Project 10 Structured sparsity in genetics

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Problem 28: Uniqueness of predictions from the lasso

Proof. To prove that $\mathbf{X}\hat{\beta}_1 = \mathbf{X}\hat{\beta}_2$, it's equivalent to prove $\hat{\beta}_1 = \hat{\beta}_2$. We'll prove by contradiction.

Let the solution set be denoted as S, we have $\hat{\beta}_1, \hat{\beta}_2 \in S$.

Since S is convex, we also have $\alpha \hat{\beta}_1 + (1-\alpha)\hat{\beta}_2 \in S$.

So that we have:

$$\frac{1}{2}\left\| \mathbf{y} - \mathbf{X} \left(\alpha \hat{\beta}_1 + (1 - \alpha) \hat{\beta}_2 \right) \right\| + \lambda \left\| \alpha \hat{\beta}_1 + (1 - \alpha) \hat{\beta}_2 \right\| = c^* \tag{1}$$

Define the loss function as $f: S \mapsto \|\mathbf{y} - \mathbf{X}\beta\|_2^2$. We know that the lineare loss function is stirctly convex, so that we have:

$$f\left(\alpha\hat{\beta}_1 + (1-\alpha)\hat{\beta}_2\right) < \alpha f(\hat{\beta}_1) + (1-\alpha)f(\hat{\beta}_2) \tag{2}$$

Similarly, since l_1 -nrom is convex, we have:

$$\left\|\alpha\hat{\beta}_1 + (1-\alpha)\hat{\beta}_2\right\|_1 \le \alpha\|\hat{\beta}_1\| + (1-\alpha)\|\hat{\beta}_2\| \tag{3}$$

Combining Equation 2 and Equation 3, we can start the following derivation:

$$\begin{split} &\frac{1}{2}f\left(\alpha\hat{\beta}_{1}+(1-\alpha)\hat{\beta}_{2}\right)+\lambda\left\|\alpha\hat{\beta}_{1}+(1-\alpha)\hat{\beta}_{2}\right\|\\ &<\alpha\left(f(\hat{\beta}_{1})+\lambda\|\hat{\beta}_{1}\|\right)+(1-\alpha)\left(f(\hat{\beta}_{2})+\lambda\|\hat{\beta}_{2}\|)\right)\\ &=\alpha c^{*}+(1-\alpha)c^{*}\\ &=c^{*} \end{split}$$

Just a re-write:

$$\frac{1}{2} f\left(\alpha \hat{\beta}_1 + (1-\alpha)\hat{\beta}_2\right) + \lambda \left\|\alpha \hat{\beta}_1 + (1-\alpha)\hat{\beta}_2\right\| < c^* \tag{4}$$

Equation 4 contradicts with Equation 1, which suggests that our assumption $\hat{eta}_1 \neq \hat{eta}_2$ is wrong.

\end{proof}

Problem 29: Ridge regression solution

Given $n \times p$ matrix **X**, response n-vector **y**, and the parameter p-vector β , consider augmenting **X** with rows corresponding to $\sqrt{\lambda}$ times a $p \times p$ identity matrix I:

$$\mathbf{X}_* = \begin{pmatrix} \mathbf{X} \\ \sqrt{\lambda} \mathbf{I} \end{pmatrix}$$

and y is similarly augmented with p zeros at its end into y_* .

Now, the least square objective function on the modified dataset turns out to be:

$$(\mathbf{y}_* - \mathbf{X}_* \beta)^T (\mathbf{y}_* - \mathbf{X}_* \beta)$$

$$= \mathbf{y}_*^T \mathbf{y}_* - 2\mathbf{y}_*^T \mathbf{X}_* \beta + \beta^T \mathbf{X}_*^T \mathbf{X}_* \beta$$
(5)

The solution to the linear regression on the modified dataset hence follows:

$$\hat{\beta} = (\mathbf{X}_*^T \mathbf{X}_*)^{-1} \mathbf{X}_*^T \mathbf{y} \tag{6}$$

Since we are just padding y with zeros, the augmented part of X_* will not have any effect when it is multiplied by y_*^T :

$$\mathbf{y}_{*}^{T}\mathbf{X}_{*} = \mathbf{y}^{T}\mathbf{X} \tag{7}$$

And finally with augmented dataset X_* we have:

$$\mathbf{X}_{*}^{T}\mathbf{X}_{*}$$

$$= (\mathbf{X}^{T} \quad \sqrt{\lambda}\mathbf{I}) \begin{pmatrix} \mathbf{X} \\ \sqrt{\lambda}\mathbf{I} \end{pmatrix}$$

$$= \mathbf{X}^{T}\mathbf{X} + \lambda\mathbf{I}$$
(8)

Plug Equation 7 and Equation 8 back into Equation 6, we delightfully sees that performing linear regression on the modified dataset gives:

$$\hat{\beta} = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y}$$
(9)

which is just the ridge regression solution.

Problem 30: Variable selection under various norms

Load data and split into training and test set

```
# load data
set.seed(42)
load("yeastStorey.rda")

sample <- createDataPartition(data$Marker, p = 0.7, list = FALSE)

train <- data[sample, ]
test <- data[-sample, ]

X_train <- as.matrix(train[, -1])
y_train <- train[, 1]

X_test <- as.matrix(test[, -1])
y_test <- test[, 1]</pre>
```

Do 10-fold cross validation

```
alphas <- seq(0, 1, 0.1)
foldid <- sample(1:10, size = length(sample), replace = TRUE)</pre>
num_cores <- detectCores()</pre>
registerDoParallel(num_cores)
start <- Sys.time()</pre>
res <- foreach (alpha = alphas, .combine = c, .options.RNG = 42) %dorng% {
  cv_res <- cv.glmnet(X_train, y_train, alpha = alpha,</pre>
                       family = "binomial", foldid = foldid,
                        type.measure = "mse")
 list(cv_res)
end <- Sys.time()</pre>
end - start
## Time difference of 0.2892711 secs
stopImplicitCluster()
df <- do.call(rbind,</pre>
               lapply(1:length(alphas),
                      function(x) {
                       cbind.data.frame(alphas[x],
                                          res[[x]]$lambda,
                                          res[[x]]$cvm)
                      })
colnames(df) <- c("alpha", "lambda", "cvm")</pre>
```

```
param <- filter(df, cvm = min(cvm))
param</pre>
```

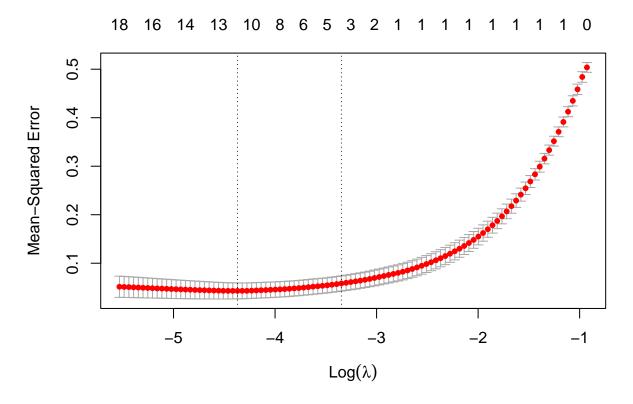
```
## alpha lambda cvm
## 1 1 0.0126508 0.04302968
```

For the optimal α (here $\alpha = 1$)...

```
opt_idx <- which(alphas = param$alpha)
res.opt_alpha <- res[[opt_idx]]</pre>
```

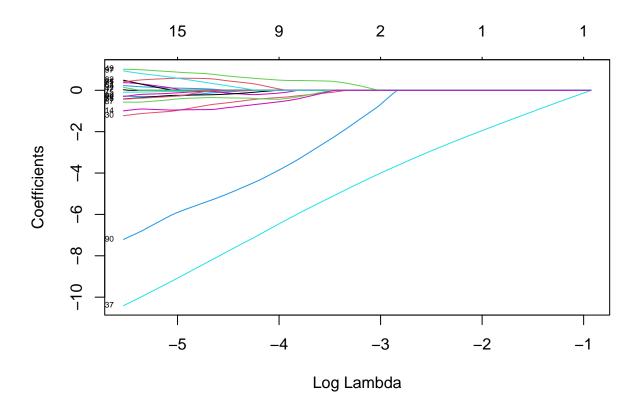
Plot the mean cross-validated error as a function of log λ .

```
plot(res.opt_alpha)
```



Plot the trace curve of coefficients as a function of $\log \lambda$.

```
plot(res.opt_alpha$glmnet.fit, "lambda", label=TRUE)
```



Fit the final model and predict the response

Report the varaibles selected.

```
selected <- coef(res.opt_alpha, param$lambda)
selected@Dimnames[[1]][selected@i]

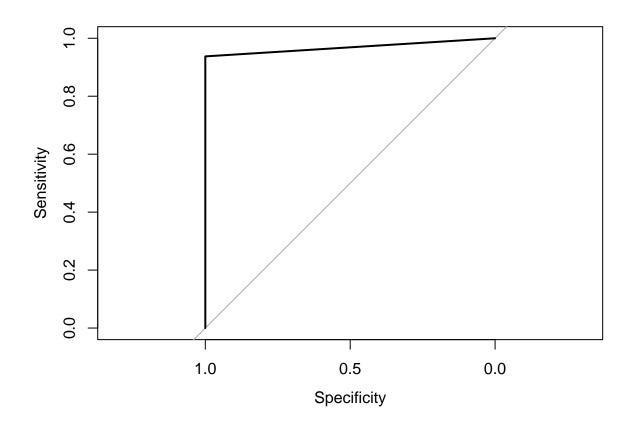
## [1] "YDL180W" "YEL007W" "YGR046W" "YHL018W" "YIR016W" "YKR096W" "YLR012C"

## [8] "YLR281C" "YNL149C" "YNL213C" "YOL057W" "YPL066W"

Plot ROC...

test_roc <- roc(y_test, as.vector(y_pred))</pre>
```

```
## Setting levels: control = 0, case = 1
## Setting direction: controls < cases
plot(test_roc)</pre>
```



... and report the corresponding AUC.

test_roc\$auc

Area under the curve: 0.9688