

Lecture 23

Last time

- ▷ Markov Property
- ▷ Consequences

Today

- ▷ Strong Markov Property
- ▷ Consequences

Question: Imagine we start $B_0 = 0$, then what is the less likely point $t \in [0, 1]$ to be the last point in $[0, 1]$ $B_t = 0$?

What's the distribution of $L = \sup\{t : B_t = 0\}$?
The answer to the first question
is $\frac{1}{2}$!

Stopping times

We say that a filtration $\{\mathcal{F}_t\}$ is right continuous if

$$\bigcap_{s > t} \mathcal{F}_s = \mathcal{F}_t.$$

The reason we like right continuous filtrations is that they make infinitesimals into the future negligible.

Def: We say that a random variable

S in $[0, \infty]$ is a stopping time w.r.t.
a filtration $\{\mathcal{F}_t\}$ if $\{S < t\} \in \mathcal{F}_t$

Ht.

This is the same as $\{S \leq t\} \in \mathcal{F}_t$
if $\{\mathcal{F}_t\}$ is r.c. Why?

We need to understand what type of things are stopping times?

Q: Say that $G \subseteq \mathbb{R}$ is an open set or a closed set. Is $T_G = \inf \{t : B_t \in G\}$ a stopping time? Yes!

Theorem: If G is an open set, then T_G is a stopping time.

Proof: Since G is open and $t \mapsto B_t$ is continuous, then

$$\{T < t\} = \bigcup_{\substack{q < t \\ q \in \mathbb{Q}}} \{B_q \in G\},$$

thus we conclude that $\{T < t\} \in \mathcal{F}_t$. \square

Theorem: Suppose that T_n is a sequence of stopping times. If either

$$T_n \downarrow T \quad \text{or} \quad T_n \uparrow T.$$

Then, T is a stopping time. \(\square\)

Proof: It suffices to note that

$$\{T < t\} = \bigcup_n \{T_n < t\} \quad \text{and} \quad \{T \leq t\} = \bigcap_n \{T_n \leq t\}$$

Theorem: If G is closed, then T_G is a stopping time.

Proof: Let $B(x, r) = \{y : |y - x| < r\}$, let $G_n = \bigcup_{x \in K} B(x, \frac{1}{n})$ and let $T_n = \inf \{t \geq 0 : B_t \in G_n\}$. Since G_n is open $\Rightarrow T_n$ is a stopping time. Next we show that $T_n \uparrow T$. Notice that by construction $T_n \leq T$ and $T_n \uparrow t^*$ for some $t^* < \infty$. Since $B_{T_n} \in \overline{G_n}$ $\forall n \Rightarrow B_{T_n} \rightarrow B_{t^*} \in G$ and so $t^* \geq T \Rightarrow \lim T_n = T$. \(\square\)

Strong Markov Property

We develop an analogue of the SMP. We now define the random shift operator. Given a nonnegative rv S in $[0, \infty]$, define

$$(\Theta_S(w))(t) = \begin{cases} w(S(w) + t) & \text{on } \{S < \infty\} \\ \Delta & \text{on } \{S = \infty\} \end{cases}$$

Extra symbol.

We also define the information known at time s :

$$\mathcal{F}_s = \{A : A \cap \{s \leq t\} \in \mathcal{F}_t \text{ for all } t \geq 0\}.$$

Proposition: If $S \leq T$ are stopping times
 $\Rightarrow \mathcal{F}_S \subseteq \mathcal{F}_T$. +

Proposition: If $T_n \downarrow T$ are stopping times
 $\Rightarrow \mathcal{F}_T = \bigcap \mathcal{F}_{T_n}$. +

Exercise: Prove these two facts!

Theorem (Strong Markov Property). Let $(s, \omega) \rightarrow Y_s(\omega)$ be bounded and $\mathbb{R} \times \Omega$ measurable. If S is a stopping time, then for all $x \in \mathbb{R}$

$$\mathbb{E}_x [Y_S \circ \theta_S | \mathcal{F}_S] = \mathbb{E}_{B_S} Y_S \text{ on } \{S < \infty\}. \quad \dagger$$

Function $\varphi(x, t) = \mathbb{E}_x Y_t$
evaluated at $x = B_S$ and $t = S$.

The proof of this result is similar to the one we covered in Lecture 16 (albeit much more technical), see Theorem

7.3.a. in Durrett.

Reflection Principle

Let $a > 0$ and $T_a = \inf \{t : B_t = a\}$

Theorem:

$$P_0(T_a < t) = 2 P_0(B_t \geq a).$$

Proof: We shall see that

$$P_0(T_a < t, B_t > a) = \frac{1}{2} P_0(T_a < t), \quad (\heartsuit)$$

which right away implies:

$$P_0(T_a < t) = 2 P_0(B_t \geq a)$$

since $\{T_a < t\} \supseteq \{B_t > a\}$.

We focus on (\heartsuit) . We will use the SMP, define

$$Y_s(w) = \begin{cases} 1 & \text{if } s < t, w(t-s) > a, \\ 0 & \text{otherwise.} \end{cases}$$

If we let $s = \inf \{s < t : B_s = a\}$ with $\inf \emptyset = \infty$, then

$$Y_s(\theta_s(w)) = \begin{cases} 1 & \text{if } s < t, B_t > a \\ 0 & \text{otherwise.} \end{cases}$$

So SMP gives

$$\begin{aligned} \mathbb{E}_0(Y_s \circ \theta_s | \mathcal{F}_s) &= \mathbb{E}_{B_s}(Y_s) = \mathbb{E}_a(Y_s) \\ \text{on } \{s < \infty\} = \{T_a < t\} &= \frac{1}{2} \quad \leftarrow \text{Gaussian centered at } a. \end{aligned}$$

Taking expectations

$$\begin{aligned}
 P_0(T_a < t, B_t \geq a) &= E_0[(Y_s^0 \theta_s) \mathbb{1}_{\{s < \infty\}}] \\
 &= E_0[E_0[Y_s^0 \theta_s | \tilde{\mathcal{F}}_s] \mathbb{1}_{\{s < \infty\}}] \\
 &= \frac{1}{2} E_0[\mathbb{1}_{\{T_a < t\}}] \\
 &= \frac{1}{2} P(T_a < t).
 \end{aligned}$$

□

But this means that we have a closed form for the dist. of T_a :

$$P_0(T_a < t) = 2 P_0(B_t \geq a) = \frac{2}{\sqrt{2\pi t}} \int_a^\infty \exp(-x^2/2t) dx.$$

Recall we were interested in $L = \inf \{t \leq 1 : B_t = 0\}$.

Then,

$$\begin{aligned}
 P_0(L \leq t) &\stackrel{\text{HW 5}}{=} \int_{-\infty}^{-\infty} P_t(0, y) P_y(T_0 > 1-t) dy \\
 &= 2 \int_0^\infty (2\pi t)^{-1/2} \exp(-y^2/2s) \\
 &\quad \int_{1-s}^\infty (2\pi u)^{-1/2} \exp(-x^2/2t) dx
 \end{aligned}$$

:

← Change of variables and simplifications

$$= \frac{1}{\pi} \int_0^t (s(1-s))^{-1/2} ds = \frac{2}{\pi} \arcsin(\sqrt{s})$$

Thus, the density of L is $\frac{1}{\pi} s(1-s)^{-1/2}$
which looks like

