

3.3 多次元確率分布

3.3.1 多項分布

二項分布の多次元化

コインの代りに, $1 \sim k$ の目のサイコロ投げ
 目の出る確率 p_i (i の目が出る確率)

$$\sum_{i=1}^k p_i = 1$$

n 回サイコロを投げる。

それぞれ目の出た回数 X_i $i=1 \sim k$

多次元確率変数 $\vec{X} = (X_1, \dots, X_k)$

$$\vec{X} \sim M_k(n; p_1, \dots, p_k)$$

$$f(\vec{x}) = P(\vec{X} = \vec{x}) = \frac{n!}{\prod_{j=1}^k x_j!} \prod_{j=1}^k p_j^{x_j}$$

$$(x_i = 0 \sim n; \sum x_i = n)$$

$$\vec{\mu} = E[\vec{X}] = n(p_1, \dots, p_k) = (np_1, \dots, np_k)$$

$$\sigma^2 = V[\vec{X}] = n(p_1(1-p_1), \dots)$$

$$\text{Cov}[X_i, X_j] = -np_i p_j \quad (i \neq j)$$

(σ_{ij}^2)

一方が増えれば一方は減る

二項の
自然な拡張 $\vec{\mu}$
 σ^2

多項分布の続き、
周辺 pdf

$\vec{X} \sim M_k(n; p_1, \dots, p_k)$ の時

X_1 の周辺確率分布は? 1 の出る回数 X_1 だけ $B(n; p)$ である

$k=3$ の時、

$$f_1(x_1) = \sum_{x_2, x_3} f(x_1, x_2, x_3)$$

$$= \sum_{x_2=0}^{n-x_1} \frac{n!}{x_1! x_2! (n-x_1-x_2)!} p_1^{x_1} p_2^{x_2} p_3^{n-x_1-x_2} \quad \begin{matrix} x_1+x_2+x_3=n \\ x_3=(n-x_1)-x_2 \end{matrix}$$

$$= \frac{n! p_1^{x_1}}{x_1! (n-x_1)!} \sum_{x_2=0}^{n-x_1} \frac{(n-x_1)!}{x_2! (n-x_1-x_2)!} p_2^{x_2} p_3^{n-x_1-x_2}$$

$$= \frac{n! p_1^{x_1}}{x_1! (n-x_1)!} \cdot (p_2 + p_3)^{n-x_1}$$

$$= \frac{n!}{x_1! (n-x_1)!} p_1^{x_1} (1-p_1)^{n-x_1}$$

$$X_1 \sim B(n; p_1)$$

3.3.2 多次元正規分布

一次元正規分布の多次元化.

$$\vec{X} = (X_1, \dots, X_k) \sim N_k(\vec{\mu}, \Sigma)$$

$$f(\vec{x}) = \frac{1}{\sqrt{(2\pi)^k} |\Sigma|^{1/2}} \exp \left\{ -\frac{1}{2} (\vec{x} - \vec{\mu})^T \Sigma^{-1} (\vec{x} - \vec{\mu}) \right\}$$

$\vec{x} \in \mathbb{R}^k$

3.5 多次元正規分布の性質 (二次元)

$$\vec{x} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}, \quad \vec{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \quad \Sigma = \begin{pmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{pmatrix}$$

$$|\Sigma| = \sigma_1^2 \sigma_2^2 - \sigma_{12}^2$$

$$\Sigma^{-1} = \frac{1}{\sigma_1^2 \sigma_2^2 - \sigma_{12}^2} \begin{pmatrix} \sigma_2^2 & -\sigma_{12} \\ -\sigma_{12} & \sigma_1^2 \end{pmatrix}$$

正規分布の期待値

$N(\mu, \sigma^2)$

教科書 3.4.3

p.53

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx$$

$$= \int_{-\infty}^{\infty} x \cdot \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \quad \begin{cases} y = x - \mu, \\ dx = dy \end{cases}$$

$$= \int_{-\infty}^{\infty} (y + \mu) \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{y^2}{2\sigma^2}} dy$$

$$= \underbrace{\int_{-\infty}^{\infty} y \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{y^2}{2\sigma^2}} dy}_{\text{奇関数} \dots \text{積分は} 0} + \mu \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{y^2}{2\sigma^2}} dy$$

$$= \mu \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2} \frac{y^2}{\sigma^2}} dy}_{N(0, \sigma^2) \text{ の p.d.f.}}$$

$$z = \frac{y}{\sigma}, \quad dy = \sigma dz$$

$$= \mu$$

正規分布 $X \sim N(\mu, \sigma^2)$ の分散

教科書 3.4.3
p54

$$V[X] = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) \cdot dx$$

$$= \int_{-\infty}^{\infty} (x - \mu)^2 \cdot \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2}} dx$$

$$\begin{cases} y = \frac{x - \mu}{\sigma} \\ dx = \sigma \cdot dy \end{cases}$$

$$= \int_{-\infty}^{\infty} \sigma^2 y^2 \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy$$

$$= \sigma^2 \int_{-\infty}^{\infty} y \cdot \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \right) dy$$

$$= \sigma^2 \left[-y \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \right]_{-\infty}^{\infty} + \sigma^2 \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy$$

$N(0,1)$ の pdf

$$= \sigma^2$$

3.5 多次元正規分布の性質

(二次元)

2変数のベクトル

$$\vec{x} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}, \quad \vec{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}$$

変換行列

$$\Sigma = \begin{pmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{pmatrix}$$

行列式

$$|\Sigma| = \sigma_1^2 \sigma_2^2 - \sigma_{12}^2$$

変換の逆行列

$$\Sigma^{-1} = \frac{1}{|\Sigma|} \begin{pmatrix} \sigma_2^2 & -\sigma_{12} \\ -\sigma_{12} & \sigma_1^2 \end{pmatrix}$$

二次元正規分布の p.d.f

$$\phi(\vec{x}; \vec{\mu}, \Sigma) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}}$$

$$\times \exp \left\{ -\frac{1}{2} (\vec{x} - \vec{\mu})^t \Sigma^{-1} (\vec{x} - \vec{\mu}) \right\}$$

$$\phi(\vec{x}; \vec{\mu}, \Sigma) = \frac{1}{\sqrt{(2\pi)^k |\Sigma|}} \exp \left\{ -\frac{1}{2} (\vec{x} - \vec{\mu})^T \overset{\text{stat 3.5-3}}{\Sigma^{-1}} (\vec{x} - \vec{\mu}) \right\}$$

$\vec{x}' = \vec{x} - \vec{\mu} \in \mathbb{R}^k$

$$= \frac{1}{\sqrt{2\pi}\eta^2} \exp \left\{ -\frac{1}{2} \left(x_2' - \frac{\sigma_{12}}{\sigma_1^2} x_1' \right)^2 \right\} \times$$

$$\frac{1}{\sqrt{2\pi}\sigma_1^2} \exp \left\{ -\frac{1}{2} \frac{x_1'^2}{\sigma_1^2} \right\}$$

$$= \phi(x_2'; \frac{\sigma_{12}}{\sigma_1^2} x_1', \eta^2) \cdot \phi(x_1'; 0, \sigma_1^2)$$

(3.1)

$$= \underbrace{\phi(x_2; \mu_2 + \frac{\sigma_{12}}{\sigma_1^2} x_1', \eta^2)}_{?} \cdot \underbrace{\phi(x_1; \mu_1, \sigma_1^2)}_{?}$$

$f_1(x_1)$

$$f_1(x_1) = \int \phi(\vec{x}; \vec{\mu}, \Sigma)$$

$$= \int \underbrace{\phi(x_2; \mu_2 + \overset{\text{stat 3.1}}{\frac{\sigma_{12}}{\sigma_1^2}} x_1', \eta^2)}_{\text{stat 3.1}} \phi(x_1; \mu_1, \sigma_1^2) dx_2$$

$$= \phi(x_1; \mu_1, \sigma_1^2)$$

$$\therefore X_1 \sim N(\mu, \sigma_1^2)$$

3.5.2 平均と分散

$$E[\vec{X}] = \vec{\mu},$$

$$V[\vec{X}] = E[\vec{X}^2] - (E[\vec{X}])^2 = \Sigma$$

確認

k=2の時.

$$f_1(x_1) \sim N(\mu_1, \sigma_1^2) \quad \text{已知},$$

$$f_2(x_2) \sim N(\mu_2, \sigma_2^2)$$

$$\begin{aligned} \text{Cov}[X_1, X_2] &= E[(X_1 - \mu_1)(X_2 - \mu_2)] \\ &= \iiint (x_1 - \mu_1)(x_2 - \mu_2) \phi(\vec{x}; \vec{\mu}, \Sigma) d\vec{x} \\ &= \iint \quad \quad \quad \phi(x_2; \nu(x_1), \eta^2) \phi(x_1; \mu_1, \sigma_1^2) d\vec{x} \\ &= \int (x_1 - \mu_1) \underbrace{\left[\int (x_2 - \mu_2) \phi_2(\cdot) dx_2 \right]}_{\nu_1(x_1) - \mu_2} \phi(x_1; \mu_1, \sigma_1^2) dx_1 \\ &= \int \underline{(x_1 - \mu_1)} (\nu(x_1) - \mu_2) \phi(x_1; \mu_1, \sigma_1^2) dx_1 \\ &\quad (x_1 - \mu_1) \left(\mu_2 + \frac{\sigma_{12}}{\sigma_1^2} (x_1 - \mu_1) \right) \\ &\quad \frac{\sigma_{12}}{\sigma_1^2} (x_1 - \mu_1)^2 + \mu_2 (x_1 - \mu_1) \\ &\quad \downarrow \qquad \qquad \qquad \downarrow \\ &\quad \frac{\sigma_{12}}{\sigma_1^2} \quad \sigma_1^2 \qquad \qquad \sigma_{12} \text{ と } \delta \end{aligned}$$

stat 3.5 - 5

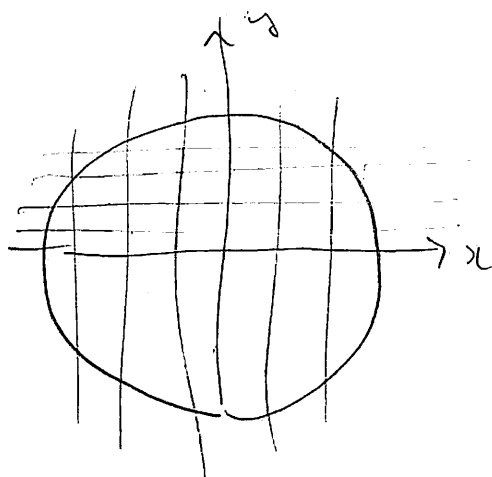
$$= \int \underbrace{(x_1 - \mu_1) \left(\mu_2 + \frac{\sigma_{12}}{\sigma_1^2} (x_1 - \mu_1) \right)} \phi(x_1; \mu_1, \sigma_1^2) dx_1$$

$$= \int \left(\underbrace{\frac{\sigma_{12}}{\sigma_1^2} (x_1 - \mu_1)^2}_{\frac{\sigma_{12}}{\sigma_1^2} \cdot (\sigma_1)^2} + \underbrace{\mu_2 (x_1 - \mu_1)}_{\int_{-\infty}^{\infty} x \phi(x) dx = 0} \right) dx_1$$

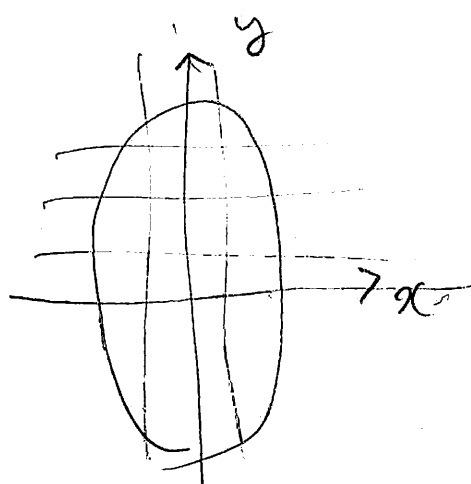
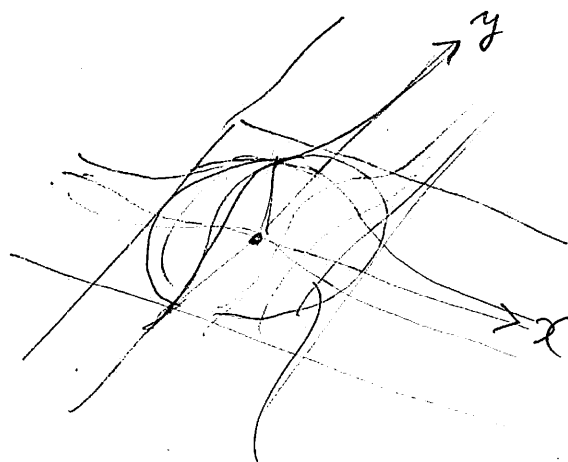
$$= \sigma_{12}$$

3.5.3 $\phi(\vec{x}; \vec{\mu}, \Sigma) \sim N_2(\vec{\mu}, \Sigma)$ のグラフ

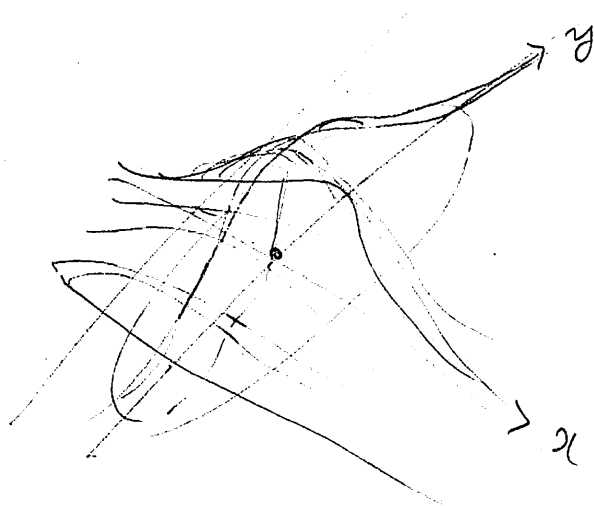
1° x - y -18, $\sigma_1, \sigma_2, \rho = \frac{\sigma_{12}}{\sqrt{\sigma_1 \sigma_2}}$ (相関係数)



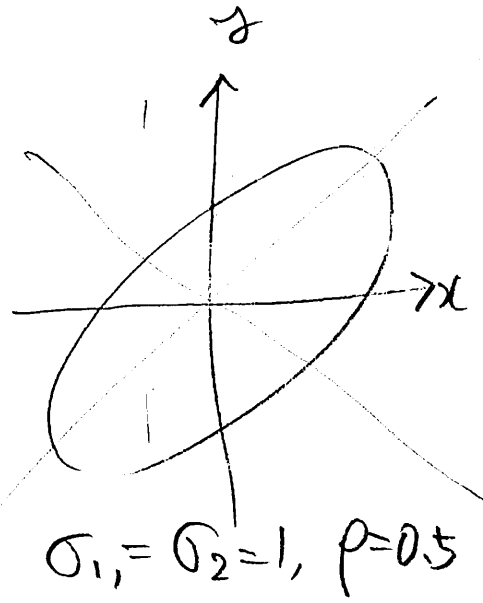
$$\sigma_1 = \sigma_2 = 1, \rho = 0$$



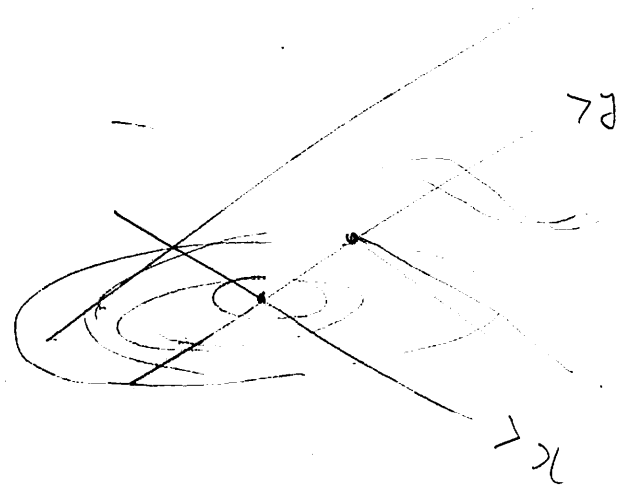
$$\sigma_1 = 1, \sigma_2 = 2, \rho = 0$$



グラフ 統計



stat 5-7



3.5.4 独立性と条件付確率分布

 X_1 と X_2 が独立 \Downarrow 無相関である $\equiv \sigma_{12} = \text{cov}[X_1, X_2] = 0$

二次元正規分布では

 X_1 と X_2 が独立 \Leftrightarrow 無相関である $\sigma_{12} = 0$ を仮定する

$$\begin{aligned}
 \phi(\vec{x}; \vec{\mu}, \Sigma) &= \frac{1}{(2\pi)^{2/2} |\Sigma|^{1/2}} e^{-\frac{1}{2} (\vec{x} - \vec{\mu})^t \Sigma^{-1} (\vec{x} - \vec{\mu})} \\
 &= \frac{1}{(2\pi)^{2/2} |\Sigma|^{1/2}} e^{\left\{ -\frac{1}{2} \frac{(x_1 - \mu_1)^2}{\sigma_1^2} - \frac{1}{2} \frac{(x_2 - \mu_2)^2}{\sigma_2^2} \right\}} \\
 &= \frac{1}{\sqrt{2\pi\sigma_1^2}} e^{\left\{ -\frac{1}{2} \frac{(x_1 - \mu_1)^2}{\sigma_1^2} \right\}} \cdot \frac{1}{\sqrt{2\pi\sigma_2^2}} e^{\left\{ -\frac{1}{2} \frac{(x_2 - \mu_2)^2}{\sigma_2^2} \right\}} \\
 &= \phi_1(x_1; \mu_1, \sigma_1^2) \phi_2(x_2; \mu_2, \sigma_2^2)
 \end{aligned}$$

独立

互いに独立な $X_j \sim N(\mu_j, \sigma_j^2)$ $j=1:k$

$$\Rightarrow \vec{X} \sim N_K(\vec{\mu}, \begin{pmatrix} \sigma_1^2 & & 0 \\ & \ddots & \\ 0 & & \sigma_k^2 \end{pmatrix})$$

35.4 統計学,

条件付密度関数

$$\frac{f_{x_2|x_1}(x_2|x_1)}{\downarrow} = \frac{f(x_1, x_2)}{f_1(x_1)}$$

$f(x_1, x_2)$ と $x_1 = x_1$
の交点曲線
 $\propto f_2(x_2)$

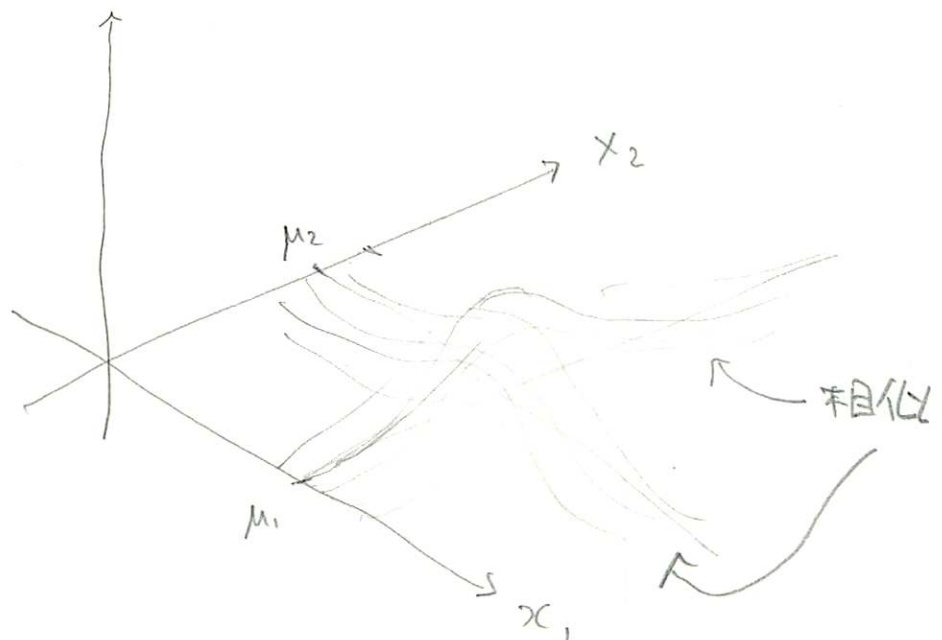
$$= \frac{\phi(\vec{x}; \vec{\mu}, \Sigma)}{\phi(x_1; \mu_1, \sigma_1^2)}$$

$$= \phi(x_2; \nu(x_1), \eta^2)$$

$$\nu(x_1) = \mu_2 + \frac{\sigma_{12}}{\sigma_1^2} (x_1 - \mu_1)$$

$$\eta^2 = \sigma_2^2 - \frac{\sigma_{12}^2}{\sigma_1^2} \left(\frac{\sigma_{11}\sigma_{22} - \sigma_{12}^2}{\sigma_{11}} \right)$$

| 1 | Σ |



モーメント母関数

モーメント母関数

確率分布を特徴づける関数 (表現)

$\circ \int_{-\infty}^{\infty} x^k f(x) dx$ x^k の期待値 ($k=0 \sim \infty$)
 の情報を用いて
 \rightarrow の式に

$\circ 1 + \frac{x^1}{1!} + \frac{x^2}{2!} + \dots = \sum_0^{\infty} \frac{x^n}{n!} = e^x$ となる

$e^{tx} = \sum_0^{\infty} \frac{1}{n!} (tx)^n$

と表す

モーメント母関数 $\psi_x(t)$ ($M_x(t)$)

$\left\{ \begin{aligned} \psi_x(t) &= E[e^{tx}] \\ &= \int_{-\infty}^{\infty} e^{tx} f(x) dx \\ \psi_x(0) &= 1 \end{aligned} \right.$

$X \sim N(\mu, \sigma^2)$ の e^{-tx} の母関数の導出. 3.6 P61

$$\psi_x(t) = E[e^{tx}]$$

$$= \int_{-\infty}^{\infty} e^{tx} \cdot \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{x^2 - 2(\mu+t\sigma^2)x + \mu^2}{2\sigma^2} \right\} dx$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{\{x - (\mu+t\sigma^2)\}^2}{2\sigma^2} \right\} \cdot \underbrace{\exp \left\{ \frac{(\mu+t\sigma^2)^2 - \mu^2}{2\sigma^2} \right\}}_{x \text{ に関与しない}} dx$$

$$= e^{\frac{(\mu+t\sigma^2)^2 - \mu^2}{2\sigma^2}} \cdot \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{\{x - (\mu+t\sigma^2)\}^2}{2\sigma^2}} dx}_{\text{正規分布の pdf の積分}} = 1$$

$N(\mu+t\sigma^2, \sigma^2)$ の pdf

$$= e^{\mu t + \frac{\sigma^2}{2} t^2}$$

$$\begin{aligned} & \frac{(\mu+t\sigma^2)^2 - \mu^2}{2\sigma^2} \\ &= \frac{(\mu^2 + 2\mu t\sigma^2 + t^2\sigma^4) - \mu^2}{2\sigma^2} \\ &= \mu t + \frac{\sigma^2}{2} t^2 \end{aligned}$$

モーメント母関数を使った

正規分布 $(N(\mu, \sigma^2))$ の平均と分散の導出,

$$E[X] = \frac{d}{dt} \psi_X(t) \Big|_{t=0}$$

$$= \left\{ e^{\mu t + \frac{1}{2}\sigma^2 t^2} \right\}' \Big|_{t=0}$$

$$= \left[(\mu + \sigma^2 t) e^{\mu t + \frac{1}{2}\sigma^2 t^2} \right]_{t=0}$$

$$= \mu$$

$$V[X] = E[X^2] - (E[X])^2 = E[X^2] - \mu^2$$

$$E[X^2] = \frac{d^2}{dt^2} \psi_X(t) \Big|_{t=0}$$

$$= \left[(\mu + \sigma^2 t)^2 e^{\mu t + \frac{1}{2}\sigma^2 t^2} + \sigma^2 e^{\mu t + \frac{1}{2}\sigma^2 t^2} \right]_{t=0}$$

$$= \mu^2 + \sigma^2$$

$$\therefore V[X] = E[X^2] - E^2[X]$$

$$= \mu^2 + \sigma^2 - \mu^2$$

$$= \sigma^2$$