

# TMA4315: Compulsory exercise 1 (title)

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24.09.2018

## Part 1

**Bold**

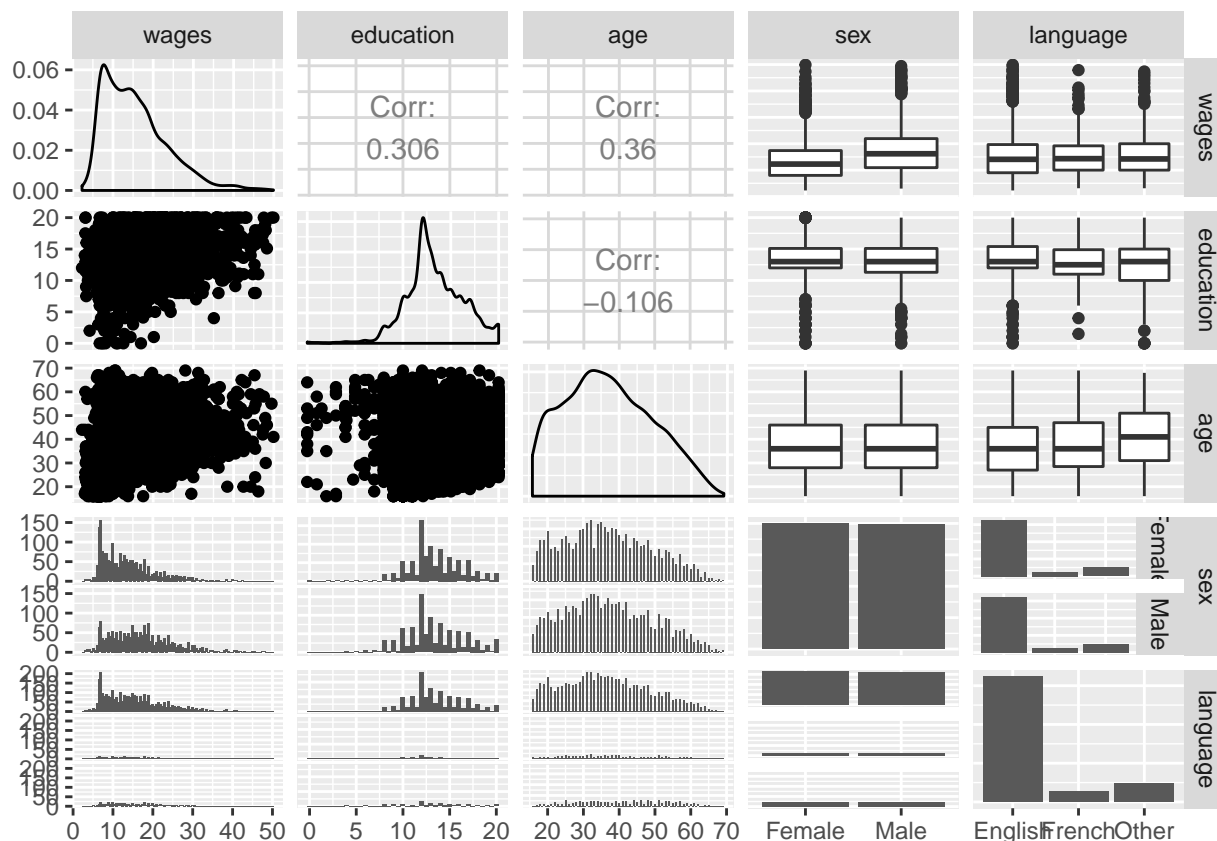
*italic*

To get a pdf file, make comments of the lines with the “html\_document” information, and make the lines with the “pdf\_document” information regular, and vice versa.

a)

Your answer for part 1a)

```
# some R code for part 1a)
library(GGally)
ggpairs(SLID, lower = list(combo = wrap(ggally_facethist, binwidth = 0.5)))
```



From the top row we can see that there is a noticeable correlation between wages, and the education, age and sex. Language on the other hand doesn't seem to have a large impact on the wages. One can see that people with high level (20) of education are distributed over the whole span of wages, but low educated people are centred around low wages, with very few or none at high level of wages.

The age correlates to wages in that there are people from all age categories that have low wages, but there is more likely to have a higher wage around age 40, and decreasing when younger or older.

There is a correlation between age and education, in that the education level decreases as the age increases. There is a known fact that the average education level has increased over the last 50 years, which correlates with the data set.

First we assume that there is a linear relationship between the covariates, i.e. the relationship can be expressed as:

$$\mathbf{Y} = \mathbf{X}\beta + \varepsilon$$

Linearity of covariates:  $\mathbf{Y} = \mathbf{X} + \cdot$ . Problem: non-linear relationship?

Homoscedastic error variance:  $\text{Cov}(\cdot) = 2\mathbf{I}$ . Problem: Non-constant variance of error terms

Uncorrelated errors:  $\text{Cov}(i, j) = 0$ .

Additivity of errors:  $\mathbf{Y} = \mathbf{X} +$

Assumption of normality:  $N(0, 2\mathbf{I})$

## Part 2

a)

```
# some R code for part 2a)
library(mylm)
modell1 <- mylm(wages ~ education, data = SLID)

## [1] 3987
print(modell1)

## Call:
## mylm(formula = wages ~ education, data = SLID)
##
## Coefficients:
## (Intercept) education
## [1,]      4.9717    0.79231

modell1b <- lm(wages ~ education, data = SLID)
print(modell1b)

##
## Call:
## lm(formula = wages ~ education, data = SLID)
##
## Coefficients:
## (Intercept)      education
##      4.9717         0.7923
```

b) Here is a print out of the covariance matrix defined as:

$$\Sigma = E[(X - E[X])(X - E[X])^T] = \frac{1}{n} \left( \sum_{i=1}^n (Y_i - \hat{Y}_i) \right) (X^T X)^{-1}$$

```
# some R code for part 2b
print.default(modell1$covariance.matrix)
```

```
##           (Intercept)      education
## (Intercept)  0.28532651 -0.020338410
## education   -0.02033841  0.001524956
```

*# some R code for part 2b*

```
summary(model1)
```

```
## Call:
## mylm(formula = wages ~ education, data = SLID)
##
## Residuals:
## Min: [1] -17.688
## 1Q: Median: [1] -1.039
## 3Q: Max: [1] 34.19
##
## Coefficients:
## Estimate           [,1]
## (Intercept)  4.97169
## education    0.79231
##
## Std. Error           [,1]
## (Intercept)  0.53429
## education    0.03906
##
## z-value           [,1]
## (Intercept)  9.305
## education   20.284
##
## Pr(>|z|)           [,1]
## (Intercept) 1.337833e-20
## education   1.774739e-91
##
## (Intercept) 1.352946e-20
## education   1.779032e-91
##
## R-squared:[1] 0.09358627
## Adjusted R-squared:[1] 0.09335881
```

```
summary(model1b)
```

```
##
## Call:
## lm(formula = wages ~ education, data = SLID)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -17.688  -5.822  -1.039   4.148  34.190
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  4.97169    0.53429   9.305  <2e-16 ***
## education    0.79231    0.03906  20.284  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
```

```
## Residual standard error: 7.492 on 3985 degrees of freedom
## Multiple R-squared:  0.09359,    Adjusted R-squared:  0.09336
## F-statistic: 411.4 on 1 and 3985 DF,  p-value: < 2.2e-16
```

The intercept estimate as shown in the print out is 4.97169 and the estimated standard error is 0.53429. For the regression coefficient the estimate is 0.79231 and the estimated standard error is 0.03906. Using a Z-test, we get:

$$P(Z \leq |z|) = 2 \cdot \Phi(-|Z|), \quad Z = \frac{x - \mu}{\sigma}$$

In our case, the  $H_0$  hypotheses is that  $\mu$  is zeros, and thus we get  $Z = x/\sigma$ :

```
cat("Z-values for the regression coefficients: ")
```

```
## Z-values for the regression coefficients:
```

```
print.default(model1$z_value)
```

```
##                [,1]
## (Intercept)  9.305166
## education   20.284158
```

Computing the p-values using

$$P(Z \leq z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-t^2} dt$$

The computation is implemented in our mylm package, and gives the values:

```
cat("P-values for the regression coefficients: ")
```

```
## P-values for the regression coefficients:
```

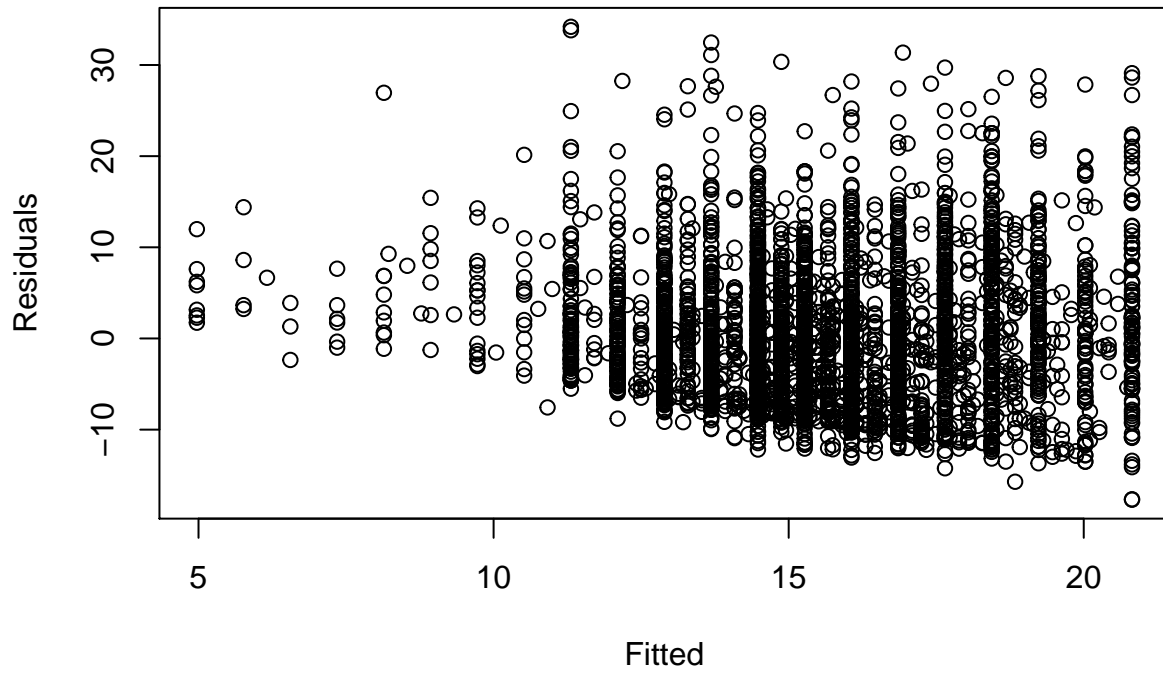
```
print.default(model1$p_value2)
```

```
##                [,1]
## (Intercept) 1.352946e-20
## education   1.779032e-91
```

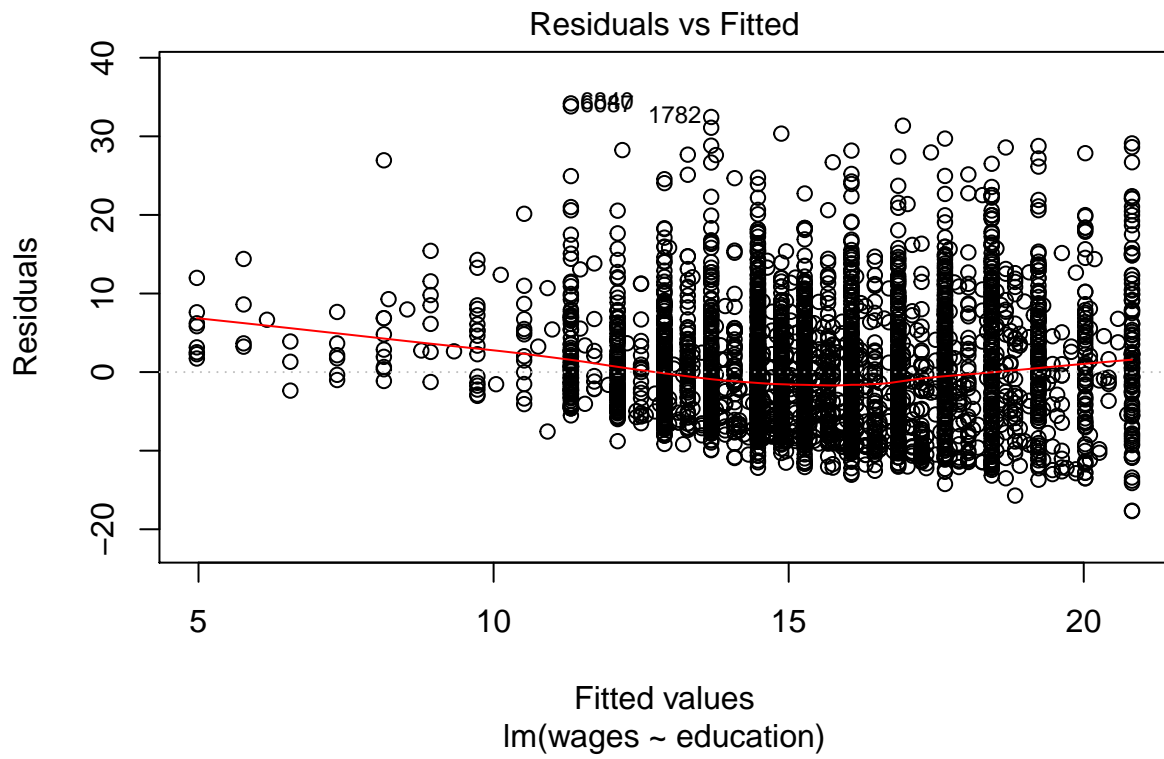
Which suggests that both the intercept and the regression coefficients are significant. The usual level to determine if a coefficient is significant or not is a 95%-confidence interval where  $P(Z \leq z) < 0.05$ . If true, the parameter in question is significant at a 5%-level, which both our parameters are in this case. **c)**

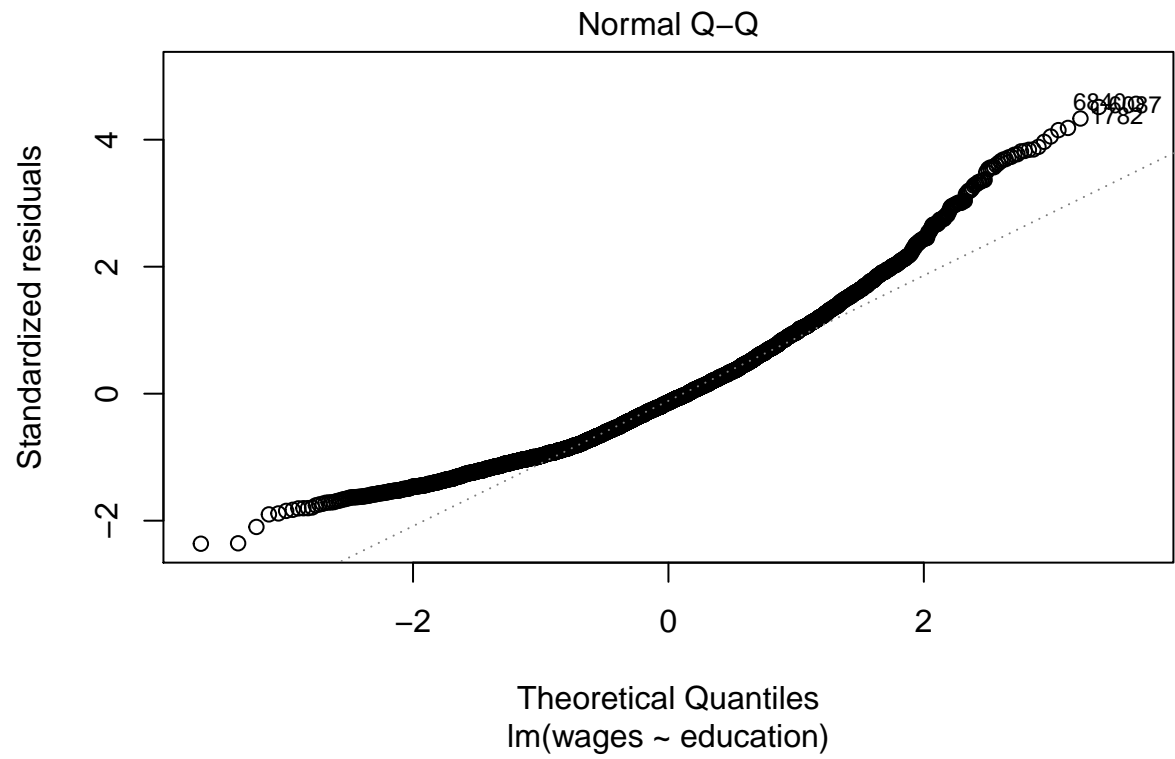
```
# some R code for part 2c)
library(ggplot2)
plot(model1)
```

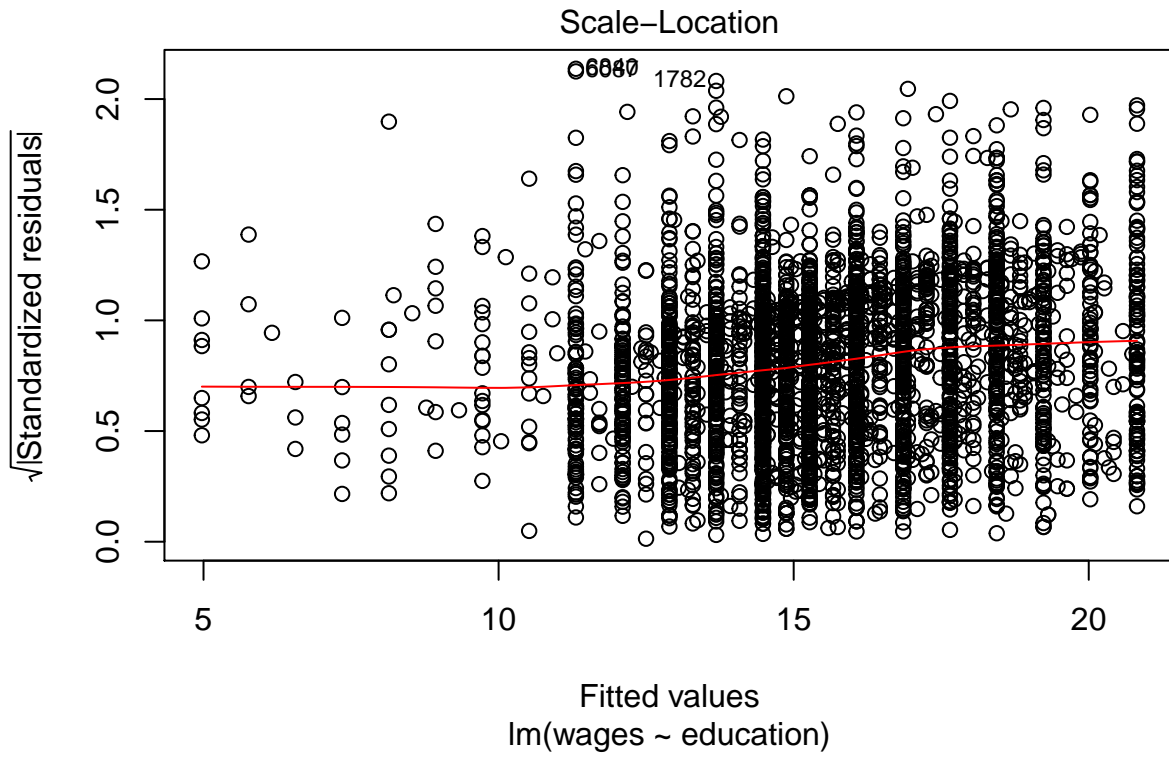
## Residual vs Fitted



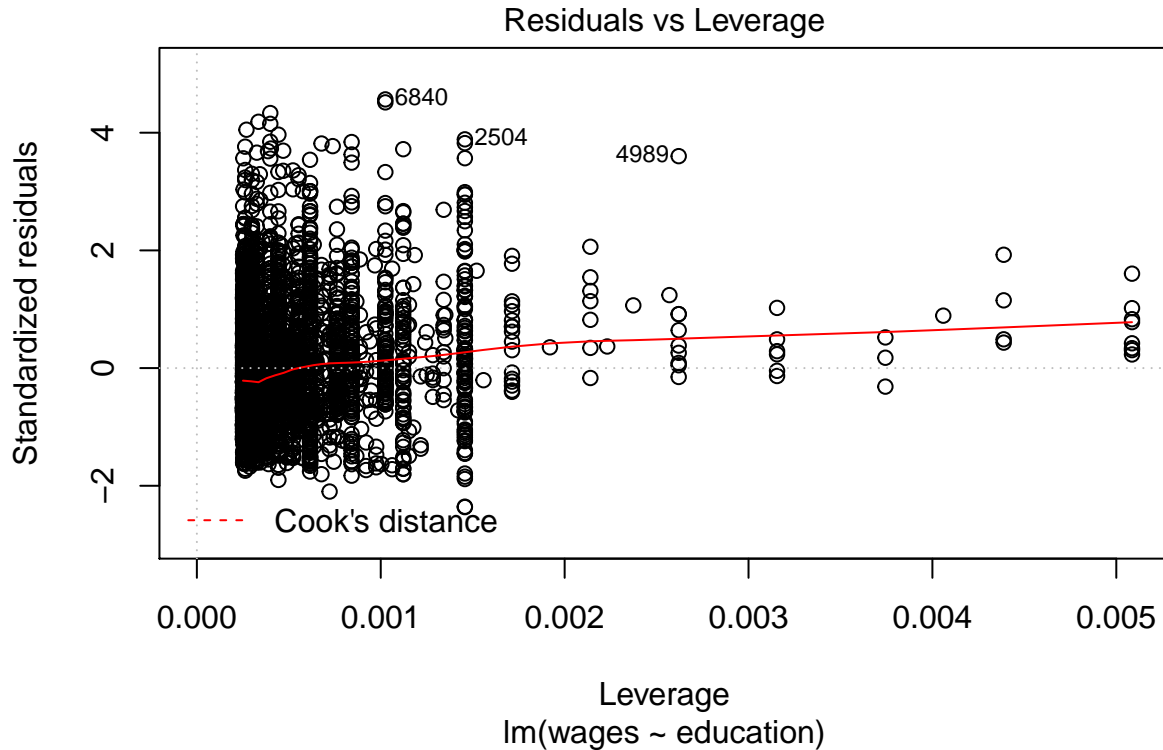
```
plot(model1b)
```











The plot shows the residuals plotted against the fitted values. A residual plot shows if the linear regression is appropriate for the data. A random distribution around the horizontal-axis suggest that there is no systematic error in the regression, and that a linear regression is appropriate. On the other side, if the residuals follow a systematic distribution around the horizontal-axis, there is likely that the relationship between the covariates and the response is non-linear. In this plot we interpret the plot as randomly distributed, and that the relationship can be described as a linear regression. **d)** After a scaling, the  $\chi^2$ -distribution is the limiting distribution of an F-distribution as the denominator degrees of freedom goes to infinity. The normalization is  $\chi^2 = (\text{numerator degrees of freedom}) \cdot F$ . • What is the residual sum of squares (SSE) and the degrees of freedom for this model? • What is total sum of squares (SST) for this model? Test the significance of the regression using a  $\chi^2$ -test. • What is the relationship between the  $\chi^2$  and z-statistic in simple linear regression? Find the critical value(s) for both tests.

- The residual sum of squares (SSE) for this model is computed as:  $SSE = \sum_{i=1}^n \epsilon_i^2$  where  $\epsilon_i = (I - H)Y$ . The degrees of freedom for this model is the number of dimensions that are free, which can be expressed by  $df = n - p - 1$ , where  $n$  is the number of datapoints (in this case 3987) and  $p$  is the number of explanatory parameters (in this case 1). Thus  $df = 3985$ .
- The total sum of squares (SST) is  $\sum_{i=1}^n (Y - \text{mean}(Y))^2$