

Homework 6 Stochastic - Matheus Raku

Exercise 1.1) (page 12 d)

$$T = \inf\{t: B_t = 0 \text{ or } B_t = 3\}$$

B_t is a standard Brownian Motion

$$E[B_0] = E[B_T] \rightarrow \text{optional sampling theorem}$$

starts
at x

$$\leftarrow X = 0 \cdot P(B_T = 0) + (3) P(B_T = 3)$$

$$P(B_T = 3) = \frac{x}{3} // \rightarrow \text{probability that Brownian motion reaches 3 before reaching 0}$$

Exercice 1.2

$$X_t = x + \int_0^t B_s dB_s = x + \frac{1}{2} [B_t^2 - t] \quad (\text{page 104})$$

$$d(X_t) = d\left(\frac{1}{2} [B_t^2 - t]\right), \quad x \text{ is a constant}$$

$$f(t, B_t) = \frac{1}{2} B_t^2 - \frac{1}{2} t$$

$$\dot{f} = -\frac{1}{2}, \quad f' = \frac{1}{2} \cdot 2 \cdot B_t = B_t, \quad f'' = 1$$

$$d(X_t) = f' dB_t + \left[\dot{f} + \frac{1}{2} f'' \right] dt \rightarrow \text{Ito's formula II}$$

$$= B_t dB_t + \left(-\frac{1}{2} + \frac{1}{2} \cdot 1 \right) dt$$

$$d(X_t) = B_t dB_t + 0 dt$$

drift term is 0 $\rightarrow X_t$ is a local martingale.

Since B_t is a standard Brownian motion, X_t is a continuous martingale.

at $t=0$,

$$X_0 = x + \int_0^0 B_s dB_s = x + 0 = x$$

Since X_t is a continuous martingale, starting at x , we can apply the same method as 1.1

$$E[X_0] = E[X_t]$$

$$x = 0 \cdot P(X_t = 0) + 3 \cdot P(X_t = 3)$$

$$P(X_t = 3) = \frac{x}{3} \rightarrow \text{Prob. that } X_t \text{ reached 3 before reaching 0}$$

1.2

Exercise 2.1) page 131-132

for $T = T(r, R) = \min\{t: X_t = r \text{ or } X_t = R\}$

for $r \leq x \leq R$, $\phi(x) = P(X_T = R | X_0 = x)$

$$\phi(x) = \frac{x^{1-2a} - r^{1-2a}}{R^{1-2a} - r^{1-2a}}, \quad a \neq \frac{1}{2}$$

in our case, $X_0 = 1$, $a < \frac{1}{2}$,

$$P(r, R) = P(X_T = R | X_0 = 1) = \frac{1^{1-2a} - r^{1-2a}}{R^{1-2a} - r^{1-2a}} = \frac{1 - r^{1-2a}}{R^{1-2a} - r^{1-2a}}$$

$$P(0, R) = \lim_{r \downarrow 0} P(r, R) = \lim_{r \rightarrow 0^+} P(r, R) = \frac{1 - 0}{R^{1-2a} - 0} = \frac{1}{R^{1-2a}} //$$

Exercise 2.2

$$\lim_{R \rightarrow \infty} P(0, R) = \lim_{R \rightarrow \infty} \frac{1}{R^{1-2a}} = \frac{1}{\infty^{1-2a}} = \frac{1}{\infty} = 0$$

the denominator diverges, the limit goes to 0 as $R \rightarrow \infty$

~~this means the Bessel process eventually reaches 0 as $R \rightarrow \infty$, with probability one.~~

The probability that the Bessel process reaches ∞ is 0.

The Bessel process will eventually reach 0 with probability one.

Exercise 2.3) page 132

$$\text{For } a = \frac{1}{2}, \phi(x) = P(X_T = R \mid X_0 = x) = \frac{\ln(x) - \ln(r)}{\ln(R) - \ln(r)}$$

$$\text{For } a = \frac{1}{2}, X_0 = 1,$$

$$P(r, R) = \frac{\ln(1) - \ln(r)}{\ln(R) - \ln(r)} = \frac{-\ln(r)}{\ln(R) - \ln(r)}$$

$$P(0, R) = \lim_{r \downarrow 0} \frac{-\ln(r)}{\ln(R) - \ln(r)} = \frac{-(-\infty)}{\ln(R) - (-\infty)} = \frac{\infty}{\infty} \text{ } \left. \vphantom{\lim_{r \downarrow 0}} \right\} \text{indeterminate}$$

L'Hopital's rule: if $\lim_{x \rightarrow a} \frac{f(x)}{g(x)} = \text{indeterminate}$, $\lim_{x \rightarrow a} \frac{f(x)}{g(x)} = \lim_{x \rightarrow a} \frac{f'(x)}{g'(x)}$

$$P(0, R) = \lim_{r \downarrow 0} \frac{-\ln(r)}{\ln(R) - \ln(r)} = \lim_{r \rightarrow 0^+} \frac{-1/r}{-1/r} = \lim_{r \downarrow 0} \frac{1/r}{1/r} = 1 //$$

here, $r \rightarrow 0^+$, $T = \min\{t: X_t = r \text{ or } R\}$

The process starts at 1 ($X_0 = 1$). The lowest possible value of X_t is 0.

The process can hit 0 or R, and the probability that $X_T = R$ is 1 ($P(0, R) = 1$). This means that X_t is always positive.

With probability one, $X_t > 0$ for all t . //

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Homework 6 – Stochastic Calculus

Exercise 2.4

For $a = \frac{1}{2}$, $X_0 = 1$,

$$\rho(r, \infty) = \lim_{R \rightarrow \infty} \rho(r, R) = \lim_{R \rightarrow \infty} \frac{\ln(1) - \ln(r)}{\ln(R) - \ln(r)} = \lim_{R \rightarrow \infty} \frac{-\ln(r)}{\ln(R) - \ln(r)} = \frac{-\ln(r)}{\infty} \rightarrow 0$$

$$\rho(r, \infty) = \lim_{R \rightarrow \infty} \rho(r, R) \rightarrow 0$$

As $R \rightarrow \infty$, the denominator diverges, and the limit goes to 0.

This means that for

$$T = T_{r,R} = \min \{t: X_t = r \text{ or } R\}$$

The probability that $X_t = \infty$ is 0.

$$\rho(r, \infty) = 0$$

which means that X_t will never reach ∞ .

This also means that process X_t will be below r for some t and it can also be above r for some t .

With probability one, for all $r > 0$, $X_t < r$ for some t .

Exercise 3.1)

$$f(t, \beta_t) = X_t = 2e^{\beta_t} \cdot e^{-t}$$

$$\dot{f}(t, \beta_t) = -1 \cdot 2e^{\beta_t - t}$$

$$f'(t, \beta_t) = f''(t, \beta_t) = 2e^{\beta_t - t} \quad \text{Using Ito's formula II}$$

$$dX_t = f'(t, \beta_t) d\beta_t + \left[\dot{f}(t, \beta_t) + \frac{1}{2} f''(t, \beta_t) \right] dt$$

$$= [-2e^{\beta_t - t} + e^{\beta_t - t}] dt + 2e^{\beta_t - t} d\beta_t$$

$$dX_t = -e^{\beta_t - t} dt + 2e^{\beta_t - t} d\beta_t = -\frac{X_t}{2} dt + X_t d\beta_t$$

$$\mu(X_t) = -\frac{X_t}{2} //$$

$$\sigma(X_t) = X_t //$$

} these functions do not depend on t .
therefore, X_t is a time-homogeneous diffusion

Exercise 3.2

$$m(x) = -\frac{1}{2}x, \quad \sigma(x) = x, \quad r = 2$$

Feynman-Hac formula: page 135

$$\dot{\phi}(t, x) = -m(t, x) \phi'(t, x) - \frac{1}{2}(\sigma(t, x))^2 \phi''(t, x) + r(t, x) \phi(t, x)$$

$$\dot{\phi}(t, x) = \frac{x}{2} \cdot \phi'(t, x) - \frac{1}{2} x^2 \phi''(t, x) + 2 \phi(t, x) \rightarrow \text{PDE} //$$

for $0 \leq t \leq T$, with terminal condition $\phi(T, x) = (x - 3)_+$
 $\hookrightarrow F(x)$

Exercise 3.3

PDE:

$$\dot{\phi}(t, x) = \frac{x}{2} \phi'(t, x) - \frac{1}{2} x^2 \phi''(t, x) + 2 \phi(t, x)$$

for $0 \leq t \leq T$, with terminal condition $\phi(T, x) = x^2 e^{-x}$ //

Exercise 4.1 Page 24: martingale M_n is square integrable if for each n , $E[M_n^2] < \infty$

Page 89: Variance Rule: $\text{Var}[Z_t] = E[Z_t^2] = \int_0^t E[A_s^2] ds$

$$\alpha = 1/4$$

$$E[Z_t^2] = \int_0^t \frac{1}{(1-s)^{1/4 \cdot 2}} ds = \int_0^t \frac{1}{(1-s)^{1/2}} ds = \int_0^t \frac{1}{u^{1/2}} \cdot (-du) = - \int u^{-1/2} du$$

$$\begin{array}{l} u = 1-s \\ du = -ds \\ ds = -du \end{array} \left| \begin{array}{l} E[Z_t^2] = -\frac{1}{1/2} u^{1/2} \Big|_{s=0}^{s=t} = -2(1-s)^{1/2} \Big|_0^t = -2[(1-t)^{1/2} - 1] \\ E[Z_t^2] = 2[1 - (1-t)^{1/2}] < \infty \text{ for } 0 \leq t < 1 \end{array} \right.$$

Since $E[Z_t^2] < \infty$ for $0 \leq t < 1$, Z_t is square integrable martingale

$$\begin{array}{l} \text{For } t=1, E[Z_t^2] = 2[1-0] = 2 \\ \text{For } t=0, E[Z_t^2] = 2[1-1^{1/2}] = 0 \end{array} \left. \vphantom{\begin{array}{l} \text{For } t=1, E[Z_t^2] = 2[1-0] = 2 \\ \text{For } t=0, E[Z_t^2] = 2[1-1^{1/2}] = 0 \end{array}} \right\} 0 \leq E[Z_t^2] \leq 2 < \infty$$

$$\text{Var}(Z_t) = E[Z_t^2] \leq 2 < \infty$$

$$\text{Var}(Z_t) \leq 2 \rightarrow C=2 //$$

b

Exercise 4.2

$$\alpha=1 \quad \left\{ \begin{array}{l} u=1-s \\ du=-ds \\ ds=-du \end{array} \right.$$
$$Z_t = \int_0^t \frac{1}{1-s} dB_s$$

$$E[Z_t^2] = \int \frac{1}{(1-s)^2} ds = \int u^{-2} (-du) = -\int u^{-2} du = u^{-1} \Big|_{s=0}^{s=t}$$

$$E[Z_t^2] = \frac{1}{1-s} \Big|_{s=0}^{s=t} = \frac{1}{1-t} - \frac{1}{1} = \frac{1}{1-t} - 1$$

$$\text{For } t=0 : E[Z_t^2] = 1-1=0$$

$$\text{For } t=0.999 : E[Z_t^2] = \frac{1}{1-0.999} - 1 = 999$$

as $t \rightarrow 1^-$, $E[Z_t^2] \rightarrow \infty$, meaning Z_t is not square integrable here

Since $E[Z_t^2]$ is between 0 and ∞ for $0 \leq t < 1$,

then with probability one, there will exist $t < 1$ with $Z_t = 1$ //

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7