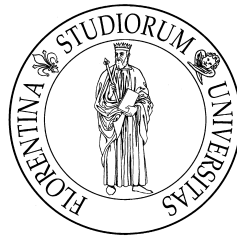UNIVERSITÀ DEGLI STUDI DI FIRENZE
Facoltà di Scienze Matematiche, Fisiche e Naturali
Corso di Laurea Magistrale in Informatica

Elaborato d'Esame

PROGETTAZIONE E ANALISI DI ALGORITMI

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CONTENTS

2   Contents

Text contents

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LECTURES NOTES

1.1 SORTING ALGORITHMS

We study two algorithms that are based on checks between keys, those are MERGESORT and QUICKSORT.

1.1.1 *MERGESORT algorithm*

The MERGESORT algorithm is independent from the keys present in the input vector and its methodology behaves always the same. Let $n = 2^m$ be the length of the vector to be ordered, we can define a function C which count the number of checks needed to order the input vector. We define the function C using the method used by the algorithm at each step:

$$C(2^m) = 2C(2^{m-1}) + 2^m$$

Solving the recurrence[1] we obtain $C(n) \in O(n \log n)$. We can observe that if we use a method (like the ones we're studying in this section) based on checks between keys, isn't possible to do better than build a "checks tree", this allow use to have a lower bound for the complexity of those algorithms.[2]

1.1.2 *QUICKSORT algorithm*

The QUICKSORT algorithm depends on the distribution of the keys in the input vector. For what follow we assume to have a probability space $\Omega = D_n$, where D_n is the set of permutation of length n without repetition over $\{1, \ldots, n\}$. We focus on the simpler variant where the pivot is chosen as the right-most key[3]. We study the behavior of an application to the vector $(20, 25, 7, 3, 30, 8, 41, 18)$, reporting in Table 1 the steps performed. We can observe that in order to move the pivot element in its final position, it is necessary for two keys $(7, 25)$ to be checked twice. Hence, given a vector of length n, the number of checks performed before recurring on left and right partitions is $(n-1) + 2$ (where $n-1$ is explained by the pivot isn't indexed neither with i nor with j).

We can analyze the number of checks performed, partitioning it in the following cases:

---

1 put here the proof
2 in the slides of the first lecture maybe there's more material about this topic
3 report here the code

Table 1: Quicksort example

20	25	7	3	30	8	41	18		
↑ i					↑ j		↑ pivot	→	$\{20, 41, 8\}$
8	25	7	3	30	20	41	18		
	↑ i		↑ j				↑ pivot	→	$\{25, 30, 3\}$
8	3	7	25	30	20	41	18		
		↑ j	↑ i				↑ pivot	→	$\{7, 25, 7\}$
8	3	7	18	30	20	41	25		
			↑ pivot					→	recursion

WORST CASE when the vector is already ordered (in one of the two directions). In this case we got:

$$C(n) = (n-1) + 2 + C(n-1)$$

recurring only on one partition because the other have to be empty. We can expand the recurrence, fixing $C(0) = 0$:

$$C(n) = (n+1) + C(n-1) = (n+1) + n + C(n-2) =$$
$$= (n+1) + n + (n-1) + \ldots + 2 + C(0) =$$
$$= \sum_{k=2}^{n+1} k + C(0) = \sum_{k=1}^{n+1} k - 1 + C(0) = \frac{(n+1)(n+2)}{2} - 1$$

so $C(n) \in O(n^2)$.

BEST CASE when the partition phase puts the pivot in the middle, hence the QUICKSORT recurs on balanced partitions. In this case we have the same complexity of MERGESORT, hence $C(n) \in O(n \log n)$

We explain the average case in a dedicated section.

1.1.3 *QUICKSORT: On the average number of checks*

To study this case we have to consider all elements of Ω (recall that $w \in \Omega \rightarrow (w[i] \in \{1, \ldots, n\}) \land (\forall i \neq j : w[i] \neq w[j])$). First of all we can suppose that j is the pivot, hence a generic w will have this structure:

$$w = (C_{j-1} \quad C_{n-j} \quad j)$$

where C_k is a vector of length k. We can consider the probability to have j as pivot considering the uniform distribution on Ω:

$$\mathbb{P}\left(w \in \Omega : w[n] = j\right) = \frac{(n-1)!}{n!} = \frac{1}{n}$$

Our goal here is to build a function $C(n)$ which counts the average number of checks during an execution of the algorithm given an input vector of length n. In order to do that observe that every keys $j \in \{1, \ldots, n\}$ can be the pivot, we can write:

$$C(n) = (n+1) + \frac{1}{n} \sum_{j=1}^{n} C(j-1) + C(n-j)$$

Observing the sum when j runs:

$$j = 1 \rightarrow C(0) + C(n-1)$$
$$j = 2 \rightarrow C(1) + C(n-2)$$
$$\cdots$$
$$j = n-1 \rightarrow C(n-2) + C(1)$$
$$j = n \rightarrow C(n-1) + C(0)$$

Hence we can rewrite:

$$C(n) = (n+1) + \frac{2}{n} \sum_{j=0}^{n-1} C(j)$$

Now we do some manipulation:

$$C(n) = (n+1) + \frac{2}{n} \sum_{j=0}^{n-1} C(j)$$

$$nC(n) = n(n+1) + 2 \sum_{j=0}^{n-1} C(j)$$

Subtract the previous $(n-1)$ term to both members:

$$nC(n) - (n-1)C(n-1) = n(n+1) + 2\sum_{j=0}^{n-1} C(j)$$

$$- \left((n-1)((n-1)+1) + 2 \sum_{j=0}^{(n-1)-1} C(j) \right)$$

$$= n(n+1) + 2\sum_{j=0}^{n-1} C(j)$$

$$- n(n-1) - 2\sum_{j=0}^{n-2} C(j)$$

$$= n(n+1-(n-1)) + 2C(n-1)$$

$$= 2(n + C(n-1))$$

Getting $nC(n) = 2n + (n+1)C(n-1)$, divide both member by $n(n+1)$:

$$\frac{C(n)}{n+1} = \frac{2}{n+1} + \frac{C(n-1)}{n}$$

We arrived at a recurrence $A(n) = b(n) + A(n-1)$, where $A(n) = \frac{C(n)}{n+1}$ and $b(n) = \frac{2}{n+1}$. So we expand, fixing $C(0) = 0$:

$$\frac{C(n)}{n+1} = \frac{2}{n+1} + \frac{C(n-1)}{n} = \frac{2}{n+1} + \frac{2}{n} + \frac{C(n-2)}{n-1}$$

$$= \frac{2}{n+1} + \frac{2}{n} + \ldots + \frac{2}{3} + \frac{2}{2} + \frac{C(0)}{1}$$

$$= \frac{2}{n+1} + \frac{2}{n} + \ldots + \frac{2}{3} + 1$$

$$= 2\left(\frac{1}{n+1} + \frac{1}{n} + \ldots + \frac{1}{3} \right) + 1$$

$$= 2\left(\frac{1}{n+1} + \frac{1}{n} + \ldots + \frac{1}{3} \right) + 2\frac{1}{2} + 2 + 1 - 2\frac{1}{2} - 2$$

$$= 2\left(\frac{1}{n+1} + \frac{1}{n} + \ldots + \frac{1}{3} + \frac{1}{2} + 1 \right) - 2$$

$$= 2(H_{n+1} - 1)$$

Having recognized the harmonic numbers $H_{n+1} = \left(\frac{1}{n+1} + \frac{1}{n} + \ldots + \frac{1}{3} + \frac{1}{2} + 1 \right)$, the final result is

$$C(n) = 2(n+1)(H_{n+1} - 1)$$

In order to bound $C(n)$ we have to recall that $H_n \sim \ln(n) + \gamma$, hence $C(n) \in O(n\log n)$.

From a practical point of view, to avoid the worst case, when a sorting problem is approached with the QUICKSORT algorithm we can choose to do one of the following actions before starting the sorting process:

- shuffling the input vector and proceed with the algorithm described above;

- choose the `pivot` element at random, move it in the right-most position and proceed with the algorithm described above.

Each of the two tricks require linear time in the dimension of the input vector and allow us to use the result of the average case and working with $O(n \log n)$ number of checks.

1.1.4 QUICKSORT: *On the average number of swaps*

Our goal here is to build a function $S(n)$ which counts the average number of swaps during an execution of the algorithm given an input vector of length n.

The recurrence for the average number of swaps is as follow:

$$S(n) = \frac{n-2}{6} + \frac{1}{n} \sum_{j=1}^{n} S(j-1) + S(n-j)$$

We proceed the proof in two stages, first studying

$$\mathbb{E}\left[K_j\right] = \frac{(j-1)(n-j)}{n-1} \tag{1.1}$$

where K_j is a random variable that depends on j, after we'll proof:

$$\frac{n-2}{6} = \frac{1}{n} \sum_{j=1}^{n} \mathbb{E}\left[K_j\right] \tag{1.2}$$

In what follow, suppose to have a probability space Ω as defined in the previous sections and an uniform distribution above it.

Proof of Equation 1.1. Again, suppose the `pivot` is $j \in \{1, \ldots, n\}$. We define the random variable $K_j : \Omega \to \mathbb{R}$ such that:

$$w[n] = j \to K_j(w) = s$$
$$w[n] \neq j \to K_j(w) = 0$$

where s is the number of swaps performed by the partitioning phase of QUICK-SORT given an input vector w of length n (hence the variable K_j counts the number of swaps). Now we can compute the probability to have k swaps when j is the `pivot`:

$$\mathbb{P}\left(K_j = k\right) = \frac{\binom{n-j}{k}\binom{j-1}{k}(n-j)!(j-1)!}{(n-1)!}$$

We can justify the above formula in the following steps:

- we have k swaps when $|\{w_i : w_i < j\}| = k$ and $|\{w_i : w_i > j\}| = k$

- consider the set $\{w_i : w_i < j\}$: we can build permutations of length $j - 1$, and foreach disposition we can choose k keys in $\binom{j-1}{k}$ ways, hence the term $\binom{j-1}{k}(j-1)!$;

- consider the set $\{w_i : w_i > j\}$: we can build permutations of length $n - j$, and foreach disposition we can choose k keys in $\binom{n-j}{k}$ ways, hence the term $\binom{n-j}{k}(n-j)!$;

- the total number of possible permutation of n keys excluding the pivot (which is fixed in the right-most position) is $(n-1)!$

Now we study the mean of the variable K_j:

$$\mathbb{E}\left[K_j\right] = \sum_{k \geqslant 0} k \mathbb{P}\left(K_j = k\right)$$

Using $\binom{n}{m} = \frac{n!}{m!(n-m)!}$, we can rewrite:

$$\mathbb{P}\left(K_j = k\right) = \binom{n-j}{k}\binom{j-1}{k}\frac{(n-j)!(j-1)!}{(n-1)!} = \binom{n-j}{k}\binom{j-1}{k}\binom{n-1}{j-1}^{-1}$$

We put the previous rewrite of $\mathbb{P}\left(K_j = k\right)$ into the definition of $\mathbb{E}\left[K_j\right]$:

$$\mathbb{E}\left[K_j\right] = \sum_{k \geqslant 0} k \mathbb{P}\left(K_j = k\right) = \sum_{k \geqslant 0} k \frac{\binom{n-j}{k}\binom{j-1}{k}}{\binom{n-1}{j-1}}$$

Using the following rewrite for $\binom{j-1}{k}$:

$$\binom{j-1}{k} = \frac{(j-1)!}{k!(j-1-k)!} = \frac{(j-1)}{k}\frac{(j-2)!}{(k-1)!(j-1-k)!} = \frac{(j-1)}{k}\binom{j-2}{k-1}$$

And $\binom{n}{m} = \binom{n}{n-m}$ implies $\binom{j-2}{k-1} = \binom{j-2}{j-2-(k-1)} = \binom{j-2}{j-k-1}$, then:

$$\mathbb{E}\left[K_j\right] = \sum_{k \geqslant 0} k \mathbb{P}\left(K_j = k\right) = \frac{1}{\binom{n-1}{j-1}}\sum_{k \geqslant 0} k \binom{n-j}{k}\binom{j-1}{k}$$

$$= \frac{1}{\binom{n-1}{j-1}}\sum_{k \geqslant 0} k \binom{n-j}{k}\frac{j-1}{k}\binom{j-2}{j-k-1} = \frac{j-1}{\binom{n-1}{j-1}}\sum_{k \geqslant 0}\binom{n-j}{k}\binom{j-2}{j-k-1}$$

Now we recognize the Vandermonde result:

$$\sum_{k \geqslant 0} \binom{r}{k}\binom{s}{n-k} = \binom{r+s}{n}$$

which can be proved directly because the ipergeometric distribution has exactly the same structure, and being a distribution, it sum up to 1. So we use this result applying it to $\sum_{k \geqslant 0} \binom{n-j}{k}\binom{j-2}{j-k-1} = \binom{n-2}{j-1}$, obtaining:

$$\mathbb{E}\left[K_j\right] = \frac{j-1}{\binom{n-1}{j-1}}\binom{n-2}{j-1} = \frac{(j-1)(n-2)!(j-1)!(n-j)!}{(n-1)!(j-1)!(n-j-1)!} =$$

$$= \frac{(j-1)(n-2)!(j-1)!(n-j)(n-j-1)!}{(n-1)(n-2)!(j-1)!(n-j-1)!} = \frac{(j-1)(n-j)}{n-1}$$

$\square$

Now we proceed with the second proof:

Proof of Equation 1.2. Let us start with:

$$\frac{1}{n}\sum_{j=1}^{n}\mathbb{E}\left[K_j\right] = \frac{1}{n}\sum_{j=1}^{n}\frac{(j-1)(n-j)}{n-1} = \frac{1}{n(n-1)}\sum_{j=1}^{n}(j-1)(n-j) =$$

$$= \frac{1}{n(n-1)}\sum_{j=1}^{n}(j(1+n)-j^2-n)$$

$$= \frac{1}{n(n-1)}\left((n+1)\sum_{j=1}^{n}j - \sum_{j=1}^{n}j^2 - n\sum_{j=1}^{n}1\right)$$

$$= \frac{1}{n(n-1)}\left(\frac{n(n+1)^2}{2} - \frac{n(n+1)(2n+1)}{6} - n^2\right) = \ldots = \frac{n-2}{6}$$

$\square$

We use `Maxima` to justify the above calculations:

```
(%i3)   expr:(1/(n*(n-1)))* ((n*(n+1)^2)/2 - n*(n+1)*(2*n+1)/6 -n^2);
        ratsimp(expr);
```

$$(\%o3) \quad \frac{\frac{n\,(n+1)^2}{2} - n^2 - \frac{n\,(n+1)\,(2\,n+1)}{6}}{(n-1)\,n}$$

$$(\%o4) \quad \frac{n-2}{6}$$

We are now ready to solve the main recurrence using the same strategy for the average number of checks, fixing $S(0) = S(1) = S(2) = 0$:

$$nS(n) - (n-1)S(n-1) = \frac{2n-3}{6} + 2S(n-1)$$

$$\frac{S(n)}{n+1} = \frac{S(n-1)}{n} + \frac{2n-3}{6n(n+1)} = \sum_{k=3}^{n}\frac{2k-3}{6k(k+1)}$$

Study the general term of the summation $\frac{2k-3}{6k(k+1)}$ and decompose it in partial fractions using Maxima:

```
(%i131) generalTerm:(2*k-3)/(6*k*(k-1));
        pf:partfrac(generalTerm,k);
        integrate(pf , k);
```

$$(\%o131) \quad \frac{2\,k-3}{6\,(k-1)\,k}$$

$$(\%o132) \quad \frac{1}{2\,k} - \frac{1}{6\,(k-1)}$$

$$(\%o133) \quad \frac{\log(k)}{2} - \frac{\log(k-1)}{6}$$

Hence $S(n) \in O(n\log n)$.

1.1.5 *Final consideration*

We've finished our treatment of sorting algorithms based on checks between keys. We've seen theoretical results under a probability space Ω composed by permutations of n objects. This allow us to remark that if we setup a simulation where we apply one algorithm seen above to the entire space Ω with vectors of a given dimension n, then the average number of checks and swaps must be equal to the theoretical results obtained in the previous formulas.

1.2   ELEMENTS OF COMBINATORICS CALCULUS

1.2.1 *Permutations* P_n

Given a vector $(a_1, \ldots, a_n)$ the number of permutations (or the possible ways to order the elements a_i) is $n!$. A simple proof of that is as follow: for the position 1 are available n elements, for the second position (after fixing the first) are available $n-1$ elements arriving to the position n where only one element is left, hence $n!$.

It is interesting to show how to add a new element a_{n+1} to existing permutations from the set $\{a_1, \ldots, a_n\}$, this is indeed another proof (by induction) to the number $n!$ justified above. Having the permutations of length n by hypothesis, we study how is it possible to add a_{n+1} into a given permutation $w = (a_1, a_2, \ldots, a_n)$:

$$(a_{n+1}, a_1, a_2, \ldots, a_n)$$
$$(a_1, a_{n+1}, a_2, \ldots, a_n)$$
$$\ldots$$
$$(a_1, a_2, \ldots, a_{n+1}, a_n)$$
$$(a_1, a_2, \ldots, a_n, a_{n+1})$$

Those new vectors are $n+1$ because in the original permutation $w = (a_1, a_2, \ldots, a_n)$ there are $n-1$ "internal holes" between the consecutive elements a_i and a_{i+1}, plus the left-most and right-most "external holes" $(\square, a_1, \ldots)$ and $(\ldots, a_n, \square)$. So far we've considered only one permutation w of n elements, repeating the same reasoning for the others w_i (which are $n! - 1$), it must exists $n+1$ new vectors for each of them. So the number of permutation of $n+1$ elements is $(n+1)n! = (n+1)!$ (what we would have expected).

1.2.2 *Dispositions* $D_{n,k}$

A special case of permutations are dispositions, that is the number of different ways in which k elements can be chosen from a set of length n, with $n \geqslant k$. We have n elements available for the first choice, $n-1$ for the second choice arriving to $n-k+1$ elements for the k-th choice, hence $D_{n,k} = n(n-1)\cdots(n-k+1)$. We can write an example to show how to build dispositions. Let $\Omega = \{a, b, c, d\}$ and we want to build $D_{4,2}$, that is the set of ordered pairs, paying attention that $\forall i : (i, i) \notin D_{n,2}$. The following pairs are the candidates (the pairs of the form (i, i) aren't listed because after we choose i for the first position it isn't possible to choose it again for the second):

$$
\begin{array}{ccc}
(a, b) & (a, c) & (a, d) \\
(b, a) & (b, c) & (b, d) \\
(c, a) & (c, b) & (c, d) \\
(d, a) & (d, b) & (d, c)
\end{array}
$$

We've build the previous matrix because we take care about the order in which the pairs are built. In the next section we'll see that there exists another set that doesn't take care of the order.

1.2.3 *Combinations* $C_{n,k}$

A special case of dispositions are combinations. With combinations, notation $C_{n,k}$, we mean a set of sets $\{\{a_{i_1}, \ldots, a_{i_k}\}\}$ without regard to the order in which $a_{i_1}, \ldots, a_{i_k}$ appear. We can see combinations like dispositions modulo the symmetric relation: $(j, i) \in C_{n,2} \rightarrow (i, j) \notin C_{n,2}$ and $(i, i) \notin C_{n,2}, \forall i$. The following sub matrix is to combinations as the previous one is to dispositions:

$$
\begin{array}{ccc}
(a, b) & (a, c) & (a, d) \\
 & (b, c) & (b, d) \\
 & & (c, d)
\end{array}
$$

The combinations matrix is obtained by the disposition matrix, breaking the symmetry, only half of them belongs to the dispositions $C_{4,2}$. Hence $C_{n,k} = \binom{n}{k} = \frac{D_{n,k}}{k!} = \frac{n(n-1)\cdots(n-k+1)}{k!} = \frac{n!}{k!(n-k)!)}$.

1.2.4   Binomial coefficients

In this section we study the binomial coefficients, and in the particular the writing $(a+b)^n, n \in \mathbb{N}$. It is known that:

$$(a+b)^n = \sum_{k=0}^{n} \binom{n}{k} a^k b^{n-k}$$

Now we study $(a+b)^r = \sum_{k=0}^{r} \binom{r}{k} a^k b^{r-k}, r \in \mathbb{R}$, in particular we have to pay attention to the binomial coefficient when $r \in \mathbb{R}$:

$$\binom{r}{k} = \frac{r(r-1)\cdots(r-k+1)}{k!}$$

The previous product make sense, instead we would have had problem with $\frac{r!}{k!(r-k)!}$. Now do simple manipulation:

$$(a+b)^r = \left(a\left(1+\frac{b}{a}\right)\right)^r = a^r \left(1+\frac{b}{a}\right)^r$$

In general we can call $z = \frac{b}{a}$ and reduce to study the function $f(z) = (1+z)^r$, which can be expanded with Taylor:

$$f(z) = \sum_{k \geqslant 0} \frac{f^{(k)}(0)}{k!} z^k$$

Study the derivatives of $f(z)$:

$f^{(0)}(0) = 1$

$f^{(1)}(z) = r(1+z)^{r-1} \quad \rightarrow \quad f^{(1)}(0) = r$

$f^{(2)}(z) = r(r-1)(1+z)^{r-2} \quad \rightarrow \quad f^{(2)}(0) = r(r-1)$

$f^{(k)}(z) = r(r-1)\cdots(r-k+1)(1+z)^{r-k} \quad \rightarrow \quad f^{(k)}(0) = r(r-1)\cdots(r-k+1)$

Now we can substitute the previous into $f(z)$:

$$f(z) = \sum_{k \geqslant 0} \frac{f^{(k)}(0)}{k!} z^k = \sum_{k \geqslant 0} \frac{r(r-1)\cdots(r-k+1)}{k!} z^k = \sum_{k \geqslant 0} \binom{r}{k} z^k$$

Exercise 1.2.1. *Compute* $\binom{\frac{1}{2}}{3}$.

Sol. It is sufficient to build a function $f(z) = (1 + z)^r$ with the right value for r and then expand $f(z)$ with Taylor, the coefficient of z^3 is the desired value $\binom{r}{3}$.

In this case $r = \frac{1}{2}$, so $f(z) = \sqrt{1 + z} = \binom{\frac{1}{2}}{0}z^0 + \binom{\frac{1}{2}}{1}z^1 + \binom{\frac{1}{2}}{2}z^2 + \binom{\frac{1}{2}}{3}z^3 + \dots$

Simple check using `Maxima`:

```
(%i27)  f(z):=sqrt(1+z);
        taylor(f(z), z, 0, 6);
        mapBinomial(x):=binomial (1/2, x)$
        map (mapBinomial, [0, 1, 2, 3, 4, 5, 6]);
```

$(\%o27) \quad f(z) := \sqrt{1 + z}$

$(\%o28) \quad 1 + \dfrac{z}{2} - \dfrac{z^2}{8} + \dfrac{z^3}{16} - \dfrac{5\,z^4}{128} + \dfrac{7\,z^5}{256} - \dfrac{21\,z^6}{1024} + \dots$

$(\%o30) \quad [1, \dfrac{1}{2}, -\dfrac{1}{8}, \dfrac{1}{16}, -\dfrac{5}{128}, \dfrac{7}{256}, -\dfrac{21}{1024}]$

Where output $(\%o28)$ is obtained using the Taylor expansion, while $(\%o30)$ is obtained computing $\binom{\frac{1}{2}}{3}$ using combinatorial `Maxima` function. $\qquad \square$

Now suppose to have $\binom{-n}{k}, n \in \mathbb{N}$:

$$\binom{-n}{k} = \frac{-n(-n-1)\cdots(-n-k+1)}{k!}$$

At numerator there are k terms, so factor -1:

$$\binom{-n}{k} = \frac{(-1)^k(n+k-1)\cdots(n+1)n}{k!} = \frac{(-1)^k(n+k-1)\cdots(n+1)n(n-1)!}{k!(n-1)!}$$

$$= \frac{(-1)^k(n+k-1)!}{k!(n-1)!} = (-1)^k\binom{n+k-1}{k}$$

1.2.5 *Central Binomial Coefficients and Catalan Numbers*

We introduce in the following paragraph a very important class of numbers, the Catalan numbers. We proceed step by step, making an observation on a rewriting for $\binom{n}{k}$ where $n, k \in \mathbb{N}$.

Given a set of n elements $a_1, \dots, a_n$ it is possible to generate the set S of the combinations of length k. Now choose an element, let say a_j, and partition S in two set:

$$S_{a_j}^+ = \left\{ (a_{\pi(1)}, \dots, a_{\pi(k)}) | \exists i : a_j = a_{\pi(i)} \right\}$$

$$S_{a_j}^- = \left\{ (a_{\pi(1)}, \dots, a_{\pi(k)}) | \forall i : a_j \neq a_{\pi(i)} \right\}$$

where $\pi : \{1, \dots, k\} \to \{1, \dots, n\}$ is one-to-one. The following holds:

$$\binom{n}{k} = \left| S_{a_j}^+ \right| + \left| S_{a_j}^- \right| = \binom{n-1}{k-1} + \binom{n-1}{k}$$

Proof. We've chosen a_j to partition S, hence:

- $\left|S_{a_j}^+\right| = \binom{n-1}{k-1}$ because $S_{a_j}^+$ contains sets which everyone contains a_j. But we can build $S_{a_j}^+$ from scratch choosing $k-1$ elements (a_j is contained in every set of $S_{a_j}^+$, hence we have to build combinations of length $k-1$) from $\{a_1, \ldots, a_n\} \setminus \{a_j\}$ (a_j is an implicit choice, so we choose from $n-1$ elements)

- $\left|S_{a_j}^-\right| = \binom{n-1}{k}$ because $S_{a_j}^-$ contains sets which everyone doesn't contains a_j. But we can build $S_{a_j}^-$ from scratch choosing k elements (a_j isn't contained in any set of $S_{a_j}^-$, hence we have to build combinations of length k) from $\{a_1, \ldots, a_n\} \setminus \{a_j\}$ (a_j is an impossible choice, so we choose from $n-1$ elements)

$\square$

The recurrence $\binom{n}{k} = \binom{n-1}{k-1} + \binom{n-1}{k}$ allow us to build an infinite matrix reported in Table 2 with interesting numbers written inside a ticket as $\boxed{j}$. The

Table 2: Central Binomial Coefficients

	0	1	2	3	4	5	6	...	k
0	$\boxed{1}$								
1	1	1							
2	1	$\boxed{2}$	1						
3	1	3	3	1					
4	1	4	$\boxed{6}$	4	1				
5	1	5	10	10	5	1			
6	1	6	15	$\boxed{20}$	15	6	1		
$\vdots$								$\ddots$	
$\vdots$								$\binom{n-1}{k-1}$	$\binom{n-1}{k}$
n									$\binom{n}{k}$

numbers $\binom{2h}{h}$ are called *Central Binomial Coefficients*, while $\frac{1}{h+1}\binom{2h}{h}$ are called *Catalan Numbers*.

We finish this section on combinatorics with some exercises.

Exercise 1.2.2. *Compute* $\binom{-\frac{1}{2}}{k}$.

Sol. Using the approach followed in Exercise 1.2.1, we can build the function $f(z) = \frac{1}{\sqrt{1+z}}$ and expand it with Taylor (see result (%o118)). Here we use another approach that follow the definition:

$$
\binom{-\frac{1}{2}}{k} = \frac{-\frac{1}{2}(-\frac{1}{2} - 1)(-\frac{1}{2} - 2)\cdots(-\frac{1}{2} - k + 1)}{k!} =
$$

$$
= \frac{-\frac{1}{2}(-\frac{3}{2})(-\frac{5}{2})\cdots(-\frac{2k-1}{2})}{k!} = \frac{(-1)^k 135\cdots(2k-1)}{2^k k!} =
$$

$$
= \frac{(-1)^k 135\cdots(2k-1)}{2^k k!} \frac{246\cdots 2k}{246\cdots 2k} = \frac{(-1)^k(2k)!}{2^k k!} \frac{1}{2^k(123\cdots k)} =
$$

$$
= \frac{(-1)^k(2k)!}{2^k k!} \frac{1}{2^k k!} = \frac{(-1)^k}{4^k}\binom{2k}{k}
$$

Doing a little check with Maxima, for the proposed method we get result (%o122) , equals to Taylor expansion method's result:

```
(%i117) f(z):=1/sqrt(1+z);
        taylor(f(z), z, 0, 6);
        mapBinomial(x):=binomial (-1/2, x);
        map (mapBinomial, [0, 1, 2, 3, 4, 5, 6]);
        byDef(k):=(((-1)^k)/(4^k))*binomial(2*k,k);
        map (byDef, [0, 1, 2, 3, 4, 5, 6]);
```

(%o117) $f(z) := \dfrac{1}{\sqrt{1+z}}$

(%o118) $1 - \dfrac{z}{2} + \dfrac{3z^2}{8} - \dfrac{5z^3}{16} + \dfrac{35z^4}{128} - \dfrac{63z^5}{256} + \dfrac{231z^6}{1024} + \ldots$

(%o119) $\mathrm{mapBinomial}(x) := \dbinom{-\frac{1}{2}}{x}$

(%o120) $[1, -\dfrac{1}{2}, \dfrac{3}{8}, -\dfrac{5}{16}, \dfrac{35}{128}, -\dfrac{63}{256}, \dfrac{231}{1024}]$

(%o121) $\mathrm{byDef}(k) := \dfrac{(-1)^k}{4^k}\dbinom{2k}{k}$

(%o122) $[1, -\dfrac{1}{2}, \dfrac{3}{8}, -\dfrac{5}{16}, \dfrac{35}{128}, -\dfrac{63}{256}, \dfrac{231}{1024}]$

□

1.3 GENERATING FUNCTIONS BASICS

In this section we approach the basics for the generating functions topic.

1.3.1 *Approaching*

The main goal of this section is to manipulate a sequence $\{a_n\}_{n\in\mathbb{N}}$ with a function $a(t)$ such that:

$$a(t) = \sum_{n\geqslant 0} a_n t^n$$

Using $a(t)$ allow us to let t be unbounded and we wont assign any value to it (it is just a "placeholder").

We introduce an operator $\mathcal{G}$ which consume a sequence $\{a_n\}_{n\in\mathbb{N}}$ and produce a function $a(t)$ as defined above. In the following examples we experiment a little with Maxima, which implement $\mathcal{G}$ with the function `ggf(s)`, for some sequence s.

Let $\{a_n = n\}$:

```
(%i109) sequence:makelist(n,n,0, 6);
        taylor(ggf(sequence), x, 0, 8);
```

(%o109) $[0,1,2,3,4,5,6]$
(%o110) $x + 2\,x^2 + 3\,x^3 + 4\,x^4 + 5\,x^5 + 6\,x^6 + ...$
So $\mathcal{G}(a_n) = a(x) = x + 2\,x^2 + 3\,x^3 + 4\,x^4 + 5\,x^5 + 6\,x^6 + ...$

Let $\{a_n = 1$ if n is even $, 0$ otherwise$\}$:

```
(%i105) onlyEvenTerms(n):=if is(mod (n,2)=0) then 1 else 0$
        sequence:makelist(onlyEvenTerms(n),n,0, 6);
        ggfOnlyEvenTerms:ggf(sequence);
        taylor(ggfOnlyEvenTerms, x, 0, 12);
```

(%o106) $[1,0,1,0,1,0,1]$
(%o107) $-\dfrac{1}{x^2 - 1}$
(%o108) $1 + x^2 + x^4 + x^6 + x^8 + x^{10} + x^{12} + ...$
So $\mathcal{G}(a_n) = a(x) = 1 + x^2 + x^4 + x^6 + x^8 + x^{10} + x^{12} + ...$

Let $\{a_n = 1$ if $n \leqslant 2, 0$ otherwise$\}$:

```
(%i97)  onlyFirstThreeTerms(n):=if is(n <= 2) then 1 else 0$
        sequence:makelist(onlyFirstThreeTerms(n),n,0, 6);
        ggfOnlyFirstThreeTerms:ggf(sequence);
        taylor(ggfOnlyFirstThreeTerms, x, 0, 100);
```

(%o98) $[1,1,1,0,0,0,0]$
(%o99) $x^2 + x + 1$
(%o100) $1 + x + x^2 + ...$

So $\mathcal{G}(a_n) = a(x) = 1 + x + x^2$.

In the following subsections we'll explore some properties of $\mathcal{G}$.

1.3.2   *Linearity of $\mathcal{G}$*

Let $\{a_n\}_{n \in \mathbb{N}}$ and $\{b_n\}_{n \in \mathbb{N}}$ be two sequences, $a(t)$ and $b(t)$ their formal power series respectively. Lets sum both of them:

$$
\begin{aligned}
\mathcal{G}(a_n) + \mathcal{G}(b_n) &= a_0 + a_1 t + a_2 t^2 + \ldots + a_n t^n + \ldots \\
&\quad + b_0 + b_1 t + b_2 t^2 + \ldots + b_n t^n + \ldots = \\
&= (a_0 + b_0) + (a_1 + b_1)t + (a_2 + b_2)t^2 + \ldots + (a_n + b_n)t^n + \ldots = \\
&= \mathcal{G}(a_n + b_n)
\end{aligned}
$$

It is possible to generalize, obtaining:

$$
\alpha \mathcal{G}(a_n) = \mathcal{G}(\alpha a_n)
$$
$$
\alpha \mathcal{G}(a_n) + \beta \mathcal{G}(b_n) = \mathcal{G}(\alpha a_n + \beta b_n)
$$

So $\mathcal{G}$ is a *linear operator*.

1.3.3   *Convolution of $\mathcal{G}$*

Let $\{a_n\}_{n \in \mathbb{N}}$ and $\{b_n\}_{n \in \mathbb{N}}$ be two sequences, $a(t)$ and $b(t)$ their formal power series respectively. We study the product $a(t)b(t)$ with Maxima:

```
(%i190) makeSequence(t,j):=sum (j[i]*t^i, i, 0, inf);
        as:taylor(makeSequence(t,a), t, 0, 3);
        bs:taylor(makeSequence(t,b), t, 0, 3);
        as*bs;
        sum ((sum(a[k]*b[n-k],k,0,n))*t^n, n, 0, inf);
        taylor(sum((sum(a[k]*b[n-k],k,0,n))*t^n,n,0,inf), t, 0, 3);
```

$$
(\%o190) \quad \text{makeSequence}(t, j) := \sum_{i=0}^{\infty} j_i \, t^i
$$

$$
(\%o191) \quad a_0 + a_1 \, t + a_2 \, t^2 + a_3 \, t^3 + \ldots
$$

$$
(\%o192) \quad b_0 + b_1 \, t + b_2 \, t^2 + b_3 \, t^3 + \ldots
$$

$$
(\%o193) \quad a_0 \, b_0 + (a_1 \, b_0 + a_0 \, b_1) \, t + (a_2 \, b_0 + a_1 \, b_1 + a_0 \, b_2) \, t^2 + (a_3 \, b_0 + a_2 \, b_1 + a_1 \, b_2 + a_0 \, b_3) \, t^3 + \ldots
$$

$$
(\%o194) \quad \sum_{n=0}^{\infty} \left(\sum_{k=0}^{n} a_k \, b_{n-k} \right) t^n
$$

$$
(\%o195) \quad \left(\sum_{k=0}^{0} b_{-k} \, a_k \right) + \left(\sum_{k=0}^{1} b_{1-k} \, a_k \right) t + \left(\sum_{k=0}^{2} b_{2-k} \, a_k \right) t^2 +
$$

$$\left(\sum_{k=0}^{3} b_{3-k}\, a_k \right) t^3 + \dots$$

The output (%o193) is the product $a(t)b(t)$, we can write it in a more compact form as done in (%o194), recognizing the sum in brackets as the general term of a new sequence, so $\mathcal{G}(a_n)\mathcal{G}(b_n) = \mathcal{G}\left(\sum_{k=0}^{n} a_k\, b_{n-k}\right)$.

1.3.4 *Translation of $\mathcal{G}$*

Let $\{a_n\}_{n\in\mathbb{N}}$ and $\{a_{n+1}\}_{n\in\mathbb{N}}$ be two sequences, $a(t)$ and $b(t)$ their formal power series, reported in (%o79) and (%o80) respectively. We study how we can obtain $\mathcal{G}(a_{n+1})$ from $\mathcal{G}(a_n)$ with Maxima:

```
(%i78)  makeSequence(t,j,offset):=if is(offset >= 0) then sum
        (j[i+offset]*t^i, i, 0, inf) else sum (j[i]*t^(i-offset), i, 0, inf)$
        as:taylor(makeSequence(t,a,0), t, 0, 4);
        bs:taylor(makeSequence(t,a,1), t, 0, 3);
        (as-a[0])/t;
```

(%o79) $a_0 + a_1\, t + a_2\, t^2 + a_3\, t^3 + a_4\, t^4 + \dots$
(%o80) $a_1 + a_2\, t + a_3\, t^2 + a_4\, t^3 + \dots$
(%o81) $a_1 + a_2\, t + a_3\, t^2 + a_4\, t^3 + \dots$

To obtain $\mathcal{G}(a_{n+1})$ we subtract a_0 from $\mathcal{G}(a_n)$, dividing the result by t, as reported in (%o80) , so:

$$\mathcal{G}(a_{n+1}) = \frac{\mathcal{G}(a_n) - a_0}{t}$$

We can generalize, allowing arbitraty translation $s \geqslant 0$:

```
(%i249) load("simplify_sum");
        sequence:makeSequence(t,a,0);
        translated:makeSequence(t,a,s)$
        (sequence - sum(a[i]*t^i,i,0,s-1))/t^s;
        (taylor(sequence, t, 0, 55) - sum(a[i]*t^i,i,0,49))/t^50;
```

(%o250) $\displaystyle\sum_{i=0}^{\infty} a_i\, t^i$

(%o252) $\dfrac{\left(\sum_{i=0}^{\infty} a_i\, t^i\right) - \sum_{i=0}^{s-1} a_i\, t^i}{t^s}$

(%o253) $a_{50} + a_{51}\, t + a_{52}\, t^2 + a_{53}\, t^3 + a_{54}\, t^4 + a_{55}\, t^5 + \dots$

where in the last line we've fixed $s = 50$. In general for an arbitrary translation $s \geqslant 0$ holds:

$$\mathcal{G}(a_{n+s}) = \frac{\mathcal{G}(a_n) - \sum_{i=0}^{s-1} a_i\, t^i}{t^s}$$

If we have a negative translation $s < 0$, then we're in the following context (in the following we use a_{n-s}, fixing the sing of $s > 0$):

$$\mathcal{G}(a_n) = a_0 + a_1 t + a_2 t^2 + \ldots + a_n t^n + \ldots$$
$$\mathcal{G}(a_{n-s}) = a_{-s} + a_{-s+1} t + \ldots + a_0 t^s + a_1 t^{s+1} + \ldots + a_n t^{n+s} + \ldots$$

With $a_k = 0$ for $k < 0$ since the sequence is defined only for $k \in \mathbb{N}$. So we've:

$$\mathcal{G}(a_{n-s}) = a_0 t^s + a_1 t^{s+1} + \ldots + a_n t^{n+s} + \ldots = t^s \mathcal{G}(a_n)$$

With Maxima:

```
(%i129) p:5$
        cs:taylor(makeSequence(t,a,-p), t, 0, 10);
        ds:taylor((t^p)*makeSequence(t,a,0), t, 0, 10);
```

$$(\%o130)\ a_0\, t^5 + a_1\, t^6 + a_2\, t^7 + a_3\, t^8 + a_4\, t^9 + a_5\, t^{10} + \ldots$$
$$(\%o131)\ a_0\, t^5 + a_1\, t^6 + a_2\, t^7 + a_3\, t^8 + a_4\, t^9 + a_5\, t^{10} + \ldots$$

1.3.5 *Identity respect to product*

There exists a function $c(t)$ such that $a(t)c(t) = a(t)$, where $a(t) = \mathcal{G}(a_n)$ and where $c(t) = \mathcal{G}(c_n)$? To answer we apply the convolution property of $\mathcal{G}$:

$$\mathcal{G}(a_n)\mathcal{G}(c_n) = \mathcal{G}\left(\sum_{k=0}^{n} a_k\, c_{n-k} \right)$$

We impose the equality as desired $\mathcal{G}(a_n) = \mathcal{G}\left(\sum_{k=0}^{n} a_k\, c_{n-k}\right)$ which holds if the sequences to which $\mathcal{G}$ is applied to are the same. Hence:

$$a_n = \sum_{k=0}^{n} a_k\, c_{n-k} = a_0 c_n + a_1 c_{n-1} + \ldots + a_{n-1} c_1 + a_n c_0$$

The above equality is true if and only if $c_0 = 1$ and $\forall i \in \{1, \ldots, n\} : c_i = 0$. The sequence $\{c_n\}_{n \in \mathbb{N}}$ is something like $(1, 0, 0, 0, 0, \ldots)$ and its formal power series $c(t) = \mathcal{G}(c_n) = 1 + 0t + 0t^2 + 0t^3 + \ldots$, hence we rename $c(t) = 1(t)$.

1.3.6 *Inverse of a formal power series*

There exists a function $b(t)$ such that $a(t)b(t) = 1(t)$, where $a(t) = \mathcal{G}(a_n)$ and where $b(t) = \mathcal{G}(b_n)$? Lets study the product $a(t)b(t)$:

$$a(t)b(t) = a_0 \, b_0 \quad + (a_1 \, b_0 + a_0 \, b_1) \, t + (a_2 \, b_0 + a_1 \, b_1 + a_0 \, b_2) \, t^2 \quad + \dots$$
$$= 1 \qquad\qquad\qquad\qquad\quad +0 + 0 \qquad\qquad\qquad\qquad\qquad\quad + \dots$$

The product $a(t)b(t)$, reported in (%o112) , has to be equal to $1(t)$, this allow to build a system, which we solve with Maxima, obtaining the solutions reported in (%o113) :

```
(%i109) makeSequence(t,j):=sum (j[i]*t^i, i, 0, inf)$
        makeSum(n):=sum (a[i]*b[n-i], i, 0, n)$
        as:taylor(makeSequence(t,a), t, 0, 4);
        bs:taylor(makeSequence(t,b), t, 0, 4);
        prod:as*bs;
        solve([makeSum(0)=1, makeSum(1)=0, makeSum(2)=0,
        makeSum(3)=0, makeSum(4)=0],[b[0],b[1],b[2],b[3],b[4]]);
```

(%o110) $a_0 + a_1 \, t + a_2 \, t^2 + a_3 \, t^3 + a_4 \, t^4 + \dots$
(%o111) $b_0 + b_1 \, t + b_2 \, t^2 + b_3 \, t^3 + b_4 \, t^4 + \dots$
(%o112) $a_0 \, b_0 + (a_1 \, b_0 + a_0 \, b_1) \, t + (a_2 \, b_0 + a_1 \, b_1 + a_0 \, b_2) \, t^2 +$
$(a_3 \, b_0 + a_2 \, b_1 + a_1 \, b_2 + a_0 \, b_3) \, t^3 + (a_4 \, b_0 + a_3 \, b_1 + a_2 \, b_2 + a_1 \, b_3 + a_0 \, b_4) \, t^4 +$
$\dots$
(%o113) $[[b_0 = \dfrac{1}{a_0}, b_1 = -\dfrac{a_1}{a_0^2}, b_2 = -\dfrac{a_0 \, a_2 - a_1^2}{a_0^3}, b_3 = -\dfrac{a_0^2 \, a_3 - 2 \, a_0 \, a_1 \, a_2 + a_1^3}{a_0^4},$
$b_4 = -\dfrac{a_0^3 \, a_4 + a_0^2 \, (-2 \, a_1 \, a_3 - a_2^2) + 3 \, a_0 \, a_1^2 \, a_2 - a_1^4}{a_0^5}]]$

So the inverse of $a(t)$ is the function $a^{\langle -1 \rangle}(t) = b(t)$ such that $b(t) = \mathcal{G}(b_n)$, where $\{b_n\}_{n \in \mathbb{N}}$ is such that b_i is the i-th (zero based indexing) solution reported in (%o113) . It is interesting to observe that it is possible to find $a^{\langle -1 \rangle}(t)$ if and only if the first term of $\{a_n\}_{n \in \mathbb{N}}$ is such that $a_0 \neq 0$.

Exercise 1.3.1. *Let* $a(t) = 1 - t$ *be a formal power series such that* $a(t) = \mathcal{G}(a_n)$. *Find the inverse* $a^{\langle -1 \rangle}(t)$ *and the "inverse" sequence* $\{b_n\}_{n \in \mathbb{N}}$ *such that* $a^{\langle -1 \rangle}(t) = \mathcal{G}(b_n)$.

Sol. We can find it immediately by setting $a(t)b(t) = 1(t)$ and resolving respect to $b(t)$, obtaining $b(t) = \frac{1}{1-t}$ (to find the term b_i of the sequence just expand $b(t)$ using Taylor and get the coefficients).

To find the sequence $\{b_n\}_{n \in \mathbb{N}}$ such that $a^{\langle -1 \rangle}(t) = \mathcal{G}(b_n)$, we use Maxima (where `sol` is the vector obtained in the previous Maxima output): in (%o91) we report the coefficients of the sequence $\{b_n\}_{n \in \mathbb{N}}$ using the solutions of the system reported in (%o113) of the above paragraph, while in (%o93) we expand the function $b(t) = \frac{1}{1-t}$ using Taylor, both methods returns the sequence $\{b_n\}_{n \in \mathbb{N}} = (1, 1, 1, 1, 1, \dots)$.

```
(%i90)   as:[1,-1,0,0,0];
         ev(sol[1],a[0]=as[1],a[1]=as[2],a[2]=as[3],a[3]=as[4],a[4]=as[5]);
         b(t):=1/(1-t);
         taylor(b(t),t,0,8);
```

(%o90) $[1, -1, 0, 0, 0]$

(%o91) $[b_0 = 1, b_1 = 1, b_2 = 1, b_3 = 1, b_4 = 1]$

(%o92) $b(t) := \dfrac{1}{1-t}$

(%o93) $1 + t + t^2 + t^3 + t^4 + t^5 + t^6 + t^7 + t^8 + \ldots$

So the sequence $\{b_n\}_{n \in \mathbb{N}} = \{1_n\}_{n \in \mathbb{N}} = (1, 1, 1, 1, 1, 1, \ldots)$ is the "inverse" of $\{a_n\}_{n \in \mathbb{N}} = (1, -1, 0, 0, 0, 0, \ldots)$, and we've found the following important result:

$$b(t) = \mathcal{G}(1_n) = \frac{1}{1-t} \tag{1.3}$$

$\square$

1.3.7 *Derivation for $\mathcal{G}$*

Let $\{a_n\}_{n \in \mathbb{N}}$ and $\{na_n\}_{n \in \mathbb{N}}$ be two sequences. We study how we can obtain $\mathcal{G}(na_n)$, reported in (%o212) from $\mathcal{G}(a_n)$, reported in (%o211), with Maxima:

```
(%i207) makeSum(n):=sum (a[i]*t^i, i, 0, n)$
        makeDerSum(n):=sum (i*a[i]*t^i, i, 0, n)$
        an(n):=taylor(makeSum(n),t,0,6)$
        deran(n):=taylor(makeDerSum(n),t,0,6)$
        an(6);
        deran(6);
        t*diff(an, t);
```

(%o211) $a_0 + a_1 t + a_2 t^2 + a_3 t^3 + a_4 t^4 + a_5 t^5 + a_6 t^6 + \ldots$

(%o212) $a_1 t + 2 a_2 t^2 + 3 a_3 t^3 + 4 a_4 t^4 + 5 a_5 t^5 + 6 a_6 t^6 + \ldots$

(%o213) $a_1 t + 2 a_2 t^2 + 3 a_3 t^3 + 4 a_4 t^4 + 5 a_5 t^5 + 6 a_6 t^6 + \ldots$

Starting from $\mathcal{G}(a_n)$ we derive term-wise, obtaining $a_1 + 2a_2 t + 3a_3 t^2 + \ldots + a_n t^{n-1} + \ldots$, which in turn we multiply by t, getting $\mathcal{G}(na_n)$. In compact form:

$$\mathcal{G}(na_n) = t \frac{\partial}{\partial t} (\mathcal{G}(a_n))$$

1.3.8 *Identity Principle*

Let $\{a_n\}_{n \in \mathbb{N}}$ and $\{b_n\}_{n \in \mathbb{N}}$ be two sequences, the *Identity Principle* states:

$$\forall n \in \mathbb{N} : (a_n = b_n) \leftrightarrow \mathcal{G}(a_n) = \mathcal{G}(b_n)$$

Exercise 1.3.2. *Suppose we forget the result shown in Equation 1.3 about* $\mathcal{G}(1_n) = \frac{1}{1-t}$, *how can be used the* Identity Principle *to prove that result?*

Sol. In order to use the principle we've to find two sequences which are equal for all $n \in \mathbb{N}$. Since $\{a_n\}_{n\in\mathbb{N}} = (1,1,1,1,\ldots)$ we can build another sequence $\{a_{n+1}\}_{n\in\mathbb{N}}$ using $\{a_n\}_{n\in\mathbb{N}}$. But the index n for the former and $n+1$ for the latter doesn't change the sequence, also $\{a_{n+1}\}_{n\in\mathbb{N}} = (1,1,1,1,\ldots)$, hence $\forall n \in \mathbb{N}$: $a_n = a_{n+1}$. By the principle $\mathcal{G}(a_n) = \mathcal{G}(a_{n+1})$ holds. We apply the translation property of $\mathcal{G}$:

$$\mathcal{G}(a_n) = \mathcal{G}(a_{n+1}) = \frac{\mathcal{G}(a_n) - a_0}{t} \quad \text{by forward translation}$$

$$\mathcal{G}(a_n) = \frac{\mathcal{G}(a_n) - 1}{t} \quad \text{where } a_0 = 1 \text{ by def of } \{a_n\}_{n\in\mathbb{N}}$$

$$\mathcal{G}(a_n) = \mathcal{G}(1_n) = \frac{1}{1-t}$$

What we now remember! □

The exercise above let us to generalize a little. Let $\{a_n\}_{n\in\mathbb{N}}$ such that $a_n = c$. When the sequence is known a priory as in this case (and in the exercise above), it is useful to study the ratio between consecutive terms:

$$\frac{a_{n+1}}{a_n} = \frac{c}{c} \quad \rightarrow \quad a_{n+1} = a_n \quad \forall n \in \mathbb{N}$$

By *Identity Principle* follows:

$$\mathcal{G}(a_n) = \mathcal{G}(a_{n+1}) = \frac{\mathcal{G}(a_n) - a_0}{t} \quad \text{where } a_0 = c \text{ by def of } \{a_n\}_{n\in\mathbb{N}}$$

$$\mathcal{G}(a_n) = \frac{\mathcal{G}(a_n) - c}{t}$$

$$\mathcal{G}(a_n) = c\,\mathcal{G}(1_n) = \frac{c}{1-t}$$

1.3.9   *Composition for $\mathcal{G}$*

Let $\{a_n\}_{n\in\mathbb{N}}$ and $\{b_n\}_{n\in\mathbb{N}}$ be two sequences, we can compose their formal power series in the following way:

$$a(t) = \sum_{k\geqslant 0} a_k t^k \quad b(t) = \mathcal{G}(b_n)$$

$$\mathcal{G}(a_n) \circ \mathcal{G}(b_n) = a(b(t)) = \sum_{k=0}^{\infty} a_k b(t)^k = \sum_{k=0}^{\infty} a_k \mathcal{G}(b_n)^k$$

1.3.10 $\mathcal{G}(\frac{1}{n}a_n) = \int \frac{\mathcal{G}(a_n)-a_0}{t}dt$

Let $b_n = \frac{1}{n}a_n$ be the generic term of a new sequence $\{b_n\}_{n\in\mathbb{N}}$. It follows that $a_n = nb_n$, but it isn't sufficient to apply the *Identity Principle* because in $n = 0$ the term b_0 isn't defined. So we've to augment $a_n = nb_n + a_0\delta_{n,0}$:

$$a_n = nb_n + a_0\delta_{n,0}$$

$$\downarrow$$

$$\mathcal{G}(a_n) = \mathcal{G}(nb_n) + \mathcal{G}(a_0\delta_{n,0})$$

$$\mathcal{G}(a_n) = t\frac{\partial}{\partial t}(\mathcal{G}(b_n)) + a_0\mathcal{G}(\delta_{n,0})$$

$$\mathcal{G}(a_n) = t\frac{\partial}{\partial t}\left(\mathcal{G}(\frac{1}{n}a_n)\right) + a_0 \quad \text{by } \mathcal{G}(\delta_{n,0}) = 1$$

$$\frac{\partial}{\partial t}\left(\mathcal{G}(\frac{1}{n}a_n)\right) = \frac{\mathcal{G}(a_n)-a_0}{t}$$

$$\mathcal{G}(\frac{1}{n}a_n) = \int \frac{\mathcal{G}(a_n)-a_0}{t}dt$$

1.3.11 $\mathcal{G}(\frac{1}{n+1}a_n) = \frac{1}{t}\int \mathcal{G}(a_n)dt$

Let $g_n = a_{n-1}$, where $g_0 = 0$ (otherwise a_{-1} hasn't meaning), to have:

$$\frac{1}{n+1}a_n = \frac{1}{n+1}g_{n+1}$$

We apply the *Identity Principle* directly because we've just renamed the general term of the sequence:

$$\frac{1}{n+1}a_n = \frac{1}{n+1}g_{n+1}$$

$$\downarrow$$

$$\mathcal{G}(\frac{1}{n+1}a_n) = \mathcal{G}(\frac{1}{n+1}g_{n+1})$$

$$\mathcal{G}(\frac{1}{n+1}a_n) = \frac{\mathcal{G}(\frac{1}{n}g_n)-g_0}{t} \quad \text{by translation of } \mathcal{G}$$

$$\mathcal{G}(\frac{1}{n+1}a_n) = \frac{1}{t}\mathcal{G}(\frac{1}{n}g_n) \quad g_0 = 0$$

$$\mathcal{G}(\frac{1}{n+1}a_n) = \frac{1}{t}\int \frac{\mathcal{G}(g_n)-g_0}{t}dt \quad \text{by previous result}$$

$$\mathcal{G}(\frac{1}{n+1}a_n) = \frac{1}{t}\int \frac{\mathcal{G}(a_{n-1})}{t}dt$$

Recalling that $\mathcal{G}(a_{n-1}) = t\mathcal{G}(a_n)$, we have:

$$\mathcal{G}(\frac{1}{n+1}a_n) = \frac{1}{t}\int \mathcal{G}(a_n)dt \tag{1.4}$$

1.4   EXAMPLES OF GENERATING FUNCTIONS

In the following sections we study some interesting sequences. Every section's title shows the sequence under study in that section and its body contains the building of the relative generating function.

1.4.1 $\{c^n\}_{n \in \mathbb{N}}$

$$\mathcal{G}(c^n) = \sum_{n \geqslant 0} c^n t^n = \sum_{n \geqslant 0} 1_n \cdot (ct)^n = \sum_{n \geqslant 0} 1_n \cdot (ct)^n$$

$$= \mathcal{G}(1_n) \circ (ct) = \frac{1}{1 - ct} \quad \text{by composition of } \mathcal{G}$$

Doing a simple check with Maxima:

```
(%i226) b(t):=1/(1-c*t)$
        taylor(b(t),t,0,8);
```

(%o227) $1 + c\,t + c^2\,t^2 + c^3\,t^3 + c^4\,t^4 + c^5\,t^5 + c^6\,t^6 + c^7\,t^7 + c^8\,t^8 + \ldots$

1.4.2 $\{(-1)^n a_n\}_{n \in \mathbb{N}}$

$$\mathcal{G}((-1)^n a_n) = \sum_{n \geqslant 0} (-1)^n a_n t^n = \sum_{n \geqslant 0} a_n (-t)^n = \mathcal{G}(a_n) \circ (-t) = a(-t)$$

1.4.3 $\{n\}_{n \in \mathbb{N}}$

$$\mathcal{G}(n) = \mathcal{G}(n \cdot 1_n) = t \frac{\partial}{\partial t} (\mathcal{G}(1_n)) = \frac{t}{(1-t)^2}$$

Doing a little check with Maxima:

```
(%i244) b(t):=1/(1-t)$
        fg(t):=t*diff(b(t),t);
        fg(t);
        taylor(fg(t),t,0,8);
```

(%o245) $\mathrm{fg}\,(t) := t\,\mathrm{diff}\,(b\,(t)\,,t)$

(%o246) $\dfrac{t}{(1-t)^2}$

(%o247) $t + 2\,t^2 + 3\,t^3 + 4\,t^4 + 5\,t^5 + 6\,t^6 + 7\,t^7 + 8\,t^8 + \ldots$

1.4.4 $\{n^2\}_{n \in \mathbb{N}}$

Using the result of the previous exercise:

$$\mathcal{G}(n^2) = \mathcal{G}(nn) = t\frac{\partial}{\partial t}\left(\mathcal{G}(n)\right) = -\frac{t^2 + t}{t^3 - 3\,t^2 + 3\,t - 1}$$

Just a little check:

```
(%i252) snder(t):=t*diff(fg(t),t)$
        ratsimp(snder(t));
        taylor(snder(t),t,0,8);
```

$(\%o253)\ -\dfrac{t^2 + t}{t^3 - 3\,t^2 + 3\,t - 1}$

$(\%o254)\ t + 4\,t^2 + 9\,t^3 + 16\,t^4 + 25\,t^5 + 36\,t^6 + 49\,t^7 + 64\,t^8 + \ldots$

1.4.5 $\{a_n = 1 \text{ if } n \text{ is even}, 0 \text{ otherwise}\}_{n \in \mathbb{N}}$

We compare $\mathcal{G}(a_n)$ with $\mathcal{G}(1_n)$:

$$\mathcal{G}(a_n) = 1 + t^2 + t^4 + t^6 + \ldots$$
$$\mathcal{G}(1_n) = 1 + t^1 + t^2 + t^3 + \ldots$$

So we can compose $\mathcal{G}(1_n)$ which is known with the function t^2 to obtaining $\mathcal{G}(a_n)$:

$$\mathcal{G}(a_n) = \frac{1}{1-t} \circ t^2 = \frac{1}{1-t^2}$$

Simple check:

```
(%i8)   b(t):=1/(1-t)$
        taylor(b(t^2),t,0,10);
```

$(\%o9)\quad 1 + t^2 + t^4 + t^6 + t^8 + t^{10} + \ldots$

1.4.6 $\{a_n = (1,0,0,1,0,0,1,0,0,1,\ldots)\}_{n \in \mathbb{N}}$

We compare $\mathcal{G}(a_n)$ with $\mathcal{G}(1_n)$ again:

$$\mathcal{G}(a_n) = 1 + t^3 + t^6 + t^9 \ldots$$
$$\mathcal{G}(1_n) = 1 + t^1 + t^2 + t^3 + \ldots$$

So we can compose $\mathcal{G}(1_n)$ which is known with the function t^3 to obtaining $\mathcal{G}(a_n)$:

$$\mathcal{G}(a_n) = \frac{1}{1-t} \circ t^3 = \frac{1}{1-t^3}$$

Simple check:

```
(%i10)  taylor(b(t^3),t,0,10);
```

$(\%o10) \quad 1 + t^3 + t^6 + t^9 + \ldots$

1.4.7 $\{a_n = (1, 0, 0, 2, 0, 0, 4, 0, 0, 8, \ldots)\}_{n \in \mathbb{N}}$

We compare $\mathcal{G}(a_n)$ with $\mathcal{G}(2^n)$ again:

$$\mathcal{G}(a_n) = 1 + 2t^3 + 4t^6 + 8t^9 \ldots$$
$$\mathcal{G}(2^n) = 1 + 2t^1 + 4t^2 + 8t^3 + \ldots$$

So we can compose $\mathcal{G}(2^n)$ which is known with the function t^3 to obtaining $\mathcal{G}(a_n)$:

$$\mathcal{G}(a_n) = \frac{1}{1 - 2t} \circ t^3 = \frac{1}{1 - 2t^3}$$

Simple check:

```
(%i13)  b(t):=1/(1-2*t)$
        taylor(b(t^3),t,0,15);
```

$(\%o14) \quad 1 + 2\,t^3 + 4\,t^6 + 8\,t^9 + 16\,t^{12} + 32\,t^{15} + \ldots$

1.4.8 $\{a_n = \frac{1}{n}\}_{n \in \mathbb{N}} \wedge a_0 = 0$

We try to understand the generating function $\mathcal{G}(a_n)$:

$$\mathcal{G}(\frac{1}{n}) = t + \frac{1}{2}t^2 + \frac{1}{3}t^3 + \frac{1}{4}t^4 + \ldots$$
$$\frac{a_{n+1}}{a_n} = \frac{n}{n+1} \quad \text{by ratio of consecutive terms}$$
$$(n+1)a_{n+1} = na_n$$

In order to apply the identity principle, it is required to check if the equivalence holds for all n. But this isn't the case for $n = 0 : a_1 = 1 \neq 0 = a_0$. So we've to augment the second member with the term $\delta_{n,k}$ such that $\delta_{n,k} = 1$ if $n = k, 0$ otherwise. Augmenting we obtain $(n+1)a_{n+1} = na_n + \delta_{n,0}, \forall n \in \mathbb{N}$, in

particular for $n = 0 : a_1 = 1 = 1 = a_0 + \delta_{0,0}$. Now it is possible to apply the identity principle, calling $b_{n+1} = (n+1)a_{n+1}$ (hence $b_n = na_n$ and $b_0 = 0$):

$$b_{n+1} = na_n + \delta_{n,0}$$

$$\downarrow$$

$$\mathcal{G}(b_{n+1}) = \mathcal{G}(na_n) + \mathcal{G}(\delta_{n,0})$$

$$\frac{\mathcal{G}(b_n) - b_0}{t} = \mathcal{G}(na_n) + \mathcal{G}(\delta_{n,0})$$

$$\frac{\mathcal{G}(na_n)}{t} = \mathcal{G}(na_n) + \mathcal{G}(\delta_{n,0})$$

$$\frac{t\frac{\partial}{\partial t}(\mathcal{G}(a_n))}{t} = \mathcal{G}(na_n) + \mathcal{G}(\delta_{n,0})$$

$$\frac{\partial}{\partial t}(\mathcal{G}(a_n)) = t\frac{\partial}{\partial t}(\mathcal{G}(a_n)) + 1$$

$$\frac{\partial}{\partial t}(\mathcal{G}(a_n)) = \frac{1}{1-t}$$

Integrating both members:

$$\mathcal{G}(\frac{1}{n}) = \log\frac{1}{1-t} \tag{1.5}$$

Simple check:

```
(%i15)  b(t):=log(1/(1-t))$
        taylor(b(t),t,0,13);
```

$$(\%\text{o}16) \quad t + \frac{t^2}{2} + \frac{t^3}{3} + \frac{t^4}{4} + \frac{t^5}{5} + \frac{t^6}{6} + \frac{t^7}{7} + \frac{t^8}{8} + \frac{t^9}{9} + \frac{t^{10}}{10} + \frac{t^{11}}{11} + \frac{t^{12}}{12} + \frac{t^{13}}{13} + \dots$$

1.4.9 $\{H_n = \sum_{k=1}^{n} \frac{1}{k}\}_{n\in\mathbb{N}}$

$$\mathcal{G}\left(\sum_{k=1}^{n} \frac{1}{k}\right) = \mathcal{G}\left(\sum_{k=1}^{n} \frac{1}{k} 1_{n-k}\right) =$$

$$= \mathcal{G}\left(\frac{1}{n}\right)\mathcal{G}(1_n) \quad \text{by convolution of } \mathcal{G}$$

$$= \left(\log\frac{1}{1-t}\right)\frac{1}{1-t} \quad \text{by Equation 1.5 and Equation 1.3}$$

Simple check:

```
(%i66)  harmonic(n) := sum(1/k,k,1,n);
        b(t):=log(1/(1-t))*(1/(1-t));
        map(harmonic, makelist(n,n,1,10));
        taylor(b(t),t,0,10);
```

$(\%o66) \quad \text{harmonic}\,(n) := \sum_{k=1}^{n} \frac{1}{k}$

$(\%o67) \quad b\,(t) := \log\left(\frac{1}{1-t}\right) \frac{1}{1-t}$

$(\%o68) \quad [1, \dfrac{3}{2}, \dfrac{11}{6}, \dfrac{25}{12}, \dfrac{137}{60}, \dfrac{49}{20}, \dfrac{363}{140}, \dfrac{761}{280}, \dfrac{7129}{2520}, \dfrac{7381}{2520}]$

$(\%o69) \quad t + \dfrac{3\,t^2}{2} + \dfrac{11\,t^3}{6} + \dfrac{25\,t^4}{12} + \dfrac{137\,t^5}{60} + \dfrac{49\,t^6}{20} + \dfrac{363\,t^7}{140} + \dfrac{761\,t^8}{280} + \dfrac{7129\,t^9}{2520} +$
$\dfrac{7381\,t^{10}}{2520} + \dots$

1.4.10 $\{F_{n+2} = F_{n+1} + F_n\}_{n \in \mathbb{N}} \wedge F_1 = 1 \wedge F_0 = 0$

It is possible to apply the *Identity Principle* directly because the equality $F_{n+2} = F_{n+1} + F_n$ holds $\forall n \in \mathbb{N}$ (would be problem if we formulate the same as $F_n = F_{n-1} + F_{n-2}$ for $n = 0$):

$$F_{n+2} = F_{n+1} + F_n$$
$$\downarrow$$
$$\mathcal{G}(F_{n+2}) = \mathcal{G}(F_{n+1}) + \mathcal{G}(F_n)$$
$$\frac{\mathcal{G}(F_n) - F_0 - F_1 t}{t^2} = \frac{\mathcal{G}(F_n) - F_0}{t} + \mathcal{G}(F_n) \quad \text{by translation of } \mathcal{G}$$
$$\frac{\mathcal{G}(F_n) - t}{t^2} = \frac{\mathcal{G}(F_n)}{t} + \mathcal{G}(F_n)$$
$$\mathcal{G}(F_n) - t = t\mathcal{G}(F_n) + t^2 \mathcal{G}(F_n)$$

Finally we get:

$$\mathcal{G}(F_n) = -\frac{t}{t^2 + t - 1} = \frac{t}{1 - t - t^2} \tag{1.6}$$

Simple check:

```
(%i779) b(t):=t/(1-t-t^2)$
        map(fib, makelist(n,n,1,10));
        taylor(b(t),t,0,10);
```

$(\%o780) \quad [1, 1, 2, 3, 5, 8, 13, 21, 34, 55]$

$(\%o781) \quad t + t^2 + 2\,t^3 + 3\,t^4 + 5\,t^5 + 8\,t^6 + 13\,t^7 + 21\,t^8 + 34\,t^9 + 55\,t^{10} + \dots$

1.4.11 $\{T_{n+3} = T_{n+2} + T_{n+1} + T_n\}_{n\in\mathbb{N}} \wedge T_2 = 1 \wedge T_1 = 1 \wedge T_0 = 0$

It is possible to apply the *Identity Principle* directly because the equality $T_{n+3} = T_{n+2} + T_{n+1} + T_n$ holds $\forall n \in \mathbb{N}$:

$$T_{n+3} = T_{n+2} + T_{n+1} + T_n$$
$$\downarrow$$
$$\mathcal{G}(T_{n+3}) = \mathcal{G}(T_{n+2}) + \mathcal{G}(T_{n+1}) + \mathcal{G}(T_n)$$
$$\frac{\mathcal{G}(T_n) - T_0 - T_1 t - T_2 t^2}{t^3} = \frac{\mathcal{G}(T_n) - T_0 - T_1 t}{t^2} + \frac{\mathcal{G}(T_n) - T_0}{t} + \mathcal{G}(T_n) \quad \text{by translation of } \mathcal{G}$$
$$\frac{\mathcal{G}(T_n) - t - t^2}{t^3} = \frac{\mathcal{G}(T_n) - t}{t^2} + \frac{\mathcal{G}(T_n)}{t} + \mathcal{G}(T_n)$$
$$\mathcal{G}(T_n) - t - t^2 = t(\mathcal{G}(T_n) - t) + t^2\mathcal{G}(T_n) + t^3\mathcal{G}(T_n)$$
$$\mathcal{G}(T_n) = \frac{t}{1 - t - t^2 - t^3}$$

Simple check:

```
(%i41)  tribonacci(n):=if is(n=0) then 0 else if is(n=1) then 1 else if
        is(n=2) then 1 else tribonacci(n-1)+tribonacci(n-2)+tribonacci(n-3)\$
        b(t):=(t/(1-t-t^2-t^3));
        map(tribonacci, makelist(n,n,0,9));
        taylor(b(t),t,0,9);
```

$$(\%o42) \quad b(t) := \frac{t}{1 - t - t^2 - t^3}$$
$$(\%o43) \quad [0, 1, 1, 2, 4, 7, 13, 24, 44, 81]$$
$$(\%o44) \quad t + t^2 + 2t^3 + 4t^4 + 7t^5 + 13t^6 + 24t^7 + 44t^8 + 81t^9 + \dots$$

1.4.12 $\{a_n = \binom{p}{n}\}_{n\in\mathbb{N}}$

There isn't a property ready for direct application, so we study the ratio between consecutive terms:

$$\frac{a_{n+1}}{a_n} = \frac{\binom{p}{n+1}}{\binom{p}{n}} = \frac{p!}{(n+1)!(p-n-1)!}\frac{n!(p-n)!}{p!} = \frac{p-n}{n+1}$$
$$\downarrow$$
$$(n+1)a_{n+1} = (p-n)a_n$$

Before applying the *Identity Principle* we check if the equivalence holds for all n, in particular for $n = 0 : a_1 = p = p = pa_0$. So we can apply the principle, calling $b_n = na_n$:

$$b_{n+1} = (p-n)a_n$$

$$\downarrow$$

$$\mathcal{G}(b_{n+1}) = \mathcal{G}((p-n)a_n)$$

$$\frac{\mathcal{G}(b_n) - b_0}{t} = \mathcal{G}((p-n)a_n)$$

$$\frac{\mathcal{G}(b_n)}{t} = \mathcal{G}((p-n)a_n)$$

$$\frac{\mathcal{G}(na_n)}{t} = \mathcal{G}((p-n)a_n)$$

$$\frac{t\frac{\partial}{\partial t}(\mathcal{G}(a_n))}{t} = \mathcal{G}((p-n)a_n)$$

$$\frac{\partial}{\partial t}(\mathcal{G}(a_n)) = p\mathcal{G}(a_n) - \mathcal{G}(na_n)$$

$$\frac{\partial}{\partial t}(\mathcal{G}(a_n)) = p\mathcal{G}(a_n) - t\frac{\partial}{\partial t}(\mathcal{G}(a_n))$$

$$(1+t)\frac{\partial}{\partial t}(\mathcal{G}(a_n)) = p\mathcal{G}(a_n)$$

$$\frac{a'(t)}{a(t)} = \frac{p}{1+t} \quad \text{by calling } a(t) = \mathcal{G}(a_n)$$

$$\log a(t) = \int \frac{p}{1+t}dt = p\log(1+t) + c$$

We compute the constant c for $t = 0 : \log 1 = p \log 1 + c$, with $a(0) = a_0 + 0t + 0t^2 + 0t^3 + \ldots = 1$, hence $c = 0$.

$$\log a(t) = \log(1+t)^p$$

$$\downarrow \quad \text{exping both members}$$

$$a(t) = \mathcal{G}\left(\binom{p}{n}\right) = (1+t)^p$$

Simple check:

```
(%i45)  b(t):=p/(1+t);
        integrate(b(t),t);
        a(t) := (1+t)^p$
        curriedBinomial(n):=binomial(p,n)$
        map(ratsimp, map(curriedBinomial, makelist(n,n,0,5)));
        taylor(a(t),t,0,5);
```

$(\%o45) \quad b(t) := \dfrac{p}{1+t}$

$(\%o46) \quad p\log(t+1)$

$(\%o49) \quad [1, p, \dfrac{p^2 - p}{2}, \dfrac{p^3 - 3p^2 + 2p}{6}, \dfrac{p^4 - 6p^3 + 11p^2 - 6p}{24},$

$$\frac{p^5 - 10\,p^4 + 35\,p^3 - 50\,p^2 + 24\,p}{120}\big]$$

(%o50) $1 + p\,t + \dfrac{(p^2 - p)\ t^2}{2} + \dfrac{(p^3 - 3\,p^2 + 2\,p)\ t^3}{6} + \dfrac{(p^4 - 6\,p^3 + 11\,p^2 - 6\,p)\ t^4}{24} +$

$\dfrac{(p^5 - 10\,p^4 + 35\,p^3 - 50\,p^2 + 24\,p)\ t^5}{120} + \ldots$

1.4.13 $\{a_n = \binom{2n}{n}\}_{n \in \mathbb{N}}$

There isn't a property ready for direct application, so we study the ratio between consecutive terms:

$$\frac{a_{n+1}}{a_n} = \frac{\binom{2n+2}{n+1}}{\binom{2n}{n}} = \frac{(2n+2)!}{(n+1)!(n+1)!}\frac{n!n!}{(2n)!} = \frac{(2n+2)(2n+1)}{(n+1)^2} = \frac{2(2n+1)}{n+1}$$

$$\downarrow$$

$$(n+1)a_{n+1} = 2(2n+1)a_n$$

Before applying the *Identity Principle* we check if the equivalence holds for all n, in particular for $n = 0 : a_1 = 2 = 2 = 2a_0$ and $a_0 = 1$. So we can apply the principle, calling $b_n = na_n$:

$$b_{n+1} = (n+1)a_{n+1}$$

$$\downarrow$$

$$\mathcal{G}(b_{n+1}) = \mathcal{G}(2(2n+1)a_n)$$

$$\frac{\mathcal{G}(b_n) - b_0}{t} = \mathcal{G}(2(2n+1)a_n)$$

$$\frac{\mathcal{G}(b_n)}{t} = \mathcal{G}(2(2n+1)a_n)$$

$$\frac{\mathcal{G}(na_n)}{t} = \mathcal{G}(2(2n+1)a_n)$$

$$\frac{t\frac{\partial}{\partial t}(\mathcal{G}(a_n))}{t} = \mathcal{G}(2(2n+1)a_n)$$

$$\frac{\partial}{\partial t}(\mathcal{G}(a_n)) = 4\mathcal{G}(na_n) + 2\mathcal{G}(a_n)$$

$$\frac{\partial}{\partial t}(\mathcal{G}(a_n)) = 4t\frac{\partial}{\partial t}(\mathcal{G}(a_n)) + 2\mathcal{G}(a_n)$$

$$(1 - 4t)\frac{\partial}{\partial t}(\mathcal{G}(a_n)) = 2\mathcal{G}(a_n)$$

$$\frac{a'(t)}{a(t)} = \frac{2}{1 - 4t} \quad \text{by calling } a(t) = \mathcal{G}(a_n)$$

$$\log a(t) = 2\int \frac{1}{1 - 4t}\,dt = -\frac{1}{2}\log(1 - 4t) + c$$

We compute the constant c for $t = 0 : \log 1 = -\frac{1}{2}\log 1 + c$, with $a(0) = a_0 + 0t + 0t^2 + 0t^3 + \ldots = 1$, hence $c = 0$.

$$\log a(t) = \log \frac{1}{\sqrt{1-4t}}$$

$$\downarrow$$

$$a(t) = \frac{1}{\sqrt{1-4t}}$$

Finally we have found:

$$\mathcal{G}\left(\binom{2n}{n}\right) = \frac{1}{\sqrt{1-4t}} \tag{1.7}$$

Simple check:

```
(%i63)  b(t):=2/(1-4*t);
        integrate(b(t),t);
        a(t) := 1/sqrt(1-4*t)$
        curriedBinomial(n):=binomial(2*n,n)$
        map(ratsimp, map(curriedBinomial, makelist(n,n,0,5)));
        taylor(a(t),t,0,5);
```

$$(\%o63)\quad b\,(t) := \frac{2}{1-4\,t}$$

$$(\%o64)\quad -\frac{\log{(1-4\,t)}}{2}$$

$$(\%o67)\quad [1,2,6,20,70,252]$$

$$(\%o68)\quad 1 + 2\,t + 6\,t^2 + 20\,t^3 + 70\,t^4 + 252\,t^5 + \ldots$$

1.4.14 $\{a_n = \frac{1}{n+1}\binom{2n}{n}\}_{n \in \mathbb{N}}$

$$\mathcal{G}(\frac{1}{n+1}\binom{2n}{n})) = \frac{1}{t}\int \mathcal{G}(\binom{2n}{n})\,dt \quad \text{by Equation 1.4}$$

$$= \frac{1}{t}\int \frac{1}{\sqrt{1-4t}}\,dt \quad \text{by Equation 1.7}$$

Calling $\sqrt{1-4t} = y$ follows $t = \frac{1-y^2}{4}$ and $dt = -\frac{1}{2}y\,dy$, so:

$$= \frac{1}{t}\int_0^t \frac{1}{y}\left(-\frac{1}{2}y\right)dy = -\frac{1}{2t}\int_0^t dy = -\frac{1}{2t}\,[y]_0^t = -\frac{1}{2t}\left[\sqrt{1-4t}\right]_0^t$$

$$= -\frac{1}{2t}(\sqrt{1-4t}-1) = \frac{1-\sqrt{1-4t}}{2t}$$

Simple check:

```
(%i167) b(t):=(1-sqrt(1-4*t))/(2*t)$
        catalan(n):=binomial(2*n,n)/(n+1)$
        map(catalan, makelist(n,n,0,10));
        taylor(b(t),t,0,10);
```

(%o169) $[1, 1, 2, 5, 14, 42, 132, 429, 1430, 4862, 16796]$

(%o170) $1 + t + 2\,t^2 + 5\,t^3 + 14\,t^4 + 42\,t^5 + 132\,t^6 + 429\,t^7 + 1430\,t^8 + 4862\,t^9 + 16796\,t^{10} + \ldots$

1.5 GENERATING FUNCTION QUICKSORT'S AVERAGE CHECKS

In this section we go back to the recurrence about the average number of checks for the basic implementation of quicksort algorithm and we'll try to find the generating function for the sequence C_n of average checks.

We start again from the known recurrence:

$$C_n = n + 1 + \frac{1}{n} \sum_{k=1}^{n} C_k + C_{n-k}$$

$$C_n = n + 1 + \frac{2}{n} \sum_{k=0}^{n-1} C_k$$

$$nC_n = n(n+1) + 2 \sum_{k=0}^{n-1} C_k$$

$$\mathcal{G}(nC_n) = \mathcal{G}(n^2 + n) + 2\mathcal{G}\left(\sum_{k=0}^{n-1} C_k\right)$$

$$\mathcal{G}(nC_n) = \mathcal{G}(n^2) + \mathcal{G}(n) + 2\mathcal{G}\left(\sum_{k=0}^{n-1} C_k\right)$$

Now we remeber the following equalities:

$$\mathcal{G}(n) = \frac{t}{(1-t)^2}$$

$$\mathcal{G}(n^2) = t\frac{\partial}{\partial t}(\mathcal{G}(n)) = \frac{t(1+t)}{(1-t)^3}$$

And it is possible to see $\sum_{k=0}^{n-1} C_k = a_{n-1}$ as a sequence, so $\mathcal{G}(a_{n-1}) = t\mathcal{G}(a_n)$; substituting those rewritings and using the convolution property we get:

$$t\frac{\partial}{\partial t}(\mathcal{G}(C_n)) = \frac{t}{(1-t)^2} + \frac{t(1+t)}{(1-t)^3} + 2t\mathcal{G}\left(\sum_{k=0}^{n} C_k\right)$$

$$t\frac{\partial}{\partial t}(\mathcal{G}(C_n)) = \frac{t - t^2 + t + t^2}{(1-t)^3} + \frac{2t}{1-t}\mathcal{G}(C_n) \quad \text{by } \mathcal{G}\left(\sum_{k=0}^{n} C_k 1_{n-k}\right) = \mathcal{G}(1_n)\mathcal{G}(C_n)$$

$$\frac{\partial}{\partial t}(\mathcal{G}(C_n)) = \frac{2}{(1-t)^3} + \frac{2}{1-t}\mathcal{G}(C_n)$$

We can write the same expression by $\mathcal{G}(C_n) = C(t)$ as a differential equation:

$$C'(t) = \frac{2}{(1-t)^3} + \frac{2}{1-t}C(t) \tag{1.8}$$

with initial condition $C(0) = 0$. Now we study the homogeneous associate equation $\mu(t)$ (excluding for a moment the term $\frac{2}{(1-t)^3}$):

$$\mu'(t) = \frac{2}{1-t}\mu(t) \tag{1.9}$$

$$\frac{\mu'(t)}{\mu(t)} = \frac{2}{1-t} \tag{1.10}$$

$$\log\mu(t) = -2\log(1-t) = \log\frac{1}{(1-t)^2} \quad \text{integrating both members} \tag{1.11}$$

$$\mu(t) = \frac{1}{(1-t)^2} \quad \text{exping both members} \tag{1.12}$$

Now we study the derivative which complies the μ function, solution of the unknown function $C(t)$ we hope to find:

$$\frac{\partial}{\partial t}\left(\frac{C(t)}{\mu(t)}\right) = \frac{C'(t)\mu(t) - C(t)\mu'(t)}{\mu(t)^2}$$

$$= \mu(t)\frac{C'(t) - C(t)\frac{\mu'(t)}{\mu(t)}}{\mu(t)^2}$$

$$= \frac{1}{\mu(t)}(C'(t) - \frac{2}{1-t}C(t)) \quad \text{by Equation 1.10}$$

$$= \frac{1}{\mu(t)}\frac{2}{(1-t)^3} \quad \text{by Equation 1.8}$$

$$= \frac{2}{1-t}$$

Now we integrate both members:

$$\frac{C(t)}{\mu(t)} = -2\log(1-t) + k$$

$$C(t) = \frac{2\log\frac{1}{1-t} + k}{(1-t)^2} \quad \text{by Equation 1.12}$$

Since $C(0) = 0$ it must have $k = 0$, hence:

$$C(t) = \frac{2}{(1-t)^2} \log \frac{1}{1-t} \qquad (1.13)$$

Simple check with Maxima:

```
(%i5)   c(t):=(2/(1-t)^2)*log(1/(1-t))$
        taylor(c(t),t,0,8);
```

$$(\%o6) \quad 2\,t + 5\,t^2 + \frac{26\,t^3}{3} + \frac{77\,t^4}{6} + \frac{87\,t^5}{5} + \frac{223\,t^6}{10} + \frac{962\,t^7}{35} + \frac{4609\,t^8}{140} + \ldots$$

1.6 GENERATING FUNCTION OF QUICKSORT'S AVERAGE SWAPS

In this section we go back again to the recurrence about the average number of swaps for the basic implementation of quicksort algorithm and we'll try to find the generating function for the sequence S_n of average swaps.

We start again from the known recurrence:

$$S_n = \frac{n-2}{6} + \frac{2}{n} \sum_{j=0}^{n-1} S_j$$

$$6nS_n = n(n-2) + 12 \sum_{j=0}^{n-1} S_j$$

with $S_0 = S_1 = 0$. We've to remember to check if the recurrence holds for all n, the previous doesn't for $n = 1$ because $0 \neq -1$. So we fix augmenting the second member with the sequence $\delta_{n,1}$ which returns 1 if and only if $n = 1$, now the recurrence holds for all $n \in \mathbb{N}$. By the *Identity principle* we get:

$$6\mathcal{G}(nS_n) = \mathcal{G}(n^2) - 2\mathcal{G}(n) + 12 \sum_{j=0}^{n-1} S_j + \mathcal{G}(\delta_{n,1})$$

$$6tS'(t) = \frac{t+t^2}{(1-t)^3} - \frac{2t}{(1-t)^2} + 12t\mathcal{G}\left(\sum_{j=0}^{n} S_j\right) + t$$

$$6tS'(t) = \frac{t+t^2 - 2t(1-t) + t(1-t)^3}{(1-t)^3} + \frac{12t}{1-t}S(t) \quad \text{by } \mathcal{G}\left(\sum_{j=0}^{n} S_j 1_{n-j}\right) = \mathcal{G}(1_n)\mathcal{G}(S_n)$$

$$6S'(t) = \frac{1+t - 2(1-t) + (1-t)^3}{(1-t)^3} + \frac{12}{1-t}S(t) \quad \text{by dividing both members by } t$$

Finally we have:

$$S'(t) = \frac{t^2(3-t)}{6(1-t)^3} + \frac{2}{1-t}S(t) \qquad (1.14)$$

with initial condition $S(0) = 0$. Now we study the homogeneous associate equation $\mu(t)$ (excluding for a moment the term $\frac{t^2(3-t)}{6(1-t)^3}$):

$$\mu'(t) \;=\; \frac{2}{1-t}\mu(t) \tag{1.15}$$

$$\frac{\mu'(t)}{\mu(t)} \;=\; \frac{2}{1-t} \tag{1.16}$$

$$\log\mu(t) \;=\; -2\log(1-t) = \log\frac{1}{(1-t)^2} \quad \text{integrating both members} \tag{1.17}$$

$$\mu(t) \;=\; \frac{1}{(1-t)^2} \quad \text{exping both members} \tag{1.18}$$

Now we study the derivative which complies the μ function, solution of the unknown function $S(t)$ we hope to find:

$$\frac{\partial}{\partial t}\left(\frac{S(t)}{\mu(t)}\right) = \frac{S'(t)\mu(t) - S(t)\mu'(t)}{\mu(t)^2}$$

$$= \mu(t)\frac{S'(t) - S(t)\frac{\mu'(t)}{\mu(t)}}{\mu(t)^2}$$

$$= \frac{1}{\mu(t)}\left(S'(t) - \frac{2}{1-t}S(t)\right) \quad \text{by Equation 1.16}$$

$$= \frac{1}{\mu(t)}\frac{t^2(3-t)}{6(1-t)^3} \quad \text{by Equation 1.14}$$

$$= \frac{t^2(3-t)}{6(1-t)} \quad \text{by Equation 1.18}$$

We do simple manipulation in order to ease the integration:

$$\frac{\partial}{\partial t}\left(\frac{S(t)}{\mu(t)}\right) = \frac{t^2(3-t)}{6(1-t)} = \frac{t^2(1-t+2)}{6(1-t)} = \frac{1}{6}\left(t^2 + \frac{2t^2}{1-t}\right) = \frac{1}{6}\left(t^2 + \frac{2t^2-2+2}{1-t}\right) =$$

$$= \frac{1}{6}\left(t^2 + \frac{2(t-1)(t+1)}{1-t} + \frac{2}{1-t}\right) = \frac{1}{6}\left(t^2 - \frac{2(1-t)(t+1)}{1-t} + \frac{2}{1-t}\right) =$$

$$= \frac{1}{6}\left(t^2 - 2(t+1) + \frac{2}{1-t}\right)$$

Now we integrate both members:

$$\frac{S(t)}{\mu(t)} = \frac{1}{6}\left(\frac{t^3}{3} - t^2 - 2t - 2\log(1-t) + k\right)$$

$$S(t) = \frac{1}{6(1-t)^2}\left(\frac{t^3}{3} - t^2 - 2t - 2\log(1-t) + k\right) \quad \text{by Equation 1.18}$$

After simplification and $k = 0$ by the initial conditions $S(0) = S(1) = 0$ we get:

$$S(t) = \frac{1}{3(1-t)^2}\log\frac{1}{1-t} + \frac{t^3 - 3t^2 - 6t}{18(1-t)^2} \tag{1.19}$$

Simple check with Maxima:

```
(%i3)   s(t):=(1/(3*(1-t)^2))*log(1/(1-t)) + (t^3-3*t^2-6*t)/(18*(1-t)^2)$
        taylor(s(t),t,0,8);
```

(%o4) $\dfrac{t^3}{6} + \dfrac{5\,t^4}{12} + \dfrac{11\,t^5}{15} + \dfrac{199\,t^6}{180} + \dfrac{961\,t^7}{630} + \dfrac{1669\,t^8}{840} + ...$

1.7 [t^n] INVERSE OPERATOR OF GENERATING FUNCTION OPERATOR

In this section we'll study the inverse operator $[t^n]$ of $\mathcal{G}$, which consume a generating function $g(t)$, where $g(t) = \mathcal{G}(\{a_n\}_{n\in\mathbb{N}})$, and returns the a_n coefficient of the sequence $\{a_n\}_{n\in\mathbb{N}}$ (in other words, $[t^n]a(t) = f(n)$, so $[t^n]$ returns a function of n), formally $[t^n]g(t) = a_n$. It is interesting to note that the function $g(t)$ to which $[t^n]$ is applied to depends only on t (in the last lectures we will see a special case of application $[t^n]g(t,n)$).

1.7.1 *Linearity of* $[t^n]$

$$[t^n](\alpha a(t) + \beta b(t)) = [t^n]\left(\sum_{n=0}^{\infty} (\alpha a_n + \beta b_n)t^n \right) = \alpha a_n + \beta b_n$$

1.7.2 *Translation of* $[t^n]$

Let $s \in \mathbb{Z}$:

$$\begin{aligned}
[t^n]t^s a(t) &= [t^n]t^s(a_0 + a_1 t + a_2 t^2 + \ldots + a_n t^n + \ldots) \\
&= [t^n](a_0 t^s + a_1 t^{s+1} + a_2 t^{s+2} + \ldots + a_n t^{n+s} + \ldots) \\
&= [t^{n-s}]a(t)
\end{aligned}$$

A simple observation:

$$[t^n]\frac{a(t)}{t^s} = [t^{n+s}]a(t)$$

1.7.3 *Derivation of* $[t^n]$

$$a(t) = a_0 + a_1 t + a_2 t^2 + \ldots + a_n t^n + \ldots$$
$$\frac{\partial}{\partial t}(a(t)) = a_1 + 2a_2 t + 3a_3 t^2 + \ldots + (n+1)a_{n+1}t^n + \ldots$$
$$[t^n]\frac{\partial}{\partial t}(a(t)) = (n+1)a_{n+1}$$

It is interesting to note the following fact:

$$[t^n]\frac{\partial}{\partial t}(a(t)) = (n+1)a_{n+1}$$

$$\downarrow$$

$$[t^{n-1}]\frac{\partial}{\partial t}(a(t)) = na_n$$

$$n[t^n]a(t) = na_n \quad \text{by def of } [t^n]$$

So the following holds:

$$[t^n]a(t) = \frac{1}{n}[t^{n-1}]a'(t) \tag{1.20}$$

1.7.4 *Convolution of $[t^n]$*

$$[t^n]a(t)b(t) = [t^n]\mathcal{G}(a_n)\mathcal{G}(b_n) = [t^n]\mathcal{G}\left(\sum_{k=0}^{n} a_k b_{n-k}\right) = \sum_{k=0}^{n} a_k b_{n-k}$$

1.7.5 *Composition of $[t^n]$*

$$[t^n]a(t) \circ b(t) = [t^n]a(b(t)) = [t^n]\sum_{k=0}^{\infty} a_k(\mathcal{G}(b_n))^k$$

$$= \sum_{k=0}^{\infty} a_k[t^n]b(t)^k \quad \text{by linearity of } [t^n]$$

1.7.6 *Netwon's rule*

Let $\alpha, r \in \mathbb{R}$:

$$[t^n](1+\alpha t)^r = \binom{r}{n}\alpha^n \tag{1.21}$$

Proof. We repeatedly apply Equation 1.20 to $[t^n](1 + \alpha t)^r$ getting the following:

$$
\begin{aligned}
[t^n](1 + \alpha t)^r &= \frac{1}{n}[t^{n-1}]r\alpha(1 + \alpha t)^{r-1} \\
&= \frac{r\alpha}{n}\frac{1}{n-1}[t^{n-2}](r-1)\alpha(1 + \alpha t)^{r-2} \\
&= \frac{r\alpha}{n}\frac{(r-1)\alpha}{n-1}\frac{1}{n-2}[t^{n-3}](r-2)\alpha(1 + \alpha t)^{r-3} \\
&\;\;\vdots \\
&= \frac{r\alpha}{n}\frac{(r-1)\alpha}{n-1}\cdots\frac{(r-n+2))\alpha}{n-n+2}\frac{1}{n-n+1}[t^{n-n}](r-n+1)\alpha(1 + \alpha t)^{r-n} \\
&= \frac{r\alpha}{n}\frac{(r-1)\alpha}{n-1}\cdots\frac{(r-n+2))\alpha}{n-n+2}\frac{(r-n+1)\alpha}{n-n+1}[t^0](1 + \alpha t)^{r-n} \\
&= \alpha^n \frac{r(r-1)\cdots(r-n+1)}{n!}[t^0](1 + \alpha t)^{r-n} \\
&= \alpha^n \binom{r}{n}[t^0](1 + \alpha t)^{r-n} \\
&= \alpha^n \binom{r}{n}
\end{aligned}
$$

where $[t^0](1 + \alpha t)^{r-n} = 1$ because, let $f(t) = (1 + \mu t)^\nu$ in:

$$
[t^0]f(t) = [t^0]\left(\sum_{k=0}^{\infty}\frac{f^k(0)}{k!}t^k\right) = f^0(0) = f(0) = 1 \quad \forall \mu, \nu \in \mathbb{R}
$$

$\square$

1.7.7 *Inverting Fibonacci generating function*

In this section we'll study the extract coefficient operator $[t^n]$ applied to the generating function of Fibonacci's sequence.

In Equation 1.6 we found the generating function for Fibonacci's sequence, in order to apply the Newton's rule, we have to reach something like $(1 + \alpha t)^r$, for some $\alpha, r \in \mathbb{R}$. So the idea is to write the term $\frac{t}{1-t-t^2}$ as partial fractions and try to apply $[t^n]$.

First we solve $1 - t - t^2 = 0$ respect t, having two solutions $\phi = \frac{\sqrt{5}-1}{2}$ and $\hat{\phi} = -\frac{\sqrt{5}+1}{2}$ (Maxima results (%o759) and (%o760)).

Those solutions are very interesting and enjoy the following properties (Maxima results (%o761) and (%o762)):

$$\phi + \hat{\phi} = \frac{\sqrt{5}}{2} - \frac{1}{2} - \frac{\sqrt{5}}{2} - \frac{1}{2} = -1$$

$$\phi\hat{\phi} = \frac{(\sqrt{5}-1)(-\sqrt{5}-1)}{4} = \frac{-5+1}{4} = -1$$

$$\downarrow$$

$$\phi + \hat{\phi} = \phi\hat{\phi} = -1$$

Now we use the previous solutions to factor $1 - t - t^2$:

$$1 - t - t^2 = -(t - \phi)\left(t - \hat{\phi}\right)$$

$$= -(-\phi)(-\hat{\phi})\left(1 - \frac{t}{\phi}\right)\left(1 - \frac{t}{\hat{\phi}}\right)$$

$$= -\phi\hat{\phi}\left(1 - \frac{t}{\phi}\right)\left(1 - \frac{t}{\hat{\phi}}\right)$$

$$= \left(1 - \frac{t}{\phi}\right)\left(1 - \frac{t}{\hat{\phi}}\right) \quad \text{by above properties}$$

Now we'll find two constants $A, B \in \mathbb{R}$ to build the partial fractions:

$$\frac{t}{1-t-t^2} = \frac{A}{\left(1 - \frac{t}{\phi}\right)} + \frac{B}{\left(1 - \frac{t}{\hat{\phi}}\right)} = \frac{A\left(1 - \frac{t}{\hat{\phi}}\right) + B\left(1 - \frac{t}{\phi}\right)}{\left(1 - \frac{t}{\phi}\right)\left(1 - \frac{t}{\hat{\phi}}\right)} = \frac{(A+B) - t\left(\frac{A}{\hat{\phi}} + \frac{B}{\phi}\right)}{\left(1 - \frac{t}{\phi}\right)\left(1 - \frac{t}{\hat{\phi}}\right)}$$

So we've to solve the system with two equations in two unknowns: $A + B = 0$ and $\left(\frac{A}{\hat{\phi}} + \frac{B}{\phi}\right) = -1$, which yields $A = \frac{1}{\sqrt{5}}$ and $B = -\frac{1}{\sqrt{5}}$ (as confirmed in (%o765)). So we've found the partial fractions decomposition (as confirmed in (%o767)):

$$\frac{t}{1-t-t^2} = \frac{1}{\sqrt{5}}\frac{1}{1 - \frac{t}{\phi}} - \frac{1}{\sqrt{5}}\frac{1}{1 - \frac{t}{\hat{\phi}}} \tag{1.22}$$

We apply the [t^n] to both members:

$$[t^n]\frac{t}{1-t-t^2} = \frac{1}{\sqrt{5}}[t^n]\frac{1}{1-\frac{t}{\phi}} - \frac{1}{\sqrt{5}}[t^n]\frac{1}{1-\frac{t}{\hat{\phi}}}$$

$$= \frac{1}{\sqrt{5}}[t^n]\left(1-\frac{t}{\phi}\right)^{-1} - \frac{1}{\sqrt{5}}[t^n]\left(1-\frac{t}{\hat{\phi}}\right)^{-1}$$

$$= \frac{1}{\sqrt{5}}\binom{-1}{n}\left(-\frac{1}{\phi}\right)^n - \frac{1}{\sqrt{5}}\binom{-1}{n}\left(-\frac{1}{\hat{\phi}}\right)^n \quad \text{by Newton's rule}$$

$$= \frac{1}{\sqrt{5}}(-1)^n(-1)^n\left(\frac{1}{\phi}\right)^n - \frac{1}{\sqrt{5}}(-1)^n(-1)^n\left(\frac{1}{\hat{\phi}}\right)^n$$

$$\text{by }\binom{-1}{n} = (-1)^n\binom{1+n-1}{n} = (-1)^n$$

$$= \frac{1}{\sqrt{5}}\left(\frac{1}{\phi}\right)^n - \frac{1}{\sqrt{5}}\left(\frac{1}{\hat{\phi}}\right)^n$$

Checking with Maxima, in (%o770) we obtain the right Fibonacci numbers.

```
(%i757) eqFib:1-k-k^2=0;
        solutions:solve(eqFib,k);
        phi:ev(k,solutions[2]);
        phiHat:ev(k,solutions[1]);
        ratsimp(phi*phiHat);
        ratsimp(phi+phiHat);
        ratsimp(-phi*phiHat*(1-t/phi)*(1-t/phiHat));
        ratsimp((1+phi*t)*(1+phiHat*t));
        ABvalues:solve([A+B=0, (A/phiHat)+(B/phi)=-1],[A,B]);
        parFracAB(A,B):=(A/(1-t/phi))+(B/(1-t/phiHat));
        ratsimp(ev(parFracAB(A,B),ABvalues));
        coef(n,A,B):=A*(1/phi)^n + B*(1/phiHat)^n;
        coefCurry(n):=ratsimp(ev(coef(n,A,B),ABvalues))$
        map(coefCurry, makelist(n,n,0,10));
```

$$(\%o757) \; -k^2 - k + 1 = 0$$

$$(\%o758) \; [k = -\frac{\sqrt{5}+1}{2}, k = \frac{\sqrt{5}-1}{2}]$$

$$(\%o759) \; \frac{\sqrt{5}-1}{2}$$

$$(\%o760) \; -\frac{\sqrt{5}+1}{2}$$

$$(\%o761) \; -1$$

$$(\%o762) \; -1$$

$$(\%o763) \; -t^2 - t + 1$$

$$(\%o764) \; -t^2 - t + 1$$

$$(\%o765) \; [[A = \frac{1}{\sqrt{5}}, B = -\frac{1}{\sqrt{5}}]]$$

(%o766) $\mathrm{parFracAB}\,(A, B) := \dfrac{A}{1 - \frac{t}{\phi}} + \dfrac{B}{1 - \frac{t}{\mathrm{phiHat}}}$

(%o767) $-\dfrac{t}{t^2 + t - 1}$

(%o768) $\mathrm{coef}\,(n, A, B) := A\left(\dfrac{1}{\phi}\right)^n + B\left(\dfrac{1}{\mathrm{phiHat}}\right)^n$

(%o770) $[0, 1, 1, 2, 3, 5, 8, 13, 21, 34, 55]$

2

SEQUENTIAL SEARCH SIMULATION

We've repeated the simulation for the *sequential search* problem as did during the class. We want to study the mean of checks performed by the sequential search algorithm before the desired element is found in the given permutation. After we compare the mean of checks obtained by simulation with the theoretical mean value for the same input dimension.

The sequential search algorithm cited above is a very simple searching method: it consume a permutation of integer of a fixed dimension n and a target integer; produce true if the target belongs to the given permutation, else otherwise (in our case we augment the information returned with the number of checks needed in the run). The searching strategy consists of starting from the very left and moving one step to right every time the target is missed.

2.1 IMPLEMENTATION

We don't report here the description for the implementation of the sequential search algorithm because is very simple. Instead, we focus the explanation on the main simulation procedure.

The simulation function consume three parameters:

- numdimensions, the number of dimensions that we would like to test;

- interval, the multiplier to build the permutation vector to be used during the search;

- attempts, the number of application of the sequential search algorithm to a given permutation.

We use the number of dimensions numdimensions to build a vector dimensions such that both of the following hold:

$$\text{length}(\text{dimensions}) = \text{numdimensions}$$
$$\text{dimensions}[i] = i * \text{interval} \quad \forall i \in \{1, \ldots, \text{numdimensions}\}$$

For a given dimension $n \in \{\text{interval}, 2 * \text{interval}, \ldots, \text{numdimensions} * \text{interval}\}$, we sample a permutation of integers from $\{1, \ldots, n\}$ using the uniform distribution. For each permutation we apply the sequential search algorithm: the number of application is proportional to n (specifically $2n$) and after every application we record the number of checks needed to hit the target. When we run out all the iterations we compute the mean and the variance of the checks, storing them into auxiliary vectors.

2.2   RESULTS

In Table 3 we report a summary table for a simulation invoked with arguments $numdimensions = 50, interval = 20, attempts = 50$. Observing the table we can see that the simulation is quite correct, the theoretical values matches the simulated ones with little differences.

We've used the standardized means to plot them against the normal distribution in order to verify the Central Limit Theorem in Figure 1. The dotted curve represent the normal distribution, the blue one represent the sampling means distribution instead. Observing the plot we can say that the approximation is quite good, however the max dimension is $n = 50$, this allow us to apply the theorem but we could gain precision if we'll repeat our simulation for a greater value of n.
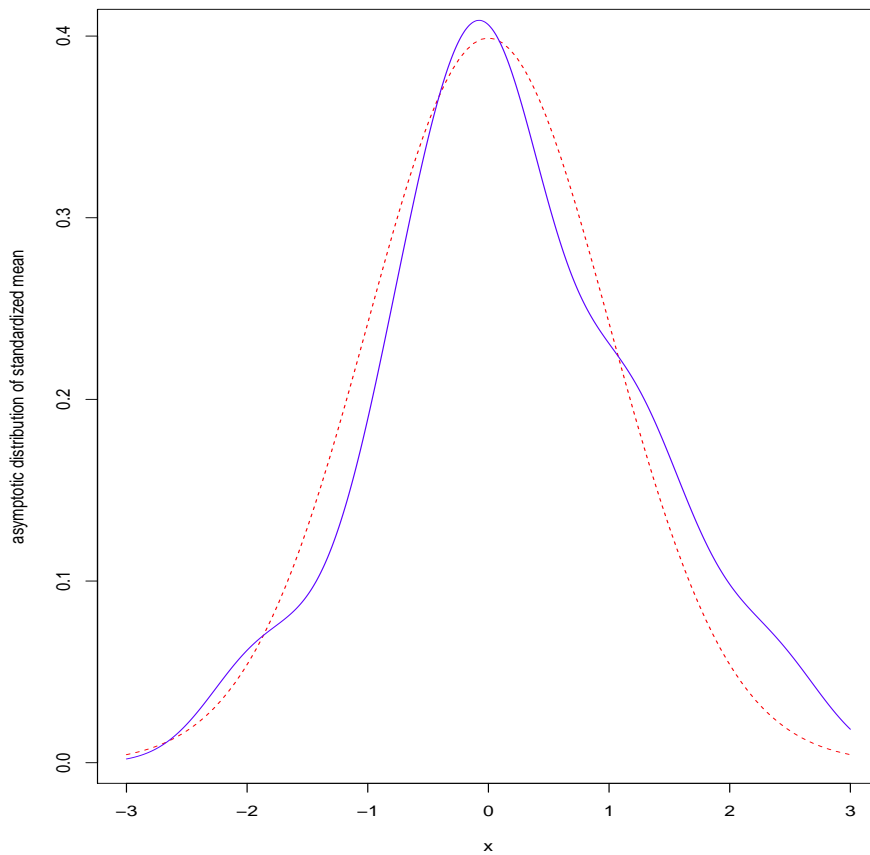

Figure 1: Plot of standardized means

Also we've build a second plot with a regression of the sampling means in Figure 2. The red line is drawn after computing the intercept and the coefficient using an hybrid model. Our attempt to explain the data consists of a mixture up to the second degree, as follow:

$$\text{dims} = \sum_{i}^{n} \text{dimensions}[i]$$

$$\text{means} = \sum_{i}^{n} \text{mean}[i]$$

$$\text{dim\_squares} = \sum_{i}^{n} \text{dimensions}[i]^2$$

$$\text{mean\_squares} = \sum_{i}^{n} \text{means}[i]^2$$

$$\text{dim\_mean} = \sum_{i}^{n} \text{means}[i] * \text{dimensions}[i]$$

$$\text{intercept} = \frac{\text{dim\_squares} * \text{means} - \text{dims} * \text{dim\_mean}}{n * \text{dim\_squares} - \text{dims}^2}$$

$$\text{coefficient} = \frac{n * \text{dim\_mean} - \text{dims} * \text{means}}{n * \text{dim\_squares} - \text{dims}^2}$$

where n is our numdimensions cited above, hence the red line is defined as $\text{means} = \text{coefficient} * \text{dimensions} + \text{intercept}$. Using an R interpreter we can see our numerical output:

```
coefficient = 0.49, intercept = 0.66
square of correlation index = 0.999
```

The regression is almost perfect, hence there exists a strong linear relation between means and distribution (from a statistical point of view we can conclude with the following observation: for an increment of the dimensions of one unit we get an increment of about a half unit on the number of checks).
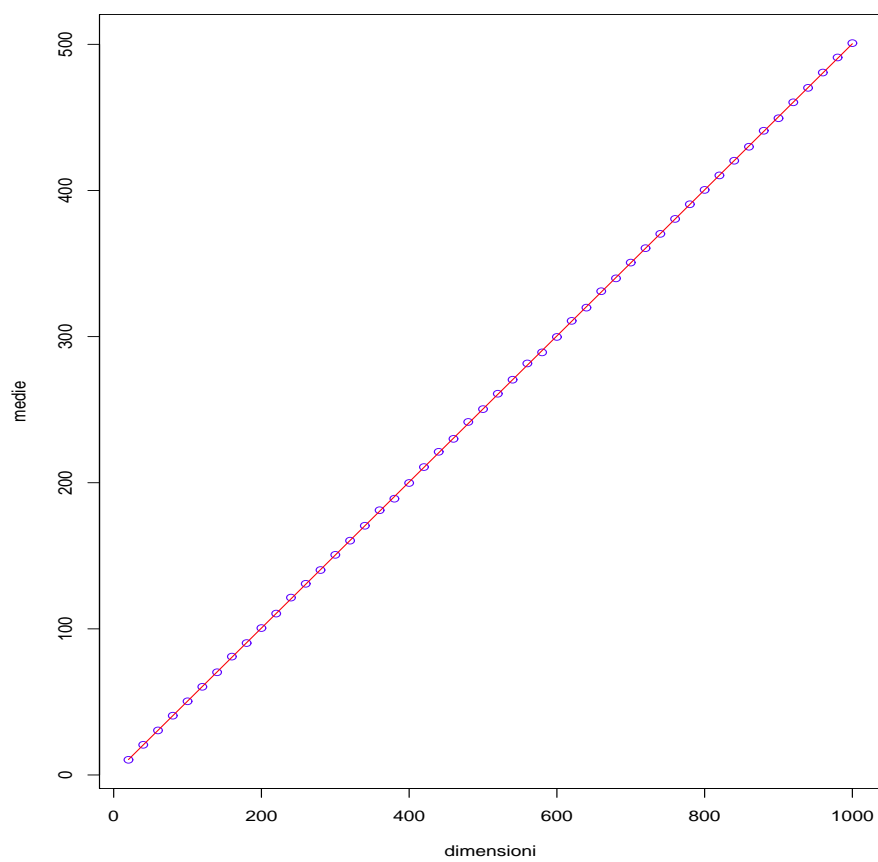
Figure 2: Plot of means regression

Table 3: Sequential search summary

	dimensions	theo means	means	theo vars	vars	theo var of vars	var of vars	stand. means	stand. vars
1	20.00	10.50	10.66	33.25	34.34	877.80	975.55	1.22	1.64
2	40.00	20.50	20.65	133.25	135.02	14177.80	14994.21	0.82	0.94
3	60.00	30.50	30.48	299.92	302.46	71900.02	74572.74	-0.10	0.74
4	80.00	40.50	41.09	533.25	531.28	227377.80	221994.03	2.28	-0.37
5	100.00	50.50	50.52	833.25	833.01	555277.80	553196.84	0.07	-0.03
6	120.00	60.50	60.13	1199.92	1189.70	1151600.02	1123368.13	-1.18	-1.04
7	140.00	70.50	70.42	1633.25	1621.51	2133677.80	2078575.00	-0.22	-0.95
8	160.00	80.50	79.80	2133.25	2155.97	3640177.80	3782358.53	-1.92	1.51
9	180.00	90.50	91.06	2699.92	2696.73	5831100.02	5837985.82	1.44	-0.18
10	200.00	100.50	100.26	3333.25	3315.86	8887777.80	8760281.48	-0.59	-0.82
...									
41	820.00	410.50	409.84	56033.25	56005.31	2511768877.80	2520823155.84	-0.80	-0.16
42	840.00	420.50	419.93	58799.92	59131.53	2765932400.02	2833053962.68	-0.68	1.83
43	860.00	430.50	429.49	61633.25	61801.47	3038913677.80	3061925029.95	-1.20	0.89
44	880.00	440.50	441.59	64533.25	64686.44	3331619377.80	3349812105.46	1.27	0.79
45	900.00	450.50	448.81	67499.92	67476.72	3644977500.02	3620068539.03	-1.95	-0.12
46	920.00	460.50	461.97	70533.25	70360.59	3979937377.80	3963315767.79	1.68	-0.83
47	940.00	470.50	470.75	73633.25	73791.33	4337469677.80	4381479940.27	0.28	0.74
48	960.00	480.50	481.49	76799.92	76684.82	4718566400.02	4704539498.17	1.11	-0.52
49	980.00	490.50	490.41	80033.25	80176.91	5124240877.80	5142289190.74	-0.10	0.63
50	1000.00	500.50	500.43	83333.25	83210.10	5555527777.80	5524755813.88	-0.08	-0.52

GENERATING BINARY TREES AT RANDOM

In this section we'll explain our work on generation of binary trees at random. We're interested to setup a simulation to study the means of the number of leaves in trees with n nodes and comparing the result of the simulation against theoretical results.

In order to do this we've implemented an algorithm to generate binary trees: it consume $n \in \mathbb{N}$ and produce a binary tree with n nodes.

We've repeated the application of that algorithm $k(>> n)$ times in order to check if the algorithm is an *uniform* binary tree generator, that is, if the generator would be perfect, each tree with n nodes should have $\frac{1}{\frac{1}{n+1}\binom{2n}{n}}$ probability to be generated.

The last point of our work is to check a theoretical result about the mean of number of leaves among binary trees with n nodes.

3.1 ATKINSON AND SACK ALGORITHM

The set of binary trees with n nodes is in one-to-one correspondence with many other sets of combinatorial objects, one of them is the set of well-formed bracket sequences with n pairs of brackets. The Atkinson and Sack algorithm focus on generating those sequences: in this section we'll explain in our words their solution (in chapter 4 is reported the original article).

3.1.1 *Definitions*

A sequence is a word $w \in \{(,)\}^*$. In the following we draw a sequence as a zigzag walk which starts from a base line (represented as - - -) and for each $(\in w$ we draw the character slash, for each $) \in w$ we draw the character backslash. For instance, the sequence of brackets ())()((()) is drawn as:

```
            /\
   /\      /  \
- - - - - - - - - - - - -
      \/\/
```
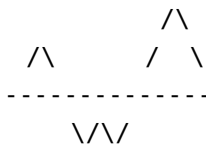
A word is *balanced* if it contains an equal number of (and), as in the sequence above. Graphically speaking, balanced words are sequences that start from the base line and, at right-most, return to the base line (in the middle it is possible

to cross zero, one or more time the base line).

A word w is *reducible* if $w = w_1 w_2$ where w_1 and w_2 are *balanced* and *non empty*. For instance the sequence reported above is reducible in four words while the following one isn't:

```
  /\
 /  \/\
/      \
----------
```
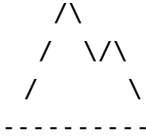
A *balanced* word w has *defect* i if w has i bracket pairs under the base line. Words with *defect* 0 are called *well-formed*. The previous word has *defect* 0, while the following one has *defect* 3:

```
 /\            /\
----------------
    \/\   /
       \/
```
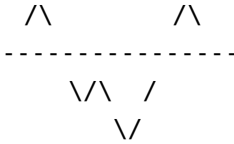
A word w^* is the complement of a word w if $\forall i \in \{0,\ldots,|w|\} : w[i] = (\leftrightarrow w^*[i] =)$, for instance the complement of the previous word is:

```
     /\
   /\/   \
--------------
 \/          \/
```
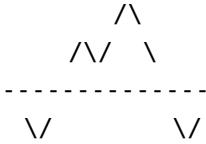
If a *balanced* word w is *non reducible* then either w or w^* is *well-formed*. We can see this result assuming w be one of the following words:

```
       ----------
  /\        \      /
 /  \/\      \  /\/
/      \      \/
----------
```
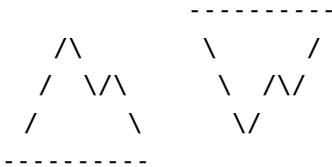
Both of them are *balanced* and *non reducible*: if w is the left-most one then w is *well-formed*, if w is the right-most one then w^* is *well-formed*.

3.1.2 *Splitting a* reducible *word*

How can we decide if a word w is *reducible* or not? We can do a cumulative summation on w where we consider each (as 1 and each) as -1. The following word has the cumulative sums $(0,-1,0,1,0,1,2,1,0,-1,0)$:
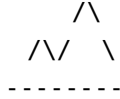
```
      /\
    /\/   \
- - - - - - - - - - - - -
   \/           \/
```
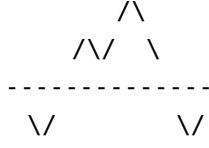
The following words have cumulative sums $(0, 1, 2, 3, 2, 1, 2, 1, 0)$ and $(0, -1, -2, -3, -2, -1, -2, -1, 0)$ respectively:

```
            - - - - - - - - - -
    /\          \       /
   /  \/\         \   /\/
  /        \         \/
- - - - - - - - -
```
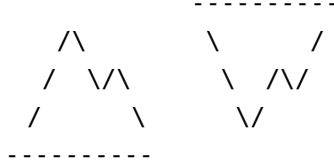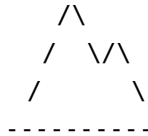
Let w be a word and $s = (0, s_1, \ldots, s_n)$ be the sequence of cumulative sums respect of w, where $n = |w|$. We split $w = uv$ such that $|u| > 0 \land |v| \geqslant 0$ using the following strategy:

$$i = \min\{k \in \{1, \ldots, n\} : s_k = 0\}$$
$$u = (w_1, \ldots, w_i) \quad v = (w_{i+1}, \ldots, w_n)$$

For instance, the following sequence as cumulative sums $(0, 1, 0, 1, 2, 1, 0)$, hence $i = 2$ so $u = ()$ and $v = (())$:

```
     /\
   /\/   \
- - - - - - - -
```

While the following one as cumulative sums $(0, 1, 2, 3, 2, 1, 2, 1, 0)$, hence $i = 8$ so $u = ((())())$ and v empty:

```
    /\
   /  \/\
  /        \
- - - - - - - - - -
```

3.1.3 *Generating a* balanced *word*

In order to generate a *balanced* word of length $2n$ (not necessary *well-formed*) we use the follow strategy:

1. generate a uniform sample L of length n from $\{1, \ldots, 2n\}$;

2. let $w = w_1 w_2 \ldots w_{2n}$ a word such that $w_i = (\leftrightarrow i \in L$;

3.1.4 *Transform a* balanced *word in a* well-formed *word*

Let $w = uv$ a *balanced* word obtained using the strategy described in the previous section. To get a *well-formed* word from w we define a function $\phi : \{(,)\}^* \to \{(,)\}^*$ inductively as follow (ϵ represent the empty string):

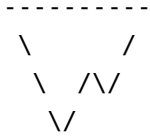$$\phi(\epsilon) = \epsilon$$
$$\phi(w) = u\phi(v) \quad \text{if } u \text{ is } \textit{well-formed}$$
$$\phi(w) = (\phi(v))t^* \quad \text{if } u =)t(\text{ is } \textit{not well-formed}$$

In order to recognize if a word w is *well-formed* it is sufficient to compute the cumulative sequence of sums s and check that $\forall i : s_i \geqslant 0$. Let's do some examples:
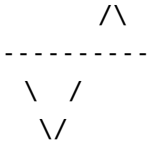
```
      /\
- - - - - - - - - -
  \   /
   \/
```
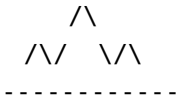
The previous sequence represents the word $w =))((()$. It is *balanced* so applying ϕ we get $\phi(w) = (\phi(v))()$ where $u =)t($, $t =)($ and $v = ()$. Now $\phi(v) = ()$ because $v = ()$ is *well-formed*, hence $\phi(w) = (())()$.

Another limit example:

```
- - - - - - - - - -
  \         /
   \   /\/
    \/
```

The previous sequence represents the word $w =)))(()((. $ It is *balanced* so applying ϕ we get $\phi(w) = (\phi(v))(())()$ where $u =)t($, $t =))(()($ and $v = \epsilon$. Now $\phi(v) = \epsilon$ hence $\phi(w) = ()(())()$, graphically:

```
     /\
  /\/   \/\
- - - - - - - - - -
```

3.1.5 *Complexity and Space*

Let w a *balanced* word and $T(n)$ the number of operation required. The decomposition $w = uv$ can be computed in $T(r)$ where $n \geqslant r = |u|$ by cumulative summations, remaining $T(n-r)$ operation for v. Hence $T(n) = O(r) + T(n-r) = O(n)$ a linear time.

For a space analysis, the algorithm use integers up to $2n$.

3.2   IMPLEMENTATION USING R

```r
generate.tree <- function(number_of_nodes){
  word_dimension <- 2 * number_of_nodes
  universe <- 1:word_dimension
  sample <- sample(universe, size=number_of_nodes)
  w = rep(0, word_dimension)
  for (i in 1:word_dimension) {
    w[i] <- ifelse(any(sample == i), 1, -1)
  }
  phi=phi(w)
  list(word=w, phi=phi, as_brackets = brackets_of_word(
    phi))
}

split.word <- function(w){
  if(length(w) == 0){
    return(list(u=c(), v=c()))
  }
  u_index_set <- 1:match(0, cumsum(w))
  list(u=w[u_index_set], v=w[-u_index_set])
}

phi <- function(w){
  if(length(w) == 0){
    return(w)
  }
  split <- split.word(w)
  if(all(cumsum(split$u) > -1)){
    return (c(split$u, phi(split$v)))
  }
  else{
    t = split$u[-c(1, length(split$u))]
    return (c(1, phi(split$v), -1, -t))
  }
}
```

3.3   CHECKING RANDOMNESS

In order to establish if the algorithm is an *uniform random* generator, we perform a χ^2 test on the generated trees. We've implemented that process in R and we're going to comment the results obtained.

In this section we often report a table with the following structure: we put on columns the dimensions 4,5,6,8,10 which represent the number of nodes inside trees under study, while we put on rows the dimensions 1000, 2000, 5000, 10000, 20000, 50000 which represent the quantities of generated trees using our generator. We begin with reporting in Table **??** the theoretical number of hits

	4	5	6	8
1000	71.43	23.81	7.58	0.70
2000	142.86	47.62	15.15	1.40
5000	357.14	119.05	37.88	3.50
10000	714.29	238.10	75.76	6.99
20000	1428.57	476.19	151.52	13.99
50000	3571.43	1190.48	378.79	34.97

Table 4: Theoretical uniform distribution of hits among tree structures

that every category (in our case a category is a tree structure) should contain; more formally, let C be the table in

$$C(row(i), col(j)) = \frac{i}{\frac{1}{j+1}\binom{2j}{j}}$$

with $i \in \{1000, 2000, 5000, 10000, 20000, 50000\}$ and $j \in \{4, 5, 6, 8\}$.

We note that for column labeled with "8" there are a sufficient number of hits only from simulation with 10000 trees in order to perform a correct χ^2 test (at least 5 hits per category are required by the test, so we'll discard the first three simulations for column "8").

In Table 5 we report two kind of informations: on the left of the doubled line there are the observed χ^2 values, while on the right there are the corresponding p-values. In order to be considered an *uniform random* generator, each p-value v should be such that $.1 \leqslant v \leqslant .9$, our values respect these bounds (the 1s for the first three rows for column "8" are meaningless). Those values have to be interpreted by reasoning on the χ^2 curve reported in Figure 3, relative to a simulation of trees with 4 nodes (therefore 13 freedom degrees). The χ^2

Figure 3: χ^2 curve with 13 freedom degrees of trees with 4 nodes

	4	5	6	8	4	5	6	8
1000	10.38	40.68	134.72	236.17	0.66	0.48	0.37	1.00
2000	13.75	50.57	123.62	595.92	0.39	0.15	0.66	1.00
5000	15.12	48.82	113.73	1116.00	0.30	0.19	0.86	1.00
10000	9.93	43.22	138.26	1435.09	0.70	0.38	0.32	0.42
20000	13.79	40.30	125.49	1413.96	0.39	0.50	0.62	0.61
50000	12.33	51.12	120.77	1466.10	0.50	0.13	0.73	0.24

Table 5: χ^2 observed statistics (on the left) and p-values (on the right)

values for column "4" fall around the value 13, which is the freedom degrees characterizing the curve and the point with maximum probability: they are quite close to it, neither to close to 0 (which means that the χ^2 test doesn't recognize differences from an uniform distribution) not to far on the right (which means that the χ^2 test recognize a counterexample which discriminate from the uniform distribution). On the other hand, considering the corresponding p-values, they measure the area under the curve from the observed value to $+\infty$, for instance let $\chi(x)$ be the represented curve, to the observed statistic 9.93 there correspond a p-value of .7, more formally $.7 = \int_{9.93}^{+\infty} \chi(x)dx$ (this is coherent with the other interpretations, get for instance a bigger observed statistic 15.12 for which correspond a smaller area of $.3 = \int_{15.12}^{+\infty} \chi(x)dx$).

3.4 MEANS OF LEAVES AND HEIGHTS

In this section we study the means of the number of leaves and of heights of all different binary trees with n nodes.

From theory we know that the mean of leaves of trees with n nodes is:

$$\frac{n\,(n+1)}{2\,(2\,n-1)}$$

Using Maxima to have the exact means for some nodes dimensions:

```
(%i178) fpprintprec:4$
        leaves(n):=(n*(n+1))/(2*(2*n-1));
        combineResult(n):=[n,leaves(n)]$
        map(combineResult, makelist(n,n,1,10)),numer;
```

(%o179) $\text{leaves}\,(n) := \frac{n\,(n+1)}{2\,(2\,n-1)}$

(%o181) $[[1,1],[2,1],[3,1.2],[4,1.429],[5,1.667],[6,1.909],[7,2.154],[8,2.4],[9,2.647],[10,2.895]]$

In Table **??** we report the results obtained with our implementation in R: the second column report the number of trees generated with the correspondent number of nodes.

Table 6: Theoretical and empirical means for leaves and height

nodes	dimensions	theo.mean.leaves	theo.mean.height	emp.mean.height	emp.mean.leaves	emp.mean.height
3	100	1.20	2.80	1.16	2.84
4	200	1.43	3.57	1.35	3.65
5	300	1.67	4.24	1.70	4.23
6	1000	1.91	4.88	1.91	4.87
7	10000	2.15	5.47	2.15	5.46
8	10000	2.40	6.03	2.41	6.02
9	50000	2.65	6.56	2.65	6.56
10	100000	2.89	7.07	2.90	7.07

3.5 STANDARDIZED MEANS AND ASYMPTOTIC DISTRIBUTIONS

In this section we study the asymptotic distribution of the means for the leaves and heights of trees with n nodes. We use the result of the *Central Limit Theorem* to check if the two means under study behave like a normal distribution in repeated sampling.

Each of the following plots are obtained repeating 1000 times the generation of 500 trees each with 5 nodes. In each plot the red dotted curve is the normal distribution (out "target"), while the blue continue curve is the inferred distribution.

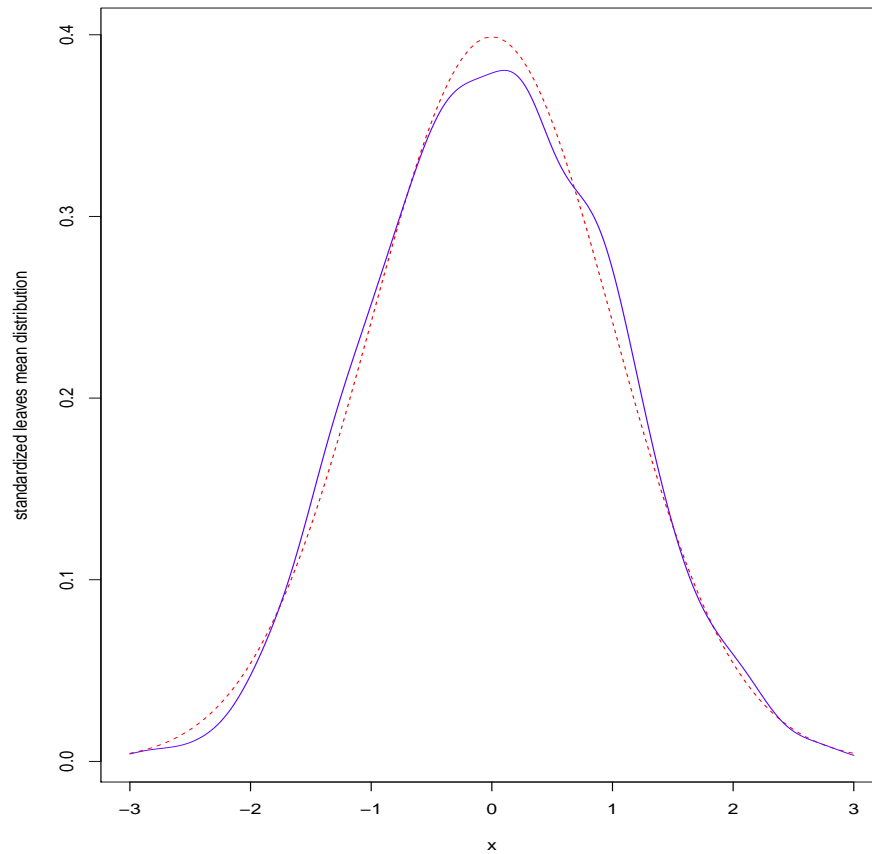In Figure 4 we report the standardized mean of leaves.

Figure 4: Standardized leaves mean distribution

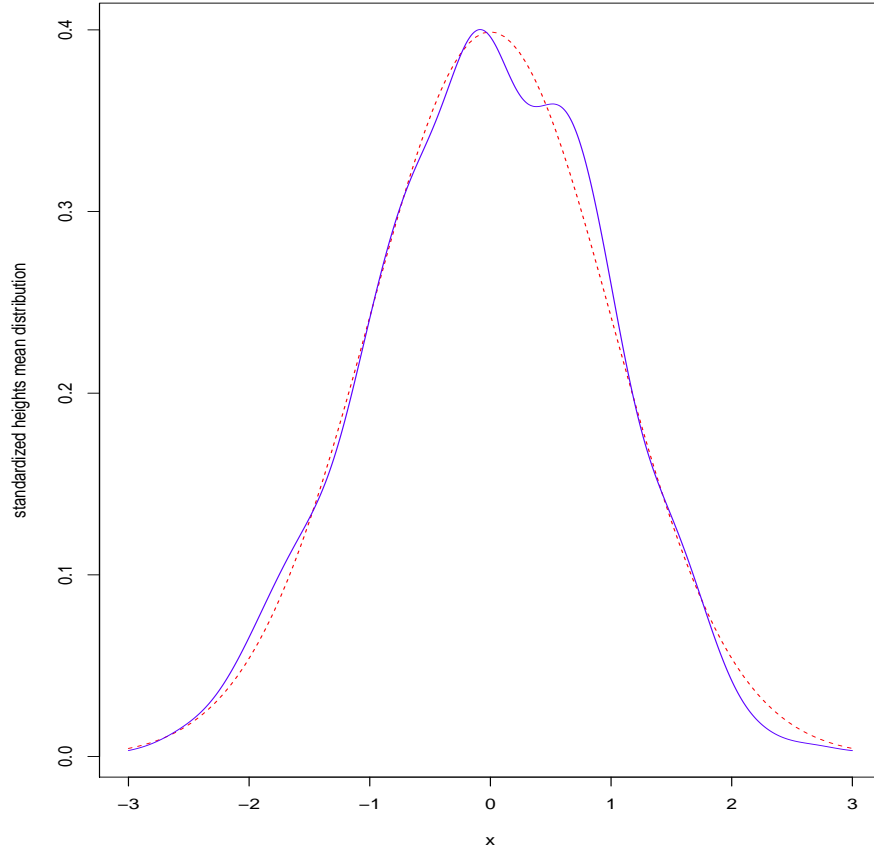In Figure 5 we report the standardized mean of leaves.

Figure 5: Standardized height mean distribution

3.6   DRAWING TREES

Just for fun, our implementation draw all different trees with n nodes (this implementation is done in OCaml). In Figure 6 we report the image with all binary trees with 4 nodes.

### 3.6.1   *From bracket representations to binary trees*

In this section we report the algorithm used to build the OCaml tree objects from bracket representations. The algorithm is recursive: given an input bracket representation, the algorithm breaks the representation in smaller representations, which are valid bracket representations themselves, recurs on each of them and

then collects the results to returns the final answer. In order to describe it more precisely we distinguish two cases, let d be a bracket representation:

- the base of the recursion is for $d = \epsilon$, the empty bracket representation. In this case the tree for d is an empty binary tree;

- d cannot be broken into two shorter bracket representations. It follows that, since d is however a valid bracket representation (ie. d is balanced as we've defined in previous sections), it must have the form $d = (M)$, where M is a shorter bracket representation (can be the case of $M = \epsilon$ empty, this is necessary to reach the base of recursion). In this case we recur on M combining the result using the following strategy:

  > let m be the tree corresponding to bracket representation M: build an empty tree t for bracket representation d attaching the root of tree m to the *right* leaf of tree t.
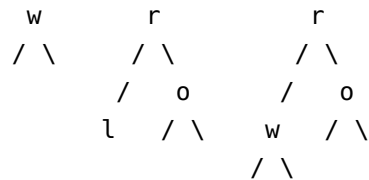
  For instance assume $d = (())$, hence $M = ()$, so we build the following tree t for bracket representation d:

  ```
      t
     / \
    /   m
   /   / \
  ```
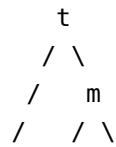
- d can be broken into two shorter bracket representations, that is $d = d_1 d_2$. Using the toy $d = ()(())$ we've $d_1 = ()$ and $d_2 = (())$. In this case we recur on both d_1, d_2 combining the results using the following strategy:

  > let w be the tree corresponding to bracket representation d_1 and let r be the tree corresponding to bracket representation d_2: build a tree t for bracket representation d attaching the root of tree t_1 to the *left-most* leaf of tree t_2.

  For example let $d = ()(())$ we get the following trees by recursive application on $d_1 = (), d_2 = (())$ respectively:

  ```
    w         r             r
   / \       / \           / \
            /   o         /   o
           l   / \       w   / \
                            / \
  ```

  We attach the tree rooted at w to the *left-most* leaf l of tree rooted at r to get the right-most tree rooted at t for d in the picture above.

  If a bracket representation d can be broken in more shorter bracket representations $d_1, \ldots, d_n$ then we break d using the *right-most* breakpoint, that is we recur on $d' = d_1, \ldots, d_{n-1}$ and d_n and proceed as explained above.
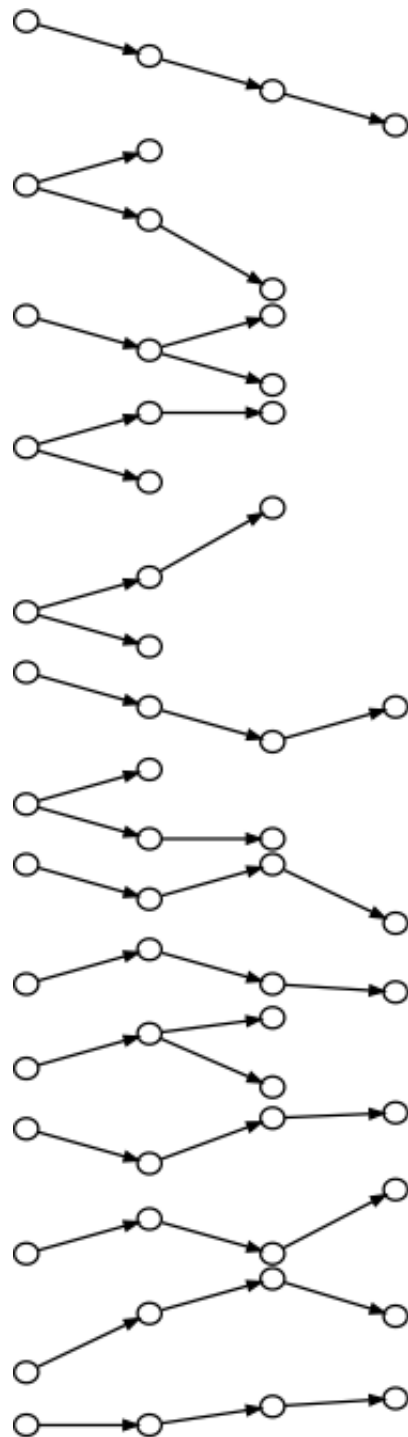
Figure 6: Binary trees with four nodes

4

APPENDIX

4.1 PROJECT DESCRIPTION

In this section we explain some details "under the hood" of the implementation of this project.

We've used the statistical environment R to use its very efficient functions to perform uniform sampling (required by the Atkinson and Sack algorithm) and χ^2 test in order to check the randomness of the generator. To produce the images with all different trees with n nodes, we've used the OCaml language to use its functional power.

The two programs works together: the entry point for the user are the R functions (which can be loaded in any R interpreter). During their computation, they invoke an OCaml program giving their argument in form a file written on the file system. The OCaml program parse the input file, compute the representation of trees and write them back in another file. The control return to R functions, which invoke the *graphviz* programs to build a representation of trees in image form.

4.2 MAXIMA INTERACTION

We've tested some the theoretical results with the engine Maxima, an open source tool similar to Maple, saving all the stuff in the file `maxima-worksheet.wxm`. It is possible to run Maxima, load the file and experiment with it. It has been useful because has a feature to export the results directly in LATEXready form, for instance:

(%o179) $\text{leaves}\,(n) := \dfrac{n\,(n+1)}{2\,(2\,n-1)}$

(%o181) $[[1,1],[2,1],[3,1.2],[4,1.429],[5,1.667],[6,1.909],[7,2.154],[8,2.4],[9,2.647]]$

4.3 PROJECT HOSTING

This entire work, including the source code for the document you're reading, the R functions and the OCaml implementations are completely available for download at:

https://github.com/massimo-nocentini/pacc

It is possible to clone the entire Git repository typing the following command in

a Unix terminal:

`git clone git@github.com:massimo-nocentini/pacc.git`

A compiled version of this document is available at:

https://github.com/massimo-nocentini/pacc/blob/master/compiled-versions/Elaborato.pdf

4.4   ORIGINAL ARTICLE ABOUT RANDOM BINARY TREES GENERATION

In the last pages of this document we report the original article about the algorithm for random trees generation written by Atkinson and Sack (this isn't our work, we report it here because it is difficult to find and it is just for getting the reading experience of this document more complete).

Generating binary trees at random

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Abstract

Atkinson, M.D. and J.-R. Sack, Generating binary trees at random, Information Processing Letters 41 (1992) 21–23.

We give a new constructive proof of the Chung–Feller theorem. Our proof provides a new and simple linear-time algorithm for generating random binary trees on n nodes; the algorithm uses integers no larger than $2n$.

Keywords: Analysis of algorithms, binary trees, bracket sequences, data structures

1. Introduction

Methods for generating binary trees on n nodes have been considered by several authors (see [4,8] and [6] also for additional references). In most cases the focus has been on generating all binary trees in some order or on ranking and unranking them. The number of binary trees on n nodes is the Catalan number $\binom{2n}{n}/(n+1)$ which is exponential in n ($\approx 4^n$) and so these computations cannot be carried out very easily unless n is small. Unranking algorithms allow binary trees to be generated uniformly at random and this is often more useful than being able to list all the possible trees. Unfortunately, numbers which are exponential in n enter these calculations and this makes them impracticable unless n is small.

The problem of generating binary trees uniformly at random without introducing exponentially large numbers was overcome by Arnold and Sleep [1] and Martin and Orr [6]; they gave linear time algorithms which used integers of size $O(n)$ for generating a random binary tree. We shall present a new and simpler solution having the same advantages. Our solution is based on a constructive version of the Chung–Feller theo-

rem on coin-tossing, which has not previously been applied to this area. Our treatment also provides a new proof of the Chung–Feller theorem.

The set of binary trees on n nodes is well known to be in one-to-one correspondence with many other sets of combinatorial objects including rooted (ordered) trees with n branches, triangulations of a convex $(n+2)$-gon, lattice paths from $(0, 0)$ to (n, n) which do not cross the diagonal, and well-formed bracket sequences with n pairs of brackets. The one-to-one correspondences are explicit and efficient to compute in linear time; thus a uniform random generator for binary trees gives rise to a uniform random generator for all these other objects, and vice versa. We shall focus on generating well-formed sequences of brackets.

2. Terminology

We begin with some terminology. It is convenient to denote the left and right bracket symbols by λ and ρ respectively. Thus a bracket sequence (well formed or not) corresponds to a word over

Fig. 1. The zigzag diagram corresponding to the word $\lambda\rho\rho\lambda\rho\lambda\lambda\lambda\rho\rho$.

the alphabet $\{\lambda, \rho\}$. A word such as $\lambda\rho\rho\lambda\rho\lambda\lambda\lambda\rho\rho$ may be pictured as a zigzag diagram drawn from some base line where each upward edge represents λ and each downward edge represents ρ (see Fig. 1).

A word is said to be *balanced* if it contains equal numbers of λ's and ρ's. Balanced words are precisely those whose diagram returns to the base line. A balanced word w is said to be *reducible* if it may be written $w = w_1 w_2$ with w_1, w_2 each balanced and nonempty, otherwise w is *irreducible*.

A balanced word is defined to have *defect* i if its diagram has precisely $2i$ links below its base line. Defect 0 words are called *well-formed* and corresponds to well-nested bracket sequences. Observe that the defect of a word is easily found by a summation: we scan the word of length k from left to right regarding each λ as $+1$, each ρ as -1, and computing the partial sums 0, $s_1, \ldots, s_k$; the final sum is zero and the number of negative interim sums s_j at odd indices j is the defect. We call this calculation *partial summation*.

For any word w let w^* denote the result of replacing all occurrences of λ by ρ and ρ by λ. The following two results are immediate consequences of these definitions.

Lemma 1. *If a balanced word w is irreducible, then one of w and w^* is well-formed; in fact, $w = \lambda u \rho$ where u is well-formed, or $w = \rho u \lambda$ where u^* is well-formed. If a balanced word w is well-formed, then $\lambda w \rho$ is irreducible.*

Lemma 2. *A balanced word w has a unique factorisation as $w = w_1 w_2 \cdots w_k$, where each w_i is irreducible. If w is well-formed, so is each w_i.*

22

3. The algorithm

Let B_n denote the set of $\binom{2n}{n}$ balanced words of length $2n$, and let B_{ni} denote the subset of balanced words of defect i. Clearly B_n is the disjoint union of $B_{n0}, B_{n1}, \ldots, B_{nn}$. The Chung–Feller Theorem, see [2, Theorem 2A] and [3, p.94], states that these subsets all have the same size. The central idea of our algorithm is to use a new constructive proof of this theorem which depends on explicit 1–1 correspondences between these sets.

Our algorithm has the following form:

Algorithm RANDOM BRACKET SEQUENCE

Input: An integer n.
Output: A well-formed word of length $2n$ over the alphabet $\{\lambda, \rho\}$.

1. Generate a uniformly random combination L of n integers from $\{1, 2, \ldots, 2n\}$
2. Define a random member $x = (x_1 x_2 \cdots x_{2n})$ of B_{2n} by the rule $x_i = \lambda$ if $i \in L$, $x_i = \rho$ if $i \notin L$.
3. Return the well-formed member of B_{2n} to which x corresponds.

Steps 1 and 2 of the algorithm are straightforward. To generate a combination of n integers from $\{1, 2, \ldots, 2n\}$ uniformly at random, and hence a member of B_n, we may use the technique described, for example, in [5, p.137] (see also [7, pp. 189–198] for more discussion on this topic). This technique takes only linear time and uses only integers less than $2n$.

To implement step 3 we need to define some suitable correspondences. We denote by $|w|$ the length of w and by card(S) the cardinality of a set S. We now define a map $\Phi_n : B_n \to B_{n0}$. The definition is inductive. For $n = 0$, we define Φ_0 in the only way possible: it maps the empty string to the empty string. For $n > 0$ and $w \in B_n$ we begin by expressing w as $w = uv$, where u is irreducible, $|u| = r > 0$, $|v| = s \geqslant 0$; then we define Φ_n by the rules

$$\Phi_n(w) = u\Phi_s(v) \qquad \text{if } u \text{ is well-formed,}$$
$$\Phi_n(w) = \lambda\Phi_s(v)\rho t^* \qquad \text{if } u = \rho t \lambda \text{ is not well-formed.}$$

Theorem 3. Φ_n is $n+1$ to 1 onto B_{n0} and is bijective on each B_{ni}.

Corollary 4. (Chung and Feller) $\text{card}(B_{ni}) = \text{card}(B_{n0})$.

Corollary 5. If w is a random variable distributed uniformly in B_n, then $\Phi_n(w)$ is distributed uniformly in B_{n0}.

Proof. It is sufficient to show that Φ_n is a bijection $B_{ni} \to B_{n0}$ for each i. Suppose that $w_1, w_2 \in B_{ni}$ and that they have the same image under Φ_n. For $k = 1,2$ put $w_k = u_k v_k$, where u_k irreducible, and let $|u_k| = r_k$, $|v_k| = s_k$. There are four possibilities:

(1) u_1, u_2 are each well-formed. Then $v_1 \in B_{s_1,i}$, $v_2 \in B_{s_2,i}$, and $u_1 \Phi_s(v_1) = \Phi_n(w_1) = \Phi_n(w_2) = u_2 \Phi_{s_2}(v_2)$. By Lemma 2, $u_1 = u_2$, $\Phi_{s_1}(v_1) = \Phi_{s_2}(v_2)$ and so $v_1 = v_2$ by induction.

(2) Neither of u_1 and u_2 are well-formed, say $u_1 = \rho t_1 \lambda$ and $u_2 = \rho t_2 \lambda$. Then $v_1 \in B_{s_1,i-r_1}$, $v_2 \in B_{s_2,i-r_2}$, and $\lambda \Phi_{s_1}(v_1)\rho t_1^* = \lambda \Phi_{s_2}(v_2)\rho t_2^*$. The leading subwords $\lambda \Phi_s(v_1)\rho t_1^*$ and $\lambda \Phi_s(v_2)\rho t_2^*$ are irreducible, therefore equal. Therefore, by induction, $v_1 = v_2$ and $t_1 = t_2$, so $u_1 = u_2$.

(3) u_1 is well-formed and u_2 is not well-formed, say $u_2 = \rho t_2 \lambda$. Then $v_1 \in B_{s_1,i}$, $v_2 \in B_{s_2,i-r_2}$, and $u_1 \Phi_s(v_1) = \lambda \Phi_{s_2}(v_2)\rho t_2^*$. By Lemma 2 again, $u_1 = \lambda \Phi_{s_2}(v_2)\rho$ and, taking lengths, $r_1 = s_2 + 2$ and $s_1 = r_2 - 2$. But $r_2 \leqslant i$ since $v_2 \in B_{s_2,i-r_2}$ and $i \leqslant s_1$ since $v_1 \in B_{s_1,i}$ from which it follows that $i \leqslant s_1 \leqslant r_2 - 2 < r_2 \leqslant i$, a contradiction.

(4) u_1 is not well-formed and u_2 is well-formed. This case is impossible for the same reasons as case 3.

This proves that Φ_n is one-to-one on B_{ni}. To prove that it maps B_{ni} onto B_{n0} it is enough to show that these sets have the same size. But we have seen that $\text{card}(B_{ni}) \leqslant \text{card}(B_{n0})$ for each i and if any of these inequalities were strict we would have the contradiction

$$\binom{2n}{n} = \text{card}(B_n) = \sum_{i=0}^{n} \text{card}(B_{ni})$$

$$< (n+1)\text{card}(B_{n0}) = \binom{2n}{n}. \qquad \square$$

Lemma 6. If $w \in B_n$, $\Phi_n(w)$ can be determined in $O(n)$ time.

Proof. We follow the inductive definition of $\Phi_n(w)$. Let $T(n)$ be the total number of operations required. The decomposition $w = uv$, where u is irreducible, can be found in $O(r)$ steps, where $r = |u|$ by partial summation; the first partial sum equal to zero defines u. Then $\Phi_{n-r}(v)$ must be calculated and so we obtain the recurrence $T(n) = O(r) + T(n-r) = O(n)$. $\square$

Note that the computation of $\Phi_n(w)$ requires no integers larger than $2n$. Thus step 3 of our algorithm can be implemented in linear time, with integers of size $O(n)$, by applying the function Φ_n.

The above discussion provides the proof of the following theorem:

Theorem 7. Algorithm RANDOM BRACKET SEQUENCE is an unbiased random binary tree generator. It is executes in linear time and uses integers of size $O(n)$.

Acknowledgment

References

[1] D.B. Arnold and M.R. Sleep, Uniform random number generation of n balanced parenthesis strings, ACM Trans. Programming Languages Systems 2 (1980) 122–128.

[2] K.L. Chung and W. Feller, Fluctuations in coin-tossing, Proc. Nat. Acad. Sci. U.S.A. 35 (1949) 605–608.

[3] W. Feller, An Introduction to Probability Theory and its Applications, Vol. 1 (Wiley, New York, 3rd ed., 1968).

[4] G.D. Knott, A numbering system for binary trees, Comm. ACM 20 (1977) 113–115.

[5] D.E. Knuth, Semi-numerical Algorithms, The Art of Computer Programming, Vol. 2 (Addision-Wesley, Reading, MA, 2nd ed., 1981).

[6] H.W. Martin and B.J. Orr, A random binary tree generator, in: Computing Trends in the 1990's, ACM Seventeenth Computer Science Conf., Louisville, KY (ACM, New York, 1989) 33–38.

[7] E.M. Reingold, J. Nievergelt and N. Deo, Combinatorial Algorithms: Theory and Practice (Prentice Hall, Englewood Cliffs, NJ, 1977).

[8] M. Solomon and R.A. Finkel, A note on enumerating binary trees, J. ACM 27 (1980) 3–5.