

Lecture 4: September 14

Lecturer: Csaba Szepesvári

Scribes: Zixin Zhong

Note: \LaTeX template courtesy of UC Berkeley EECS dept. ([link](#) to directory)**Disclaimer:** These notes have not been subjected to the usual scrutiny reserved for formal publications. They may be distributed outside this class only with the permission of the Instructor.

4.1 Outline

1. Concentration inequalities:
Chernoff's inequality, multiplicative Chernoff's inequality; Bennett's inequality, Bernstein inequality
2. PAC-learning:
PAC learnability based on 'fitness'/union bounds

4.2 Concentration inequalities

Theorem 4.1 (Chernoff's Inequality). Let $X_1, \dots, X_n \in [0, 1]$ be i.i.d. random variables, $\bar{X}_n = \frac{1}{n}(X_1 + \dots + X_n)$, $\mu = \mathbb{E}X_1$. We have(a) $\forall \delta \in (0, 1)$, with probability $1 - \delta$,

$$\bar{X}_n \leq \mu + \sqrt{\frac{\log(1/\delta)}{2n}};$$

(b) $\forall \delta \in (0, 1)$, with probability $1 - \delta$,

$$\bar{X}_n \geq \mu - \sqrt{\frac{\log(1/\delta)}{2n}}.$$

Proof. Since $X_1 \in [a, b]$ implies that X_1 is $\sigma(X_1)$ -SG with $\sigma(X_1) = \frac{b-a}{n}$, $X_1 \in [0, 1]$ indicates that

$$\sigma(\bar{X}_n) = \frac{\sigma(X_1)}{\sqrt{n}} = \frac{1}{2\sqrt{n}}.$$

Applying this fact with Hoeffding inequality, the Chernoff's inequality is proven. \square **Theorem 4.2** (Multiplicative Chernoff's Inequality). Let $X_1, \dots, X_n \in [0, 1]$ be i.i.d. random variables, $\bar{X}_n = \frac{1}{n}(X_1 + \dots + X_n)$, $\mu = \mathbb{E}X_1$. We have(a) $\forall \delta \in (0, 1)$, with probability $1 - \delta$,

$$\bar{X}_n \leq \mu + \sqrt{\frac{2\mu \log(1/\delta)}{n}} + \frac{1}{3n};$$

(b) $\forall \delta \in (0, 1)$, with probability $1 - \delta$,

$$\bar{X}_n \geq \mu - \sqrt{\frac{2\mu \log(1/\delta)}{n}}. \quad (*)$$

Remark 4.3.(a) How big can μ be?

By (*): $\mu \leq \bar{X}_n + \sqrt{\frac{2\mu \log(1/\delta)}{n}}$.

(b) Let

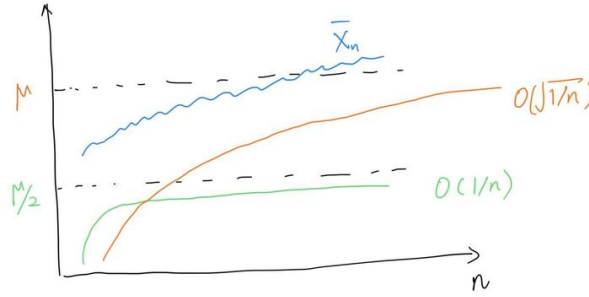
$$f(a, c) = \max\{u : u \leq a + \sqrt{u \cdot c}\}, \text{ where } a = \bar{X}_n, c = \frac{2 \log(1/\delta)}{n}.$$

Then

$$\begin{aligned} \mu + \frac{\log(1/\delta)}{2n} &\geq \sqrt{\frac{2\mu \log(1/\delta)}{n}} \text{ and equality holds when } \mu = \frac{\log(1/\delta)}{2n}, \\ \Rightarrow \inf_{0 < \gamma < 1} \gamma\mu + \frac{\log(1/\delta)}{2\gamma n} &= \sqrt{\frac{2\mu \log(1/\delta)}{n}}, \\ \Rightarrow \mu - \sqrt{\frac{2\mu \log(1/\delta)}{n}} &= \sup_{0 < \gamma < 1} (1 - \gamma)\mu - \frac{\log(1/\delta)}{2\gamma n}. \end{aligned}$$

Let $\gamma = 1/2$, then with (*) we have

$$\bar{X}_n \geq \frac{\mu}{2} - \frac{\log(1/\delta)}{n}.$$

**Figure 4.1:** Example: set $\gamma = 1/2$.

(c) When we apply γ that does not maximize the term $(1 - \gamma)\mu - \frac{\log(1/\delta)}{2\gamma n}$, we cannot claim that we get a better ‘convergence’ rate, because when $n \rightarrow \infty$, $(1 - \gamma)\mu - \frac{\log(1/\delta)}{2\gamma n}$ and \bar{X}_n converges to different values. In detail, \bar{X}_n converges to μ regardless of the value of γ , and $(1 - \gamma)\mu - \frac{\log(1/\delta)}{2\gamma n}$ converges to $(1 - \gamma)\mu \neq \mu$ when $0 < \gamma < 1$.

To say something about the convergence of \bar{X}_n , we need to have the coefficient of μ be 1.

Theorem 4.4 (Bernett’s Inequality). Let X_1, \dots, X_n be i.i.d. random variables. Set $\bar{X}_n = \frac{1}{n}(X_1 + \dots + X_n)$ and $\mu = \mathbb{E}X_1$. If $X_1 - \mu \leq b$, with probability $1 - \delta$, we have

$$\bar{X}_n \leq \mu + \sqrt{\frac{2 \text{Var}(X_1) \log(1/\delta)}{n}} + \frac{b}{3n}.$$

4.3 PAC-learning (L. Valiant)

Let function $f_* : \{0, 1\}^d \rightarrow \{0, 1\}$, $X_1, X_2, \dots, X_n \in \{0, 1\}^d := \underline{2}^d$ be i.i.d. random variables drawn from distribution P_X , data set $D_n = \{(X_1, f_*(X_1)), \dots, (X_n, f_*(X_n))\}$.

Let $f_* \in \mathcal{F} \subset \underline{2}^{\underline{2}^d}$ and $f \in \underline{2}^{\underline{2}^d}$, $P_X^{f_*} := P(X_1, f_* X_1)$, and

$$\begin{aligned} L(f) &= \mathbb{P}(f(X) \neq f_*(X)) = L(P_X^{f_*}, f), \\ l : \underline{2} \times \underline{2} &\rightarrow \underline{2}, \quad l(y, y') = \mathbf{1}(y \neq y'), \\ L(P_X^{f_*}, f) &= \int P(dx, dy) l(f(x), y). \end{aligned}$$

Definition 4.5 (PAC-Learning). Fix $\mathcal{C} = (\mathcal{C}_d)_{d \geq 1}$, where $\mathcal{C}_d \subset \underline{2}^{\underline{2}^d}$. \mathcal{C} is **PAC-learnable** (**Proably Approximately Correctly**) if \exists polynomial $p \in \mathbb{R}[x, y, z]$ and $\mathcal{A} = (\mathcal{A}_{n,d})_{n \geq 1, d \geq 1}$ where $\mathcal{A}_{n,d} : (\underline{2}^d \times \underline{2})^n \rightarrow \underline{2}^{\underline{2}^d}$

$$\begin{aligned} \text{s.t. } \forall \varepsilon \in (0, 1), \delta \in (0, 1), d \geq 1, P \in \mathcal{M}_1(\underline{2}^d), f_* \in \mathcal{C}_d, \\ n \geq \lceil p(1/\varepsilon, 1/\delta, d) \rceil, \\ X_1, X_2, \dots, X_n \sim P_X, \\ f_n = \mathcal{A}_{n,d} \left(\underbrace{(X_1, f_*(X_1)), \dots, (X_n, f_*(X_n))}_{D_n} \right), \end{aligned}$$

we have

$$\mathbb{P} \left(L \left(P_X^{f_*}, f_n \right) \geq \varepsilon \right) \leq \delta.$$

In other words, with probability $1 - \delta$, $\mathbb{P}(f_n(X) \neq f_*(x) | D_n) \leq \varepsilon$.

Remark 4.6. (a) Example:

$$\mathcal{C}_{\text{AND}}^d = \left\{ f : \underline{2}^d \rightarrow \underline{2} \mid \exists u \subset [d], \forall x \in \underline{2}^d : f(x) = \min_{j \in u} X_j \right\}.$$

(b) (i) $L(f_*) = 0$. (ii) When $Y_i = f_*(X_i)$, there is **NO** noise and this will make learning **faster**.