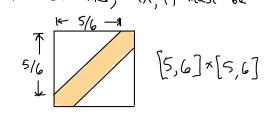
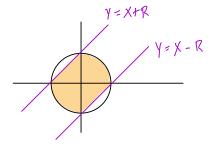
(c) For them to arrive within 10 minutes of each other, (X, Y) must be a point in the orange shaded region, The 5/6 - which has area (= probability) 5/6 $1-(2)(\frac{1}{2})(\frac{1}{2})^2 = \frac{11}{2}$



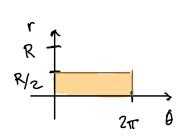
3.57 (a) Let $A = \{(x,y) \mid x^2 + y^2 \leq R^2\}$. For the pdf to be a constant k, we require $1 = k \cdot Area(A) = \pi R^2 k \implies k = \frac{1}{\pi R^2}$ $f(x,y) = \begin{cases} \frac{1}{\pi R^2}, & (x,y) \in A, \\ 0, & \text{otherwise} \end{cases}$

(b) Let
$$H = \left\{ (x,y) \mid x^2 + y^2 \le \frac{R}{2} \right\}$$
. $P_r((x,y) \in H) = \frac{Area(H)}{Area(A)} = \frac{\pi \left(\frac{R}{2} \right)^2}{\pi R^2} = \frac{1}{4}$.

(c) $|X-Y| \leq R \iff X-R \leq Y \leq X+R$ The desired probability corresponds to the area of the orange shaded region: $\left(\frac{1}{2}\pi R^2 + R^2\right) \cdot \frac{1}{\pi R^2} = \frac{1}{2} + \frac{1}{\pi}.$



- (d) Let $x \in [-R, R]$. Then $f_{X(x)} = \int_{X}^{\infty} f_{X,Y}(x,y) dy = \int_{X}^{R^{2}-x^{2}} \frac{1}{\pi R^{2}} dy = \frac{2}{\pi R^{2}} \sqrt{R^{2}-x^{2}}$
- (e) X and Y are not independent. By symmetry, $f_{\gamma}(y) = \frac{2}{\pi R^2} \sqrt{R^2 y^2}$ for $y \in [-R, R]$. But $f_{\chi(x)}f_{\gamma(y)} = \frac{4}{\pi^2 R^4} \sqrt{(R^2 - \chi^2)(R^2 - y^2)} \neq \frac{1}{\pi R^2}$
- 3.58 This gives a different distribution. Under this approach, Pr ((x,y) & H) = 1/2TR · Area (shaded region of ro-plane) $= \frac{1}{2\pi R} \cdot \left(2\pi \left(\frac{1}{2}R\right) = \frac{1}{2}$



a different result than in part (b) of Exercise 3.57.

3.63
$$\times \sim \text{Pois}(\lambda_1)$$
 and $\times \sim \text{Pois}(\lambda_2)$, so $M_{\chi}(t) = e^{t\lambda_1 - \lambda_1}$, and $M_{\chi}(t) = e^{t\lambda_2 - \lambda_2}$. By independence of \times , \times .

$$M_{X+Y}(t) = M_{X}(t)M_{Y}(t) = e^{e^{t}\lambda_{1}} \cdot e^{-\lambda_{1}} \cdot e^{e^{t}\lambda_{2}} \cdot e^{-\lambda_{2}}$$

$$= e^{e^{t}\lambda_{1} + e^{t}\lambda_{2}} \cdot e^{-(\lambda_{1} + \lambda_{2})} = e^{e^{t}(\lambda_{1} + \lambda_{2}) - (\lambda_{1} + \lambda_{2})}$$

This is the mgf for another Poisson r.v. with parameter λ , $\pm \lambda_z$. Thus $X + Y \sim \text{Pois}(\lambda, \pm \lambda_z)$.

3.67 Each
$$\chi_i \sim Genma(\alpha, \lambda)$$
. By independence,
$$M_S(t) = \prod_{i=1}^{\infty} \left(\frac{\lambda}{\lambda - t}\right)^{\alpha} = \left(\frac{\lambda}{\lambda - t}\right)^{n\alpha} \implies S \sim Gomma(n\alpha, \lambda).$$

$$M_{\widetilde{X}}(t) = M_{\frac{1}{n}}S(t) = M_{S}(\frac{1}{n}t) = \left(\frac{\lambda}{\lambda - t/n}\right)^{n\alpha} = \left(\frac{n\lambda}{n\lambda - t}\right)^{n\alpha} \Rightarrow \widetilde{X} \wedge Gamma(n\alpha, n\lambda).$$

4.1 Let $X = \frac{1}{n} \sum X_i$ be the first sample moment about the origin (a.k.a. the sample mean). The population mean for Binom (1, π) is $1 \cdot \pi = \pi$. Our estimate is $\hat{\pi} = X$.

4.4 For
$$X \sim NBinom$$
, $E(X) = \frac{\Delta}{\pi} - \Delta$. So, we set
$$\frac{\Delta}{\hat{\pi}} - \Delta = \overline{X} \qquad \Longrightarrow \qquad \hat{\pi} = \frac{\Delta}{\Delta + \overline{X}}.$$

4.7 favstats reveals X = 0.6091, S = 0.248, $n = 134 \implies v = 0.06105$.

From the formulas:

$$\hat{\alpha} = \overline{x} \left(\frac{\overline{x}(1-\overline{x})}{\sigma} - 1 \right) \doteq 1.7665$$

$$\hat{\beta} = \left(1 - \overline{x} \right) \left(\frac{\overline{x}(1-\overline{x})}{\sigma} - 1 \right) \doteq 1.1337$$

The beta distribution using these shape parameters gives a very poor fit to the data. By filtering out the players with FTPct = 0, the new parameter estimates from remaining players are

$$\hat{a} = 4.824, \quad \hat{\beta} = 2.387,$$

and the fit is vastly improved.

4.9 Take
$$\bar{\chi} = \frac{1}{n} \sum_{i} \sum_{j=1}^{n} \sum_{i} \sum_{j=1}^{n} \sum_{j=1}^{n} \sum_{i} \sum_{j=1}^{n} \sum_{j=1}^{n} \sum_{i} \sum_{j=1}^{n} \sum_{i} \sum_{j=1}^{n} \sum_{i} \sum_{j=1}^{n} \sum_{j=1}^{n} \sum_{i} \sum_{j=1}^{n} \sum_{i} \sum_{j=1}^{n} \sum_{i} \sum_{j=1}^{n} \sum_{j=1}^{n} \sum_{i} \sum_{j=1}^{n} \sum_{j=1}^{n} \sum_{j=1}^{n} \sum_{i} \sum_{j=1}^{n} \sum_{j$$

$$\frac{\hat{\lambda}}{\hat{\lambda}} = \bar{\lambda}$$
 \Rightarrow $\hat{\lambda} = \hat{\lambda} \bar{x}$

$$\frac{\hat{\lambda}}{\hat{\lambda}^2} = V \implies \frac{\hat{\lambda} \times}{\hat{\lambda}^2} = \frac{\overline{\lambda}}{\hat{\lambda}} = V$$

$$\Rightarrow \hat{\lambda} = \frac{\overline{x}}{x} = 1.858 \quad \text{and} \quad \hat{\lambda} = \hat{\lambda} = \frac{\overline{x}}{x} = 0.0141.$$

$$4.14$$
 We know $SE_{\overline{\chi}} = \frac{\sigma}{\sqrt{n}} = \frac{10}{\sqrt{25}} = 2$.

Thus,
$$Pr(|X-\mu|<3) = pnorm(3,0,2) - pnorm(-3,0,2)$$

= 0.866.

4.16 (a) The 10 different SRS, along w/resulting sample means:

$$\Rightarrow \bigvee_{X} = (3.5 + 7 + 7.5)(2/10) + (4.5 + 5 + 6 + 8.5)(1/10) = 6.$$

$$V_{ar}(\overline{X}) = (3.5^{2} + 7^{2} + 7.5^{2})(2/10) + (4.5^{2} + 5^{2} + 6^{2} + 8.5^{2})(1/10) - \bigvee_{X}^{2}$$

$$= 2.85.$$

(b) From Coro. 4.3.3, we have
$$E(\overline{X}) = \mu = (1+6+6+8+9)(1/5) = 6,$$
a match with part (a). Furthermore,
$$Var(\overline{X}) = \frac{\sigma^2}{n} \cdot \frac{N-n}{N-1} = \frac{\sigma^2}{2} \cdot \frac{5-2}{5-1} = \frac{3}{8} \sigma^2,$$
where
$$\sigma^2 = (1^2+6^2+6^2+8^2+9^2)(1/5)-6^2 = 7.6.$$
 So,
$$Var(\overline{X}) = \frac{3}{8} \sigma^2 = 2.85, \text{ also matching part (a-)}.$$

(c) We may treat an iid sample as if we were rolling 5-sided dice, yielding pairings:
$$(1,1)$$
 $(1,6)$ $(1,6)$ $(1,8)$ $(1,9)$ $(6,1)$ $(6,6)$ $(6,6)$ $(6,8)$ $(6,9)$ $(6,1)$ $(6,6)$ $(6,6)$ $(6,8)$ $(6,9)$ $(8,1)$ $(8,6)$ $(8,6)$ $(8,8)$ $(8,9)$ $(9,9)$

The 5-by-5 table of means corresponds directly to these pairings

$$S_{D_{1}}$$

$$\mu = (1+8+9)(1/25) + (4.5+5+8.5)(2/25) + (3.5+6+7+7.5)(4/25) = 6,$$
and
$$\sigma^{2} = (1^{2}+8^{2}+9^{2})(1/25) + (4.5^{2}+5^{2}+8.5^{2})(2/25) + (3.5^{2}+6^{2}+7^{2}+7.5^{2})(4/25) - 6^{2}$$

$$= 3.8.$$

$$4.39 (a) |\vec{u}_{1}|^{2} = (\frac{1}{\sqrt{n}})^{2} + \cdots + (\frac{1}{\sqrt{n}})^{2} = n \cdot (\frac{1}{n}) = | \implies |\vec{u}_{1}| = 1.$$

For
$$i = 2, 3, ..., \%$$
,
$$|\vec{u}_{i}|^{2} = \frac{1}{i(i-1)} \left[(i-1)^{2} + \sum_{j=1}^{i-1} |^{2} \right] = \frac{1}{i(i-1)} \left[(i-1)^{2} + (i-1) \right]$$

$$= \frac{i-1}{i(i-1)} \left[(i-1) + 1 \right] = \frac{1}{i} (i) = 1 \implies |\vec{u}_{i}| = 1.$$

Moreover, for
$$|\langle i \langle j \rangle |$$

 $\vec{u}_i \circ \vec{u}_j = |-i| + \sum_{m=1}^{i-1} | = (|-i|) + (i-1) = 0.$

It is more transparent that each $\vec{u}_i \circ \vec{u}_i = 0$, i > 1.

(b) Let
$$\vec{X} = \langle x_1, x_2, ..., x_n \rangle$$
. Then,

$$\vec{X} \cdot \vec{u}_1 = \frac{1}{\sqrt{n}} (X_1 + X_2 + ... + X_n) = \sqrt{n} \cdot \frac{1}{n} \sum_{i=1}^{n} X_i = \sqrt{n}.$$

(c) With
$$\vec{x}$$
 as in part (b), $(\vec{x} \circ \vec{u},)\vec{u} = (\vec{x} \sqrt{n}) \cdot \frac{1}{\sqrt{n}} \langle 1, 1, ..., 1 \rangle = \langle \vec{x}, \vec{x}, ..., \vec{x} \rangle$.

(d) For
$$\vec{\chi} = \langle 3, 4, 4, 7, 7 \rangle$$
, $\vec{\chi} = \frac{1}{5}(25) = 5$, and $\vec{\tau} = \langle 3, 4, 4, 7, 7 \rangle - \langle 5, 5, 5, 5, 5 \rangle = \langle -2, -1, -1, 2, 2 \rangle$.

Thus,
$$\frac{1}{1}$$
 $\frac{1}{1/\sqrt{5}} < 1, 1, 1, 1, 1 > \frac{1}{5^{3/2}} = \frac{1}{1/\sqrt{5}} = \frac$

For $i \geq 2$, $\vec{x} \cdot \vec{u}_i = \vec{v} \cdot \vec{u}_i$. This is true for other $\vec{x} \in \mathbb{R}^5$, too, since $\vec{v} \cdot \vec{u}_i = (\vec{x} - \vec{x}) \cdot \vec{u}_i = \vec{x} \cdot \vec{u}_i - \vec{x} \cdot \vec{u}_i = \vec{x} \cdot \vec{u}_i = 0$

since \$\overline{\pi}\$, being parallel to \$\overline{\pi}\$, is orthogonal to each \$\overline{\pi}\$; with \$i\ge 2\$.

4.40 (a)
$$\overline{X} = (3+4+5+8)/4 = 5$$

$$S^{2} = \frac{1}{3} \left[(3-5)^{2} + (4-5)^{2} + (5-5)^{2} + (8-5)^{2} \right] = \frac{14}{3}.$$

(b)
$$\vec{p}_1 = \langle 5, 5, 5, 5 \rangle$$
, as determined in Exercise 4.39.
 $\vec{p}_2 = \frac{1}{\sqrt{2}} (3-4) \vec{v}_2 = \langle -\frac{1}{2}, \frac{1}{2}, 0, 0 \rangle$
 $\vec{p}_3 = \frac{1}{\sqrt{6}} (3+4-10) \vec{v}_3 = \langle -\frac{1}{2}, -\frac{1}{2}, 1, 0 \rangle$

$$\vec{P}_{4} = \frac{1}{\sqrt{12}} (3+4+5-24) \vec{u}_{4} = \langle -1, -1, -1, 3 \rangle$$

and
$$\sum \vec{p}_i = \langle 3, 4, 5, 8 \rangle$$
 as predicted.

(c)
$$l_1 = |\vec{p}_1| = \sqrt{4(5^1)} = 10$$

$$l_2 = |\vec{p}_2| = \sqrt{2(\frac{1}{2})^2} = \frac{1}{\sqrt{2}}$$

$$l_3 = |\vec{p}_3| = \sqrt{2(\frac{1}{2})^2 + 1^2} = \sqrt{\frac{3}{2}}$$

$$l_4 = |\vec{p}_4| = \sqrt{3(-1)^2 + 3^2} = 2\sqrt{3}$$
(d) $\sum_{i=3}^{4} l_i^2 = \frac{1}{2} + \frac{3}{2} + 12 = 14 = 3(\frac{14}{3}) = 3s^2$.

C.10
$$\text{proj}(\langle 1,2,3\rangle \to \langle 1,1,1\rangle) = \frac{\langle 1,2,3\rangle \circ \langle 1,1,1\rangle}{|\langle 1,1,1\rangle|^2} \langle 1,1,1\rangle = \langle 2,2,2\rangle$$
This is the closest vector to $\langle 1,2,3\rangle$ in $W = \text{span}(\langle 1,1,1\rangle)$, so $\langle 1,2,3\rangle - \text{proj}(\langle 1,2,3\rangle \to \langle 1,1,1\rangle) = \langle -1,0,1\rangle$
is purpudicular/orthogonal to $\langle 1,1,1\rangle$, and $\text{span}\{\langle 1,1,1\rangle,\langle 1,2,3\rangle\} = \text{span}\{\langle 1,1,1\rangle,\langle -1,0,1\rangle\}$.

To get these orthogonal vectors to be of length 1, we divide by their lengths:

$$\vec{u}_1 = \frac{1}{\sqrt{3}} \langle 1, 1, 1 \rangle = \langle \frac{1}{\sqrt{13}}, \frac{1}{\sqrt{13}}, \frac{1}{\sqrt{13}} \rangle, \quad \vec{u}_2 = \frac{1}{\sqrt{2}} \langle -1, 0, 1 \rangle = \langle -\frac{1}{\sqrt{2}}, 0, \frac{1}{\sqrt{2}} \rangle.$$