

Lecture 4

Last time

- ▷ Robust mean estimation
- ▷ Chernoff's inequality

Today

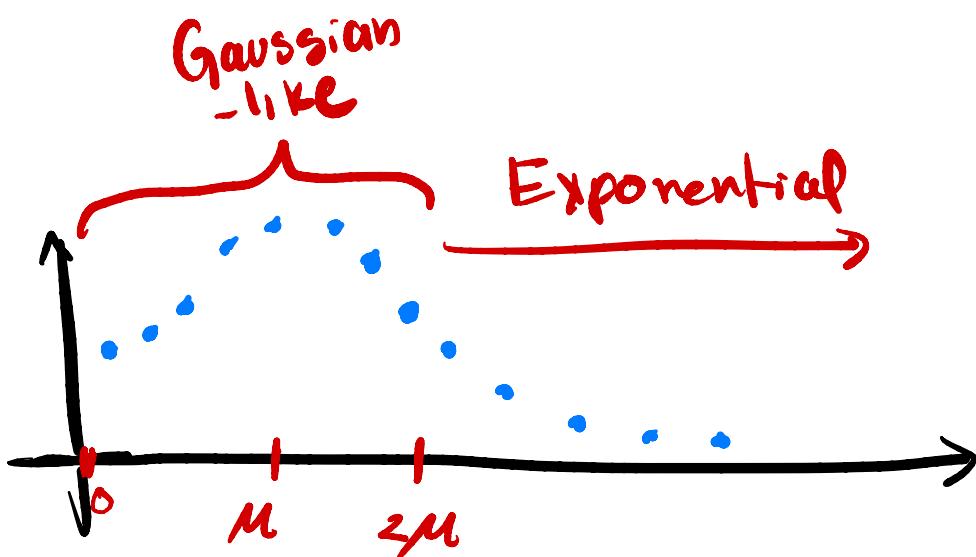
- ▷ Degree of random graphs.
- ▷ Subexponentials

Erdős-Renyi Graphs

Last time we finished with:

Corollary: Take $\delta \in [0, 1]$, then

$$P(|S_n - \mu| \geq \delta \mu) \leq \exp\left(-\frac{\delta^2 \mu}{6}\right).$$



Def $G(n, p)$: Fix a set of n nodes connect each pair independently with probability p .

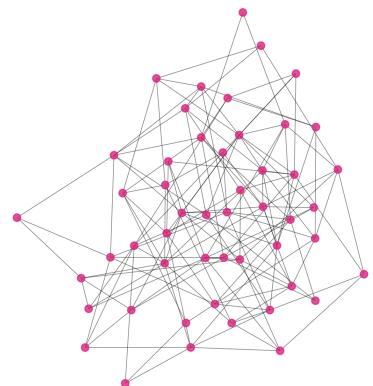
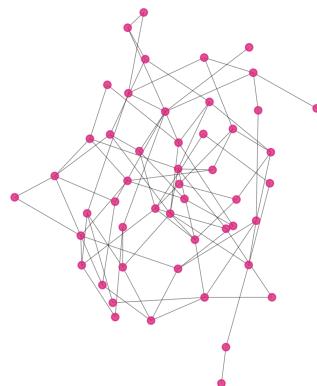
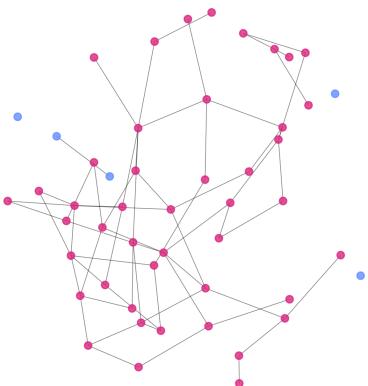
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Erdős-Rényi Graphs at Different p Values (n = 50)

$0.7 \cdot p_c$
 $p = 0.0548$
Largest component: 45/50

p_c
 $p = 0.0782$
Largest component: 50/50

$1.5 \cdot p_c$
 $p = 0.1174$
Largest component: 50/50



The degree $\deg(i)$ of node i is equal to the number of edges connected to i .

Then,

$$\deg(i) = S_{n-1} = (n-1)p =: d$$

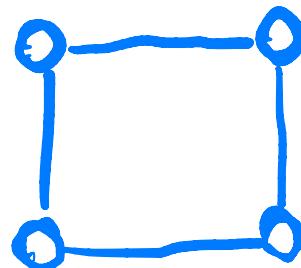
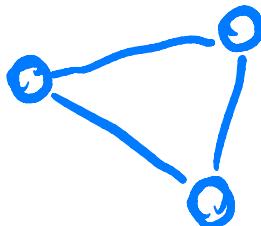
There is an interesting phase transition at $p = \log n / n$

$P \geq (1+\epsilon) \log n / n \rightarrow$ Large component.

$P \leq (1-\epsilon) \log n / n \rightarrow$ Disconnected.

Something similar happens with their degree.

Def: A graph is d -regular if $\deg(i) = d$ for all nodes. +



Proposition: Take $\delta > 0$, and $(n-1)p \geq \frac{8}{\delta^2} \log(2n)$. Then, $G \sim G(n, p)$ is almost $\frac{(n-1)p}{d}$ -regular, i.e.,

$$P(\forall i \in [n], (1-\delta)d \leq \deg(i) \leq (1+\delta)d) \leq 1 - \frac{1}{2n}$$

Proof: For each i , we have

$$\deg(i) = \sum \mathbf{1}_{ij \in E} \xleftarrow{\text{Random edges of } G.}$$

where $E \deg(i) = (n-1)p$ and Chernoff's inequality yields

$$P(|\deg(i) - d| \geq \delta d) \leq 2 \exp(-\delta^2 d / 4).$$

Taking union bound

$$P(\bigcap \mathcal{E}_i) = 1 - P(\bigcup \mathcal{E}_i^c)$$

$$\begin{aligned}
 &\geq 1 - \sum_i P(\epsilon_i^c) \\
 &\geq 1 - 2n \exp(-\delta^2 d/4) \\
 &= 1 - \exp(\log(2n) - \delta^2 d/4) \\
 &\approx 1 - \frac{1}{2n}.
 \end{aligned}$$

□

Sub-exponential

Not all random variables are sub-Gaussian, indeed one often encounters squared Gaussians (χ^2).

Let's see what the MGF looks like in that case. If $Z \sim N(0,1)$,

$$\begin{aligned}
 \mathbb{E}[e^{\lambda(Z^2-1)}] &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\lambda(t^2-1)} e^{-t^2/2} dt \\
 &= \frac{e^{-\lambda}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-(1-2\lambda)t^2/2} dt
 \end{aligned}$$

If $|\lambda| \geq \frac{1}{2}$, the integrant diverges.

If $|\lambda| < \frac{1}{2}$, this is a Gaussian with different variance.

$$\mathbb{E} e^{\lambda(z^2 - 1)} = \begin{cases} \frac{e^{-\lambda}}{\sqrt{1-2\lambda}} & \text{if } |\lambda| < \frac{1}{2}, \\ +\infty & \text{if } |\lambda| \geq \frac{1}{2}. \end{cases}$$

We can further upper bound

$$\mathbb{E} e^{\lambda(z^2 - 1)} \leq \frac{e^{-\lambda}}{\sqrt{1-2\lambda}} \leq e^{4\lambda^2/2} \quad \forall |\lambda| < \frac{1}{4}$$

We follow the same strategy as with sub-Gaussians.

Def (sub-exponential): A r.v. X

(σ, α) -subexponential if

$$\mathbb{E} e^{\lambda(X-\mu)} \leq e^{\sigma^2 \lambda^2/2} \quad \forall |\lambda| \leq \frac{1}{\alpha}.$$

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Thus, if $Z \sim N(0, 1)$, then, Z^2 is $(2, 4)$ -subexponential.

Theorem: Let X be a (σ, α) -sub-exponential r.v. with $\mathbb{E}X = \mu$. Then,

$$P(X - \mu \geq t) \leq \begin{cases} e^{-t^2/2\sigma^2} & \text{if } |t| \leq \frac{\sigma^2}{\alpha} \\ e^{-\frac{t}{\alpha} + \frac{\sigma^2}{2\alpha^2}} & \text{otherwise.} \end{cases}$$

Proof: Applying the Chernoff bound

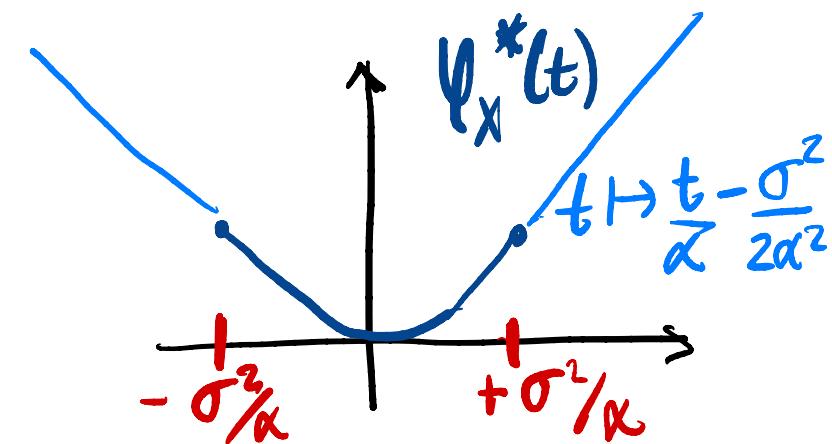
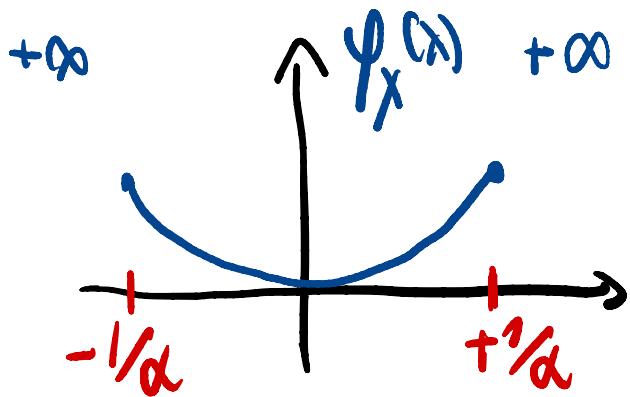
$$\log P(X - \mu \geq t) \leq -\psi_X^*(t)$$

where

$$\begin{aligned}\psi_X(\lambda) &= \log \mathbb{E} e^{\lambda(X-\mu)} \\ &\leq \begin{cases} \sigma^2 \lambda^2 / 2 & \text{if } |\lambda| \leq \frac{1}{\alpha} \\ +\infty & \text{otherwise.} \end{cases}\end{aligned}$$

Thus,

$$\psi_X^*(t) = \sup_{\lambda} \{t\lambda - \psi_X(\lambda)\}$$



Clearly, $|\lambda| \leq 1/\alpha$ in the sup.

Thus, there are two cases:

Case 1: Suppose $\lambda \mapsto t\lambda - \sigma^2\lambda^2/2$

has a maximizer with $|\lambda| \leq 1/\alpha$.

Taking derivatives equal zero

$$\lambda = t/\sigma^2 \text{ so } |t| \leq \frac{\sigma^2}{\alpha} \text{ and}$$

$$\Psi_x^*(t) = \frac{t^2}{2\sigma^2}$$

Case 2: The maximizer λ^* of
 $\lambda \mapsto t\lambda - \sigma^2\lambda^2/2$ satisfying $|\lambda^*| > 1/\alpha$

and so if $\lambda^* > 0$ we take

the extreme $\lambda = 1/\alpha$

$$\Psi_x^*(t) = \frac{t}{\alpha} - \frac{\sigma^2}{2\alpha^2}.$$

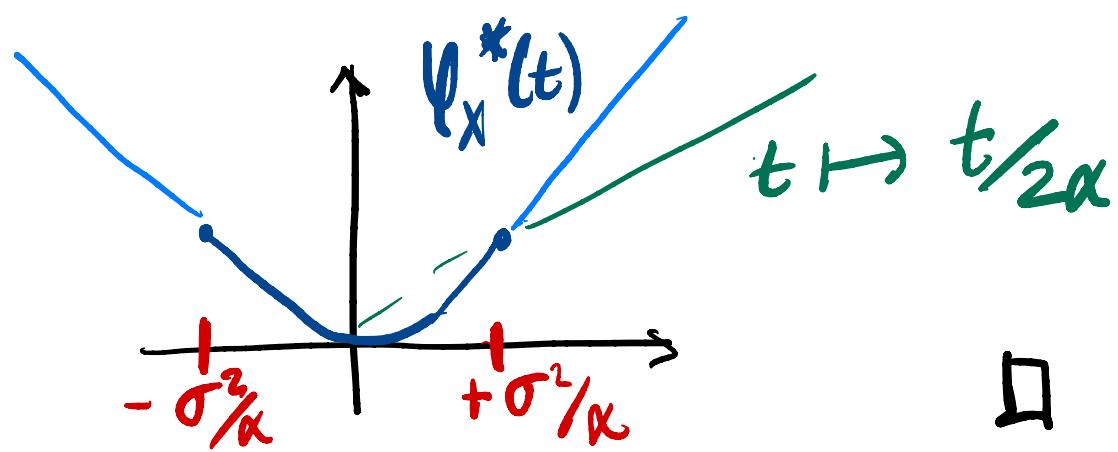
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Corollary : Let X be (σ, α) -subexponential r.v. with $\mathbb{E} X = \mu$.

Then,

$$P(|X - \mu| \geq t) \leq 2 \exp\left(-\left(\frac{t^2}{2\sigma^2} \wedge \frac{t}{2\alpha}\right)\right)$$

Proof: Proof by picture:



Lemma (sum rule): Let X_1, \dots, X_n be independent r.v. with X_i (σ_i, α_i) -subexponential. Then, $\sum_{i=1}^n X_i$ is $(\|\sigma\|_2, \|\alpha\|_\infty)$

↑
max | α_i |

- subexponential.

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Combining the sum rule with
Corollary :) yields:

Theorem (Bernstein's Inequality):

Let X_1, \dots, X_n be independent r.v.s
with $X_i (\sigma_i, \alpha_i)$ subexponential
and $E X_i = \mu$. Then,

$$P\left(\left|\sum_{i=1}^n (X_i - \mu_i)\right| \geq t\right) \leq 2 \exp\left(-\left(\frac{t^2}{2\sigma^2 \tau^2} \wedge \frac{t}{\|\alpha\|_\infty}\right)\right)$$

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