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## STAT3023 Statistical Inference

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Lab Week 7

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Consider the test of  $H_0 : \theta = \theta_0$  against  $H_1 : \theta \neq \theta_0$  based on  $X \sim f(., \gamma_0, \theta)$  where

$$f(x; \gamma, \theta) = \begin{cases} \frac{1}{\Gamma(\gamma)\theta^\gamma} x^{\gamma-1} e^{-\frac{x}{\theta}} & \text{for } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

that is  $X$  has a gamma distribution with known shape parameter  $\gamma_0$  but unknown scale parameter  $\theta$ .

1. Consider first the exponential case where  $\gamma_0 = 1$  and suppose the hypothesised value of the scale parameter (also the mean in this case) is  $\theta_0 = 1$

- (a) The “equal-tailed” test at level  $\alpha$  rejects for  $X \leq a$  or  $X \geq b$  where

$$P_1(X \leq a) = P_1(X \geq b) = \frac{\alpha}{2}$$

Taking  $\alpha = 0.05$ , determine the value of **a** and **b** satisfying (2) above (**hint**: use `qexp()`).

### Solution

Since we take  $\gamma_0 = 1$  and under the null  $\theta_0 = 1$  we have that the pdf is given by:

$$f(x) = e^{-x} \quad \text{for } x > 0$$

```
a = qexp(0.025, rate = 1)
b = qexp(0.975, rate = 1)
c(a, b)

## [1] 0.02531781 3.68887945
```

- (b) We shall plot the power function of the equal-tailed test. Define a vector of  $\theta$ -values: `th=(250:1500)/1000` and obtain a corresponding vector of values of the power (the probability of rejecting) for each such  $\theta$ -value; that is:

$$P_\theta(X \leq a) + P_\theta(X \geq b)$$

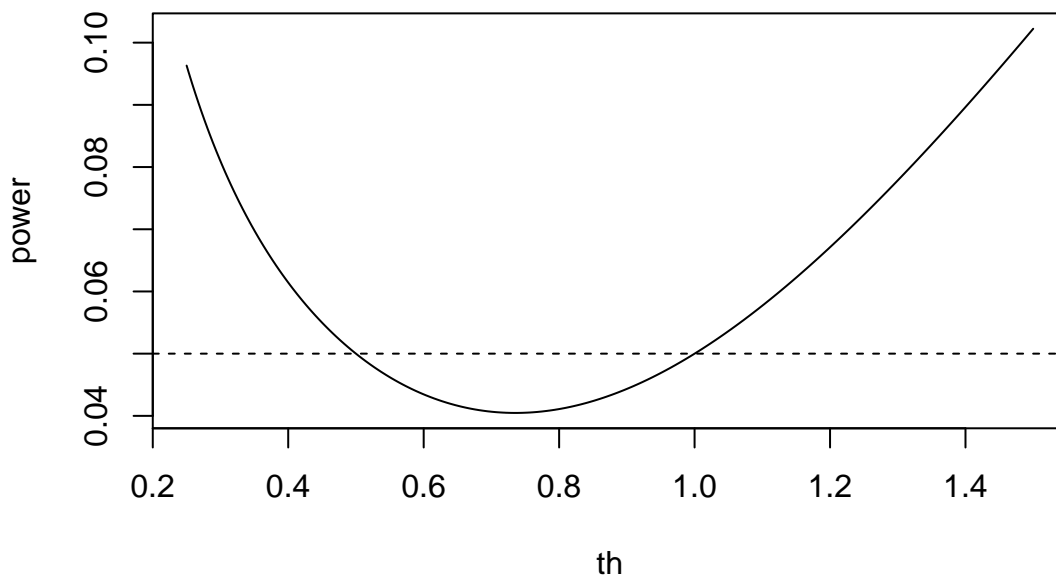
Finally plot the power against `th` and add a horizontal dashed line at  $\gamma = 0.05$ . Add an informative heading, etc and remember to use `type = 'l'`

### Solution

We plot the power function of the equal-tailed test. Note that for  $P_\theta(X \geq b)$  we note that  $P_\theta(X \geq b) = 1 - P_\theta(X \leq b)$ . We also use the `abline(.)` function to plot a horizontal line

```
th = (250:1500)/1000
power = pexp(a, rate = 1/th, lower.tail = TRUE) + pexp(b, rate = 1/th, lower.tail = FALSE)
# Plot theta against the power:
plot(th, power, type = "l", main = "Power of two tail test")
# add a horizontal dashed line at gamma = 0.05
abline(h = 0.05, lty = 2)
```

### Power of two tail test



- (c) This is a 1-parameter exponential family with sufficient statistic  $X$  and so (since it is continuous) the *uniformly most powerful unbiased* (UMPU) test is of the form

$$\delta(X) = \begin{cases} 1 & \text{for } X \geq d \\ 0 & \text{for } c < X < d \\ 1 & \text{for } X \leq c \end{cases}$$

Where  $c$  and  $d$  are chosen so that:

$$\mathbb{E}_{\theta_0}[\delta(X)] = \alpha$$

$$\mathbb{E}_{\theta_0}[X\delta(X)] = \alpha \mathbb{E}_{\theta_0}(X) = \alpha$$

Since  $\mathbb{E}_{\theta_0}(X) = 1$ . We show in a tutorial exercise these are equivalent to

$$\begin{aligned} 1 - e^{-c} + e^{-d} &= \alpha \\ ce^{-c} &= de^{-d} \end{aligned}$$

From the first equation we get:

$$\begin{aligned} e^{-d} &= \alpha - 1 + e^{-c} \\ \implies d &= -\log[\alpha - 1 + e^{-c}] \end{aligned}$$

Hence, once  $c$  is determined we can compute  $d$ . To determine  $c$  we need to solve the equation:

$$ce^{-c} - de^{-d} = ce^{-c} + \{\log[\alpha - 1 + e^{-c}]\}[\alpha - 1 + e^{-c}] = 0$$

We can use the R function `uniroot()` to determine  $c$  *numerically*.

- (i) Write an R function which computes the middle member of the equation above (i.e. the function whose root we wish to find).

**Solution**

We write a function which computes the root  $c$  to the equation

$$ce^{-c} + \{\log[\alpha - 1 + e^{-c}]\} [\alpha - 1 + e^{-c}] = 0$$

```
fn = function(c, alpha) {
  term = alpha - 1 + exp(-1 * c)
  return(c * exp(-c) + (log(term) * term))
}
```

- (ii) Noting that  $c$  can be no bigger than the lower 0.05-quantile of the exponential(1) distribution, execute a certain command involving `eps = 1e-5`. Note: the use of `eps` here is to stay away from the upper bound, since there the function is trying to evaluate `log(0)`

```
eps = 1e-5
uniroot(f = fn, lower = 0, upper = qexp(0.05) - eps, alpha = 0.05)
## $root
## [1] 0.04235629
##
## $f.root
## [1] -3.187136e-05
##
## $iter
## [1] 4
##
## $init.it
## [1] NA
##
## $estim.prec
## [1] 6.103516e-05
```

- (iii) Write an R function which takes as an argument the level `alpha` and returns a list with elements `c` and `d`, corresponding to the desired values  $c$  and  $d$  defining the UMPU test (3) above for  $\theta_0 = 1$  and  $\alpha = 0.05$

**Solution**

We write a function `expon.umpu` (exponential uniformly most powerful unbiased test) which takes in  $\alpha$  and returns the optimal  $c$  and  $d$  which solve the above equations.

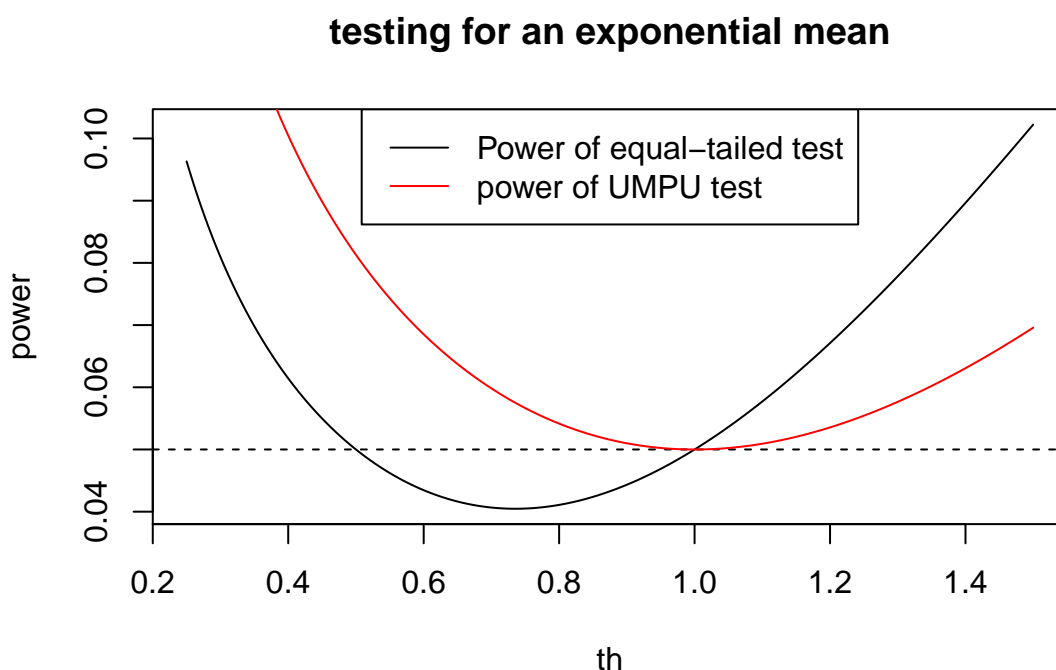
```
expon.umpu = function(alpha) {
  eps = 1e-8
  c = uniroot(f = fn, lower = 0, upper = qexp(alpha) - eps, alpha = alpha)$root
  term = alpha - 1 + exp(-c)
  d = -log(term)
  list(c_val = c, d_val = d)
}
expon.umpu(0.05)
## $c_val
## [1] 0.04235611
##
## $d_val
## [1] 4.764356
```

- (d) Recreate your plot from part (b) above and add to it a red curve of the power of the UMPU test. Add an informative heading and legend. **Comment** on what feature of the plot indicates that the UMPU test is unbiased. The power function of the UMPU test never goes below the 0.05 level, thus it is unbiased.

### Solution

We now plot the power of the UMPU test in red:

```
th = (250:1500)/1000
# for the two tailed test
power = pexp(a, 1/th) + pexp(b, 1/th, lower.tail = FALSE)
plot(th, power, type = "l", main = "testing for an exponential mean")
abline(h = 0.05, lty = 2)
# for the uniformly most powerful test:
umpu = expon.umpu(0.05)
umpu.power = pexp(umpu$c, 1/th) + pexp(umpu$d, 1/th, lower.tail = FALSE)
lines(th, umpu.power, col = "red")
legend("top", legend = c("Power of equal-tailed test", "power of UMPU test"),
      col = c("black", "red"),
      lty = c(1, 1))
```



2. Consider now the case where  $Y_1, \dots, Y_n$  are iid exponential with mean  $\theta$  and we again wish to test  $H_0 : \theta = \theta_0$  against a two sided  $H_1 : \theta \neq \theta_0$ . The likelihood is:

$$\prod_{i=1}^n \left[ \frac{1}{\theta} e^{-\frac{Y_i}{\theta}} \right] = \exp \left( -\frac{1}{\theta} \sum_{i=1}^n Y_i - n \log \theta \right)$$

and so clearly  $X = \sum_{i=1}^n Y_i$  is a sufficient statistic; indeed this is a 1-parameter exponential family. The UMPU test is thus of the form (3) where  $c$  and  $d$  are chosen to satisfy the two conditions (4) and (5).

However since  $X$  itself has a gamma distribution with shape parameter  $n$  and scale parameter  $\theta$ , the UMPU above is the same as for a single observation from the density (1) above with  $\gamma_0 = n$ . We show in the tutorial that in the present case the two conditions (4) and (5) are equivalent to

$$\int_0^c f(x; n, 1) dx + \int_d^\infty f(x; n, 1) dx = \alpha = \int_0^c f(x; n+1, 1) dx + \int_d^\infty f(x; n+1, 1) dx$$

Below we shall write a function to determine the UMPU test, for the case  $n = 5$  and  $\alpha = 0.05$ .

- (a) By adapting your solution to part (c) of the previous question, write a function (playing the same role as the function `fn()` above; it will use `pgamma()` and `qgamma()`) the root of which gives the desired value of  $c$  to solve the above equations (**hint**: in the body of this function you will need to first find  $d$  in terms of  $c$  and  $\alpha$  using one of the two constraints). Then use `uniroot()` to actually find the root. Wrap this all in an appropriate function which takes as input values of `alpha` and `n` and outputs a list with elements `c` and `d`, the lower and upper critical values of the desired UMPU test.

### Solution

```
gamma.root = function(c, n, alpha) {
  lower = pgamma(q = c, shape = n, scale = 1) # P (X < c)
  # P(X < c) + P(X > d) = alpha
  # P(X > d) = alpha - P(x < c)
  # P(X < d) = 1 - alpha + P(X < c)
  # d = F^-1(1 - alpha + P(X < c))
  d = qgamma(p = 1 - alpha + lower, shape = n, scale = 1)

  # return the equation to solve--> the latter half of the integral equation
  return(pgamma(c, shape = (n + 1), scale = 1) +
         pgamma(d, shape = (n + 1), scale = 1, lower.tail = FALSE) -
         alpha)
}
```

Show the roots:

```
eps = 1e-8
upper = qgamma(0.05, shape = 5, scale = 1) - eps
uniroot(f = gamma.root, lower = 0, upper = upper, n = 5, alpha = 0.05)

## $root
## [1] 1.758069
##
## $f.root
## [1] 1.352949e-06
##
## $iter
## [1] 6
##
## $init.it
## [1] NA
##
## $estim.prec
## [1] 6.103516e-05
```

And finally wrap all this up in a function:

```
gamma.umpu = function(alpha, n) {
  eps = 1e-8
  upper = qgamma(alpha, shape = n, scale = 1) - eps
  c = uniroot(f = gamma.root, lower = 0, upper = upper, n = n, alpha = alpha)$root
  lower = pgamma(c, shape = n, scale = 1)
  d = qgamma(1 - (alpha - lower), shape = n, scale = 1)
  return(list(c_val = c, d_val = d))
}
```

- (b) Use your `gamma.umpu()` function to determine the appropriate  $c$  and  $d$  for the UMPU test for this problem with  $n = 5$  and  $\alpha = 0.05$ . Plot the power as a function of  $\theta$  and graphically verify that the test is unbiased and of level 0.05.

```
gamma.umpu(0.05, 5)

## $c_val
## [1] 1.758069
##
## $d_val
## [1] 10.86438

gu = gamma.umpu(0.05, 5) # obtain c and d
g.power = pgamma(gu$c, shape = 5, scale = th) +
  1 - pgamma(gu$d, shape = 5, scale = th)
plot(th, g.power, type = "l")
abline(h = 0.05, lty = 2)
```

