Measure Transformation via the Radon-Nikodym Derivative and the Cameron-Martin-Girsanov Theorem

Masaru Okada

October 15, 2025

Contents

1	Current Status and Problem Formulation	2
2	Measure Transformation and the Radon-Nikodym Derivative	2
3	Equivalence	5
4	The Relationship Between Expectation and the Radon-Nikodym Derivative	5
5	The Radon-Nikodym Process	6
6	Example: Discrete Process	7
	6.1 Problem	7
	6.2 Solution	7
7	The Joint Density Function of Brownian Motion	9
8	The Radon-Nikodym Derivative - Continuous Version	9
9	Reviewing the Moment Generating Function	10
10	Simple Measure Transformation (Brownian Motion $+$ Constant Drift)	10
	10.1 Verification for Time $t(< T)$	11
11	Example: Continuous Process	13
	11.1 Solution	14
12	The Cameron-Martin-Girsanov Theorem	14
	12.1 The Cameron-Martin-Girsanov Theorem	15
	12.2 The Converse of the Cameron-Martin-Girsanov Theorem	15
13	The Cameron-Martin-Girsanov Theorem and Stochastic Differentials	15
14	Example - Measure Transformation	16
	14.1 Example 1: A Constant Multiple of a Q-Brownian Motion	16

1 Current Status and Problem Formulation

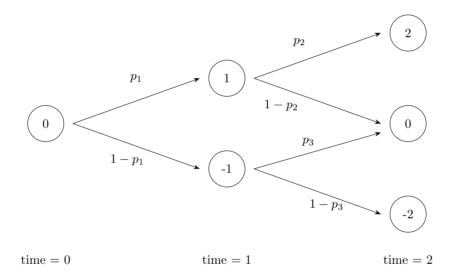
While the previous chapter explored Itō calculus, it didn't explicitly consider the probability measure.

More specifically, we've focused on Itō calculus for a \mathbb{P} -Brownian motion W_t , but we still don't have a way to handle a Brownian motion (written as \tilde{W}_t in the textbook) under a measure \mathbb{Q} that isn't independent of \mathbb{P} . So, we'll now expand our knowledge on methods for moving back and forth between \mathbb{P} and \mathbb{Q} .

We'll be using the textbook: Financial Calculus - An Introduction to Derivative Pricing - Martin Baxter, Andrew Rennie [1].

2 Measure Transformation and the Radon-Nikodym Derivative

To build some intuition for the effects of measure transformation, let's start by looking at a discretetime process.



Let's consider a two-period recombining random walk represented by this diagram.

In this figure, the nodes represent values, and the paths show the transition probabilities.

The four possible paths from time 0 to the final time can be expressed by their values as follows:

$$\left\{ \; 0 \;, \; 1 \;, \; 2 \; \right\} \;, \; \left\{ \; 0 \;, \; 1 \;, \; 0 \; \right\} \;, \; \left\{ \; 0 \;, \; -1 \;, \; 0 \; \right\} \;, \; \left\{ \; 0 \;, \; -1 \;, \; -2 \; \right\}$$

Table 1 summarizes these paths and their corresponding probabilities, which we'll call $\pi_1, \pi_2, \pi_3, \pi_4$ in order.

Path	Probability of reaching the final point	
{ 0 , 1 , 2 }	$p_1p_2=\pi_1$	
{ 0 , 1 , 0 }	$p_1(1 - p_2) = \pi_2$	
{ 0 , -1 , 0 }	$(1-p_1)p_3 = \pi_3$	
{ 0 , -1 , -2 }	$(1 - p_1)(1 - p_3) = \pi_4$	

Table 1: Paths and their corresponding probabilities

Following the logic of our previous explanations (and the probability columns in Table 3.1 of the textbook), if p_1, p_2, p_3 are given, then $\pi_1, \pi_2, \pi_3, \pi_4$ can be determined.

Conversely, if p_1, p_2, p_3 are unknown and $\pi_1, \pi_2, \pi_3, \pi_4$ are given, we can solve for p_1, p_2, p_3 by working

backward.

To make this clearer, let's do it with an example:

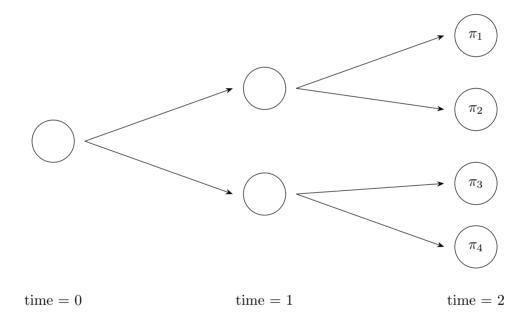
$$\begin{cases}
 p_1 p_2 = \pi_1 \\
 p_1 (1 - p_2) = \pi_2 \\
 (1 - p_1) p_3 = \pi_3 \\
 (1 - p_1) (1 - p_3) = \pi_4
\end{cases}$$

Solving for p_1, p_2, p_3 gives:

$$\begin{cases} p_1 = \pi_1 + \pi_2 \\ p_2 = \frac{\pi_1}{\pi_1 + \pi_2} \\ p_3 = \frac{\pi_3}{\pi_3 + \pi_4} \end{cases}$$

This demonstrates that specifying the final path probabilities $\pi_1, \pi_2, \pi_3, \pi_4$ is equivalent to defining p_1, p_2, p_3 , which in turn defines the measure \mathbb{P} .

The previous diagram showed the path of values (the nodes in the tree represent values), but if we focus on the probabilities of reaching the final point, the tree looks like the following diagram.



Note that while the previous diagram recombined the paths for $\{0, 1, 0\}$ and $\{0, -1, 0\}$ at time 2, this diagram shows them as distinct paths. We've seen that specifying the final path probabilities $\pi_1, \pi_2, \pi_3, \pi_4$ defines the measure \mathbb{P} .

Next, let's consider a measure \mathbb{Q} expressed by transition probabilities q_1, q_2, q_3 instead of p_1, p_2, p_3 . We'll denote the final path probabilities under this measure as $\pi'_1, \pi'_2, \pi'_3, \pi'_4$.

The exact same logic we used for measure \mathbb{P} applies here: defining q_1, q_2, q_3 is the same as determining $\pi'_1, \pi'_2, \pi'_3, \pi'_4$.

■A Summary So Far

For measure \mathbb{P} , if the transition probabilities between nodes, p_1, p_2, p_3 , are given, the final path probabilities $\pi_1, \pi_2, \pi_3, \pi_4$ are determined.

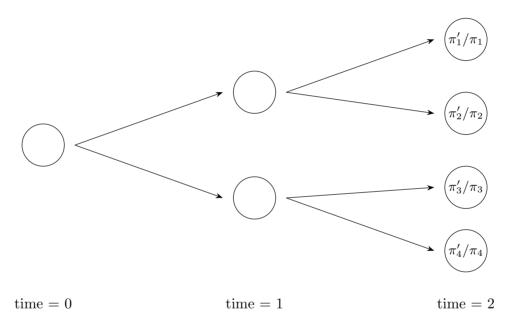
Conversely, if the final path probabilities $\pi_1, \pi_2, \pi_3, \pi_4$ are given, it's equivalent to having the measure \mathbb{P} specified.

The same reasoning applies to measure \mathbb{Q} by simply changing the notation. If the transition probabilities between nodes, q_1, q_2, q_3 , are given, the final path probabilities $\pi'_1, \pi'_2, \pi'_3, \pi'_4$ are determined.

And again, the reverse is also true: if the final path probabilities $\pi_1', \pi_2', \pi_3', \pi_4'$ are given, the measure $\mathbb Q$ is also defined. Let's now consider the ratio of the final path probabilities for measures $\mathbb P$ and $\mathbb Q$: $\frac{\pi_1'}{\pi_1}, \frac{\pi_2'}{\pi_2}, \frac{\pi_3'}{\pi_3}, \frac{\pi_4'}{\pi_4}$. Even if the individual transition probabilities q_1, q_2, q_3 for measure $\mathbb Q$ aren't initially given, if we know

Even if the individual transition probabilities q_1, q_2, q_3 for measure \mathbb{Q} aren't initially given, if we know measure \mathbb{P} and the values of these ratios $\frac{\pi'_i}{\pi_i}$ (for i = 1, 2, 3, 4), we can work backward to determine q_1, q_2, q_3 .

This ratio $\frac{\pi_i'}{\pi_i}$ takes on four different values, one for each final path. If we were to represent this on a tree like before, it would look like the following diagram.



This can be viewed as a random variable whose value at the *i*-th final point is $\frac{\pi'_i}{\pi_i}$.

This random variable is denoted as $\frac{d\mathbb{Q}}{d\mathbb{P}}$ and is called the **Radon-Nikodym derivative** of \mathbb{Q} with respect to \mathbb{P} .

Using this new notation, let's reiterate what we've learned.

■A Summary So Far (Revisited)

If \mathbb{P} and \mathbb{Q} are given, $\frac{d\mathbb{Q}}{d\mathbb{P}}$ is determined. If \mathbb{P} and $\frac{d\mathbb{Q}}{d\mathbb{P}}$ are given, \mathbb{Q} is determined.

3 Equivalence

By definition of probability, each p_i , q_i is a value between 0 and 1, inclusive.

A problem arises when these values are on the boundaries, 0 or 1, since the Radon-Nikodym derivative $\frac{d\mathbb{Q}}{d\mathbb{P}}$ would be undefined.

The values of $\frac{d\mathbb{Q}}{d\mathbb{P}}$ are the ratios of the final path probabilities: $\frac{\pi_1'}{\pi_1}$, $\frac{\pi_2'}{\pi_2}$, $\frac{\pi_3'}{\pi_3}$, $\frac{\pi_4'}{\pi_4}$. If any π_i is 0, we'd have a division by zero, making the expression undefined.

For example, consider the case where $p_1 = 0$ and $q_i > 0$.

Since $\pi_1 = p_1 p_2$ and $\pi_2 = p_1 (1 - p_2)$, multiplying by $p_1 = 0$ makes the final path probabilities π_1 and π_2 both 0.

In this case, the final path probability ratios $\frac{\pi'_1}{\pi_1}$ and $\frac{\pi'_2}{\pi_2}$ don't exist and are therefore undefined.

The textbook [1] states:

'If a path is possible under \mathbb{Q} but impossible under \mathbb{P} , $\frac{d\mathbb{Q}}{d\mathbb{P}}$ cannot be defined.'

To prevent this issue, the concept of equivalence of probability measures is introduced.

■Equivalence of Probability Measures

Two measures \mathbb{P} and \mathbb{Q} on the same probability space are said to be equivalent if for any event in that space, its probability is never 0 under both \mathbb{P} and \mathbb{O} .

If \mathbb{P} and \mathbb{Q} are equivalent, then $\frac{d\mathbb{Q}}{d\mathbb{P}}$ can be defined.

The same logic holds if we swap the symbols \mathbb{P} and \mathbb{Q} . For $\frac{d\mathbb{P}}{d\mathbb{Q}}$ to exist, \mathbb{P} and \mathbb{Q} must be equivalent probability measures.

4 The Relationship Between Expectation and the Radon-Nikodym Derivative

Two sections ago, we saw that if \mathbb{P} and $\frac{d\mathbb{Q}}{d\mathbb{P}}$ are given, \mathbb{Q} can be determined. Let's revisit this by focusing on expectation.

Let's look at the discrete probability measure tree diagram again.

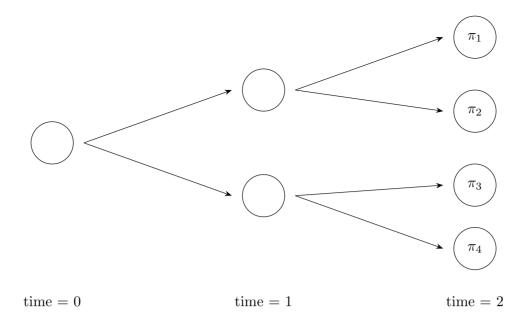
To consider expectations, let's introduce a random variable X. Let x_i be the value of X at the i-th final point (i = 1, 2, 3, 4).

The expectation of X under measure \mathbb{P} , $\mathbf{E}_{\mathbb{P}}(X)$, is calculated as follows:

$$\mathbf{E}_{\mathbb{P}}(X) = \pi_1 x_1 + \pi_2 x_2 + \pi_3 x_3 + \pi_4 x_4$$

Conversely, the expectation of X under measure \mathbb{Q} , $\mathbf{E}_{\mathbb{Q}}(X)$, is calculated as:

$$\mathbf{E}_{\mathbb{Q}}(X) = \pi_1' x_1 + \pi_2' x_2 + \pi_3' x_3 + \pi_4' x_4$$



If we rearrange this to include the ratios $\frac{\pi'_i}{\pi_i}$, we get:

$$\mathbf{E}_{\mathbb{Q}}(X) = \pi_1' x_1 + \pi_2' x_2 + \pi_3' x_3 + \pi_4' x_4$$

$$= \pi_1 \frac{\pi_1'}{\pi_1} x_1 + \pi_2 \frac{\pi_2'}{\pi_2} x_2 + \pi_3 \frac{\pi_3'}{\pi_3} x_3 + \pi_4 \frac{\pi_4'}{\pi_4} x_4$$

This expression is the expectation under \mathbb{P} of the random variable $\frac{d\mathbb{Q}}{d\mathbb{P}}X$, whose value at the *i*-th final point is $\frac{\pi'_i}{\pi_i}x_i$. In summary, we've confirmed that:

$$\mathbf{E}_{\mathbb{Q}}(X) = \mathbf{E}_{\mathbb{P}} \left(\frac{d\mathbb{Q}}{d\mathbb{P}} X \right)$$

5 The Radon-Nikodym Process

The expectation we just looked at wasn't a conditional expectation.

In the two-period binomial model we're considering, time t can only be t = 0, 1, 2. If we set the final time to T(=2) and the random variable at that time to $X = X_T$, we can express our result as a conditional expectation:

$$\mathbf{E}_{\mathbb{Q}}(X_T|\mathcal{F}_0) = \mathbf{E}_{\mathbb{P}}\left(\frac{d\mathbb{Q}}{d\mathbb{P}}X_T\Big|\mathcal{F}_0\right)$$

 (\mathcal{F}_t) , which was introduced in Chapter 2, represents the filtration at time t. In a binomial model, this is specifically the set of all possible paths up to time t.)

From here, we'll see how to express the conditional expectation for more general times $t(\neq T)$ and $s(\neq 0)$, $\mathbf{E}_{\mathbb{Q}}(X_t|\mathcal{F}_s)$, as an expectation under measure \mathbb{P} .

Up to this point, our thinking about the Radon-Nikodym derivative has focused only on the ratio of probabilities at the final time t=T, such as $\frac{\pi'_i}{\pi_i}$ (i=1,2,3,4).

Let's expand this concept and consider the ratio of transition probabilities for each path at times other than t = T.

At time t = 1, the possible ratios of transition probabilities are $\frac{q_1}{p_1}$ or $\frac{1 - q_1}{1 - p_1}$. At time t = 0, since both \mathbb{P} and \mathbb{Q} only have a single starting point with a probability of 1, the ratio

At time t = 0, since both \mathbb{P} and \mathbb{Q} only have a single starting point with a probability of 1, the ratio of transition probabilities is $\frac{1}{1} = 1$.

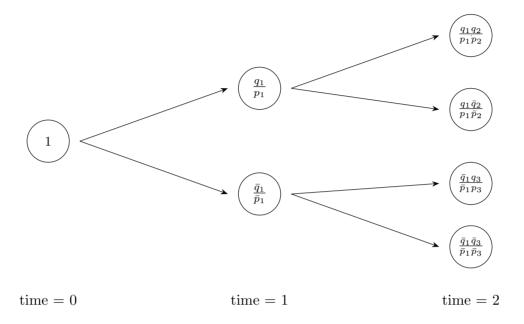
The following diagram shows this in a tree. For simplicity, we've used the notation $1 - p_i = \bar{p}_i$ and $1 - q_i = \bar{q}_i$.

Let's call this stochastic process ζ_t .

At the final time, $\zeta_T = \frac{d\mathbb{Q}}{d\mathbb{P}}$, meaning this is a stochastic process that extends the Radon-Nikodym derivative (which was only defined at time t = T).

At time zero, since the only possible path under both \mathbb{P} and \mathbb{Q} is the single starting point (in the value-based path, this is $\{0\}$) with a probability of 1, we set $\zeta_0 = 1$.

This is known as the Radon-Nikodym process or derivative process.



6 Example: Discrete Process

6.1 Problem

Show for t = 0, 1, 2 that the stochastic process we just defined, ζ_t , can be expressed as the expectation of $\frac{d\mathbb{Q}}{d\mathbb{P}}$ conditioned on \mathcal{F}_t under measure \mathbb{P} :

$$\zeta_t = \mathbf{E}_{\mathbb{P}} \left(\frac{d\mathbb{Q}}{d\mathbb{P}} \middle| \mathcal{F}_t \right)$$

6.2 Solution

6.2.1 Case for t = 2

Since t = 2(=T) is the final point in time, the relationship holds by our very definition.

To double-check, we can refer to the table below. (Note that \mathcal{F}_2 is expressed using the values at each node from the diagram.)

\mathcal{F}_2	ζ_2	$igg \mathbf{E}_{\mathbb{P}} \left(rac{d \mathbb{Q}}{d \mathbb{P}} \Big \mathcal{F}_2 ight)$
$\{1, \frac{q_1}{m}, \frac{\pi'_1}{m}\}$	$\frac{\pi_1'}{}$	$\frac{\pi_1'}{}$
p_1 π_1 p_2	π_1	π_1
$\{1, \frac{q_1}{2}, \frac{\pi'_2}{2}\}$	$\frac{\pi_2'}{}$	$\frac{\pi_2'}{}$
$\lfloor \mid \mid \mid \mid \mid \mid \mid \mid \mid $	π_2	π_2
$\{1, \frac{1-q_1}{2}, \frac{\pi_3'}{2}\}$	$\frac{\pi_3'}{}$	$\frac{\pi_3'}{}$
$\lfloor \lfloor \lfloor \lceil 1 - p_1 \rceil \pi_3 \rfloor \rfloor$	π_3	π_3
$\{1, \frac{1-q_1}{2}, \frac{\pi'_4}{2}\}$	$\frac{\pi_4'}{}$	$\frac{\pi_4'}{}$
$[\ \ \ \ \ \ \ \ \ \ \]$	π_4	π_4

6.2.2 Case for t = 1

Let's check for t = 1.

For the case where $\mathcal{F}_1 = \left\{1, \frac{q_1}{p_1}\right\}$:

$$\mathbf{E}_{\mathbb{P}}\left(\frac{d\mathbb{Q}}{d\mathbb{P}}\Big|\mathcal{F}_{1}\right) = p_{2}\frac{\pi'_{1}}{\pi_{1}} + (1-p_{2})\frac{\pi'_{2}}{\pi_{2}}$$

$$= p_{2}\frac{q_{1}q_{2}}{p_{1}p_{2}} + (1-p_{2})\frac{q_{1}(1-q_{2})}{p_{1}(1-p_{2})}$$

$$= \frac{q_{1}}{p_{1}}\{q_{2} + (1-q_{2})\}$$

$$= \frac{q_{1}}{p_{1}}$$

This matches the value of ζ_1 for the case where $\mathcal{F}_1 = \left\{1, \frac{q_1}{p_1}\right\}$.

Next, for the case where $\mathcal{F}_1 = \left\{1, \frac{1-q_1}{1-p_1}\right\}$:

$$\mathbf{E}_{\mathbb{P}}\left(\frac{d\mathbb{Q}}{d\mathbb{P}}\Big|\mathcal{F}_{1}\right) = p_{3}\frac{\pi_{3}'}{\pi_{3}} + (1-p_{3})\frac{\pi_{4}'}{\pi_{4}}$$

$$= p_{3}\frac{(1-q_{1})q_{3}}{(1-p_{1})p_{3}} + (1-p_{3})\frac{(1-q_{1})(1-q_{3})}{(1-p_{1})(1-p_{3})}$$

$$= \frac{1-q_{1}}{1-p_{1}}\{q_{3}+(1-q_{3})\}$$

$$= \frac{1-q_{1}}{1-p_{1}}$$

This also matches the value of ζ_1 for the case where $\mathcal{F}_1 = \left\{1, \frac{1-q_1}{1-p_1}\right\}$. In summary:

\mathcal{F}_1	ζ_1	$oxed{\mathbf{E}_{\mathbb{P}}\left(rac{d\mathbb{Q}}{d\mathbb{P}}\Big \mathcal{F}_1 ight)}$
$(1, \frac{q_1}{})$	$\frac{q_1}{}$	$\frac{q_1}{}$
p_1	p_1	p_1
$(1 \frac{1-q_1}{1-q_1})$	$\frac{1 - q_1}{}$	$\frac{1-q_1}{}$
$(1, 1-p_1)$	$1 - p_1$	$1 - p_1$

6.2.3 Case for t = 0

$$\mathbf{E}_{\mathbb{P}}\left(\frac{d\mathbb{Q}}{d\mathbb{P}}\middle|\mathcal{F}_{0}\right) = \pi_{1}\frac{\pi'_{1}}{\pi_{1}} + \pi_{2}\frac{\pi'_{2}}{\pi_{2}} + \pi_{3}\frac{\pi'_{3}}{\pi_{3}} + \pi_{4}\frac{\pi'_{4}}{\pi_{4}}$$
$$= \pi'_{1} + \pi'_{2} + \pi'_{3} + \pi'_{4}$$
$$= 1$$

This indeed matches the value of ζ_0 .

We've now shown that for all t, the relationship

$$\zeta_t = \mathbf{E}_{\mathbb{P}} \left(\frac{d\mathbb{Q}}{d\mathbb{P}} \middle| \mathcal{F}_t \right)$$

holds true.

7 The Joint Density Function of Brownian Motion

We've been looking at measure transformations for discrete-time stochastic processes. Now, let's look at continuous-time measure transformations so we can work with Brownian motion.

The text begins by using the standard normal distribution as an example to explain why the concept of "path likelihood" can't be grasped simply by integrating a density function over an interval to find a marginal distribution.

Instead of jumping directly into continuous time, we'll first consider a finite number of discrete time points for simplicity. The probability density function for a \mathbb{P} -Brownian motion W_{t_i} at a given time t_i , denoted as $f_{\mathbb{P}}^i(x)$, is the probability density function for a normal distribution $N(0, t_i)$, which can be written as:

$$f_{\mathbb{P}}^{i}(x) = \frac{1}{\sqrt{2\pi t_{i}}} \exp\left(-\frac{x^{2}}{2t_{i}}\right)$$

Moving forward, let's consider the joint density function $f_{\mathbb{P}}^n(x_1, x_2, ..., x_n)$ for a \mathbb{P} -Brownian motion taking values $x_1, x_2, ..., x_n$ at times $t_1, t_2, ..., t_n$.

Recalling the independent increments property of Brownian motion, the increments $W_{t_2} - W_{t_1}$, $W_{t_3} - W_{t_2}$, ..., $W_{t_n} - W_{t_{n-1}}$ are independent and follow a normal distribution under \mathbb{P} .

For any i greater than 1 (i < n), the density function of $W_{t_i} - W_{t_{i-1}}$ is:

$$\frac{1}{\sqrt{2\pi(t_i - t_{i-1})}} \exp\left(-\frac{(x_i - x_{i-1})^2}{2(t_i - t_{i-1})}\right)$$

where t_0 and x_0 are taken to be 0.

To simplify, if we write $\Delta x_i = x_i - x_{i-1}$ and $\Delta t_i = t_i - t_{i-1}$, the same equation becomes:

$$\frac{1}{\sqrt{2\pi\Delta t_i}}\exp\left(-\frac{\Delta x_i^2}{2\Delta t_i}\right)$$

Since each $W_{t_i} - W_{t_{i-1}}$ (for i = 1, 2, ..., i, ..., n) is independent, their joint probability density function $f_{\mathbb{P}}^n(x_1, x_2, ..., x_n)$ is the product of their individual density functions:

$$f_{\mathbb{P}}^{n}(x_1, x_2, ..., x_n) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\Delta t_i}} \exp\left(-\frac{\Delta x_i^2}{2\Delta t_i}\right)$$

8 The Radon-Nikodym Derivative - Continuous Version

Let's assume \mathbb{P} and \mathbb{Q} are equivalent measures. Given a path ω , and for times $(t_1, t_2, ..., t_n)$ (where $t_n = T$), if we define $x_i = W_{t_i}(\omega)$, then the Radon-Nikodym derivative up to time t, $\frac{d\mathbb{Q}}{d\mathbb{P}}$, is obtained

as the continuous limit of the ratio of the joint probability densities of the Brownian motion under each measure:

$$\frac{d\mathbb{Q}}{d\mathbb{P}}(\omega) = \lim_{n \to \infty} \frac{f_{\mathbb{Q}}^{n}(x_1, x_2, ..., x_n)}{f_{\mathbb{P}}^{n}(x_1, x_2, ..., x_n)}$$

This limit is taken by increasing the number of divisions within the interval [0, T] while keeping $t_n = T$ fixed.

Just like in the discrete-time case, the following relationship holds for the continuous Radon-Nikodym derivative:

$$\mathbf{E}_{\mathbb{Q}}(X_t) = \mathbf{E}_{\mathbb{P}}\left(\frac{d\mathbb{Q}}{d\mathbb{P}}X_t\right)$$

With the derivative process defined as $\zeta_t = \mathbf{E}_{\mathbb{P}}\left(\frac{d\mathbb{Q}}{d\mathbb{P}}\Big|\mathcal{F}_t\right)$, the following holds for $s \leq t$:

$$\mathbf{E}_{\mathbb{Q}}\left(X_{t}\middle|\mathcal{F}_{s}\right) = \zeta_{s}^{-1}\mathbf{E}_{\mathbb{P}}\left(\zeta_{t}X_{t}\middle|\mathcal{F}_{s}\right)$$

9 Reviewing the Moment Generating Function

The moment generating function of a random variable X under a measure \mathbb{P} is defined as $\mathbf{E}_{\mathbb{P}}[\exp(\theta X)]$ with a parameter θ .

Specifically, when X follows a normal distribution $N(\mu, \sigma)$, the moment generating function is:

$$\mathbf{E}_{\mathbb{P}}[\exp(\theta X)] = \exp(\theta \mu + \frac{1}{2}\theta^2 \sigma^2)$$

10 Simple Measure Transformation (Brownian Motion + Constant Drift)

In a previous discrete-time section, we saw that if \mathbb{P} and $\frac{d\mathbb{Q}}{d\mathbb{P}}$ are specified, then \mathbb{Q} is also determined. This time, let's see what happens to a \mathbb{P} -Brownian motion W_T under \mathbb{Q} when the Radon-Nikodym derivative is given as:

$$\frac{d\mathbb{Q}}{d\mathbb{P}} = \exp\left(-\gamma W_T - \frac{1}{2}\gamma^2 T\right)$$

To find out, we'll examine the moment generating function under \mathbb{Q} , $\mathbf{E}_{\mathbb{Q}}[\exp(\theta W_T)]$. Let Z be a

standard normal random variable. The calculation proceeds as follows:

$$\mathbf{E}_{\mathbb{Q}}[\exp(\theta W_T)] = \mathbf{E}_{\mathbb{P}} \left[\frac{d\mathbb{Q}}{d\mathbb{P}} \exp(\theta W_T) \right]$$

$$= \mathbf{E}_{\mathbb{P}} \left[\exp\left(-\gamma W_T - \frac{1}{2} \gamma^2 T + \theta W_T \right) \right]$$

$$= \exp\left(-\frac{1}{2} \gamma^2 T \right) \mathbf{E}_{\mathbb{P}} \left[\exp\left\{ (\theta - \gamma) W_T \right\} \right]$$

$$= \exp\left(-\frac{1}{2} \gamma^2 T \right) \mathbf{E} \left[\exp\left\{ (\theta - \gamma) \sqrt{T} Z \right\} \right]$$

$$= \exp\left(-\frac{1}{2} \gamma^2 T \right) \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \exp\left\{ (\theta - \gamma) \sqrt{T} x \right\} \exp\left(-\frac{x^2}{2} \right) dx$$

$$= \exp\left(-\frac{1}{2} \gamma^2 T \right) \frac{1}{\sqrt{2\pi}} \times \sqrt{2\pi} \exp\left(\frac{1}{2} (\theta - \gamma)^2 T \right)$$

$$= \exp\left(-\theta \gamma T + \frac{1}{2} \theta^2 T \right)$$

(The transformation from the 3rd to 5th line utilizes the fact that W_T follows a normal distribution N(0,T) under measure \mathbb{P} and returns to the expectation calculation (integral) weighted by the standard normal probability density.) Comparing this result to the moment generating function for a normal distribution $N(\mu,\sigma)$, which is $\exp(\theta\mu + \frac{1}{2}\theta^2\sigma^2)$, we can see that it's a normal distribution $N(-\gamma T,T)$ with mean $\mu = -\gamma T$ and variance $\sigma = T$.

So, from the perspective of measure \mathbb{Q} , the \mathbb{P} -Brownian motion becomes a Brownian motion with a constant drift of $(-\gamma)$.

If we write the Q-Brownian motion as \tilde{W}_T , we can see that the P-Brownian motion can be expressed as $\tilde{W}_T - \gamma T$.

This shows that when the Radon-Nikodym derivative is given as $\frac{d\mathbb{Q}}{d\mathbb{P}} = \exp\left(-\gamma W_T - \frac{1}{2}\gamma^2 T\right)$, there is a relationship between the \mathbb{P} -Brownian motion W_T and the \mathbb{Q} -Brownian motion \tilde{W}_T such that $\tilde{W}_T = W_T + \gamma T$.

10.1 Verification for Time t(< T)

Let's similarly verify this for a time t(< T), given the Radon-Nikodym derivative: $\frac{d\mathbb{Q}}{d\mathbb{P}} = \exp\left(-\gamma W_T - \frac{1}{2}\gamma^2 T\right)$. The claim is that the \mathbb{P} -Brownian motion W_t will be a Brownian motion with a constant drift of $-\gamma t$ under measure \mathbb{Q} . Let's prove this.

10.1.1 Finding the Radon-Nikodym Derivative Process

First, we need to find the Radon-Nikodym derivative process. *1

$$\zeta_{t} = \mathbf{E}_{\mathbb{P}} \left(\frac{d\mathbb{Q}}{d\mathbb{P}} \middle| \mathcal{F}_{t} \right)
= \mathbf{E}_{\mathbb{P}} \left[\exp \left(-\gamma W_{T} - \frac{1}{2} \gamma^{2} T \right) \middle| \mathcal{F}_{t} \right]
= \exp \left(-\frac{1}{2} \gamma^{2} T \right) \mathbf{E}_{\mathbb{P}} \left[\exp \left(-\gamma W_{T} \right) \middle| \mathcal{F}_{t} \right]
= \exp \left(-\frac{1}{2} \gamma^{2} T \right) \mathbf{E}_{\mathbb{P}} \left[\exp \left\{ -\gamma (W_{T} - W_{t}) \right\} \exp \left(-\gamma W_{t} \right) \middle| \mathcal{F}_{t} \right]$$

If we focus on one of the factors inside the expectation:

$$\exp\left\{-\gamma(W_T - W_t)\right\} = \exp\left(-\gamma\sqrt{T - t}\frac{W_T - W_t}{\sqrt{T - t}}\right)$$

The term

$$\frac{W_T - W_t}{\sqrt{T - t}}$$

is a random variable that follows a standard normal distribution N(0,1) under measure \mathbb{P} . If we let this be \mathbb{Z} , the expression inside the expectation can be broken down into a factor that is \mathcal{F}_t -measurable,

$$\exp\left(-\gamma W_t\right)$$

and a factor that is \mathcal{F}_t -independent,

$$\exp\left(-\gamma\sqrt{T-t}Z\right)$$

resulting in:

$$\mathbf{E}_{\mathbb{P}}\left[\exp\left\{-\gamma(W_{T}-W_{t})\right\}\exp\left(-\gamma W_{t}\right)\Big|\mathcal{F}_{t}\right] = \exp\left(-\gamma W_{t}\right)\mathbf{E}\left[\exp\left(-\gamma \sqrt{T-t}Z\right)\right]$$

$$= \exp\left(-\gamma W_{t}\right)\frac{1}{\sqrt{2\pi}}\int_{-\infty}^{\infty}\exp\left(-\gamma \sqrt{T-t}x\right)e^{-\frac{x^{2}}{2}}dx$$

$$= \exp\left(-\gamma W_{t}\right)\exp\left\{\frac{1}{2}(-\gamma \sqrt{T-t})^{2}\right\}$$

Putting it all together, we get:

$$\zeta_t = \exp\left(-\frac{1}{2}\gamma^2 T\right) \exp\left(-\gamma W_t\right) \exp\left\{\frac{1}{2}(-\gamma\sqrt{T-t})^2\right\}$$
$$= \exp\left(-\gamma W_t - \frac{1}{2}\gamma^2 t\right)$$

(It turns out that the derivative process at time t(< T) is the same as the expression for the Radon-Nikodym derivative, just with the capital letter T (the expiration) replaced with the lowercase t (a time before expiration).)

^{*1} The method for calculating the expectation is based on Chapter 2 and Chapter 5 of "Stochastic Calculus for Finance II: Continuous-Time Models" by S. E. Shreve.

10.1.2 Relationship Between \tilde{W}_t and W_t at Time t(< T)

Similar to the t = T case, we'll calculate the moment generating function under \mathbb{Q} . We'll use the measure transformation formula for conditional expectations to perform the calculation as an expectation under \mathbb{P} .

$$\mathbf{E}_{\mathbb{Q}} \left[\exp(\theta W_t) \right] = \mathbf{E}_{\mathbb{Q}} \left[\exp(\theta W_t) \middle| \mathcal{F}_0 \right]$$

$$= \zeta_0^{-1} \mathbf{E}_{\mathbb{P}} \left[\zeta_t \exp(\theta W_t) \middle| \mathcal{F}_0 \right]$$

$$= 1 \times \mathbf{E}_{\mathbb{P}} \left[\exp\left(-\gamma W_t - \frac{1}{2} \gamma^2 t \right) \exp(\theta W_t) \right]$$

$$= \exp\left(-\frac{1}{2} \gamma^2 t \right) \mathbf{E}_{\mathbb{P}} \left[\exp\left((\theta - \gamma) W_t \right) \right]$$

$$= \exp\left(-\frac{1}{2} \gamma^2 t \right) \mathbf{E}_{\mathbb{P}} \left[\exp\left((\theta - \gamma) \sqrt{t} \frac{W_t}{\sqrt{t}} \right) \right]$$

$$= \exp\left(-\frac{1}{2} \gamma^2 t \right) \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \exp\left\{ (\theta - \gamma) \sqrt{t} x \right\} \exp\left(-\frac{x^2}{2} \right) dx$$

$$= \exp\left(-\frac{1}{2} \gamma^2 t \right) \frac{1}{\sqrt{2\pi}} \times \sqrt{2\pi} \exp\left(\frac{1}{2} (\theta - \gamma)^2 t \right)$$

$$= \exp\left(-\theta \gamma t + \frac{1}{2} \theta^2 t \right)$$

Just as we saw for t = T, W_t follows a normal distribution $N(-\gamma t, t)$ under \mathbb{Q} with a mean of $(-\gamma t)$ and variance of t. Rewriting this shows that the relationship between the \mathbb{Q} -Brownian motion \tilde{W}_t and the \mathbb{P} -Brownian motion is:

$$\tilde{W}_t = W_t + \gamma t$$

11 Example: Continuous Process

Using the Radon-Nikodym derivative process, show that

$$\mathbf{E}_{\mathbb{Q}} \left[\exp \left(\theta (\tilde{W}_{t+s} - \tilde{W}_s) \right) \middle| \mathcal{F}_s \right] = \exp \left(\frac{1}{2} \theta^2 t \right)$$

and verify that the independent increments property of Brownian motion holds under \mathbb{Q} as well.

11.1 Solution

We'll use the relationship $\tilde{W}_t = W_t + \gamma t$. We'll transform the measure and calculate the expectation under \mathbb{P} . Using the measure transformation formula for conditional expectations:

$$\begin{split} \mathbf{E}_{\mathbb{Q}} \left[\exp\left(\theta(\tilde{W}_{t+s} - \tilde{W}_s)\right) \middle| \mathcal{F}_s \right] &= \mathbf{E}_{\mathbb{Q}} \left[\exp\left\{\theta\left(W_{t+s} + \gamma(t+s) - W_s - \gamma s\right)\right\} \middle| \mathcal{F}_s \right] \\ &= \zeta_s^{-1} \mathbf{E}_{\mathbb{P}} \left[\zeta_{t+s} \exp\left\{\theta\left(W_{t+s} - W_s + \gamma t\right)\right\} \middle| \mathcal{F}_s \right] \\ &= \zeta_s^{-1} \mathbf{E}_{\mathbb{P}} \left[\exp\left(-\gamma W_{t+s} - \frac{1}{2}\gamma^2(t+s)\right) \exp\left\{\theta\left(W_{t+s} - W_s + \gamma t\right)\right\} \middle| \mathcal{F}_s \right] \\ &= \zeta_s^{-1} \exp\left(\theta \gamma t - \frac{1}{2}\gamma^2(t+s)\right) \mathbf{E}_{\mathbb{P}} \left[\exp\left((\theta - \gamma)W_{t+s} - \theta W_s\right) \middle| \mathcal{F}_s \right] \\ &= \zeta_s^{-1} \exp\left(\theta \gamma t - \frac{1}{2}\gamma^2(t+s)\right) \mathbf{E}_{\mathbb{P}} \left[\exp\left((\theta - \gamma)(W_{t+s} - W_s) - \gamma W_s\right) \middle| \mathcal{F}_s \right] \\ &= \zeta_s^{-1} \exp\left(\theta \gamma t - \frac{1}{2}\gamma^2(t+s)\right) \exp(-\gamma W_s) \mathbf{E}_{\mathbb{P}} \left[\exp\left((\theta - \gamma)(W_{t+s} - W_s)\right) \middle| \mathcal{F}_s \right] \\ &= \zeta_s^{-1} \exp\left(\theta \gamma t - \frac{1}{2}\gamma^2(t+s)\right) \exp(-\gamma W_s) \exp\left(\frac{1}{2}(\theta - \gamma)^2 t\right) \\ &= \exp\left(\gamma W_s + \frac{1}{2}\gamma^2 s\right) \exp\left(\theta \gamma t - \frac{1}{2}\gamma^2(t+s)\right) \exp(-\gamma W_s) \exp\left(\frac{1}{2}(\theta - \gamma)^2 t\right) \\ &= \exp\left(\frac{1}{2}\theta^2 t\right) \end{split}$$

Let's assume a constant γ and a \mathbb{P} -Brownian motion W_t . Given equivalent measures \mathbb{P} and \mathbb{Q} and a Radon-Nikodym derivative:

$$\frac{d\mathbb{Q}}{d\mathbb{P}} = \exp\left(-\gamma W_T - \frac{1}{2}\gamma^2 T\right)$$

The derivative process becomes:

$$\mathbf{E}_{\mathbb{P}}\left(\frac{d\mathbb{Q}}{d\mathbb{P}}\Big|\mathcal{F}_t\right) = \exp\left(-\gamma W_t - \frac{1}{2}\gamma^2 t\right)$$

Furthermore, if we denote the \mathbb{Q} -Brownian motion as \tilde{W}_t , the relationship with the \mathbb{P} -Brownian motion W_t is:

$$\tilde{W}_t = W_t + \gamma t$$

12 The Cameron-Martin-Girsanov Theorem

We've seen that when a specific relationship (the Radon-Nikodym derivative) is given for equivalent measures, a stochastic process that's a Brownian motion under one measure is transformed into a Brownian motion with drift under the other.

The textbook states that 'for the stochastic processes handled in this textbook, any measure transformation will result in a Brownian motion with drift and nothing else.'

This suggests that the result of a Brownian motion simply transforming in this way wasn't a coincidence due to a special relationship being arbitrarily given. The following sections will expand on this idea.

12.1 The Cameron-Martin-Girsanov Theorem

Let W_t be a \mathbb{P} -Brownian motion, and let γ_t be an \mathcal{F} -adapted process that satisfies the condition:

$$\mathbf{E}_{\mathbb{P}} \exp \left(\frac{1}{2} \int_{0}^{T} \gamma_{t}^{2} dt \right) < \infty$$

Then there exists a measure \mathbb{Q} that satisfies the following conditions:

1. \mathbb{Q} is equivalent to \mathbb{P} .

2.
$$\frac{d\mathbb{Q}}{d\mathbb{P}} = \exp\left(-\int_0^T \gamma_t dW_t - \frac{1}{2} \int_0^T \gamma_t^2 dt\right)$$

3. $\tilde{W}_t = W_t + \int_0^t \gamma_s ds$ becomes a \mathbb{Q} -Brownian motion. Conversely, W_t becomes a Brownian motion with a drift of $(-\gamma_t)$ at time t under \mathbb{Q} .

The example we saw in the previous section was a special case where $\gamma_t = \gamma = \text{const.}$

12.2 The Converse of the Cameron-Martin-Girsanov Theorem

The converse is also true.

If W_t is a \mathbb{P} -Brownian motion and \mathbb{Q} is a measure equivalent to \mathbb{P} , then there exists an \mathcal{F} -adapted process γ_t such that:

$$\tilde{W}_t = W_t + \int_0^t \gamma_s ds$$

becomes a Q-Brownian motion. Furthermore, the Radon-Nikodym derivative will be:

$$\frac{d\mathbb{Q}}{d\mathbb{P}} = \exp\left(-\int_0^T \gamma_t dW_t - \frac{1}{2} \int_0^T \gamma_t^2 dt\right)$$

13 The Cameron-Martin-Girsanov Theorem and Stochastic Differentials

Let's look at a practical application of the Cameron-Martin-Girsanov theorem.

All stochastic processes handled in this textbook are variations of Brownian motion, so the Cameron-Martin-Girsanov theorem is a useful tool for manipulating drift.

Let's consider a stochastic process X whose differential form is given as:

$$dX_t = \sigma_t dW_t + \mu_t dt$$

where W_t is a \mathbb{P} -Brownian motion.

The drift of this stochastic process has the functional form μ_t , but we'd like to change it to another functional form, ν_t .

Let \mathbb{Q} be a probability measure under which the drift of the stochastic process X_t is ν_t .

If we denote the \mathbb{Q} -Brownian motion as W_t , the differential form of the stochastic process X_t can be written as:

$$dX_t = \sigma_t d\tilde{W}_t + \nu_t dt$$

Note that the volatility, σ_t , does not change with the measure transformation.

Setting the two right-hand sides equal, we get:

$$\sigma_t dW_t + \mu_t dt = \sigma_t d\tilde{W}_t + \nu_t dt$$

From this, we can see the relationship:

$$d\tilde{W}_t = dW_t + \frac{\mu_t - \nu_t}{\sigma_t} dt$$

If we write the coefficient of dt as γ_t , then:

$$\gamma_t = \frac{\mu_t - \nu_t}{\sigma_t}$$

If this satisfies the condition of the Cameron-Martin-Girsanov theorem,

$$\mathbf{E}_{\mathbb{P}} \exp \left(\frac{1}{2} \int_{0}^{T} \gamma_{t}^{2} dt \right) < \infty$$

then a measure \mathbb{Q} exists such that

$$\tilde{W}_t = W_t + \int_0^t \frac{\mu_s - \nu_s}{\sigma_s} ds$$

is a Q-Brownian motion.

14 Example - Measure Transformation

14.1 Example 1: A Constant Multiple of a Q-Brownian Motion

Let's consider a stochastic process like:

$$X_t = \sigma W_t + \mu t$$

where W_t is a \mathbb{P} -Brownian motion and σ and μ are constants.

If we set $\gamma = \frac{\mu}{\sigma}$ and use the Cameron-Martin-Girsanov theorem, then

$$\tilde{W}_t = W_t + \frac{\mu}{\sigma}t$$

becomes a \mathbb{Q} -Brownian motion up to time T.

If we express X_t using the Q-Brownian motion \tilde{W}_t , it becomes:

$$X_t = \sigma \tilde{W}_t$$

The drift term disappears, and we can say that X_t is a \mathbb{Q} -martingale (a concept that will be introduced in a later section).

Using different measures results in different expectations. Let's find the expectation of X_t^2 under each measure.

Using the \mathbb{P} -Brownian motion,

$$X_t^2 = \sigma^2 W_t^2 + \mu^2 t^2 + 2\sigma \mu W_t t$$

Using the Q-Brownian motion,

$$X_t^2 \ = \ \sigma^2 \tilde{W}_t^2$$

Therefore:

$$\mathbf{E}_{\mathbb{P}}(X_t^2) = \mu^2 t^2 + \sigma^2 t$$
$$\mathbf{E}_{\mathbb{Q}}(X_t^2) = \sigma^2 t$$

The results are indeed different.

14.2 Example 2: Geometric Brownian Motion

Let's consider a stochastic process X_t whose stochastic differential equation is:

$$dX_t = X_t(\sigma dW_t + \mu dt)$$

where W_t is a \mathbb{P} -Brownian motion and σ and μ are constants.

We want to see if we can transform this stochastic differential equation into:

$$dX_t = X_t(\sigma d\tilde{W}_t + \nu dt)$$

using a different measure, \mathbb{Q} , with its Brownian motion \tilde{W}_t .

Similar to the previous examples, if we consider a \tilde{W}_t that satisfies:

$$\sigma W_t + \mu t = \sigma \tilde{W}_t + \nu t$$

we can see that we need:

$$d\tilde{W}_t = dW_t + \frac{\mu - \nu}{\sigma} dt$$

to hold.

The condition for the Cameron-Martin-Girsanov theorem to be applicable,

$$\mathbf{E}_{\mathbb{P}} \exp\left(\frac{1}{2} \int_{0}^{T} \left(\frac{\mu - \nu}{\sigma}\right)^{2} dt\right) = \frac{1}{2} \left(\frac{\mu - \nu}{\sigma}\right)^{2} T$$

is satisfied, as long as μ , ν , and σ are constants.

This confirms that a measure \mathbb{Q} exists under which \tilde{W}_t is a Brownian motion, just as we hypothesized.

(Therefore, we've shown that we can transform the stochastic differential equation into:

$$dX_t = X_t(\sigma \tilde{W}_t + \nu t)$$

using a Brownian motion \tilde{W}_t under measure \mathbb{Q} .)

References

[1] Martin Baxter, Andrew Rennie. Financial Calculus - An Introduction to Derivative Pricing.