

Stat 343, Fri 9-Oct-2020 -- Fri 9-Oct-2020
 Probability and Statistics
 Fall 2020

If A, B ind. counts
 then A^c and B are, too.

$$\text{Var}(aX) = a^2 \text{Var}(X)$$

Friday, October 09th 2020

Wk 6, Fr

Topic:: Moment generating functions (mgf)

To each sequence a_0, a_1, a_2, \dots
 there is a corresponding power series centered
 at 0
 $a_0 + a_1 t + a_2 t^2 + \dots = \sum_{j=0}^{\infty} a_j t^j$

Power series

$$\sum_{j=0}^{\infty} c_j t^j = c_0 + c_1 t + c_2 t^2 + \dots$$

Generating Functions

Power series

Have seq. $a_0, a_1, a_2, a_3, \dots$

Notes and reminders:

- The appearance of any power series centered at $t = a$ is ...
- Radius and interval of convergence
- Any function f that is differentiable to arbitrary order at a point $t = a$ has a formal power series at $t = a$, called its **Taylor series**:

$$\sum_{k=0}^{\infty} \frac{f^{(k)}(a)}{k!} (x - a)^k.$$

There are things we would like to be true, but are not generally:

- It is *not* generally true that the domain (interval of convergence) of the power series is the same as that for f .
- When t is in both domains, the value to which the power series converges need not be the same as $f(t)$.

Despite the uncertainties of such facts, there are some pairings of functions with their Taylor series about which we have a good understanding of when they are equal:

$$\frac{1}{1-t} = \sum_{k=0}^{\infty} t^k, \quad -1 < t < 1$$

$$e^x = \sum_{k=0}^{\infty} \frac{x^k}{k!}, \quad -\infty < x < \infty$$

- Term-by-term differentiation

Generating functions

Want to generate a sequence
 Given r.v. X , $\mu_1, \mu_2, \mu_3, \mu_4, \dots$

Definition 1: Let $(a_k)_{k=0}^{\infty}$ be a sequence. The (formal) power series

$$A(t) = \sum_{k=0}^{\infty} a_k t^k$$

is called the **ordinary generating function** for (a_k) . Another power series,

$$B(t) = \sum_{k=0}^{\infty} \frac{a_k}{k!} t^k,$$

is called the **exponential generating function** for (a_k) .

Example 1: Generating functions for exponential sequences

For $r \neq 0$, let $a_k = r^k$, $k = 0, 1, 2, \dots$. Then generating functions for (a_k) have well-known closed-form expressions. In particular, the ordinary generating function for (a_k) is

$$A(t) = 1 + rt + (rt)^2 + (rt)^3 + \dots + (rt)^k + \dots = \frac{1}{1 - rt},$$

and the exponential generating function is

$$B(t) = 1 + rt + \frac{(rt)^2}{2!} + \dots + \frac{(rt)^k}{k!} + \dots = e^{rt}.$$

■

As these generating functions are power series centered at 0,

$$A(t) = \sum_{k=0}^{\infty} a_k t^k \quad \text{and} \quad B(t) = \sum_{k=0}^{\infty} b_k t^k,$$

with each $b_k = a_k/k!$, we have the relationship between coefficients and derivatives of $A(t)$:

$$a_k = \frac{A^{(k)}(0)}{k!} \quad \text{and} \quad b_k = \frac{A^{(k)}(0)}{(k!)^2}, \quad k = 0, 1, 2, \dots$$

Similarly,

$$b_k = \frac{B^{(k)}(0)}{k!} \quad \text{and} \quad a_k = B^{(k)}(0), \quad k = 0, 1, 2, \dots$$

Moment generating functions

Definition 2: Let X be a discrete or continuous random variable, and define the function $M_X(t) := E(e^{tX})$. When this function exists for some nontrivial interval $|t| \leq b$ about the origin, we call it the **moment generating function** (or **mgf**) of X .

Why is M_X called by this name?

- **Case: X is discrete.**

(X might be Pois, Binom, DUnif) $(\mu_0, \mu_1, \mu_2, \dots)$
 so it has a pmf $f_X(x)$

$$\begin{aligned} E(e^{tX}) &= \sum_x e^{tx} f_X(x) = \sum_x f_X(x) \sum_{j=0}^{\infty} \frac{(tx)^j}{j!} = \sum_{j=0}^{\infty} \sum_x \frac{(tx)^j}{j!} f_X(x) \\ &= \sum_{j=0}^{\infty} \frac{t^j}{j!} \underbrace{\sum_x x^j f_X(x)}_{\mu_j} = \sum_{j=0}^{\infty} \frac{\mu_j}{j!} t^j \end{aligned}$$

- **Case: X is continuous.**

$$\begin{aligned} E(e^{tX}) &= \int_{-\infty}^{\infty} e^{tx} f_X(x) dx = \dots \\ &= \sum_{j=0}^{\infty} \frac{\mu_j}{j!} t^j \end{aligned}$$

In either case, if the sequence $(\mu_k)_{k=1}^{\infty}$ of k^{th} moments of X about the origin exists, then $M_X(t)$ is the exponential generating function for the sequence of (μ_k) .

$$\mu_0, \mu_1, \mu_2, \mu_3, \dots$$

Then

$$\underline{E(e^{tX})} = M_X(t) \quad \text{mgf}$$

$$e^{bx} = (e^t)^x$$

Example 2: Moment generating function for $X \sim \text{Pois}(\lambda)$

$$\Rightarrow \text{pmf } f_X(x) = e^{-\lambda} \frac{\lambda^x}{x!}$$

To calculate mgf

$$\begin{aligned} M_X(t) &= E(e^{tx}) = \sum_{x=0}^{\infty} e^{tx} \cdot e^{-\lambda} \frac{\lambda^x}{x!} = e^{-\lambda} \sum_{x=0}^{\infty} \underbrace{\frac{(\lambda e^t)^x}{x!}}_{e^{\lambda e^t}} \\ &= e^{-\lambda} \cdot e^{\lambda e^t} = e^{\lambda e^t - \lambda} \\ &= e^{-\lambda(1 - e^t)} \end{aligned}$$

Example 3: Moment generating function for $X \sim \text{Binom}(n, \pi)$

Here,

$$\begin{aligned} M_X(t) &= \sum_{x=0}^n e^{tx} \binom{n}{x} \pi^x (1-\pi)^{n-x} = \sum_{x=0}^n \binom{n}{x} (e^t \pi)^x (1-\pi)^{n-x} \\ &= (\pi e^t + 1 - \pi)^n. \end{aligned}$$

Example 4: Moment generating function for $X \sim \text{Unif}(a, b)$

Since X is a continuous random variable, we have

$$M_X(t) = \int_a^b \frac{e^{tx}}{b-a} dx = \frac{e^{tx}}{t(b-a)} \Big|_a^b = \frac{e^{bt} - e^{at}}{t(b-a)}.$$

Example 5: Moment generating function for $X \sim \text{Exp}(\lambda)$

Here,

$$\begin{aligned} M_X(t) &= \int_0^{\infty} e^{tx} \lambda e^{-\lambda x} dx = \lim_{A \rightarrow \infty} \int_0^A \lambda e^{(t-\lambda)x} dx = \lim_{A \rightarrow \infty} \frac{\lambda}{t-\lambda} e^{(t-\lambda)x} \Big|_0^A \\ &= \lim_{A \rightarrow \infty} \frac{\lambda}{t-\lambda} (e^{A(t-\lambda)} - 1) = \frac{\lambda}{\lambda - t} = \frac{1}{1 - \frac{t}{\lambda}}, \end{aligned}$$

if $t < \lambda$ (so, in particular, the limit exists for $|t| < \lambda$).



Theorem 1: For X a discrete or continuous r.v., a, b constants,

- (i) $M_{aX}(t) = M_X(at)$.
- (ii) $M_{X+b}(t) = e^{bt}M_X(t)$.
- (iii) $M_{aX+b}(t) = e^{bt}M_X(at)$.

Two applications of mgfs:

- Since $M_X(t)$ is the exponential generating function of the sequence (μ_k) of k^{th} moments, we have, in general, $\mu_k = M_X^{(k)}(0)$.
- mgfs are unique, providing a signature for recognizing a random variable, as specified in this theorem:

Theorem 2: Let X, Y be r.v.s with mgfs M_X, M_Y . Then X, Y are **identically distributed** (i.e., they have the same cdfs) if and only if $M_X(t) = M_Y(t)$ for all t in some nontrivial interval containing 0.

Say have seq.

1, 1, 1, 1, 1, 1, ...

Corresp. power series

$$1 + 1 \cdot t + 1 \cdot t^2 + 1 \cdot t^3 + \dots = \sum_{j=0}^{\infty} t^j$$

geometric series

$$= \frac{1}{1-t} \leftarrow \text{somehow a generator of orig. seq.}$$

Since $f(t) = \frac{1}{1-t}$ produces Maclaurin series

$$\sum_{j=0}^{\infty} \frac{f^{(j)}(0)}{j!} t^j$$

$$f(0) = 1$$

$$\begin{aligned} f'(t) &= \frac{d}{dt} (1-t)^{-1} = -1(1-t)^{-2}(-1) \\ &= \frac{1}{(1-t)^2} \end{aligned}$$

$$\Rightarrow f'(0) = 1$$

$$f''(t) = -2(1-t)^{-3}(-1) = \frac{2!}{(1-t)^3}$$

$$\Rightarrow \frac{f''(0)}{2!} = 1$$

If we had gone with the exponential generating fn approach

1, 1, 1, 1, 1, ...

$$\frac{\text{Corresp. Series}}{(\text{modified})} \quad \frac{1}{0!} + \frac{1}{1!}t + \frac{1}{2!}t^2 + \frac{1}{3!}t^3 + \dots$$

$$= e^t$$