Some finer points concerning inference on one proportion

Confidence intervals for *p*:

critical value closer for level of confidence

• Centered interval approach (Check the rules of thumb before applying)

• Formula
$$SE_{\widehat{p}} = \underbrace{\sqrt{p(1-p)}}_{\sqrt{n}}$$

$$\widehat{p} \pm (z^*)(SE_{\widehat{p}})$$

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propertion

Specifically,

$$SE_{p} = \sqrt{\frac{p(1-p)}{n}}$$

Insurt as P

Note that $SE_{\widehat{p}}$ decreases as n grows. Specifically,

- can cut $SE_{\widehat{p}}$ in half by quadrupling (4×) sample size
- can cut $SE_{\widehat{p}}$ to one third its size at n by increasing sample size by a factor of 9
- other similar statements?
- Can "tailor" a margin of error

Typical issue at election time: Want to estimate p = population proportion χ who will vote for χ using a CI.

Useless-sounding: We estimate

p to be 0.44 ± 0.2

W/ 95% Confidence

Better: we estimate p to be _____ W/M.of E, 0.03 0.44 + 0.03 or 0.75 ± 0.03

better m. of E

Q: Can I decide ahead of time what sample site in is sufficient to give one a m. it error no bigger than U.O.3 (or some specified anount)?

M. of error =
$$(Z^*)$$
 SE
 $0.03 \ge ME = (Z^*) \frac{\sqrt{p(1-p)}}{\sqrt{n}}$

Muttiply by In, divide by 0.03:

$$\sqrt{n} \geq \left(z^*\right) \frac{\sqrt{p(1-p)}}{0.03}$$

From calculus, a fact about numerator is that it con't ever get larger than 0.5.

So, ve cen replace numerator v/ 1/2:

$$\sqrt{n} \geq \left(z^*\right) \frac{0.5}{0.03}$$

Now pick 2* for Jesived level of confidence

$$95\% \text{ conf.}$$
: $2^{*} = 1.96$
 $90\% \text{ conf.}$: $2^{*} = 1.645$

$$\sqrt{n} \geq \left(1.96\right) \frac{.5}{0.03} \qquad \Longrightarrow \qquad n \geq \left[1.96 \cdot \frac{0.5}{0.03}\right]$$

= 1067.111

Hypothesis tests involving null hypothesis \mathbf{H}_0 : $p = p_0$:

- several **test statistics** one might use from the sample
 - 1. the sample proportion \widehat{p}
 - natural to use, as an unbiased estimator of p
 - approximating normal distribution: Norm $(p_0, \sqrt{p_0(1-p_0)/n})$
 - 2. the count of successes *X*
 - natural to use, as an unbiased estimator of p
 - approximating normal distribution $(Norm(np_0, \sqrt{np_0(1-p_0)}))$
 - 3. the Z-score of either of the previous
 - called the **standardized test statistic**
 - approximating normal distribution: Norm(0,1)
 - one advantage: immediate comparison with critical z^* value

Ho: p=1/2

On Monday: hyp. fest for home-won soccer genes

X = 70 times in n = 120 games the home team won

 $7. \quad \chi = 70$

1. $\hat{p} = 70/120$ has roughly a normal

Norm (2, (1/2) = 0.0456

hus roughly a normal dist

Norm (120. 2) (120(12)(1-12))

= 5.477 ~ (°D

3. Z-SLore from P.

$$=\frac{70/120}{0.0456}$$

= 1.83

from X:

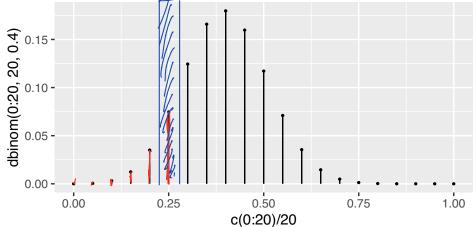
= 183

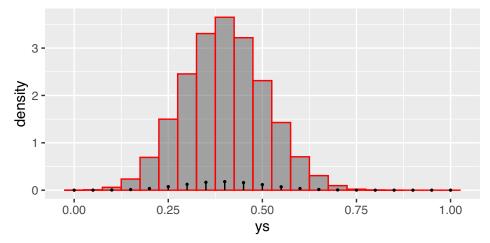
has Norm (D ,)

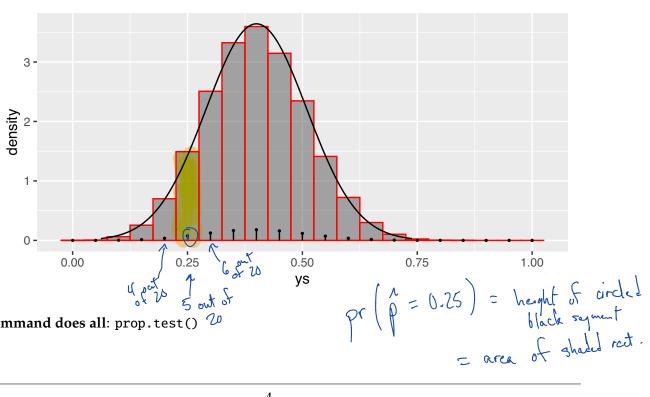
continuity correction

Here is the null distribution that goes with $H_0: p = .4$ when n = 20:A normal model for single proportion replaces heights U areas

5 cg 0.25







one-command does all: prop.test()

Pr (3=0.25) =

btucen 5.5 1 6.5

Monday: $H_0: p = 0.5$ is. $H_a: p > 0.5$ $\hat{p} = \frac{70}{120} = 1.583$ Usil wormed dist Norm (0.5, 0.0456) 1,5t side 70.5 is right side 120 se rectage $arca = 1 - pnovm \left(0.583, 0.5, 0.0456\right)$ better approach using continuity correction 1 - provin (69.5)

e a