ )

\* We know the distribution of  $R = Y - X\hat{B}$

\* Residuals  $R$  and  $\hat{B}$  are independent

all can use  
of this  
to do  
testing  
etc.

There are equivalent definitions of AIC/AICC/BIC

$$\hat{\Sigma}_w = \frac{R'R}{n} \quad \left. \vphantom{\frac{R'R}{n}} \right\} \begin{array}{l} \text{maximum likelihood} \\ \text{estimator of the noise variance} \end{array}$$

$$\text{AIC: } \ln \left( |\hat{\Sigma}_w| \right) + \frac{2}{n} \left( rk + \frac{r(r+1)}{2} \right)$$

$$\text{AICC: } \ln \left( |\hat{\Sigma}_w| \right) + \frac{r(n+q)}{n-r-q-1}$$

$$\text{BIC/SIC: } \ln \left( |\hat{\Sigma}_w| \right) + (kr + r(r+1)/2) \times \log(n)/n$$

# Multivariate Time Series Concepts

Again, suppose we observe  $Y = [y_1, \dots, y_r]$

We can always decompose  $y = \underset{\substack{\uparrow \\ \text{fixed mean}}}{M} y + \underset{\substack{\downarrow \\ \text{random errors}}}{W}$

Each column of  $Y$  comprised of  $n$  equally spaced observations ordered in time

characterized by mean function  $E[y_{ij}] = m_{ij}$  and covariance function  $\gamma_{ij}(s, t) = \text{cov}[y_{si}, y_{tj}]$

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when  $i \neq j$ , we call  $\gamma_{ij}(s,t)$  the cross-covariance function of  $y_i$  and  $y_j$

Assume joint stationarity to simplify things in the multivariate setting

\* Second moments of  $y_{ti}$  are finite for all  $t$  and  $i$ ,  $E[y_{ti}^2] < \infty$

\* The mean function is constant for each time series and doesn't depend on time

$$\mu_{ti} = \mu_i$$

\* The autocovariance function  $\gamma_{ii}(s, t)$  depends on  $s$  and  $t$  only through their absolute

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