

$$p(\Theta|D) = \frac{p(D|\Theta)p(\Theta)}{p(D|\Theta)p(\Theta) + p(D|\neg\Theta)p(\neg\Theta)}$$

Bayesian Learning 732A46: Lecture 2

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- ▶ The Poisson model
- ▶ Conjugate priors
- ▶ Prior elicitation
- ▶ Non-informative priors

The Poisson model with a Gamma prior

► Model:

$$y_1, \dots, y_n | \theta \stackrel{iid}{\sim} \text{Poisson}(y_i | \theta) = \frac{1}{y_i!} \theta^{y_i} \exp(-\theta), \quad \theta > 0.$$

► Likelihood

$$p(y|\theta) = \prod_{i=1}^n p(y_i|\theta) \propto \theta^{\sum_{i=1}^n y_i} \exp(-\theta n),$$

► Prior

$$p(\theta) \propto \theta^{\alpha_0 - 1} \exp(-\theta \beta_0) \propto \text{Gamma}(\theta | \alpha_0, \beta_0)$$

Interpretation: contains the info: $\alpha_0 - 1$ counts in β_0 observations.

► Posterior

$$\begin{aligned} p(\theta|y) &\propto \left[\prod_{i=1}^n p(y_i|\theta) \right] p(\theta) \\ &\propto \theta^{\sum_{i=1}^n y_i} \exp(-\theta n) \theta^{\alpha_0 - 1} \exp(-\theta \beta_0) \\ &= \theta^{(\alpha_0 + \sum_{i=1}^n y_i) - 1} \exp[-\theta(\beta_0 + n)] \propto \text{Gamma}(\theta | \underbrace{\alpha_0 + \sum_{i=1}^n y_i}_{\alpha_n}, \underbrace{\beta_0 + n}_{\beta_n}). \end{aligned}$$

Poisson example - Bomb hits in London

$$n = 576, \sum_{i=1}^n y_i = 229 \cdot 0 + 211 \cdot 1 + 93 \cdot 2 + 35 \cdot 3 + 7 \cdot 4 + 1 \cdot 5 = 537.$$

Average number of hits per region $= \bar{y} = 537/576 \approx 0.9323$.

$$p(\theta|y) \propto \theta^{\alpha_0+537-1} \exp[-\theta(\beta_0 + 576)]$$

$$E(\theta|y) = \frac{\alpha_0 + \sum_{i=1}^n y_i}{\beta_0 + n} \approx \bar{y} \approx 0.9323,$$

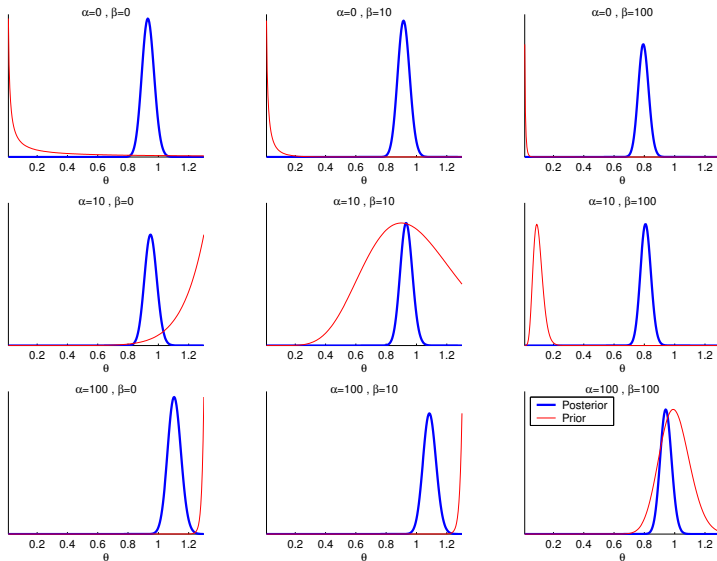
and

$$SD(\theta|y) = \left(\frac{\alpha_0 + \sum_{i=1}^n y_i}{(\beta_0 + n)^2} \right)^{1/2} = \frac{(\alpha_0 + \sum_{i=1}^n y_i)^{1/2}}{(\beta_0 + n)} \approx \frac{(537)^{1/2}}{576} \approx 0.0402.$$

if α and β **are small compared** to $\sum_{i=1}^n y_i$ and n .

Poisson bomb hits in London

Analysis of bomb hits in regions of London – Poisson model with Gamma prior



Poisson example - posterior intervals

- ▶ **Bayesian 95% interval**: the probability that the **unknown parameter** θ lies in the interval is 0.95. **What an easy and logical interpretation!**
- ▶ *Approximate* 95% **credible interval** for θ (for small α_0 and β_0):

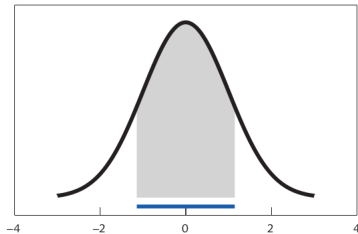
$$E(\theta|y) \pm 1.96 \cdot SD(\theta|y) = [0.8535; 1.0111]$$

Assumes that $p(\theta|y)$ is (approximately) normal.

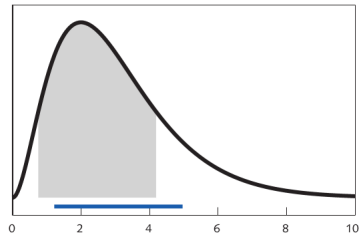
- ▶ An exact 95% **equal-tail interval** is $[0.8550; 1.0125]$ (assuming $\alpha_0 = \beta_0 = 0$)
- ▶ **Highest Posterior Density (HPD)** interval contains the θ values with highest pdf. Here $[0.8525; 1.0144]$, assuming $\alpha = \beta = 0$.

Illustration of different interval types

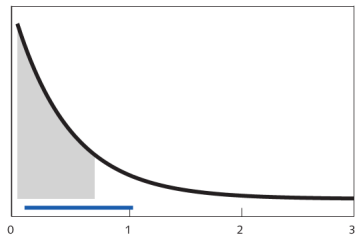
Symmetrical distribution



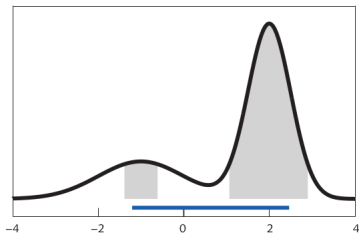
Skewed distribution



Skewed monotonous distribution



Bimodal distribution



Conjugate priors

- **Models** we have seen

Model	Prior	→	Posterior
Bernoulli	$\theta \sim \text{Beta}(\alpha_0, \beta_0)$	→	$\theta y \sim \text{Beta}(\alpha_n, \beta_n)$
Normal (σ^2 known)	$\theta \sim \mathcal{N}(\mu_0, \tau_0^2)$	→	$\theta y \sim \mathcal{N}(\mu_n, \tau_n^2)$
Poisson	$\theta \sim \text{Gamma}(\alpha_0, \beta_0)$	→	$\theta y \sim \text{Gamma}(\alpha_n, \beta_n)$

- **Conjugate priors:** A prior is conjugate to a model (likelihood) if the prior and posterior belong to the same distributional family.
- **Formally:** Let $\mathcal{F} = \{p(y|\theta), \theta \in \Theta\}$ be a class of sampling distributions. A family of distributions \mathcal{P} is conjugate for \mathcal{F} if

$$p(\theta) \in \mathcal{P} \Rightarrow p(\theta|y) \in \mathcal{P}$$

holds for all $p(y|\theta) \in \mathcal{F}$.

- A Conjugate prior is **computationally convenient**.

Prior elicitation

- ▶ The prior should (ideally) be elicited by an **expert** (\neq statistician, often)
- ▶ Elicit the prior on a **quantity that she knows well** (maybe log odds $\log \frac{\theta}{1-\theta}$ when the model is $\text{Bern}(\theta)$).
- ▶ The statistician can compute the **implied prior** on θ by transformation of variables.

Recall: Let $p_u(u)$ be continuous and let $v = h(u)$ be a one-to-one transform.

$$p_v(v) = p_u(h^{-1}(v))|J|, \quad |J| = \text{determinant of } h^{-1}(v) \left[1 - \dim : \frac{d}{dv} h^{-1}(v)\right].$$

- ▶ **Example:** expert believes $\phi = \log \frac{\theta}{1-\theta} \sim \mathcal{N}(0, 20)$. The implied prior on θ is $[u = \phi, v = \theta, h^{-1}(v) = \log \frac{v}{1-v}]$

$$p_\theta(\theta) = \mathcal{N}\left(\log \frac{\theta}{1-\theta} \middle| 0, 20\right) \frac{1}{\theta(1-\theta)}, \quad 0 < \theta < 1.$$

- ▶ The example works out a **full distribution**.

- ▶ Working out **hyper-parameters from expert information**.
- ▶ Elicit the prior by asking the expert simple questions: What is $E(\theta)$? or $V(\theta)$?
- ▶ The hyper-parameters are "backed out". **Example:** The prior is

$p(\theta) = \text{Gamma}(\theta|\alpha_0, \beta_0)$, expert believes $E(\theta) = 2$ and $V(\theta) = 0.25$.

$$E(\theta) = \frac{\alpha_0}{\beta_0}, \quad V(\theta) = \frac{\alpha_0}{\beta_0^2} \implies p(\theta) = \text{Gamma}(\theta|16, 8).$$

- ▶ **Show the expert some consequences** of her elicited prior.

Prior elicitation - AR(p) example

- ▶ **Autoregressive process** of order p

$$y_t = \mu + \phi_1 \cdot (y_{t-1} - \mu) + \dots + \phi_p \cdot (y_{t-p} - \mu) + \varepsilon_t, \quad \varepsilon_t \stackrel{iid}{\sim} N(0, \sigma^2)$$

- ▶ **Informative prior** on the unconditional mean: $\mu \sim N(\mu_0, \tau_0^2)$.
- ▶ **"Non-informative"** prior on σ^2 :

$$p(\sigma^2) \propto 1/\sigma^2 \quad [\text{uniform in the parameterization } p(\log(\sigma^2)) \propto c]$$

- ▶ **Assume** for simplicity that all $\phi_i, i = 1, \dots, p$ are independent a priori, and $\phi_i \sim N(\mu_i, \psi_i^2)$.
- ▶ Prior on $\phi = (\phi_1, \dots, \phi_p)$ centered on a persistent AR(1) process:

$$\mu_1 = 0.8, \mu_2 = \dots = \mu_p = 0.$$

- ▶ **Prior variance** ψ_i^2 of the ϕ_i decay towards zeros: $\text{Var}(\phi_i) = \frac{c}{i^\lambda}$, so that "longer" lags are **more concentrated around zero** (less likely a priori).
- ▶ λ is a parameter that can be used to determine the rate of decay.
Shrinkage/regularization/smoothness prior.

Different types of prior information

- ▶ Real **expert information**. Combo of previous studies and experience.
- ▶ Vague prior information, or even **non-informative priors**. **Beware of improper priors - make sure the posterior is proper!**
- ▶ **Smoothness priors**. Regularization. Shrinkage. Big thing in modern statistics/machine learning.
- ▶ **Hierarchical priors**. Model the uncertainty in the hyper-parameters. **Bayesian estimation of hyper-parameters.**

Non-informative priors

- ▶ **Do not exist!** The "flatness" depends on the parametrization of the model.
- ▶ Can be improper but still lead to a **proper posterior**.
- ▶ **Reference prior**: A prior that plays a "minimal role". "Let the data speak for themselves".
- ▶ Jeffreys' **invariance principle**: The prior should contain the same information **regardless of the parametrization** of the model.
- ▶ **Jeffreys'** prior (1-dim)

$$p(\theta) \propto |I(\theta)|^{1/2}, \quad I(\theta) = -E_y \left(\frac{d^2}{d\theta^2} \log p(y|\theta) \right),$$

where $I(\theta)$ is the **Fisher information** for θ .

- ▶ The expectation is **w.r.t data**... an **unconditional** (frequentist) feature!
- ▶ ... consequently, Jeffreys' prior **does not respect** the likelihood principle.
- ▶ Can give **dubious results** in multivariate (parameter) models.

Jeffreys' prior for Bernoulli trial data

Let $y = (y_1, \dots, y_n)$

$$y_1, \dots, y_n | \theta \stackrel{iid}{\sim} \text{Bern}(\theta) \quad \text{and} \quad \log p(y|\theta) = s \log \theta + f \log(1 - \theta).$$

$$\begin{aligned} \frac{d \log p(y|\theta)}{d\theta} &= \frac{s}{\theta} - \frac{f}{(1-\theta)} \\ \frac{d^2 \log p(y|\theta)}{d\theta^2} &= -\frac{s}{\theta^2} - \frac{f}{(1-\theta)^2} \\ I(\theta) &= \frac{E_y(s)}{\theta^2} + \frac{E_{y|\theta}(f)}{(1-\theta)^2} \\ &= \frac{n\theta}{\theta^2} + \frac{n(1-\theta)}{(1-\theta)^2} = \frac{n}{\theta(1-\theta)} \end{aligned}$$

Thus, **the Jeffreys' prior** is

$$p(\theta) = |I(\theta)|^{1/2} \propto \theta^{-1/2} (1-\theta)^{-1/2} \propto \text{Beta}(\theta|1/2, 1/2).$$

Non-informative priors - my two cents

- ▶ **OVERRATED.** Likelihood **dominates the prior** as more data becomes available.
- ▶ **State-of-the-art** models are **very complex** these days.
Regularization/shrinkage/smoothness priors to avoid over-fitting.
- ▶ Non-informative priors **do not shrink**.

Non-informative prior \implies no shrinkage \implies no fun.