# 1179: Probability Lecture 25 — Concentration Inequalities

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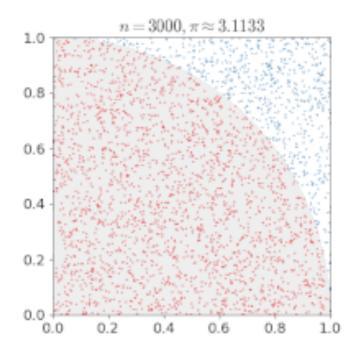
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#### Monte-Carlo Method?



- What is "Monte-Carlo method"?
- "... computational algorithms that rely on repeated random sampling to obtain numerical results... use randomness to solve problems that might be deterministic in principle." (by Wikipedia)
- Math principle behind Monte-Carlo?
- Use Monte-Carlo to estimate  $\pi$





(Rafael Nadal: 11 titles at Monte Carlo Masters)

#### This Lecture

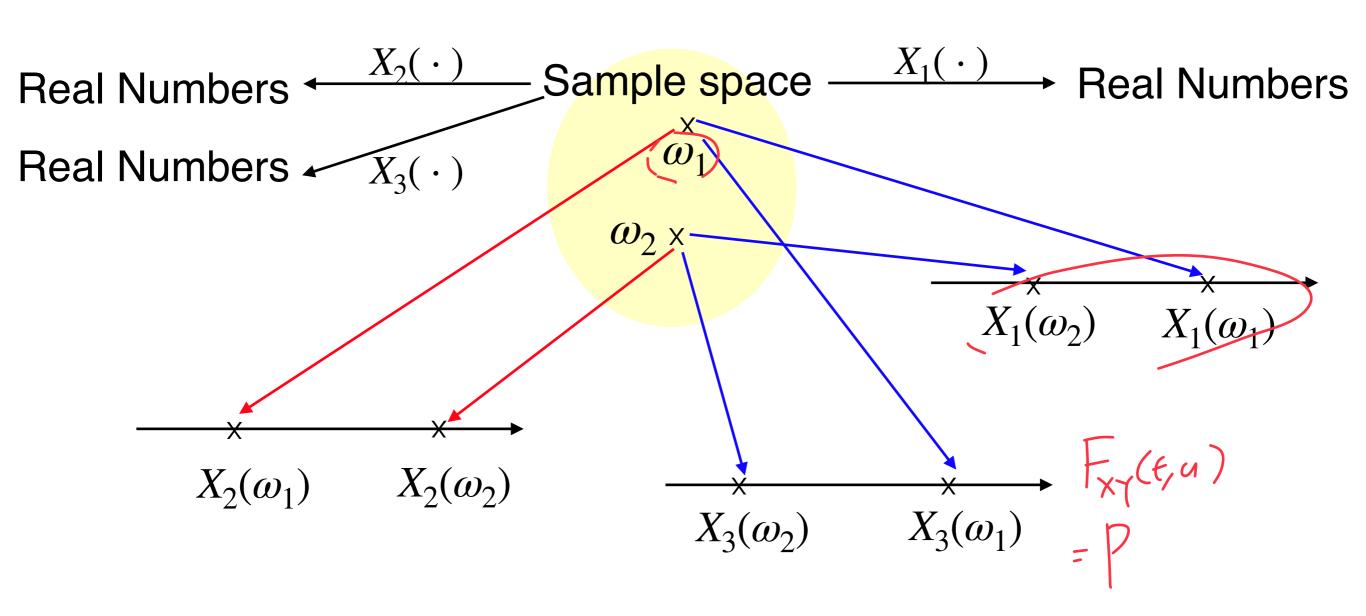
1. Multivariate Random Variables

2. Concentration Inequalities

3. Weak Law of Large Numbers (WLLN)

Reading material: Chapter 9.1 and 11.3-11.4

## A Primer on Multiple Random Variables



• Could we study the CDF regarding  $X_1, X_2$ , and  $X_3$ ?

#### From Bivariate To Multivariate

- Key Idea: "Bivariate" definitions and properties can be directly extended to the "multivariate" cases
- For example:
  - 1. Joint CDF / PMF / PDF
  - 2. Expected value
  - 3. Marginal CDF / PMF / PDF
  - 4. Independence

#### Joint CDF of Multivariate R.V.s

Joint CDF of 2 Random Variables: Let X and Y be two random variables defined on the same sample space  $\Omega$ . The joint CDF  $F_{XY}(t,u)$  is defined as

$$F_{XY}(t, u) = P(X \le t, Y \le u), \ \forall t, u \in \mathbb{R}$$

Joint CDF of n Random Variables: Let  $X_1, \dots, X_n$  be random variables defined on the same sample space  $\Omega$ . The joint CDF  $F(x_1, x_2, \dots, x_n)$  is defined as

$$F(x_1, x_2, \dots, x_n) = P(X_1 \le x_1, X_2 \le x_2, \dots, X_n \le x_n), \ \forall x_i \in \mathbb{R}$$

#### Joint PDF Multivariate R.V.s

Joint PDF of 2 Random Variables: Let X and Y be two continuous random variables. Then,  $f_{XY}(x,y)$  is the joint PDF of X and Y if for every subset B of  $\mathbb{R}^2$ , we have

$$P((X,Y) \in B) = \iint_{B} f_{XY}(x,y) dxdy$$

**Joint PDF of** n **Random Variables**: Let  $X_1, \dots, X_n$  be n continuous random variables. Then,  $f(x_1, x_2, \dots, x_n)$  is the joint PDF of  $X_1, \dots, X_n$  if for every subset B of  $\mathbb{R}^n$ , we have

$$P((X_1, X_2, \dots, X_n) \in B) = \int \dots \int_B f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n$$

## **Expected Value**

#### **Expected Value of a Function of 2 Continuous RVs:**

Let X, Y be 2 continuous random variables with joint PDF  $f_{XY}(x, y)$ . Let  $g(\cdot, \cdot)$  be a function from  $\mathbb{R}^2 \to \mathbb{R}$ The expected value of g(X, Y) is

$$E[g(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f_{XY}(x,y) dx dy$$

#### **Expected Value of a Function of** *n* **Continuous RVs:**

Let  $X_1, \dots, X_n$  be n continuous random variables with joint PDF  $f(x_1, x_2, \dots, x_n)$ . Let g be a function from  $\mathbb{R}^n \to \mathbb{R}$ . The expected value of  $g(X_1, X_2, \dots, X_n)$  is

$$E[g(X_1, X_2, \dots, X_n)] = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} \underline{g(x_1, \dots, x_n)} \underline{f(x_1, \dots, x_n)} dx_1 dx_2 \dots dx_n$$

# Concentration Inequalities

# Motivating Example: Tossing Moon Blocks



- 3 possible outcomes: Yes / No / Laughing
- - Each toss is independent from other tosses
- Question: Suppose p is unknown
  - How to learn p?
  - Could we learn anything useful after n experiments?

Suppose the true p = 0.5

Loo times

P(see ) Yes

experiments?

Concentration Inequalities

# Markov's Inequality

# For all west, X(w) >, 0

Markov's Inequality: Let X be a nonnegative random

variable. Then, for any t > 0,

continuous 
$$\pm \infty$$

Visualization:

$$Y(\omega) = \begin{cases} 0, & \text{if } X(\omega) < t \\ t & \text{if } X(\omega) > t \end{cases} PMF of Y$$

$$P(X \ge t) \le E[X]$$

$$= 0. P(X = 0)$$

$$+ t.P(Y=t)$$

$$= t.P(X>t)$$

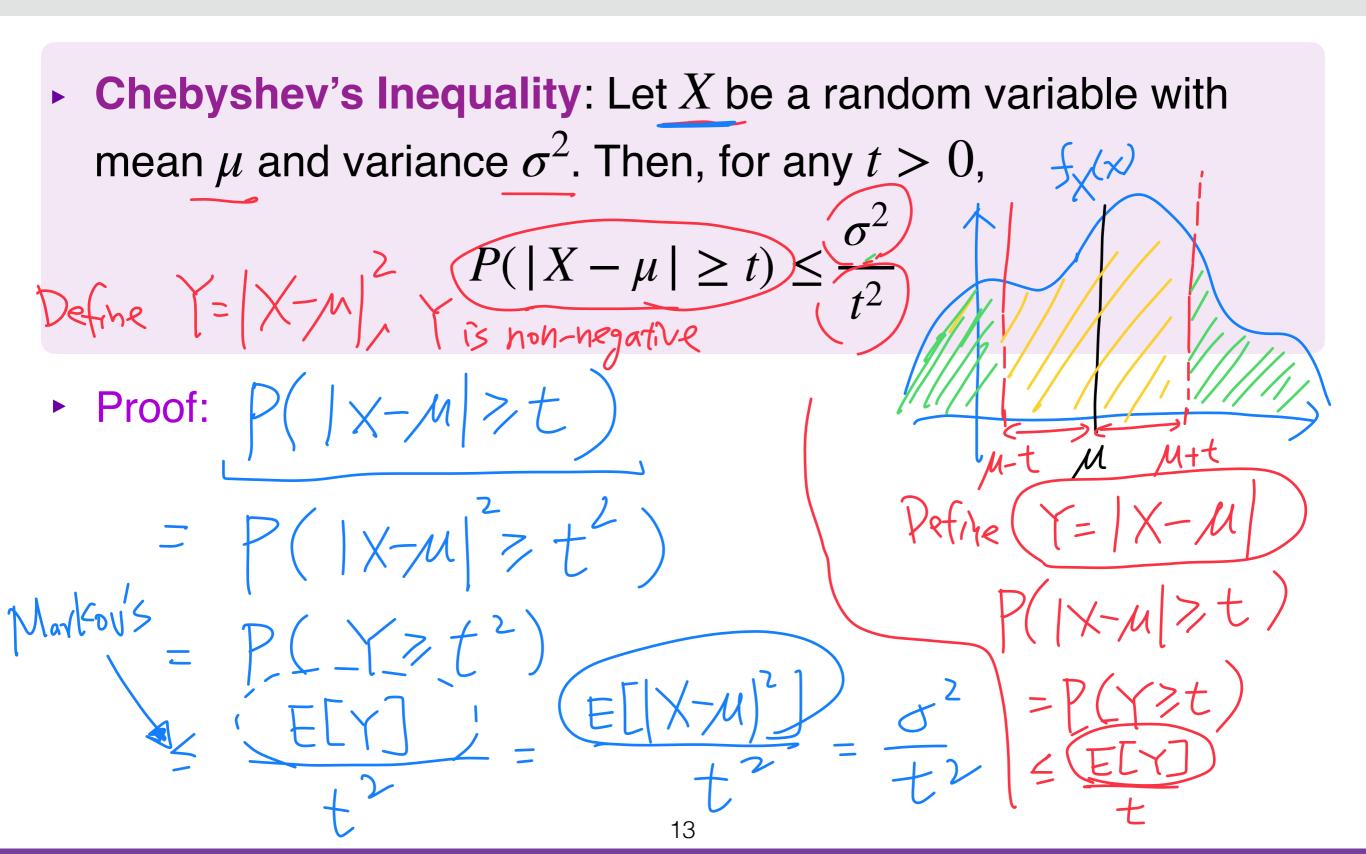
# Proof of Markov's Inequality

• Markov's Inequality: Let X be a nonnegative random variable. Then, for any t>0,

$$P(X \ge t) \le \frac{E[X]}{t}$$

Proof: (Please refer to the previous page)

# Chebyshev's Inequality

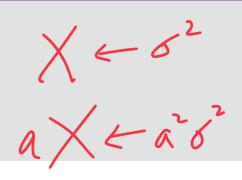


# Quick Review: Mean and Variance of Sum of Independent Random Variables

- Example: Each  $X_i$  has mean  $\mu_i$  and variance  $\sigma_i^2$ 
  - $X_1, X_2, \dots, X_n$  are assumed to be independent
  - Question 1:  $E[X_1 + X_2 + \dots + X_n] = \mathcal{M}_1 + \mathcal{M}_2 + \dots + \mathcal{M}_n$  (  $\sum_{i \neq j} \mathcal{M}_i$ )
  - Question 2:  $E\left[\frac{1}{n}(X_1 + X_2 + \dots + X_n)\right] = \frac{1}{n}(M_1 + M_2 + \dots + M_n)$

empinical mean

# Quick Review: Mean and Variance of Sum of Independent Random Variables (Cont.)



- **Example:** Each  $X_i$  has mean  $\mu_i$  and variance  $\sigma_i^2$ 
  - $X_1, X_2, \dots, X_n$  are assumed to be independent
  - $\Lambda_1, \Lambda_2, \cdots, \Lambda_n$  are assumed to be independent Question 3:  $Var[X_1 + X_2 + \cdots + X_n] = \sigma_1^2 + \sigma_2^2 + \cdots + \sigma_n^2$   $(= \sum_{i=1}^n \sigma_i^2)$

Question 4: 
$$Var\left[\frac{1}{n}(X_1 + X_2 + \dots + X_n)\right] = \frac{1}{n}(S_1 + S_2 + \dots + S_n)$$

$$V_{qv}[X_1+X_2] = V_{av}[X_1] + V_{av}[X_2] + 2 \cdot \underbrace{Cov(X_1,X_2)}_{0}$$

$$= \sigma_1^2 + \sigma_2^2$$

$$V_{av}[x_1+x_2+\cdots+x_n] = E[(x_1+x_2+\cdots+x_n) - E[x_1+x_2+\cdots+x_n]$$

$$= E[(x_1+x_2+\cdots+x_n) + (x_2-x_2) + \cdots + (x_n-x_n)$$

$$= e_1+e_2+e_1+e_2+e_2+e_3+e_4$$

$$= e_1+e_2+e_3+e_4+e_4$$

$$= E/(X_1 - M_1) + (X_2 - M_2) + (X_1 - M_1) + (X_2 - M_2) + (X_1 - M_1) + (X_2 - M_2) + (X_1 - M_2) + (X_2 - M_2$$

# Chebyshev's Inequality and Sample Mean

Example: Tossing moon blocks



- Each toss  $X_i$  is Bernoulli with P(outcome is "Yes")
- Each toss is independent from other tosses
- Question: Can we say anything about the sample mean

of 
$$n$$
 tosses  $\frac{1}{n}(X_1 + \dots + X_n)$ ?

Define  $X = \frac{1}{n}(X_1 + \dots + X_n)$ ?

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#### Chebyshev's Inequality and Sample Mean (Formally)

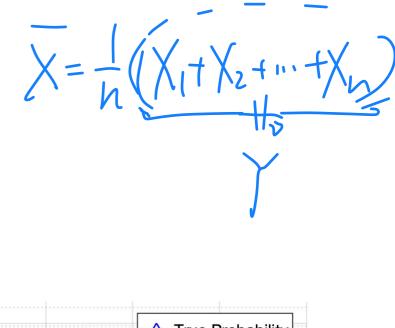
• Chebyshev's and Sample Mean: Let  $X_1, \dots, X_n$  be a sequence of independent and identically distributed (i.i.d.) random variables with mean  $\mu$  and variance  $\sigma^2$ . Define

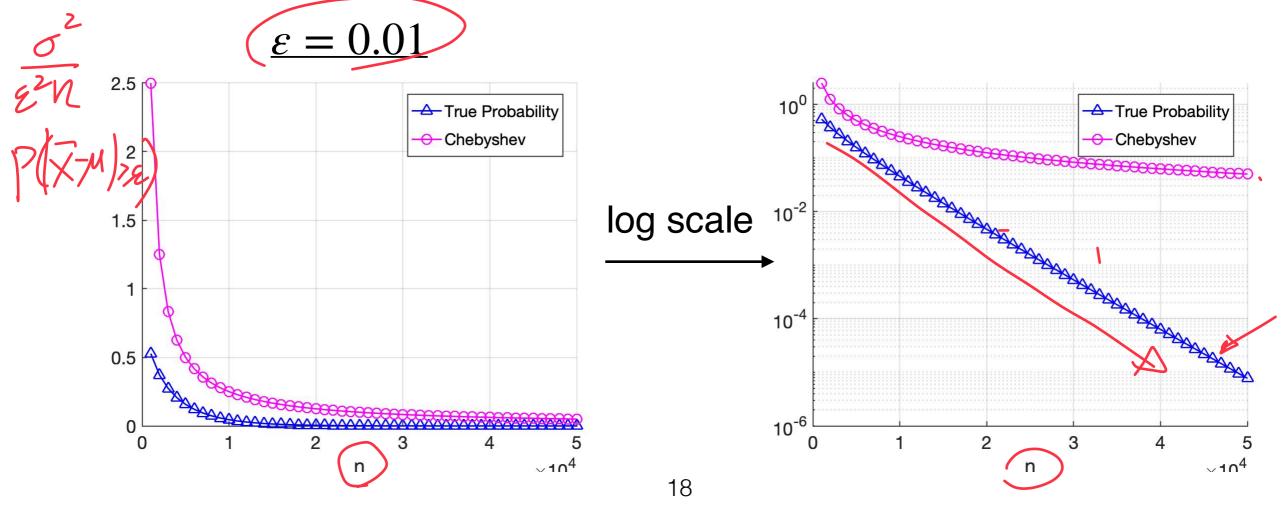
random variables with mean 
$$\mu$$
 and variance  $\sigma^2$ . Defining  $\bar{X} = \frac{1}{n}(X_1 + \dots + X_n)$ . Then, for any  $\varepsilon > 0$ , we have 
$$P(|\bar{X} - \mu| \ge \varepsilon) \le \frac{\sigma^2}{n}$$

$$P(|\bar{X} - \mu| \ge \varepsilon) \le \frac{\sigma^2}{\varepsilon^2 n} = \mathcal{O}(\frac{1}{n})$$

# Any Issue With Chebyshev's Inequality?

- Example:  $X_1, \dots, X_n$  are i.i.d. Bernoulli with parameter 0.5
  - $E[X_i] = \underline{\hspace{1cm}}$  and  $Var[X_i]$
  - Chebyshev's  $P(|\bar{X} \mu| \ge \varepsilon) \le \frac{\delta}{\varepsilon^{2} \hbar}$ Let's plot  $P(|\bar{X} \mu| \ge \varepsilon)$  for small  $\varepsilon$





#### Chernoff Bound

Proof:

• Chernoff Bound: Let X be a random variable with MGF  $M_X(t)$  Suppose  $M_X(t)$  exists for all t in some set S. Then, for any t>0 and  $t\in S$ , for any  $a\in \mathbb{R}$ , we have

$$P(X \ge a) \times e^{-ia} M_X(t)$$

$$= P(X \ge a) \times e^{-ia} M_X(t)$$

# Optimizing the Chernoff Bound

• Chernoff Bound Let X be a random variable with MGF  $M_X(t)$ Suppose  $M_X(t)$  exists for all t in some set S. Then, for any t > 0 and  $t \in S$ , for any  $a \in \mathbb{R}$ , we have

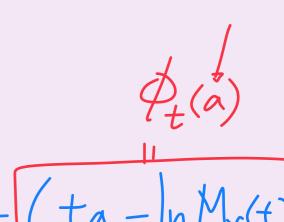
$$P(X \ge a) \le e^{-\phi(a)},$$

where 
$$\phi(a) = \max_{t>0, t\in S} (ta - \ln M_X(t))$$

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$$\phi(a) = \max_{t>0, t \in S} (ta - \ln M_X(t))$$

Proof:  $\gamma(x) = 0$ 
 $\gamma(x) = 0$ 

$$P(X>a) \leq \frac{-(\max_{t>0} \phi_{t}(a))}{t \in S}$$



## Example: Chernoff Bound for Bernoulli R.V.s

- Example: Suppose  $X \sim \text{Bernoulli}(p)$ 
  - What is  $M_X(t)$ ?
  - What is the Chernoff bound for X?  $(P(X \ge a) \le e^{-ta} \cdot M_X(t))$

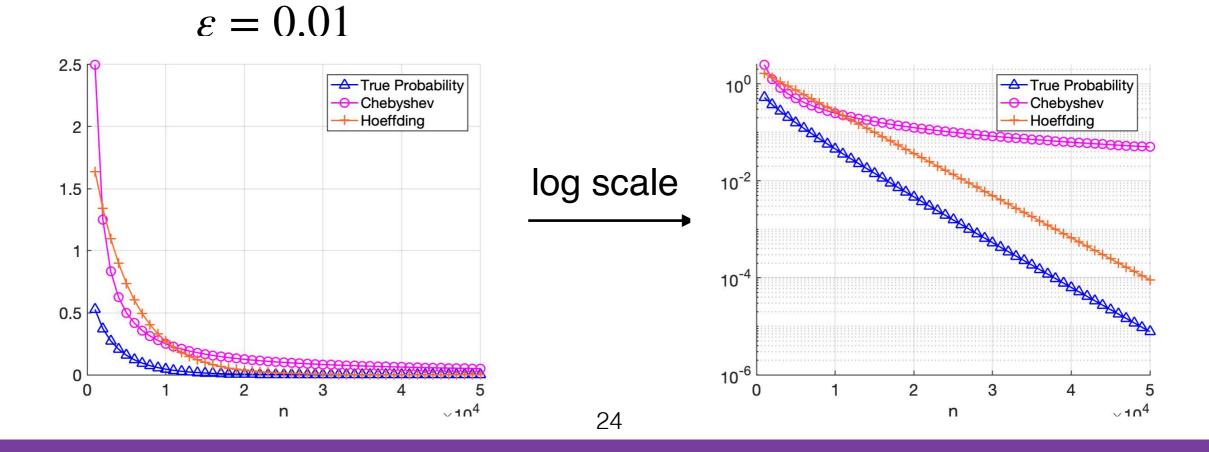
#### Example: Optimizing Chernoff Bound for Bernoulli R.V.s

- Example: Suppose  $X \sim \text{Bernoulli}(p)$ 
  - How to optimize the Chernoff bound for X?  $(P(X \ge a) \le e^{-\phi(a)}, \phi(a) = \max_{t>0, t \in S} (ta \ln M_X(t)))$

# How about applying Chernoff bound to "sum of independent random variables"?

# Hoeffding's Inequality (Formally)

► Hoeffding's Inequality (For Bernoulli): Let  $X_1, \dots, X_n$  be a sequence of i.i.d. Bernoulli random variables with parameter p. Define  $\bar{X} = \frac{1}{n}(X_1 + \dots + X_n)$ . Then, for any  $\varepsilon > 0$ , we have  $P(|\bar{X} - p| \ge \varepsilon) \le 2 \exp(-2n\varepsilon^2)$ 



# Proof of Hoeffding's Inequality (Positive Part)

$$P(\bar{X} - p \ge \varepsilon) \le \exp(-2n\varepsilon^2)$$

• [Hint] Chernoff bound:  $P(X \ge a) \le e^{-ta} \cdot M_X(t)$ 

$$P(\bar{X} - p \ge \varepsilon) \le$$

# Hoeffding's Lemma

► Hoeffding's Lemma: Let Z be a random variable with E[Z] = 0, and  $Z \in [a,b]$  with probability 1. Then, for any t > 0, we have  $E[e^{tZ}] \le \exp\left(\frac{t^2(b-a)^2}{2}\right)$ 

• Question: If  $Z \sim \text{Bernoulli}(p)$ , then  $E[e^{t(Z-p)}] \leq$ 

# Proof of Hoeffding's Inequality (Negative Part)

$$P(\bar{X} - p \le -\varepsilon) = P(p - \bar{X} \ge \varepsilon) \le \exp(-2n\varepsilon^2)$$

• [Hint] Chernoff bound:  $P(X \ge a) \le e^{-ta} \cdot M_X(t)$ 

$$P(p - \bar{X} \ge \varepsilon) \le$$

# Weak Law of Large Numbers (WLLN)

#### Review: Chebyshev's and Sample Mean: $n \to \infty$

• Chebyshev's and Sample Mean: Let  $X_1, \dots, X_n$  be a sequence of independent and identically distributed (i.i.d.) random variables with mean  $\mu$  and variance  $\sigma^2$ . Define  $S_n = (X_1 + \dots + X_n)$ . Then, for any  $\varepsilon > 0$ , we have

$$P\Big(\left|\frac{S_n}{n} - \mu\right| \ge \varepsilon\Big) \le \frac{\sigma^2}{\varepsilon^2 n}$$

• What if we let  $n \to \infty$ ?

# The Weak Law of Large Numbers (WLLN)

► The Weak Law of Large Numbers (Khinchin's Law): Let  $X_1, \dots, X_n$  be a sequence of independent and identically distributed (i.i.d.) random variables with mean  $\mu$ . Define  $S_n = (X_1 + \dots + X_n)$ . Then, for every  $\varepsilon > 0$ , we have  $P\left(\left|\frac{S_n}{n} - \mu\right| \ge \varepsilon\right) \to 0$  as  $n \to \infty$ 

Question: Any change in technical conditions (cf: Chebyshev's)?

Question: What does "convergence" mean here?

# Convergence in Probability

• Convergence of a Deterministic Sequence: Let  $a_1, a_2 \cdots$  be a sequence of real numbers. We say that  $a_n$  converges to a if for every  $\varepsilon > 0$ , there exists  $N_0$  such that

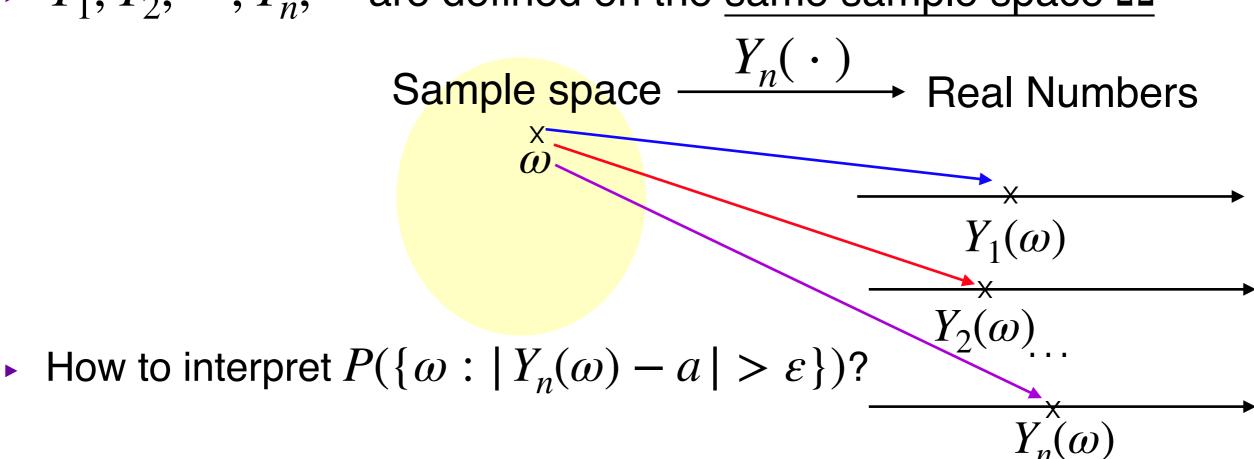
$$|a_n - a| \le \varepsilon$$
 for all  $n \ge N_0$ 

• Convergence to a <u>Scalar</u> in <u>Probability</u>: Let  $Y_1, Y_2 \cdots$  be a sequence of random variables, and let a be a real number. We say that  $Y_n$  converges to a in probability if for every  $\varepsilon > 0$ ,

Question: How to interpret this definition?

#### Recall: Random Variables Defined on $\Omega$

•  $Y_1, Y_2, \dots, Y_n, \dots$  are defined on the same sample space  $\Omega$ 



How about  $\lim_{n\to\infty} P(\{\omega: | Y_n(\omega) - a | > \varepsilon\}) = 0$ ?

# Example: Convergence in Probability

ightharpoonup Example: Consider a sequence of r.v.s  $Y_n$ 

$$P(Y_n = y) = \begin{cases} 1 - \frac{1}{n} & \text{, if } y = 0\\ \frac{1}{n} & \text{, if } y = n^2\\ 0 & \text{, otherwise} \end{cases}$$

- For every  $\varepsilon > 0$ , can we find  $P(|Y_n 0| > \varepsilon)$ ?
- How about  $\lim_{n\to\infty} P(|Y_n 0| > \varepsilon)$ ?