

# Non-Monotonicity of the Tversky-Kahneman Probability-Weighting Function: A Cautionary Note

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## Abstract

*Cumulative Prospect Theory has gained a great deal of support as an alternative to Expected Utility Theory as it accounts for a number of anomalies in the observed behavior of economic agents. Expected Utility Theory uses a utility function and subjective or objective probabilities to compare risky prospects. Cumulative Prospect Theory alters both of these aspects. The concave utility function is replaced by a loss-averse utility function and probabilities are replaced by decision weights. The latter are determined with a weighting function applied to the cumulative probability of the outcomes. Several different probability weighting functions have been suggested. The two most popular are the original proposal of Tversky and Kahneman and the compound-invariant form proposed by Prelec. This note shows that the Tversky-Kahneman probability weighting function is not increasing for all parameter values and therefore can assign negative decision weights to some outcomes. This in turn implies that Cumulative Prospect Theory could make choices not consistent with first-order stochastic dominance.*

**Keywords:** *prospect theory, decision weights, probability-weighting function*

**JEL classification:** C91, D10, D81, G19

Cumulative Prospect Theory (CPT) has been used in a wide variety of economic models as an alternative to Expected Utility Theory (EUT). It is used to account for a number of anomalies in the observed behaviour of economic agents. Like EUT, CPT uses a utility or value function to rate any given outcome. In CPT, however, the utility function is loss averse (S-shaped) rather than risk averse (concave). The other difference between the two theories is the use of decision weights rather than probabilities in computing the expectation.

This note provides a caution on the use of a popular form of decision weights – those proposed by Tversky and Kahneman (1992) in their introduction of the

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theory. In particular, the weights generated by their proposed probability weighting function can be negative for some parameter values. This could lead to evaluations that choose first-order stochastically dominated gambles rather than the dominating one.

The parameter values that can cause such problems are small relative to estimated values. Nevertheless, failure to recognise this problem may prevent the proofs of some theorems or allow the proof of some theorems which would be invalid with properly restricted probability weights.

### The Tversky-Kahneman Probability-Weighting Function

In Prospect Theory, the ‘expected’ utility of any risky prospect with outcomes  $x_i$  is computed as

$$\sum v(x_i)\omega_i(\mathbf{p}) \quad (1)$$

where  $v$  is a value (utility) function and  $\omega_i$  are a set of decision weights. In the original application of Prospect Theory, each decision weight was a function of the probability of the outcome; in CPT, decisions weights depend on the probability distribution. They are determined as first differences of a probability-weighting function applied to the cumulative probabilities,  $p_i$ ,  $\omega_i = \Omega(p_i) - \Omega(p_{i-1})$ .<sup>1</sup>

The particular probability-weighting function originally proposed by Tversky and Kahneman (1992) is

$$\Omega(p) = \frac{p^\gamma}{[p^\gamma + (1-p)^\gamma]^{1/\gamma}} \quad 0 < \gamma \leq 1 \quad (2)$$

where  $p$  is the cumulative probability of the distribution of gains or losses. As shown in Figure 1, this weighting function has an inverted S-shape and is subproportional.<sup>2</sup> The latter property allows for the Allais paradox and can explain the preference for lottery tickets.

Since the decision weights take the place of probabilities in the computation of ‘expected’ utility, they must be positive. If they are not positive, then a first-order stochastically dominated gamble might be preferred to the dominating gamble.<sup>3</sup> The decision weights are first differences and will be positive for all possible risky prospects only if the weighting function is strictly increasing. But the Tversky-Kahneman probability-weighting function does not possess this property for all parameter values.

Let  $F(p) \equiv \ln[\Omega(p)]$ . Since the logarithm is a strictly-increasing function, the probability-weighting function,  $\Omega$ , is strictly increasing in  $p$  if and only if  $F$  is strictly

<sup>1</sup> In CPT, separate weighting functions are used for the probability distribution of losses and the complementary distribution of gains, but this is irrelevant to the issue presented here which is about the properties of the weighting function itself. If  $p(x)$  is a cumulative probability function for a continuous distribution with density  $p'(x)$ , then the continuous decision weighting density function is  $\omega(x) = \Omega'(p)p'(x)$ .

<sup>2</sup> A function  $\Omega$  is subproportional if for  $p > q$  and  $0 < \lambda < 1$ ,  $\Omega(\lambda q)/\Omega(q) > \Omega(\lambda p)/\Omega(p)$ .

<sup>3</sup> Let  $x > x'$  and consider two risky prospects which are identical except for the inclusion of either  $x$  or  $x'$  as the  $i$ th outcome. Now suppose that  $\omega_i < 0$ , then the dominated gamble including  $x'$  will be preferred when using decision weights.

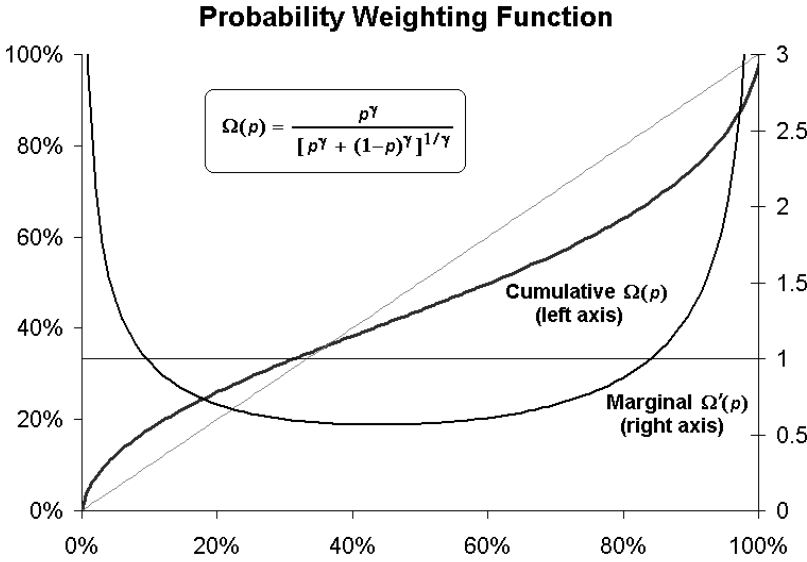


Fig. 1. Tversky-Kahneman probability weighting function and marginal

increasing in  $p$ . The derivative of  $F(p)$  is

$$F'(p) = \frac{\gamma}{p} - \frac{p^{\gamma-1} - (1-p)^{\gamma-1}}{p^{\gamma} + (1-p)^{\gamma}} = \frac{\gamma p^{\gamma} + \gamma(1-p)^{\gamma} - p^{\gamma} + p(1-p)^{\gamma-1}}{p[p^{\gamma} + (1-p)^{\gamma}]}. \quad (3)$$

Since the denominator of the last fraction is positive,  $F'$  is negative whenever the numerator is negative. Dividing the numerator by  $(1-p)^{\gamma}$ , we see that  $F'$  is negative when

$$(\gamma - 1)\frac{p^{\gamma}}{(1-p)^{\gamma}} + \gamma + \frac{p}{1-p} < 0. \quad (4)$$

Define  $x \equiv p/(1-p)$ , then we are looking for values of the weighting function parameter,  $\gamma$ , for which the function

$$f(x) \equiv (\gamma - 1)x^{\gamma} + x + \gamma \quad (5)$$

can take on negative values for some positive  $x$ . Note that  $f(0) = \gamma > 0$ ,  $f(\infty) = \infty$ , and  $f$  is strictly convex

$$\begin{aligned} f'(x) &= \gamma(\gamma - 1)x^{\gamma-1} + 1 \geq 0 \\ f''(x) &= \gamma(\gamma - 1)^2 x^{\gamma-2} > 0. \end{aligned} \quad (6)$$

Therefore, the function  $f$  will take on negative values over some range when it is negative at its minimum. The minimum occurs at

$$f'(x) = \gamma(\gamma - 1)x^{\gamma-1} + 1 = 0 \Rightarrow x_{\min} = [\gamma(1 - \gamma)]^{1/(1-\gamma)} \quad (7)$$

So the minimum value of  $f$  for a given value of the parameter  $\gamma$  is

$$\begin{aligned} f_{\min}(\gamma) &\equiv f(x_{\min}) = (\gamma - 1)[\gamma(1 - \gamma)]^{\gamma/(1-\gamma)} + [\gamma(1 - \gamma)]^{1/(1-\gamma)} + \gamma \\ &= \gamma [1 - (1 - \gamma)^{(2-\gamma)/(1-\gamma)} \gamma^{(2\gamma-1)/(1-\gamma)}]. \end{aligned} \quad (8)$$

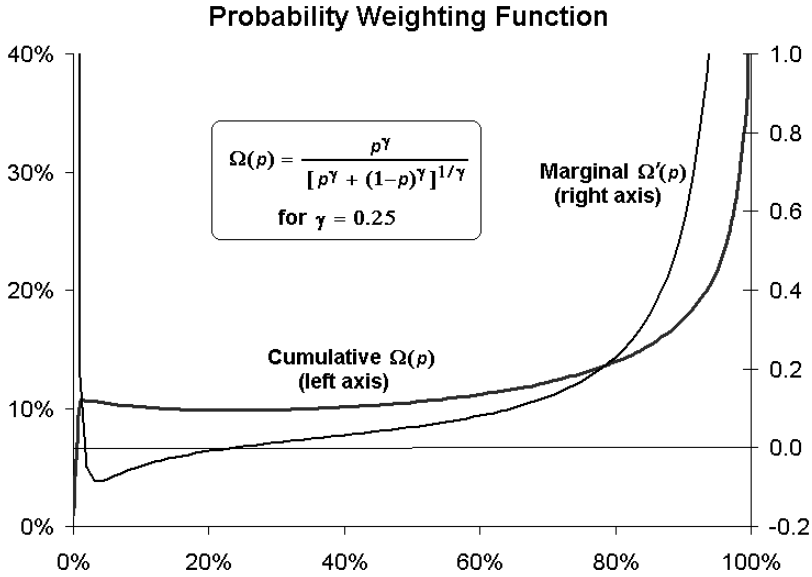


Fig. 2. Tversky-Kahneman probability weighting function illustrating non-monotonicity for  $\gamma < 0.279$

By inspection, there is a critical value,  $\gamma_{\text{critical}}$ , where  $f_{\min} = 0$ . For all values of  $\gamma$  less than this critical value,  $f_{\min} < 0$ , and the probability weighting-function is nonmonotonic. The critical value satisfies

$$1 = (1 - \gamma)^{(2-\gamma)} \gamma^{(2\gamma-1)} \Rightarrow \gamma_{\text{critical}} \approx 0.279. \quad (9)$$

Figure 2 illustrates the Tversky-Kahneman probability-weighting function for a parameter value of  $\gamma = 0.25$ . Note that the derivative is negative over the approximate range  $1.6\% < p < 23.6\%$ . Any continuous probability density will have a corresponding decision weight density which is negative over this range. A discrete distribution with two (or more) cumulative probabilities in this range will have a negative decision weight(s) for the corresponding outcome(s), and could have others. If just one cumulative probability is in this range, the corresponding outcome may have a positive or negative weight.<sup>4</sup>

Tversky and Kahneman (1992) estimated the parameters  $\gamma = 0.69$  for losses and  $\gamma = 0.61$  for gains. Other estimates have varied widely; Camerer and Ho (1994) report on eleven tests whose estimates for  $\gamma$  range from 0.28 to 1.87 with an average value of 0.55. All of these point estimates for  $\gamma$  are above the critical value and would lead to a monotone probability-weighting function and generate only positive decision weights. Nevertheless, the range of estimates reported does not engender great confidence that all future studies would continue to do this.

Probably of more concern, however, is the use of this probability-weighting function in theory. If the Tversky-Kahneman weighting function is posited, and no restriction is placed on the parameter, it is very possible, that some results might not be provable due

<sup>4</sup> Precisely if the cumulative probability,  $p_i$ , for a discrete distribution lies in the range of decreasing  $\Omega$ , then either or both  $\omega_i$  and  $\omega_{i+1}$  could be negative. If  $p_{i-1}$  ( $p_{i+1}$ ) is also in the decreasing range, then  $\omega_i$  ( $\omega_{i+1}$ ) will definitely be negative.

to the possibility of first-order stochastic dominance. Conversely some results might be obtained that could not hold if the weighting function did not generate negative decision weights.

For example, in another paper (Ingersoll, 2007), I have extended the standard result in portfolio theory that in a complete market all risk-averse investors hold portfolios with perfect rank correlation to include investors who use decision weights. This result is only true if the probability-weighting function is monotone.

### Other Probability-Weighting Functions

Several other probability weighting functions have also been proposed. Some of them are:<sup>5</sup>

$$\begin{aligned}
 \Omega(p) &= \frac{p^\gamma}{[p^\gamma + (1-p)^\gamma]^\alpha} && \text{Wu and Gonzalez (1996)} \\
 \Omega(p) &= \frac{\alpha p^\gamma}{\alpha p^\gamma + (1-p)^\gamma} && \text{Lattimore, Baker, and Witte (1992)} \\
 \Omega(p) &= \gamma \exp[-\beta(-\ln p)^\alpha] && \text{Compound-Invariant } \Omega \text{ Prelec (1998)} \\
 \Omega(p) &= \gamma \exp[-\beta(1-p^\eta)/\eta] && \text{Conditionally-Invariant } \Omega \text{ Prelec (1998)} \\
 \Omega(p) &= \gamma (1 - \alpha \ln p)^{-\beta/\alpha} && \text{Projection-Invariant } \Omega \text{ Prelec (1998)}
 \end{aligned} \tag{10}$$

In each case,  $0 < \gamma \leq 1$ ,  $\alpha > 0$ ,  $\beta > 0$ , and  $\eta$  is unconstrained.<sup>6</sup>

The LBW and all of the Prelec probability weighting functions are strictly increasing for all relevant parameter values

$$\frac{d}{dp}[\ln \Omega(p)] = \begin{cases} \frac{\gamma(1-p)^{\gamma-1}}{\alpha p^{\gamma+1} + p(1-p)^\gamma} > 0 & \text{Lattimore, Baker, and Witte (1992)} \\ \frac{\alpha\beta}{p}(-\ln p)^{\alpha-1} > 0 & \text{Compound Invariant} \\ \beta p^{\eta-1} > 0 & \text{Conditionally Invariant} \\ \frac{\gamma\beta}{p\Omega(p)}(1 - \alpha \ln p)^{-\beta/\alpha-1} > 0 & \text{Projection Invariant} \end{cases} \tag{11}$$

They cannot, therefore, assign negative decision weights to any outcomes for any probability distribution.

The Wu and Gonzalez weighting function is also strictly increasing and cannot assign negative decision weights provided  $\alpha \leq 1$ . However, for  $\alpha > 1$ , negative decision weights are possible in some cases. This is clearly true since the TK weighting function is the special case,  $\alpha = 1/\gamma$ .

<sup>5</sup> Karmarkar's (1978) weighting function is included as a special case with  $\alpha = 1$  of either the Wu and Gonzalez (1996) or Lattimore, Baker, and Witte (1992) weighting function. It is strictly increasing for all parameter values just like the latter.

<sup>6</sup> For  $\eta = 0$ , the conditional-invariance probability weighting function is  $\Omega(p) = \gamma p^\beta$ .

## Conclusion

The purpose of this note is not to comment directly on Cumulative Prospect Theory. Rather it is to point out a possible obstacle in its use. Some parameter values of the commonly used probability-weighting function, originally proposed by Tversy and Kahneman, can lead to negative decision weights and the preference for first-order stochastically dominated prospects. While the parameter values which allow this seem to be empirically not relevant, the problem could lead to awkward interpretations of results. In addition, in theoretical work, certain consequences might be found which would not obtain were only positive decision weights allowed, or conversely 'general' propositions might remain unverifiable.

## References

- Camerer, C. F. and Ho, T.-H., 'Violations of the betweenness axiom and nonlinearity in probability', *Journal of Risk and Uncertainty*, 1994, Vol. 8, pp. 167–96.
- Ingersoll, J., 'Theoretical prospects for cumulative prospect theory in optimal portfolios', unpublished working paper (Yale School of Management, 2007).
- Kahneman, D. and Tversky, A., 'Prospect theory: an analysis of decision under risk', *Econometrica*, Vol. 47, 1979, pp. 263–91.
- Karmarkar, U., 'Subjective weighted utility: a descriptive extension of the expected utility model', *Organizational Behavior and Human Performance*, Vol. 21, 1978, pp. 61–72.
- Lattimore, P. K., Baker, J. R. and White, A. D., 'The influence of probability on risky choice: a parametric examination', *Journal of Economic Behavior and Organization*, Vol. 17, 1992, pp. 377–400.
- Prelec, D., 'The probability weighting function', *Econometrica*, Vol. 66, 1998, pp. 497–527.
- Tversky, A. and Kahneman, D., 'Advances in prospect theory: cumulative representation of uncertainty', *Journal of Risk and Uncertainty*, Vol. 5, 1992, pp. 297–323.
- Wu, G. and Gonzalez, R., 'Curvature of the probability weighting function', *Management Science*, Vol. 42, 1996, pp. 1676–90.