# Test for Association of Categorical Variables

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You may click here to access the .qmd file.

# The Big Picture

This test uses the chi-square statistic

$$\chi^2 = \sum_i \frac{(O_{i,j} - E_{i,j})^2}{E_{i,j}},$$

and quite closely resembles the goodness-of-fit test. The differences include

- Goodness-of-fit uses univariate data; the observed counts come from a frequency table of that variable. The test for association involves bivariate data; the observed counts come from a two-way table.
- The expected counts for the test for association are found using the totals in the margins of the two-way table. Specifically, the expected count in row i, column j is

$$E_{i,j} \, = \, \frac{(\text{total of row } i) \times (\text{total of column } j)}{\text{grand total}}.$$

• If the rule of thumb (all  $E_{i,j} \ge 5$ ) is met, so that you may obtain a P-value using a theoretical chi-square distribution, you select the one with

$$df = [\#(\text{rows}) - 1] \times [\#(\text{columns}) - 1].$$

for choosing the degrees of freedom.

We start with a simple example.

**Example**: Is there an association between the dominant hand of a child and that of the child's father?

Step 1: State hypotheses

**H**<sub>0</sub>: The dominant hand of a child and that of the child's father are independent.

 $\mathbf{H}_a$ : There is an association between the dominant hand of a child and that of the child's father.

Step 2: Compute a test statistic

We have data on these variables in the survey results found in the file

```
ssurv <- read.csv("https://scofield.site/teaching/data/csv/ssurv.csv")
names(ssurv) # output has been suppressed
tally(selfhandedness ~ dadhandedness, data=ssurv) # summary two-way table</pre>
```

#### dadhandedness

```
      selfhandedness
      L
      R

      1
      0
      0

      L
      0
      5
      26

      R
      2
      23
      223
```

Some respondents offer missing data, and the next line uses filter() to clean the data up a bit, giving us a better two-way table. I've used a pipe to addmargins() to give us marginal totals.

#### dadhandedness

The two-way table gives us four observed counts:

$$O_{1,1}=5, \quad O_{1,2}=26, \quad O_{2,1}=23, \quad O_{2,2}=223.$$

We get the corresponding four expected counts using the formula above:

$$E_{1,1} = \frac{(28)(31)}{277} = 3.13, \quad E_{1,2} = \frac{(249)(31)}{277} = 27.87, \quad E_{2,1} = \frac{(28)(246)}{277} = 24.87, \quad E_{2,2} = \frac{(249)(246)}{277} = 221.13.$$

Note that one is less than 5.

We can now compute the test statistic:

$$\chi^2 = \frac{(5-3.13)^2}{3.13} + \frac{(26-27.87)^2}{27.87} + \frac{(23-24.87)^2}{24.87} + \frac{(223-221.13)^2}{221.13} \doteq 1.399.$$

If we build lists in R from these numbers, we can use R to do this calculation:

```
obs = c(5, 26, 23, 223)
expected = c(3.13, 27.87, 24.87, 221.13)
sum((obs-expected)^2/expected)
```

[1] 1.399113

### **Step 3**: Compute a *P*-value

Because one expected count is under 5, we should use simulation to obtain a P-value. In class, I recommended another app, found at <a href="https://www.lock5stat.com/StatKey/advanced\_association/advanced\_association.html">https://www.lock5stat.com/StatKey/advanced\_association/advanced\_association.html</a>. Using that app, the approximate P-value is 0.328. Had we blundered forward and used a theoretical  $\chi^2$  distribution with df=1, we would have obtained a noticeably different P-value:

```
1 - pchisq(1.393, df=1)
```

[1] 0.2378991

### Step 4: Draw a conclusion.

At any of the usual significance levels, alpha = 0.1,  $\alpha = 0.05$ , or  $\alpha = 0.01$ , we would fail to reject the null hypothesis. This data is consistent with the two variables having independence.

# Using the chisq.test() command

The command is meant as a short-cut to the various calculations involved in testing for association between categorical variables. It requires you to supply it with the two-way table, which means you have to build that table (without marginal totals). So, in order to use it on the dominant-hand data above:

#### dadhandedness

```
selfhandedness L R
L 5 26
R 23 223
```

```
chisq.test(domHandTable)
```

Warning in chisq.test(domHandTable): Chi-squared approximation may be incorrect

Pearson's Chi-squared test with Yates' continuity correction

data: domHandTable
X-squared = 0.74638, df = 1, p-value = 0.3876

There are several things to take note of, here.

- 1. The most glaring thing is the warning: "Chi-squared approximation may be incorrect." This indicates that it obtained its *P*-value from a chi-square distribution, but it feels guilty about doing so, as not all expected counts were high enough.
- 2. The  $\chi^2$  statistic reported here is 0.74638. But we calculate it above to be 1.393. This discrepancy is due to "Yates' continuity correction." If we run, instead, the command

```
chisq.test(domHandTable, correct=FALSE) # I've suppressed the output
```

then the test statistic (as well as the corresponding \$P\$-value), will match calculations we have performed above.

3. We can use the dollar-sign notation to make filter down results of the command to things we may want particularly.

Warning in chisq.test(domHandTable): Chi-squared approximation may be incorrect

```
dadhandedness
```

```
selfhandedness L R
L 3.133574 27.86643
R 24.866426 221.13357
```

```
chisq.test(domHandTable)$statistic  # produces the chi-square statistic
```

Warning in chisq.test(domHandTable): Chi-squared approximation may be incorrect

X-squared 0.7463755

```
chisq.test(domHandTable)$p.value  # produces the P-value
```

Warning in chisq.test(domHandTable): Chi-squared approximation may be incorrect

### [1] 0.3876262

4. Finally, one can get chisq.test() to forego using a theoretical chi-square distribution and, in lieu of that, compute the *P*-value directly from a simulation. One simply adds the simulate.p.value switch:

```
chisq.test(domHandTable, simulate.p.value=TRUE)

Pearson's Chi-squared test with simulated p-value (based on 2000 replicates)

data: domHandTable
X-squared = 1.3925, df = NA, p-value = 0.3538
```

That comprises the essentials for chi-square tests for an association. There are still some scenarios where extra tips may be helpful. Below, I have included a few.

## Extra R Tips

### Building a two-way table when you only have frequency data

The website https://www.datacamp.com/tutorial/contingency-analysis-r gives a two-way table, already built, offering sport-choices by gender. When you have a ready table, but not yet available in R, you can build it directly. I build the one from the website above using commands like these:

```
sexAndSportTable <- rbind( c(35, 15, 50), c(10, 30, 60) )
rownames(sexAndSportTable) <- c('Female', 'Male')
colnames(sexAndSportTable) <- c('Archery', 'Boxing', 'Cycling')
sexAndSportTable</pre>
```

	Archery	Boxing	Cycling
${\tt Female}$	35	15	50
Male	10	30	60

Such a table can be handed directly to the chisq.test() command for its various computations (the results of which I suppress here):

```
chisq.test(sexAndSportTable)$expected
chisq.test(sexAndSportTable)$statistic
chisq.test(sexAndSportTable)$p.value
```

```
chisq.test(sexAndSportTable)
```

A more round-about method that has its uses is to build a raw data frame containing this data, and then using tally:

```
sexAndSportRawData <- rbind(</pre>
    do(35) * c(sex = "Female", sport="Archery"),
    do(15) * c(sex = "Female", sport="Boxing"),
    do(50) * c(sex = "Female", sport="Cycling"),
    do(10) * c(sex = "Male", sport="Archery"),
    do(30) * c(sex = "Male", sport="Boxing"),
    do(60) * c(sex = "Male", sport="Cycling")
  head(sexAndSportRawData)
           sport
     sex
1 Female Archery
2 Female Archery
3 Female Archery
4 Female Archery
5 Female Archery
6 Female Archery
  tally(sex ~ sport, data=sexAndSportRawData)
        sport
         Archery Boxing Cycling
sex
              35
  Female
                     15
                              50
  Male
              10
                     30
                              60
```